ECON 613: Applied Econometrics Methods for Panel Data

March 28, 2019

Linear Models Applications

Econometrics of Policy Evaluation

Applications

Nonlinear Pane

Introduction (1)

- Data on cross section that is observed over several unit of time.
- In microeconometrics, panel are usually short.

Introduction (2)

- ▶ The error is correlated over time..
- Examples
- Open possibilities..

Introduction (3)

Consider the following Model

$$Y_{it} = \alpha_i + \gamma_{j(t)} + \beta X_{it} + \epsilon_{it}$$
 (1)

- Estimation of fixed effects
- Correlation between the fixed effects
- Estimation issues

Introduction (4)

Consider the following DGP:

- ▶ 1,000 individuals over 10 periods.
- $Y_{it} = \alpha_i + \beta X_{it} + \epsilon_{it}$
- Parametrization
 - $\beta = 1$
 - $ightharpoonup lpha_i \sim \textit{uniform}(0,1)$
 - $ightharpoonup \epsilon_i \sim \mathbb{N}(0,1)$

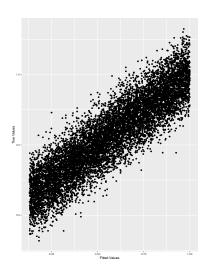
Pooled Estimation

	Model 1	
(Intercept)	0.49***	
,	(0.02)	
c(xMat)	0.93***	
	(0.00)	
R^2	0.87	
Adj. R ²	0.87	
Num. obs.	10000	
RMSE	1.05	

^{***}p < 0.001, **p < 0.01, *p < 0.05

Table: Statistical models

Fitted Values (1)



Introduction (5)

Consider the following DGP:

- ▶ 1,000 individuals over 10 periods.
- $Y_{it} = \alpha_i + \beta X_{it} + \epsilon_{it}$
- Parametrization
 - $\beta = 1$
 - $ightharpoonup lpha_i \sim \textit{uniform}(-10, 10)$
 - $ightharpoonup \epsilon_i \sim \mathbb{N}(0,1)$

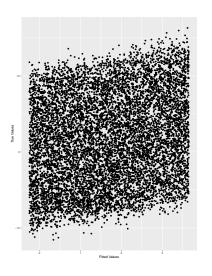
Pooled Estimation

	Model 1	
/1		
(Intercept)	-0.26^{*}	
	(0.11)	
c(xMat)	0.40***	
	(0.02)	
R^2	0.04	
Adj. R ²	0.04	
Num. obs.	10000	
RMSE	5.73	
*** .0.001 **	0.01 # 0.05	

^{***}p < 0.001, **p < 0.01, *p < 0.05

Table: Statistical models

Fitted Values (2)



Effects

- ▶ Pooled Estimation is a good starting point.
- ▶ Individual VS Time Effect.

Individual Effects

- Fixed Effects
- ► Random Effects
- Examples: Return to Education

Time Effects

- ► Long Panel Case
- ► Example: Seasonality?

Some Models (1)

Pooled Estimator

$$Y_{it} = \alpha + \beta X_{it} + \epsilon_{it}$$
 (2)

► Problems

Some Models (2)

► Between Estimator

$$\bar{y}_i = \alpha_i + \beta \bar{x}_i + \bar{\epsilon}_i \tag{3}$$

► Problems

Some Models (3)

► Within Estimator

$$y_{it} - \bar{y}_i = \beta(x_{it} - \bar{x}_i) + (\epsilon_{it} - \bar{\epsilon}_i)$$
 (4)

Problems

Some Models (4)

► First Difference Estimator

$$y_{it} - y_{i,t-1} = \beta(x_{it} - x_{i,t-1}) + (\epsilon_{it} - \epsilon_{i,t-1})$$
 (5)

Problems

More Guns, Less Crime

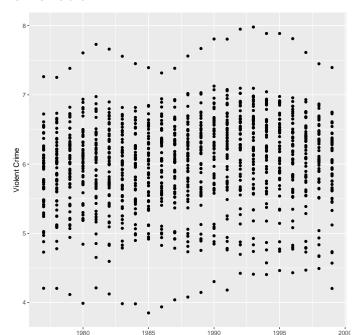
In a remarkable paper published in 1997, John Lott and David Mustard managed to set the agenda for much subsequent work on the impact of guns on crime in America by creating a massive data set of crime across all U.S. counties from 1977 through 1992 and amassing a powerful statistical argument that state laws enabling citizens to carry concealed handguns had reduced crime.1 The initial paper was followed a year later by an even more comprehensive and sustained argument to the same effect in a book solely authored by John Lott entitled More Guns, Less Crime (now in its second edition).

