



Tarea 3

Modelación ARIMA

Modelos de series de Tiempo y Supervivencia

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Analizar los datos Quarterly U.S. new plant/equip. expenditures 64 76 billions de la liberia tsdl de R.

1. Análisis descriptivo.

Grafique los datos, describa lo que observe (varianza constante o no constante, descomposición clásica, tendencia, ciclos estacionales, periodicidad de los ciclos).

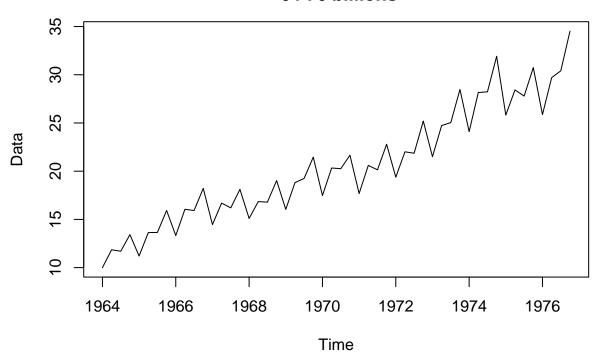
Primero carguemos la librería tsdl, ya que los datos necesarios se encuentran en dicha librería; así como otras necesarias para esta tarea.

```
library(tsdl); library(ggplot2); library(itsmr); library(forecast); library(TSA); library(lmtest)
library(timeSeries); library(timeSeries); library(astsa); library(dygraphs);
library(tseries); library(forecast); library(nortest); library(dplyr); library(imputeTS)
```

Cargamos la base de datos, nos aseguramos de que es la deseada.

```
Data <- tsdl[[12]]</pre>
attributes(Data)
## $tsp
## [1] 1964.00 1976.75
                           4.00
##
## $class
## [1] "ts"
##
## $source
## [1] "Abraham & Ledolter (1983)"
##
## $description
## [1] "Quarterly U.S. new plant/equip. expenditures -64 - -76 billions"
##
## $subject
## [1] "Microeconomic"
Ahora que noa aseguramos de que es la base que queriamos, procedemos a graficar:
plot(Data, main = "Quarterly U.S. new plant/equip. expenditures \n 64 76 billions")
```

Quarterly U.S. new plant/equip. expenditures 64 76 billions



Varianza Al menos de manera gráfica, la intuición nos dice que no hay varianza constante, pero probémoslo con un test de homocedasticidad:

```
#Los pasamos a series de tiempo
Serie<-ts(data=Data,start=c(1964,01),end=c(1976,4),frequency=4)
tiempo<-seq(1964+0/4, 1976+3/4, by = 1/4)
bptest(Serie~tiempo)

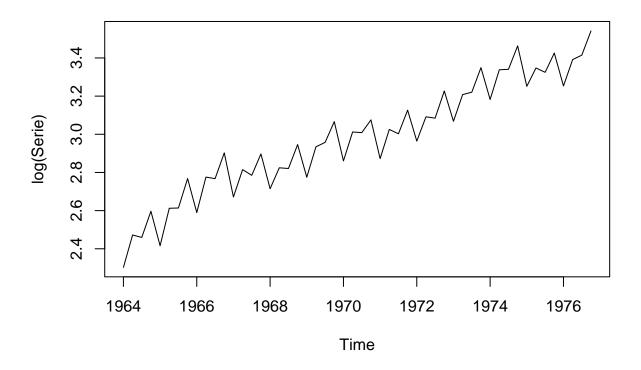
##
## studentized Breusch-Pagan test
##
## data: Serie ~ tiempo
## BP = 5.3713, df = 1, p-value = 0.02047</pre>
```

Efectivamente, no pasa el test de homocedasticidad.

Veamos qué pasa si aplicamos la transformación logaritmo:

```
bptest(log(Serie)~tiempo)
```

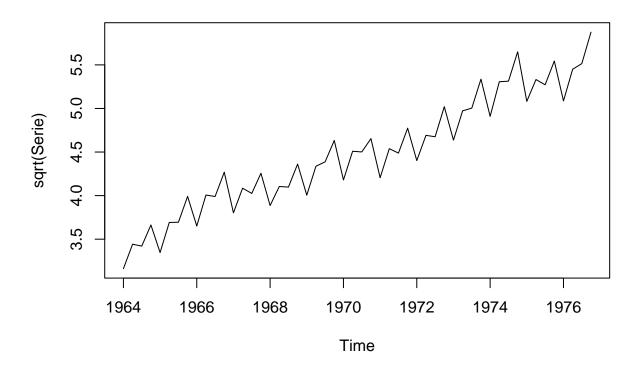
```
##
## studentized Breusch-Pagan test
##
## data: log(Serie) ~ tiempo
## BP = 1.965, df = 1, p-value = 0.161
```



Tampoco ayudó, aunque mejoró un poco. Intentemos con la raíz cuadrada

```
bptest(sqrt(Serie)~tiempo)
```

```
##
## studentized Breusch-Pagan test
##
## data: sqrt(Serie) ~ tiempo
## BP = 0.8651, df = 1, p-value = 0.3523
plot(sqrt(Serie))
```



¡Logramos estabilizarla!

Veamos si es estacionaria.

```
adf.test(sqrt(Serie))
##
##
    Augmented Dickey-Fuller Test
##
## data: sqrt(Serie)
## Dickey-Fuller = -1.415, Lag order = 3, p-value = 0.8097
## alternative hypothesis: stationary
kpss.test(sqrt(Serie))
## Warning in kpss.test(sqrt(Serie)): p-value smaller than printed p-value
##
    KPSS Test for Level Stationarity
##
##
## data: sqrt(Serie)
## KPSS Level = 1.389, Truncation lag parameter = 3, p-value = 0.01
```

Tendencia

Podemos observar una tendencia creciente que se presenta de manera lineal (al parecer), de manera general a lo largo de la serie

Ciclos estacionales

Los ciclos están bastante marcados , y tiene sentido puesto que son los datos de gastos de una empresa en maquinaria por trimestre, y en ese contexto es lógico que se presenten ciclos: Generalmente decae del último trimestre del año anterior al primero del año siguiente, para después crecer en el segundo semestre, en el tercero se mantiene casi al mismo nivel que el segundo, pero en el último aumenta; y es un comportamiento que se repite año con año

Periodicidad de los ciclos

Como comentamos en el apartado anterior, parece (al menos de manera gráfica) que se tienen ciclos anuales.

Descomposición clásica

Descomponeremos la serie por medio de filtros lineales:

Estabilización de la varianza

Aplicamos la transformación raíz cuadrada

```
Serie_sq<-sqrt(Serie)
bptest(Serie_sq~tiempo)</pre>
```

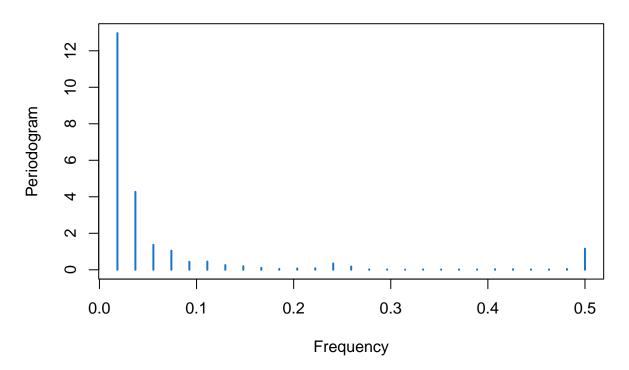
```
##
## studentized Breusch-Pagan test
##
## data: Serie_sq ~ tiempo
## BP = 0.8651, df = 1, p-value = 0.3523
```

Podemos asumir varianza constante

Periodicidad de ciclos

```
#Veamos la tendencia y los ciclos
Xt = Serie_sq
p = periodogram(Xt, main="Periodograma", col=4) # Obtenemos el periodograma
```

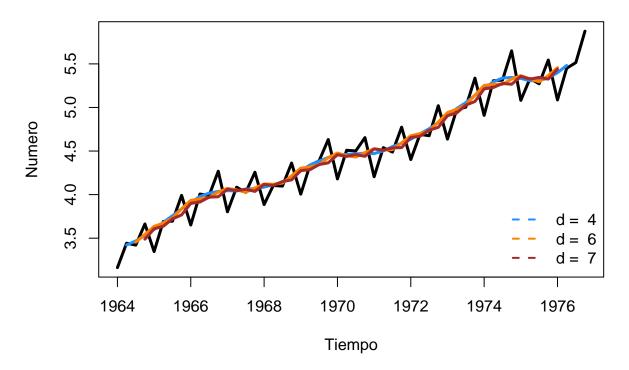
Periodograma



```
names(p)
                                                                    "df"
    [1] "freq"
                    "spec"
                                "coh"
                                            "phase"
                                                        "kernel"
   [7] "bandwidth" "n.used"
                                "orig.n"
                                            "series"
                                                        "snames"
                                                                    "method"
## [13] "taper"
                    "pad"
                                "detrend"
                                            "demean"
# Ordenamos de mayor a menor las estimaciones del periodograma.
spec = sort(p$spec, decreasing = TRUE)
(spec = spec[1:10]) # Nos quedamos con los 8 coeficientes de mayor frecuencia.
  [1] 12.9646281 4.2701894 1.3676068 1.1581310 1.0460373 0.4486585
   [7] 0.4395386 0.3466661 0.2593484 0.1970454
i = match(spec, p$spec) # Buscamos sus indices en el periodograma.
d = p$freq # Vemos las frecuencias del periodograma.
d = d[i] # Nos quedamos con las frecuencias que nos interesan.
cbind(spec,d,i)#
##
                               i
               spec
##
   [1,] 12.9646281 0.01851852
   [2,] 4.2701894 0.03703704
   [3,] 1.3676068 0.05555556
   [4,] 1.1581310 0.50000000 27
##
        1.0460373 0.07407407
##
   [5,]
   [6,] 0.4486585 0.11111111
##
  [7,] 0.4395386 0.09259259 5
##
  [8,] 0.3466661 0.24074074 13
```

```
## [9,] 0.2593484 0.12962963 7
## [10,] 0.1970454 0.14814815 8
d = 1 / d # Obtenemos los parametros para utilizar en promedios moviles.
d = floor(d) #
(d = sort(d))
## [1] 2 4 6 7 9 10 13 18 27 54
# Quitamos los periodos mas grandes
d = d[-length(d)]
d = d[-length(d)]
# Quitamos el más pequeño
d = d[-1]
d #Posibles periodos del ciclo
## [1] 4 6 7 9 10 13 18
#Realizamos la grafica:
col = c("dodgerblue1", "darkorange1", "brown")
plot(Serie_sq, lwd = 3, xlab = "Tiempo", col = "gray0",
     main = "Serie con varianza Homocedastica",
    ylab = "Numero", col.main = "burlywood")
for (i in 1:3) {
 lines(tiempo, stats::filter(Serie_sq, rep(1 / d[i], d[i])), col = col[i],
legend("bottomright", col = col, lty = 2, lwd = 2, bty = "n",
       legend = c(paste("d = ", d[1]), paste("d = ", d[2]),
                 paste("d = ", d[3])), cex = 1)
```

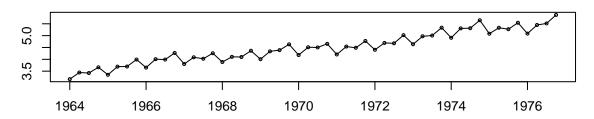
Serie con varianza Homocedastica

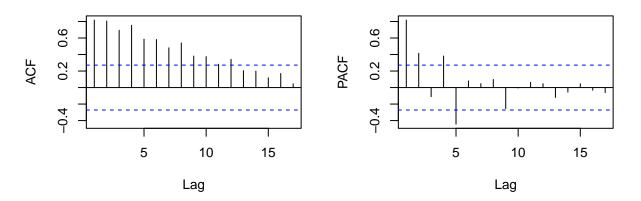


Notemos que d=2 parece sobreajustar un poco nuestra gráfica, de hecho, bastante. Sin embargo, con d=4 podemos obtener un buen suavizamiento sin pagar el costo de otros 2 datos al elegir d=6. Veamos el ACF y PACF:

tsdisplay(Serie_sq)

Serie_sq



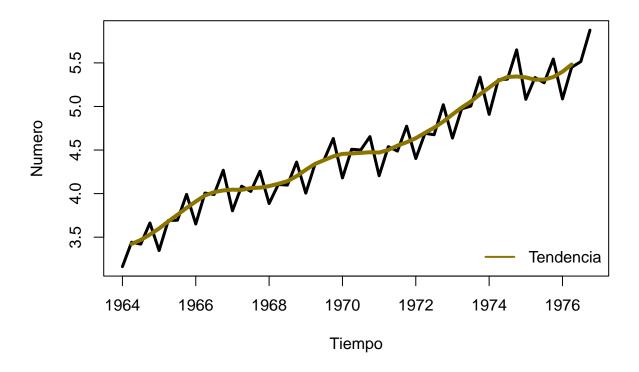


Y, junto con este último resultado, nos parece ideal concluir que el ciclo es d=4

Tendencia

Ahora, aislemos la tendencia:

Tendencia

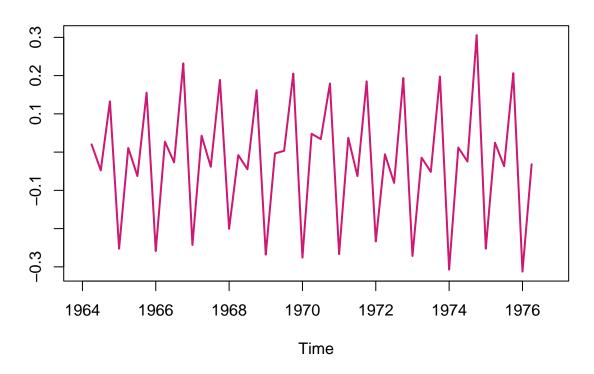


Lo que refuerza lo que creíamos: Tiene tebdebcua creciente casi de manera general.

```
# Quitamos la tendencia
# Solo trabajamos con la serie cuya varianza es cte.

datosSinTendencia = Serie_sq - tendencia # Serie sin tendencia
plot(datosSinTendencia, main="Serie sin tendencia", lwd=2, ylab="", col=14)
```

