

# Privacy Amplification

**Borja Balle**

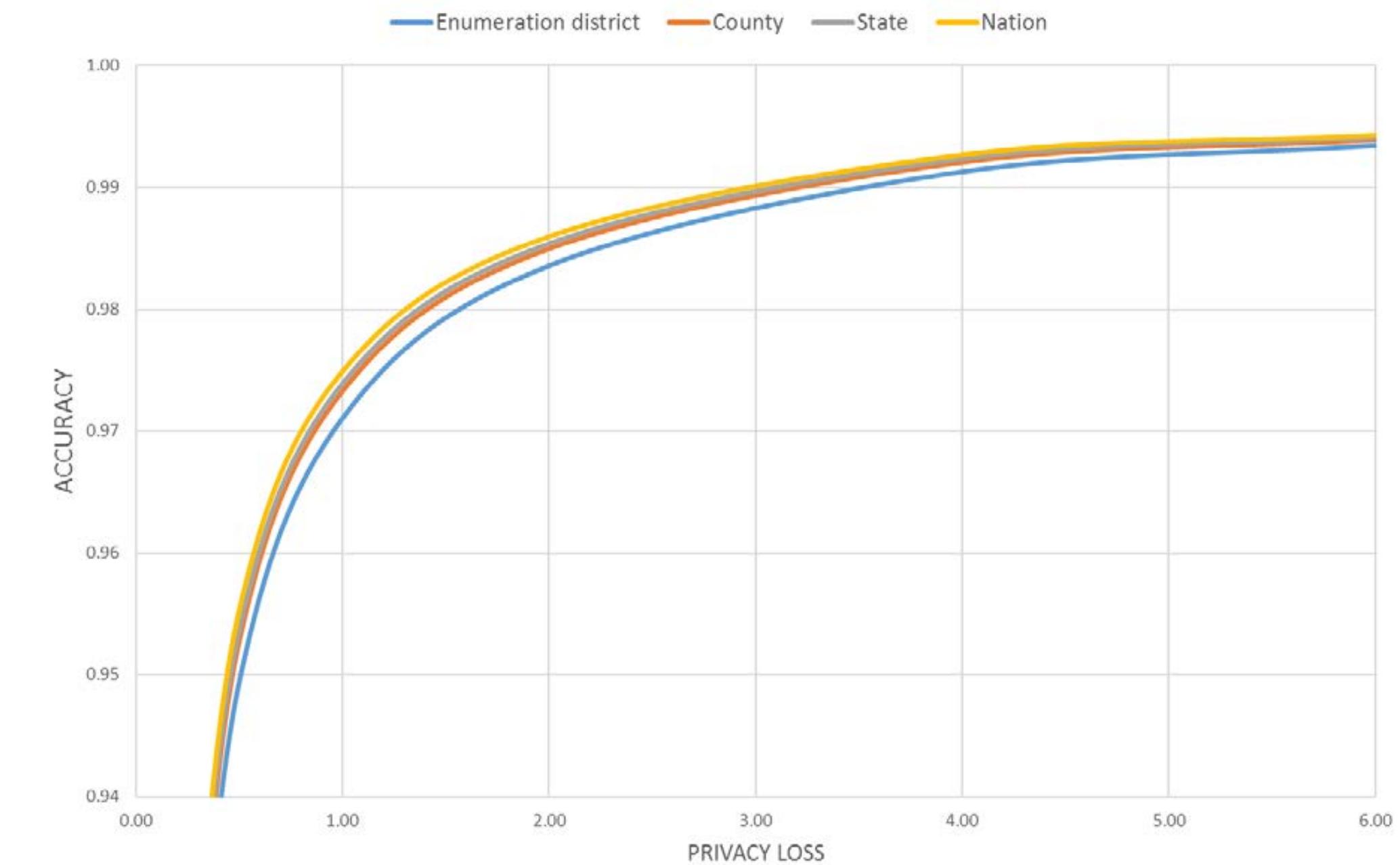
# A Fundamental Trade-off



# Bridging Theory and Practice

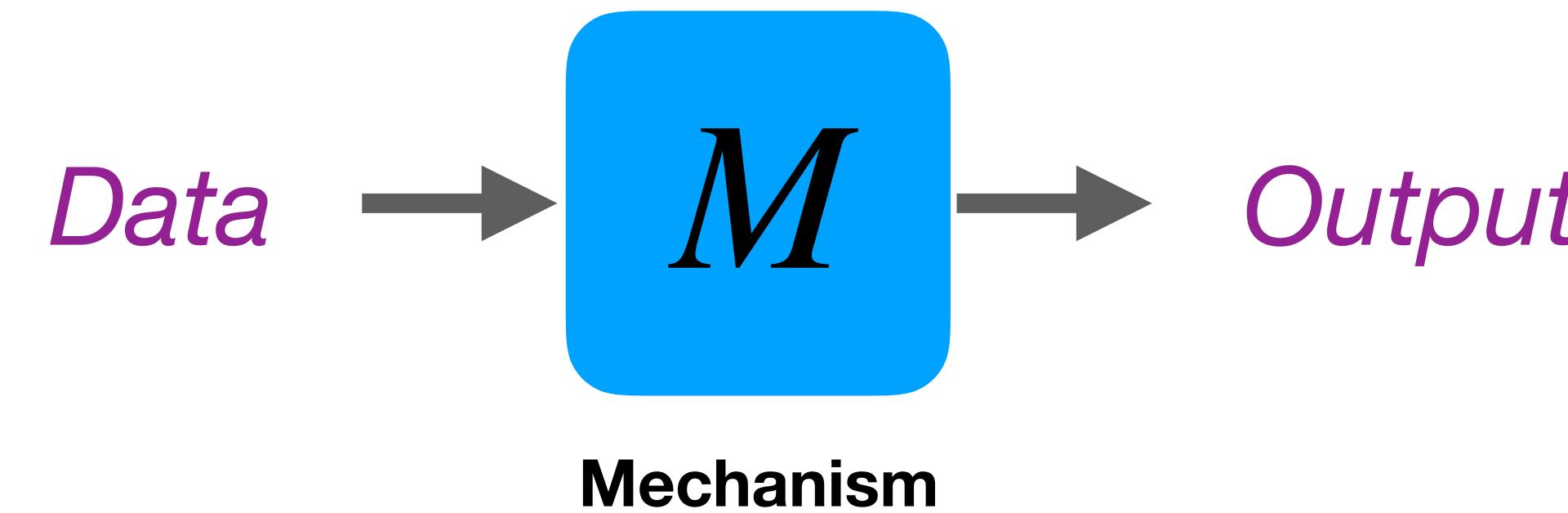
$$O\left(\frac{\sqrt{d \log(1/\delta)}}{n\epsilon}\right)$$