Data: Guns

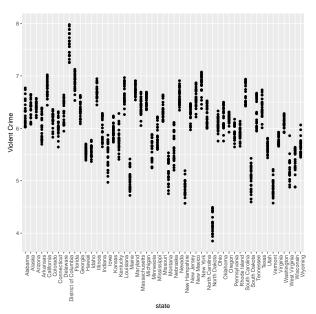
A data frame containing 1,173 observations on 13 variables.

- state: factor indicating state.
- year: factor indicating year.
- ▶ violent: violent crime rate (incidents per 100,000 members of the population).
- murder: murder rate (incidents per 100,000).
- robbery: robbery rate (incidents per 100,000).
- prisoners: incarceration rate in the state in the previous year
- afam: percent of state population that is African-American
- cauc: percent of state population that is Caucasian,
- ▶ male: percent of state population that is male
- population: state population, in millions of people.
- ▶ income: real per capita personal income in the state (US \$).
- density population per square mile of land area, divided by 1.000.
- ▶ law factor. Does the state have a shall carry law in effect in that year?

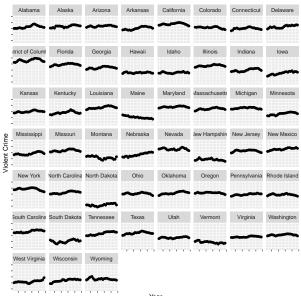
Overtime Variation



Cross-sectional Variation (1)



Cross-sectional Variation (2)



First regressions

	Violent	Violent Crime		bery
	Model 1	Model 2	Model 1	Model 2
(Intercept)	6.13***	2.98***	4.87***	0.90
	(0.02)	(0.54)	(0.03)	(0.77)
lawyes	-0.44***	-0.37***	-0.77***	-0.53***
	(0.04)	(0.03)	(0.06)	(0.05)
prisoners		0.00***		0.00***
		(0.00)		(0.00)
density		0.03*		0.09***
		(0.01)		(0.02)
income		0.00		0.00***
		(0.00)		(0.00)
population		0.04***		0.08***
		(0.00)		(0.00)
afam		0.08***		0.10***
		(0.02)		(0.02)
cauc		0.03***		0.03*
		(0.01)		(0.01)
male		0.01		0.03
		(0.01)		(0.02)
R ²	0.09	0.56	0.12	0.60
Adj. R ²	0.09	0.56	0.12	0.59
Num. obs.	1173	1173	1173	1173
RMSE	0.62	0.43	0.90	0.61

***p < 0.001, **p < 0.01, *p < 0.05

Table: Statistical models

Exploiting the Panel Structure

	Model 1	Model 2	Model 3
(Intercept)	4.04***	3.09***	3.97***
	(0.39)	(0.58)	(0.47)
lawyes	-0.05*	-0.29***	-0.03
	(0.02)	(0.03)	(0.02)
prisoners	-0.00	0.00***	0.00
	(0.00)	(0.00)	(0.00)
density	-0.17^*	-0.01	-0.09
	(0.09)	(0.01)	(80.0)
income	-0.00	0.00	0.00
	(0.00)	(0.00)	(0.00)
population	0.01	0.04***	-0.00
	(0.01)	(0.00)	(0.01)
afam	0.10***	0.10***	0.03
	(0.02)	(0.02)	(0.02)
cauc	0.04***	0.04***	0.01
	(0.01)	(0.01)	(0.01)
male	-0.05***	-0.04*	0.07***
	(0.01)	(0.02)	(0.02)
State FE	YES	NO	YES
TIME FE	NO	YES	YES
R ²	0.94	0.59	0.96
Adj. R ²	0.94	0.58	0.95
Num. obs.	1173	1173	1173
RMSE	0.16	0.42	0.14

***p < 0.001, **p < 0.01, *p < 0.05

Table: Statistical models

Data: EmplUK

Employment and Wages in the United Kingdom

An unbalanced panel of 140 observations from 1976 to 1984

▶ firm: firm index

year: year

sector: the sector of activity

emp: employment

wage: wages

capital: capital

output: output

Unbalanced Panel: Definitions

- ► Unbalanced panel: Definition
- ▶ What to do: Missing at random?

