

# metaSEM: An R Package for Meta-Analysis Using Structural Equation Modeling

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## Abstract

The **metaSEM** package provides functions to conducting univariate and multivariate meta-analysis using a structural equation modeling approach via the **OpenMx** package. It also implemented the two-stage structural equation modeling (TSSEM) approach to conducting fixed- and random-effects meta-analytic structural equation modeling (MASEM) on correlation/covariance matrices. This paper outlines the basic theories. Examples are used to illustrate the procedures.

*Keywords:* meta-analysis, structural equation modeling, meta-analytic structural equation modeling, **metaSEM**, R.

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## 1. Introduction

Meta-analysis is a popular technique to synthesizing research findings in social, behavioral, educational and medical sciences (Borenstein, Hedges, Higgins, and Rothstein 2009; Hedges and Olkin 1985; Hunter and Schmidt 2004; Whitehead 2002). There are several standalone packages for meta-analysis, e.g., Comprehensive Meta-Analysis and RevMan. Many standard statistical packages, for instance, SPSS (Lipsey and Wilson 2000), SAS (Arthur, Bennett, and Huffcutt 2001) and STATA (Sterne 2009), have macros or packages to fitting some meta-analytic models. Even in the R community, there are already several packages to conducting meta-analysis, for instance, **meta** (Schwarzer 2010), **rmeta** (Lumley 2009), **mvmeta** (Gasparini 2012), **metaLik** (Guolo and Varin 2011) and **metafor** (Viechtbauer 2010).

The **metaSEM** package is yet another R package to conducting univariate and multivariate meta-analysis. It formulates meta-analytic models as structural equation models (Cheung 2008, 2011b) via the **OpenMx** package (Boker, Neale, Maes, Wilde, Spiegel, Brick, Spies, Estabrook, Kenny, Bates, Mehta, and Fox 2011). It also implemented the two-stage structural equation modeling (TSSEM) approach (Cheung and Chan 2005b, 2009) to conducting fixed- and random-effects meta-analytic structural equation modeling (MASEM) on correlation/covariance matrices. The main functions in this package are:

- **meta()** and **reml()**: **meta()** fits univariate and multivariate meta-analysis with maximum likelihood (ML) estimation method while **reml()** estimates the variance components of the random-effects with restricted (residual) maximum likelihood (REML)

estimation method. Mixed-effects meta-analysis can be fitted by including study characteristics as predictors. Equality constraints on the intercepts, regression coefficients and variance components can be imposed.

- **meta3()** and **rem13()**: They fit 3-level meta-analysis by including clustering effect. **meta3()** fits the 3-level meta-analysis with ML estimation method while **rem13()** estimates the variance components with REML estimation method.
- **tssem1()**: It fits the first stage analysis of TSSEM by pooling correlation/covariance matrices with either a fixed- or random-effects model.
- **tssem2()**: It fits the second stage analysis of TSSEM by fitting structural models on the pooled correlation/covariance matrix. It is a wrapper of **wls()**.
- **wls()**: It fits a correlation/covariance structure analysis with weighted least squares (WLS) estimation method.

Besides reporting Wald confidence intervals (CIs) based on z statistic, likelihood-based CIs on the parameter estimates may also be requested (Cheung 2009a; Neale and Miller 1997). Several generic functions, such as **anova()**, **coef()**, **vcov()**, **print()**, **summary()** and **plot()**, have been implemented.

This paper was based on the **metaSEM** package version 0.7-1 and the **OpenMx** package version 1.2.4-2063. The paper is organized as follows. The next section presents basic ideas on structural equation models and how they are linked to meta-analytic models. Basic theory on the TSSEM are then presented. Several examples are used to illustrate these procedures.

## 2. Structural Equation Modeling Based Meta-Analysis

In this section, basic structural equation models are introduced. Univariate and multivariate meta-analysis are treated as special cases of SEM (Cheung 2008, 2011b).

### 2.1. Structural equation model

Structural equation modeling is a multivariate technique to fitting and testing hypothesized models. Let  $\mathbf{y}$  be a  $p \times 1$  vector of the sample data where  $p$  is the number of variables. It is hypothesized that the model for the first and second moments are  $\mu = \mu(\theta)$  and  $\Sigma = \Sigma(\theta)$ , respectively, where  $\theta$  is a vector of parameters.

The  $-2 \times \log$ -likelihood of the  $i$ th case is:

$$-2 * \log L_i(\theta; \mathbf{y}_i)_{\text{ML}} = p_i * \log(2\pi) + \log |\Sigma_i(\theta)| + (\mathbf{y}_i - \mu_i(\theta))' \Sigma_i(\theta)^{-1} (\mathbf{y}_i - \mu_i(\theta)), \quad (1)$$

where  $p_i$  is the number of variables in the  $i$ th case,  $\mu_i(\theta)$  and  $\Sigma_i(\theta)$  are the model implied mean vector and the model implied covariance matrix for the  $i$ th case, respectively. Since there is a subscript  $i$  in these quantities, the model implied mean vector and covariance matrix may vary across cases. In other words, this model handles incomplete data automatically by selecting only the complete variables in the log-likelihood function.

To obtain the parameter estimates, we may take the sum of the  $-2 \times \log$ -likelihood over all cases and minimize it. This is known as the full information maximum likelihood (FIML or simply

ML) estimation method. Iterative methods are usually used to obtain the solutions. If it is convergent, the asymptotic covariance matrix (thus the standard errors) of the parameter estimates may be obtained from the inverse of the Hessian matrix. The parameter estimates divided by their standard errors follow a  $z$  distribution under the null hypothesis. Moreover, likelihood ratio statistic may also be used to compare nested models.

## 2.2. Univariate fixed-effects model

When there is only one effect size, the univariate fixed-effects model for the  $i$ th study is:

$$y_i = \beta_{\text{fixed}} + e_i, \quad (2)$$

where  $\beta_{\text{fixed}}$  is the common effect under a fixed-effects model and  $\text{var}(e_i) = v_i$  is the known sampling variance. To fit the univariate fixed-effects meta-analysis in SEM, we may use the following model:

$$\mu_i(\theta) = \beta_{\text{fixed}} \quad (3)$$

and

$$\Sigma_i(\theta) = v_i \quad (4)$$

Since  $v_i$  is known, the only parameter in the univariate fixed-effects model is  $\beta_{\text{fixed}}$ .

## 2.3. Univariate random-effects model

A random-effects model allows studies to have their own study specific effect. The model for the  $i$ th study is:

$$y_i = \beta_{\text{random}} + u_i + e_i. \quad (5)$$

where  $\beta_{\text{random}}$  is the average effect under a random-effects model and  $\text{var}(u_i) = \tau^2$  is the heterogeneity variance that has to be estimated. To fit the univariate fixed-effects meta-analysis in SEM, we may use the following model:

$$\mu_i(\theta) = \beta_{\text{random}} \quad (6)$$

and

$$\Sigma_i(\theta) = \tau^2 + v_i \quad (7)$$

In this model we have to estimate both  $\beta_{\text{random}}$  and  $\tau^2$ .

## 2.4. Univariate mixed-effects model

The mixed-effects meta-analysis extends the random-effects meta-analysis by including study characteristics as predictors. Assuming that  $\mathbf{x}_i$  is a  $m \times 1$  vector of predictors where  $m$  is the number predictors in the  $i$ th study, the model is:

$$y_i = \beta_0 + \beta' \mathbf{x}_i + u_i + e_i, \quad (8)$$

where  $\beta$  is a vector of regression coefficients.

To fit the univariate mixed-effects meta-analysis in SEM, we may use the following model for the conditional mean and variance:

$$\mu_i(\theta|\mathbf{x}_i) = \beta_0 + \beta' \mathbf{x}_i \quad (9)$$

and

$$\Sigma_i(\theta|\mathbf{x}_i) = \tau^2 + v_i. \quad (10)$$

Since  $\mathbf{x}_i$  is specified via definition variables (Boker *et al.* 2011), the means and covariance matrix of  $\mathbf{x}$  are not estimated. That is,  $\mathbf{x}$  is treated as a design matrix rather than random variables.

## 2.5. Multivariate mixed-effects model

Let us assume that there are  $p$  effect sizes with  $m$  predictors in  $k$  studies. The model for the multivariate effect sizes in the  $i$ th study is:

$$\mathbf{y}_i = \mathbf{B}\mathbf{x}_i + \mathbf{u}_i + \mathbf{e}_i, \quad (11)$$

where  $\mathbf{y}_i$  is a  $p \times 1$  effect sizes,  $\mathbf{B}$  is a  $p \times (m+1)$  regression coefficients including the intercepts,  $\mathbf{x}_i$  is a  $(m+1) \times 1$  predictors including 1 in the first column,  $\mathbf{u}_i$  is a  $p \times 1$  study specific random effects, and  $\mathbf{e}_i$  is a  $p \times 1$  sampling error. We assume that  $\text{var}(\mathbf{e}_i) = V_i$  is known and given in the  $i$ th study and  $\text{var}(\mathbf{u}_i) = T^2$  is the variance component of the between-study heterogeneity that has to be estimated.

The -2\*log-likelihood of the above model is:

$$-2 * \log L_i(\mathbf{B}, T^2; \mathbf{y}_i)_{\text{ML}} = p_i * \log(2\pi) + \log|T^2 + V_i| + (\mathbf{y}_i - \mathbf{B}\mathbf{x}_i)'(T^2 + V_i)^{-1}(\mathbf{y}_i - \mathbf{B}\mathbf{x}_i), \quad (12)$$

where  $p_i$  is the number of effect sizes in the  $i$ th study.

In applied research, different studies may report different effect sizes, that is,  $p_i$  may vary across studies. The above -2\*log-likelihood may handle missing effect sizes by using different dimensions of the elements in the above equation. It is expected that there is no missing data in  $\mathbf{x}_i$ . When there are missing data in  $\mathbf{x}_i$ , the whole study will be deleted before the analysis.

## 2.6. Restricted Maximum Likelihood (REML) Estimation Method

Since both the fixed- and random-effects are estimated simultaneously, it is well-known that  $\hat{T}_{\text{ML}}^2$  based on the ML estimation is negatively biased. It is because it does not take the uncertainty in estimating  $\hat{\mathbf{B}}_{\text{ML}}$  into account. If the unbiasedness of the variance component is crucial in the research questions, we may obtain the variance component  $\hat{T}_{\text{REML}}^2$  based on the REML estimation method (Cheung 2011a; Harville 1977; Patterson and Thompson 1971).

The -2log-likelihood of the model is:

$$-2 \log L_i(T^2; \mathbf{y}_i)_{\text{REML}} = p_i * \log(2\pi) + \log|T^2 + V_i| + (\mathbf{y}_i - \alpha \mathbf{X}_i)'(T^2 + V_i)^{-1}(\mathbf{y}_i - \alpha \mathbf{X}_i) + |\mathbf{X}_i' V_i^{-1} \mathbf{X}_i|, \quad (13)$$

where  $\alpha = (X'V^{-1}X)^{-1}X'V^{-1}\mathbf{y}$ .

Since the fixed effects  $\mathbf{B}$  is not involved in the above -2log-likelihood function, it has to be calculated in a second stage.

## 2.7. Three-level meta-analysis

Observed effect sizes may be related or dependent. For example, effect sizes reported by the same research team may be more similar when comparing to effect sizes reported by other

research teams. Effect sizes reported by studies from the same country may be more similar when comparing to studies across countries. If the degree of dependence is known, multivariate meta-analysis as introduced before may be applied. When the degree of dependence is unknown, a three-level meta-analytic model may be used (Konstantopoulos 2011). The model is:

$$y_i = \beta_0 + \beta' \mathbf{x}_i + u_{(2)i} + u_{(3)i} + e_i, \quad (14)$$

where  $u_{(2)i}$  and  $u_{(3)i}$  are the random-effects at level-2 and level-3, respectively.

