

metaSEM: An R Package for Meta-Analysis Using Structural Equation Modeling

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Abstract

The **metaSEM** package provides functions to conducting univariate and multivariate meta-analysis using a structural equation modeling approach via the **OpenMx** package. It also implemented the two-stage structural equation modeling (TSSEM) approach (Cheung and Chan 2005b, 2009) to conducting fixed- and random-effects meta-analytic structural equation modeling (MASEM) on correlation/covariance matrices. This paper outlines the basic theories. Examples are used to illustrate the procedures.

Keywords: meta-analysis, structural equation modeling, meta-analytic structural equation modeling, **metaSEM**, R.

1. Introduction

Meta-analysis is a popular technique to synthesizing research findings in social, behavioral, educational and medical sciences (Borenstein, Hedges, Higgins, and Rothstein 2009; Hedges and Olkin 1985; Hunter and Schmidt 2004; Whitehead 2002). There are several standalone packages for meta-analysis, e.g., Comprehensive Meta-Analysis and RevMan. Many standard statistical packages, for instance, SPSS (Lipsey and Wilson 2000), SAS (Arthur, Bennett, and Huffcutt 2001) and STATA (Sterne 2009), have macros or packages to fitting some meta-analytic models. Even in the R community, there are already several packages to conducting meta-analysis, for instance, **meta** (Schwarzer 2010), **rmeta** (Lumley 2009), **mvmeta** (Gasparini 2012), **metaLik** (Guolo and Varin 2011) and **metafor** (Viechtbauer 2010).

The **metaSEM** package is yet another R package to conducting univariate and multivariate meta-analysis. It formulates meta-analytic models as structural equation models (Cheung 2008, 2011b) via the **OpenMx** package (Boker, Neale, Maes, Wilde, Spiegel, Brick, Spies, Estabrook, Kenny, Bates, Mehta, and Fox 2011). It also implemented the two-stage structural equation modeling (TSSEM) approach (Cheung and Chan 2005b, 2009) to conducting fixed- and random-effects meta-analytic structural equation modeling (MASEM) on correlation/covariance matrices. The main functions in this package are:

- **meta()** and **reml()**: **meta()** fits univariate and multivariate meta-analysis with maximum likelihood (ML) estimation method while **reml()** estimates the variance components of the random-effects with restricted (residual) maximum likelihood (REML)

estimation method. Mixed-effects meta-analysis can be fitted by including study characteristics as predictors. Equality constraints on the intercepts, regression coefficients and variance components can be imposed.

- **meta3()**: It fits 3-level meta-analysis by considering cluster effect.
- **tssem1()**: It fits the first stage analysis of TSSEM by pooling correlation/covariance matrices with either a fixed- or random-effects model.
- **tssem2()**: It fits the second stage analysis of TSSEM by fitting structural models on the pooled correlation/covariance matrix. It is a wrapper of **wls()**.
- **wls()**: It fits a correlation/covariance structure analysis with weighted least squares (WLS) estimation method.

Besides reporting Wald confidence intervals (CIs) based on z statistic, likelihood-based CIs on the parameter estimates may also be requested (Cheung 2009a; Neale and Miller 1997). Several generic functions, such as **anova()**, **coef()**, **vcov()**, **print()**, **summary()** and **plot()**, have been implemented.

This paper was based on the **metaSEM** package version 0.7-1 and the **OpenMx** package version 1.2.3-2011. The paper is organized as follows. The next section introduces general meta-analytic models. Basic theory of the TSSEM are then presented. Several examples are used to illustrate these procedures.

2. Structural Equation Modeling Based Meta-Analysis

In this section, basic structural equation models are introduced. Univariate and multivariate meta-analysis are treated as special cases of SEM (Cheung 2008, 2011b).

2.1. Structural equation model

Structural equation modeling is a multivariate technique to fitting and testing hypothesized models. Let \mathbf{y} be a $p \times 1$ vector of the sample data where p is the number of variables. It is hypothesized that the model for the first and second moments are $\mu = \mu(\theta)$ and $\Sigma = \Sigma(\theta)$, respectively, where θ is a vector of parameters.

The $-2 \times \log$ -likelihood of the i th case is:

$$-2 * \log L_i(\theta; \mathbf{y}_i)_{\text{ML}} = p_i * \log(2\pi) + \log |\Sigma_i(\theta)| + (\mathbf{y}_i - \mu_i(\theta))' \Sigma_i(\theta)^{-1} (\mathbf{y}_i - \mu_i(\theta)), \quad (1)$$

where p_i is the number of variables in the i th case, $\mu_i(\theta)$ and $\Sigma_i(\theta)$ are the model implied mean vector and the model implied covariance matrix for the i th case, respectively. Since there is a subscript i in these quantities, the model implied mean vector and covariance matrix may vary across cases. In other words, this model handles incomplete data automatically by selecting only the complete variables in the log-likelihood function.

To obtain the parameter estimates, we may take the sum of the $-2 \times \log$ -likelihood over all cases and minimize it. This is known as the ML estimation method. After the optimization, the asymptotic covariance matrix (thus the standard errors) of the parameter estimates may be obtained from the inverse of the Hessian matrix. The parameter estimates divided by their

standard errors follow a z distribution under the null hypothesis. Moreover, likelihood ratio statistic may also be used to compare nested models.

2.2. Univariate fixed-effects model

When there is only one effect size, the univariate fixed-effects model for the i th study is:

$$y_i = \beta_{\text{fixed}} + e_i, \quad (2)$$

where β_{fixed} is the common effect under a fixed-effects model and $\text{var}(e_i) = v_i$ is the known sampling variance. To fit the univariate fixed-effects meta-analysis in SEM, we may use the following model:

$$\mu_i(\theta) = \beta_{\text{fixed}} \quad (3)$$

and

$$\Sigma_i(\theta) = v_i \quad (4)$$

Since v_i is known, the only parameter in the univariate fixed-effects model is β_{fixed} .

2.3. Univariate random-effects model

A random-effects model allows studies having their own study specific effect. The model for the i th study is:

$$y_i = \beta_{\text{random}} + u_i + e_i. \quad (5)$$

where β_{random} is the average effect under a random-effects model and $\text{var}(u_i) = \tau^2$ is the heterogeneity variance that has to be estimated. To fit the univariate fixed-effects meta-analysis in SEM, we may use the following model:

$$\mu_i(\theta) = \beta_{\text{random}} \quad (6)$$

and

$$\Sigma_i(\theta) = \tau^2 + v_i \quad (7)$$

In this model we have to estimate both β_{random} and τ^2 .

2.4. Univariate mixed-effects model

The mixed-effects meta-analysis extends the random-effects meta-analysis by including predictors. Assuming that \mathbf{x}_i is a $m \times 1$ vector of predictors where m is the number predictors in the i th study, the model is:

$$y_i = \beta_0 + \beta' \mathbf{x}_i + u_i + e_i, \quad (8)$$

where β is a vector of regression coefficients.

To fit the univariate mixed-effects meta-analysis in SEM, we may use the following model:

$$\mu_i(\theta|\mathbf{x}_i) = \beta_0 + \beta' \mathbf{x}_i \quad (9)$$

and

$$\Sigma_i(\theta|\mathbf{x}_i) = \tau^2 + v_i. \quad (10)$$

Since \mathbf{x}_i is specified via definition variables, the means and covariance matrix of \mathbf{x} are not estimated. That is, \mathbf{x} is treated as a design matrix rather than a random variable.

2.5. Multivariate mixed-effects model

Let us assume that there are p effect sizes with m predictors in k studies. The model for the multivariate effect sizes in the i th study is:

$$\mathbf{y}_i = \mathbf{B}\mathbf{x}_i + \mathbf{u}_i + \mathbf{e}_i, \quad (11)$$

where \mathbf{y}_i is a $p \times 1$ effect sizes, \mathbf{B} is a $p \times (m+1)$ regression coefficients including the intercepts, \mathbf{x}_i is a $(m+1) \times 1$ predictors including 1 in the first column, \mathbf{u}_i is a $p \times 1$ study specific random effects, and \mathbf{e}_i is a $p \times 1$ sampling error. We assume that $\text{var}(\mathbf{e}_i) = V_i$ is known and given in the i th study and $\text{var}(\mathbf{u}_i) = T^2$ is the variance component of the between-study heterogeneity that has to be estimated.

The -2*log-likelihood of the above model is:

$$-2 * \log L_i(\mathbf{B}, T^2; \mathbf{y}_i)_{\text{ML}} = p_i * \log(2\pi) + \log|T^2 + V_i| + (\mathbf{y}_i - \mathbf{B}\mathbf{x}_i)'(T^2 + V_i)^{-1}(\mathbf{y}_i - \mathbf{B}\mathbf{x}_i), \quad (12)$$

where p_i is the number of effect sizes in the i th study.

In applied research, different studies may report different effect sizes, that is, p_i may vary across studies. The above -2*log-likelihood may handle missing effect sizes by using different dimensions of the elements in the above equation. It is expected that there is no missing data in \mathbf{x}_i . When there are missing data in \mathbf{x}_i , the whole study will be deleted before the analysis.

2.6. Restricted Maximum Likelihood (REML) Estimation Method

Since both the fixed- and random-effects are estimated simultaneously, it is well-known that \hat{T}_{ML}^2 based on the ML estimation is under-estimated. It is because it does not take the uncertainty in estimating $\hat{\mathbf{B}}_{\text{ML}}$ into account. If the unbiasedness of the variance component is crucial to the research questions, it is possible to obtain the variance component \hat{T}_{REML}^2 based on the REML estimation method (Cheung 2011a; Harville 1977; Patterson and Thompson 1971).

The -2log-likelihood of the model is:

$$-2 \log L_i(T^2; \mathbf{y}_i)_{\text{REML}} = p_i * \log(2\pi) + \log|T^2 + V_i| + (\mathbf{y}_i - \alpha \mathbf{X}_i)'(T^2 + V_i)^{-1}(\mathbf{y}_i - \alpha \mathbf{X}_i) + |X_i' V_i^{-1} X_i|, \quad (13)$$

where $\alpha = (X' V^{-1} X)^{-1} X' V^{-1} \mathbf{y}$.

Since the fixed effects \mathbf{B} is not involved in the above -2log-likelihood function, it has to be calculated in a second stage.

2.7. Three-level meta-analysis

Observed effect sizes may be dependent. For example, effect sizes reported by the same research team may be more similar when comparing to effect sizes reported by other teams. Effect sizes reported by studies from the same country may be more similar when comparing to studies across countries. If the degree of dependence is known, multivariate meta-analysis as

introduced before may be applied. When the degree of dependence is unknown, a three-level meta-analytic model may be used (Konstantopoulos 2011). The model is:

$$y_i = \beta_0 + \beta' \mathbf{x}_i + u_{(2)i} + u_{(3)i} + e_i, \quad (14)$$

where $u_{(2)i}$ and $u_{(3)i}$ are the random-effects at level-2 and level-3, respectively.

To fit the three-level meta-analytic model in SEM, we may use the following model:

$$\mu_i(\theta|\mathbf{x}_i) = \beta_0 + \beta' \mathbf{x}_i \quad (15)$$

and

$$\Sigma_i(\theta|\mathbf{x}_i) = \tau_{(2)}^2 + \tau_{(3)}^2 + v_i. \quad (16)$$

where $\tau_{(2)}^2 = \text{var}(u_{(2)i})$ and $\tau_{(3)}^2 = \text{var}(u_{(3)i})$ are the heterogeneity at level-2 and level-3, respectively.

2.8. Examples

Two example data sets are used to demonstrate the procedures of fitting univariate and multivariate meta-analyses. The first data set was taken from Becker (1983) who reported 10 studies on sex differences in conformity using the fictitious norm group paradigm. di and vi are the standardized mean difference and its sampling variance, respectively. Becker hypothesized that the logarithm of the number of items (*items*) predicted the effect size.

The second data set is adapted from Berkey, Hoaglin, Antczak-Bouckoms, Mosteller, and Colditz (1998). They summarized five published trials comparing surgical and non-surgical treatments for medium-severity periodontal disease, one year after treatment. Publication year *pub_year* was hypothesized as a predictor.

Univariate random-effects model The function `meta()` is used to conduct the analyses. The arguments `y` and `v` are used to specify the effect sizes and its sampling variances (and covariances for multivariate meta-analysis), respectively. By default, a random-effects meta-analysis is fitted. After running the analysis, `summary()` may be used to report the results. The estimated fixed- and random-effects are represented by the **Intercept** and **Tau2** parameters. `coef()` and `vcov()` may be used to extract the parameter estimates and their asymptotic sampling covariance matrix, respectively.