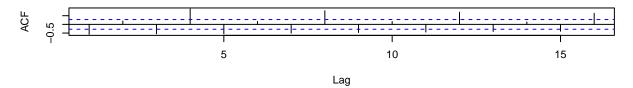
Serie sin tendencia



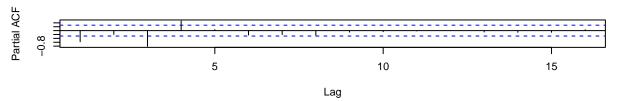
Serie de tiempo sin tendencia



Series datos.ts4[2:50]

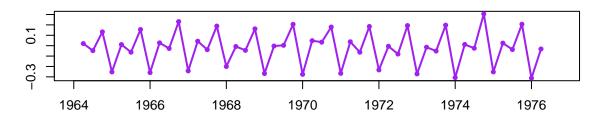


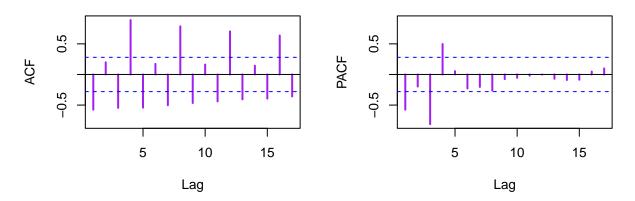
Series datos.ts4[2:50]



```
par(mfrow = c(1,1))
tsdisplay(datos.ts4, col="purple", lwd=2)
```

datos.ts4



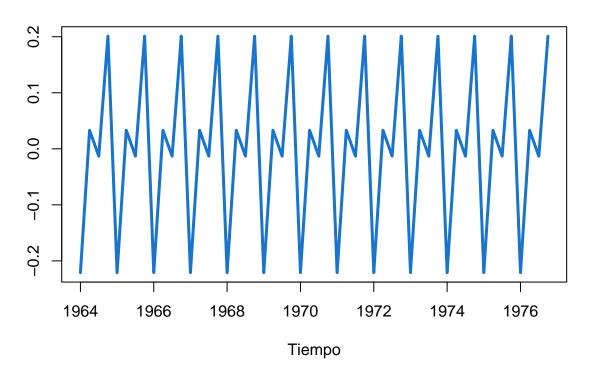


Parece que eliminamos la tendencia, ahora tratemos de verificar los ciclos

Ciclos o parte estacional

Ahora, estimaremos la parte estacional. Tenemos que d=4. Originalmente contábamos con 52 datos, pero ahora tenemos 48 (por los NA), entonces $\frac{48}{4} = 12$ ciclos.

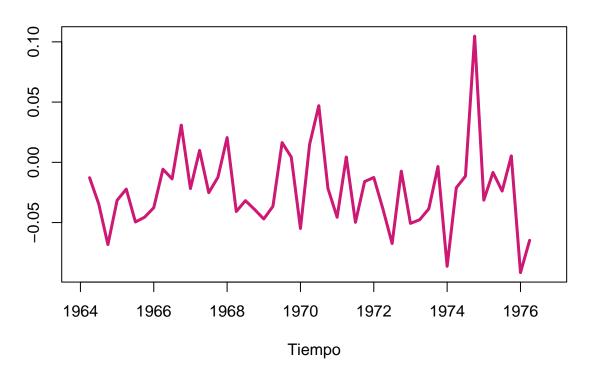
Ciclos de la serie



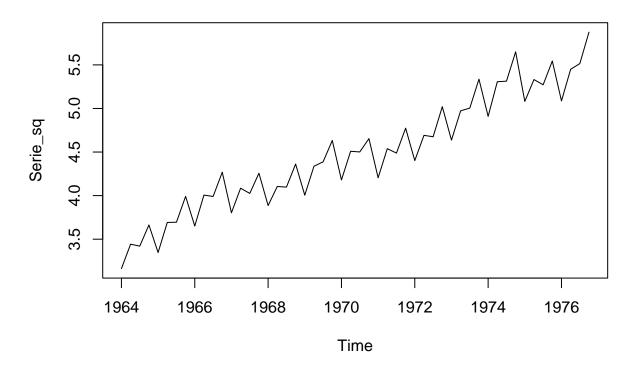
Ciclos anuales

Ahora verifiquemos de manera gráfica

Parte aleatoria



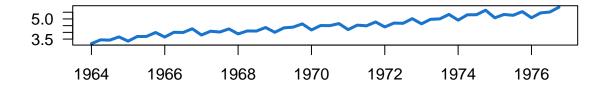
plot(Serie_sq)



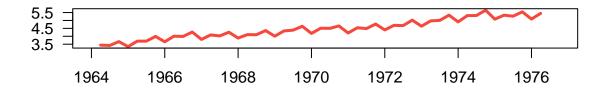
```
# Con esto, ya tenemos nuestras series

componentes = tendencia + ciclo+parte_aleatoria
componentes = ts(componentes, start = start(Serie_sq), frequency = 4)
par(mfrow = c(2,1))
plot(Serie_sq, col=28, las=1, main="Serie con varianza constante", lwd=3, xlab="",ylab="")
plot(componentes, col = 18, lwd = 3, las=1, main="Yt=tendencia+ciclos+aleatoria", xlab="",ylab="")
```

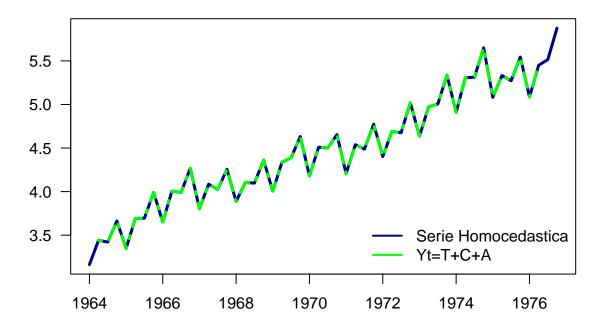
Serie con varianza constante



Yt=tendencia+ciclos+aleatoria



Serie_In



¡Logramos identificar los componentes de la serie!

2. Missing data

Suponga que las observaciones de 1971 Qtr1, 1973 Qtr2 y 1973 Qtr3, son datos faltantes NA, es decir, sustituya estas observaciones por NA.

Use al menos dos métodos de imputación de la paquetería imputeTS. ¿Cuál método es adecuado para estos datos? (Note que el valor imputado debe aproximarse al valor omitido).

Vamos a poner los datos que se piden, como NA:

```
Original=Serie[29]
Original[2]=Serie[38]
Original[3]=Serie[39]
Serie_2=Serie
Serie_2[29]=NA
Serie_2[38:39]=NA
Serie_2

## Qtr1 Qtr2 Qtr3 Qtr4
## 1964 10.00 11.85 11.70 13.42
## 1965 11.20 13.63 13.65 15.93
## 1966 13.33 16.05 15.92 18.22
## 1967 14.46 16.69 16.20 18.12
## 1968 15.10 16.85 16.79 19.03
## 1969 16.04 18.81 19.25 21.46
```

```
## 1970 17.47 20.33 20.26 21.66
           NA 20.60 20.14 22.79
## 1971
## 1972 19.38 22.01 21.86 25.20
## 1973 21.50
                 NA
                        NA 28.48
## 1974 24.10 28.16 28.23 31.92
## 1975 25.82 28.43 27.79 30.74
## 1976 25.87 29.70 30.41 34.52
Ahora, veamos método por método:
Un vistazo a la documentación de la paquetería mencionada nos menciona los siguientes métodos:
na interpolation:
Missing Value Imputation by Interpolation. Acepta 3 tipos:
"linear" - for linear interpolation using approx (default choice)
Comparaciones=data.frame(Original)
na_interpolation_lineal=na_interpolation(Serie_2, option='linear')[29]-Original[1]
na_interpolation_lineal[2:3]=na_interpolation(Serie_2,option='linear')[38:39]-Original[2:3]
Comparaciones['na_interpolation_lineal']=na_interpolation_lineal
"spline" - for spline interpolation using spline
na_interpolation(Serie_2,option='spline')
                      Qtr2
                               Qtr3
            Qtr1
                                         Otr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 21.91352 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 22.89979 27.71251 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_interpolation_spline=na_interpolation(Serie_2,option='spline')[29]-Original[1]
na_interpolation_spline[2:3]=na_interpolation(Serie_2,option='spline')[38:39]-Original[2:3]
Comparaciones['na_interpolation_spline']=na_interpolation_spline
"stine" - for Stineman interpolation using stinterp
na_interpolation(Serie_2,option='stine')
            Qtr1
                      Qtr2
                               Qtr3
                                         Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
```

```
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 21.13000 20.60000 20.14000 22.79000
```

```
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 23.10575 26.86069 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_interpolation_stine=na_interpolation(Serie_2,option='stine')[29]-Original[1]
na_interpolation_stine[2:3]=na_interpolation(Serie_2,option='stine')[38:39]-Original[2:3]
Comparaciones['na_interpolation_stine']=na_interpolation_stine
```

na kalman:

Missing Value Imputation by Kalman Smoothing and State Space Models. Acepta los siguientes modelos: \begin{enumerate} "auto.arima" - For using the state space representation of arima model (using auto.arima) (default choice)

```
na_kalman(Serie_2,model='auto.arima')
            Qtr1
                     Qtr2
                              Qtr3
                                       Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 17.88402 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.99381 25.24268 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_kalman_auto.arima=na_kalman(Serie_2, model = 'auto.arima')[29]-Original[1]
na_kalman_auto.arima[2:3]=na_kalman(Serie_2,model='auto.arima')[38:39]-Original[2:3]
Comparaciones['na_kalman_auto.arima']=na_kalman_auto.arima
```

"StructTS" - For using a structural model fitted by maximum likelihood (using StructTS)

```
na_kalman(Serie_2,model='StructTS')
```

```
Qtr1
                     Qtr2
                              Qtr3
                                       Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 17.98114 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.82428 24.92793 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_kalman_StructTS=na_kalman(Serie_2, model='StructTS')[29]-Original[1]
na_kalman_StructTS[2:3]=na_kalman(Serie_2,model='StructTS')[38:39]-Original[2:3]
```

```
Comparaciones['na_StructTS']=na_kalman_StructTS
na locf:
Missing Value Imputation by Last Observation Carried Forward. Acepta los siguientes métodos:
"locf" - for Last Observation Carried Forward (default choice)
na_locf(Serie_2, option = 'locf')
##
         Qtr1 Qtr2 Qtr3 Qtr4
## 1964 10.00 11.85 11.70 13.42
## 1965 11.20 13.63 13.65 15.93
## 1966 13.33 16.05 15.92 18.22
## 1967 14.46 16.69 16.20 18.12
## 1968 15.10 16.85 16.79 19.03
## 1969 16.04 18.81 19.25 21.46
## 1970 17.47 20.33 20.26 21.66
## 1971 21.66 20.60 20.14 22.79
## 1972 19.38 22.01 21.86 25.20
## 1973 21.50 21.50 21.50 28.48
## 1974 24.10 28.16 28.23 31.92
## 1975 25.82 28.43 27.79 30.74
## 1976 25.87 29.70 30.41 34.52
na_locf=na_locf(Serie_2,option='locf')[29]-Original[1]
na_locf[2:3]=na_locf(Serie_2,option='locf')[38:39]-Original[2:3]
Comparaciones['na_locf']=na_locf
"noch" - for Next Observation Carried Backward
na_locf(Serie_2, option = 'nocb')
         Qtr1 Qtr2 Qtr3 Qtr4
## 1964 10.00 11.85 11.70 13.42
## 1965 11.20 13.63 13.65 15.93
## 1966 13.33 16.05 15.92 18.22
## 1967 14.46 16.69 16.20 18.12
## 1968 15.10 16.85 16.79 19.03
## 1969 16.04 18.81 19.25 21.46
## 1970 17.47 20.33 20.26 21.66
## 1971 20.60 20.60 20.14 22.79
## 1972 19.38 22.01 21.86 25.20
## 1973 21.50 28.48 28.48 28.48
## 1974 24.10 28.16 28.23 31.92
## 1975 25.82 28.43 27.79 30.74
## 1976 25.87 29.70 30.41 34.52
na_locf_nocb=na_locf(Serie_2,option='nocb')[29]-Original[1]
na_locf_nocb[2:3]=na_locf(Serie_2,option='nocb')[38:39]-Original[2:3]
Comparaciones['na_locf_nocb']=na_locf_nocb
na ma Missing:
```

Value Imputation by Weighted Moving Average. Acepta los siguientes métodos:

"simple" - Simple Moving Average (SMA)

```
na_ma(Serie_2, weighting = 'simple')
            Qtr1
                     Qtr2
                              Qtr3
                                        Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 20.32875 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.47286 25.36143 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na ma simple=na ma(Serie 2, weighting='simple')[29]-Original[1]
na_ma_simple[2:3]=na_ma(Serie_2, weighting='simple')[38:39]-Original[2:3]
Comparaciones['na_ma_simple']=na_ma_simple
"linear" - Linear Weighted Moving Average (LWMA)
na_ma(Serie_2, weighting = 'linear')
##
                     Qtr2
            Qtr1
                              Qtr3
                                        Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 20.55065 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.27452 25.54742 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_ma_linear=na_interpolation(Serie_2,option='linear')[29]-Original[1]
na_ma_linear[2:3]=na_interpolation(Serie_2,option='linear')[38:39]-Original[2:3]
Comparaciones['na_ma_linear']=na_ma_linear
"exponential" - Exponential Weighted Moving Average (EWMA) (default choice)
na_ma(Serie_2, weighting = 'exponential')
            Qtr1
                     Qtr2
                              Qtr3
                                        Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 20.75900 20.60000 20.14000 22.79000
```

```
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.03682 25.77500 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_ma_exponential=na_ma(Serie_2, weighting='exponential')[29]-Original[1]
na_ma_exponential[2:3] = na_ma(Serie_2, weighting='exponential')[38:39] - Original[2:3]
Comparaciones['na exponential']=na ma exponential
na mean Missing:
Value Imputation by Mean Value. Acepta los siguientes métodos:
"mean" - take the mean for imputation (default choice)
na_mean(Serie_2, option = 'mean')
##
         Qtr1 Qtr2 Qtr3 Qtr4
## 1964 10.00 11.85 11.70 13.42
## 1965 11.20 13.63 13.65 15.93
## 1966 13.33 16.05 15.92 18.22
## 1967 14.46 16.69 16.20 18.12
## 1968 15.10 16.85 16.79 19.03
## 1969 16.04 18.81 19.25 21.46
## 1970 17.47 20.33 20.26 21.66
## 1971 20.43 20.60 20.14 22.79
## 1972 19.38 22.01 21.86 25.20
## 1973 21.50 20.43 20.43 28.48
## 1974 24.10 28.16 28.23 31.92
## 1975 25.82 28.43 27.79 30.74
## 1976 25.87 29.70 30.41 34.52
na mean=na mean(Serie 2, option='mean')[29]-Original[1]
na_mean[2:3]=na_mean(Serie_2,option='mean')[38:39]-Original[2:3]
Comparaciones['na mean']=na mean
"median" - take the median for imputation
na_mean(Serie_2,option = 'median')
         Qtr1 Qtr2 Qtr3 Qtr4
## 1964 10.00 11.85 11.70 13.42
## 1965 11.20 13.63 13.65 15.93
## 1966 13.33 16.05 15.92 18.22
## 1967 14.46 16.69 16.20 18.12
## 1968 15.10 16.85 16.79 19.03
## 1969 16.04 18.81 19.25 21.46
## 1970 17.47 20.33 20.26 21.66
## 1971 19.38 20.60 20.14 22.79
## 1972 19.38 22.01 21.86 25.20
## 1973 21.50 19.38 19.38 28.48
## 1974 24.10 28.16 28.23 31.92
## 1975 25.82 28.43 27.79 30.74
## 1976 25.87 29.70 30.41 34.52
na_mean_median=na_mean(Serie_2,option='median')[29]-Original[1]
```

na_mean_median[2:3]=na_mean(Serie_2,option='median')[38:39]-Original[2:3]

```
Comparaciones['na_mean_median']=na_interpolation_lineal
"mode" - take the mode for imputation
na_mean(Serie_2, option = 'mode')
         Qtr1 Qtr2 Qtr3 Qtr4
## 1964 10.00 11.85 11.70 13.42
## 1965 11.20 13.63 13.65 15.93
## 1966 13.33 16.05 15.92 18.22
## 1967 14.46 16.69 16.20 18.12
## 1968 15.10 16.85 16.79 19.03
## 1969 16.04 18.81 19.25 21.46
## 1970 17.47 20.33 20.26 21.66
## 1971 10.00 20.60 20.14 22.79
## 1972 19.38 22.01 21.86 25.20
## 1973 21.50 10.00 10.00 28.48
## 1974 24.10 28.16 28.23 31.92
## 1975 25.82 28.43 27.79 30.74
## 1976 25.87 29.70 30.41 34.52
na mean mode=na mean(Serie 2, option='mode')[29]-Original[1]
na_mean_mode[2:3]=na_mean(Serie_2,option='mode')[38:39]-Original[2:3]
Comparaciones['na_mean_mode']=na_mean_mode
"harmonic" - take the harmonic mean
na_mean(Serie_2, option = 'harmonic')
##
            Qtr1
                     Qtr2
                                        Qtr4
                              Qtr3
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 18.67181 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 18.67181 18.67181 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_mean_harmonic=na_mean(Serie_2,option='harmonic')[29]-Original[1]
na mean harmonic[2:3]=na mean(Serie 2,option='harmonic')[38:39]-Original[2:3]
Comparaciones['na_mean_harmonic']=na_mean_harmonic
"geometric" - take the geometric mean
na_mean(Serie_2, option = 'geometric')
##
            Qtr1
                     Qtr2
                              Qtr3
                                        Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
```

```
## 1968 15.10000 16.85000 16.79000 19.03000

## 1969 16.04000 18.81000 19.25000 21.46000

## 1970 17.47000 20.33000 20.26000 21.66000

## 1971 19.54094 20.60000 20.14000 22.79000

## 1972 19.38000 22.01000 21.86000 25.20000

## 1973 21.50000 19.54094 19.54094 28.48000

## 1974 24.10000 28.16000 28.23000 31.92000

## 1975 25.82000 28.43000 27.79000 30.74000

## 1976 25.87000 29.70000 30.41000 34.52000

na_mean_geometric=na_mean(Serie_2,option='geometric')[29]-Original[1]

na_mean_geometric[2:3]=na_mean(Serie_2,option='geometric')[38:39]-Original[2:3]

Comparaciones['na_mean_geometric']=na_mean_geometric
```

na_random:

Missing Value Imputation by Random Sample

```
na random(Serie 2)
```

```
##
                     Qtr2
                              Qtr3
            Qtr1
                                       Otr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 28.22092 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 16.31476 26.27779 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_random=na_random(Serie_2)[29]-Original[1]
na_random[2:3]=na_random(Serie_2)[38:39]-Original[2:3]
Comparaciones['na_random']=na_random
```

na_seadec:

Seasonally Decomposed Missing Value Imputation. Admite los siguiente métodos:

"interpolation" - Imputation by Interpolation (default choice)

```
na_seadec(Serie_2,algorithm = 'interpolation')
```

```
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 18.02499 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.90596 25.54880 28.48000
```

```
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_seadec_interpolation=na_seadec(Serie_2, algorithm='interpolation')[29]-Original[1]
na_seadec_interpolation[2:3]=na_seadec(Serie_2,algorithm='interpolation')[38:39]-Original[2:3]
Comparaciones['na_seadec_interpolation']=na_seadec_interpolation
"locf" - Imputation by Last Observation Carried Forward
na_seadec(Serie_2,algorithm = 'locf')
##
            Qtr1
                     Qtr2
                              Qtr3
                                       Otr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 17.78802 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 23.93409 23.60507 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_seadec_locf=na_seadec(Serie_2,algorithm='locf')[29]-Original[1]
na_seadec_locf[2:3]=na_seadec(Serie_2,algorithm='locf')[38:39]-Original[2:3]
Comparaciones['na_seadec_locf']=na_seadec_locf
"mean" - Imputation by Mean Value
na_seadec(Serie_2,algorithm = 'mean')
            Qtr1
                     Qtr2
                              Qtr3
                                       Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 18.33885 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 20.67408 20.34506 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_seadec_mean=na_seadec(Serie_2,algorithm = 'mean')[29]-Original[1]
na_seadec_mean[2:3]=na_seadec(Serie_2,algorithm = 'mean')[38:39]-Original[2:3]
Comparaciones['na_seadec_mean']=na_seadec_mean
"random" - Imputation by Random Sample
na_seadec(Serie_2,algorithm = 'random')
```

```
##
            Qtr1
                     Qtr2
                              Qtr3
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 15.94081 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.06344 23.36207 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_seadec_random=na_seadec(Serie_2, algorithm = 'random')[29]-Original[1]
na_seadec_random[2:3]=na_seadec(Serie_2, algorithm = 'random')[38:39]-Original[2:3]
Comparaciones['na_seadec_random']=na_seadec_random
"kalman" - Imputation by Kalman Smoothing and State Space Models
na_seadec(Serie_2,algorithm = 'kalman')
            Qtr1
                     Qtr2
                              Qtr3
                                       Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 17.86519 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.88983 24.98804 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_seadec_kalman=na_seadec(Serie_2,algorithm = 'kalman')[29]-Original[1]
na seadec kalman[2:3]=na seadec(Serie 2,algorithm = 'kalman')[38:39]-Original[2:3]
Comparaciones['na_seadec_kalman']=na_seadec_kalman
"ma" - Imputation by Weighted Moving Average
na_seadec(Serie_2,algorithm = 'ma')
##
            Qtr1
                     Qtr2
                              Qtr3
                                       Otr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 18.16021 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.58540 25.62559 28.48000
```

```
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_seadec_ma=na_seadec(Serie_2,algorithm = 'ma')[29]-Original[1]
na_seadec_ma[2:3]=na_seadec(Serie_2, algorithm = 'ma')[38:39]-Original[2:3]
Comparaciones['na_seadec_ma']=na_seadec_ma
Seasonally Splitted Missing Value Imputation. Admite los siguiente métodos:
"interpolation" - Imputation by Interpolation (default choice)
na_seasplit(Serie_2,algorithm = 'interpolation')
          Qtr1
                 Qtr2
                        Qtr3
                                Qtr4
## 1964 10.000 11.850 11.700 13.420
## 1965 11.200 13.630 13.650 15.930
## 1966 13.330 16.050 15.920 18.220
## 1967 14.460 16.690 16.200 18.120
## 1968 15.100 16.850 16.790 19.030
## 1969 16.040 18.810 19.250 21.460
## 1970 17.470 20.330 20.260 21.660
## 1971 18.425 20.600 20.140 22.790
## 1972 19.380 22.010 21.860 25.200
## 1973 21.500 25.085 25.045 28.480
## 1974 24.100 28.160 28.230 31.920
## 1975 25.820 28.430 27.790 30.740
## 1976 25.870 29.700 30.410 34.520
na_seasplit_interpolation=na_seasplit(Serie_2, algorithm = 'interpolation')[29]-Original[1]
na_seasplit_interpolation[2:3]=na_seasplit(Serie_2,algorithm = 'interpolation')[38:39]-Original[2:3]
Comparaciones['na seadec interpolation'] = na seasplit interpolation
"locf" - Imputation by Last Observation Carried Forward
na_seasplit(Serie_2,algorithm = 'locf')
##
         Qtr1 Qtr2 Qtr3 Qtr4
## 1964 10.00 11.85 11.70 13.42
## 1965 11.20 13.63 13.65 15.93
## 1966 13.33 16.05 15.92 18.22
## 1967 14.46 16.69 16.20 18.12
## 1968 15.10 16.85 16.79 19.03
## 1969 16.04 18.81 19.25 21.46
## 1970 17.47 20.33 20.26 21.66
## 1971 17.47 20.60 20.14 22.79
## 1972 19.38 22.01 21.86 25.20
## 1973 21.50 22.01 21.86 28.48
## 1974 24.10 28.16 28.23 31.92
## 1975 25.82 28.43 27.79 30.74
## 1976 25.87 29.70 30.41 34.52
na_seadec_locf=na_seasplit(Serie_2,algorithm = 'locf')[29]-Original[1]
na_seadec_locf[2:3]=na_seasplit(Serie_2,algorithm = 'locf')[38:39]-Original[2:3]
Comparaciones['na_seadec_locf']=na_seadec_locf
```

[&]quot;mean" - Imputation by Mean Value

```
na_seasplit(Serie_2,algorithm = 'mean')
            Qtr1
                     Qtr2
                                       Qtr4
                              Otr3
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 17.85583 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 20.25917 20.18333 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na seadec mean=na seasplit(Serie 2,algorithm = 'mean')[29]-Original[1]
na_seadec_mean[2:3]=na_seasplit(Serie_2, algorithm = 'mean')[38:39]-Original[2:3]
Comparaciones['na_seadec_mean']=na_seadec_mean
"random" - Imputation by Random Sample
na_seasplit(Serie_2,algorithm = 'random')
##
                     Qtr2
            Qtr1
                              Qtr3
                                       Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 14.33650 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 21.55354 26.53126 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_seadec_random=na_seasplit(Serie_2,algorithm = 'random')[29]-Original[1]
na_seadec_random[2:3]=na_seasplit(Serie_2,algorithm = 'random')[38:39]-Original[2:3]
Comparaciones['na_seadec_random']=na_seadec_random
"kalman" - Imputation by Kalman Smoothing and State Space Models
na_seasplit(Serie_2,algorithm = 'kalman')
            Qtr1
                     Qtr2
                              Qtr3
                                       Qtr4
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 18.37287 20.60000 20.14000 22.79000
```

```
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 25.07742 24.68497 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_seadec_kalman=na_seasplit(Serie_2,algorithm = 'kalman')[29]-Original[1]
na_seadec_kalman[2:3]=na_seasplit(Serie_2,algorithm = 'kalman')[38:39]-Original[2:3]
Comparaciones['na seadec kalman']=na seadec kalman
"ma" - Imputation by Weighted Moving Average
na seasplit(Serie 2,algorithm = 'ma')
##
            Qtr1
                      Qtr2
                               Qtr3
## 1964 10.00000 11.85000 11.70000 13.42000
## 1965 11.20000 13.63000 13.65000 15.93000
## 1966 13.33000 16.05000 15.92000 18.22000
## 1967 14.46000 16.69000 16.20000 18.12000
## 1968 15.10000 16.85000 16.79000 19.03000
## 1969 16.04000 18.81000 19.25000 21.46000
## 1970 17.47000 20.33000 20.26000 21.66000
## 1971 18.78800 20.60000 20.14000 22.79000
## 1972 19.38000 22.01000 21.86000 25.20000
## 1973 21.50000 24.70172 24.58724 28.48000
## 1974 24.10000 28.16000 28.23000 31.92000
## 1975 25.82000 28.43000 27.79000 30.74000
## 1976 25.87000 29.70000 30.41000 34.52000
na_seadec_ma=na_seasplit(Serie_2,algorithm = 'ma')[29]-Original[1]
na_seadec_ma[2:3]=na_seasplit(Serie_2,algorithm = 'ma')[38:39]-Original[2:3]
Comparaciones['na_seadec_ma']=na_seadec_ma
En el dataframe Comparaciones, lo que hacemos es guardar las diferencias respecto a la observación original,
acorde al método respectivo de ajuste.
Veamos cuál es el que minimiza dicho error:
Qr1 1971<-Comparaciones[1,]
Qr2_1973<-Comparaciones[2,]</pre>
Qr3_1973<-Comparaciones[3,]
#Para el primer trimiestre de 1971
dif_minima_Qr1_1971<-min(abs(Qr1_1971))
dif_minima_Qr1_1971
## [1] 0.1758333
metodo_minimiza_Qr1_1971<-which.min(abs(Qr1_1971))</pre>
metodo_minimiza_Qr1_1971
## na_seadec_mean
##
#Los 5 mejores
```

