

# Bayesian Statistics and Data Analysis Lecture 6

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### Section 1

### Introduction



# Extra material for dynamic HMC

• Michael Betancourt (2018). A Conceptual Introduction to Hamiltonian Monte Carlo.

https://arxiv.org/abs/1701.02434

 Cole C. Monnahan, James T. Thorson, and Trevor A. Branch (2016) Faster estimation of Bayesian models in ecology using Hamiltonian Monte Carlo.

https://dx.doi.org/10.1111/2041-210X.12681



#### Extra material for Stan

- Andrew Gelman, Daniel Lee, and Jiqiang Guo (2015)
   Stan: A probabilistic programming language for Bayesian inference and optimization.
  - http://www.stat.columbia.edu/~gelman/research/published/stan\_jebs\_2.pdf
- Carpenter et al (2017). Stan: A probabilistic programming language. Journal of Statistical Software 76(1).
   https://dox.doi.org/10.18637/jss.v076.i01
- Stan User's Guide, Language Reference Manual, and Language Function Reference (in html and pdf) https://mc-stan.org/users/documentation/
  - easiest to start from Example Models in User's guide
- Basics of Bayesian inference and Stan, part 1 Jonah Gabry & Lauren Kennedy (StanCon 2019 Helsinki tutorial)
  - https://www.youtube.com/watch?v=ZRpo41102KQ& index=6&list=PLuwyh42iHquU4hUBQs20hkBsKSMrp6H0J
  - https://www.youtube.com/watch?v=6cc4N1vT8pk& index=7&list=PLuwyh42iHquU4hUBQs20hkBsKSMrp6H0J



### Chapter 12 demos

- demo12 1: HMC
- http://elevanth.org/blog/2017/11/28/ build-a-better-markov-chain/
- rstan\_demo
- rstanarm\_demo
- http://sumsar.net/blog/2017/01/ bayesian-computation-with-stan-and-farmer-jons/
- http:
  - //mc-stan.org/documentation/case-studies.html
- https://cran.r-project.org/package=rstan
- https://cran.r-project.org/package=rstanarm



#### Hamiltonian Monte Carlo

- Uses log density (negative log density is called energy)
- Uses gradient of log density for more efficient sampling
- Augments parameter space with momentum variables



#### Hamiltonian Monte Carlo

- Uses log density (negative log density is called energy)
- Uses gradient of log density for more efficient sampling
- Augments parameter space with momentum variables
- Simulation of Hamiltonian dynamics reduces random walk
  - Explanation of HMC with black board
  - http://elevanth.org/blog/2017/11/28/ build-a-better-markov-chain/



### Hamiltonian Monte Carlo

- Uses gradient of log density for more efficient sampling
- Alternating dynamic simulation and sampling energy level
- Parameters: step size, number of steps in each chain
- No U-Turn Sampling (NUTS) and dynamic HMC
  - adaptively selects number of steps to improve robustness and efficiency
  - dynamic HMC refers to dynamic trajectory length
  - to keep reversibility of Markov chain, need to simulate in two directions
  - http://elevanth.org/blog/2017/11/28/ build-a-better-markov-chain/
- Dynamic simulation is discretized
  - small step size gives accurate simulation, but requires more log density evaluations
  - large step size reduces computation, but increases simulation error which needs to be taken into account in the Markov chain
  - black board explanation of the effect of step size



# Adaptive dynamic HMC in Stan

- Dynamic HMC using growing tree to increase simulation trajectory until no-U-turn criterion stopping
  - max treedepth to keep computation in control
  - pick a draw along the trajectory with probabilities adjusted to take into account the error in the discretized dynamic simulation
  - give bigger weight for tree parts further away to increase probability of jumping further away
- Mass matrix and step size adaptation in Stan
  - mass matrix refers to having different scaling for different parameters and optionally also rotation to reduce correlations
  - mass matrix and step size adjustment and are estimated during initial adaptation phase
  - step size is adjusted to be as big as possible while keeping discretization error in control (adapt\_delta)
- After adaptation the algorithm parameters are fixed
- After warmup store iterations for inference
- See more details in Stan reference manual