To fit the three-level meta-analytic model in SEM, we may use the following model for the conditional mean and variance:

$$\mu_i(\theta|\mathbf{x}_i) = \beta_0 + \beta' \mathbf{x}_i \quad (15)$$

and

$$\Sigma_i(\theta|\mathbf{x}_i) = \tau_{(2)}^2 + \tau_{(3)}^2 + v_i. \quad (16)$$

where  $\tau_{(2)}^2 = \text{var}(u_{(2)i})$  and  $\tau_{(3)}^2 = \text{var}(u_{(3)i})$  are the heterogeneity at level-2 and level-3, respectively.

## 2.8. Examples

Two example data sets are used to demonstrate the procedures of fitting univariate and multivariate meta-analyses. The first data set was taken from Becker (1983) who reported 10 studies on sex differences in conformity using the fictitious norm group paradigm.  $d_i$  and  $v_i$  are the standardized mean difference and its sampling variance, respectively. Becker hypothesized that the logarithm of the number of items (*items*) predicted the effect size.

The second data set is adapted from Berkey, Hoaglin, Antczak-Bouckoms, Mosteller, and Colditz (1998). They summarized five published trials comparing surgical and non-surgical treatments for medium-severity periodontal disease, one year after treatment. Publication year *pub\_year* was hypothesized as a predictor.

**Univariate random-effects model** The function `meta()` is used to conduct the analyses. The arguments `y` and `v` are used to specify the effect sizes and its sampling variances (and covariances for multivariate meta-analysis), respectively. By default, a random-effects meta-analysis is fitted. After running the analysis, `summary()` is used to extract the results. The estimated fixed- and random-effects are represented by the **Intercept** and **Tau2** parameters. `coef()` and `vcov()` may be used to extract the parameter estimates and their asymptotic sampling covariance matrix, respectively.

From the following analyses, the  $Q$  statistic ( $df = 9$ ) is 30.6495,  $p < .001$  and the  $I^2$  based on the  $Q$  statistic is .6718. The pooled effect size with its 95% Wald confidence interval (CI) based on the random-effects model is 0.1747 (-0.0475, 0.3970). The estimated heterogeneity variance is 0.0774.

```
R> ## Load the library
R> library(metaSEM)
R> ## Show the first few studies of the data set
R> head(Becker83)
```

	study	di	vi	percentage	items
1	1	-0.33	0.03	25	2
2	2	0.07	0.03	25	2
3	3	-0.30	0.02	50	2
4	4	0.35	0.02	100	38
5	5	0.69	0.07	100	30
6	6	0.81	0.22	100	45

```
R> ## Random-effects meta-analysis with ML
R> summary( random1 <- meta(y=di, v=vi, data=Becker83) )
```

Running Meta analysis with ML

```
Call:
meta(y = di, v = vi, data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound	z value
Intercept1	0.174734	0.113378	-0.047482	0.396950	1.5412
Tau2_1_1	0.077376	0.054108	-0.028674	0.183426	1.4300

Pr(>|z|)

Intercept1	0.1233
Tau2_1_1	0.1527

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Heterogeneity indices (based on the estimated Tau2):

	Estimate
Intercept1: I2 (Q statistic)	0.6718

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 2

Degrees of freedom: 8

-2 log likelihood: 7.928307

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Exact the coefficients
```

```
R> coef(random1)
```

Intercept1	Tau2_1_1
0.17473402	0.07737594

```
R> ## Exact the sampling variance covariance matrix
```

```
R> vcov(random1)
```

```

      Intercept1      Tau2_1_1
Intercept1 0.012854471 0.001240975
Tau2_1_1    0.001240975 0.002927667

```

**Univariate mixed-effects model** We may include a predictor to conduct a mixed-effects meta-analysis. The argument `x` is used to specify the predictors. If there are more than one predictor, `cbind()` may be used to specify them. The estimated regression coefficients are represented by the **Slope** parameter. The following analysis suggests that  $\log(items)$  is a significant predictor with the estimated regression coefficient and its 95% CI of 0.2109 (0.1225, 2.9924) with  $R^2 = 1$ .

```

R> ## Mixed-effects meta-analysis with "log(items)" as the predictor
R> summary( mixed1 <- meta(y=di, v=vi, x=log(items), data=Becker83) )

```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, x = log(items), data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound
Intercept1	-3.2015e-01	1.0981e-01	-5.3539e-01	-1.0492e-01
Slope1_1	2.1088e-01	4.5084e-02	1.2251e-01	2.9924e-01
Tau2_1_1	1.0000e-10	2.0095e-02	-3.9386e-02	3.9386e-02

	z value	Pr(> z )
Intercept1	-2.9154	0.003552 **
Slope1_1	4.6774	2.905e-06 ***
Tau2_1_1	0.0000	1.000000

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Explained variances (R2):

	Estimate
y1: Tau2 (no predictor)	0.0774
y1: Tau2 (with predictors)	0.0000
y1: R2	1.0000

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 3

Degrees of freedom: 7

-2 log likelihood: -4.208024

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Exact the coefficients
```

```
R> coef(mixed1)
```

```
      Intercept1      Slope1_1      Tau2_1_1
-0.3201549197  0.2108782268  0.0000000001
```

```
R> ## Exact the sampling variance covariance matrix
```

```
R> vcov(mixed1)
```

```
      Intercept1      Slope1_1      Tau2_1_1
Intercept1  0.0120593278 -3.940014e-03 3.887682e-04
Slope1_1   -0.0039400143  2.032587e-03 8.408773e-05
Tau2_1_1    0.0003887682  8.408773e-05 4.038148e-04
```

**Univariate fixed-effects model** Mathematically, fixed-effects meta-analysis is a special case of the random-effects meta-analysis by fixing the variance of the random-effects at 0. The argument `RE.constraints`, which expects a matrix as input, is used to constrain the variance component of the random effects.

```
R> ## Fixed-effects meta-analysis
```

```
R> summary( fixed1 <- meta(y=di, v=vi, data=Becker83,
                          RE.constraints=matrix(0, ncol=1, nrow=1)) )
```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, data = Becker83, RE.constraints = matrix(0,
  ncol = 1, nrow = 1))
```

95% confidence intervals: z statistic approximation

Coefficients:

```
      Estimate Std.Error    lbound    ubound z value
Intercept1  0.100640  0.060510 -0.017957  0.219237  1.6632
```

```
Pr(>|z|)
```

```
Intercept1  0.09627 .
```

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Heterogeneity indices (based on the estimated Tau2):

```
      Estimate
Intercept1: I2 (Q statistic)      0
```

Number of studies (or clusters): 10



```

Number of observed statistics: 10
Number of estimated parameters: 1
Degrees of freedom: 9
-2 log likelihood: 17.86043
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

```

R> ## Exact the coefficients
R> coef(fixed1)

```

```

Intercept1
0.1006404

```

```

R> ## Exact the sampling variance covariance matrix
R> vcov(fixed1)

```

```

Intercept1
Intercept1 0.003661436

```

**Multivariate random-effects model** Multivariate meta-analysis can be fitted by specifying the multivariate effect sizes and its sampling covariance matrix in the arguments `y` and `v` with `cbind()`, respectively. Only the lower triangle of the sampling covariance matrix arranged by the column major is used in `v`.

The  $Q$  statistic ( $df = 8$ ) of the following example is 128.2267,  $p < .001$ . The pooled effect sizes with their 95% Wald CIs based on the random-effects model for *PD* and *AL* are 0.3448 (0.2397, 0.4500) and -0.3379 (-0.4972, -0.1787), respectively. The estimated variance component is  $\begin{bmatrix} 0.0070 & \\ 0.0095 & 0.02614 \end{bmatrix}$ . The  $I^2$  based on the  $Q$  statistic for *PD* and *AL* are .6021 and .9250, respectively.

```

R> ## Show the data set
R> Berkey98

```

	trial	pub_year	no_of_patients	PD	AL	var_PD	cov_PD_AL
1	1	1983	14	0.47	-0.32	0.0075	0.0030
2	2	1982	15	0.20	-0.60	0.0057	0.0009
3	3	1979	78	0.40	-0.12	0.0021	0.0007
4	4	1987	89	0.26	-0.31	0.0029	0.0009
5	5	1988	16	0.56	-0.39	0.0148	0.0072

	var_AL
1	0.0077
2	0.0008
3	0.0014
4	0.0015
5	0.0304

```
R> ## Multivariate meta-analysis with a random-effects model
R> summary( mult1 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL),
                        data=Berkey98) )
```

Running Meta analysis with ML

Call:

```
meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
     data = Berkey98)
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound
Intercept1	0.3448392	0.0536312	0.2397239	0.4499544
Intercept2	-0.3379381	0.0812480	-0.4971812	-0.1786951
Tau2_1_1	0.0070020	0.0090497	-0.0107351	0.0247391
Tau2_2_1	0.0094607	0.0099698	-0.0100797	0.0290010
Tau2_2_2	0.0261445	0.0177409	-0.0086270	0.0609161

	z value	Pr(> z )
Intercept1	6.4298	1.278e-10 ***
Intercept2	-4.1593	3.192e-05 ***
Tau2_1_1	0.7737	0.4391
Tau2_2_1	0.9489	0.3427
Tau2_2_2	1.4737	0.1406

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 128.2267

Degrees of freedom of the Q statistic: 8

P value of the Q statistic: 0

Heterogeneity indices (based on the estimated Tau2):

	Estimate
Intercept1: I2 (Q statistic)	0.6021
Intercept2: I2 (Q statistic)	0.9250

Number of studies (or clusters): 5

Number of observed statistics: 10

Number of estimated parameters: 5

Degrees of freedom: 5

-2 log likelihood: -11.68131

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Exact the coefficients
```

```
R> coef(mult1)
```

Intercept1	Intercept2	Tau2_1_1	Tau2_2_1
0.344839167	-0.337938117	0.007001998	0.009460665

```

    Tau2_2_2
0.026144517

```

```

R> ## Exact the sampling variance covariance matrix
R> vcov(mult1)

```

```

              Intercept1  Intercept2  Tau2_1_1
Intercept1  2.876307e-03  2.215623e-03  1.241471e-04
Intercept2  2.215623e-03  6.601230e-03  6.203168e-05
Tau2_1_1    1.241471e-04  6.203168e-05  8.189758e-05
Tau2_2_1    -1.684464e-05  1.220432e-04  5.825516e-05
Tau2_2_2    1.315804e-06  3.246241e-05  4.503124e-05
              Tau2_2_1    Tau2_2_2
Intercept1 -1.684464e-05  1.315804e-06
Intercept2  1.220432e-04  3.246241e-05
Tau2_1_1    5.825516e-05  4.503124e-05
Tau2_2_1    9.939612e-05  1.187750e-04
Tau2_2_2    1.187750e-04  3.147402e-04

```

**Multivariate mixed-effects model** As an illustration, we use *pub\_year* as a predictor. To make the intercept more interpretable, we center the publication year at 1979, the first record of publication year in the data set. The estimated regression coefficients and their 95% CIs on *PD* and *AL* are 0.0064 (-0.2048, 0.2177) and -0.0706 (-0.3883, 0.2471), respectively. The  $R^2$  in predicting *PD* and *AL* are .0000 and .0433, respectively. The likelihood ratio statistic on testing both regression coefficient is  $\chi^2(df = 2) = 0.3273, p = .8490$ . Thus, both regression coefficients are non-significant.

Sometimes, we may want to test the equality of the regression coefficients and see if they are different. We may impose the equality constraint on the regression coefficients with the argument `coef.constraints`. The average regression coefficient is 0.0017 (-0.1991, 0.2025). The likelihood ratio statistic on testing the equality of the regression coefficients is  $\chi^2(df = 1) = 0.3270, p = .5674$ . There is no evidence that one regression coefficient is stronger from the other.