Warning in xtfrm.data.frame(x): cannot xtfrm data frames

aux_indices1<-(sort.list(abs(Qr1_1971)))[1:5]</pre>

```
for (indice in aux_indices1){
  print(Qr1_1971[indice])
  i=i+1
}
##
     na_seadec_mean
## 1
          0.1758333
##
    na_kalman_auto.arima
## 1
               0.2040217
##
   na_seadec_locf
## 1
          -0.21
##
   na_StructTS
## 1
        0.301137
##
   na_seadec_kalman
            0.6928704
## 1
#Para el segundo trimiestre de 1973
{\tt dif\_minima\_Qr2\_1973 \leftarrow min(abs(Qr2\_1973))}
dif_minima_Qr2_1973
## [1] 0.02827586
metodo_minimiza_Qr2_1973<-which.min(abs(Qr2_1973))
metodo_minimiza_Qr2_1973
## na_seadec_ma
##
#Los 5 mejores
aux\_indices2 < -(sort.list(abs(Qr2\_1973)))[1:5]
## Warning in xtfrm.data.frame(x): cannot xtfrm data frames
i=1
for (indice in aux_indices2){
  print(Qr2_1973[indice])
  i=i+1
}
##
   na_seadec_ma
## 2 -0.02827586
## na StructTS
## 2 0.09427935
## na ma simple
## 2 -0.2571429
   na_kalman_auto.arima
## 2
                0.2638126
##
   na_seadec_kalman
            0.3474205
## 2
#Para el tercer trimiestre de 1973
dif_minima_Qr3_1973<-min(abs(Qr3_1973))</pre>
dif_minima_Qr3_1973
## [1] 0.005
metodo_minimiza_Qr3_1973<-which.min(abs(Qr3_1973))</pre>
metodo_minimiza_Qr3_1973
```

```
## na_seadec_interpolation
##
#Los 5 mejores
aux_indices3<-(sort.list(abs(Qr3_1973)))[1:5]</pre>
## Warning in xtfrm.data.frame(x): cannot xtfrm data frames
i=1
for (indice in aux_indices3){
  print(Qr2_1973[indice])
  i=i+1
}
##
     na_seadec_interpolation
## 2
                        0.355
##
     na_StructTS
## 2 0.09427935
##
     na_kalman_auto.arima
## 2
                0.2638126
##
     na_ma_simple
## 2
       -0.2571429
##
     na_seadec_kalman
## 2
            0.3474205
```

3. Ajuste

Con los datos observados completos, ajuste un modelo ARIMA o SARIMA adecuado.

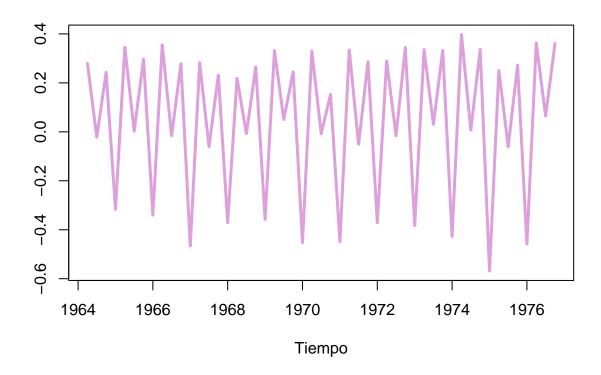
Obtenga correlogramas, revise si los parámetros son significativos, compruebe los supuestos (que los residuales sean ruido blanco con distribución normal). Obtenga dos o más posibles modelos, realice análisis de residuales y calcule medidas de bondad de ajuste. Haga la comparación para decidir cuál modelo sería el más adecuado.

Estacionariedad

En la primera parte vimos que si le aplicamos la transformación raíz cuadrada a la serie original, pasa la prueba para varianza constnte:

```
bptest(Serie_sq~tiempo)
##
##
    studentized Breusch-Pagan test
##
## data: Serie_sq ~ tiempo
## BP = 0.8651, df = 1, p-value = 0.3523
Ahora, las pruebas para estacionariedad:
adf.test(Serie_sq)
##
    Augmented Dickey-Fuller Test
##
##
## data: Serie_sq
## Dickey-Fuller = -1.415, Lag order = 3, p-value = 0.8097
## alternative hypothesis: stationary
```

Serie homoscedastica con una diferencia



```
adf.test(diff(Serie_sq))

##

## Augmented Dickey-Fuller Test

##

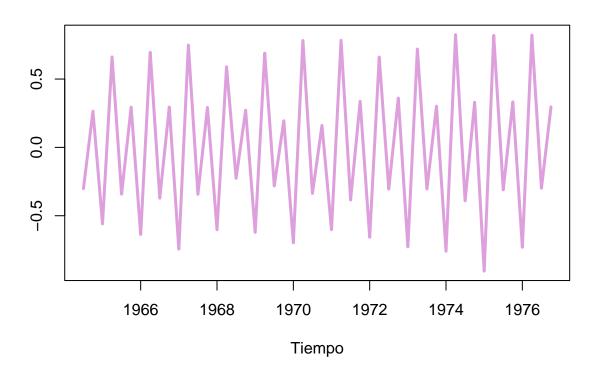
## data: diff(Serie_sq)

## Dickey-Fuller = -1.6911, Lag order = 3, p-value = 0.6986

## alternative hypothesis: stationary

kpss.test(diff(Serie_sq))
```

Serie homoscedastica con dos diferencias



```
adf.test(diff(diff(Serie_sq)))
## Warning in adf.test(diff(diff(Serie_sq))): p-value smaller than printed p-value
##
##
    Augmented Dickey-Fuller Test
##
## data: diff(diff(Serie_sq))
## Dickey-Fuller = -4.2222, Lag order = 3, p-value = 0.01
## alternative hypothesis: stationary
kpss.test(diff(diff(Serie_sq)))
##
##
    KPSS Test for Level Stationarity
##
## data: diff(diff(Serie_sq))
## KPSS Level = 0.44788, Truncation lag parameter = 3, p-value = 0.05652
```

Pasa las dos pruebas si aplicamos dos diferencias; trabajaremos con esa.

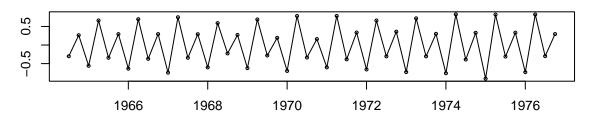
Correlogramas

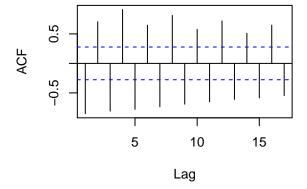
```
Serie_est<-diff(diff(Serie_sq))</pre>
```

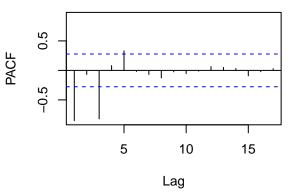
Visualizamos ACF y PACF

tsdisplay(Serie_est)

Serie_est







```
k=length(Serie_est)
banda<-qnorm(0.95)/(sqrt(k))
auxacf=acf(Serie_est,plot = F)#MA(6)
ACF_superior<-sum(auxacf$acf>banda)
ACF_inferior<-sum(auxacf$acf< -banda)
superan_banda_acf<-ACF_superior+ACF_inferior
superan_banda_acf</pre>
```

```
## [1] 16
```

```
pauxacf=pacf(Serie_est,plot = F)#AR(6)
PACF_superior<-sum(pauxacf$acf>banda)
PACF_inferior<-sum(pauxacf$acf< -banda)
superan_banda_pacf<-PACF_superior+PACF_inferior
superan_banda_pacf</pre>
```

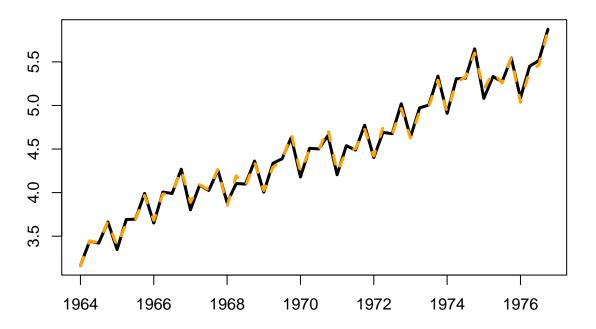
[1] 3

Parece, a simple vista, que puede que tengamos un ARIMA(3,2,16). Sin embargo, no hemos considerado la parte de los ciclos aún.

Veamos el ARIMA(3,2,16)...

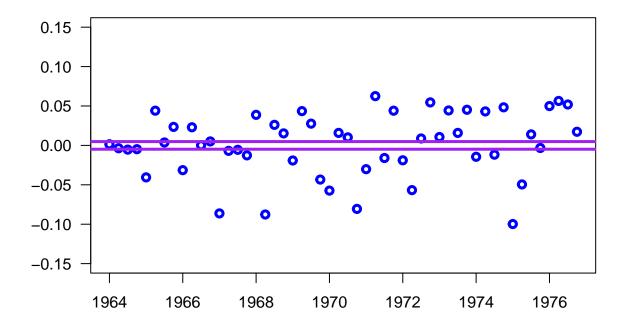
```
ARIMA<-arima(Serie_sq,order=c(3,2,16))
## Warning in stats::arima(x = x, order = order, seasonal = seasonal, xreg =
## xreg, : possible convergence problem: optim gave code = 1
ARIMA
##
## Call:
## arima(x = Serie_sq, order = c(3, 2, 16))
##
## Coefficients:
##
             ar1
                      ar2
                               ar3
                                       ma1
                                               ma2
                                                        ma3
                                                                 ma4
                                                                          ma5
                                                    0.1531
##
         -0.9944
                  -0.9932
                           -0.9987
                                    0.0916
                                            0.3508
                                                             -0.5243
                                                                      -0.3060
## s.e.
         0.0072
                   0.0074
                            0.0017
                                    0.1824
                                            0.2205
                                                    0.2262
                                                              0.2668
                                                                       0.2554
##
                                                 ma10
             ma6
                      ma7
                               ma8
                                        ma9
                                                         ma11
                                                                 ma12
                                                                         ma13
##
         -0.2524
                  -0.2833
                           -0.6654
                                    -0.1581
                                             -0.1254
                                                      0.0519
                                                              0.5361
                                                                       0.3603
          0.1789
                   0.2009
                            0.2805
                                     0.3226
                                              0.2876 0.2392 0.2432 0.2316
## s.e.
##
           ma14
                    ma15
                            ma16
         0.1052 -0.0599 0.0633
##
## s.e. 0.2092
                  0.2192 0.2458
##
## sigma^2 estimated as 0.001648: log likelihood = 75.1, aic = -112.19
Ahora veamos la gráfica
ARIMA ajuste <- (Serie sq - residuals(ARIMA))
ts.plot(Serie_sq, lwd=3, main="Comparación ", ylab="", xlab="")
points(ARIMA_ajuste, type = "1", col ="orange", lty = 2, lwd=3)
```

Comparación



```
###Y ahora los residuales
plot(ARIMA$residuals, type="p", col="blue", ylim=c(-0.15,0.15), ylab="", xlab="", main="Datos discrepant
abline(h=3*(var(ARIMA$residuals)), col="purple", lwd=3)
abline(h=-3*(var(ARIMA$residuals)), col="purple", lwd=3)
```

Datos discrepantes



¡Tenemos muchísimos datos discrepantes! Estamos sobreajustando los datos

Ahora atacaremos el problema con un SARIMA(p,d,q)x(P,D,Q), usando la estrategia definida en la página 180 de Introduction to Time Series and Forecasting de Peter Brockwell y Richard Davis.

Por el análisis hecho en la primera parte, sabemos que:

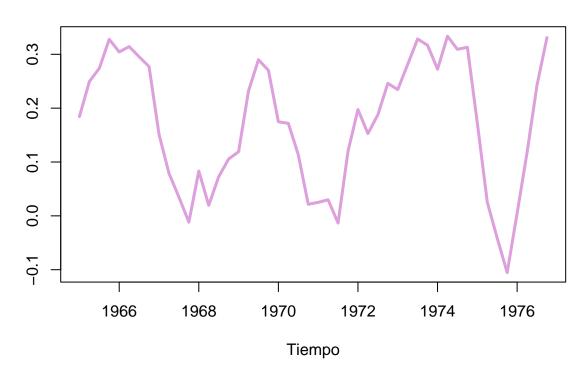
$$s = 4$$

Podemos estimar d y D como aquellos que hacen que

$$Y_t = (1 - B)^d (1 - B^s)^D$$

Sea estacionaria. Probemos con d=0 y D=1

Serie homoscedastica con una diferencia de lag=4



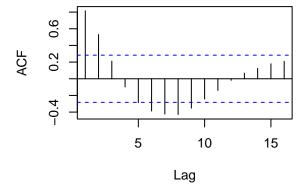
```
adf.test(diff(Serie_sq,lag=4))
## Warning in adf.test(diff(Serie_sq, lag = 4)): p-value smaller than printed p-
## value
##
   Augmented Dickey-Fuller Test
##
##
## data: diff(Serie_sq, lag = 4)
## Dickey-Fuller = -4.3985, Lag order = 3, p-value = 0.01
## alternative hypothesis: stationary
kpss.test(diff(Serie_sq,lag=4))
## Warning in kpss.test(diff(Serie_sq, lag = 4)): p-value greater than printed p-
## value
##
   KPSS Test for Level Stationarity
##
##
## data: diff(Serie_sq, lag = 4)
## KPSS Level = 0.076816, Truncation lag parameter = 3, p-value = 0.1
¡Las pasa! Entonces:
                                        d = 0, D = 1
```

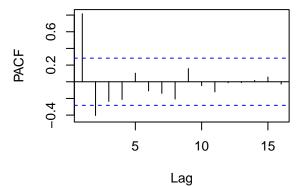
Trabajaremos con esta serie. Veamos su ACF y PACF

Serie_est<-diff(Serie_sq,lag=4)
tsdisplay(Serie_est)</pre>

Serie_est







¡Se ve mucho mejor que el anterior!

