ECON 613: Applied Econometrics

Methods

Overview: Linear Models

- Study the relationship between an outcome variable y and a set of regressors x.
  - Conditional Prediction.
  - Causal inference.
  - Example: propensity to consume.
- Loss function approach

$$L(e) = L(y - \hat{y})$$

where $\hat{y} = E(y \mid x)$ is a predictor of y, and the error $e = y - \hat{y}$

Squared Loss Function

- Squared error loss: $L(e) = e^2$
- Optimization problem

$$\min_{\beta} \sum_{i}^{N} (y_i - f(x_i, \beta))^2$$

Linear Prediction

- $E[y \mid x] = x'\beta$
- OLS

$$y = x\beta + e$$

- Derivation

$$\begin{aligned} L(\beta) &= (y - x\beta)'(y - x\beta) \\ &= y'y - 2y'x\beta + \beta'x'x\beta \end{aligned}$$

Then

$$\frac{\partial L(\beta)}{\partial \beta} = -2x'y + 2x'x\beta = 0$$

- Formula

$$\hat{\beta} = (x'x)^{-1}x'y$$

Other loss functions?

- Median regression, also known as least-absolute-deviations (LAD) regression,minimizes

$$\sum_i |e_i|$$

- No close form solution.

Properties

see 4.4.4 and 4.4.5.

Properties of an estimator

- Unbiasedness: $E(\hat{\theta}) = \theta$.
- Consistency: $plim \hat{\theta}_n = \theta$.
- Efficiency: Reach Cramer-Rao lower bound asymptotically.

Application: More guns, less crime?

- Using cross sectional time series data for US counties from 1977 to 1992, we find that allowing citizens to carry concealed weapons deters violent crimes, without increasing accidental deaths.
- Thoughts?

A dataset: Guns in the AER package

- state: factor indicating state.
- year: factor indicating year.
- violent: violent crime rate (incidents per 100,000).
- murder: murder rate (incidents per 100,000).
- robbery: robbery rate (incidents per 100,000).
- prisoners: incarceration rate in the state in the previous year.
- afam: percent of state population that is African-American.
- cauc: percent of state population that is Caucasian.
- male: percent of state population that is male.
- population: state population.
- income: real per capita personal income in the state ($).
- density: population per square mile of land area.
- law: factor. Does the state have a shall carry law in effect in that year?

Table: Statistical models

	Model 1	Model 2	Model 3	Model 4
(Intercept)	6.13^{***}	2.98^{***}	4.04^{***}	3.97^{***}
	(0.02)	(0.54)	(0.39)	(0.47)
lawyes	-0.44^{***}	-0.37^{***}	-0.05^{*}	-0.03
	(0.04)	(0.03)	(0.02)	(0.02)
prisoners		0.00^{***}	-0.00	0.00
		(0.00)	(0.00)	(0.00)
density		0.03^{*}	-0.17^{*}	-0.09
		(0.01)	(0.09)	(0.08)
income		0.00	-0.00	0.00
		(0.00)	(0.00)	(0.00)
population		0.04^{***}	0.01	-0.00
		(0.00)	(0.01)	(0.01)
afam		0.08^{***}	0.10^{***}	0.03
		(0.02)	(0.02)	(0.02)
cauc		0.03^{***}	0.04^{***}	0.01
		(0.01)	(0.01)	(0.01)
Year and state FE				X
R^2	0.09	0.56	0.94	0.96
Adj. R^2	0.09	0.56	0.94	0.95
Num. obs.	1173	1173	1173	1173
RMSE	0.62	0.43	0.16	0.14

$^{***}p < 0.001$, $^{**}p < 0.01$, $^{*}p < 0.05$

Model Selection

- $R - squared$
- Model selection criterias: AIC, BIC
- Endogeneity concerns.

Principle of a monte-carlo study

- Generate S independent data sets under the conditions of interest
- Compute the numerical value of the estimator/test statisticT(data)for each data set $T_1, ..., T_S$
- If S is large enough,summary statistics across $T_1, ..., T_S$.

Monte-carlo

- Simulate *xvec* drawing from a normal distribution
- Set a=2, and b=0.1, and construct $yvec = a + bxvec + \epsilon$ with ϵ $Normal(0, 1)$.