```

R> ## Multivariate meta-analysis with "publication year-1979" as a predictor
R> mult2 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL), data=Berkey98,
               x=scale(pub_year, center=1979), model.name="No constraint")

```

Running No constraint

```

R> summary(mult2)

```

Call:

```

meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
     x = scale(pub_year, center = 1979), data = Berkey98, model.name = "No constraint")

```

95% confidence intervals: z statistic approximation

## Coefficients:

	Estimate	Std.Error	lbound	ubound
Intercept1	0.3440001	0.0857659	0.1759020	0.5120982
Slope1_1	0.0063540	0.1078235	-0.2049761	0.2176842
Intercept2	-0.2918175	0.1312797	-0.5491208	-0.0345141
Slope2_1	-0.0705887	0.1620966	-0.3882922	0.2471147
Tau2_1_1	0.0080405	0.0101206	-0.0117955	0.0278766
Tau2_2_1	0.0093413	0.0105515	-0.0113392	0.0300218
Tau2_2_2	0.0250135	0.0170788	-0.0084603	0.0584873

	z value	Pr(> z )
Intercept1	4.0109	6.048e-05 ***
Slope1_1	0.0589	0.95301
Intercept2	-2.2229	0.02622 *
Slope2_1	-0.4355	0.66322
Tau2_1_1	0.7945	0.42692
Tau2_2_1	0.8853	0.37599
Tau2_2_2	1.4646	0.14303

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 128.2267

Degrees of freedom of the Q statistic: 8

P value of the Q statistic: 0

Explained variances (R2):

	Estimate
y1: Tau2 (no predictor)	0.0070
y1: Tau2 (with predictors)	0.0080
y1: R2	0.0000
y2: Tau2 (no predictor)	0.0261
y2: Tau2 (with predictors)	0.0250
y2: R2	0.0433

Number of studies (or clusters): 5

Number of observed statistics: 10

Number of estimated parameters: 7

Degrees of freedom: 3

-2 log likelihood: -12.00859

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

R&gt; ## Exact the coefficients

R&gt; coef(mult2)

Intercept1	Intercept2	Slope1_1	Slope2_1
0.344000113	-0.291817459	0.006354034	-0.070588746
Tau2_1_1	Tau2_2_1	Tau2_2_2	
0.008040540	0.009341319	0.025013487	

```
R> ## Exact the sampling variance covariance matrix
R> vcov(mult2)
```

	Intercept1	Intercept2	Slope1_1
Intercept1	7.355789e-03	6.628461e-03	-7.148841e-03
Intercept2	6.628461e-03	1.723435e-02	-6.290031e-03
Slope1_1	-7.148841e-03	-6.290031e-03	1.162591e-02
Slope2_1	-6.692079e-03	-1.692509e-02	9.550531e-03
Tau2_1_1	-1.288708e-04	-2.379371e-05	3.800441e-04
Tau2_2_1	-1.307349e-04	-5.734270e-05	1.944444e-04
Tau2_2_2	-2.517793e-05	-4.989090e-06	4.768716e-05

	Slope2_1	Tau2_1_1	Tau2_2_1
Intercept1	-0.0066920794	-1.288708e-04	-1.307349e-04
Intercept2	-0.0169250890	-2.379371e-05	-5.734270e-05
Slope1_1	0.0095505309	3.800441e-04	1.944444e-04
Slope2_1	0.0262752958	1.400374e-04	2.907864e-04
Tau2_1_1	0.0001400374	1.024271e-04	7.161024e-05
Tau2_2_1	0.0002907864	7.161024e-05	1.113334e-04
Tau2_2_2	0.0001142197	4.726642e-05	1.186402e-04

	Tau2_2_2
Intercept1	-2.517793e-05
Intercept2	-4.989090e-06
Slope1_1	4.768716e-05
Slope2_1	1.142197e-04
Tau2_1_1	4.726642e-05
Tau2_2_1	1.186402e-04
Tau2_2_2	2.916845e-04

```
R> ## Multivariate meta-analysis with both regression coefficients fixed at 0
R> mult0 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL), data=Berkey98,
  x=scale(pub_year, center=1979),
  model.name="Both regression coefficients fixed at 0",
  coef.constraints=matrix(c("0", "0"), nrow=2))
```

Running Both regression coefficients fixed at 0

```
R> summary(mult0)
```

Call:

```
meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
  x = scale(pub_year, center = 1979), data = Berkey98, coef.constraints = matrix(c("0",
  "0"), nrow = 2), model.name = "Both regression coefficients fixed at 0")
```

95% confidence intervals: z statistic approximation

Coefficients:

Estimate	Std.Error	lbound	ubound
----------	-----------	--------	--------

```

Intercept1  0.3448392  0.0536312  0.2397239  0.4499544
Intercept2 -0.3379381  0.0812480 -0.4971812 -0.1786951
Tau2_1_1    0.0070020  0.0090497 -0.0107351  0.0247391
Tau2_2_1    0.0094607  0.0099698 -0.0100797  0.0290010
Tau2_2_2    0.0261445  0.0177409 -0.0086270  0.0609161

```

```
      z value  Pr(>|z|)
```

```

Intercept1  6.4298 1.278e-10 ***
Intercept2 -4.1593 3.192e-05 ***
Tau2_1_1    0.7737   0.4391
Tau2_2_1    0.9489   0.3427
Tau2_2_2    1.4737   0.1406

```

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Q statistic on homogeneity of effect sizes: 128.2267
```

```
Degrees of freedom of the Q statistic: 8
```

```
P value of the Q statistic: 0
```

```
Explained variances (R2):
```

	Estimate
y1: Tau2 (no predictor)	0.0070
y1: Tau2 (with predictors)	0.0070
y1: R2	0.0000
y2: Tau2 (no predictor)	0.0261
y2: Tau2 (with predictors)	0.0261
y2: R2	0.0000

```
Number of studies (or clusters): 5
```

```
Number of observed statistics: 10
```

```
Number of estimated parameters: 5
```

```
Degrees of freedom: 5
```

```
-2 log likelihood: -11.68131
```

```
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)
```

```
R> ## Likelihood ratio test on testing both regression coefficients are 0
```

```
R> anova(mult2, mult0)
```

	base	comparison	ep
1	No constraint		<NA> 7
2	No constraint Both regression coefficients fixed at 0		5
	minus2LL	df	AIC
1	-12.00859	3	-18.00859
2	-11.68131	5	-21.68131
			diffLL
1			NA
2			0.3272789
			diffdf
1			NA
2			2
			p
1			NA
2			0.8490481

```
R> ## Multivariate meta-analysis with an equality constraint on the slopes
```

```

R> mult3 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL), data=Berkey98,
  x=scale(pub_year, center=1979), model.name="With equality constraint",
  coef.constraints=matrix(c("0.3*Equal_Slope", "0.3*Equal_Slope"), nrow=2))

```

Running With equality constraint

```
R> summary(mult3)
```

Call:

```
meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
      x = scale(pub_year, center = 1979), data = Berkey98, coef.constraints = matrix(c("0.3*
      "0.3*Equal_Slope"), nrow = 2), model.name = "With equality constraint")
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound
Intercept1	0.3437612	0.0849828	0.1771980	0.5103245
Equal_Slope	0.0016745	0.1024442	-0.1991124	0.2024614
Intercept2	-0.3390010	0.1041005	-0.5430343	-0.1349678
Tau2_1_1	0.0070474	0.0094638	-0.0115013	0.0255961
Tau2_2_1	0.0095164	0.0105668	-0.0111940	0.0302269
Tau2_2_2	0.0261979	0.0180773	-0.0092330	0.0616287

	z value	Pr(> z )
Intercept1	4.0451	5.231e-05 ***
Equal_Slope	0.0163	0.986958
Intercept2	-3.2565	0.001128 **
Tau2_1_1	0.7447	0.456472
Tau2_2_1	0.9006	0.367800
Tau2_2_2	1.4492	0.147278

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 128.2267

Degrees of freedom of the Q statistic: 8

P value of the Q statistic: 0

Explained variances (R2):

	Estimate
y1: Tau2 (no predictor)	0.0070
y1: Tau2 (with predictors)	0.0070
y1: R2	0.0000
y2: Tau2 (no predictor)	0.0261
y2: Tau2 (with predictors)	0.0262
y2: R2	0.0000

Number of studies (or clusters): 5

Number of observed statistics: 10

Number of estimated parameters: 6

Degrees of freedom: 4

-2 log likelihood: -11.68158

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Exact the coefficients
R> coef(mult3)
```

```
Intercept1 Intercept2 Equal_Slope Tau2_1_1
0.343761239 -0.339001019 0.001674542 0.007047407
Tau2_2_1 Tau2_2_2
0.009516440 0.026197856
```

```
R> ## Exact the sampling variance covariance matrix
R> vcov(mult3)
```

```
Intercept1 Intercept2 Equal_Slope
Intercept1 7.222075e-03 0.0065104549 -0.0067506847
Intercept2 6.510455e-03 0.0108369149 -0.0066601781
Equal_Slope -6.750685e-03 -0.0066601781 0.0104948091
Tau2_1_1 -6.400934e-05 -0.0001221578 0.0002892786
Tau2_2_1 -2.447019e-04 -0.0001027742 0.0003530272
Tau2_2_2 -2.159290e-04 -0.0001825541 0.0003374792
Tau2_1_1 Tau2_2_1 Tau2_2_2
Intercept1 -6.400934e-05 -2.447019e-04 -2.159290e-04
Intercept2 -1.221578e-04 -1.027742e-04 -1.825541e-04
Equal_Slope 2.892786e-04 3.530272e-04 3.374792e-04
Tau2_1_1 8.956372e-05 6.810625e-05 5.473181e-05
Tau2_2_1 6.810625e-05 1.116565e-04 1.308804e-04
Tau2_2_2 5.473181e-05 1.308804e-04 3.267891e-04
```

```
R> ## Likelihood ratio test on the equality of regression coefficients
R> anova(mult2, mult3)
```

	base	comparison	ep	minus2LL	df
1	No constraint	<NA>	7	-12.00859	3
2	No constraint With equality constraint		6	-11.68158	4

	AIC	diffLL	diffdf	p
1	-18.00859	NA	NA	NA
2	-19.68158	0.3270107	1	0.5674246

**Multivariate fixed-effects model** A multivariate fixed-effects meta-analysis can be conducted by fixing the variance component at a zero matrix. The following code illustrates the syntax.

```
R> ## Multivariate meta-analysis with a fixed-effects model
R> summary( mult4 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL),
                        data=Berkey98,
                        RE.constraints=matrix(0, nrow=2, ncol=2)) )
```



Running Meta analysis with ML

Call:

```
meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
     data = Berkey98, RE.constraints = matrix(0, nrow = 2, ncol = 2))
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound	z value
Intercept1	0.307219	0.028575	0.251212	0.363225	10.751
Intercept2	-0.394377	0.018649	-0.430929	-0.357825	-21.147

Pr(>|z|)

Intercept1 < 2.2e-16 \*\*\*

Intercept2 < 2.2e-16 \*\*\*

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 128.2267

Degrees of freedom of the Q statistic: 8

P value of the Q statistic: 0

Heterogeneity indices (based on the estimated Tau2):

	Estimate
Intercept1: I2 (Q statistic)	0
Intercept2: I2 (Q statistic)	0

Number of studies (or clusters): 5

Number of observed statistics: 10

Number of estimated parameters: 2

Degrees of freedom: 8

-2 log likelihood: 90.88326

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

R> ## Exact the coefficients

R> coef(mult4)

Intercept1 Intercept2

0.3072186 -0.3943770

R> ## Exact the sampling variance covariance matrix

R> vcov(mult4)

	Intercept1	Intercept2
Intercept1	0.0008165393	0.0002072041
Intercept2	0.0002072041	0.0003477936

**REML** The `reml()` function may be used to estimate the variance component with the REML estimation method. It should be noted that it does not estimate the fixed-effects. The

fixed-effects estimates can be obtained via the `meta()` function by specifying the estimated variance component from `reml()` as fixed values in the `RE.constraints` argument. This approach is consistent with the idea of REML that removes the fixed-effects parameter when estimating the variance component.