¿Cómo elegimos P y Q? La metodología seguida nos dice que veamos los lags que son múltiplos del ciclo (de s=4), y ver aquellos que se ajustan a un ARMA(P,Q). Es decir, en los lags ks $s=4,k\in\mathbb{N}\setminus\{0\}$

 $\c i$ Cómo elegimos p
 y q? Debemos fijarnos en aquellos lags entre 1 y s-1, es decir: Nos fijar
emos en los lags 1,2,3 y los ajustaremos a un ARMA(p,q)

Veamos el ACF

```
k=length(diff(Serie_sq,lag=4))
banda<-qnorm(0.95)/(sqrt(k))
auxacf=acf(diff(Serie_sq,lag=4),plot = F)#MA(6)
ACF_superior<-sum(auxacf$acf>banda)
ACF_inferior<-sum(auxacf$acf< -banda)
superan_banda_acf<-ACF_superior+ACF_inferior
superan_banda_acf</pre>
```

[1] 8

```
#Los que superan las bandas del acf son:
which(abs(auxacf$acf) > banda)
```

[1] 1 2 5 6 7 8 9 10

Por lo que:

$$q=2, Q=1$$

Para el PACF:

```
pauxacf=pacf(diff(Serie_sq,lag=4),plot = F)
PACF_superior<-sum(pauxacf$acf>banda)
PACF_inferior<-sum(pauxacf$acf< -banda)
superan_banda_pacf<-PACF_superior+PACF_inferior
superan_banda_pacf</pre>
```

[1] 2

```
#Los que superan las bandas del pacf son:
which(abs(pauxacf$acf) > banda)
```

[1] 1 2

Por lo que:

$$p = 2, P = 0$$

Entonces tenemos un modelo

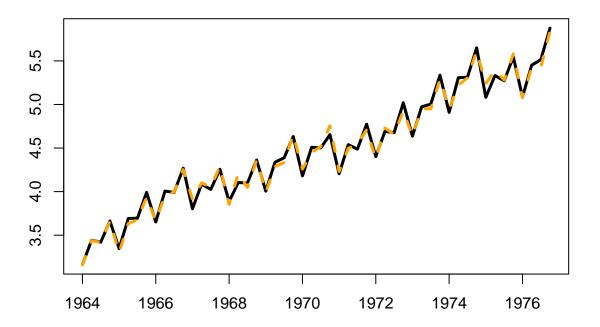
$$SARIMA(2,0,2) \times (0,1,1)_{[4]}$$

Ajustémoslo:

```
SARIMA<-arima(Serie_sq,order=c(2,0,2),seasonal=list(order=c(0,1,1),period=4), include.mean=F) SARIMA
```

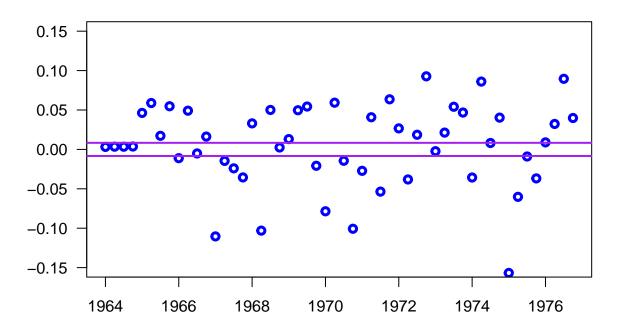
```
##
## Call:
## arima(x = Serie_sq, order = c(2, 0, 2), seasonal = list(order = c(0, 1, 1),
##
       period = 4), include.mean = F)
##
## Coefficients:
##
            ar1
                    ar2
                                            sma1
                           ma1
                                    ma2
##
        0.0779 0.9168 1.1411 0.2977 -0.8051
## s.e. 0.5505 0.5574 0.4604 0.2173 0.1948
##
## sigma^2 estimated as 0.002979: log likelihood = 68.54, aic = -127.08
SARIMA_ajuste <- Serie_sq - residuals(SARIMA)</pre>
ts.plot(Serie_sq, lwd=3, main="Comparación ", ylab="", xlab="")
points(SARIMA_ajuste, type = "1", col ="orange", lty = 2, lwd=3)
```

Comparación



```
###Y ahora los residuales
plot(SARIMA$residuals, type="p", col="blue", ylim=c(-.15,.15), ylab="", xlab="", main="Datos discrepant
abline(h=3*(var(SARIMA$residuals)), col="purple", lwd=2)
abline(h=-3*(var(SARIMA$residuals)), col="purple", lwd=2)
```

Datos discrepantes



Comparemos con el que ajusta R:

```
modelo_automatico<-auto.arima((diff(Serie_sq,lag=4)))</pre>
modelo_automatico
## Series: (diff(Serie_sq, lag = 4))
## ARIMA(2,0,1)(0,0,2)[4] with non-zero mean
##
## Coefficients:
##
                     ar2
                               ma1
                                       sma1
                                                sma2
                                                         mean
##
         1.6142
                -0.7911
                           -0.5142
                                    -0.3092
                                             -0.3808
                                                      0.1685
                                     0.2012
## s.e. 0.1300
                  0.1123
                           0.1760
                                              0.2298
## sigma^2 estimated as 0.002567: log likelihood=75.7
## AIC=-137.41
                 AICc=-134.61
                                BIC=-124.31
Probemos con otra diferencia:
adf.test(diff(diff(Serie_sq,lag=4)))
##
    Augmented Dickey-Fuller Test
##
## data: diff(diff(Serie_sq, lag = 4))
## Dickey-Fuller = -3.5973, Lag order = 3, p-value = 0.0434
## alternative hypothesis: stationary
```

```
kpss.test((diff(diff(Serie_sq,lag=4))))

## Warning in kpss.test((diff(diff(Serie_sq, lag = 4)))): p-value greater than
## printed p-value

##

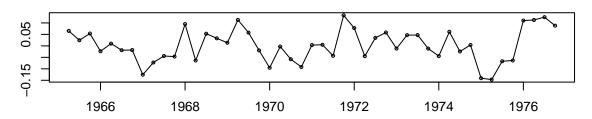
## KPSS Test for Level Stationarity
##

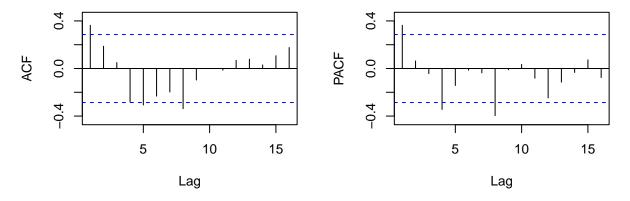
## data: (diff(diff(Serie_sq, lag = 4)))

## KPSS Level = 0.06211, Truncation lag parameter = 3, p-value = 0.1

tsdisplay(diff(diff(Serie_sq,lag=4)))
```

diff(diff(Serie_sq, lag = 4))





Sí pasa la prueba; ahora intentemos ajustar con dichas diferencias

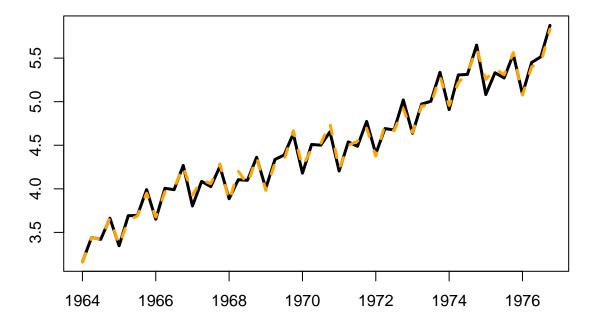
```
modelo_automatico<-auto.arima((diff(diff(Serie_sq,lag=4))))
modelo_automatico</pre>
```

```
## Series: (diff(diff(Serie_sq, lag = 4)))
## ARIMA(1,0,0)(0,0,1)[4] with zero mean
##
## Coefficients:
##
            ar1
                    sma1
                -0.6739
##
         0.2682
## s.e. 0.1468
                  0.1292
##
## sigma^2 estimated as 0.003262: log likelihood=67.64
## AIC=-129.27
                 AICc=-128.72
                               BIC=-123.72
```

```
Entonces es un SARIMA(1,1,0)x(0,1,1)[4]
```

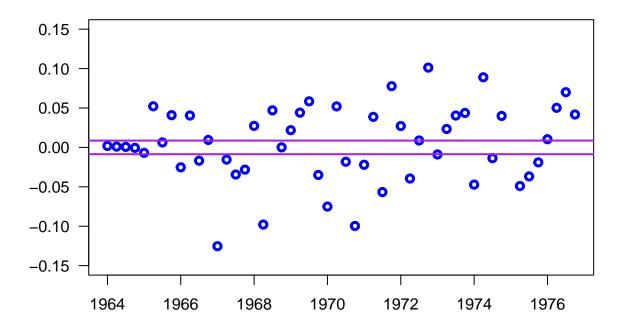
```
modelo_automatico<-arima(Serie_sq,order=c(1,1,0),seasonal=list(order=c(0,1,1),period=4))</pre>
modelo_automatico
##
## Call:
## arima(x = Serie_sq, order = c(1, 1, 0), seasonal = list(order = c(0, 1, 1),
       period = 4))
##
##
## Coefficients:
            ar1
                    sma1
##
         0.2682 -0.6739
## s.e. 0.1468
                  0.1292
##
## sigma^2 estimated as 0.003123: log likelihood = 67.64, aic = -131.27
AUTOMATICO_ajuste <- Serie_sq - residuals(modelo_automatico)</pre>
ts.plot(Serie sq, lwd=3, main="Comparación ", ylab="", xlab="")
points(AUTOMATICO_ajuste, type = "1", col ="orange", lty = 2, lwd=3)
```

Comparación



```
###Y ahora los residuales
plot(modelo_automatico$residuals, type="p", col="blue", ylim=c(-.15,.15), ylab="", xlab="", main="Datos
abline(h=3*(var(modelo_automatico$residuals)), col="purple", lwd=2)
abline(h=-3*(var(modelo_automatico$residuals)), col="purple", lwd=2)
```

Datos discrepantes



Tratemos de ajustar otro SARIMA manualmente; con el enfoque antes mencionado:

Veamos el ACF

```
k=length(diff(Serie_sq,lag=4)))
banda<-qnorm(0.95)/(sqrt(k))
auxacf=acf(diff(Serie_sq,lag=4)),plot = F)#MA(6)
ACF_superior<-sum(auxacf$acf>banda)
ACF_inferior<-sum(auxacf$acf< -banda)
superan_banda_acf<-ACF_superior+ACF_inferior
superan_banda_acf</pre>
```

[1] 4

#Los que superan las bandas del acf son:
which(abs(auxacf\$acf) > banda)

[1] 1 4 5 8

Por lo que:

$$q = 1, Q = 2$$

Para el PACF:

```
pauxacf=pacf(diff(diff(Serie_sq,lag=4)),plot = F)
PACF_superior<-sum(pauxacf$acf>banda)
PACF_inferior<-sum(pauxacf$acf< -banda)
superan_banda_pacf<-PACF_superior+PACF_inferior
superan_banda_pacf</pre>
```

```
## [1] 4
```

```
#Los que superan las bandas del pacf son:
which(abs(pauxacf$acf) > banda)
```

```
## [1] 1 4 8 12
```

Por lo que:

$$p=1, P=3$$

Entonces tenemos un modelo

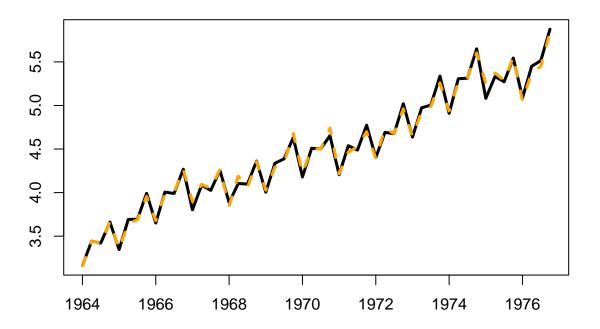
$$SARIMA(1,1,1) \times (3,1,2)_{[4]}$$

Ajustémoslo:

```
SARIMA2<-arima(Serie_sq,order=c(1,1,1),seasonal=list(order=c(3,1,2),period=4), include.mean=F)
SARIMA2
##
## Call:
## arima(x = Serie_sq, order = c(1, 1, 1), seasonal = list(order = c(3, 1, 2),
      period = 4), include.mean = F)
##
## Coefficients:
##
           ar1
                    ma1
                            sar1
                                     sar2
                                              sar3
                                                      sma1
                                                               sma2
        0.5047 -0.2597 -0.9560 -0.5185 -0.3367 0.5136 -0.4861
## s.e. 0.3125 0.3233
                         0.2893
                                  0.3190 0.2335 0.4041
                                                             0.2919
## sigma^2 estimated as 0.002429: log likelihood = 71.29, aic = -128.59
SARIMA_ajuste2 <- Serie_sq - residuals(SARIMA2)</pre>
```

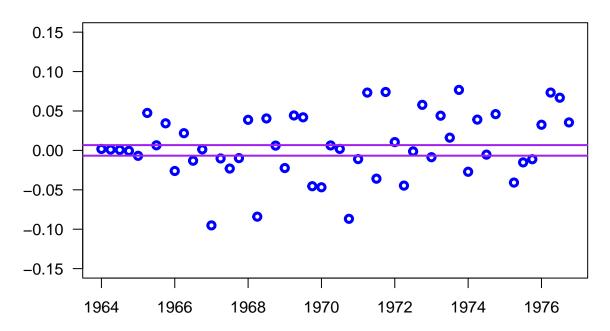
ts.plot(Serie_sq, lwd=3, main="Comparación ", ylab="", xlab="")
points(SARIMA_ajuste2, type = "l", col ="orange", lty = 2, lwd=3)

Comparación



```
###Y ahora los residuales
plot(SARIMA2$residuals, type="p", col="blue", ylim=c(-.15,.15), ylab="", xlab="", main="Datos discrepantabline(h=3*(var(SARIMA2$residuals)), col="purple", lwd=2)
abline(h=-3*(var(SARIMA2$residuals)), col="purple",lwd=2)
```