Unbalanced Panel: Solutions

- ► Testing for missingness at random.
- Missing at random
 - Imputation
 - ► Full sample
 - Non missing sample
- Not missing at random
 - Understand why?
 - Find an instrument

Description

Table:

Ν	Mean	St. Dev.	Min	Pctl(25)	Pctl(75)	Max
1,031	73.204	41.233	1	37	110	140
1,031	1,979.651	2.216	1,976	1,978	1,981	1,984
1,031	5.123	2.678	1	3	8	9
1,031	7.892	15.935	0.104	1.180	7.020	108.562
1,031	23.919	5.648	8.017	20.636	27.494	45.232
1,031	2.507	6.249	0.012	0.221	1.501	47.108
1,031	103.801	9.938	86.900	97.098	110.603	128.365
	1,031 1,031 1,031 1,031 1,031 1,031	1,031 73.204 1,031 1,979.651 1,031 5.123 1,031 7.892 1,031 23.919 1,031 2.507	1,031 73.204 41.233 1,031 1,979.651 2.216 1,031 5.123 2.678 1,031 7.892 15.935 1,031 23.919 5.648 1,031 2.507 6.249	1,031 73.204 41.233 1 1,031 1,979.651 2.216 1,976 1,031 5.123 2.678 1 1,031 7.892 15.935 0.104 1,031 23.919 5.648 8.017 1,031 2.507 6.249 0.012	1,031 73.204 41.233 1 37 1,031 1,979.651 2.216 1,976 1,978 1,031 5.123 2.678 1 3 1,031 7.892 15.935 0.104 1.180 1,031 23.919 5.648 8.017 20.636 1,031 2.507 6.249 0.012 0.221	1,031 73.204 41.233 1 37 110 1,031 1,979.651 2.216 1,976 1,978 1,981 1,031 5.123 2.678 1 3 8 1,031 7.892 15.935 0.104 1.180 7.020 1,031 23.919 5.648 8.017 20.636 27.494 1,031 2.507 6.249 0.012 0.221 1.501

Linear VS Log Specifications

	Log	Linear	
(Intercept)	0.34	8.25**	
	(0.86)	(3.11)	
log(wage)	-0.37^{***}		
	(0.06)		
log(capital)	0.81***		
	(0.01)		
log(output)	0.48**		
	(0.18)		
wage		-0.32***	
		(0.05)	
capital		2.11***	
		(0.04)	
output		0.02	
		(0.03)	
R ²	0.84	0.69	
Adj. R ²	0.84	0.69	
Num. obs.	1031	1031	
*** $p < 0.001$, ** $p < 0.01$, * $p < 0.05$			

Table: Statistical models

Fixed VS Random Effects

-	Random Effect	Fixed Effects	
(Intercept)	2.20**		
((0.15)		
log(wage)	-0.24***	-0.61^{***}	
J(J)	(0.05)	(0.03)	
log(capital)	0.61***	0.56***	
	(0.07)	(0.02)	
R ²	0.78	0.99	
Num. obs.	1031	1031	
*** $p < 0.001$, ** $p < 0.01$, * $p < 0.05$			

Table: Statistical models

Specification Problem

- Choosing between random and fixed effects;
- Durbin Wu Hausman Test

$$H = (\beta_{FE} - \beta_{RE})'(Var(\beta_{FE}) - Var(\beta_{RE}))'(\beta_{FE} - \beta_{RE})$$
 (6)

 $ightharpoonup H \sim \chi_2(rank(Var(\beta_{FE}) - Var(\beta_{RE}))$

Data: US STATES PRODUCTION

- state: stateyear: year
- region: the region
- pcap: public capital stock
- hwy: highway and streets
- water: water and sewer facilities
- util: other public buildings and structures
- pc:private capital stock
- gsp: gross state product
- emp: labor input measured by the employment in nonagricultural payrolls
- unemp: state unemployment rate