```
R> ## Random-effects meta-analysis with ML
R> summary( meta(y=di, v=vi, data=Becker83) )
```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound	z value
Intercept1	0.174734	0.113378	-0.047482	0.396950	1.5412
Tau2_1_1	0.077376	0.054108	-0.028674	0.183426	1.4300

Pr(>|z|)

Intercept1	0.1233
Tau2_1_1	0.1527

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Heterogeneity indices (based on the estimated Tau2):

	Estimate
Intercept1: I2 (Q statistic)	0.6718

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 2

Degrees of freedom: 8

-2 log likelihood: 7.928307

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Random-effects meta-analysis with REML
R> summary( VarComp <- reml(y=di, v=vi, data=Becker83) )
```

Running Variance component with REML

Call:

```
reml(y = di, v = vi, data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound	z value
--	----------	-----------	--------	--------	---------

```
Tau2_1_1  0.091445  0.064228 -0.034439  0.217329  1.4238
          Pr(>|z|)
Tau2_1_1  0.1545
```

```
Number of studies (or clusters): 10
Number of observed statistics: 9
Number of estimated parameters: 1
Degrees of freedom: 8
-2 log likelihood: -6.110579
OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)
```

```
R> ## Extract the variance component
R> VarComp_REML <- matrix( coef(VarComp), ncol=1, nrow=1 )
R> ## Meta-analysis by treating the variance component as fixed
R> summary( meta(y=di, v=vi, data=Becker83, RE.constraints=VarComp_REML) )
```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, data = Becker83, RE.constraints = VarComp_REML)
```

95% confidence intervals: z statistic approximation

Coefficients:

```
          Estimate Std.Error    lbound    ubound z value
Intercept1 0.180189  0.117535 -0.050176  0.410555  1.5331
          Pr(>|z|)
Intercept1  0.1253
```

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Heterogeneity indices (based on the estimated Tau2):

```
          Estimate
Intercept1: I2 (Q statistic)  0.7075
```

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 1

Degrees of freedom: 9

-2 log likelihood: 7.986161

OpenMx status1: 1 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Estimate variance components with REML
```

```
R> summary( reml(y=di, v=vi, x=log(items), data=Becker83) )
```

Running Variance component with REML

Call:

```
reml(y = di, v = vi, x = log(items), data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound
Tau2_1_1	0.0052656	0.0212014	-0.0362884	0.0468196

z value Pr(>|z|)

Tau2_1_1	0.2484	0.8039
----------	--------	--------

Number of studies (or clusters): 10

Number of observed statistics: 8

Number of estimated parameters: 1

Degrees of freedom: 7

-2 log likelihood: -10.84567

OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

R> ## Estimate variance components with REML

R> summary( reml(y=cbind(PD, AL), v=cbind(var\_PD, cov\_PD\_AL, var\_AL), data=Berkey98) )

Running Variance component with REML

Call:

```
reml(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
     data = Berkey98)
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound	z value
Tau2_1_1	0.011733	0.013645	-0.015011	0.038477	0.8599
Tau2_2_1	0.011916	0.014416	-0.016340	0.040172	0.8266
Tau2_2_2	0.032651	0.024402	-0.015176	0.080479	1.3380

Pr(>|z|)

Tau2_1_1	0.3899
Tau2_2_1	0.4085
Tau2_2_2	0.1809

Number of studies (or clusters): 5

Number of observed statistics: 8

Number of estimated parameters: 3

Degrees of freedom: 5

-2 log likelihood: -18.86768

OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

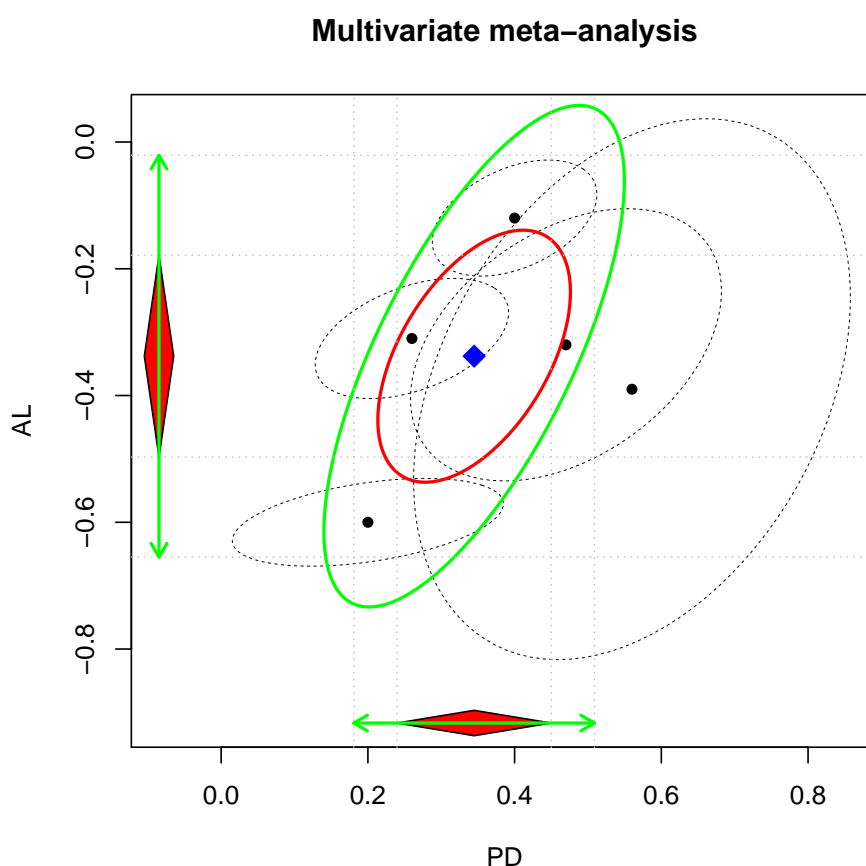
**Plots of multivariate effect sizes** If multivariate meta-analysis is conducted, pairwise plots on the pooled effect sizes and their confidence ellipses can be obtained via the `plot()`

function. By default, 95% confidence intervals on the average effect sizes and confidence ellipses on the random effects are plotted. For example,

```
R> Berkey98.ma <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL), data=Berkey98)
```

Running Meta analysis with ML

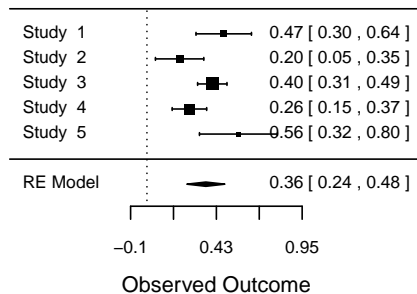
```
R> plot(Berkey98.ma, main="Multivariate meta-analysis", axis.label=c("PD", "AL"))
```



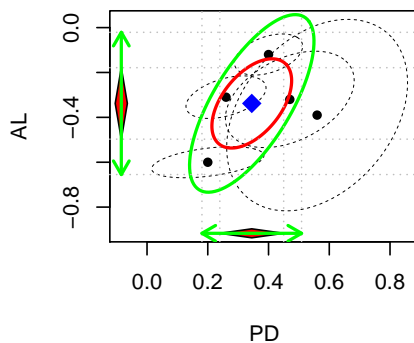
By combining with the forest plots from the **metafor** package, we may combine the univariate and multivariate natures of the effect sizes in a single figure. This will be very useful for multivariate meta-analysis.

```
R> ## Load the metafor package to display forest plots
R> library(metafor)
R> plot(Berkey98.ma, diag.panel=TRUE, main="Multivariate meta-analysis",
       axis.label=c("PD", "AL"))
R> forest( rma(yi=PD, vi=var_PD, data=Berkey98) )
R> title("Forest plot for PD")
R> forest( rma(yi=AL, vi=var_AL, data=Berkey98) )
R> title("Forest plot for AL")
```

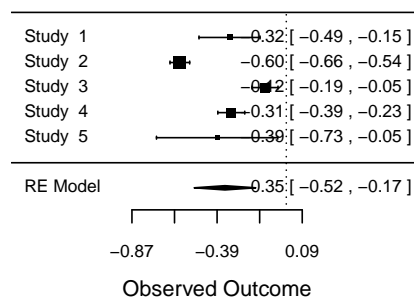
Forest plot for PD



Multivariate meta-analysis



Forest plot for AL



**Three-level meta-analysis** The `meta3()` function may be used to fit three-level meta-analytic models. It is assumed that effect sizes within `cluster` are dependent. The  $Q$  statistic ( $df = 55$ ) of the following example is 578.864,  $p < .001$ . The  $I^2$  at level-2 and level-3 are .3440 and .6043, respectively. The estimated coefficient of *Year* of publication is 0.0051,  $p = .5518$ . The  $R^2$  at level-2 and level-3 are .0000 and .0221, respectively.

```
R> ## ML estimation method
R> ## No predictor
R> summary( meta3(y=y, v=v, cluster=District, data=Cooper03) )
```

Running Meta analysis with ML

```
Call:
meta3(y = y, v = v, cluster = District, data = Cooper03)
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound
Intercept	0.1844553	0.0805411	0.0265977	0.3423130

```

Tau2_2      0.0328648  0.0111397  0.0110314  0.0546982
Tau2_3      0.0577384  0.0307423 -0.0025154  0.1179921

```

```

      z value Pr(>|z|)

```

```

Intercept  2.2902 0.022010 *
Tau2_2     2.9502 0.003175 **
Tau2_3     1.8781 0.060362 .

```

```

---

```

```

Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

```

Q statistic on homogeneity of effect sizes: 578.864

```

```

Degrees of freedom of the Q statistic: 55

```

```

P value of the Q statistic: 0

```

```

Heterogeneity indices (based on the estimated Tau2):

```

```

      Estimate

```

```

I2_2 (Q statistic)  0.3440
I2_3 (Q statistic)  0.6043

```

```

Number of studies (or clusters): 11

```

```

Number of observed statistics: 56

```

```

Number of estimated parameters: 3

```

```

Degrees of freedom: 53

```

```

-2 log likelihood: 16.78987

```

```

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

```

R> ## Year as a predictor

```

```

R> summary( meta3(y=y, v=v, cluster=District, x=scale(Year, scale=FALSE), data=Cooper03) )

```

```

Running Meta analysis with ML

```

```

Call:

```

```

meta3(y = y, v = v, cluster = District, x = scale(Year, scale = FALSE),
      data = Cooper03)

```

```

95% confidence intervals: z statistic approximation

```

```

Coefficients:

```

```

      Estimate Std.Error    lbound    ubound
Slope_1  0.0050737  0.0085266 -0.0116382  0.0217856
Intercept 0.1780268  0.0805219  0.0202067  0.3358469
Tau2_2    0.0329390  0.0111620  0.0110618  0.0548162
Tau2_3    0.0564628  0.0300330 -0.0024007  0.1153264

```

```

      z value Pr(>|z|)

```

```

Slope_1    0.5950 0.551814
Intercept  2.2109 0.027042 *
Tau2_2     2.9510 0.003168 **
Tau2_3     1.8800 0.060104 .