Potential problems: sample size

Table: Statistical models

	$n = 10$	$n = 100$	$n = 1000$	$n = 10000$	$n = 10000$
(Intercept)	1.69^{***}	1.90^{***}	2.02^{***}	1.99^{***}	2.00^{***}
	(0.28)	(0.10)	(0.03)	(0.01)	(0.00)
xvec	1.01^{*}	-0.05	0.04	0.11^{***}	0.10^{***}
	(0.39)	(0.11)	(0.03)	(0.01)	(0.00)
R^2	0.45	0.00	0.00	0.01	0.01
Adj. R^2	0.38	-0.01	0.00	0.01	0.01
Num. obs.	10	100	1000	10000	100000
RMSE	0.85	0.96	1.01	1.00	1.00

$^{***}p < 0.001$, $^{**}p < 0.01$, $^{*}p < 0.05$

Miss-specification (1)

- Simulate *xvec* drawing from a normal distribution
- Set a=2, and b=0.1, and construct $yvec = a + bxvec + \epsilon$ with ϵ $uniformly(0, 5)$.
- Thoughts?

Error terms misspecification

Table: Statistical models

	$n = 10$	$n = 100$	$n = 1000$	$n = 10000$	$n = 10000$
(Intercept)	5.06^{***}	4.37^{***}	4.50^{***}	4.48^{***}	4.50^{***}
	(0.41)	(0.15)	(0.05)	(0.01)	(0.00)
xvec	0.40	0.26	0.11^{*}	0.10^{***}	0.10^{***}
	(0.45)	(0.17)	(0.05)	(0.01)	(0.00)
R^2	0.09	0.02	0.01	0.00	0.00
Adj. R^2	-0.03	0.01	0.00	0.00	0.00
Num. obs.	10	100	1000	10000	100000
RMSE	1.30	1.53	1.45	1.44	1.45

$^{***}p < 0.001$, $^{**}p < 0.01$, $^{*}p < 0.05$

Miss-specification (2)

- Simulate *xvec* drawing from a normal distribution
- Set a=2, and b=0.1, and construct $yvec = a + b\exp(xvec) + \epsilon$ with ϵ $Normal(0, 1)$.

Functional form misspecification

Table: Statistical models

	$n = 10$	$n = 100$	$n = 1000$	$n = 10000$	$n = 10000$
(Intercept)	5.21^{***}	4.52^{***}	4.67^{***}	4.64^{***}	4.66^{***}
	(0.41)	(0.15)	(0.05)	(0.01)	(0.00)
xvec	0.48	0.31	0.18^{***}	0.17^{***}	0.17^{***}
	(0.45)	(0.17)	(0.05)	(0.01)	(0.00)
R^2	0.12	0.03	0.02	0.01	0.01
Adj. R^2	0.02	0.03	0.01	0.01	0.01
Num. obs.	10	100	1000	10000	100000
RMSE	1.29	1.52	1.46	1.45	1.45

$^{***}p < 0.001$, $^{**}p < 0.01$, $^{*}p < 0.05$

Maximum Likelihood

Game 1

GMM

Numerical Optimization

Inference

Introduction to MLE

Consider a parametric model in which the joint distribution of $Y = (Y_1, \ldots, Y_n)$ has a density $\ell(y, \theta)$ with respect to a measure μ. Then consider $P_\theta = \ell(y, \theta)\mu$ where $\theta \in \Theta \in \mathbb{R}^p$. Once $y = (y_1, \ldots, y_n)$ is observed, the maximum likelihood method consists of estimating the parameter θ a value $\hat{\theta}(y)$ that maximizes the likelihood function $\theta \to \ell(y, \theta)$. Formally, a maximum likelihood estimator of θ is a solution to the maximization problem

$$\max_\theta \ell(Y; \theta)$$

or

$$\max_\theta \log(\ell(Y; \theta))$$

Feasible examples: Poisson distribution

Consider a dependent variable that takes only non negative integer values $0, 1, 2, \ldots$, and one assumes that the dependent variable follows a Poisson distribution, and we wishes to estimate the Poisson parameter.