Specifications

	Within	Between	First Difference
log(pcap)	-0.03	0.18*	-0.01
	(0.03)	(0.07)	(0.05)
log(pc)	0.29***	0.30***	-0.03
	(0.03)	(0.04)	(0.02)
log(emp)	0.77***	0.58***	0.83***
	(0.03)	(0.06)	(0.04)
unemp	-0.01***	-0.00	-0.01^{***}
	(0.00)	(0.01)	(0.00)
(Intercept)		1.59***	0.01***
		(0.23)	(0.00)
R^2	0.94	0.99	0.69
Adj. R ²	0.94	0.99	0.69
Num. obs.	816	48	768

^{***}p < 0.001, **p < 0.01, *p < 0.05

Table: Statistical models

Linear Models Applications

Econometrics of Policy Evaluation

Applications

Nonlinear Pane

Statement

- Identify the causal effect of a policy
 - ► Minimum Wages on Employment
 - ► Training on Wages
 - Class size on Student Outcomes
 - Welfare on Labor Supply
- ► Essentially a self selection problem

Evaluation Problem

- Let y₁ denote the outcome with treatment
- \triangleright Let y_0 denote the outcome without treatment
- ▶ We are interested in the average treatment effect

$$ATE = E(y_1 - y_0) \tag{7}$$

- ▶ Evaluation problem: An individual can not be in both states, we can not observe both y_0 and y_1 .
- ► Another quantity is the average treatment effect of the treated (Let w be an indicator of treatment).

$$ATT = E(y_1 - y_0 \mid w = 1)$$
 (8)

Assumptions

Under which assumptions can you do a DiD?

- ➤ Stable Unit Treatment Value Assumption (SUTVA): potential outcomes for each person i are unrelated to the treatment status of other individuals
- Random Assignment. The treatment assignment is random i.e we have an independent, identically distributed sample from the population

Threat to Validity

- Pre-trend
- ► Placebo Effect
- ► Rubin Effect

DiD as a Linear Regression

Consider

$$Y_{it} = \alpha + \delta Post_t + \gamma D_i + \beta Post_t D_i + \epsilon_{it}$$
 (9)

Where

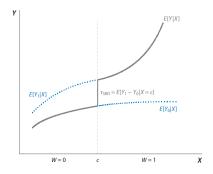
- \triangleright $D_i = 1$ if treated, 0 otherwise
- ▶ $Post_t = 1$ after the implementation of the policy

Alternative Method: Regression discontinuity

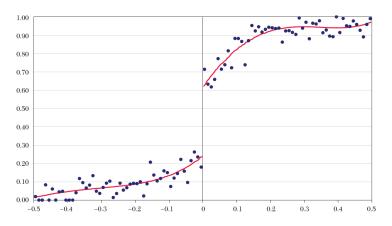
The Key intuition for RDD is that we have an understanding of the mechanism which underlies the assignment of treatment. Specifically, assignment to treatment depends on a single variable. In the sharp, regression discontinuity design, the running variable fully determines the treatment

$$D_{i} = \begin{cases} 1, & \text{If } X_{i} > X_{0} \\ 0, & \text{If } X_{i} < X_{0} \end{cases}$$
 (10)

Idea

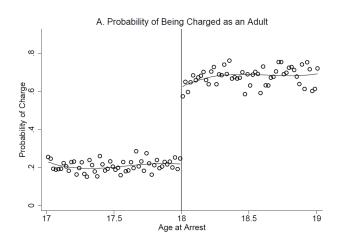


Some examples (1)

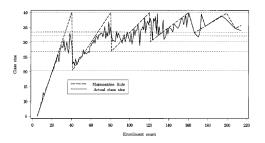


Figure~10.~Winning~the~Next~Election,~Bandwidth~of~0.01~(100~bins)

Some examples (2)



Some examples (3)



Explanation (1)

Twenty five children may be put in charge of one teacher. If the number in the class exceeds twenty five but is not more than forty, he should have an assistant to help with the instruction. If there are more than forty, two teachers must be appointed.

Explanation (2)

- ► A law prevents class size to exceed say 30
- ▶ If cohorts are of average size 90 but fluctuates
- ▶ If cohort size is 91-96, we end up four classrooms of size 22 to 24, while if cohort size is 85-90, we end up with three classrooms of size 28 to 30.