```

```

---

```

```

Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Q statistic on homogeneity of effect sizes: 578.864

Degrees of freedom of the Q statistic: 55

P value of the Q statistic: 0

Explained variances (R2):

	Estimate
Tau2_2 (no predictor)	0.0329
Tau2_2 (with predictors)	0.0329
R2_2 (level-2)	0.0000
Tau2_3 (no predictor)	0.0577
Tau2_3 (with predictors)	0.0565
R2_3 (level-3)	0.0221

Number of studies (or clusters): 11

Number of observed statistics: 56

Number of estimated parameters: 4

Degrees of freedom: 52

-2 log likelihood: 16.43629

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

R> ## REML estimation method

R> ## No predictor; with LBCI

R> summary( reml3(y=y, v=v, cluster=District, data=Cooper03, intervals.type="LB") )

Running Variance component with REML

Call:

reml3(y = y, v = v, cluster = District, data = Cooper03, intervals.type = "LB")

95% confidence intervals: Likelihood-based statistic

Coefficients:

	Estimate	Std.Error	lbound	ubound	z value
Tau2_2	0.032737	0.011092	0.016264	0.062842	2.9513
Tau2_3	0.065062	0.035510	0.022234	0.207846	1.8322
	Pr(> z )				
Tau2_2	0.003164	**			
Tau2_3	0.066921	.			

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Number of studies (or clusters): 56

Number of observed statistics: 55

Number of estimated parameters: 2

Degrees of freedom: 53

-2 log likelihood: -81.14044

OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)



```

R> ## Year as a predictor
R> summary( reml3(y=y, v=v, cluster=District, x=scale(Year, scale=FALSE), data=Cooper03) )

Running Variance component with REML

Call:
reml3(y = y, v = v, cluster = District, x = scale(Year, scale = FALSE),
      data = Cooper03)

95% confidence intervals: z statistic approximation
Coefficients:
      Estimate Std.Error    lbound    ubound z value
Tau2_2  0.0326502  0.0110529  0.0109870  0.0543134  2.9540
Tau2_3  0.0722656  0.0405349 -0.0071813  0.1517125  1.7828
      Pr(>|z|)
Tau2_2 0.003137 **
Tau2_3 0.074619 .
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Number of studies (or clusters): 56
Number of observed statistics: 54
Number of estimated parameters: 2
Degrees of freedom: 52
-2 log likelihood: -72.09405
OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

```

### 3. Meta-analytic structural equation modeling

MASEM combines the idea of meta-analysis and SEM by pooling correlation/covariance matrices and testing structural equation models on the pooled correlation/covariance matrix. There are two stages in conducting a MASEM. In the first stage the correlation/covariance matrices are pooled together. In the second stage, the pooled correlation/covariance matrix is used to fit structural equation models.

[Cheung and Chan \(2005b, 2009\)](#) proposed a two-stage structural equation modeling (TSSEM) based on a fixed-effects model. The `metaSEM` package has implemented the TSSEM approach. Moreover, the TSSEM approach has been extended to the random-effects model by using a multivariate meta-analysis ([Cheung 2011b](#)) in the first stage analysis. Regardless of whether a fixed- or random-effects model is used, the `tssem2()` function will handle this automatically. In other words, parameter estimates, standard errors and goodness-of-fit indices in the stage 2 analysis has already taken the stage 1 model into account.

An example from [Cheung \(2009b\)](#) is used to illustrate the procedures. In this example, [Digman \(1997\)](#) reported a second-order factor analysis on a five-factor model with 14 studies. He proposed that there were two second-order factors for the five-factor model: an alpha factor for agreeableness, conscientiousness, and emotional stability, and a beta factor for extroversion and intellect.

### 3.1. Fixed-effects model

The `tssem1()` function is used to pool the correlation matrices with a fixed-effects model in the first stage by specifying `method='FEM'` in the argument. `tssem2()` is then used to fit a factor analytic model on the pooled correlation matrix with the inverse of its asymptotic covariance matrix as the weight matrix (Cheung and Chan 2005b, 2009).

The fit indices for testing the homogeneity of the correlation matrices in Stage 1 analysis are  $\chi^2(130, N = 4496) = 1499.73, p < .001$ , CFI=0.6825, TLI=0.6581, SRMR=0.1750 and RMSEA=0.1812. This indicates that it is not reasonable to assume that the correlation matrices are homogenous. Sub-group analysis or random-effects model that will be illustrated later are more appropriate. As an exercise, we continue to fit the stage 2 model. The fit indices for fitting the structural model in Stage 2 are  $\chi^2(4, N = 4496) = 67.89, p < .001$ , CFI=0.9845, TLI=0.9613, SRMR=0.0285 and RMSEA=0.0596.

```
R> ## Show the first 2 studies in Digman97
R> head(Digman97$data, n=2)
```

```
$`Digman 1 (1994)`
      A      C    ES      E      I
A  1.00  0.62 0.41 -0.48 0.00
C  0.62  1.00 0.59 -0.10 0.35
ES 0.41  0.59 1.00  0.27 0.41
E -0.48 -0.10 0.27  1.00 0.37
I  0.00  0.35 0.41  0.37 1.00
```

```
$`Digman 2 (1994)`
      A      C    ES      E      I
A  1.00  0.39 0.53 -0.30 -0.05
C  0.39  1.00 0.59  0.07  0.44
ES 0.53  0.59 1.00  0.09  0.22
E -0.30 0.07 0.09  1.00  0.45
I -0.05 0.44 0.22  0.45  1.00
```

```
R> ## Show the first 2 sample sizes in Digman97
R> head(Digman97$n, n=2)
```

```
[1] 102 149
```

```
R> ## Example of Fixed-effects TSSEM
R> fixed1 <- tssem1(Digman97$data, Digman97$n, method="FEM")
```

Running TSSEM1 Analysis of Correlation Matrix

```
R> summary(fixed1)
```

Call:

```
tssem1FEM(my.df = my.df, n = n, cor.analysis = cor.analysis,
```

```
model.name = model.name, cluster = cluster, suppressWarnings = suppressWarnings)
```

Coefficients:

	Estimate	Std.Error	z value	Pr(> z )
S[1,2]	0.363116	0.013390	27.1174	< 2.2e-16 ***
S[1,3]	0.390176	0.012903	30.2394	< 2.2e-16 ***
S[1,4]	0.103751	0.015070	6.8846	5.794e-12 ***
S[1,5]	0.092246	0.015071	6.1207	9.316e-10 ***
S[2,3]	0.415999	0.012540	33.1744	< 2.2e-16 ***
S[2,4]	0.135208	0.014799	9.1363	< 2.2e-16 ***
S[2,5]	0.141213	0.014891	9.4834	< 2.2e-16 ***
S[3,4]	0.244505	0.014175	17.2488	< 2.2e-16 ***
S[3,5]	0.138167	0.014858	9.2992	< 2.2e-16 ***
S[4,5]	0.424514	0.012396	34.2466	< 2.2e-16 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Goodness-of-fit indices:

	Value
Sample size	4496.0000
Chi-square of target model	1499.7340
DF of target model	130.0000
p value of target model	0.0000
Chi-square of independent model	4454.5995
DF of independent model	140.0000
RMSEA	0.1812
SRMR	0.1750
TLI	0.6581
CFI	0.6825
AIC	1239.7340
BIC	406.3114

OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Extract the pooled correlation matrix
```

```
R> coef(fixed1)
```

	x1	x2	x3	x4	x5
x1	1.00000000	0.3631157	0.3901765	0.1037511	0.09224586
x2	0.36311572	1.0000000	0.4159987	0.1352076	0.14121294
x3	0.39017646	0.4159987	1.0000000	0.2445051	0.13816667
x4	0.10375113	0.1352076	0.2445051	1.0000000	0.42451421
x5	0.09224586	0.1412129	0.1381667	0.4245142	1.00000000

```
R> ## Exact the sampling variance covariance matrix
```

```
R> vcov(fixed1)
```

	x2 x1	x3 x1	x4 x1	x5 x1
x2 x1	1.469907e-04	3.838613e-05	1.459870e-05	1.738852e-05

```

x3 x1 3.838613e-05 1.380993e-04 3.291322e-05 1.593040e-05
x4 x1 1.459870e-05 3.291322e-05 2.131330e-04 8.828891e-05
x5 x1 1.738852e-05 1.593040e-05 8.828891e-05 2.160834e-04
x3 x2 3.188759e-05 2.548896e-05 6.819538e-06 5.068077e-06
x4 x2 7.181122e-06 5.714476e-06 7.178708e-05 2.902241e-05
x5 x2 5.973882e-06 2.980482e-06 2.891491e-05 7.375722e-05
x4 x3 1.147757e-06 9.431105e-07 7.186433e-05 2.930447e-05
x5 x3 3.714609e-06 5.377907e-06 2.994837e-05 7.926245e-05
x5 x4 4.522226e-07 8.036684e-07 7.138436e-06 9.757525e-06
      x3 x2      x4 x2      x5 x2      x4 x3
x2 x1 3.188759e-05 7.181122e-06 5.973882e-06 1.147757e-06
x3 x1 2.548896e-05 5.714476e-06 2.980482e-06 9.431105e-07
x4 x1 6.819538e-06 7.178708e-05 2.891491e-05 7.186433e-05
x5 x1 5.068077e-06 2.902241e-05 7.375722e-05 2.930447e-05
x3 x2 1.292690e-04 2.905738e-05 1.189870e-05 4.659802e-06
x4 x2 2.905738e-05 2.083179e-04 8.423570e-05 7.495569e-05
x5 x2 1.189870e-05 8.423570e-05 2.092633e-04 3.026949e-05
x4 x3 4.659802e-06 7.495569e-05 3.026949e-05 1.859687e-04
x5 x3 1.248480e-05 3.164628e-05 8.223973e-05 7.646729e-05
x5 x4 1.707692e-06 1.252276e-05 1.126839e-05 4.745070e-06
      x5 x3      x5 x4
x2 x1 3.714609e-06 4.522226e-07
x3 x1 5.377907e-06 8.036684e-07
x4 x1 2.994837e-05 7.138436e-06
x5 x1 7.926245e-05 9.757525e-06
x3 x2 1.248480e-05 1.707692e-06
x4 x2 3.164628e-05 1.252276e-05
x5 x2 8.223973e-05 1.126839e-05
x4 x3 7.646729e-05 4.745070e-06
x5 x3 2.077769e-04 2.841400e-05
x5 x4 2.841400e-05 1.270258e-04

```

```
R> ## S matrix
```

```
R> Phi <- matrix(c(1,"0.3*cor","0.3*cor",1), ncol=2, nrow=2)
```

```
R> S1 <- bdiagMat(list(diag(c("0.2*e1","0.2*e2","0.2*e3","0.2*e4","0.2*e5")), Phi))
```

```
R> S1 <- as.mxMatrix(S1)
```

```
R> ## A matrix
```

```
R> Lambda <- matrix(c(0,".3*f1_x2",".3*f1_x3",".3*f1_x4",0,".3*f2_x1",0,0,0,".3*f2_x5"),
                    ncol=2, nrow=5)
```

```
R> A1 <- rbind( cbind(matrix(0,ncol=5,nrow=5), Lambda),
               matrix(0, ncol=7, nrow=2) )
```

```
R> A1 <- as.mxMatrix(A1)
```

```
R> F1 <- create.Fmatrix(c(1,1,1,1,1,0,0), name="F1")
```

```
R> fixed2 <- tssem2(fixed1, Amatrix=A1, Smatrix=S1, Fmatrix=F1, diag.constraints=TRUE, int
                    model.name="TSSEM2 Digman97")
```

```
Running TSSEM2 Digman97
```

```
R> summary(fixed2)
```