- Given $y_i \sim f(\lambda, y_i) = \frac{\exp(-\lambda)\lambda^{y_i}}{y_i!}$
- Likelihood $\mathcal{L}(y;\lambda) = \prod_{i=1}^{N} \frac{\exp(-\lambda)\lambda^{y_i}}{y_i!} = \frac{\exp(-N\lambda)\lambda^{\sum_i^N y_i}}{\prod_{i=1}^{N} y_i!}$
- Log likelihood
  $\log \mathcal{L}(y;\lambda) = -N\lambda + \sum_i^N y_i \log(\lambda) - \sum_i^N \log(y_i!)$
- Estimate

$$\frac{\partial \log \mathcal{L}(y;\lambda)}{\partial \lambda} = 0 \Longrightarrow \widehat{\lambda} = \frac{\sum_i^N y_i}{N}$$

Feasible examples: Least Squares

- Normality assumption $e \sim \mathbb{N}(0, \sigma^2)$, then $y \sim \mathbb{N}(x\beta, \sigma^2)$.
- Likelihood $L(\beta) = (2\pi\sigma^2)^{-\frac{N}{2}} \exp(-0.5\sigma^{-2}(y - x\beta)'(y - x\beta))$
- log likelihood $\log L(\beta) = -\frac{N}{2} \log \sigma^2 - \frac{1}{2\sigma^2}(y - x\beta)'(y - x\beta)$
- $\beta = (x'x)^{-1}x'y$

Some difficulties

- Non-uniqueness of the Likelihood Function
- Non-existence of a solution to the Maximization Problem
- Multiple Solutions to the Maximization Problem

Asymptotic Properties (1): Convergence

Definition

Under a set of regularity conditions, there exists a sequence of maximum likelihood estimators converging almost surely to the true parameter value θ_0

- The variables $Y_i, i = 1, 2, \ldots$ are independent and identically distributed with density $f(y;\theta), \theta \in \Theta \in \mathbb{R}^p$
- The parameter space Θ is compact.
- The log likelihood function $\mathcal{L}(y, \theta)$ is continuous in θ and is a measurable function of y.
- The log-likelihood function is such that $(1/n)\mathcal{L}_n(y, \theta)$ converges surely to $E_{\theta_0} log(f(Y_i; \theta))$ uniformly in $\theta \in \Theta$. $E_{\theta_0} log(f(Y_i; \theta))$ exists.

Asymptotic Properties (2): Asymptotic Normality

- The log likelihood function $\mathcal{L}_n(\theta)$ is twice continuously differentiable in an open neighborhood of θ_0
- The matrix (Fisher Information Matrix) $\mathcal{I}_1(\theta_0) = E_{\theta_0}\left(-\frac{\partial^2 \log f(Y_1;\theta_0)}{\partial\theta\partial\theta'}\right)$

Definition

$$\sqrt{n}(\hat{\theta}_n - \theta_0) \to \mathbb{N}(0, \mathcal{I}_1(\theta_0)^{-1}).$$

Concentrated Likelihood Function

Definition

Let the parameter set $\theta = (\alpha, \beta)$. The solutions $\hat{\theta} = (\hat{\alpha}, \hat{\beta})$ to the mazimization problem $\max_{\alpha,\beta} \log \mathcal{L}(y; \alpha, \beta)$ can be obtained via the following two-step procedure:

a) Maximize the log-likelihood function with respect α given β. The maximum value is attained for values of α in a set $A(\beta)$ depending on the parameter β. Thus, if $\alpha \in A(\beta)$, the log-likelihood value is

$$\log \mathcal{L}_c(y; \beta) = \max_{\alpha} \log \mathcal{L}(y; \alpha, \beta)$$

The mapping $\log \mathcal{L}_c$ is called the concentrated (in α) log likelihood function.

b) In a second step, maximize the concentrated log-likelihood function with respect to β.

Application

Consider the likelihood

$$\mathcal{L}(y,\beta,\sigma) = -\frac{n}{2}\log 2\pi - \frac{n}{2}\log\sigma^2 - \frac{1}{2\sigma^2}(y - x\beta)'(y - x\beta)$$

- First step

$$\frac{\partial \mathcal{L}(y;\beta,\sigma)}{\partial\sigma} = -\frac{n}{\sigma} + \frac{1}{\sigma^3}(y - x\beta)'(y - x\beta) = 0$$

  Then

$$\sigma^2(\beta) = \frac{1}{n}(y - x\beta)'(y - x\beta)$$

- Substituting $\sigma^2(\beta)$ into the likelihood

$$\mathcal{L}_c(y,\beta,\sigma) = -\frac{n}{2}\log 2\pi - \frac{n}{2}\log\frac{1}{n}(y - x\beta)'(y - x\beta) - \frac{n}{2}$$

Hypothesis Testing

Three procedures to do tests

Likelihood Ratio

- The likelihood ratio statistic is

$$LR = 2(\ell(\theta) - \ell(\tilde{\theta}))$$

  where $\hat{\theta}$ and $\tilde{\theta}$ are the restricted and unrestricted maximum likelihood estimates of θ.
- Wilk's theorem shows that

$$LR \sim \chi^2(r)$$

  where r is the number of restrictions.