From the following analyses, the Q statistic ($df = 9$) is 30.6495, $p < .001$. The pooled effect size with its 95% Wald confidence interval (CI) based on the random-effects model is 0.1747 (-0.0475, 0.3970). The estimated heterogeneity variance is 0.0774.

```
R> ## Load the library
R> library(metaSEM)
R> ## Show the first few studies of the data set
R> head(Becker83)
```

| | study | di | vi | percentage | items |
|---|-------|-------|------|------------|-------|
| 1 | 1 | -0.33 | 0.03 | 25 | 2 |
| 2 | 2 | 0.07 | 0.03 | 25 | 2 |

| | | | | | |
|---|---|-------|------|-----|----|
| 3 | 3 | -0.30 | 0.02 | 50 | 2 |
| 4 | 4 | 0.35 | 0.02 | 100 | 38 |
| 5 | 5 | 0.69 | 0.07 | 100 | 30 |
| 6 | 6 | 0.81 | 0.22 | 100 | 45 |

```
R> ## Random-effects meta-analysis with ML
R> summary( random1 <- meta(y=di, v=vi, data=Becker83) )
```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z value |
|------------|----------|-----------|-----------|----------|---------|
| Intercept1 | 0.174734 | 0.113378 | -0.047482 | 0.396950 | 1.5412 |
| Tau2_1_1 | 0.077376 | 0.054108 | -0.028674 | 0.183426 | 1.4300 |

Pr(>|z|)

| | |
|------------|--------|
| Intercept1 | 0.1233 |
| Tau2_1_1 | 0.1527 |

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 2

Degrees of freedom: 8

-2 log likelihood: 7.928307

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Exact the coefficients
```

```
R> coef(random1)
```

| | |
|------------|------------|
| Intercept1 | Tau2_1_1 |
| 0.17473402 | 0.07737594 |

```
R> ## Exact the sampling variance covariance matrix
```

```
R> vcov(random1)
```

| | Intercept1 | Tau2_1_1 |
|------------|-------------|-------------|
| Intercept1 | 0.012854471 | 0.001240975 |
| Tau2_1_1 | 0.001240975 | 0.002927667 |

Univariate mixed-effects model We may include a predictor to conduct a mixed-effects meta-analysis. The argument `x` is used to specify the predictors. If there are more than one predictor, `cbind()` may be used to specify them. The estimated regression coefficients are represented by the **Slope** parameter. The result suggests that *log(items)* is a significant predictor with the estimated regression coefficient and its 95% CI of 0.2109 (0.1225, 2.9924).

```
R> ## Mixed-effects meta-analysis with "log(items)" as the predictor
R> summary( mixed1 <- meta(y=di, v=vi, x=log(items), data=Becker83) )
```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, x = log(items), data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound |
|------------|-------------|------------|-------------|-------------|
| Intercept1 | -3.2015e-01 | 1.0981e-01 | -5.3539e-01 | -1.0492e-01 |
| Slope1_1 | 2.1088e-01 | 4.5084e-02 | 1.2251e-01 | 2.9924e-01 |
| Tau2_1_1 | 1.0000e-10 | 2.0095e-02 | -3.9386e-02 | 3.9386e-02 |

| | z value | Pr(> z) |
|------------|---------|---------------|
| Intercept1 | -2.9154 | 0.003552 ** |
| Slope1_1 | 4.6774 | 2.905e-06 *** |
| Tau2_1_1 | 0.0000 | 1.000000 |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 3

Degrees of freedom: 7

-2 log likelihood: -4.208024

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Exact the coefficients
```

```
R> coef(mixed1)
```

| Intercept1 | Slope1_1 | Tau2_1_1 |
|---------------|--------------|--------------|
| -0.3201549197 | 0.2108782268 | 0.0000000001 |

```
R> ## Exact the sampling variance covariance matrix
```

```
R> vcov(mixed1)
```

| | Intercept1 | Slope1_1 | Tau2_1_1 |
|------------|---------------|---------------|--------------|
| Intercept1 | 0.0120593278 | -3.940014e-03 | 3.887682e-04 |
| Slope1_1 | -0.0039400143 | 2.032587e-03 | 8.408773e-05 |
| Tau2_1_1 | 0.0003887682 | 8.408773e-05 | 4.038148e-04 |

Univariate fixed-effects model Mathematically, fixed-effects meta-analysis is a special case of the random-effects meta-analysis by fixing the variance of the random-effects at 0. The argument `RE.constraints`, which expects a matrix as input, is used to constrain the variance component of the random effects.

```
R> ## Fixed-effects meta-analysis
R> summary( fixed1 <- meta(y=di, v=vi, data=Becker83,
                          RE.constraints=matrix(0, ncol=1, nrow=1)) )
```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, data = Becker83, RE.constraints = matrix(0,
  ncol = 1, nrow = 1))
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z value |
|------------|----------|-----------|-----------|----------|---------|
| Intercept1 | 0.100640 | 0.060510 | -0.017957 | 0.219237 | 1.6632 |
| | Pr(> z) | | | | |
| Intercept1 | 0.09627 | . | | | |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 1

Degrees of freedom: 9

-2 log likelihood: 17.86043

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Exact the coefficients
```

```
R> coef(fixed1)
```

```
Intercept1
0.1006404
```



```
R> ## Exact the sampling variance covariance matrix
R> vcov(fixed1)
```

```
Intercept1
Intercept1 0.003661436
```

Multivariate random-effects model Multivariate meta-analysis can be fitted by specifying the multivariate effect sizes and its sampling covariance matrix in the arguments `y` and `v` with `cbind()`, respectively. Only the lower triangle of the sampling covariance matrix arranged by the column major is used in `v`.

The Q statistic ($df = 8$) of the following example is 128.2267, $p < .001$. The pooled effect sizes with their 95% Wald CIs based on the random-effects model for *PD* and *AL* are 0.3448 (0.2397, 0.4500) and -0.3379 (-0.4972, -0.1787), respectively. The estimated variance component is $\begin{bmatrix} 0.0070 & \\ 0.0095 & 0.02614 \end{bmatrix}$.

```
R> ## Show the data set
R> Berkey98
```

| | trial | pub_year | no_of_patients | PD | AL | var_PD | cov_PD_AL |
|---|-------|----------|----------------|------|-------|--------|-----------|
| 1 | 1 | 1983 | 14 | 0.47 | -0.32 | 0.0075 | 0.0030 |
| 2 | 2 | 1982 | 15 | 0.20 | -0.60 | 0.0057 | 0.0009 |
| 3 | 3 | 1979 | 78 | 0.40 | -0.12 | 0.0021 | 0.0007 |
| 4 | 4 | 1987 | 89 | 0.26 | -0.31 | 0.0029 | 0.0009 |
| 5 | 5 | 1988 | 16 | 0.56 | -0.39 | 0.0148 | 0.0072 |

| | var_AL |
|---|--------|
| 1 | 0.0077 |
| 2 | 0.0008 |
| 3 | 0.0014 |
| 4 | 0.0015 |
| 5 | 0.0304 |

```
R> ## Multivariate meta-analysis with a random-effects model
R> summary( mult1 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL),
                        data=Berkey98) )
```

Running Meta analysis with ML

Call:

```
meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
     data = Berkey98)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound |
|------------|-----------|-----------|-----------|-----------|
| Intercept1 | 0.3448392 | 0.0536312 | 0.2397239 | 0.4499544 |

```

Intercept2 -0.3379381  0.0812480 -0.4971812 -0.1786951
Tau2_1_1    0.0070020  0.0090497 -0.0107351  0.0247391
Tau2_2_1    0.0094607  0.0099698 -0.0100797  0.0290010
Tau2_2_2    0.0261445  0.0177409 -0.0086270  0.0609161

```

```

      z value  Pr(>|z|)

```

```

Intercept1  6.4298 1.278e-10 ***
Intercept2 -4.1593 3.192e-05 ***
Tau2_1_1    0.7737    0.4391
Tau2_2_1    0.9489    0.3427
Tau2_2_2    1.4737    0.1406

```

```

---

```

```

Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

```

Q statistic on homogeneity of effect sizes: 128.2267

```

```

Degrees of freedom of the Q statistic: 8

```

```

P value of the Q statistic: 0

```

```

Number of studies (or clusters): 5

```

```

Number of observed statistics: 10

```

```

Number of estimated parameters: 5

```

```

Degrees of freedom: 5

```

```

-2 log likelihood: -11.68131

```

```

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

```

R> ## Exact the coefficients

```

```

R> coef(mult1)

```

```

Intercept1 Intercept2 Tau2_1_1 Tau2_2_1
0.344839167 -0.337938117 0.007001998 0.009460665
Tau2_2_2
0.026144517

```

```

R> ## Exact the sampling variance covariance matrix

```

```

R> vcov(mult1)

```

```

Intercept1 Intercept2 Tau2_1_1
Intercept1 2.876307e-03 2.215623e-03 1.241471e-04
Intercept2 2.215623e-03 6.601230e-03 6.203168e-05
Tau2_1_1    1.241471e-04 6.203168e-05 8.189758e-05
Tau2_2_1    -1.684464e-05 1.220432e-04 5.825516e-05
Tau2_2_2    1.315804e-06 3.246241e-05 4.503124e-05
Tau2_2_1    Tau2_2_2
Intercept1 -1.684464e-05 1.315804e-06
Intercept2 1.220432e-04 3.246241e-05
Tau2_1_1    5.825516e-05 4.503124e-05
Tau2_2_1    9.939612e-05 1.187750e-04
Tau2_2_2    1.187750e-04 3.147402e-04

```

Multivariate mixed-effects model As an illustration, we use *pub_year* as a predictor. To make the intercept more interpretable, we center the publication year at 1979, the first record of publication year in the data set. The estimated regression coefficients and their 95% CIs on *PD* and *AL* are 0.0064 (-0.2048, 0.2177) and -0.0706 (-0.3883, 0.2471), respectively. The likelihood ratio statistic on testing both regression coefficient is $\chi^2(df = 2) = 0.3273, p = .8490$. Thus, both regression coefficients are non-significant.

Sometimes, we may want to test the equality of the regression coefficients and see if they are different. We may impose the equality constraint on the regression coefficients with the argument `coef.constraints`. The average regression coefficient is 0.0017 (-0.1991, 0.2025). The likelihood ratio statistic on testing the equality of the regression coefficients is $\chi^2(df = 1) = 0.3270, p = .5674$. There is no evidence that one regression coefficient is stronger from the other.

```
R> ## Multivariate meta-analysis with "publication year-1979" as a predictor
R> mult2 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL), data=Berkey98,
               x=scale(pub_year, center=1979), model.name="No constraint")
```

Running No constraint

```
R> summary(mult2)
```

Call:

```
meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
     x = scale(pub_year, center = 1979), data = Berkey98, model.name = "No constraint")
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound |
|------------|------------|-----------|------------|------------|
| Intercept1 | 0.3440001 | 0.0857659 | 0.1759020 | 0.5120982 |
| Slope1_1 | 0.0063540 | 0.1078235 | -0.2049761 | 0.2176842 |
| Intercept2 | -0.2918175 | 0.1312797 | -0.5491208 | -0.0345141 |
| Slope2_1 | -0.0705887 | 0.1620966 | -0.3882922 | 0.2471147 |
| Tau2_1_1 | 0.0080405 | 0.0101206 | -0.0117955 | 0.0278766 |
| Tau2_2_1 | 0.0093413 | 0.0105515 | -0.0113392 | 0.0300218 |
| Tau2_2_2 | 0.0250135 | 0.0170788 | -0.0084603 | 0.0584873 |

| | z value | Pr(> z) |
|------------|---------|---------------|
| Intercept1 | 4.0109 | 6.048e-05 *** |
| Slope1_1 | 0.0589 | 0.95301 |
| Intercept2 | -2.2229 | 0.02622 * |
| Slope2_1 | -0.4355 | 0.66322 |
| Tau2_1_1 | 0.7945 | 0.42692 |
| Tau2_2_1 | 0.8853 | 0.37599 |
| Tau2_2_2 | 1.4646 | 0.14303 |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 128.2267

Degrees of freedom of the Q statistic: 8

P value of the Q statistic: 0

Number of studies (or clusters): 5

Number of observed statistics: 10

Number of estimated parameters: 7

Degrees of freedom: 3

-2 log likelihood: -12.00859

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

R> ## Exact the coefficients

R> coef(mult2)

| Intercept1 | Intercept2 | Slope1_1 | Slope2_1 |
|-------------|--------------|-------------|--------------|
| 0.344000113 | -0.291817459 | 0.006354034 | -0.070588746 |
| Tau2_1_1 | Tau2_2_1 | Tau2_2_2 | |
| 0.008040540 | 0.009341319 | 0.025013487 | |

R> ## Exact the sampling variance covariance matrix

R> vcov(mult2)

| | Intercept1 | Intercept2 | Slope1_1 |
|------------|---------------|---------------|---------------|
| Intercept1 | 7.355789e-03 | 6.628461e-03 | -7.148841e-03 |
| Intercept2 | 6.628461e-03 | 1.723435e-02 | -6.290031e-03 |
| Slope1_1 | -7.148841e-03 | -6.290031e-03 | 1.162591e-02 |
| Slope2_1 | -6.692079e-03 | -1.692509e-02 | 9.550531e-03 |
| Tau2_1_1 | -1.288708e-04 | -2.379371e-05 | 3.800441e-04 |
| Tau2_2_1 | -1.307349e-04 | -5.734270e-05 | 1.944444e-04 |
| Tau2_2_2 | -2.517793e-05 | -4.989090e-06 | 4.768716e-05 |
| | Slope2_1 | Tau2_1_1 | Tau2_2_1 |
| Intercept1 | -0.0066920794 | -1.288708e-04 | -1.307349e-04 |
| Intercept2 | -0.0169250890 | -2.379371e-05 | -5.734270e-05 |
| Slope1_1 | 0.0095505309 | 3.800441e-04 | 1.944444e-04 |
| Slope2_1 | 0.0262752958 | 1.400374e-04 | 2.907864e-04 |
| Tau2_1_1 | 0.0001400374 | 1.024271e-04 | 7.161024e-05 |
| Tau2_2_1 | 0.0002907864 | 7.161024e-05 | 1.113334e-04 |
| Tau2_2_2 | 0.0001142197 | 4.726642e-05 | 1.186402e-04 |
| | Tau2_2_2 | | |
| Intercept1 | -2.517793e-05 | | |
| Intercept2 | -4.989090e-06 | | |
| Slope1_1 | 4.768716e-05 | | |
| Slope2_1 | 1.142197e-04 | | |
| Tau2_1_1 | 4.726642e-05 | | |
| Tau2_2_1 | 1.186402e-04 | | |
| Tau2_2_2 | 2.916845e-04 | | |