# Max tree depth diagnostic

- Dynamic HMC specific diagnostic
- Indicates inefficiency in sampling leading to higher autocorrelations and lower ESS  $(n_{\rm eff})$
- Different parameterizations matter



### Divergences

- HMC specific: indicates that Hamiltonian dynamic simulation has problems with unexpected fast changes in log-density
  - indicates possibility of biased estimates
- Different parameterizations matter
- http://mc-stan.org/users/documentation/ case-studies/divergences\_and\_bias.html



#### Problematic distributions

- Nonlinear dependencies
  - simple mass matrix scaling doesn't help
- Funnels
  - optimal step size depends on location
- Multimodal
  - difficult to move from one mode to another
- Long-tailed with non-finite variance and mean
  - efficiency of exploration is reduced
  - central limit theorem doesn't hold for mean and variance



# Probabilistic programming language

- Wikipedia "A probabilistic programming language (PPL) is a programming language designed to describe probabilistic models and then perform inference in those models"
- To make probabilistic programming useful
  - inference has to be as automatic as possible
  - diagnostics for telling if the automatic inference doesn't work
  - easy workflow (to reduce manual work)
  - fast enough (manual work replaced with automation)



# Probabilistic programming

- Enables agile workflow for developing probabilistic models
  - language
  - automated inference
  - diagnostics
- Many frameworks Stan, PyMC3, Pyro (Uber), Edward (Google), Birch, ELFI, ...



# Stan - probabilistic programming framework

- Language, inference engine, user interfaces, documentation, case studies, diagnostics, packages, ...
  - autodiff to compute gradients of the log density
- More than ten thousand users in social, biological, and physical sciences, medicine, engineering, and business
- Several full time developers, 40+ developers, more than 100 contributors
- R, Python, Julia, Scala, Stata, Matlab, command line interfaces
- More than 100 R packages using Stan







- Stanislaw Ulam (1909-1984)
  - Monte Carlo method
  - H-Bomb



```
data {
  int < lower = 0 > N; // number of experiments
  int<lower=0,upper=N> y; // number of successes
parameters {
  real < lower=0, upper=1> theta; // parameter of the
model {
  theta \sim beta (1,1); //prior
  y ~ binomial(N, theta); // observation model
```



- Data type and size are declared
- Stan checks that given data matches type and constraints
  - If you are not used to strong typing, this may feel annoying, but it will reduce the probability of coding errors, which will reduce probability of data analysis errors



```
parameters {
   real < lower=0, upper=1> theta;
}
```

- Parameters may have constraints
- Stan makes transformation to unconstrained space and samples in unconstrained space
  - e.g. log transformation for ¡lower=a¿
  - e.g. logit transformation for ¡lower=a,upper=b¿
- For these declared transformation Stan automatically takes into account the Jacobian of the transformation (see BDA3 p. 21)



```
model {
  theta ~ beta(1,1); // prior
  y ~ binomial(N,theta); // likelihood
}
```



- target is the log posterior density
- \_lpdf for continuous, \_lpmf for discrete distributions (discrete for the left hand side of |)
- for Stan sampler there is no difference between prior and likelihood, all that matters is the final target
- you can write in Stan language any program to compute the log density (Stan language is Turing complete)





- ullet Stan compiles (transplies) the model written in Stan language to C++
  - this makes the sampling for complex models and bigger data faster
  - also makes Stan models easily portable, you can use your own favorite interface

### **RStan**

```
RStan
```

```
library (rstan)
rstan_options (auto_write = TRUE)
options (mc.cores = parallel::detectCores())

d_bin <- list (N = 10, y = 7)
fit_bin <- stan (file = 'binom.stan', data = d_bin)</pre>
```



# PyStan

```
PyStan
import pystan
import stan_utility

data = dict(N=10, y=8)
model = stan_utility.compile_model('binom.stan')
fit = model.sampling(data=data)
```



### Stan

- Compilation (unless previously compiled model available)
- Warm-up including adaptation
- Sampling
- Generated quantities
- Save posterior draws
- Report divergences,  $n_{E_{\text{max}}}$ ,  $\hat{R}$