DISTRICT-BY-DISTRICT DIFFERENTIAL PRIVACY ALGORITHMS  
(1940 CENSUS DATA)

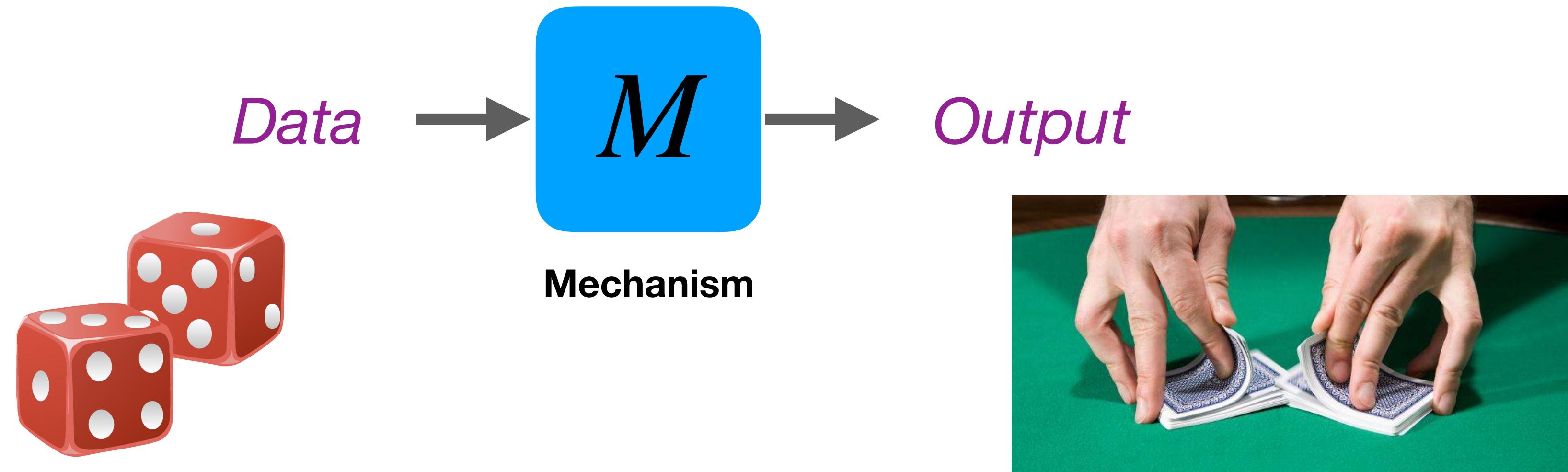


*Plot from J. M. Abowd “Disclosure Avoidance for Block Level Data and Protection of Confidentiality in Public Tabulations”  
(CSAC Meeting, December 2018)*