```
R> ## Multivariate meta-analysis with both regression coefficients fixed at 0
R> mult0 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL), data=Berkey98,
               x=scale(pub_year, center=1979),
               model.name="Both regression coefficients fixed at 0",
               coef.constraints=matrix(c("0", "0"), nrow=2))
```

Running Both regression coefficients fixed at 0

```
R> summary(mult0)
```

Call:

```
meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
     x = scale(pub_year, center = 1979), data = Berkey98, coef.constraints = matrix(c("0",
     "0"), nrow = 2), model.name = "Both regression coefficients fixed at 0")
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound |
|------------|------------|-----------|------------|------------|
| Intercept1 | 0.3448392 | 0.0536312 | 0.2397239 | 0.4499544 |
| Intercept2 | -0.3379381 | 0.0812480 | -0.4971812 | -0.1786951 |
| Tau2_1_1 | 0.0070020 | 0.0090497 | -0.0107351 | 0.0247391 |
| Tau2_2_1 | 0.0094607 | 0.0099698 | -0.0100797 | 0.0290010 |
| Tau2_2_2 | 0.0261445 | 0.0177409 | -0.0086270 | 0.0609161 |

| | z value | Pr(> z) |
|------------|---------|---------------|
| Intercept1 | 6.4298 | 1.278e-10 *** |
| Intercept2 | -4.1593 | 3.192e-05 *** |
| Tau2_1_1 | 0.7737 | 0.4391 |
| Tau2_2_1 | 0.9489 | 0.3427 |
| Tau2_2_2 | 1.4737 | 0.1406 |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 128.2267

Degrees of freedom of the Q statistic: 8

P value of the Q statistic: 0

Number of studies (or clusters): 5

Number of observed statistics: 10

Number of estimated parameters: 5

Degrees of freedom: 5

-2 log likelihood: -11.68131

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Likelihood ratio test on testing both regression coefficients are 0
```

```
R> anova(mult2, mult0)
```

| base | comparison | ep |
|-----------------|------------|----|
| 1 No constraint | <NA> | 7 |

```

2 No constraint Both regression coefficients fixed at 0 5
      minus2LL df      AIC      diffLL diffdf      p
1 -12.00859 3 -18.00859      NA      NA      NA
2 -11.68131 5 -21.68131 0.3272789      2 0.8490481

```

```

R> ## Multivariate meta-analysis with an equality constraint on the slopes
R> mult3 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL), data=Berkey98,
      x=scale(pub_year, center=1979), model.name="With equality constraint",
      coef.constraints=matrix(c("0.3*Equal_Slope", "0.3*Equal_Slope"), nrow=2))

```

Running With equality constraint

```
R> summary(mult3)
```

Call:

```

meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
      x = scale(pub_year, center = 1979), data = Berkey98, coef.constraints = matrix(c("0.3*
      "0.3*Equal_Slope"), nrow = 2), model.name = "With equality constraint")

```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound |
|-------------|------------|-----------|------------|------------|
| Intercept1 | 0.3437612 | 0.0849828 | 0.1771980 | 0.5103245 |
| Equal_Slope | 0.0016745 | 0.1024442 | -0.1991124 | 0.2024614 |
| Intercept2 | -0.3390010 | 0.1041005 | -0.5430343 | -0.1349678 |
| Tau2_1_1 | 0.0070474 | 0.0094638 | -0.0115013 | 0.0255961 |
| Tau2_2_1 | 0.0095164 | 0.0105668 | -0.0111940 | 0.0302269 |
| Tau2_2_2 | 0.0261979 | 0.0180773 | -0.0092330 | 0.0616287 |

| | z value | Pr(> z) |
|-------------|---------|---------------|
| Intercept1 | 4.0451 | 5.231e-05 *** |
| Equal_Slope | 0.0163 | 0.986958 |
| Intercept2 | -3.2565 | 0.001128 ** |
| Tau2_1_1 | 0.7447 | 0.456472 |
| Tau2_2_1 | 0.9006 | 0.367800 |
| Tau2_2_2 | 1.4492 | 0.147278 |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 128.2267

Degrees of freedom of the Q statistic: 8

P value of the Q statistic: 0

Number of studies (or clusters): 5

Number of observed statistics: 10

Number of estimated parameters: 6

Degrees of freedom: 4

-2 log likelihood: -11.68158

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Exact the coefficients
```

```
R> coef(mult3)
```

```
Intercept1 Intercept2 Equal_Slope Tau2_1_1
0.343761239 -0.339001019 0.001674542 0.007047407
Tau2_2_1 Tau2_2_2
0.009516440 0.026197856
```

```
R> ## Exact the sampling variance covariance matrix
```

```
R> vcov(mult3)
```

```
Intercept1 Intercept2 Equal_Slope
Intercept1 7.222075e-03 0.0065104549 -0.0067506847
Intercept2 6.510455e-03 0.0108369149 -0.0066601781
Equal_Slope -6.750685e-03 -0.0066601781 0.0104948091
Tau2_1_1 -6.400934e-05 -0.0001221578 0.0002892786
Tau2_2_1 -2.447019e-04 -0.0001027742 0.0003530272
Tau2_2_2 -2.159290e-04 -0.0001825541 0.0003374792
Tau2_1_1 Tau2_2_1 Tau2_2_2
Intercept1 -6.400934e-05 -2.447019e-04 -2.159290e-04
Intercept2 -1.221578e-04 -1.027742e-04 -1.825541e-04
Equal_Slope 2.892786e-04 3.530272e-04 3.374792e-04
Tau2_1_1 8.956372e-05 6.810625e-05 5.473181e-05
Tau2_2_1 6.810625e-05 1.116565e-04 1.308804e-04
Tau2_2_2 5.473181e-05 1.308804e-04 3.267891e-04
```

```
R> ## Likelihood ratio test on the equality of regression coefficients
```

```
R> anova(mult2, mult3)
```

```
base comparison ep minus2LL df
1 No constraint <NA> 7 -12.00859 3
2 No constraint With equality constraint 6 -11.68158 4
AIC diffLL diffdf p
1 -18.00859 NA NA NA
2 -19.68158 0.3270107 1 0.5674246
```

Multivariate fixed-effects model A multivariate fixed-effects meta-analysis can be conducted by fixing the variance component at a zero matrix. The following code illustrates the syntax.

```
R> ## Multivariate meta-analysis with a fixed-effects model
```

```
R> summary( mult4 <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL),
data=Berkey98,
RE.constraints=matrix(0, nrow=2, ncol=2)) )
```

Running Meta analysis with ML

Call:

```
meta(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
     data = Berkey98, RE.constraints = matrix(0, nrow = 2, ncol = 2))
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z value |
|------------|-----------|-----------|-----------|-----------|---------|
| Intercept1 | 0.307219 | 0.028575 | 0.251212 | 0.363225 | 10.751 |
| Intercept2 | -0.394377 | 0.018649 | -0.430929 | -0.357825 | -21.147 |

Pr(>|z|)

Intercept1 < 2.2e-16 ***

Intercept2 < 2.2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 128.2267

Degrees of freedom of the Q statistic: 8

P value of the Q statistic: 0

Number of studies (or clusters): 5

Number of observed statistics: 10

Number of estimated parameters: 2

Degrees of freedom: 8

-2 log likelihood: 90.88326

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

R> ## Exact the coefficients

R> coef(mult4)

Intercept1 Intercept2

0.3072186 -0.3943770

R> ## Exact the sampling variance covariance matrix

R> vcov(mult4)

Intercept1 Intercept2

Intercept1 0.0008165393 0.0002072041

Intercept2 0.0002072041 0.0003477936

REML The `reml()` function may be used to estimate the variance component with the REML estimation method. It should be noted that it does not estimate the fixed-effects. The fixed-effects estimates can be obtained via the `meta()` function by specifying the estimated variance component from `reml()` as fixed values in the `RE.constraints` argument. This approach is consistent with the idea of REML that removes the fixed-effects parameter when estimating the variance component.


```
R> ## Random-effects meta-analysis with ML
R> summary( meta(y=di, v=vi, data=Becker83) )
```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z value |
|------------|----------|-----------|-----------|----------|---------|
| Intercept1 | 0.174734 | 0.113378 | -0.047482 | 0.396950 | 1.5412 |
| Tau2_1_1 | 0.077376 | 0.054108 | -0.028674 | 0.183426 | 1.4300 |

Pr(>|z|)

| | |
|------------|--------|
| Intercept1 | 0.1233 |
|------------|--------|

| | |
|----------|--------|
| Tau2_1_1 | 0.1527 |
|----------|--------|

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 2

Degrees of freedom: 8

-2 log likelihood: 7.928307

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Random-effects meta-analysis with REML
```

```
R> summary( VarComp <- reml(y=di, v=vi, data=Becker83) )
```

Running Variance component with REML

Call:

```
reml(y = di, v = vi, data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z value |
|----------|----------|-----------|-----------|----------|---------|
| Tau2_1_1 | 0.091445 | 0.064228 | -0.034439 | 0.217329 | 1.4238 |

Pr(>|z|)

| | |
|----------|--------|
| Tau2_1_1 | 0.1545 |
|----------|--------|

Number of studies (or clusters): 10

Number of observed statistics: 9

Number of estimated parameters: 1

Degrees of freedom: 8

-2 log likelihood: -6.110579 OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Extract the variance component
R> VarComp_REML <- matrix( coef(VarComp), ncol=1, nrow=1 )
R> ## Meta-analysis by treating the variance component as fixed
R> summary( meta(y=di, v=vi, data=Becker83, RE.constraints=VarComp_REML) )
```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, data = Becker83, RE.constraints = VarComp_REML)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z value |
|------------|----------|-----------|-----------|----------|---------|
| Intercept1 | 0.180189 | 0.117535 | -0.050176 | 0.410555 | 1.5331 |
| | Pr(> z) | | | | |
| Intercept1 | 0.1253 | | | | |

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

Number of studies (or clusters): 10

Number of observed statistics: 10

Number of estimated parameters: 1

Degrees of freedom: 9

-2 log likelihood: 7.986161

OpenMx status1: 1 ("0" and "1": considered fine; other values indicate problems)

```
R> ## Estimate variance components with REML
```

```
R> summary( reml(y=di, v=vi, x=log(items), data=Becker83) )
```

Running Variance component with REML

Call:

```
reml(y = di, v = vi, x = log(items), data = Becker83)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound |
|----------|------------------|-----------|------------|-----------|
| Tau2_1_1 | 0.0052656 | 0.0212014 | -0.0362884 | 0.0468196 |
| | z value Pr(> z) | | | |
| Tau2_1_1 | 0.2484 | 0.8039 | | |

Number of studies (or clusters): 10

```

Number of observed statistics: 8
Number of estimated parameters: 1
Degrees of freedom: 7
-2 log likelihood: -10.845670openMx status: 0 ("0" and "1": considered fine; other values i

```

```

R> ## Estimate variance components with REML
R> summary( reml(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL), data=Berkey98) )

```

Running Variance component with REML

Call:

```

reml(y = cbind(PD, AL), v = cbind(var_PD, cov_PD_AL, var_AL),
     data = Berkey98)

```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z value |
|----------|----------|-----------|-----------|----------|---------|
| Tau2_1_1 | 0.011733 | 0.013645 | -0.015011 | 0.038477 | 0.8599 |
| Tau2_2_1 | 0.011916 | 0.014416 | -0.016340 | 0.040172 | 0.8266 |
| Tau2_2_2 | 0.032651 | 0.024402 | -0.015176 | 0.080479 | 1.3380 |

Pr(>|z|)

| | |
|----------|--------|
| Tau2_1_1 | 0.3899 |
| Tau2_2_1 | 0.4085 |
| Tau2_2_2 | 0.1809 |

Number of studies (or clusters): 5

Number of observed statistics: 8

Number of estimated parameters: 3

Degrees of freedom: 5

```

-2 log likelihood: -18.867680openMx status: 0 ("0" and "1": considered fine; other values i

```

Plots of multivariate effect sizes If multivariate meta-analysis is conducted, pairwise plots on the pooled effect sizes and their confidence ellipses can be obtained via the `plot()` function. By default, 95% confidence intervals on the average effect sizes and confidence ellipses on the random effects are plotted. For example,