Additional Tests

- Wald Test
- LM test

We will see in GMM.

In practice

- The regularity conditions are strong.
- What happens if we weaken them?

Problem 1

Number of parameters increases with the number of observations

- Convergence holds
- Estimates may be biased

Problem 2

True parameter value θ_0 does not belong to Θ: The model is misspecified

- Convergence holds to a parameter that is not the true parameter.

Problem 3

Correlated Observations

- Convergence does not hold.

Problem 4

Discontinuity of the likelihood function

- Numerical problems.

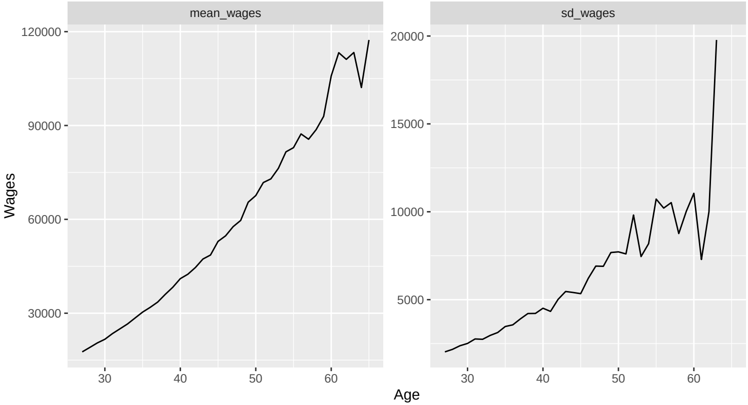
Problem 5

Known parameter space

- Constrained Optimization

Game 1 (1)

Figure: Wages



Game 1 (2)

$$y_i = m(x_i) + \epsilon \tag{1}$$

- Find a way to approximate this relationship
- The winning team gets $+5$ on assignment.

Method of Moments

- Orthogonality condition in Linear Models

$$E(x(y' - x)) = 0 \tag{2}$$

- Moment Condition

$$\frac{1}{N}\sum_i x_i(y_i - x'\beta) \tag{3}$$

- Moment Estimator

$$\hat{\beta}_{\mathrm{MM}} = (\sum_i x_i x_i')^{-1}(\sum_i x_i y_i) \tag{4}$$

Nonlinear Model

- Consider

$$Y_i = g(X_i, b_0) + u_i$$

- Orthogonality Condition

$$E[X'(y - g(X, b_0))] = 0$$

- Moments condition

$$E_0 h(Y, X, a_0) = 0$$

- The function h is H-dimensional and the parameter a is of size K.

Formal Idea

Definition

The basic idea of generalized method of moments is to choose a value for a such that the sample mean is closest to zero.

$$\frac{1}{n}\sum_{i=1}^{n} h(Y_i, X_i, a)$$

Formal Definition

Definition

Let $\mathbb{S}_n$ be an $(H \times H)$ symmetric positive definite matrix that may depend on the observations. The generalized method of moments (GMM) estimator associated with $\mathbb{S}_n$ is a solution $\tilde{a}_n(\mathbb{S}_n)$ to the problem

$$min_a \left[\sum_{i=1}^{n} h(Y_i, X_i, a)\right]' \mathbb{S}_n \left[\sum_{i=1}^{n} h(Y_i, X_i, a)\right]$$

Assumptions

- H1 The variables (Y_i, X_i) are independent and identically distributed.
- H2 The expectation $E_0 h(Y, X, a)$ exists and is zero when a is equal to the true value a_0 of the parameter of interest.
- H3 The matrix $\mathbb{S}_n$ converges almost surely to a nonrandom matrix $\mathbb{S}_0$
- H4 The parameter a_0 is identified from the equality constraints, i.e. $E_0 h(Y, X, a)' \mathbb{S}_0 E_0 h(Y, X, a) = 0$
- H5 The parameter value a_0 is known to belong to a compact set $\mathcal{A}$
- H6 The quantity $(1/n) \sum_{i=1}^{n} h(Y_i, X_i, a)$ converges almost surely and uniformly in a to $E_0 h(Y, X, a)$
- H7 The function $h(Y, X, a)$ is continuous in a
- H8 The matrix $\left[E_0 \dfrac{h(Y, X, a)}{\partial a}\right]' \mathbb{S}_0 \left[E_0 \dfrac{h(Y, X, a)}{\partial a'}\right]$ is nonsingular, which implies $H \geq K$.