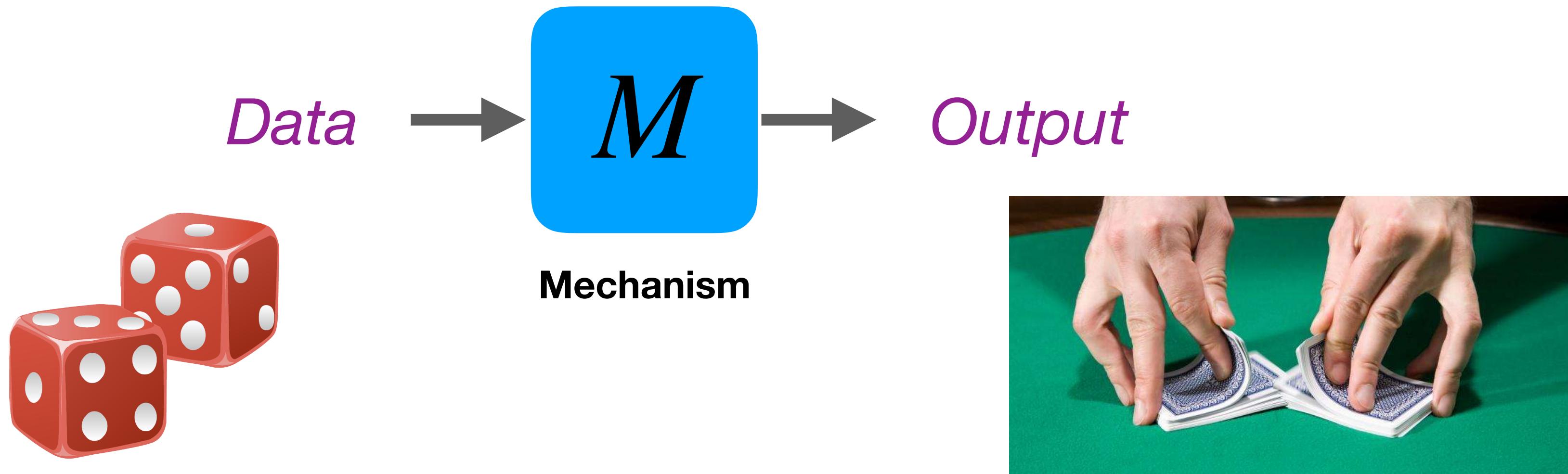
# Privacy Amplification



# Privacy Amplification



# Privacy Amplification



- Subsampling the input dataset [B, Barthe, Gaboardi NeurIPS'18]
- Shuffling outputs coming from different data records [B, Bell, Gascón, Nissim CRYPTO'19]
- Applying a stochastic post-processing to the output [B, Barthe, Gaboardi, Geumlek arXiv'19]

# Motivations

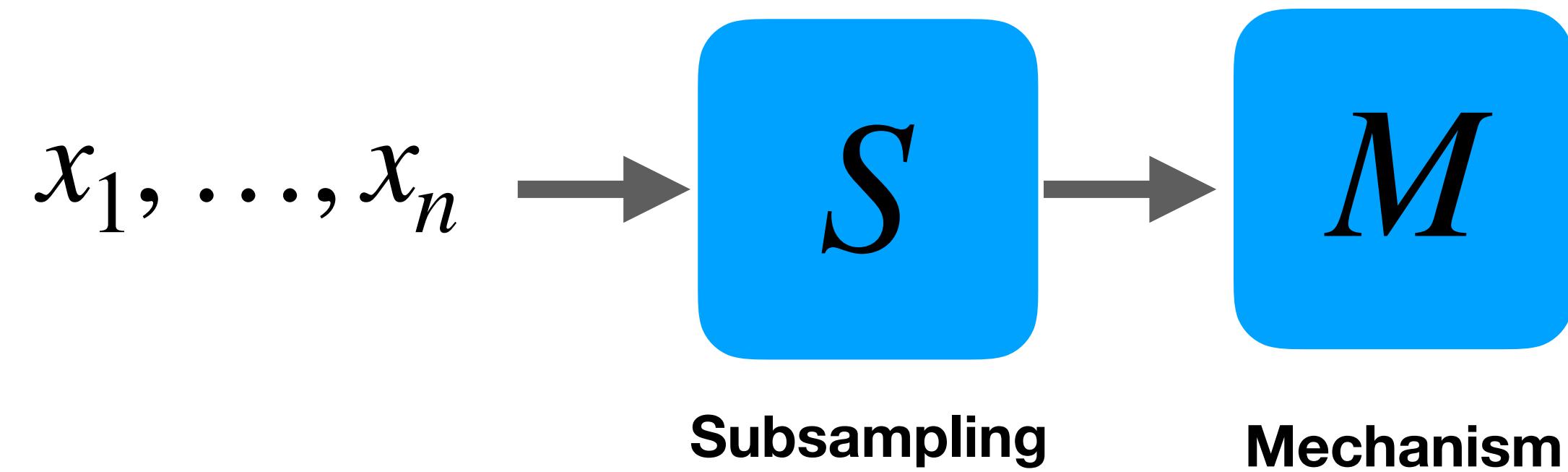
- **New building blocks** for DP mechanisms
- Account for privacy induced by **additional sources of randomness** in existing mechanisms
- **Better utility** through tighter privacy bounds
- Gain a **deeper understanding** of differential privacy
- **Cool, fun math** 😊

# In This Talk...

- **The method of overlapping mixtures**
  - Amplification by subsampling
  - Amplification by stochastic post-processing
- **Couplings beyond overlapping mixtures**
  - Amplification by shuffling
  - Amplification by iteration