```

R> Berkey98.ma <- meta(y=cbind(PD, AL), v=cbind(var_PD, cov_PD_AL, var_AL), data=Berkey98)

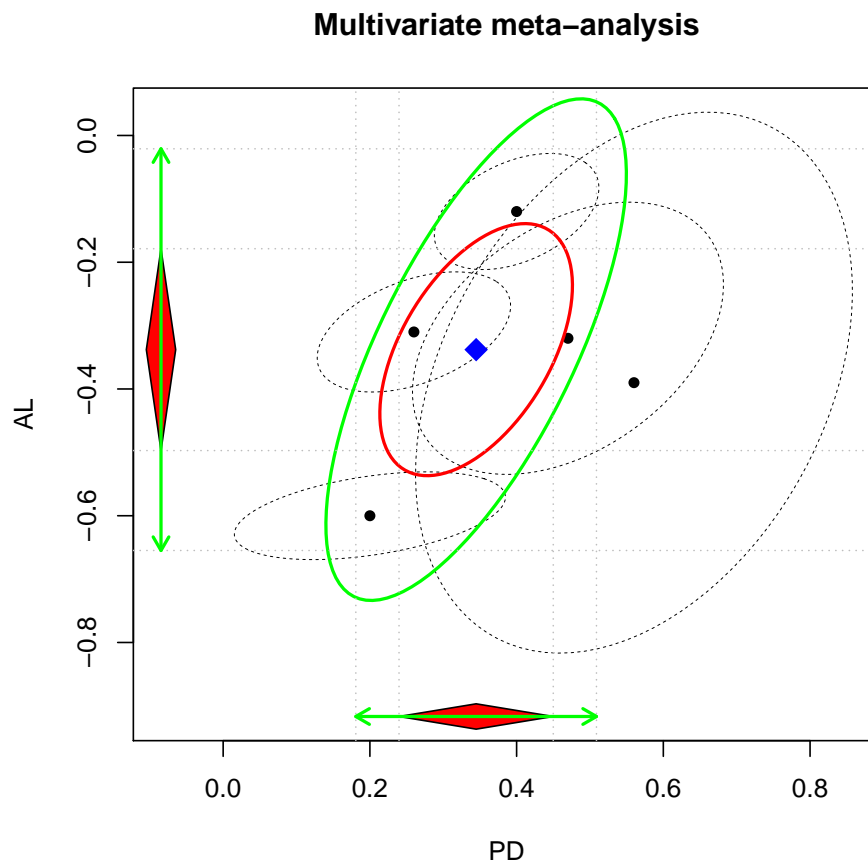
```

Running Meta analysis with ML

```

R> plot(Berkey98.ma, main="Multivariate meta-analysis", axis.label=c("PD", "AL"))

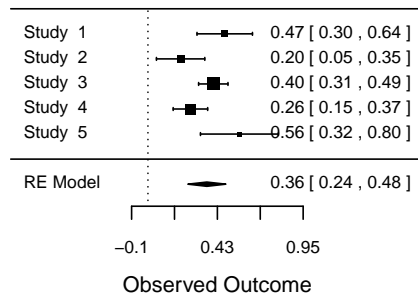
```



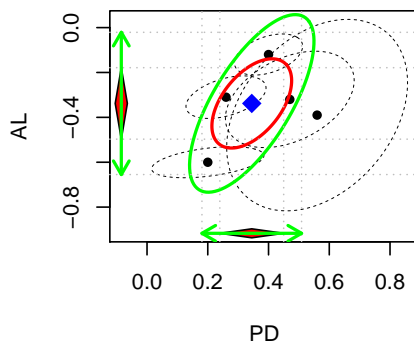
By combining with the forest plots from the **metafor** package, we may combine the univariate and multivariate natures of the effect sizes in a single figure. This will be very useful for multivariate meta-analysis.

```
R> ## Load the metafor package to display forest plots
R> library(metafor)
R> plot(Berkey98.ma, diag.panel=TRUE, main="Multivariate meta-analysis",
       axis.label=c("PD", "AL"))
R> forest( rma(yi=PD, vi=var_PD, data=Berkey98) )
R> title("Forest plot for PD")
R> forest( rma(yi=AL, vi=var_AL, data=Berkey98) )
R> title("Forest plot for AL")
```

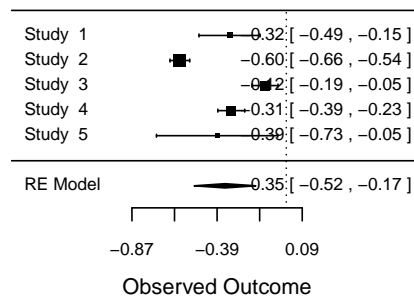
Forest plot for PD



Multivariate meta-analysis



Forest plot for AL



Three-level meta-analysis The `meta3()` function may be used to fit three-level meta-analytic models. It is assumed that effect sizes within `cluster` are dependent.

```
R> ## No predictor
R> summary( meta3(y=y, v=v, cluster=District, data=Cooper03) )
```

Running Meta analysis with ML

Call:

```
meta3(y = y, v = v, cluster = District, data = Cooper03)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound |
|-----------|-----------|-----------|------------|-----------|
| Intercept | 0.1844553 | 0.0805411 | 0.0265977 | 0.3423130 |
| Tau2_2 | 0.0328648 | 0.0111397 | 0.0110314 | 0.0546982 |
| Tau2_3 | 0.0577384 | 0.0307423 | -0.0025154 | 0.1179921 |

z value Pr(>|z|)

| | | | |
|-----------|--------|----------|---|
| Intercept | 2.2902 | 0.022010 | * |
|-----------|--------|----------|---|

```
Tau2_2      2.9502 0.003175 **
Tau2_3      1.8781 0.060362 .
```

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Q statistic on homogeneity of effect sizes: 578.864
```

```
Degrees of freedom of the Q statistic: 55
```

```
P value of the Q statistic: 0
```

```
Number of studies (or clusters): 11
```

```
Number of observed statistics: 56
```

```
Number of estimated parameters: 3
```

```
Degrees of freedom: 53
```

```
-2 log likelihood: 16.78987
```

```
Heterogeneity indices:
```

```
              Estimate
I2_2 (harmonic mean)  0.3447
I2_3 (harmonic mean)  0.6057
```

```
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)
```

```
R> ## Year as a predictor
```

```
R> summary( meta3(y=y, v=v, cluster=District, x=scale(Year, scale=FALSE), data=Cooper03) )
```

```
Running Meta analysis with ML
```

```
Call:
```

```
meta3(y = y, v = v, cluster = District, x = scale(Year, scale = FALSE),
      data = Cooper03)
```

```
95% confidence intervals: z statistic approximation
```

```
Coefficients:
```

| | Estimate | Std.Error | lbound | ubound |
|-----------|-----------|-----------|------------|-----------|
| Slope_1 | 0.0050737 | 0.0085266 | -0.0116382 | 0.0217856 |
| Intercept | 0.1780268 | 0.0805219 | 0.0202067 | 0.3358469 |
| Tau2_2 | 0.0329390 | 0.0111620 | 0.0110618 | 0.0548162 |
| Tau2_3 | 0.0564628 | 0.0300330 | -0.0024007 | 0.1153264 |

```
      z value Pr(>|z|)
```

| | | |
|-----------|--------|-------------|
| Slope_1 | 0.5950 | 0.551814 |
| Intercept | 2.2109 | 0.027042 * |
| Tau2_2 | 2.9510 | 0.003168 ** |
| Tau2_3 | 1.8800 | 0.060104 . |

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Q statistic on homogeneity of effect sizes: 578.864
```

```
Degrees of freedom of the Q statistic: 55
```

P value of the Q statistic: 0

Number of studies (or clusters): 11

Number of observed statistics: 56

Number of estimated parameters: 4

Degrees of freedom: 52

-2 log likelihood: 16.43629

R2 (untruncated):

| | Value |
|--------------------------|---------|
| Tau2_2 (no predictor) | 0.0329 |
| Tau2_2 (with predictors) | 0.0329 |
| R2_2 | -0.0023 |
| Tau2_3 (no predictor) | 0.0577 |
| Tau2_3 (with predictors) | 0.0565 |
| R2_3 | 0.0221 |
| -2LL (no predictor) | 16.7899 |
| -2LL (with predictors) | 16.4363 |
| R2 (pseudo) | 0.0211 |

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

3. Meta-analytic structural equation modeling

MASEM combines the idea of meta-analysis and SEM by pooling correlation/covariance matrices and testing structural equation models on the pooled matrix. There are two stages in conducting a MASEM. In the first stage the correlation/covariance matrices are pooled together. In the second stage, the pooled correlation/covariance matrix is used to fit structural equation models.

Cheung and Chan (2005b, 2009) proposed a two-stage structural equation modeling (TSSEM) based on a fixed-effects model. The `metaSEM` package has implemented the TSSEM approach. Moreover, the TSSEM approach has been extended to the random-effects model by using a multivariate meta-analysis (Cheung 2011b) in the first stage analysis. Regardless of whether a fixed- or random-effects model is used, the `tssem2()` function will handle this automatically. In other words, parameter estimates, standard errors and goodness-of-fit indices in the stage 2 analysis has already taken the stage 1 model into account.

An example from Cheung (2009b) is used to illustrate the procedures. In this example, Digman (1997) reported a second-order factor analysis on a five-factor model with 14 studies. He proposed that there were two second-order factors for the five-factor model: an alpha factor for agreeableness, conscientiousness, and emotional stability, and a beta factor for extroversion and intellect.

3.1. Fixed-effects model

The `tssem1()` function is used to pool the correlation matrices with a fixed-effects model in the first stage by specifying `method='FEM'` in the argument. `tssem2()` is then used to fit

a factor analytic model on the pooled correlation matrix with the inverse of its asymptotic covariance matrix as the weight matrix (Cheung and Chan 2005b, 2009).

The fit indices for testing the homogeneity of the correlation matrices in Stage 1 analysis are $\chi^2(130, N = 4496) = 1499.73, p < .001$, CFI=0.6825, TLI=0.6581, SRMR=0.1750 and RMSEA=0.1812. This indicates that it is not reasonable to assume that the correlation matrices are homogenous. Sub-group analysis or random-effects model that will be illustrated later are more appropriate. As an exercise, we continue to fit the stage 2 model. The fit indices for fitting the structural model in Stage 2 are $\chi^2(4, N = 4496) = 67.89, p < .001$, CFI=0.9845, TLI=0.9613, SRMR=0.0285 and RMSEA=0.0596.

```
R> ## Show the first 2 studies in Digman97
R> head(Digman97$data, n=2)
```

```
$`Digman 1 (1994)`
      E      A      C      ES      I
E  1.00 -0.48 -0.10  0.27  0.37
A -0.48  1.00  0.62  0.41  0.00
C -0.10  0.62  1.00  0.59  0.35
ES  0.27  0.41  0.59  1.00  0.41
I  0.37  0.00  0.35  0.41  1.00
```

```
$`Digman 2 (1994)`
      E      A      C      ES      I
E  1.00 -0.30  0.07  0.09  0.45
A -0.30  1.00  0.39  0.53 -0.05
C  0.07  0.39  1.00  0.59  0.44
ES  0.09  0.53  0.59  1.00  0.22
I  0.45 -0.05  0.44  0.22  1.00
```

```
R> ## Show the first 2 sample sizes in Digman97
R> head(Digman97$n, n=2)
```

```
[1] 102 149
```

```
R> ## Example of Fixed-effects TSSEM
R> fixed1 <- tssem1(Digman97$data, Digman97$n, method="FEM")
```

Running TSSEM1 Analysis of Correlation Matrix

```
R> summary(fixed1)
```

Call:

```
tssem1FEM(my.df = my.df, n = n, cor.analysis = cor.analysis,
  model.name = model.name, cluster = cluster, suppressWarnings = suppressWarnings)
```

Coefficients:

| | Estimate | Std.Error | z value | Pr(> z) |
|--------|----------|-----------|---------|---------------|
| S[1,2] | 0.103751 | 0.015070 | 6.8846 | 5.796e-12 *** |
| S[1,3] | 0.135208 | 0.014799 | 9.1363 | < 2.2e-16 *** |
| S[1,4] | 0.244505 | 0.014175 | 17.2487 | < 2.2e-16 *** |
| S[1,5] | 0.424514 | 0.012396 | 34.2463 | < 2.2e-16 *** |
| S[2,3] | 0.363116 | 0.013391 | 27.1169 | < 2.2e-16 *** |
| S[2,4] | 0.390176 | 0.012903 | 30.2387 | < 2.2e-16 *** |
| S[2,5] | 0.092246 | 0.015071 | 6.1207 | 9.319e-10 *** |
| S[3,4] | 0.415999 | 0.012540 | 33.1736 | < 2.2e-16 *** |
| S[3,5] | 0.141213 | 0.014891 | 9.4834 | < 2.2e-16 *** |
| S[4,5] | 0.138167 | 0.014858 | 9.2991 | < 2.2e-16 *** |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Goodness-of-fit indices:

| | Value |
|---------------------------------|-----------|
| Sample size | 4496.0000 |
| Chi-square of target model | 1499.7340 |
| DF of target model | 130.0000 |
| p value of target model | 0.0000 |
| Chi-square of independent model | 4454.5995 |
| DF of independent model | 140.0000 |
| RMSEA | 0.1812 |
| SRMR | 0.1750 |
| TLI | 0.6581 |
| CFI | 0.6825 |
| AIC | 1239.7340 |
| BIC | 406.3114 |

OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

R> ## Extract the pooled correlation matrix

R> coef(fixed1)

| | x1 | x2 | x3 | x4 | x5 |
|----|-----------|------------|------------|------------|------------|
| x1 | 1.0000000 | 0.10375113 | 0.1352076 | 0.2445051 | 0.42451422 |
| x2 | 0.1037511 | 1.00000000 | 0.3631157 | 0.3901765 | 0.09224586 |
| x3 | 0.1352076 | 0.36311571 | 1.00000000 | 0.4159987 | 0.14121294 |
| x4 | 0.2445051 | 0.39017646 | 0.4159987 | 1.00000000 | 0.13816667 |
| x5 | 0.4245142 | 0.09224586 | 0.1412129 | 0.1381667 | 1.00000000 |

R> ## Exact the sampling variance covariance matrix

R> vcov(fixed1)

| | x2 x1 | x3 x1 | x4 x1 | x5 x1 |
|-------|--------------|--------------|--------------|--------------|
| x2 x1 | 2.131326e-04 | 7.178645e-05 | 7.186422e-05 | 7.138907e-06 |
| x3 x1 | 7.178645e-05 | 2.083180e-04 | 7.495557e-05 | 1.252304e-05 |
| x4 x1 | 7.186422e-05 | 7.495557e-05 | 1.859689e-04 | 4.745543e-06 |

```

x5 x1 7.138907e-06 1.252304e-05 4.745543e-06 1.270257e-04
x3 x2 1.459845e-05 7.180707e-06 1.147490e-06 4.523611e-07
x4 x2 3.291319e-05 5.714323e-06 9.431060e-07 8.038123e-07
x5 x2 8.828850e-05 2.902149e-05 2.930451e-05 9.758175e-06
x4 x3 6.819703e-06 2.905740e-05 4.659928e-06 1.707844e-06
x5 x3 2.891476e-05 8.423574e-05 3.026973e-05 1.126867e-05
x5 x4 2.994808e-05 3.164603e-05 7.646784e-05 2.841440e-05
      x3 x2      x4 x2      x5 x2      x4 x3
x2 x1 1.459845e-05 3.291319e-05 8.828850e-05 6.819703e-06
x3 x1 7.180707e-06 5.714323e-06 2.902149e-05 2.905740e-05
x4 x1 1.147490e-06 9.431060e-07 2.930451e-05 4.659928e-06
x5 x1 4.523611e-07 8.038123e-07 9.758175e-06 1.707844e-06
x3 x2 1.469907e-04 3.838604e-05 1.738798e-05 3.188754e-05
x4 x2 3.838604e-05 1.380993e-04 1.593018e-05 2.548899e-05
x5 x2 1.738798e-05 1.593018e-05 2.160830e-04 5.068022e-06
x4 x3 3.188754e-05 2.548899e-05 5.068022e-06 1.292688e-04
x5 x3 5.973456e-06 2.980306e-06 7.375682e-05 1.189884e-05
x5 x4 3.714154e-06 5.377597e-06 7.926225e-05 1.248484e-05
      x5 x3      x5 x4
x2 x1 2.891476e-05 2.994808e-05
x3 x1 8.423574e-05 3.164603e-05
x4 x1 3.026973e-05 7.646784e-05
x5 x1 1.126867e-05 2.841440e-05
x3 x2 5.973456e-06 3.714154e-06
x4 x2 2.980306e-06 5.377597e-06
x5 x2 7.375682e-05 7.926225e-05
x4 x3 1.189884e-05 1.248484e-05
x5 x3 2.092634e-04 8.223961e-05
x5 x4 8.223961e-05 2.077774e-04

```

```
R> ## S matrix
```

```
R> Phi <- matrix(c(1,"0.3*cor","0.3*cor",1), ncol=2, nrow=2)
```

```
R> S1 <- bdiagMat(list(diag(c("0.2*e1","0.2*e2","0.2*e3","0.2*e4","0.2*e5")), Phi))
```

```
R> S1 <- as.mxMatrix(S1)
```

```
R> ## A matrix
```

```
R> Lambda <- matrix(c(0,".3*f1_x2",".3*f1_x3",".3*f1_x4",0,".3*f2_x1",0,0,0,".3*f2_x5"),
                    ncol=2, nrow=5)
```

```
R> A1 <- rbind( cbind(matrix(0,ncol=5,nrow=5), Lambda),
               matrix(0, ncol=7, nrow=2) )
```

```
R> A1 <- as.mxMatrix(A1)
```

```
R> F1 <- create.Fmatrix(c(1,1,1,1,1,0,0), name="F1")
```

```
R> fixed2 <- tssem2(fixed1, Amatrix=A1, Smatrix=S1, Fmatrix=F1, diag.constraints=TRUE, int
                    model.name="TSSEM2 Digman97")
```

```
Running TSSEM2 Digman97
```

```
R> summary(fixed2)
```

Call:

```
wls(Cov = tssem1.obj$pooledS, asyCov = tssem1.obj$acovS, n = tssem1.obj$total.n,
     Amatrix = Amatrix, Smatrix = Smatrix, Fmatrix = Fmatrix,
     diag.constraints = diag.constraints, cor.analysis = cor.analysis,
     intervals.type = intervals.type, model.name = model.name,
     suppressWarnings = suppressWarnings)
```

95% confidence intervals: Likelihood-based statistic

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z | value |
|--------------|----------|-----------|---------|---------|---|-------|
| Amatrix[1,7] | 0.77957 | NA | 0.71885 | 0.85082 | | NA |
| Amatrix[2,6] | 0.56449 | NA | 0.53661 | 0.59263 | | NA |
| Amatrix[3,6] | 0.60714 | NA | 0.57926 | 0.63535 | | NA |
| Amatrix[4,6] | 0.71713 | NA | 0.68854 | 0.74652 | | NA |
| Amatrix[5,7] | 0.55421 | NA | 0.50472 | 0.60383 | | NA |
| Smatrix[1,1] | 0.39228 | NA | 0.27605 | 0.48328 | | NA |
| Smatrix[2,2] | 0.68136 | NA | 0.64879 | 0.71205 | | NA |
| Smatrix[3,3] | 0.63138 | NA | 0.59633 | 0.66446 | | NA |
| Smatrix[4,4] | 0.48572 | NA | 0.44271 | 0.52592 | | NA |
| Smatrix[5,5] | 0.69286 | NA | 0.63539 | 0.74526 | | NA |
| Smatrix[7,6] | 0.36269 | NA | 0.31955 | 0.40574 | | NA |

Pr(>|z|)

| | |
|--------------|----|
| Amatrix[1,7] | NA |
| Amatrix[2,6] | NA |
| Amatrix[3,6] | NA |
| Amatrix[4,6] | NA |
| Amatrix[5,7] | NA |
| Smatrix[1,1] | NA |
| Smatrix[2,2] | NA |
| Smatrix[3,3] | NA |
| Smatrix[4,4] | NA |
| Smatrix[5,5] | NA |
| Smatrix[7,6] | NA |

Goodness-of-fit indices:

| | Value |
|---|-----------|
| Sample size | 4496.0000 |
| Chi-square of target model | 67.8897 |
| DF of target model | 4.0000 |
| p value of target model | 0.0000 |
| Chi-square of independent model | 4132.8505 |
| DF of independent model | 10.0000 |
| No. of constraints imposed on "Smatrix" | 5.0000 |
| RMSEA | 0.0000 |
| SRMR | 0.0285 |
| TLI | 0.9613 |
| CFI | 0.9845 |

```

AIC                      59.8897
BIC                      34.2459
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

Example: Fixed-effects model with sub-group analysis Studies may not share the same population correlation matrix. If the studies can be grouped into various subgroups, we may pool the correlation matrices separately by the subgroups (Cheung and Chan 2005a). This is similar to the subgroup analysis in conventional meta-analysis (Hedges and Olkin 1985). For example, Digman (1997) groups the 14 studies into several groups according to their sample characteristics. These include children, adolescents, young adults, and mature adults. This variable is stored in the variable `Digman97$cluster`.

The following R code may be used to conduct the analysis. Users have to supply the cluster (a vector of labels) to the `cluster` argument in `tssem1()`. The correlation/covariance matrices will be pooled separately for each cluster. The structural models will also be fitted separately for each cluster.

```
R> ## Show the clusters in Digman97
```

```
R> Digman97$cluster
```

```

[1] "Children"      "Children"      "Children"
[4] "Children"      "Adolescents"   "Young adults"
[7] "Young adults"  "Young adults"  "Mature adults"
[10] "Mature adults" "Mature adults" "Mature adults"
[13] "Mature adults" "Mature adults"

```

```
R> ## Example of Fixed-effects TSSEM with several clusters
```

```
R> fixed1.cluster <- tssem1(Digman97$data, Digman97$n, method="FEM",
                           cluster=Digman97$cluster)
```

```

Running TSSEM1 Analysis of Correlation Matrix
Running TSSEM1 Analysis of Correlation Matrix
Running TSSEM1 Analysis of Correlation Matrix
Running TSSEM1 Analysis of Correlation Matrix

```

```
R> summary(fixed1.cluster)
```

```
$Adolescents
```

```
Call:
```

```
tssem1FEM(my.df = data.cluster[[i]], n = n.cluster[[i]], cor.analysis = cor.analysis,
          model.name = model.name, suppressWarnings = suppressWarnings)
```

```
Coefficients:
```

```

      Estimate Std.Error z value Pr(>|z|)
S[1,2] 0.290000  0.096544  3.0038 0.0026661 **
S[1,3] 0.160000  0.102710  1.5578 0.1192859

```

```

S[1,4] 0.320000 0.094615 3.3821 0.0007193 ***
S[1,5] 0.530000 0.075800 6.9921 2.708e-12 ***
S[2,3] 0.640000 0.062234 10.2838 < 2.2e-16 ***
S[2,4] 0.350000 0.092496 3.7839 0.0001544 ***
S[2,5] 0.220000 0.100307 2.1933 0.0282882 *
S[3,4] 0.270000 0.097725 2.7629 0.0057296 **
S[3,5] 0.220000 0.100307 2.1933 0.0282887 *
S[4,5] 0.360000 0.091748 3.9238 8.717e-05 ***

```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Goodness-of-fit indices:

| | Value |
|---------------------------------|--------|
| Sample size | 91.00 |
| Chi-square of target model | 0.00 |
| DF of target model | 0.00 |
| p value of target model | 0.00 |
| Chi-square of independent model | 109.63 |
| DF of independent model | 10.00 |
| RMSEA | Inf |
| SRMR | 0.00 |
| TLI | -Inf |
| CFI | 1.00 |
| AIC | 0.00 |
| BIC | 0.00 |

OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

\$Children

Call:

```

tssem1FEM(my.df = data.cluster[[i]], n = n.cluster[[i]], cor.analysis = cor.analysis,
  model.name = model.name, suppressWarnings = suppressWarnings)

```

Coefficients:

| | Estimate | Std.Error | z value | Pr(> z) |
|--------|-----------|-----------|---------|---------------|
| S[1,2] | -0.071259 | 0.037983 | -1.8761 | 0.06065 . |
| S[1,3] | -0.084678 | 0.036505 | -2.3196 | 0.02036 * |
| S[1,4] | 0.158313 | 0.035949 | 4.4038 | 1.064e-05 *** |
| S[1,5] | 0.473158 | 0.028765 | 16.4489 | < 2.2e-16 *** |
| S[2,3] | 0.600192 | 0.023695 | 25.3302 | < 2.2e-16 *** |
| S[2,4] | 0.479811 | 0.028723 | 16.7049 | < 2.2e-16 *** |
| S[2,5] | 0.043055 | 0.036728 | 1.1723 | 0.24109 |
| S[3,4] | 0.526708 | 0.026960 | 19.5366 | < 2.2e-16 *** |
| S[3,5] | 0.331623 | 0.032726 | 10.1333 | < 2.2e-16 *** |
| S[4,5] | 0.298135 | 0.033718 | 8.8419 | < 2.2e-16 *** |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Goodness-of-fit indices:

| | Value |
|---------------------------------|-----------|
| Sample size | 747.0000 |
| Chi-square of target model | 311.3516 |
| DF of target model | 30.0000 |
| p value of target model | 0.0000 |
| Chi-square of independent model | 1352.0398 |
| DF of independent model | 40.0000 |
| RMSEA | 0.2242 |
| SRMR | 0.1401 |
| TLI | 0.7141 |
| CFI | 0.7856 |
| AIC | 251.3516 |
| BIC | 112.8696 |

OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

\$`Mature adults`

Call:

```
tssem1FEM(my.df = data.cluster[[i]], n = n.cluster[[i]], cor.analysis = cor.analysis,
  model.name = model.name, suppressWarnings = suppressWarnings)
```

Coefficients:

| | Estimate | Std.Error | z value | Pr(> z) | |
|--------|----------|-----------|---------|-----------|-----|
| S[1,2] | 0.076227 | 0.018725 | 4.0708 | 4.686e-05 | *** |
| S[1,3] | 0.170404 | 0.018269 | 9.3274 | < 2.2e-16 | *** |
| S[1,4] | 0.191577 | 0.018047 | 10.6156 | < 2.2e-16 | *** |
| S[1,5] | 0.366062 | 0.016326 | 22.4217 | < 2.2e-16 | *** |
| S[2,3] | 0.196629 | 0.018031 | 10.9051 | < 2.2e-16 | *** |
| S[2,4] | 0.305076 | 0.017177 | 17.7604 | < 2.2e-16 | *** |
| S[2,5] | 0.013859 | 0.018955 | 0.7312 | 0.4647 | |
| S[3,4] | 0.385234 | 0.016189 | 23.7955 | < 2.2e-16 | *** |
| S[3,5] | 0.030844 | 0.018907 | 1.6314 | 0.1028 | |
| S[4,5] | 0.037283 | 0.018830 | 1.9800 | 0.0477 | * |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Goodness-of-fit indices:

| | Value |
|---------------------------------|-----------|
| Sample size | 2862.0000 |
| Chi-square of target model | 420.5247 |
| DF of target model | 50.0000 |
| p value of target model | 0.0000 |
| Chi-square of independent model | 1707.9108 |
| DF of independent model | 60.0000 |
| RMSEA | 0.1247 |

```

SRMR                      0.1522
TLI                       0.7302
CFI                       0.7752
AIC                       320.5247
BIC                       22.5609
OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)

```

```
$`Young adults`
```

```
Call:
```

```
tssem1FEM(my.df = data.cluster[[i]], n = n.cluster[[i]], cor.analysis = cor.analysis,
  model.name = model.name, suppressWarnings = suppressWarnings)
```

```
Coefficients:
```

| | Estimate | Std.Error | z value | Pr(> z) |
|--------|----------|-----------|---------|---------------|
| S[1,2] | 0.322154 | 0.031894 | 10.1009 | < 2.2e-16 *** |
| S[1,3] | 0.219371 | 0.033847 | 6.4813 | 9.093e-11 *** |
| S[1,4] | 0.471710 | 0.027645 | 17.0630 | < 2.2e-16 *** |
| S[1,5] | 0.554691 | 0.024744 | 22.4175 | < 2.2e-16 *** |
| S[2,3] | 0.613646 | 0.022452 | 27.3309 | < 2.2e-16 *** |
| S[2,4] | 0.560195 | 0.024403 | 22.9555 | < 2.2e-16 *** |
| S[2,5] | 0.351222 | 0.031207 | 11.2545 | < 2.2e-16 *** |
| S[3,4] | 0.424434 | 0.029180 | 14.5454 | < 2.2e-16 *** |
| S[3,5] | 0.286843 | 0.032639 | 8.7883 | < 2.2e-16 *** |
| S[4,5] | 0.276926 | 0.032878 | 8.4227 | < 2.2e-16 *** |

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Goodness-of-fit indices:
```

| | Value |
|---------------------------------|-----------|
| Sample size | 796.0000 |
| Chi-square of target model | 66.8333 |
| DF of target model | 20.0000 |
| p value of target model | 0.0000 |
| Chi-square of independent model | 1285.0187 |
| DF of independent model | 30.0000 |
| RMSEA | 0.0940 |
| SRMR | 0.1511 |
| TLI | 0.9440 |
| CFI | 0.9627 |
| AIC | 26.8333 |
| BIC | -66.7587 |

```
OpenMx status: 0 ("0" and "1": considered fine; other values indicate problems)
```

```
R> ## Extract the pooled correlation matrices
```

```
R> coef(fixed1.cluster)
```

```
$Adolescents
```

```
      x1  x2  x3  x4  x5
x1 1.00 0.29 0.16 0.32 0.53
x2 0.29 1.00 0.64 0.35 0.22
x3 0.16 0.64 1.00 0.27 0.22
x4 0.32 0.35 0.27 1.00 0.36
x5 0.53 0.22 0.22 0.36 1.00
```

```
$Children
```

```
      x1      x2      x3      x4      x5
x1 1.00000000 -0.07125880 -0.08467822 0.1583127 0.47315830
x2 -0.07125880 1.00000000 0.60019239 0.4798110 0.04305501
x3 -0.08467822 0.60019239 1.00000000 0.5267076 0.33162339
x4 0.15831269 0.47981099 0.52670763 1.00000000 0.29813501
x5 0.47315830 0.04305501 0.33162339 0.2981350 1.00000000
```

```
$`Mature adults`
```

```
      x1      x2      x3      x4      x5
x1 1.00000000 0.0762269 0.17040394 0.19157746 0.36606178
x2 0.0762269 1.00000000 0.19662879 0.30507564 0.01385940
x3 0.1704039 0.1966288 1.00000000 0.38523398 0.03084429
x4 0.1915775 0.3050756 0.38523398 1.00000000 0.03728294
x5 0.3660618 0.0138594 0.03084429 0.03728294 1.00000000
```

```
$`Young adults`
```

```
      x1      x2      x3      x4      x5
x1 1.00000000 0.3221536 0.2193714 0.4717095 0.5546912
x2 0.3221536 1.00000000 0.6136457 0.5601947 0.3512219
x3 0.2193714 0.6136457 1.00000000 0.4244336 0.2868427
x4 0.4717095 0.5601947 0.4244336 1.00000000 0.2769260
x5 0.5546912 0.3512219 0.2868427 0.2769260 1.00000000
```

```
R> fixed2.cluster <- tssem2(fixed1.cluster, Amatrix=A1, Smatrix=S1, Fmatrix=F1, diag.const
```

```
Running TSSEM2 (Fixed Effects Model) Analysis of Correlation Structure
```

```
Running TSSEM2 (Fixed Effects Model) Analysis of Correlation Structure
```

```
Running TSSEM2 (Fixed Effects Model) Analysis of Correlation Structure
```

```
Running TSSEM2 (Fixed Effects Model) Analysis of Correlation Structure
```

```
R> ## Estimatio problems in the "Children" group
```

```
R> summary(fixed2.cluster)
```

```
$Adolescents
```

```
Call:
```

```
wls(Cov = tssem1.obj$pooledS, asyCov = tssem1.obj$acovS, n = tssem1.obj$total.n,
     Amatrix = Amatrix, Smatrix = Smatrix, Fmatrix = Fmatrix,
```



```
diag.constraints = diag.constraints, cor.analysis = cor.analysis,
intervals.type = intervals.type, model.name = model.name,
suppressWarnings = suppressWarnings)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z | value |
|--------------|----------|-----------|--------|--------|----|-------|
| Amatrix[1,7] | 0.73828 | NA | NA | NA | NA | NA |
| Amatrix[2,6] | 0.86734 | NA | NA | NA | NA | NA |
| Amatrix[3,6] | 0.74250 | NA | NA | NA | NA | NA |
| Amatrix[4,6] | 0.52604 | NA | NA | NA | NA | NA |
| Amatrix[5,7] | 0.73400 | NA | NA | NA | NA | NA |
| Smatrix[1,1] | 0.45494 | NA | NA | NA | NA | NA |
| Smatrix[2,2] | 0.24772 | NA | NA | NA | NA | NA |
| Smatrix[3,3] | 0.44869 | NA | NA | NA | NA | NA |
| Smatrix[4,4] | 0.72328 | NA | NA | NA | NA | NA |
| Smatrix[5,5] | 0.46124 | NA | NA | NA | NA | NA |
| Smatrix[7,6] | 0.54808 | NA | NA | NA | NA | NA |

Pr(>|z|)

| | |
|--------------|----|
| Amatrix[1,7] | NA |
| Amatrix[2,6] | NA |
| Amatrix[3,6] | NA |
| Amatrix[4,6] | NA |
| Amatrix[5,7] | NA |
| Smatrix[1,1] | NA |
| Smatrix[2,2] | NA |
| Smatrix[3,3] | NA |
| Smatrix[4,4] | NA |
| Smatrix[5,5] | NA |
| Smatrix[7,6] | NA |

Goodness-of-fit indices:

| | Value |
|---|----------|
| Sample size | 91.0000 |
| Chi-square of target model | 10.7341 |
| DF of target model | 4.0000 |
| p value of target model | 0.0297 |
| Chi-square of independent model | 270.6747 |
| DF of independent model | 10.0000 |
| No. of constraints imposed on "Smatrix" | 5.0000 |
| RMSEA | 0.0000 |
| SRMR | 0.1028 |
| TLI | 0.9354 |
| CFI | 0.9742 |
| AIC | 2.7341 |
| BIC | -7.3094 |

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

\$Children

Call:

```
wls(Cov = tssem1.obj$pooledS, asyCov = tssem1.obj$acovS, n = tssem1.obj$total.n,
     Amatrix = Amatrix, Smatrix = Smatrix, Fmatrix = Fmatrix,
     diag.constraints = diag.constraints, cor.analysis = cor.analysis,
     intervals.type = intervals.type, model.name = model.name,
     suppressWarnings = suppressWarnings)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z | value |
|--------------|-------------|-----------|--------|--------|----|-------|
| Amatrix[1,7] | 0.028309 | NA | NA | NA | NA | NA |
| Amatrix[2,6] | 0.737157 | NA | NA | NA | NA | NA |
| Amatrix[3,6] | 0.913906 | NA | NA | NA | NA | NA |
| Amatrix[4,6] | 0.689528 | NA | NA | NA | NA | NA |
| Amatrix[5,7] | 18.161955 | NA | NA | NA | NA | NA |
| Smatrix[1,1] | 0.999199 | NA | NA | NA | NA | NA |
| Smatrix[2,2] | 0.456599 | NA | NA | NA | NA | NA |
| Smatrix[3,3] | 0.164777 | NA | NA | NA | NA | NA |
| Smatrix[4,4] | 0.524552 | NA | NA | NA | NA | NA |
| Smatrix[5,5] | -328.856546 | NA | NA | NA | NA | NA |
| Smatrix[7,6] | 0.021225 | NA | NA | NA | NA | NA |

Pr(>|z|)

| | |
|--------------|----|
| Amatrix[1,7] | NA |
| Amatrix[2,6] | NA |
| Amatrix[3,6] | NA |
| Amatrix[4,6] | NA |
| Amatrix[5,7] | NA |
| Smatrix[1,1] | NA |
| Smatrix[2,2] | NA |
| Smatrix[3,3] | NA |
| Smatrix[4,4] | NA |
| Smatrix[5,5] | NA |
| Smatrix[7,6] | NA |

Goodness-of-fit indices:

| | Value |
|---|-----------|
| Sample size | 747.0000 |
| Chi-square of target model | 150.9150 |
| DF of target model | 4.0000 |
| p value of target model | 0.0000 |
| Chi-square of independent model | 3583.7700 |
| DF of independent model | 10.0000 |
| No. of constraints imposed on "Smatrix" | 5.0000 |
| RMSEA | 0.0000 |

```

SRMR                0.1075
TLI                 0.8972
CFI                 0.9589
AIC                 142.9150
BIC                 124.4508
OpenMx status1: 4 ("0" and "1": considered fine; other values indicate problems)

```

```
$`Mature adults`
```

```
Call:
```

```

wls(Cov = tssem1.obj$pooledS, asyCov = tssem1.obj$acovS, n = tssem1.obj$total.n,
     Amatrix = Amatrix, Smatrix = Smatrix, Fmatrix = Fmatrix,
     diag.constraints = diag.constraints, cor.analysis = cor.analysis,
     intervals.type = intervals.type, model.name = model.name,
     suppressWarnings = suppressWarnings)

```

```
95% confidence intervals: z statistic approximation
```

```
Coefficients:
```

| | Estimate | Std.Error | lbound | ubound | z | value |
|--------------|----------|-----------|--------|--------|----|-------|
| Amatrix[1,7] | 1.38648 | NA | NA | NA | NA | NA |
| Amatrix[2,6] | 0.39788 | NA | NA | NA | NA | NA |
| Amatrix[3,6] | 0.52301 | NA | NA | NA | NA | NA |
| Amatrix[4,6] | 0.74612 | NA | NA | NA | NA | NA |
| Amatrix[5,7] | 0.26460 | NA | NA | NA | NA | NA |
| Smatrix[1,1] | -0.92232 | NA | NA | NA | NA | NA |
| Smatrix[2,2] | 0.84169 | NA | NA | NA | NA | NA |
| Smatrix[3,3] | 0.72646 | NA | NA | NA | NA | NA |
| Smatrix[4,4] | 0.44331 | NA | NA | NA | NA | NA |
| Smatrix[5,5] | 0.92999 | NA | NA | NA | NA | NA |
| Smatrix[7,6] | 0.19241 | NA | NA | NA | NA | NA |

```
Pr(>|z|)
```

| | |
|--------------|----|
| Amatrix[1,7] | NA |
| Amatrix[2,6] | NA |
| Amatrix[3,6] | NA |
| Amatrix[4,6] | NA |
| Amatrix[5,7] | NA |
| Smatrix[1,1] | NA |
| Smatrix[2,2] | NA |
| Smatrix[3,3] | NA |
| Smatrix[4,4] | NA |
| Smatrix[5,5] | NA |
| Smatrix[7,6] | NA |

```
Goodness-of-fit indices:
```

| | Value |
|----------------------------|-----------|
| Sample size | 2862.0000 |
| Chi-square of target model | 8.9336 |

```

DF of target model                4.0000
p value of target model            0.0628
Chi-square of independent model    1704.2578
DF of independent model            10.0000
No. of constraints imposed on "Smatrix"  5.0000
RMSEA                             0.0000
SRMR                              0.0148
TLI                              0.9927
CFI                              0.9971
AIC                              0.9336
BIC                             -22.9036
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