Asymptotic Normality

Under the assumptions, we have

$$\sqrt{n}(\tilde{a}_n(\mathbb{S}_n) - a_0) \sim \mathbb{N}(0, \Sigma(\mathbb{S}_0))$$

where

$$\Sigma(\mathbb{S}_0) = \left(\left[E_0 \frac{h(Y,X,a)}{\partial a} \right]' \mathbb{S}_0 \left[E_0 \frac{h(Y,X,a)}{\partial a'} \right] \right)^{-1}$$
$$\left(\left[E_0 \frac{h(Y,X,a)}{\partial a} \right]' \mathbb{S}_0 V_0(h(Y,X,a_0)) \mathbb{S}_0 \left[E_0 \frac{h(Y,X,a)}{\partial a'} \right] \right)^{-1}$$
$$\left(\left[E_0 \frac{h(Y,X,a)}{\partial a} \right]' \mathbb{S}_0 \left[E_0 \frac{h(Y,X,a)}{\partial a'} \right] \right)^{-1}$$

Optimal GMM

- $\mathbb{S}_0$ is not known.
- Two-step procedure
  - Estimate

$$min_a \left[\sum_{i=1}^{n} h(Y_i, X_i, a)\right]' I \left[\sum_{i=1}^{n} h(Y_i, X_i, a)\right]$$

    where I is the identity matrix, and recover $\hat{a}$.
  - Matrix of variance/covariance

$$\hat{\mathbb{S}} = \frac{1}{N}\sum_{i=1}^{n} h(Y_i, X_i, \hat{a}) h(Y_i, X_i, \hat{a})'$$

Relationship to IV.

- Nonlinear 2SLS is a very good application of GMM.

Inference

- Over identification test see iv section.

Applications

- Matrix of variance/covariance in practice
- Indirect Inference
- Simulated method of moments

Maximum Likelihood

Game 1

GMM

Numerical Optimization

Inference

Numerical Optimization

Most maximum likelihood estimates require numerical optimization.

Primer on optimization

Definition

$$\min_x f(x)$$

- $x \in \mathbb{R}^n$
- f is a smooth function.

Existence: Weierstrass theorem

A point or a vector x^* is a global minimizer if $f(x^*) \leq f(x) \, \forall \, x$.

Maximization Vs Minimization

Let $-f$ denote the function whose value at any value at any x is $-f(x)$. Then,

1. x is the maximum of f if and only if x is a minimum of $-f$
2. z is a minimum of f if and only if z is a maximum of $-f$

Necessary conditions

1. If x^* is local minimizer and f is continuously differentiable in an open neighborhood of x^*, then $\nabla f(x^*) = 0$
2. If x^* is local minimizer and $\nabla^2 f$ exists and is continuous in an open neighborhood of x^*, then $\nabla f(x^*) = 0$ and $\nabla^2 f(x^*) = 0$

Likelihood setup

The likelihood function is defined by:

$$\mathcal{L}_n(\theta) = \prod_{i=1}^{n} f(X_i; \theta)$$

The necessary conditions for optimization yield the regularity conditions

Unfeasible example:

Any nonlinear model

Numerical optimization

- Local optimization: the best minimum/maximum in a vicinity - usually defined by a convergence criteria.
- Global optimization: Best of all local minimas/maximas.

Numerical optimization - Local Optimization

Overview

1. **Line Search**: Starting from an initial value, choose a direction and search along this direction to find a new iterate
2. **Trust region**: Use previous estimates of the objective function, to construct a **synthetic** or **model** function whose behavior near the current point is similar to the objective function, and search only over a region, *trust region*, with the underlying idea that the model function is a good approximate over the trust region.

Line Search

Idea:

$$x_{k+1} = x_k + \alpha_k d_k$$

where d_k is a direction to be evaluated, and α_k a scaling parameter.