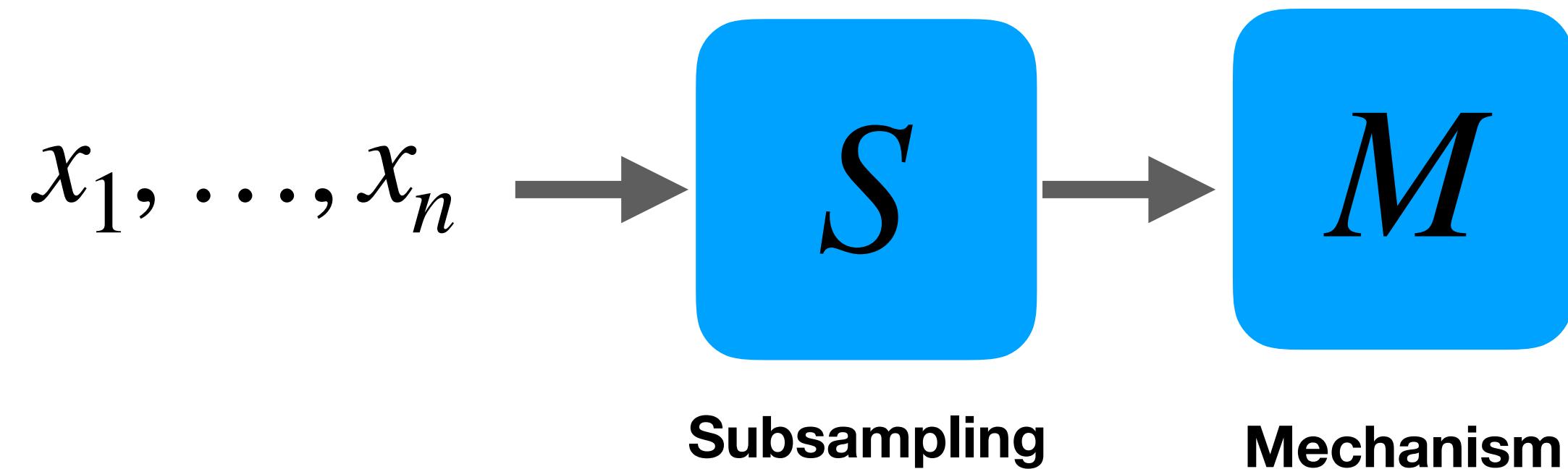
# The Method of Overlapping Mixtures

# Amplification by Subsampling



**Subsampling**  
Select a random  $\gamma$  fraction  
of the original dataset

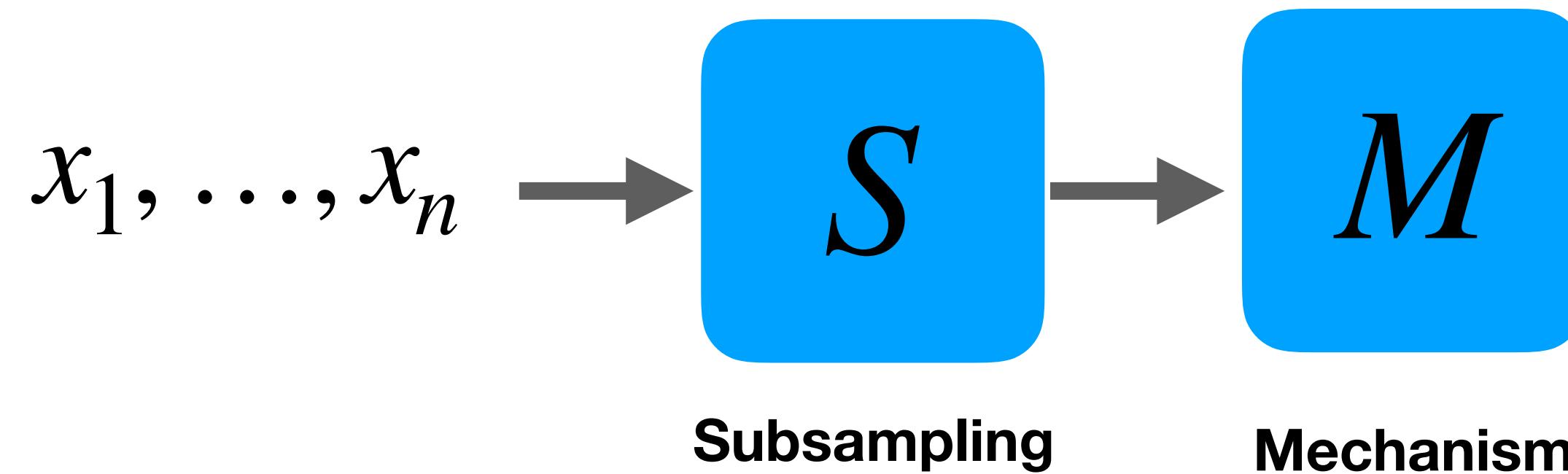
# Amplification by Subsampling



**Subsampling**  
Select a random  $\gamma$  fraction  
of the original dataset

- Secrecy of the sample: each  $x_i$  is used only w.p.  $\gamma$

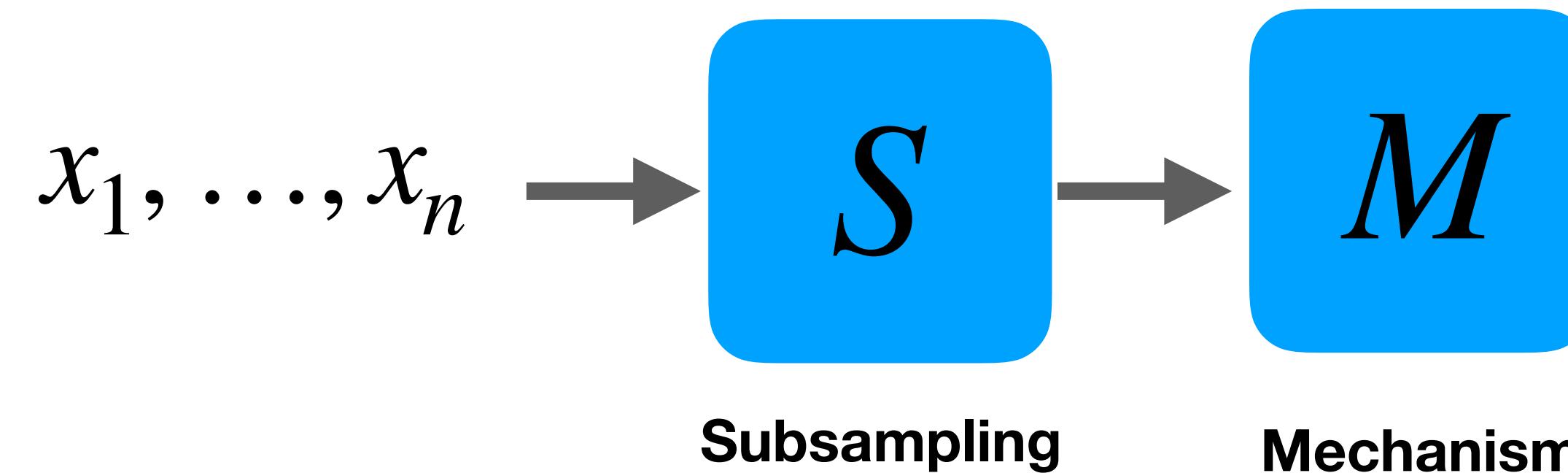