### Difference between proportions

- An experiment was performed to estimate the effect of beta-blockers on mortality of cardiac patients
- A group of patients were randomly assigned to treatment and control groups:
  - out of 674 patients receiving the control, 39 died
  - out of 680 receiving the treatment, 22 died



data {

int < lower = 0 > N1;

Introduction

# Difference between proportions

```
int < lower = 0 > v1:
  int < lower = 0 > N2;
  int < lower = 0 > y2;
parameters {
  real < lower=0, upper=1> theta1;
  real<lower=0,upper=1> theta2;
model {
  theta1 \sim beta(1,1);
  theta2 ~ beta(1,1);
  y1 ~ binomial(N1, theta1);
  y2 ~ binomial(N2, theta2);
generated quantities {
  real oddsratio:
  oddsratio = \frac{(theta2/(1-theta2))}{(theta1/(1-theta2))}
```



# Difference between proportions

```
generated quantities {
  real oddsratio;
  oddsratio = (theta2/(1-theta2))/(theta1/(1-theta2))
}
```

• generated quantities is run after the sampling



#### uppsala universitet

Introduction

# Difference between proportions

```
starting worker pid=10151 on localhost:11783 at 10:03:27.872
starting worker pid=10164 on localhost:11783 at 10:03:28.087
starting worker pid=10176 on localhost:11783 at 10:03:28.295
starting worker pid=10185 on localhost:11783 at 10:03:28.461
SAMPLING FOR MODEL 'binom2' NOW (CHAIN 1).
Gradient evaluation took 6e-06 seconds
1000 transitions using 10 leapfrog steps per transition would take 0.06 seconds.
Adjust your expectations accordingly!
Iteration: 1 / 2000 [
                               (Warmup)
Iteration: 200 / 2000 [ 10%]
                               (Warmup)
Iteration: 1000 / 2000 [
                         50%1
                               (Warmup)
Iteration: 1001 / 2000 [ 50%]
                               (Sampling)
Iteration: 2000 / 2000 [100%] (Sampling)
Elapsed Time: 0.012908 seconds (Warm-up)
               0.017027 seconds (Sampling)
               0.029935 seconds (Total)
```

SAMPLING FOR MODEL 'binom2' NOW (CHAIN 2). . . .



# Difference between proportions

monitor(fit\_bin2, probs = c(0.1, 0.5, 0.9))

Inference for the input samples
(4 chains: each with iter=1000; warmup=0):

	mean	se_mean	sd	10%	50%	90%	n_eff	Rhat
theta1	0.1	0	0.0	0.0	0.1	0.1	3280	1
theta2	0.0	0	0.0	0.0	0.0	0.0	3171	1
oddsratio	0.6	0	0.2	0.4	0.6	0.8	3108	1
l p	-253.5	0	1.0	-254.8	-253.2	-252.6	1922	1

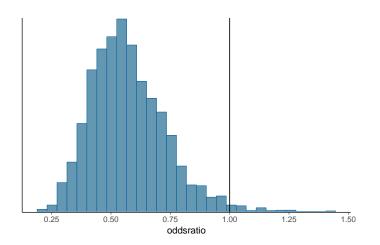
For each parameter,  $n_{-}$ eff is a crude measure of effective samp and Rhat is the potential scale reduction factor on split chair convergence, Rhat=1).

• lp\_\_ is the log density, ie, same as target



# Difference between proportions

```
\begin{array}{lll} draws < & - as.data.frame(fit\_bin2) \\ mcmc\_hist(draws, pars = 'oddsratio') + \\ geom\_vline(xintercept = 1) + \\ scale\_x\_continuous(breaks = c(seq(0.25,1.5,by=0.00))) \end{array}
```





# HMC specific diagnostics

check\_treedepth (fit\_bin2)

```
check_energy(fit_bin2)
check_div(fit_bin2)
```

```
[1] "0 of 4000 iterations saturated the maximum tree depth of [1] "0 of 4000 iterations ended with a divergence (0\%)"
```



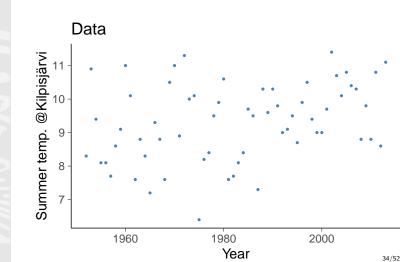
# Shinystan

• Graphical user interface for analysing MCMC results



# Kilpisjärvi summer temperature

- Temperature at Kilpisjärvi in June, July and August from 1952 to 2013
- Is there change in the temperature?