Call:

```
wls(Cov = tssem1.obj$pooledS, asyCov = tssem1.obj$acovS, n = tssem1.obj$total.n,
     Amatrix = Amatrix, Smatrix = Smatrix, Fmatrix = Fmatrix,
     diag.constraints = diag.constraints, cor.analysis = cor.analysis,
     intervals.type = intervals.type, model.name = model.name,
     suppressWarnings = suppressWarnings)
```

95% confidence intervals: Likelihood-based statistic

Coefficients:

	Estimate	Std.Error	lbound	ubound	z	value
Amatrix[1,7]	0.48934	NA	0.45743	0.52098	NA	
Amatrix[2,6]	0.57156	NA	0.54692	0.59628	NA	
Amatrix[3,6]	0.66392	NA	0.64030	0.68769	NA	
Amatrix[4,6]	0.58220	NA	0.55545	0.60895	NA	
Amatrix[5,7]	0.48476	NA	0.45134	0.51790	NA	
Smatrix[1,1]	0.76055	NA	0.72858	0.79076	NA	
Smatrix[2,2]	0.67332	NA	0.64445	0.70088	NA	
Smatrix[3,3]	0.55921	NA	0.52708	0.59001	NA	
Smatrix[4,4]	0.66105	NA	0.62917	0.69147	NA	
Smatrix[5,5]	0.76501	NA	0.73178	0.79629	NA	
Smatrix[7,6]	1.08026	NA	1.02716	1.14018	NA	

	Pr(> z )
Amatrix[1,7]	NA
Amatrix[2,6]	NA
Amatrix[3,6]	NA
Amatrix[4,6]	NA
Amatrix[5,7]	NA
Smatrix[1,1]	NA
Smatrix[2,2]	NA
Smatrix[3,3]	NA
Smatrix[4,4]	NA
Smatrix[5,5]	NA
Smatrix[7,6]	NA

Goodness-of-fit indices:

	Value
Sample size	4496.0000
Chi-square of target model	706.3461
DF of target model	4.0000
p value of target model	0.0000
Number of constraints imposed on "Smatrix"	5.0000
DF manually adjusted	0.0000
Chi-square of independent model	4132.8584
DF of independent model	10.0000
RMSEA	0.1976

```

SRMR                0.1455
TLI                 0.5741
CFI                 0.8296
AIC                 698.3461
BIC                 672.7023
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

**Example: Fixed-effects model with sub-group analysis** Studies may not share the same population correlation matrix. If the studies can be grouped into various subgroups, we may pool the correlation matrices separately by the subgroups (Cheung and Chan 2005a). This is similar to the subgroup analysis in conventional meta-analysis (Hedges and Olkin 1985). For example, Digman (1997) groups the 14 studies into several groups according to their sample characteristics. These include children, adolescents, young adults, and mature adults. This variable is stored in the variable `Digman97$cluster`. We may further group these studies into younger participants versus older participants. Separate fixed-effects analysis may be applied into these two groups.

The following R code may be used to conduct the analysis. Users have to supply the cluster (a vector of labels) to the `cluster` argument in `tssem1()`. The correlation/covariance matrices will be pooled separately for each cluster. The structural models will also be fitted separately for each cluster.

```

R> ##### Create a variable for different samples
R> ##### Younger participants: Children and Adolescents
R> ##### Older participants: others
R> sample <- ifelse(Digman97$cluster %in% c("Children","Adolescents"),
                    yes="Younger participants", no="Older participants")
R> ##### Show the sample
R> sample

```

```

[1] "Younger participants" "Younger participants"
[3] "Younger participants" "Younger participants"
[5] "Younger participants" "Older participants"
[7] "Older participants"   "Older participants"
[9] "Older participants"   "Older participants"
[11] "Older participants"   "Older participants"
[13] "Older participants"   "Older participants"

```

```

R> ## Example of Fixed-effects TSSEM with several clusters
R> fixed1.cluster <- tssem1(Digman97$data, Digman97$n, method="FEM",
                           cluster=sample)

```

```

Running TSSEM1 Analysis of Correlation Matrix
Running TSSEM1 Analysis of Correlation Matrix

```

```

R> ##### Please note that the estimates for the younger participants are problematic.
R> summary(fixed1.cluster)

```

\$`Older participants`

Call:

```
tssem1FEM(my.df = data.cluster[[i]], n = n.cluster[[i]], cor.analysis = cor.analysis,
  model.name = model.name, suppressWarnings = suppressWarnings)
```

Coefficients:

	Estimate	Std.Error	z value	Pr(> z )
S[1,2]	0.297484	0.015455	19.2487	< 2.2e-16 ***
S[1,3]	0.370088	0.014552	25.4321	< 2.2e-16 ***
S[1,4]	0.137688	0.016423	8.3838	< 2.2e-16 ***
S[1,5]	0.097971	0.016744	5.8510	4.886e-09 ***
S[2,3]	0.393709	0.014163	27.7981	< 2.2e-16 ***
S[2,4]	0.182984	0.016075	11.3831	< 2.2e-16 ***
S[2,5]	0.092664	0.016664	5.5609	2.684e-08 ***
S[3,4]	0.260756	0.015573	16.7443	< 2.2e-16 ***
S[3,5]	0.096063	0.016594	5.7890	7.079e-09 ***
S[4,5]	0.411753	0.013917	29.5857	< 2.2e-16 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Goodness-of-fit indices:

	Value
Sample size	3658.0000
Chi-square of target model	823.8769
DF of target model	80.0000
p value of target model	0.0000
Chi-square of independent model	2992.9294
DF of independent model	90.0000
RMSEA	0.1513
SRMR	0.1528
TLI	0.7117
CFI	0.7437
AIC	663.8769
BIC	167.5032

OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

\$`Younger participants`

Call:

```
tssem1FEM(my.df = data.cluster[[i]], n = n.cluster[[i]], cor.analysis = cor.analysis,
  model.name = model.name, suppressWarnings = suppressWarnings)
```

Coefficients:

	Estimate	Std.Error	z value	Pr(> z )
S[1,2]	0.604396	0.022189	27.2391	< 2.2e-16 ***
S[1,3]	0.465441	0.027579	16.8768	< 2.2e-16 ***

```

S[1,4] -0.030869  0.036047 -0.8563   0.39181
S[1,5]  0.061581  0.034650  1.7772   0.07553 .
S[2,3]  0.501309  0.026431 18.9668 < 2.2e-16 ***
S[2,4] -0.060834  0.034660 -1.7552   0.07923 .
S[2,5]  0.321019  0.031156 10.3035 < 2.2e-16 ***
S[3,4]  0.175422  0.033776  5.1937 2.061e-07 ***
S[3,5]  0.305214  0.031679  9.6345 < 2.2e-16 ***
S[4,5]  0.478573  0.026966 17.7474 < 2.2e-16 ***

```

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Goodness-of-fit indices:
```

	Value
Sample size	838.0000
Chi-square of target model	344.1826
DF of target model	40.0000
p value of target model	0.0000
Chi-square of independent model	1461.6701
DF of independent model	50.0000
RMSEA	0.2131
SRMR	0.1508
TLI	0.7307
CFI	0.7845
AIC	264.1826
BIC	74.9419

```
OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)
```

```
R> ## Extract the pooled correlation matrices
```

```
R> coef(fixed1.cluster)
```

```
$`Older participants`
```

	x1	x2	x3	x4	x5
x1	1.00000000	0.29748426	0.3700879	0.1376877	0.09797102
x2	0.29748426	1.00000000	0.3937095	0.1829840	0.09266441
x3	0.37008785	0.39370949	1.00000000	0.2607561	0.09606280
x4	0.13768766	0.18298404	0.2607561	1.00000000	0.41175343
x5	0.09797102	0.09266441	0.0960628	0.4117534	1.00000000

```
$`Younger participants`
```

	x1	x2	x3	x4	x5
x1	1.00000000	0.60439588	0.4654414	-0.03086891	0.06158054
x2	0.60439588	1.00000000	0.5013091	-0.06083381	0.32101905
x3	0.46544142	0.50130913	1.00000000	0.17542223	0.30521426
x4	-0.03086891	-0.06083381	0.1754222	1.00000000	0.47857275
x5	0.06158054	0.32101905	0.3052143	0.47857275	1.00000000

```
R> fixed2.cluster <- tssem2(fixed1.cluster, Amatrix=A1, Smatrix=S1, Fmatrix=F1,
                             diag.constraints=TRUE, intervals.type="LB")
```



Running TSSEM2 (Fixed Effects Model) Analysis of Correlation Structure  
 Running TSSEM2 (Fixed Effects Model) Analysis of Correlation Structure

```
R> summary(fixed2.cluster)
```

```
$`Older participants`
```

```
Call:
```

```
wls(Cov = tssem1.obj$pooledS, asyCov = tssem1.obj$acovS, n = tssem1.obj$total.n,  
    Amatrix = Amatrix, Smatrix = Smatrix, Fmatrix = Fmatrix,  
    diag.constraints = diag.constraints, cor.analysis = cor.analysis,  
    intervals.type = intervals.type, model.name = model.name,  
    suppressWarnings = suppressWarnings)
```

95% confidence intervals: Likelihood-based statistic

Coefficients:

	Estimate	Std.Error	lbound	ubound	z value
Amatrix[1,7]	0.48671	NA	0.44908	0.52432	NA
Amatrix[2,6]	0.52460	NA	0.49543	0.55387	NA
Amatrix[3,6]	0.65180	NA	0.62355	0.68029	NA
Amatrix[4,6]	0.63182	NA	0.60170	0.66210	NA
Amatrix[5,7]	0.47139	NA	0.43253	0.51019	NA
Smatrix[1,1]	0.76312	NA	0.72509	0.79832	NA
Smatrix[2,2]	0.72479	NA	0.69323	0.75455	NA
Smatrix[3,3]	0.57516	NA	0.53720	0.61119	NA
Smatrix[4,4]	0.60080	NA	0.56162	0.63796	NA
Smatrix[5,5]	0.77779	NA	0.73970	0.81292	NA
Smatrix[7,6]	0.99100	NA	0.93156	1.05849	NA

Pr(>|z|)

Amatrix[1,7]	NA
Amatrix[2,6]	NA
Amatrix[3,6]	NA
Amatrix[4,6]	NA
Amatrix[5,7]	NA
Smatrix[1,1]	NA
Smatrix[2,2]	NA
Smatrix[3,3]	NA
Smatrix[4,4]	NA
Smatrix[5,5]	NA
Smatrix[7,6]	NA

Goodness-of-fit indices:

	Value
Sample size	3658.0000
Chi-square of target model	510.6333
DF of target model	4.0000
p value of target model	0.0000

```

Number of constraints imposed on "Smatrix"      5.0000
DF manually adjusted                             0.0000
Chi-square of independent model                 2940.1270
DF of independent model                         10.0000
RMSEA                                           0.1861
SRMR                                           0.1334
TLI                                             0.5677
CFI                                             0.8271
AIC                                           502.6333
BIC                                           477.8146
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