# Amplification by Subsampling



**Subsampling**  
Select a random  $\gamma$  fraction  
of the original dataset

- Secrecy of the sample: each  $x_i$  is used only w.p.  $\gamma$
- Used in: learning algorithms/theory, stochastic optimization, ...

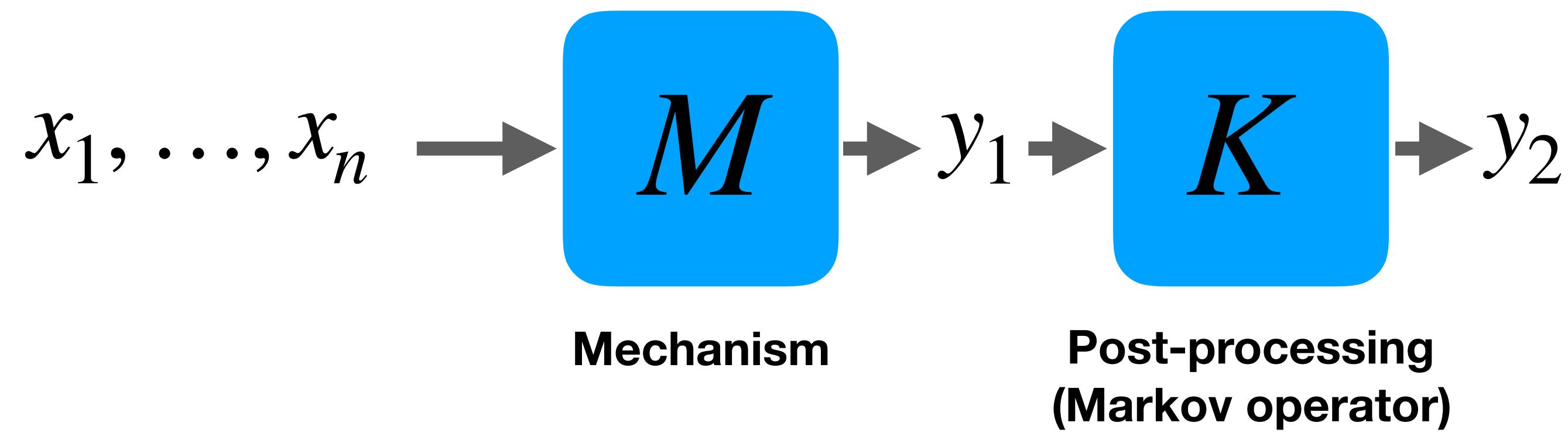
# Amplification by Subsampling



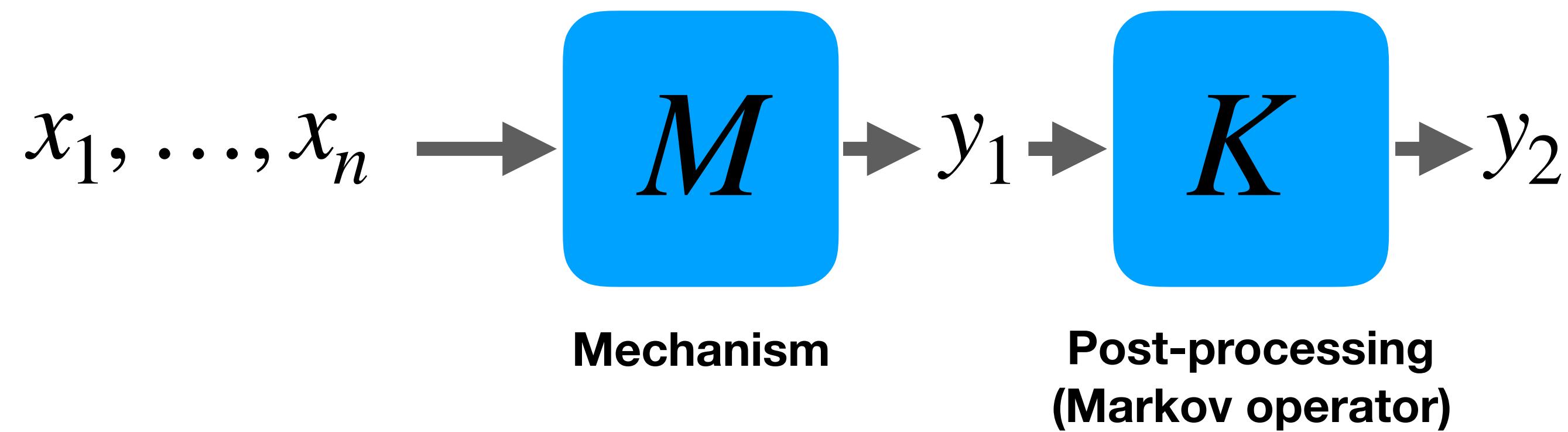
**Subsampling**  
Select a random  $\gamma$  fraction  
of the original dataset