Datos discrepantes



Por el AIC y BIC

```
AIC <- c (SARIMA$aic, ARIMA$aic, modelo_automatico$aic, SARIMA2$aic)
BIC<-c(BIC(SARIMA),BIC(ARIMA),BIC(modelo_automatico), BIC(SARIMA2))
loglik<-c(SARIMA$loglik,ARIMA$loglik,modelo_automatico$loglik, SARIMA2$loglik)
Comparar<-data.frame('AIC'=AIC,'BIC'=BIC,'Loglik'=loglik,row.names = c('SARIMA','ARIMA','Automatico', '
Comparar
```

```
##
                    AIC
                               BIC
                                     Loglik
## SARIMA
              -127.0795 -113.85233 68.53977
## ARIMA
              -112.1905 -71.95005 75.09525
## Automatico -131.2738 -123.72339 67.63692
## SARIMA2
              -128.5861 -111.78489 71.29303
```

```
Comparemos los errores:
comparar_=cbind("ARIMA",ARIMA$aic,BIC(ARIMA), mean(ARIMA$residuals),
                mean(abs(ARIMA$residuals)),sqrt(mean((ARIMA$residuals)^2)),
                length(ARIMA$coef))
comparar_2=cbind("SARIMA",SARIMA$aic,BIC(SARIMA), mean(SARIMA$residuals),
                 mean(abs(SARIMA$residuals)),sqrt(mean((SARIMA$residuals)^2)),
                 length(SARIMA$coef))
comparar_3=cbind("SARIMA AUTOMÁTICO", modelo_automatico$aic,BIC(modelo_automatico), mean(modelo_automati
                 mean(abs(modelo_automatico$residuals)),sqrt(mean((modelo_automatico$residuals)^2)),
                 length(modelo_automatico$coef))
comparar_4=cbind("SARIMA2",SARIMA2$aic,BIC(SARIMA2), mean(SARIMA2$residuals),
```

```
##
  AJUSTE
                     AIC
                                       BIC
                                                         ME
##
  ARIMA
                     -112.190505919999 -71.950045811436 0.00113404901103327
## SARIMA
                     -127.079539350415 -113.852333284968 0.00480342160710416
## SARIMA AUTOMÁTICO -131.273837494611 -123.723394689481 0.000394350469091238
## SARIMA2
                     -128.586067490355 -111.784886676675 0.00191986834666404
## MAE
                      RMSE
                                         #Parametros
## 0.0313590810867887 0.0398103705554394 19
## 0.0408775073139969 0.0524452560585087 5
## 0.0406078930009833 0.0531376953726595 2
## 0.0342715205550597 0.0468668451395007 7
```

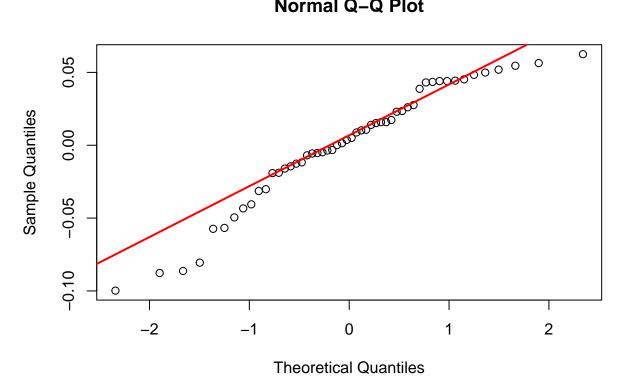
Parece que el SARIMA ajustado de manera automática es quien mejores indicadores tiene, después sería el segundo SARIMA que ajustamos nosotros, después el primer SARIMA y finalmente, el ARIMA

Hagamos la comprobación de supuestos:

Normalidad

ARIMA

```
#ARIMA
qqnorm(ARIMA$residuals)
qqline(ARIMA$residuals, col="red", lwd=2)
```

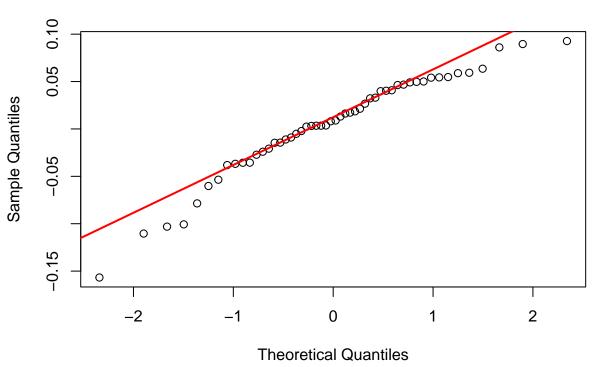


```
#Prueba Anderson-Darling
ad.test(ARIMA$residuals)
##
    Anderson-Darling normality test
##
##
## data: ARIMA$residuals
## A = 0.74525, p-value = 0.04904
#Prueba de Shapiro
shapiro.test(ARIMA$residuals)
##
    Shapiro-Wilk normality test
##
##
## data: ARIMA$residuals
## W = 0.94576, p-value = 0.01933
#Jarque-Bera Test
jarque.bera.test(ARIMA$residuals)
##
##
    Jarque Bera Test
##
## data: ARIMA$residuals
## X-squared = 3.9831, df = 2, p-value = 0.1365
Solo pasa Jarque-Bera
```

SARIMA

Jarque Bera Test

```
#SARIMA
qqnorm(SARIMA$residuals)
qqline(SARIMA$residuals, col="red", lwd=2)
```

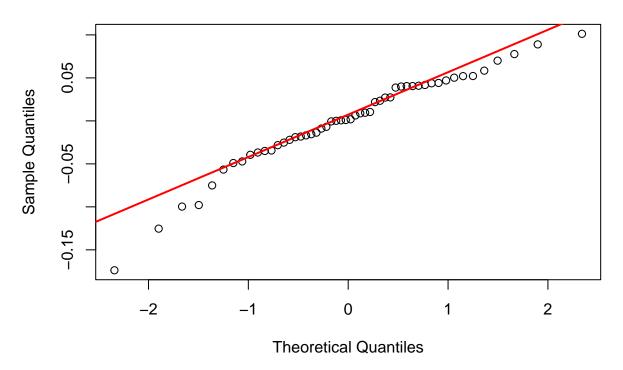


```
#Prueba Anderson-Darling
ad.test(SARIMA$residuals)
##
##
    Anderson-Darling normality test
##
## data: SARIMA$residuals
## A = 0.74048, p-value = 0.0504
#Prueba de Shapiro
shapiro.test(SARIMA$residuals)
##
##
    Shapiro-Wilk normality test
##
## data: SARIMA$residuals
## W = 0.94837, p-value = 0.02487
#Jarque-Bera Test. tseries
jarque.bera.test(SARIMA$residuals)
##
```

```
##
## data: SARIMA$residuals
## X-squared = 7.3243, df = 2, p-value = 0.02568
Solo pasa Anderson-Darling
```

Apenas pasamos Anderson-Darling ### SARIMA AUTOMÁTICO

```
#SARIMA AUTOMÁTICO
qqnorm(modelo_automatico$residuals)
qqline(modelo_automatico$residuals, col="red", lwd=2)
```



```
#Prueba Anderson-Darling
ad.test(modelo_automatico$residuals)

##
## Anderson-Darling normality test
##
## data: modelo_automatico$residuals
## A = 0.61243, p-value = 0.1055

#Prueba de Shapiro
shapiro.test(modelo_automatico$residuals)

##
## Shapiro-Wilk normality test
##
## data: modelo_automatico$residuals
##
## action = 0.95524, p-value = 0.04873
```

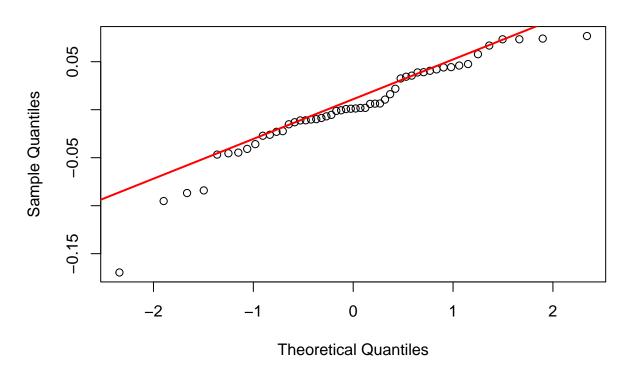
```
#Jarque-Bera Test. tseries
jarque.bera.test(modelo_automatico$residuals)

##
## Jarque Bera Test
##
## data: modelo_automatico$residuals
## X-squared = 8.8165, df = 2, p-value = 0.01218
```

Pasa Anderson-Darling con un p-value no tan cerca de 00.05, pero las demás pruebas las rechaza

SARIMA2

```
#SARIMA2
qqnorm(SARIMA2$residuals)
qqline(SARIMA2$residuals, col="red", lwd=2)
```



```
#Prueba Anderson-Darling
ad.test(SARIMA2$residuals)
```

```
##
## Anderson-Darling normality test
##
## data: SARIMA2$residuals
## A = 0.76981, p-value = 0.04256
```

```
#Prueba de Shapiro
shapiro.test(SARIMA2$residuals)
##
##
   Shapiro-Wilk normality test
##
## data: SARIMA2$residuals
## W = 0.93269, p-value = 0.005712
#Jarque-Bera Test. tseries
jarque.bera.test(SARIMA2$residuals)
##
##
    Jarque Bera Test
##
## data: SARIMA2$residuals
## X-squared = 17.165, df = 2, p-value = 0.0001874
No pasa ningún test.
El modelo ajustado por auto.arima tiene el mejor ajuste acorde a las pruebas
Varianza constante
ARIMA
#ARIMA
Y <- as.numeric(ARIMA$residuals)</pre>
X <- 1:length(ARIMA$residuals)</pre>
bptest(Y ~ X)
## studentized Breusch-Pagan test
##
## data: Y ~ X
## BP = 1.2041, df = 1, p-value = 0.2725
Lo pasa
SARIMA
#SARIMA
Y <- as.numeric(SARIMA$residuals)</pre>
X <- 1:length(SARIMA$residuals)</pre>
bptest(Y ~ X)
##
## studentized Breusch-Pagan test
## data: Y ~ X
## BP = 1.6613, df = 1, p-value = 0.1974
```

SARIMA AUTOMÁTICO

Lo pasa

```
#SARIMA AUTOMÁTICO
Y <- as.numeric(modelo_automatico$residuals)</pre>
X <- 1:length(modelo_automatico$residuals)</pre>
bptest(Y ~ X)
##
##
    studentized Breusch-Pagan test
##
## data: Y ~ X
## BP = 1.8803, df = 1, p-value = 0.1703
Lo pasa
SARIMA2
#SARIMA2
Y <- as.numeric(SARIMA2$residuals)
X <- 1:length(SARIMA2$residuals)</pre>
bptest(Y ~ X)
##
  studentized Breusch-Pagan test
##
##
## data: Y ~ X
## BP = 2.3289, df = 1, p-value = 0.127
Lo pasa
Todos pasan las pruebas; el mayor p-value lo obtuvo el ARIMA
Media cero
ARIMA
t.test(ARIMA$residuals,mu=0)
##
##
    One Sample t-test
##
## data: ARIMA$residuals
## t = 0.20352, df = 51, p-value = 0.8395
\#\# alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
## -0.01005282 0.01232092
## sample estimates:
## mean of x
## 0.001134049
Lo pasa
SARIMA
t.test(SARIMA$residuals,mu=0)
##
```

One Sample t-test

```
##
## data: SARIMA$residuals
## t = 0.65684, df = 51, p-value = 0.5142
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
## -0.009877914 0.019484757
## sample estimates:
## mean of x
## 0.004803422
Lo pasa
```

SARIMA AUTOMÁTICO

```
t.test(modelo_automatico$residuals,mu=0)

##

## One Sample t-test

##

## data: modelo_automatico$residuals

## t = 0.053, df = 51, p-value = 0.9579

## alternative hypothesis: true mean is not equal to 0

## 95 percent confidence interval:

## -0.0145432 0.0153319

## sample estimates:

## mean of x

## 0.0003943505
Lo pasa
```

SARIMA2

```
t.test(SARIMA2$residuals,mu=0)
```

```
##
## One Sample t-test
##
## data: SARIMA2$residuals
## t = 0.29279, df = 51, p-value = 0.7709
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
## -0.01124418    0.01508392
## sample estimates:
## mean of x
## 0.001919868
Lo pasa
```

Todos pasan esta prueba

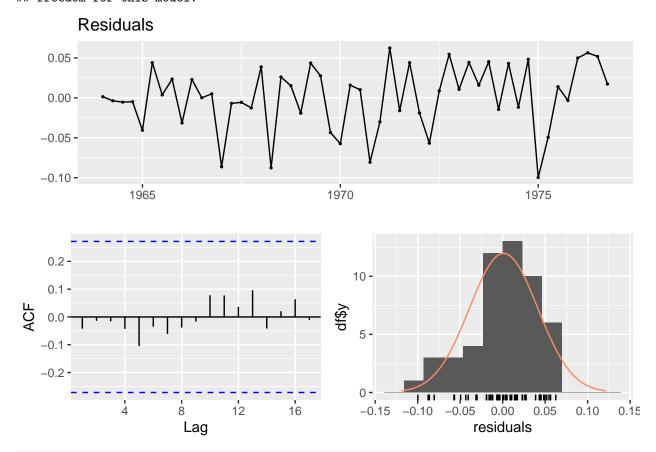
Residuales no correlacionados

ARIMA

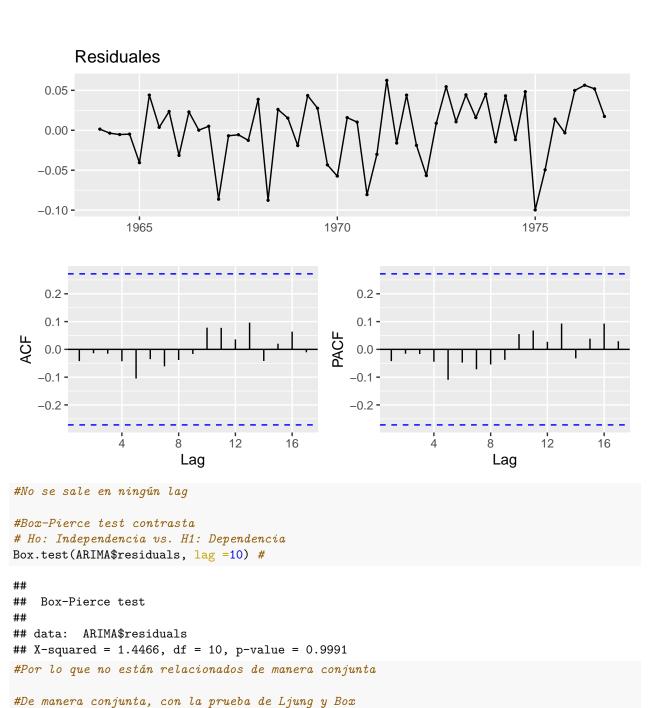
```
checkresiduals(ARIMA$residuals)
```

```
## Warning in modeldf.default(object): Could not find appropriate degrees of
```

freedom for this model.