```
$`Young adults`
```

```
Call:
```

```

wls(Cov = tssem1.obj$pooledS, asyCov = tssem1.obj$acovS, n = tssem1.obj$total.n,
     Amatrix = Amatrix, Smatrix = Smatrix, Fmatrix = Fmatrix,
     diag.constraints = diag.constraints, cor.analysis = cor.analysis,
     intervals.type = intervals.type, model.name = model.name,
     suppressWarnings = suppressWarnings)

```

```
95% confidence intervals: z statistic approximation
```

```
Coefficients:
```

| | Estimate | Std.Error | lbound | ubound | z | value |
|--------------|----------|-----------|--------|--------|----|-------|
| Amatrix[1,7] | 0.86518 | NA | NA | NA | NA | NA |
| Amatrix[2,6] | 0.84474 | NA | NA | NA | NA | NA |
| Amatrix[3,6] | 0.69998 | NA | NA | NA | NA | NA |
| Amatrix[4,6] | 0.76537 | NA | NA | NA | NA | NA |
| Amatrix[5,7] | 0.70899 | NA | NA | NA | NA | NA |
| Smatrix[1,1] | 0.25146 | NA | NA | NA | NA | NA |
| Smatrix[2,2] | 0.28642 | NA | NA | NA | NA | NA |
| Smatrix[3,3] | 0.51003 | NA | NA | NA | NA | NA |
| Smatrix[4,4] | 0.41421 | NA | NA | NA | NA | NA |
| Smatrix[5,5] | 0.49733 | NA | NA | NA | NA | NA |
| Smatrix[7,6] | 0.59588 | NA | NA | NA | NA | NA |

```
Pr(>|z|)
```

| | |
|--------------|----|
| Amatrix[1,7] | NA |
| Amatrix[2,6] | NA |
| Amatrix[3,6] | NA |
| Amatrix[4,6] | NA |
| Amatrix[5,7] | NA |
| Smatrix[1,1] | NA |
| Smatrix[2,2] | NA |
| Smatrix[3,3] | NA |
| Smatrix[4,4] | NA |
| Smatrix[5,5] | NA |

Smatrix[7,6] NA

Goodness-of-fit indices:

| | Value |
|---|-----------|
| Sample size | 796.0000 |
| Chi-square of target model | 85.9696 |
| DF of target model | 4.0000 |
| p value of target model | 0.0000 |
| Chi-square of independent model | 3125.1714 |
| DF of independent model | 10.0000 |
| No. of constraints imposed on "Smatrix" | 5.0000 |
| RMSEA | 0.0000 |
| SRMR | 0.0805 |
| TLI | 0.9342 |
| CFI | 0.9737 |
| AIC | 77.9696 |
| BIC | 59.2512 |

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

3.2. Random-effects model

TSSEM using a random-effects model may be requested by specifying the `method='REM'` argument in `tssem1()`. By default (`RE.diag.only=FALSE`), a positive definite covariance matrix among the random-effects is used. For practical reasons, e.g., there are not enough studies, it may not be feasible to estimate the full covariance matrix. A diagonal matrix on the random-effects may also be used by specifying `RE.diag.only=TRUE`.

The fit indices for fitting the structural model in Stage 2 are $\chi^2(4, N = 4496) = 8.28, p < .001$, CFI=0.9920, TLI=0.9801, SRMR=0.0154 and RMSEA=0.0154. This indicates that the model fits the data quite well.

```
R> random1 <- tssem1(Digman97$data, Digman97$n, method="REM", RE.diag.only=TRUE)
```

Running TSSEM1 (Random Effects Model) Analysis of Correlation Matrix

```
R> summary(random1)
```

Call:

```
meta(y = ES, v = acovR, RE.constraints = diag(x = paste(RE.startvalues,
  "*Tau2_", 1:no.es, "_", 1:no.es, sep = ""), nrow = no.es,
  ncol = no.es), model.name = model.name)
```

95% confidence intervals: z statistic approximation

Coefficients:

| | Estimate | Std.Error | lbound | ubound |
|------------|------------|------------|-------------|------------|
| Intercept1 | 0.05444806 | 0.06316914 | -0.06936118 | 0.17825730 |
| Intercept2 | 0.12867832 | 0.04174105 | 0.04686736 | 0.21048929 |

| | | | | |
|-------------|------------|------------|------------|------------|
| Intercept3 | 0.24064413 | 0.03220610 | 0.17752133 | 0.30376693 |
| Intercept4 | 0.44713459 | 0.03211664 | 0.38418713 | 0.51008204 |
| Intercept5 | 0.39981602 | 0.05455526 | 0.29288968 | 0.50674236 |
| Intercept6 | 0.44433496 | 0.04168028 | 0.36264310 | 0.52602681 |
| Intercept7 | 0.10138326 | 0.04681343 | 0.00963062 | 0.19313590 |
| Intercept8 | 0.43415276 | 0.04000886 | 0.35573683 | 0.51256868 |
| Intercept9 | 0.20732487 | 0.04973238 | 0.10985120 | 0.30479854 |
| Intercept10 | 0.19296057 | 0.04340498 | 0.10788838 | 0.27803277 |
| Tau2_1_1 | 0.05115982 | 0.02059752 | 0.01078943 | 0.09153022 |
| Tau2_2_2 | 0.01977637 | 0.00914599 | 0.00185056 | 0.03770217 |
| Tau2_3_3 | 0.01030041 | 0.00505945 | 0.00038407 | 0.02021675 |
| Tau2_4_4 | 0.01122092 | 0.00494563 | 0.00152767 | 0.02091418 |
| Tau2_5_5 | 0.03815848 | 0.01523929 | 0.00829002 | 0.06802693 |
| Tau2_6_6 | 0.02132564 | 0.00868725 | 0.00429894 | 0.03835235 |
| Tau2_7_7 | 0.02571721 | 0.01094038 | 0.00427445 | 0.04715997 |
| Tau2_8_8 | 0.01901265 | 0.00820035 | 0.00294026 | 0.03508504 |
| Tau2_9_9 | 0.02995569 | 0.01234177 | 0.00576625 | 0.05414512 |
| Tau2_10_10 | 0.02172540 | 0.00934584 | 0.00340789 | 0.04004291 |

| | z value | Pr(> z) |
|--|---------|----------|
|--|---------|----------|

| | | |
|-------------|---------|---------------|
| Intercept1 | 0.8619 | 0.388720 |
| Intercept2 | 3.0828 | 0.002051 ** |
| Intercept3 | 7.4720 | 7.905e-14 *** |
| Intercept4 | 13.9222 | < 2.2e-16 *** |
| Intercept5 | 7.3286 | 2.325e-13 *** |
| Intercept6 | 10.6606 | < 2.2e-16 *** |
| Intercept7 | 2.1657 | 0.030335 * |
| Intercept8 | 10.8514 | < 2.2e-16 *** |
| Intercept9 | 4.1688 | 3.062e-05 *** |
| Intercept10 | 4.4456 | 8.765e-06 *** |
| Tau2_1_1 | 2.4838 | 0.012999 * |
| Tau2_2_2 | 2.1623 | 0.030595 * |
| Tau2_3_3 | 2.0359 | 0.041763 * |
| Tau2_4_4 | 2.2689 | 0.023277 * |
| Tau2_5_5 | 2.5040 | 0.012281 * |
| Tau2_6_6 | 2.4548 | 0.014095 * |
| Tau2_7_7 | 2.3507 | 0.018740 * |
| Tau2_8_8 | 2.3185 | 0.020421 * |
| Tau2_9_9 | 2.4272 | 0.015217 * |
| Tau2_10_10 | 2.3246 | 0.020093 * |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 2694.379

Degrees of freedom of the Q statistic: 130

P value of the Q statistic: 0

Number of studies (or clusters): 14

```

Number of observed statistics: 140
Number of estimated parameters: 20
Degrees of freedom: 120
-2 log likelihood: -109.6846
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

```

R> ##### Extract the fixed-effects (pooled correlation matrix)
R> coef(random1, select="fixed")

```

| Intercept1 | Intercept2 | Intercept3 | Intercept4 | Intercept5 |
|------------|------------|------------|------------|-------------|
| 0.05444806 | 0.12867832 | 0.24064413 | 0.44713459 | 0.39981602 |
| Intercept6 | Intercept7 | Intercept8 | Intercept9 | Intercept10 |
| 0.44433496 | 0.10138326 | 0.43415276 | 0.20732487 | 0.19296057 |

```

R> ##### Extract the sampling variance covariance matrix
R> vcov(random1, select="fixed")

```

| | Intercept1 | Intercept2 | Intercept3 |
|-------------|--------------|---------------|---------------|
| Intercept1 | 3.988380e-03 | 1.069194e-04 | 1.208244e-04 |
| Intercept2 | 1.069194e-04 | 1.739234e-03 | 1.017262e-04 |
| Intercept3 | 1.208244e-04 | 1.017262e-04 | 1.032979e-03 |
| Intercept4 | 1.324164e-05 | 3.451108e-05 | 1.889352e-05 |
| Intercept5 | 3.569683e-05 | -5.731580e-06 | 3.230825e-06 |
| Intercept6 | 4.983666e-05 | 6.374661e-06 | -1.405957e-05 |
| Intercept7 | 1.239068e-04 | 3.989823e-05 | 4.672623e-05 |
| Intercept8 | 1.547122e-05 | 4.616225e-05 | 1.193891e-05 |
| Intercept9 | 4.208683e-05 | 1.194520e-04 | 3.805643e-05 |
| Intercept10 | 4.465454e-05 | 4.064810e-05 | 1.102273e-04 |

| | Intercept4 | Intercept5 | Intercept6 |
|-------------|---------------|---------------|---------------|
| Intercept1 | 1.324164e-05 | 3.569683e-05 | 4.983666e-05 |
| Intercept2 | 3.451108e-05 | -5.731580e-06 | 6.374661e-06 |
| Intercept3 | 1.889352e-05 | 3.230825e-06 | -1.405957e-05 |
| Intercept4 | 1.023029e-03 | 2.837136e-06 | 3.261140e-06 |
| Intercept5 | 2.837136e-06 | 2.971648e-03 | 4.882023e-05 |
| Intercept6 | 3.261140e-06 | 4.882023e-05 | 1.735263e-03 |
| Intercept7 | -7.762411e-06 | 3.692114e-05 | 2.871902e-05 |
| Intercept8 | 3.784316e-06 | 6.644398e-05 | 4.542279e-05 |
| Intercept9 | 1.192774e-05 | 5.077767e-07 | 4.444937e-06 |
| Intercept10 | 3.490287e-05 | 6.771715e-06 | -1.496549e-06 |

| | Intercept7 | Intercept8 | Intercept9 |
|------------|---------------|--------------|--------------|
| Intercept1 | 1.239068e-04 | 1.547122e-05 | 4.208683e-05 |
| Intercept2 | 3.989823e-05 | 4.616225e-05 | 1.194520e-04 |
| Intercept3 | 4.672623e-05 | 1.193891e-05 | 3.805643e-05 |
| Intercept4 | -7.762411e-06 | 3.784316e-06 | 1.192774e-05 |
| Intercept5 | 3.692114e-05 | 6.644398e-05 | 5.077767e-07 |
| Intercept6 | 2.871902e-05 | 4.542279e-05 | 4.444937e-06 |
| Intercept7 | 2.189113e-03 | 1.511042e-05 | 1.146071e-04 |

```

Intercept8  1.511042e-05  1.589621e-03  2.227663e-05
Intercept9  1.146071e-04  2.227663e-05  2.471764e-03
Intercept10 1.349437e-04  2.750578e-05  1.056950e-04
Intercept10
Intercept1  4.465454e-05
Intercept2  4.064810e-05
Intercept3  1.102273e-04
Intercept4  3.490287e-05
Intercept5  6.771715e-06
Intercept6 -1.496549e-06
Intercept7  1.349437e-04
Intercept8  2.750578e-05
Intercept9  1.056950e-04
Intercept10 1.882447e-03

```

```

R> ##### Extract the random-effects (variance component)
R> coef(random1, select="random")

```

```

Tau2_1_1  Tau2_2_2  Tau2_3_3  Tau2_4_4  Tau2_5_5
0.05115982 0.01977637 0.01030041 0.01122092 0.03815848
Tau2_6_6  Tau2_7_7  Tau2_8_8  Tau2_9_9  Tau2_10_10
0.02132564 0.02571721 0.01901265 0.02995569 0.02172540

```

```

R> random2 <- tssem2(random1, Amatrix=A1, Smatrix=S1, Fmatrix=F1, diag.constraints=TRUE, i

```

Running TSSEM2 (Random Effects Model) Analysis of Correlation Structure

```

R> summary(random2)