- Secrecy of the sample: each  $x_i$  is used only w.p.  $\gamma$
- Used in: learning algorithms/theory, stochastic optimization, ...
- Agnostic to the internals of the mechanism

# Amplification by Postprocessing

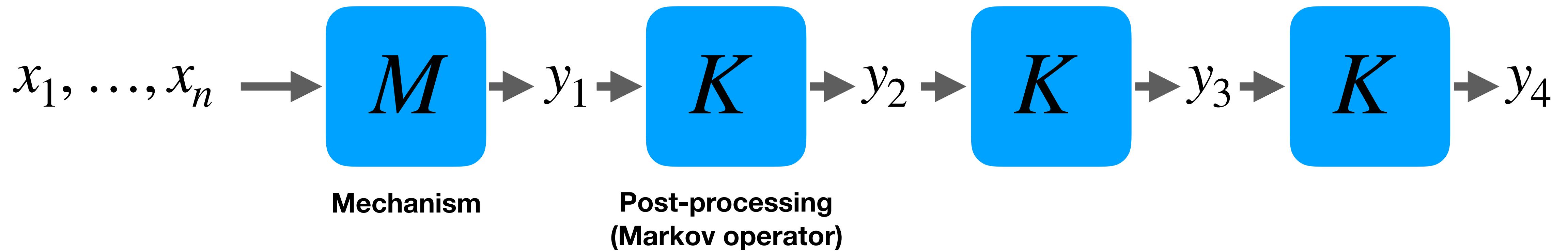


# Amplification by Postprocessing



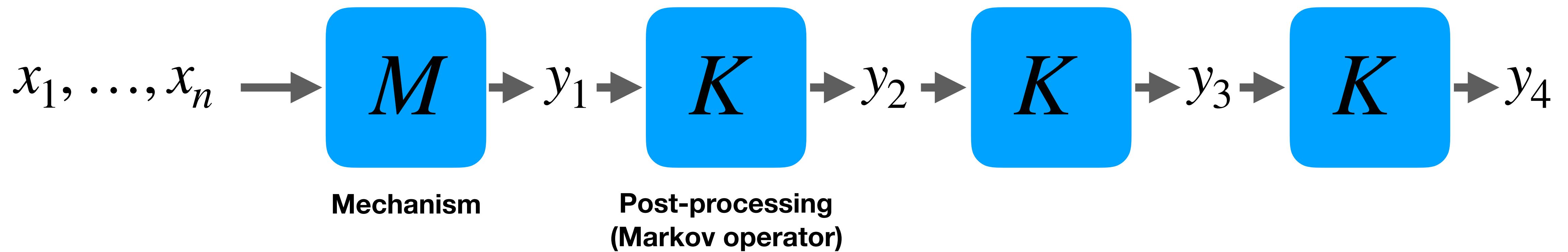
- When is  $K \circ M$  more private than  $M$ ?

# Amplification by Postprocessing



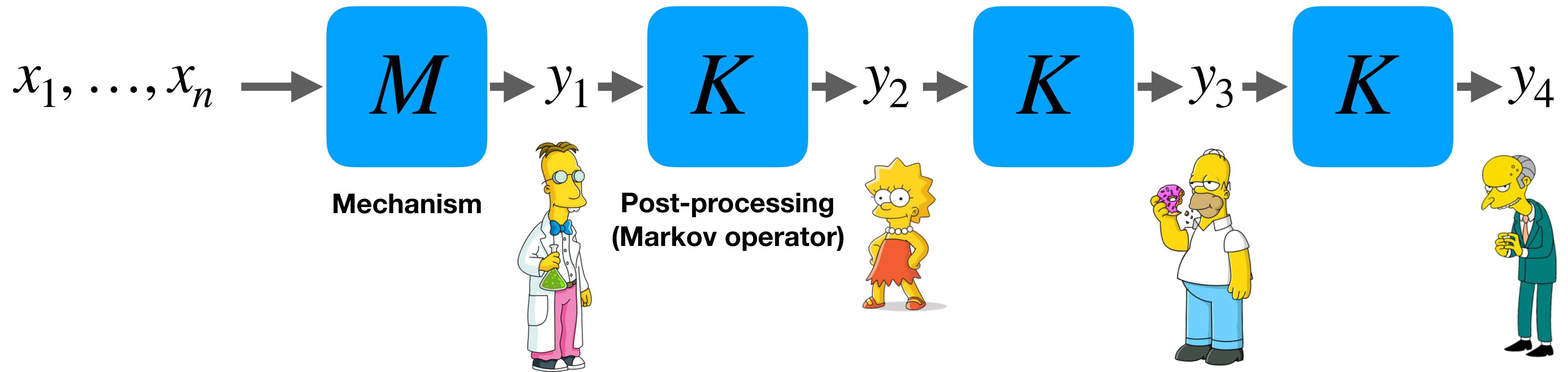
- When is  $K \circ M$  more private than  $M$ ?