The variants of numerical optimization

1. Steepest descent: $d_k = -\nabla f(x_k)$
2. Newton direction: $d_k = -(\nabla^2 f(x_k))^{-1} \nabla f(x_k)$
3. Quasi-Newton direction: $d_k = -(B(x_k))^{-1} \nabla f(x_k)$
4. Derivative free.

Properties

- Robustness: Perform well for various problems and starting values.
- Efficiency:
- Accuracy: Identify a solution with precision, not sensitive to starting values

One parameter optimization

- Bissection
- Secant Method

Codes

```
bisection <- function(f, a, b, n , tol ) {
  # Check the signs of the function.
  if (!(f(a) < 0) && (f(b) > 0)) {
    stop()} else if ((f(a) > 0) && (f(b) < 0)) {
    stop()}
  for (i in 1:n) {
    c <- (a + b) / 2 # Calculate midpoint
    # If the function equals 0 at the midpoint
    if ((f(c) == 0) || ((b - a) / 2) < tol) {
      return(c)  }
    # If another iteration is required,
    # check the signs of the function
    ifelse(sign(f(c)) == sign(f(a)),
           a <- c,
           b <- c)
  }
  # If the max number of iterations is reached
  print('Too␣many␣iterations')
```

Recover the scaling parameter

- Solve the function $\phi(\alpha) = f(x_k + \alpha d_k)$

Quasi-Newton methods

How to approximate the hessian such that:

- Reduce the computation time (Use only gradient instead of hessian)
- Increase convergence rate

Conjugate Gradient- FR

- Given x0
- Evaluate $f_0 = f(x0)$, $\nabla f_0 = \nabla f(x_0)$
- Set $d_0 = -\nabla f_0$, $k = 0$
- **While** $\nabla f_k \neq 0$
  - Set $x_{k+1} = x_k + \alpha_k d_k$
  - Evaluate ∇f_{k+1}, then:
    - $\beta^{FR}_{k+1} = \dfrac{\nabla f'_{k+1} \nabla f_{k+1}}{\nabla f'_k \nabla f_k}$
    - $d_{k+1} = -\nabla f_{k+1} + \beta^{FR}_{k+1} d_k$
    - $k = k + 1$
- **end(while)**

BFGS

- Given x0
- Evaluate $f_0 = f(x0)$, $\nabla f_0 = \nabla f(x_0)$, $H_0 = I$
- Set $k = 0$
- **While** $||\nabla f_k|| > \epsilon$
  - Compute direction $d_k = -H_k \nabla f_k$
  - Set $x_{k+1} = x_k + \alpha_k d_k$
  - Evaluate ∇f_{k+1}, then:
    - set $s_k = x_{k+1} - x_k$, $y_k = \nabla f_{k+1} - \nabla f_k$ and $\rho_k = \dfrac{1}{y_k' s_k}$
    - Update $H_{k+1} = (I - \rho_k s_k y_k') H_k (I - \rho_k y_k s_k') + \rho_k s_k s_k'$
- **end(while)**

Problems

- non differentiable functions
- disconnected and non-convex feasible space
- discrete feasible space
- large dimensionality
- multiple local minimas

Derivative Free

- Nelder-Mead
- Simulated annealing
- Divided Rectangles Method
- Genetic Algorithms
- Particle Optimization

Maximum Likelihood

Game 1

GMM

Numerical Optimization

Inference

How to get standard error

- Fisher Information Matrix
- Sandwich formula

Introduction to Bootstrap

- Inference for small samples basically...
- Inference for unknown distribution...
- Computer intensive resampling method...
  - Using data to generate new data

Applications: Inference in estimation

- Bootstrap samples
- Parallel implementation

Applications: Inference post estimation

- Bootstrap samples
- Marginal effects

Applications: Testing

- Bootstrap samples
- Approach to testing

Applications: Choosing the number of R

- Objectives and Constraints
- R=49,99,199,499,999,9999....

Delta Method

- Consider $X \sim \mathbb{N}(\mu, \sigma^2)$, and assume you are interested in $E(g(X))$ and $Var(g(X)$
- Approximation

$$g(x) = g(\mu) + g'(\mu)(x - \mu) \tag{5}$$

and then

$$E(g(x)) \approx g(E(x)) \tag{6}$$

$$Var(g(x)) \approx g'(E(x))^2 Var(x) \tag{7}$$

- Usage?