```

Call:

```

wls(Cov = pooledS, asyCov = asyCov, n = tssem1.obj$total.n, Amatrix = Amatrix,
    Smatrix = Smatrix, Fmatrix = Fmatrix, diag.constraints = diag.constraints,
    cor.analysis = cor.analysis, intervals.type = intervals.type,
    model.name = model.name, suppressWarnings = suppressWarnings)

```

95% confidence intervals: Likelihood-based statistic

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z | value |
|--------------|----------|-----------|---------|---------|---|-------|
| Amatrix[1,7] | 0.69030 | NA | 0.56049 | 0.86627 | | NA |
| Amatrix[2,6] | 0.57707 | NA | 0.47802 | 0.68181 | | NA |
| Amatrix[3,6] | 0.59498 | NA | 0.49540 | 0.70017 | | NA |
| Amatrix[4,6] | 0.77087 | NA | 0.66020 | 0.90519 | | NA |
| Amatrix[5,7] | 0.64777 | NA | 0.51558 | 0.79609 | | NA |
| Smatrix[1,1] | 0.52348 | NA | 0.24905 | 0.68587 | | NA |
| Smatrix[2,2] | 0.66699 | NA | 0.53508 | 0.77149 | | NA |
| Smatrix[3,3] | 0.64600 | NA | 0.50971 | 0.75458 | | NA |


```

Smatrix[4,4] 0.40576      NA 0.18035 0.56417      NA
Smatrix[5,5] 0.58039      NA 0.36603 0.73418      NA
Smatrix[7,6] 0.39476      NA 0.30442 0.49078      NA
      Pr(>|z|)
Amatrix[1,7]      NA
Amatrix[2,6]      NA
Amatrix[3,6]      NA
Amatrix[4,6]      NA
Amatrix[5,7]      NA
Smatrix[1,1]      NA
Smatrix[2,2]      NA
Smatrix[3,3]      NA
Smatrix[4,4]      NA
Smatrix[5,5]      NA
Smatrix[7,6]      NA

```

Goodness-of-fit indices:

| | Value |
|---|-----------|
| Sample size | 4496.0000 |
| Chi-square of target model | 8.2809 |
| DF of target model | 4.0000 |
| p value of target model | 0.0818 |
| Chi-square of independent model | 546.8075 |
| DF of independent model | 10.0000 |
| No. of constraints imposed on "Smatrix" | 5.0000 |
| RMSEA | 0.0000 |
| SRMR | 0.0465 |
| TLI | 0.9801 |
| CFI | 0.9920 |
| AIC | 0.2809 |
| BIC | -25.3628 |

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

4. Other Useful Functions

4.1. Analysis of Correlation/Covariance Structure with Weighted Least Squares

The `wls()` function may be used to fit a correlation/covariance structure with weighted least squares (WLS) estimation method. The following example fits a one-factor CFA model on the correlation matrix with WLS estimation method. It should be noted that the only off-diagonal elements are used when a correlation structure is fitted.

```

R> ##### Analysis of correlation structure
R> R1 <- matrix(c(1.00, 0.22, 0.24, 0.18,

```

```

0.22, 1.00, 0.30, 0.22,
0.24, 0.30, 1.00, 0.24,
0.18, 0.22, 0.24, 1.00), ncol=4, nrow=4)

R> n <- 1000
R> acovR1 <- asyCov(R1, n)
R> ## One-factor CFA model
R> (A1 <- cbind(matrix(0, nrow=5, ncol=4),
                matrix(c("0.2*a1", "0.2*a2", "0.2*a3", "0.2*a4", 0),
                        ncol=1)))

      [,1] [,2] [,3] [,4] [,5]
[1,] "0"  "0"  "0"  "0"  "0.2*a1"
[2,] "0"  "0"  "0"  "0"  "0.2*a2"
[3,] "0"  "0"  "0"  "0"  "0.2*a3"
[4,] "0"  "0"  "0"  "0"  "0.2*a4"
[5,] "0"  "0"  "0"  "0"  "0"

R> A1 <- as.mxMatrix(A1)
R> (S1 <- diag(c("0.2*e1", "0.2*e2", "0.2*e3", "0.2*e4", 1)))

      [,1] [,2] [,3] [,4] [,5]
[1,] "0.2*e1" "0"  "0"  "0"  "0"
[2,] "0"      "0.2*e2" "0"  "0"  "0"
[3,] "0"      "0"      "0.2*e3" "0"  "0"
[4,] "0"      "0"      "0"      "0.2*e4" "0"
[5,] "0"      "0"      "0"      "0"      "1"

R> S1 <- as.mxMatrix(S1)
R> ## The first 4 variables are observed while the last one is latent.
R> (F1 <- create.Fmatrix(c(1,1,1,1,0), name="F1"))

FullMatrix 'F1'

@labels: No labels assigned.

@values
      [,1] [,2] [,3] [,4] [,5]
[1,] 1 0 0 0 0
[2,] 0 1 0 0 0
[3,] 0 0 1 0 0
[4,] 0 0 0 1 0

@free: No free parameters.

@lbound: No lower bounds assigned.

@ubound: No upper bounds assigned.

```

```
R> wls.fit1 <- wls(Cov=R1, asyCov=acovR1, n=n, Fmatrix=F1, Smatrix=S1, Amatrix=A1,
  cor.analysis=TRUE, diag.constraints=TRUE, intervals="LB")
```

Running WLS Analysis of Correlation Structure

```
R> summary(wls.fit1)
```

Call:

```
wls(Cov = R1, asyCov = acovR1, n = n, Amatrix = A1, Smatrix = S1,
  Fmatrix = F1, diag.constraints = TRUE, cor.analysis = TRUE,
  intervals.type = "LB")
```

95% confidence intervals: Likelihood-based statistic

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z | value |
|--------------|----------|-----------|---------|---------|---|-------|
| Amatrix[1,5] | 0.42159 | NA | 0.34632 | 0.49869 | | NA |
| Amatrix[2,5] | 0.52376 | NA | 0.44829 | 0.60309 | | NA |
| Amatrix[3,5] | 0.57092 | NA | 0.49431 | 0.65292 | | NA |
| Amatrix[4,5] | 0.42159 | NA | 0.34632 | 0.49869 | | NA |
| Smatrix[1,1] | 0.82226 | NA | 0.75131 | 0.88005 | | NA |
| Smatrix[2,2] | 0.72567 | NA | 0.63627 | 0.79903 | | NA |
| Smatrix[3,3] | 0.67405 | NA | 0.57367 | 0.75566 | | NA |
| Smatrix[4,4] | 0.82226 | NA | 0.75131 | 0.88013 | | NA |

| | Pr(> z) |
|--------------|----------|
| Amatrix[1,5] | NA |
| Amatrix[2,5] | NA |
| Amatrix[3,5] | NA |
| Amatrix[4,5] | NA |
| Smatrix[1,1] | NA |
| Smatrix[2,2] | NA |
| Smatrix[3,3] | NA |
| Smatrix[4,4] | NA |

Goodness-of-fit indices:

| | Value |
|---|-----------|
| Sample size | 1000.0000 |
| Chi-square of target model | 0.0134 |
| DF of target model | 2.0000 |
| p value of target model | 0.9933 |
| Chi-square of independent model | 243.9817 |
| DF of independent model | 6.0000 |
| No. of constraints imposed on "Smatrix" | 4.0000 |
| RMSEA | 0.0000 |
| SRMR | 0.0012 |
| TLI | 1.0250 |
| CFI | 1.0000 |
| AIC | -3.9866 |

```

BIC -13.8021
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

R> ## Extract the parameter estimates
R> coef(wls.fit1)

Amatrix[1,5] Amatrix[2,5] Amatrix[3,5] Amatrix[4,5]
      0.4215923      0.5237644      0.5709210      0.4215923
Smatrix[1,1] Smatrix[2,2] Smatrix[3,3] Smatrix[4,4]
      0.8222599      0.7256709      0.6740492      0.8222599

R> ## Extract the sampling variance covariance matrix
R> vcov(wls.fit1)

[1] NA

```

4.2. Likelihood-based Confidence Intervals

Most CIs are based on the estimated standard errors. These are known as Wald CIs. Wald CIs are symmetric around the estimates. The Wald CIs might be outside of the meaningful boundaries, for example, a negative lower limit for the variance or larger than 1 for a correlation coefficient. A preferable approach is to construct the CIs based on the likelihood. This is known as the likelihood based CI (Cheung 2009a; Neale and Miller 1997). Likelihood based CIs on the parameter estimates can be required by specifying `intervals.type='LB'` argument.

```

R> ## Random-effects meta-analysis with ML
R> summary( meta(y=di, v=vi, data=Becker83, intervals.type="LB") )

```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, data = Becker83, intervals.type = "LB")
```

95% confidence intervals: Likelihood-based statistic

Coefficients:

| | Estimate | Std.Error | lbound | ubound | z value |
|------------|----------|-----------|-----------|----------|---------|
| Intercept1 | 0.174734 | 0.113378 | -0.052165 | 0.437627 | 1.5412 |
| Tau2_1_1 | 0.077376 | 0.054108 | 0.015124 | 0.302999 | 1.4300 |

Pr(>|z|)

| | |
|------------|--------|
| Intercept1 | 0.1233 |
| Tau2_1_1 | 0.1527 |

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

```

Number of studies (or clusters): 10
Number of observed statistics: 10
Number of estimated parameters: 2
Degrees of freedom: 8
-2 log likelihood: 7.928307
OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

```

R> ## Mixed-effects meta-analysis with "log(items)" as a predictor
R> summary( meta(y=di, v=vi, x=log(items), data=Becker83, intervals.type="LB") )

```

Running Meta analysis with ML

Call:

```
meta(y = di, v = vi, x = log(items), data = Becker83, intervals.type = "LB")
```

95% confidence intervals: Likelihood-based statistic

Coefficients:

| | Estimate | Std.Error | lbound | ubound |
|------------|-------------|------------|-------------|-------------|
| Intercept1 | -3.2015e-01 | 1.0981e-01 | -5.4408e-01 | -7.7598e-02 |
| Slope1_1 | 2.1088e-01 | 4.5084e-02 | 1.1838e-01 | 3.0789e-01 |
| Tau2_1_1 | 1.0000e-10 | 2.0095e-02 | 9.9937e-11 | 5.7947e-02 |

| | z value | Pr(> z) |
|------------|---------|---------------|
| Intercept1 | -2.9154 | 0.003552 ** |
| Slope1_1 | 4.6774 | 2.905e-06 *** |
| Tau2_1_1 | 0.0000 | 1.000000 |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Q statistic on homogeneity of effect sizes: 30.64949

Degrees of freedom of the Q statistic: 9

P value of the Q statistic: 0.0003399239

```

Number of studies (or clusters): 10

```

```

Number of observed statistics: 10

```

```

Number of estimated parameters: 3

```

```

Degrees of freedom: 7

```

```

-2 log likelihood: -4.208024

```

```

OpenMx status1: 0 ("0" and "1": considered fine; other values indicate problems)

```

4.3. Reading External Data Files

Data sets are most likely stored externally. `metaSEM` reads three types of data formats. The first type is full correlation/covariance matrices, for example, `fullmat.dat` is the same as the built-in data set `Cheung09`. Missing values are represented by `NA` (the default option). Suppose you have saved it at `d:\fullmat.dat`, you may read it by using the following command in R:

```
my.df <- readFullMat(file="d:/fullmat.dat")
```

The second type is lower triangle correlation/covariance matrices, for example, `lowertriangle.dat`. Missing values are represented by the strings NA. Suppose you have saved it at `d:\lowertriangle.dat`, you may read it by using the following command in R:

```
my.df <- readLowTriMat(file = "d:/lowertriangle.dat", no.var = 9, na.strings="NA")
```

The third type is vectors of correlation/covariance elements based on column vectorization. One row represents one study, for example, `stackvec.dat`. Suppose you have saved it at `d:\stackvec.dat`, you may read it by using the following R command:

```
my.df <- readStackVec(file="d:/stackvec.dat")
```

5. Installation

First of all, you need R to run it. Since metaSEM uses OpenMx as the workhorse, OpenMx has to be installed first. To install OpenMx, run the following command inside an R session:

```
install.packages('OpenMx', repos='http://openmx.psyc.virginia.edu/packages/')
```

See <http://openmx.psyc.virginia.edu/installing-openmx> for the details on how to install OpenMx. Moreover, metaSEM also depends on the ellipse package that can be installed by the following command inside an R session:

```
install.packages('ellipse')
```

5.1. Windows platform

Download the **Windows binary** of metaSEM. If the file is saved at `d:\`. Run the following command inside an R session:

```
install.packages(pkgs="d:/metaSEM_0.7-1.zip", repos=NULL)
```

Please note that `d:\` in Windows is represented by either `d:/` or `d:\\` in R.

5.2. Linux and Mac OS X platform

Download the **source package** of metaSEM. Run the following command (as Root) inside an R session:

```
install.packages(pkgs="metaSEM_0.7-1.zip", repos=NULL, type="source")
```

6. Acknowledgements

This package cannot be written without R and **OpenMx**. Contributions by the R Development Core Team and the OpenMx Core Development Team are highly appreciated.

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