# Amplification by Postprocessing



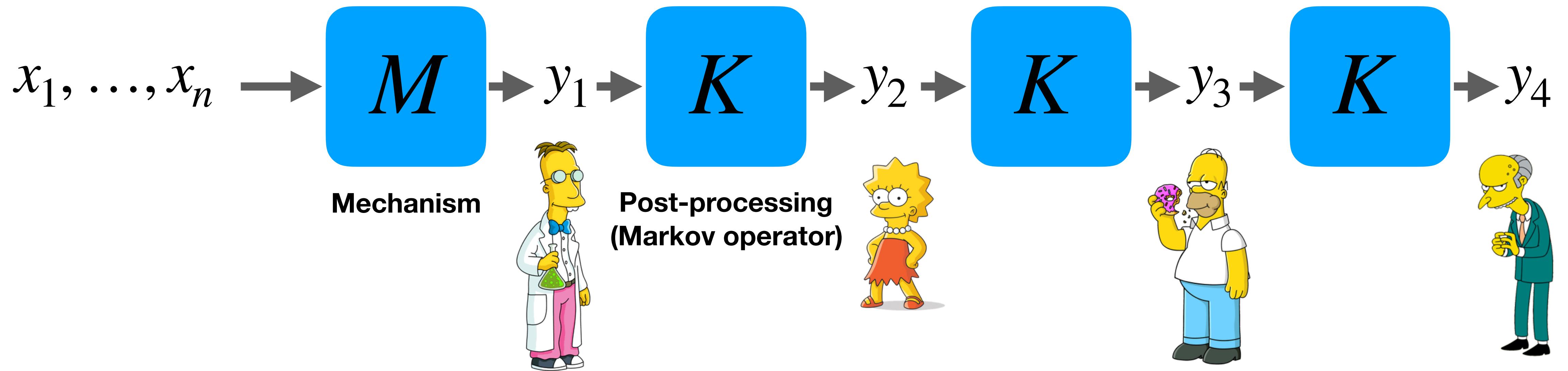
- When is  $K \circ M$  more private than  $M$ ?
- How does privacy relate to mixing in the Markov chain?

# Amplification by Postprocessing



- When is  $K \circ M$  more private than  $M$ ?
- How does privacy relate to mixing in the Markov chain?

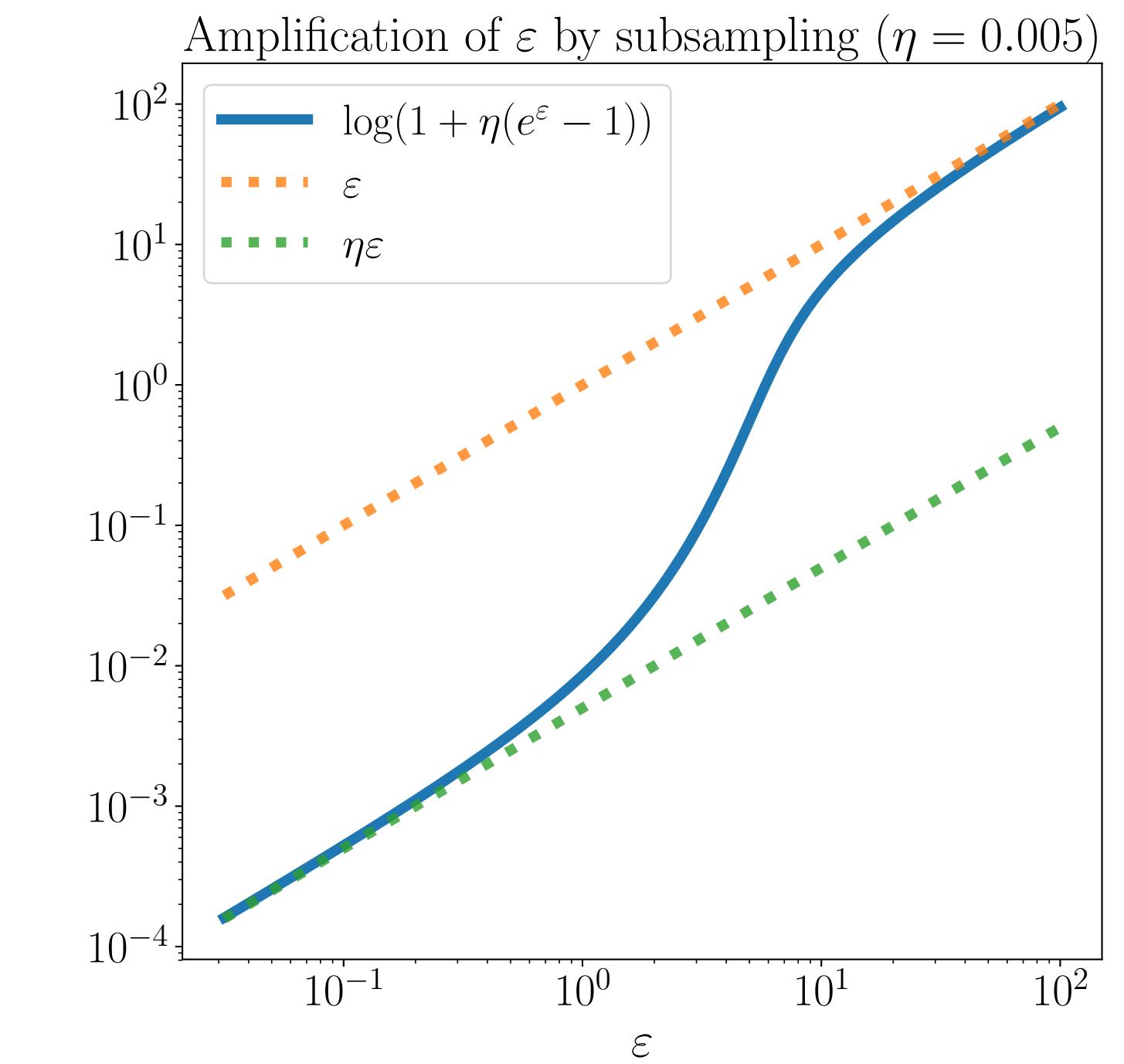
# Amplification by Postprocessing



- When is  $K \circ M$  more private than  $M$ ?
- How does privacy relate to mixing in the Markov chain?
- Starting point for “Hierarchical DP”