#### Gaussian linear model

```
data {
    int<lower=0> N; // number of data points
    vector[N] x; //
    vector[N] y; //
parameters {
    real alpha;
    real beta:
    real < lower = 0 > sigma;
transformed parameters {
    vector[N] mu;
    mu <- alpha + beta*x;
model {
    y ~ normal(mu, sigma);
```



#### Gaussian linear model

```
data {
    int < lower = 0 > N; // number of data points
    vector[N] x; //
    vector[N] y; //
}
```

• difference between vector[N] x and real x[N]



## Gaussian linear model

```
parameters {
    real alpha;
    real beta;
    real<lower=0> sigma;
}
transformed parameters {
    vector[N] mu;
    mu <- alpha + beta*x;
}</pre>
```

 transformed parameters are deterministic transformations of parameters and data



## Priors for Gaussian linear model

```
data {
    int < lower = 0 > N; // number of data points
    vector[N] x; //
    vector[N] y; //
    real pmualpha; // prior mean for alpha
    real psalpha; // prior std for alpha
    real pmubeta; // prior mean for beta
    real psbeta; // prior std for beta
transformed parameters {
    vector[N] mu;
    mu \leftarrow alpha + beta*x;
model {
    alpha ~ normal(pmualpha, psalpha);
    beta ~ normal(pmubeta, psbeta);
    y ~ normal(mu, sigma);
```



## Student-t linear model

parameters {
 real alpha;
 real beta;
 real <lower=0> sigma;
 real <lower=1, upper=80> nu;
}
transformed parameters {
 vector[N] mu;
 mu <- alpha + beta\*x;
}
model {
 nu ~ gamma(2,0.1);</pre>

y ~ student\_t(nu, mu, sigma);



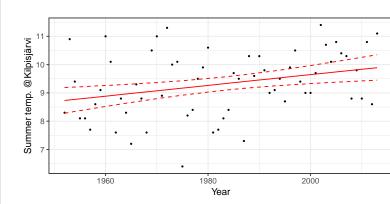
## Priors

• Prior for temperature increase?



# Kilpisjärvi summer temperature

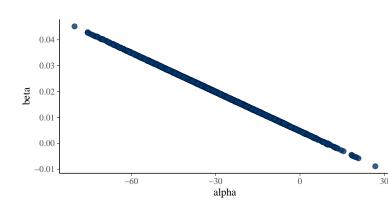
#### Posterior fit





# Kilpisjärvi summer temperature

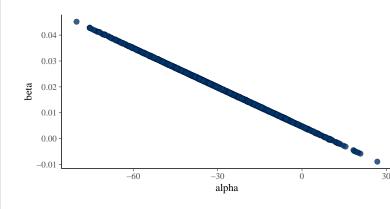
Posterior draws of alpha and beta





## Kilpisjärvi summer temperature

Posterior draws of alpha and beta



There were 14 transitions after warmup that exceeded the maximular treedepth. Increase <code>max\_treedepth</code> above 10. See <code>http://mc-stan.org/misc/warnings.html#maximum-treedepth-exceedeft Examine the pairs() plot to diagnose sampling problems</code>



## Linear regression model in Stan

```
data {
  int<lower=0> N; // number of data points
  vector[N] x; //
  vector[N] y; //
  real xpred; // input location for prediction
}
transformed data {
  vector[N] x.std;
  vector[N] y.std;
  real xpred_std;
  x.std = (x - mean(x)) / sd(x);
  y.std = (y - mean(y)) / sd(y);
  xpred_std = (xpred - mean(x)) / sd(x);
}
```