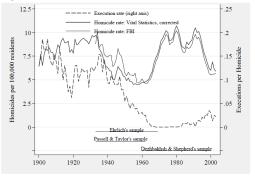
Angrist and Lavy (1999): Comparing test outcomes between students who are randomly assigned to the small vs large classes gives you a credible estimate of the effect of class size on academic performance. 10% decrease in class size increases test score by about 0.2 to 0.3 standard deviations.

Uses and Abuses of Empirical Evidence in the Death Penalty Debate

- Deterrance Effect of the death penalty
- ► What can we say?

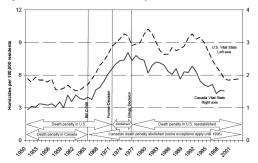
Evidence and Abuse (1)

Figure 1. Homicides and Execution in the United States



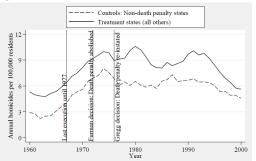
Evidence and Abuse (2)

Figure 2. Homicide Rates and the Death Penalty in the United States and Canada



Evidence and Abuse (3)

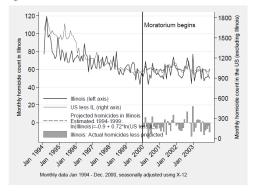
Figure 3. Homicide Rates in the United States



Non-death penalty states are those without a death penalty throughout 1960-2000: AK HI ME MI MN WI

Evidence and Abuse (4)

Figure 5. Homicides Before and After the Illinois Moratorium



Evidence and Abuse (5)

Table 6: Estimating the Impact of Executions on Murder Rates: Reanalyzing

Mocan and Gitti	ngs: 19//-	1997						
	4.0	unal Han		pendent V 100,000 Re		Log Hom	icida Pata	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	
				and Gittin		1.7	- 1.7	
Executions, per	-0.60			-0.63	-0.63	-0.05*	-0.05*	
Death Sentence.	(.35)			(0.34)	(.29)	(.03)	(.03)	
Pardons, per	(.55)	0.69**		0.73	(.23)	0.11	(.03)	
Death Sentence,		(.32)		(.30)		(.03)		
Death Row			0.17**	,	0.18**	,	0.02**	
Removals, per								
Death Sentence,			(.07)		(.07)		(0.01)	
Sample	680	693	695	679	690	679	690	
(1984-1997)	080	093	093	079	090	079	090	
	Panel B: Correcting Programming Errors							
Executions, per	-0.50			-0.52	-0.59	-0.01	-0.02	
Death Sentence,7	(.34)			(.33)	(0.39)	(0.03)	(0.02)	
Pardons, per		0.63		0.71**		0.09***		
Death Sentence ₁₋₇		(.34)		(.30)		(0.03)		
Death Row			0.24***		0.17		0.01	
Removals, per			(.08)		(0.09)		(0.01)	
Death Sentence, s Sample								
1984-1997)	679	692	691	677	636	677	636	
1501-1551)	Panel C: Measuring Deterrence Variables with a One-Year Lag on							
				Full San	ple			
executions, per	0.03			0.01	0.01	0.01	0.01	
Death Sentence,	(0.14)			0.13)	(0.14)	(0.01)	(0.01)	
Pardons, per		0.41***		0.41		0.05		
Death Sentence, 1		(.13)		(0.13)		(0.01)		
Jeath Row Removals., per			0.02		0.02		0.002	
Death Sentence,			(0.03)		(0.03)		(.002)	
Sample								
(1978-1997)	986	984	921	977	918	977	918	
		Impli	ed Life-	Life Trade	off for Exe	cutions ^(a)		
Panel A:			[959	6 Confidenc		2.2	2.2	
Replication	4.4 [-1.8, 10.5]			4.6 [-1.4, 10.6]	4.6	[-1.2, 5.7]	2.3 [-1.3, 6.0	
Panel B:	3.4			3.6	4.2	-0.2	0.5	
Corrected	[-2.6, 9.4]			[-2.2, 9.5]	[-2.6,11.1]	[-3.7, 3.4]	[-2.7, 3.7	
Panel C: Full	-1.2			-1.1	-1.1	-1.6	-1.6	
Sample	£3 1 0 71				L3 0 0 81			