ggtsdisplay(ARIMA\$residuals,main="Residuales")



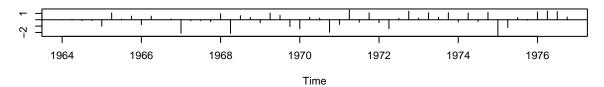
#HO:No estan correlacionados de manera conjunta

#H1:Estan correlacionados de manera conjunta

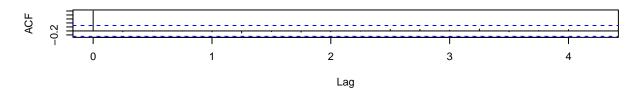
vs

tsdiag(ARIMA)

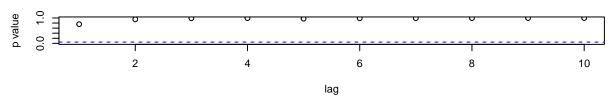
Standardized Residuals



ACF of Residuals



p values for Ljung-Box statistic

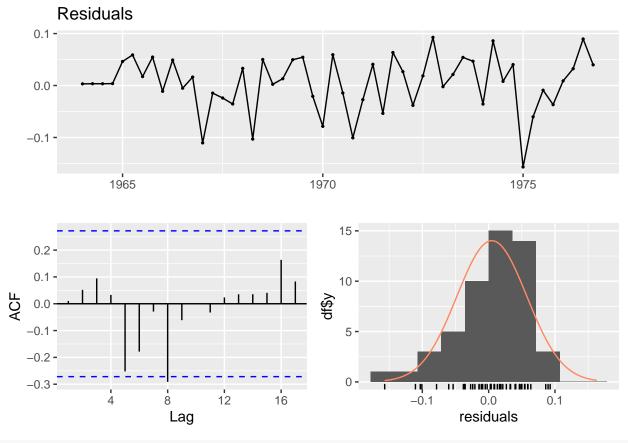


Pasa las pruebas

SARIMA

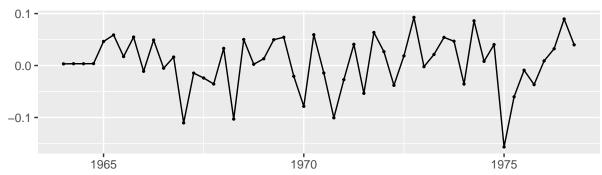
checkresiduals(SARIMA\$residuals)

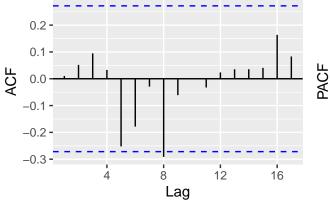
Warning in modeldf.default(object): Could not find appropriate degrees of
freedom for this model.

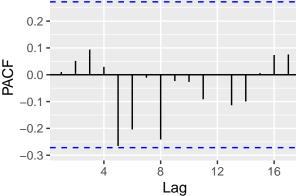


ggtsdisplay(SARIMA\$residuals,main="Residuales")









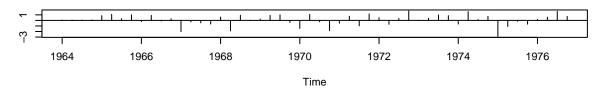
```
#Se sale en un lag cerca de 8, por poco.
#Box-Pierce test contrasta
# Ho: Independencia vs. H1: Dependencia
Box.test(SARIMA$residuals, lag =10)
```

tsdiag(SARIMA)

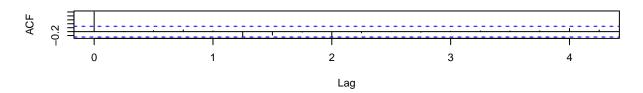
```
##
## Box-Pierce test
##
## data: SARIMA$residuals
## X-squared = 10.287, df = 10, p-value = 0.4157

#De manera conjunta, con la prueba de Ljung y Box
#HO:No estan correlacionados de manera conjunta
# vs
#H1:Estan correlacionados de manera conjunta
```

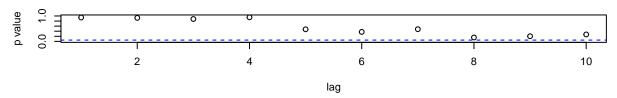
Standardized Residuals



ACF of Residuals



p values for Ljung-Box statistic

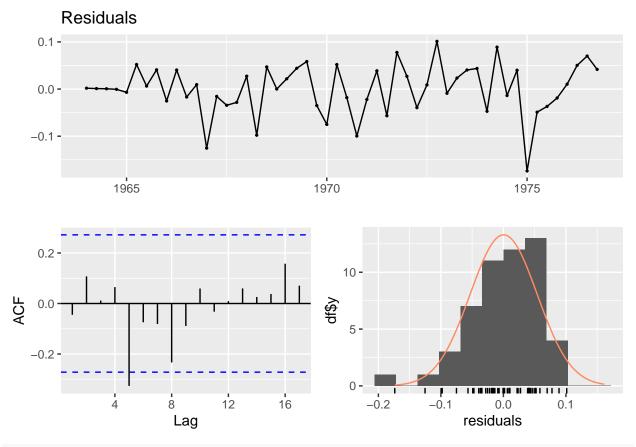


Pasa las pruebas

SARIMA AUTOMÁTICO

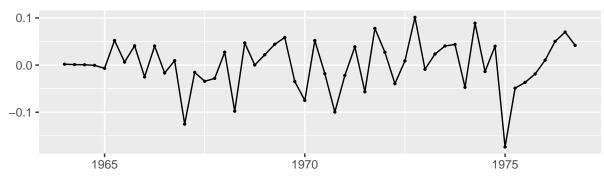
checkresiduals(modelo_automatico\$residuals)

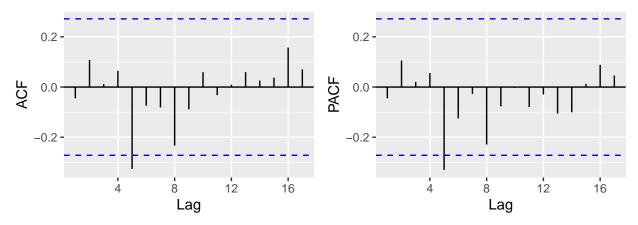
Warning in modeldf.default(object): Could not find appropriate degrees of ## freedom for this model.



ggtsdisplay(modelo_automatico\$residuals,main="Residuales")





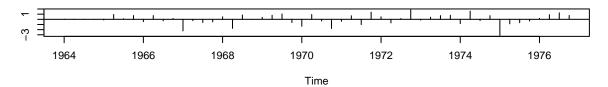


```
#Se sale cerca del lag 5
#Box-Pierce test contrasta
# Ho: Independencia vs. H1: Dependencia
Box.test(modelo_automatico$residuals, lag =10)
```

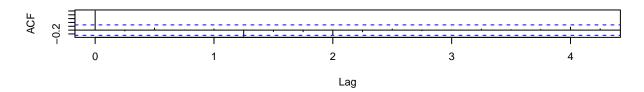
```
##
## Box-Pierce test
##
## data: modelo_automatico$residuals
## X-squared = 10.52, df = 10, p-value = 0.3962

#De manera conjunta, con la prueba de Ljung y Box
#HO:No estan correlacionados de manera conjunta
# vs
#H1:Estan correlacionados de manera conjunta
tsdiag(modelo_automatico)
```

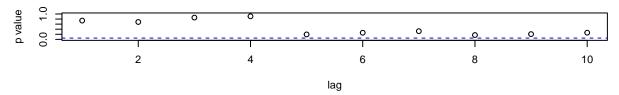
Standardized Residuals



ACF of Residuals



p values for Ljung-Box statistic

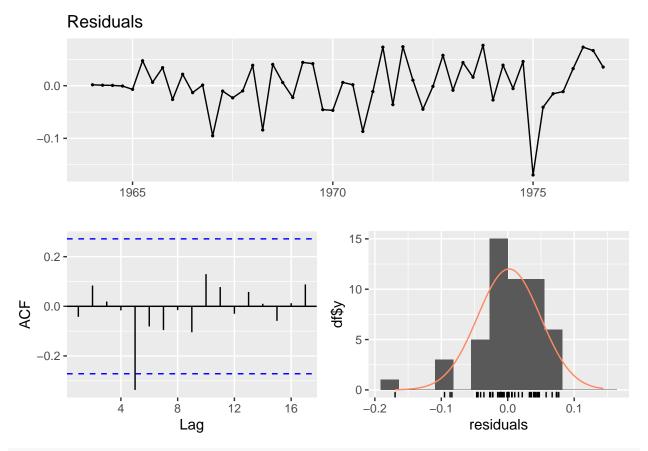


Pasa las pruebas

SARIMA2

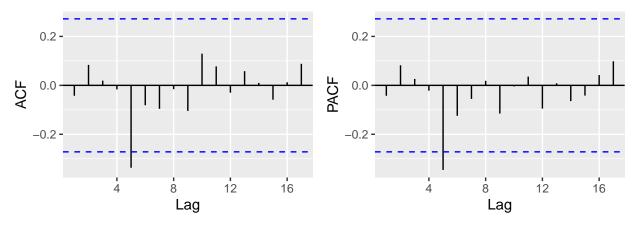
checkresiduals(SARIMA2\$residuals)

Warning in modeldf.default(object): Could not find appropriate degrees of
freedom for this model.



Residuales



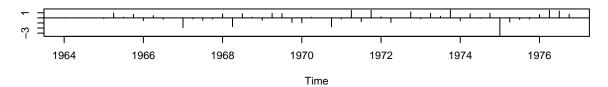


```
#Se sale en el lag 5;
#Box-Pierce test contrasta
# Ho: Independencia vs. H1: Dependencia
Box.test(SARIMA2$residuals, lag =10)
```

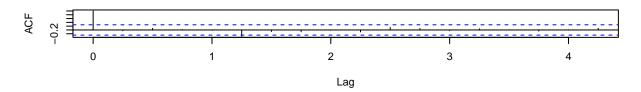
```
##
## Box-Pierce test
##
## data: SARIMA2$residuals
## X-squared = 8.6676, df = 10, p-value = 0.5639

#De manera conjunta, con la prueba de Ljung y Box
#HO:No estan correlacionados de manera conjunta
# vs
#H1:Estan correlacionados de manera conjunta
tsdiag(SARIMA2)
```

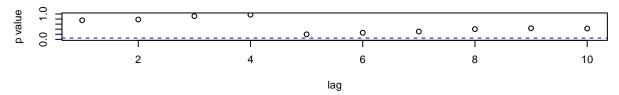
Standardized Residuals



ACF of Residuals



p values for Ljung-Box statistic



Pasa los test

Significancia de los coeficientes

ARIMA

ARIMA_int<-confint(ARIMA)
ARIMA_int

```
##
              2.5 %
                           97.5 %
## ar1
        -1.00862641 -0.980272709
        -1.00769140 -0.978709545
        -1.00209914 -0.995346337
##
  ar3
## ma1
        -0.26602909
                     0.449156386
## ma2
        -0.08125679
                     0.782926858
        -0.29018023
## ma3
                     0.596448780
##
  ma4
        -1.04713433 -0.001366895
        -0.80658372
                     0.194540758
##
  ma5
## ma6
        -0.60300938
                     0.098162739
        -0.67709203
                     0.110569510
## ma7
## ma8
        -1.21515818 -0.115601890
## ma9
        -0.79038082
                     0.474220236
## ma10 -0.68907677
                     0.438275283
  ma11 -0.41681849
                     0.520709182
  ma12 0.05951077
                     1.012705480
## ma13 -0.09375056
                     0.814269836
## ma14 -0.30496484
                     0.515268122
```

```
## ma15 -0.48960995 0.369774301
## ma16 -0.41846529 0.544986781
k=length(confint(ARIMA))/2
no_sign<-c()
for(i in 1:k){
  no_sign[i] <- (ARIMA_int[i] <0 & ARIMA_int[i+k] >0)
#No significatives
sum(no_sign)
## [1] 13
#Porcentaje no significativos
sum(no_sign)/k
## [1] 0.6842105
SARIMA
SARIMA_int<-confint(SARIMA)</pre>
SARIMA int
##
             2.5 %
                       97.5 %
## ar1 -1.0011218 1.1569074
## ar2 -0.1756916 2.0093433
        0.2387241 2.0433917
## ma2 -0.1281688 0.7236648
## sma1 -1.1870104 -0.4232753
k=length(confint(SARIMA))/2
no_sign<-c()
for(i in 1:k){
  no_sign[i]<-(SARIMA_int[i]<0 & SARIMA_int[i+k]>0)
#No significativos
sum(no_sign)
## [1] 3
#Porcentaje no significativos
sum(no_sign)/k
## [1] 0.6
SARIMA AUTOMÁTICO
SARIMA_DRIFT_int<-confint(modelo_automatico)</pre>
SARIMA_DRIFT_int
##
             2.5 %
                       97.5 %
## ar1 -0.0195618 0.5560452
## sma1 -0.9272078 -0.4205821
k=length(confint(modelo_automatico))/2
no_sign<-c()
for(i in 1:k){
  no_sign[i] <- (SARIMA_DRIFT_int[i] <0 & SARIMA_DRIFT_int[i+k] >0)
```