## **RStanARM**

- RStanARM provides simplified model description with pre-compiled models
  - no need to wait for compilation
  - a restricted set of models

#### Two group Binomial model:

```
d_bin2 <- data frame(N = c(674, 680), y = c(39,22), grp2 = c(0 fit_bin2 <- stan_glm(y/N ~ grp2, family = binomial(), data = d weights = N)
```



### **RStanARM**

- RStanARM provides simplified model description with pre-compiled models
  - no need to wait for compilation
  - a restricted set of models

#### Two group Binomial model:

```
d_bin2 <- data frame(N = c(674, 680), y = c(39,22), grp2 = c(0 fit_bin2 <- stan_glm(y/N ~ grp2, family = binomial(), data = d weights = N)
```

#### Gaussian linear model

```
fit_lin <- stan_glm(temp ~ year, data = d_lin)
```



### BRMS

- BRMS provides simplified model description
  - a larger set of models than RStanARM, but still restricted
  - need to wait for the compilation

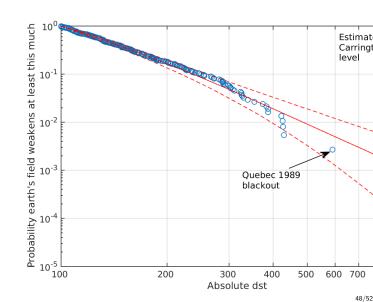
```
 \begin{array}{lll} fit\_bin2 < & brm(y/N \ \ ^{\circ} \ grp2 \,, \ family = binomial() \,, \ data = d\_bin2 \\ & weights = N) \end{array}
```

```
fit_lin_t \leftarrow brm(temp ~ year, data = d_lin, family = student())
```



## Extreme value analysis

### Geomagnetic storms





# Extreme value analysis

```
data {
  int < lower = 0 > N:
  vector<lower=0>[N] y;
  int <lower=0> Nt;
  vector<lower=0>[Nt] yt;
transformed data {
  real ymax;
  ymax < - max(y);
parameters {
  real < lower=0> sigma;
  real < lower = - sigma / ymax > k;
model {
  y ~ gpareto(k, sigma);
generated quantities {
  vector[Nt] predccdf;
  predccdf<-gpareto_ccdf(yt,k,sigma);</pre>
```



### Functions

```
functions {
  real gpareto_lpdf(vector y, real k, real sigma) {
    // generalised Pareto log pdf with mu=0
    // should check and give error if k<0
    // and max(y)/sigma > -1/k
    int N:
   N \leftarrow dims(y)[1];
    if (fabs(k) > 1e-15)
      return -(1+1/k)*sum(log1pv(y*k/sigma)) -N*log(sigma)
    else
      return -sum(y/sigma) -N*log(sigma); // limit k->0
  vector gpareto_ccdf(vector y, real k, real sigma) {
    // generalised Pareto log ccdf with mu=0
    // should check and give error if k<0
    // and max(y)/sigma < -1/k
    if (fabs(k) > 1e-15)
      return exp((-1/k)*log1pv(y/sigma*k));
    else
      return \exp(-y/\text{sigma}); // limit k = > 0
```



## Other packages

- R
- shinystan interactive diagnostics
- bayesplot visualization and model checking (see model checking in Ch 6)
- loo cross-validation model assessment, comparison and averaging (see Ch 7)
- projpred projection predictive variable selection
- Python
  - ArviZ visualization, and model checking and assessment (see Ch 6 and 7)



### Different interfaces

- RStan / PyStan
  - C++ functions of Stan are called directly from R / Python
  - Higher integration between R/Python and Stan, but maybe more difficult to install due to more requirements of compatible C++ compilers and libraries
- CmdStanR / CmdStanPy
  - Lightweight interface on top of commandline program CmdStan
  - Lacks some features that are not needed in this course, but is usually easier to install
- More recent useful R packages
  - posterior: for handling posterior draws, convergence diagnostics, and summaries
  - tidybayes + ggdist: pretty plots