```
$`Younger participants`
```

```
Call:
```

```

wls(Cov = tssem1.obj$pooledS, asyCov = tssem1.obj$acovS, n = tssem1.obj$total.n,
     Amatrix = Amatrix, Smatrix = Smatrix, Fmatrix = Fmatrix,
     diag.constraints = diag.constraints, cor.analysis = cor.analysis,
     intervals.type = intervals.type, model.name = model.name,
     suppressWarnings = suppressWarnings)

```

```
95% confidence intervals: Likelihood-based statistic
```

```
Coefficients:
```

	Estimate	Std.Error	lbound	ubound	z	value
Amatrix[1,7]	0.52484	NA	0.47174	0.57535	NA	
Amatrix[2,6]	0.78495	NA	0.74995	0.82035	NA	
Amatrix[3,6]	0.66595	NA	0.62729	0.70432	NA	
Amatrix[4,6]	0.45613	NA	0.40515	0.50694	NA	
Amatrix[5,7]	0.59209	NA	0.53203	0.64914	NA	
Smatrix[1,1]	0.72454	NA	0.66897	0.77746	NA	
Smatrix[2,2]	0.38385	NA	0.32701	0.43758	NA	
Smatrix[3,3]	0.55651	NA	0.50393	0.60651	NA	
Smatrix[4,4]	0.79194	NA	0.74302	0.83585	NA	
Smatrix[5,5]	0.64942	NA	0.57860	0.71694	NA	
Smatrix[7,6]	1.36830	NA	1.27802	1.48517	NA	

	Pr(> z )
Amatrix[1,7]	NA
Amatrix[2,6]	NA
Amatrix[3,6]	NA
Amatrix[4,6]	NA
Amatrix[5,7]	NA
Smatrix[1,1]	NA
Smatrix[2,2]	NA
Smatrix[3,3]	NA
Smatrix[4,4]	NA
Smatrix[5,5]	NA
Smatrix[7,6]	NA

Goodness-of-fit indices:

	Value
Sample size	838.0000
Chi-square of target model	314.7823
DF of target model	4.0000
p value of target model	0.0000
Number of constraints imposed on "Smatrix"	5.0000
DF manually adjusted	0.0000
Chi-square of independent model	3632.8021
DF of independent model	10.0000
RMSEA	0.3047
SRMR	0.2347
TLI	0.7855
CFI	0.9142
AIC	306.7823
BIC	287.8582
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)	

### 3.2. Random-effects model

TSSEM using a random-effects model may be requested by specifying the `method='REM'` argument in `tssem1()`. By default (`RE.diag.only=FALSE`), a positive definite covariance matrix among the random-effects is used. For practical reasons, e.g., there are not enough studies, it may not be feasible to estimate the full covariance matrix of the random effects. A diagonal matrix on the random-effects may also be used by specifying `RE.diag.only=TRUE`.

The fit indices for fitting the structural model in Stage 2 are  $\chi^2(4, N = 4496) = 8.28, p < .001$ , CFI=0.9920, TLI=0.9801, SRMR=0.0154 and RMSEA=0.0154. This indicates that the model fits the data quite well.

```
R> random1 <- tssem1(Digman97$data, Digman97$n, method="REM", RE.diag.only=TRUE)
```

Running TSSEM1 (Random Effects Model) Analysis of Correlation Matrix

```
R> summary(random1)
```

Call:

```
meta(y = ES, v = acovR, RE.constraints = diag(x = paste(RE.startvalues,
  "*Tau2_", 1:no.es, "_", 1:no.es, sep = ""), nrow = no.es,
  ncol = no.es), I2 = I2, model.name = model.name)
```

95% confidence intervals: z statistic approximation

Coefficients:

	Estimate	Std.Error	lbound	ubound
Intercept1	0.39981616	0.05455521	0.29288991	0.50674240
Intercept2	0.44433503	0.04168026	0.36264323	0.52602683

Intercept3	0.05444821	0.06316912	-0.06936099	0.17825740
Intercept4	0.10138329	0.04681343	0.00963065	0.19313593
Intercept5	0.43415295	0.04000883	0.35573708	0.51256881
Intercept6	0.12867856	0.04174106	0.04686759	0.21048953
Intercept7	0.20732505	0.04973237	0.10985139	0.30479870
Intercept8	0.24064433	0.03220608	0.17752158	0.30376708
Intercept9	0.19296080	0.04340494	0.10788867	0.27803292
Intercept10	0.44713470	0.03211662	0.38418729	0.51008211
Tau2_1_1	0.03815841	0.01523926	0.00829000	0.06802682
Tau2_2_2	0.02132562	0.00868724	0.00429894	0.03835230
Tau2_3_3	0.05115980	0.02059751	0.01078943	0.09153017
Tau2_4_4	0.02571721	0.01094038	0.00427445	0.04715996
Tau2_5_5	0.01901262	0.00820033	0.00294026	0.03508498
Tau2_6_6	0.01977637	0.00914599	0.00185056	0.03770218
Tau2_7_7	0.02995567	0.01234177	0.00576625	0.05414510
Tau2_8_8	0.01030039	0.00505944	0.00038407	0.02021671
Tau2_9_9	0.02172535	0.00934582	0.00340788	0.04004282
Tau2_10_10	0.01122090	0.00494562	0.00152766	0.02091414

	z value	Pr(> z )
--	---------	----------

Intercept1	7.3287	2.325e-13 ***
Intercept2	10.6606	< 2.2e-16 ***
Intercept3	0.8619	0.388719
Intercept4	2.1657	0.030335 *
Intercept5	10.8514	< 2.2e-16 ***
Intercept6	3.0828	0.002051 **
Intercept7	4.1688	3.062e-05 ***
Intercept8	7.4720	7.905e-14 ***
Intercept9	4.4456	8.765e-06 ***
Intercept10	13.9222	< 2.2e-16 ***
Tau2_1_1	2.5040	0.012281 *
Tau2_2_2	2.4548	0.014095 *
Tau2_3_3	2.4838	0.012999 *
Tau2_4_4	2.3507	0.018740 *
Tau2_5_5	2.3185	0.020421 *
Tau2_6_6	2.1623	0.030595 *
Tau2_7_7	2.4272	0.015217 *
Tau2_8_8	2.0359	0.041763 *
Tau2_9_9	2.3246	0.020093 *
Tau2_10_10	2.2689	0.023277 *

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 2694.373

Degrees of freedom of the Q statistic: 130

P value of the Q statistic: 0

Heterogeneity indices (based on the estimated Tau2):

Estimate

```

Intercept1: I2 (Q statistic)    0.9597
Intercept2: I2 (Q statistic)    0.9283
Intercept3: I2 (Q statistic)    0.9496
Intercept4: I2 (Q statistic)    0.9008
Intercept5: I2 (Q statistic)    0.9193
Intercept6: I2 (Q statistic)    0.8750
Intercept7: I2 (Q statistic)    0.9184
Intercept8: I2 (Q statistic)    0.8008
Intercept9: I2 (Q statistic)    0.8877
Intercept10: I2 (Q statistic)   0.8696

```

Number of studies (or clusters): 14

Number of observed statistics: 140

Number of estimated parameters: 20

Degrees of freedom: 120

-2 log likelihood: -109.6847

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

R> ##### Extract the fixed-effects (pooled correlation matrix)

R> coef(random1, select="fixed")

```

Intercept1 Intercept2 Intercept3 Intercept4 Intercept5
0.39981616 0.44433503 0.05444821 0.10138329 0.43415295
Intercept6 Intercept7 Intercept8 Intercept9 Intercept10
0.12867856 0.20732505 0.24064433 0.19296080 0.44713470

```

R> ##### Extract the sampling variance covariance matrix

R> vcov(random1, select="fixed")

```

                Intercept1 Intercept2 Intercept3
Intercept1 2.971643e-03 4.882045e-05 3.569778e-05
Intercept2 4.882045e-05 1.735261e-03 4.983673e-05
Intercept3 3.569778e-05 4.983673e-05 3.988377e-03
Intercept4 3.692228e-05 2.871922e-05 1.239071e-04
Intercept5 6.644416e-05 4.542286e-05 1.547136e-05
Intercept6 -5.730980e-06 6.374690e-06 1.069199e-04
Intercept7 5.089118e-07 4.445662e-06 4.208838e-05
Intercept8 3.231296e-06 -1.405955e-05 1.208248e-04
Intercept9 6.773077e-06 -1.495886e-06 4.465584e-05
Intercept10 2.837357e-06 3.261207e-06 1.324198e-05
                Intercept4 Intercept5 Intercept6
Intercept1 3.692228e-05 6.644416e-05 -5.730980e-06
Intercept2 2.871922e-05 4.542286e-05 6.374690e-06
Intercept3 1.239071e-04 1.547136e-05 1.069199e-04
Intercept4 2.189113e-03 1.511075e-05 3.989855e-05
Intercept5 1.511075e-05 1.589619e-03 4.616246e-05
Intercept6 3.989855e-05 4.616246e-05 1.739234e-03

```

```

Intercept7  1.146084e-04 2.227720e-05  1.194529e-04
Intercept8  4.672660e-05 1.193903e-05  1.017265e-04
Intercept9  1.349447e-04 2.750647e-05  4.064882e-05
Intercept10 -7.762461e-06 3.784392e-06  3.451132e-05
      Intercept7      Intercept8      Intercept9
Intercept1  5.089118e-07  3.231296e-06  6.773077e-06
Intercept2  4.445662e-06 -1.405955e-05 -1.495886e-06
Intercept3  4.208838e-05  1.208248e-04  4.465584e-05
Intercept4  1.146084e-04  4.672660e-05  1.349447e-04
Intercept5  2.227720e-05  1.193903e-05  2.750647e-05
Intercept6  1.194529e-04  1.017265e-04  4.064882e-05
Intercept7  2.471763e-03  3.805730e-05  1.056961e-04
Intercept8  3.805730e-05  1.032978e-03  1.102280e-04
Intercept9  1.056961e-04  1.102280e-04  1.882443e-03
Intercept10 1.192793e-05  1.889368e-05  3.490300e-05
      Intercept10
Intercept1  2.837357e-06
Intercept2  3.261207e-06
Intercept3  1.324198e-05
Intercept4 -7.762461e-06
Intercept5  3.784392e-06
Intercept6  3.451132e-05
Intercept7  1.192793e-05
Intercept8  1.889368e-05
Intercept9  3.490300e-05
Intercept10 1.023027e-03

```

```

R> ##### Extract the random-effects (variance component)
R> coef(random1, select="random")

```

```

      Tau2_1_1  Tau2_2_2  Tau2_3_3  Tau2_4_4  Tau2_5_5
0.03815841 0.02132562 0.05115980 0.02571721 0.01901262
      Tau2_6_6  Tau2_7_7  Tau2_8_8  Tau2_9_9  Tau2_10_10
0.01977637 0.02995567 0.01030039 0.02172535 0.01122090

```

```

R> random2 <- tssem2(random1, Amatrix=A1, Smatrix=S1, Fmatrix=F1, diag.constraints=TRUE, i

```

Running TSSEM2 (Random Effects Model) Analysis of Correlation Structure

```

R> summary(random2)

```

Call:

```

wls(Cov = pooledS, asyCov = asyCov, n = tssem1.obj$total.n, Amatrix = Amatrix,
    Smatrix = Smatrix, Fmatrix = Fmatrix, diag.constraints = diag.constraints,
    cor.analysis = cor.analysis, intervals.type = intervals.type,
    model.name = model.name, suppressWarnings = suppressWarnings)

```

95% confidence intervals: Likelihood-based statistic  
Coefficients:

	Estimate	Std.Error	lbound	ubound	z	value
Amatrix[1,7]	0.36594	NA	0.19372	0.49145	NA	
Amatrix[2,6]	0.52776	NA	0.44394	0.61420	NA	
Amatrix[3,6]	0.58872	NA	0.50989	0.67281	NA	
Amatrix[4,6]	0.48963	NA	0.41993	0.56199	NA	
Amatrix[5,7]	0.34919	NA	0.18457	0.46998	NA	
Smatrix[1,1]	0.86609	NA	0.75848	0.96235	NA	
Smatrix[2,2]	0.72147	NA	0.62274	0.80291	NA	
Smatrix[3,3]	0.65341	NA	0.54730	0.74001	NA	
Smatrix[4,4]	0.76026	NA	0.68417	0.82366	NA	
Smatrix[5,5]	0.87806	NA	0.77913	0.96583	NA	
Smatrix[7,6]	1.76399	NA	1.30134	3.49213	NA	

	Pr(> z )
Amatrix[1,7]	NA
Amatrix[2,6]	NA
Amatrix[3,6]	NA
Amatrix[4,6]	NA
Amatrix[5,7]	NA
Smatrix[1,1]	NA
Smatrix[2,2]	NA
Smatrix[3,3]	NA
Smatrix[4,4]	NA
Smatrix[5,5]	NA
Smatrix[7,6]	NA

Goodness-of-fit indices:

	Value
Sample size	4496.0000
Chi-square of target model	81.2723
DF of target model	4.0000
p value of target model	0.0000
Number of constraints imposed on "Smatrix"	5.0000
DF manually adjusted	0.0000
Chi-square of independent model	546.8082
DF of independent model	10.0000
RMSEA	0.0656
SRMR	0.1323
TLI	0.6401
CFI	0.8561
AIC	73.2723
BIC	47.6285

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

## 4. Other Useful Functions

### 4.1. Analysis of Correlation/Covariance Structure with Weighted Least Squares

The `wls()` function may be used to fit a correlation/covariance structure with weighted least squares (WLS) estimation method. The following example fits a one-factor CFA model on the correlation matrix with WLS estimation method. It should be noted that the only off-diagonal elements are used when a correlation structure is fitted.

```
R> ##### Analysis of correlation structure
R> R1 <- matrix(c(1.00, 0.22, 0.24, 0.18,
                  0.22, 1.00, 0.30, 0.22,
                  0.24, 0.30, 1.00, 0.24,
                  0.18, 0.22, 0.24, 1.00), ncol=4, nrow=4)

R> n <- 1000
R> acovR1 <- asyCov(R1, n)
R> ## One-factor CFA model
R> (A1 <- cbind(matrix(0, nrow=5, ncol=4),
                matrix(c("0.2*a1", "0.2*a2", "0.2*a3", "0.2*a4", 0),
                        ncol=1)))

      [,1] [,2] [,3] [,4] [,5]
[1,] "0"  "0"  "0"  "0"  "0.2*a1"
[2,] "0"  "0"  "0"  "0"  "0.2*a2"
[3,] "0"  "0"  "0"  "0"  "0.2*a3"
[4,] "0"  "0"  "0"  "0"  "0.2*a4"
[5,] "0"  "0"  "0"  "0"  "0"

R> A1 <- as.mxMatrix(A1)
R> (S1 <- diag(c("0.2*e1", "0.2*e2", "0.2*e3", "0.2*e4", 1)))

      [,1]      [,2]      [,3]      [,4]      [,5]
[1,] "0.2*e1" "0"      "0"      "0"      "0"
[2,] "0"      "0.2*e2" "0"      "0"      "0"
[3,] "0"      "0"      "0.2*e3" "0"      "0"
[4,] "0"      "0"      "0"      "0.2*e4" "0"
[5,] "0"      "0"      "0"      "0"      "1"

R> S1 <- as.mxMatrix(S1)
R> ## The first 4 variables are observed while the last one is latent.
R> (F1 <- create.Fmatrix(c(1,1,1,1,0), name="F1"))

FullMatrix 'F1'

@labels: No labels assigned.
```



@values

	[,1]	[,2]	[,3]	[,4]	[,5]
[1,]	1	0	0	0	0
[2,]	0	1	0	0	0
[3,]	0	0	1	0	0
[4,]	0	0	0	1	0

@free: No free parameters.

@lbound: No lower bounds assigned.

@ubound: No upper bounds assigned.

```
R> wls.fit1 <- wls(Cov=R1, asyCov=acovR1, n=n, Fmatrix=F1, Smatrix=S1, Amatrix=A1,
  cor.analysis=TRUE, diag.constraints=TRUE, intervals="LB")
```

Running WLS Analysis of Correlation Structure

```
R> summary(wls.fit1)
```

Call:

```
wls(Cov = R1, asyCov = acovR1, n = n, Amatrix = A1, Smatrix = S1,
  Fmatrix = F1, diag.constraints = TRUE, cor.analysis = TRUE,
  intervals.type = "LB")
```

95% confidence intervals: Likelihood-based statistic

Coefficients:

	Estimate	Std.Error	lbound	ubound	z value
Amatrix[1,5]	0.42159	NA	0.34632	0.49869	NA
Amatrix[2,5]	0.52376	NA	0.44829	0.60309	NA
Amatrix[3,5]	0.57092	NA	0.49431	0.65292	NA
Amatrix[4,5]	0.42159	NA	0.34632	0.49869	NA
Smatrix[1,1]	0.82226	NA	0.75131	0.88005	NA
Smatrix[2,2]	0.72567	NA	0.63627	0.79903	NA
Smatrix[3,3]	0.67405	NA	0.57367	0.75566	NA
Smatrix[4,4]	0.82226	NA	0.75131	0.88006	NA

Pr(>|z|)

Amatrix[1,5]	NA
Amatrix[2,5]	NA
Amatrix[3,5]	NA
Amatrix[4,5]	NA
Smatrix[1,1]	NA
Smatrix[2,2]	NA
Smatrix[3,3]	NA
Smatrix[4,4]	NA

Goodness-of-fit indices:

	Value
Sample size	1000.0000
Chi-square of target model	0.0134
DF of target model	2.0000
p value of target model	0.9933
Number of constraints imposed on "Smatrix"	4.0000
DF manually adjusted	0.0000
Chi-square of independent model	243.9826
DF of independent model	6.0000
RMSEA	0.0000
SRMR	0.0012
TLI	1.0250
CFI	1.0000
AIC	-3.9866
BIC	-13.8021

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Extract the parameter estimates
R> coef(wls.fit1)
```

```
Amatrix[1,5] Amatrix[2,5] Amatrix[3,5] Amatrix[4,5]
  0.4215923    0.5237644    0.5709210    0.4215923
Smatrix[1,1] Smatrix[2,2] Smatrix[3,3] Smatrix[4,4]
  0.8222599    0.7256709    0.6740492    0.8222599
```

```
R> ## Extract the sampling variance covariance matrix
R> vcov(wls.fit1)
```

```
[1] NA
```

## 4.2. Likelihood-based Confidence Intervals

Most CIs are based on the estimated standard errors. These are known as Wald CIs. Wald CIs are symmetric around the estimates. The Wald CIs might be outside of the meaningful boundaries, for example, a negative lower limit for the variance or larger than 1 for a correlation coefficient. A preferable approach is to construct the CIs based on the likelihood. This is known as the likelihood based CI (Cheung 2009a; Neale and Miller 1997). Likelihood based CIs on the parameter estimates can be required by specifying `intervals.type='LB'` argument.

```
R> ## Random-effects meta-analysis with ML
R> summary( meta(y=di, v=vi, data=Becker83, intervals.type="LB") )
```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, data = Becker83, intervals.type = "LB")
```

95% confidence intervals: Likelihood-based statistic

Coefficients:

	Estimate	Std.Error	lbound	ubound	z value
Intercept1	0.174734	0.113378	-0.052165	0.437627	1.5412
Tau2_1_1	0.077376	0.054108	0.015124	0.302999	1.4300

Pr(>|z|)

Intercept1	0.1233
------------	--------

Tau2_1_1	0.1527
----------	--------

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Heterogeneity indices (I2) and their 95% likelihood-based CIs:

	lbound	Estimate	ubound
Intercept1: I2 (Q statistic)	0.28410	0.67182	0.8888

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 2

Degrees of freedom: 8

-2 log likelihood: 7.928307

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

R> ## Mixed-effects meta-analysis with "log(items)" as a predictor

R> summary( meta(y=di, v=vi, x=log(items), data=Becker83, intervals.type="LB") )

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, x = log(items), data = Becker83, intervals.type = "LB")
```

95% confidence intervals: Likelihood-based statistic

Coefficients:

	Estimate	Std.Error	lbound	ubound
Intercept1	-3.2015e-01	1.0981e-01	-5.4408e-01	-7.7598e-02
Slope1_1	2.1088e-01	4.5084e-02	1.1838e-01	3.0789e-01
Tau2_1_1	1.0000e-10	2.0095e-02	9.9937e-11	5.7947e-02

z value Pr(>|z|)

Intercept1	-2.9154	0.003552	**
------------	---------	----------	----

Slope1_1	4.6774	2.905e-06	***
----------	--------	-----------	-----

Tau2_1_1	0.0000	1.000000	
----------	--------	----------	--

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Explained variances (R2):

	Estimate
y1: Tau2 (no predictor)	0.0774
y1: Tau2 (with predictors)	0.0000
y1: R2	1.0000

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 3

Degrees of freedom: 7

-2 log likelihood: -4.208024

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

### 4.3. Reading External Data Files

Data sets are most likely stored externally. `metaSEM` reads three types of data formats. The first type is full correlation/covariance matrices, for example, `fullmat.dat` is the same as the built-in data set `Cheung09`. Missing values are represented by `NA` (the default option). Suppose you have saved it at `d:\fullmat.dat`, you may read it by using the following command in R:

```
my.df <- readFullMat(file="d:/fullmat.dat")
```

The second type is lower triangle correlation/covariance matrices, for example, `lowertriangle.dat`. Missing values are represented by the strings `NA`. Suppose you have saved it at `d:\lowertriangle.dat`, you may read it by using the following command in R:

```
my.df <- readLowTriMat(file = "d:/lowertriangle.dat", no.var = 9, na.strings="NA")
```

The third type is vectors of correlation/covariance elements based on column vectorization. One row represents one study, for example, `stackvec.dat`. Suppose you have saved it at `d:\stackvec.dat`, you may read it by using the following R command:

```
my.df <- readStackVec(file="d:/stackvec.dat")
```

## 5. Installation

First of all, you need R to run it. Since `metaSEM` uses `OpenMx` as the workhorse, `OpenMx` has to be installed first. To install `OpenMx`, run the following command inside an R session:

```
install.packages('OpenMx', repos='http://openmx.psyc.virginia.edu/packages/')
```

See <http://openmx.psyc.virginia.edu/installing-openmx> for the details on how to install OpenMx. Moreover, metaSEM also depends on the ellipse package that can be installed by the following command inside an R session:

```
install.packages('ellipse')
```

### 5.1. Windows platform

Download the **Windows binary** of metaSEM. If the file is saved at d:\. Run the following command inside an R session:

```
install.packages(pkgs="d:/metaSEM_0.7-1.zip", repos=NULL)
```

Please note that d:\ in Windows is represented by either d:/ or d:\\ in R.

### 5.2. Linux and Mac OS X platform

Download the **source package** of metaSEM. Run the following command (as Root) inside an R session:

```
install.packages(pkgs="metaSEM_0.7-1.zip", repos=NULL, type="source")
```

## 6. Acknowledgements

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