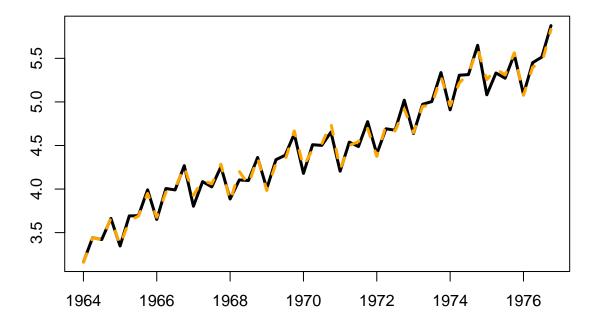
```
}
#No significativos
sum(no_sign)
## [1] 1
#Porcentaje no significativos
sum(no_sign)/k
## [1] 0.5
SARIMA2
SARIMA2_int<-confint(SARIMA2)
SARIMA2 int
##
              2.5 %
                          97.5 %
## ar1 -0.1077112 1.11718746
## ma1 -0.8933044 0.37395118
## sar1 -1.5230961 -0.38888560
## sar2 -1.1437314 0.10671376
## sar3 -0.7943175 0.12091960
## sma1 -0.2783978 1.30557906
## sma2 -1.0582232 0.08599028
k=length(confint(SARIMA2))/2
no_sign<-c()
for(i in 1:k){
  no_sign[i] <- (SARIMA2_int[i] <0 & SARIMA2_int[i+k] >0)
}
#No significativos
sum(no_sign)
## [1] 6
#Porcentaje no significativos
sum(no_sign)/k
## [1] 0.8571429
Por la simplicidad del modelo (Al tener solo 2 coeficientes), ser el que poseeel menor porcentaje de coeficientes
no significativos y además, pasar todos los tests (En normalidad pasó al menos Anderson-Darling) y tener los
mejores índices, elegimos el ajustado por auto.arima
                            \therefore Elegimos el SARIMA(1,1,0) \times (1,1,0)_{[4]}
¿Ajustará mejor si le quitamos el coeficiente no significativo?
modelo_automatico2<-arima(Serie_sq,order=c(0,1,0),seasonal=list(order=c(0,1,1),period=4))
modelo_automatico2
```

```
##
## Call:
## arima(x = Serie_sq, order = c(0, 1, 0), seasonal = list(order = c(0, 1, 1),
## period = 4))
##
## Coefficients:
```

```
## sma1
## -0.7373
## s.e. 0.1242
##
## sigma^2 estimated as 0.003296: log likelihood = 66.04, aic = -130.09
Gráfica
```

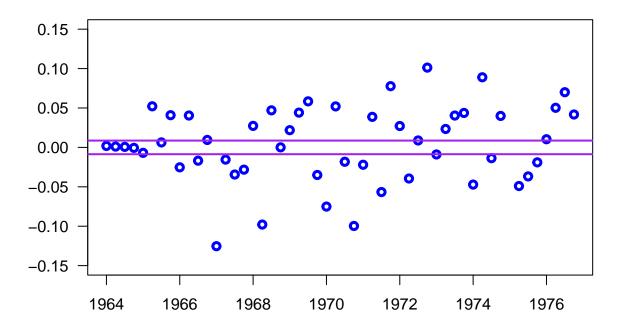
```
AUTOMATICO_ajuste <- Serie_sq - residuals(modelo_automatico)
ts.plot(Serie_sq, lwd=3, main="Comparación ", ylab="", xlab="")
points(AUTOMATICO_ajuste, type = "l", col ="orange", lty = 2, lwd=3)
```

Comparación



```
###Y ahora los residuales
plot(modelo_automatico$residuals, type="p", col="blue", ylim=c(-.15,.15), ylab="", xlab="", main="Datos
abline(h=3*(var(modelo_automatico$residuals)), col="purple", lwd=2)
abline(h=-3*(var(modelo_automatico$residuals)), col="purple",lwd=2)
```

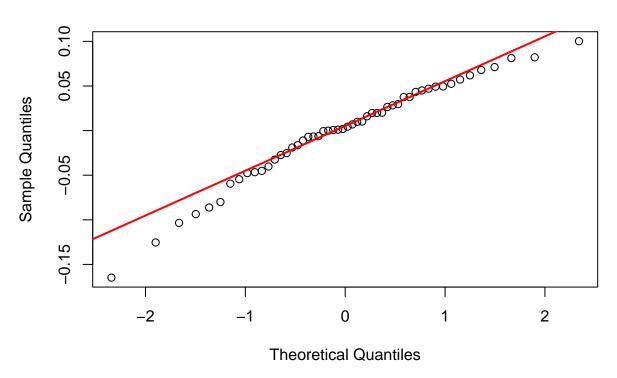
Datos discrepantes



Normalidad

#SARIMA AUTOMÁTICO

qqnorm(modelo_automatico2\$residuals)
qqline(modelo_automatico2\$residuals, col="red", lwd=2)

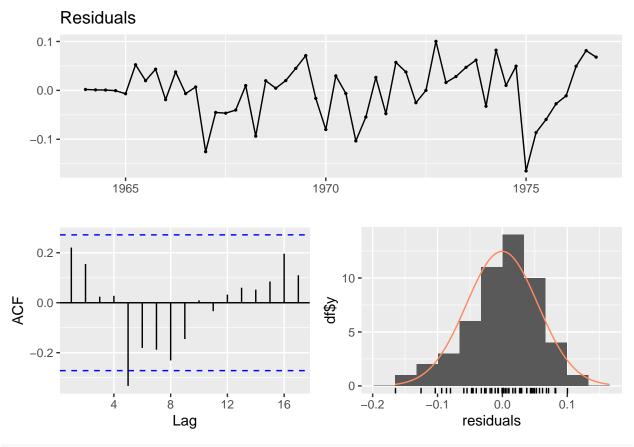


```
#Prueba Anderson-Darling
ad.test(modelo_automatico2$residuals)
##
##
    Anderson-Darling normality test
##
## data: modelo_automatico2$residuals
## A = 0.50391, p-value = 0.1954
#Prueba de Shapiro
shapiro.test(modelo_automatico2$residuals)
##
##
    Shapiro-Wilk normality test
## data: modelo_automatico2$residuals
## W = 0.9657, p-value = 0.1382
#Jarque-Bera Test. tseries
jarque.bera.test(modelo_automatico2$residuals)
##
    Jarque Bera Test
##
##
## data: modelo_automatico2$residuals
## X-squared = 5.0001, df = 2, p-value = 0.08208
```

Varianza constante

```
#SARIMA AUTOMÁTICO
Y <- as.numeric(modelo_automatico2$residuals)
X <- 1:length(modelo_automatico2$residuals)</pre>
bptest(Y ~ X)
##
## studentized Breusch-Pagan test
## data: Y ~ X
## BP = 3.2205, df = 1, p-value = 0.07272
Media 0
t.test(modelo_automatico2$residuals,mu=0)
##
## One Sample t-test
## data: modelo_automatico2$residuals
## t = -0.052502, df = 51, p-value = 0.9583
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
## -0.01574752 0.01494485
## sample estimates:
##
       mean of x
## -0.0004013333
Residuales no correlacionados
checkresiduals(modelo_automatico2$residuals)
```

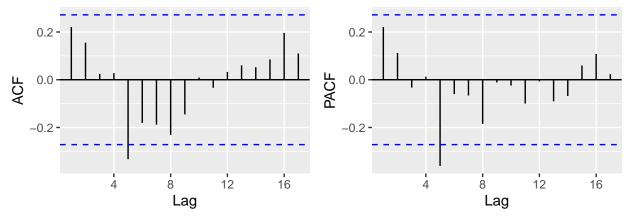
```
## Warning in modeldf.default(object): Could not find appropriate degrees of
## freedom for this model.
```



ggtsdisplay(modelo_automatico2\$residuals,main="Residuales")





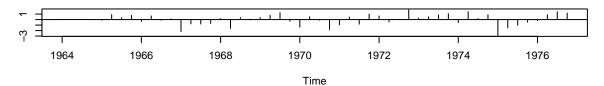


```
#Se sale cerca del lag 5
#Box-Pierce test contrasta
# Ho: Independencia vs. H1: Dependencia
Box.test(modelo_automatico2$residuals, lag =10)
```

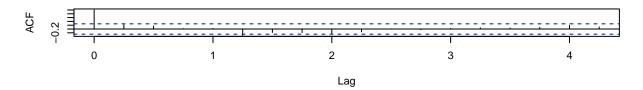
```
##
## Box-Pierce test
##
## data: modelo_automatico2$residuals
## X-squared = 17.046, df = 10, p-value = 0.07335

#De manera conjunta, con la prueba de Ljung y Box
#HO:No estan correlacionados de manera conjunta
# vs
#H1:Estan correlacionados de manera conjunta
tsdiag(modelo_automatico2)
```

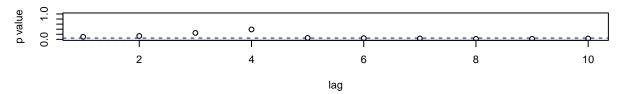
Standardized Residuals



ACF of Residuals



p values for Ljung-Box statistic



Coeficiente Significativo

```
SARIMA_DRIFT_int
## 2.5 % 97.5 %
## sma1 -0.9806686 -0.4939985
k=length(confint(modelo_automatico2))/2
no_sign<-c()
for(i in 1:k){
    no_sign[i]<-(SARIMA_DRIFT_int[i]<0 & SARIMA_DRIFT_int[i+k]>0)
}
#No significativos
sum(no_sign)
## [1] 0
#Porcentaje no significativos
sum(no_sign)/k
```

4. Forecasting

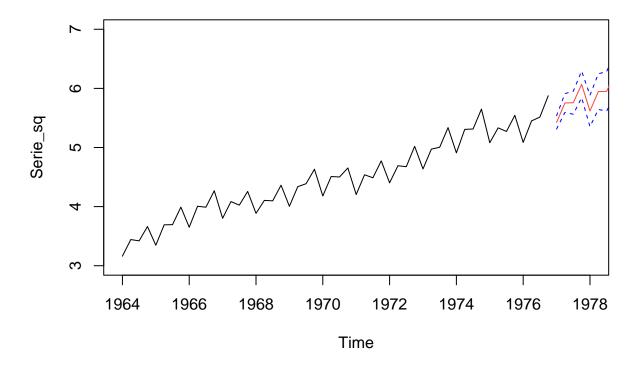
[1] 0

Con el modelo estimado, pronostique $n_{new} = 8$ (2 años) valores futuros (obtenga intervalos de predicción).

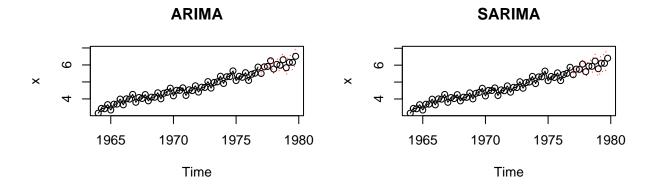
Usando el modelo mencionado anteiormente hacemos los pronosticos con la serie a la que le aplicamos la raíz cuadrada.

```
inicio<-start(Serie_sq)[1]
final<-end(Serie_sq)[1]
SARIMA_forecast <- predict(modelo_automatico2, n.ahead =8)$pred
SARIMA_forecast_se <- predict(modelo_automatico2, n.ahead = 8)$se
lower<-SARIMA_forecast - qnorm(0.975)*SARIMA_forecast_se
upper<-SARIMA_forecast + qnorm(0.975)*SARIMA_forecast_se
ts.plot(Serie_sq, xlim=c(inicio,final+2),ylim=c(3,7), main="Prediccion")
points(SARIMA_forecast, type = "l", col = 2)
points(lower, type = "l", col ="blue", lty = 2)
points(upper, type = "l", col ="blue", lty = 2)</pre>
```

Prediccion

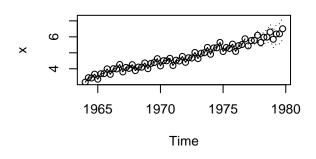


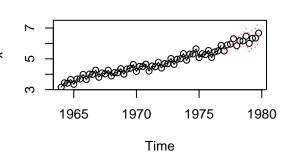
```
par(mfrow=c(2,2))
plot(ARIMA, col="red", main="ARIMA")
plot(SARIMA, col="red", main="SARIMA")
plot(modelo_automatico, main="SARIMA AUTOMÁTICO")
plot(SARIMA2, col="red", main="SARIMA2")
```



SARIMA AUTOMÁTICO

SARIMA2





par(mfrow=c(1,1))

Entonces, usando la transformación inversa que es elevar al cuadrado, tenemos:

```
##
      Puntual Banda_inf Banda_sup
## 1 5.425457
              5.312915 5.538000
## 2 5.752631
                        5.911785
              5.593478
## 3 5.757476 5.562556 5.952396
## 4 6.068250
              5.843177
                        6.293324
## 5 5.618335
              5.352146
                        5.884525
## 6 5.945509
              5.643764
                        6.247255
## 7 5.950354
              5.616822
                        6.283887
## 8 6.261128 5.898585
                        6.623672
```

Predicciones_normales<-Predicciones_raiz**2
Predicciones_normales</pre>

```
##
      Puntual Banda_inf Banda_sup
               28.22707
                         30.66944
## 1 29.43559
## 2 33.09277
               31.28699
                         34.94920
## 3 33.14853
               30.94203
                         35.43102
## 4 36.82366
               34.14272
                         39.60592
## 5 31.56569
               28.64547
                         34.62763
## 6 35.34908 31.85207
                         39.02819
```

```
## 7 35.40671 31.54869 39.48723
## 8 39.20173 34.79331 43.87302

ts.plot(Serie, xlim=c(inicio[1],final[1]+2.5),ylim=c(9,50), main="Predicción normal")
points(SARIMA_forecast**2, type = "l", col = 3)
points(lower**2, type = "l", col = "blue", lty = 2)
points(upper**2, type = "l", col = "blue", lty = 2)
```

Predicción normal