# Two Amplification Results

- Let  $M$  be an  $(\varepsilon, \delta)$ -DP mechanism between “replace one” datasets
- Let  $\varepsilon' = \log(1 + \gamma(e^\varepsilon - 1))$
- **Theorem:** If  $S$  subsamples a  $\gamma$  fraction w.o.r., then  $M \circ S$  is  $(\varepsilon', \gamma\delta)$ -DP
- **Theorem:** If  $K$  is  $\log(1/(1-\gamma))$ -LDP, then  $K \circ M$  is  $(\varepsilon', \gamma\delta e^{(\varepsilon' - \varepsilon)})$ -DP



# DP via Hockey-Stick Divergence

$$\sup_{x \simeq x'} \sup_E \left( \Pr[M(x) \in E] - e^\epsilon \Pr[M(x') \in E] \right) \leq \delta$$

# DP via Hockey-Stick Divergence

$$\sup_{x \simeq x'} \sup_E \left( \Pr[M(x) \in E] - e^\epsilon \Pr[M(x') \in E] \right) \leq \delta$$

||

$$D_{e^\epsilon}(M(x) \mid M(x')) = \int \left[ p_{M(x)}(y) - e^\epsilon p_{M(x')}(y) \right]_+ dy$$

# DP via Hockey-Stick Divergence

$$\sup_{x \simeq x'} \sup_E \left( \Pr[M(x) \in E] - e^\epsilon \Pr[M(x') \in E] \right) \leq \delta$$

||

$$D_{e^\epsilon}(M(x) \mid M(x')) = \int \left[ p_{M(x)}(y) - e^\epsilon p_{M(x')}(y) \right]_+ dy$$

$$D_{e^\epsilon}(M(x) \mid M(x')) \leq \delta$$

# DP via Hockey-Stick Divergence

$$\sup_{x \simeq x'} \sup_E (\Pr[M(x) \in E] - e^\epsilon \Pr[M(x') \in E]) \leq \delta$$

||

$$D_{e^\epsilon}(M(x) \mid M(x')) = \int [p_{M(x)}(y) - e^\epsilon p_{M(x')}(y)]_+ dy$$

$$D_{e^\epsilon}(M(x) \mid M(x')) \leq \delta$$

This is an f-divergence, so it satisfies:

- Data processing inequality (ie. postprocessing)
- Joint convexity

# Advanced Joint Convexity of DP

**Joint Convexity:**

$$\begin{aligned}\nu &= p_1\nu_1 + p_2\nu_2 \\ \mu &= p_1\mu_1 + p_2\mu_2\end{aligned} \Rightarrow D_{e^\epsilon}(\mu \mid \nu) \leq p_1D_{e^\epsilon}(\mu_1 \mid \nu_1) + p_2D_{e^\epsilon}(\mu_2 \mid \nu_2)$$

# Advanced Joint Convexity of DP

**Joint Convexity:**

$$\begin{aligned}\nu &= p_1\nu_1 + p_2\nu_2 \\ \mu &= p_1\mu_1 + p_2\mu_2\end{aligned} \qquad \Rightarrow \qquad D_{e^\epsilon}(\mu \mid \nu) \leq p_1 D_{e^\epsilon}(\mu_1 \mid \nu_1) + p_2 D_{e^\epsilon}(\mu_2 \mid \nu_2)$$

$$\mu_1 = \nu_1 \qquad \Rightarrow \qquad D_{e^\epsilon}(\mu \mid \nu) \leq p_2 D_{e^\epsilon}(\mu_2 \mid \nu_2)$$

# Advanced Joint Convexity of DP

**Joint Convexity:**

$$\begin{aligned}\nu &= p_1\nu_1 + p_2\nu_2 \\ \mu &= p_1\mu_1 + p_2\mu_2\end{aligned} \Rightarrow D_{e^\epsilon}(\mu | \nu) \leq p_1D_{e^\epsilon}(\mu_1 | \nu_1) + p_2D_{e^\epsilon}(\mu_2 | \nu_2)$$

$$\mu_1 = \nu_1 \Rightarrow D_{e^\epsilon}(\mu | \nu) \leq p_2D_{e^\epsilon}(\mu_2 | \nu_2)$$

**Advanced Joint Convexity:**  $\tilde{\epsilon} = \log(1 + p_2(e^\epsilon - 1))$

$$\mu_1 = \nu_1 \Rightarrow D_{e^{\tilde{\epsilon}}}(\mu | \nu) \leq p_2 \left( \left(1 - \frac{e^{\tilde{\epsilon}}}{e^\epsilon}\right) D_{e^\epsilon}(\mu_2 | \mu_1) + \frac{e^{\tilde{\epsilon}}}{e^\epsilon} D_{e^\epsilon}(\mu_2 | \nu_2) \right)$$

# Advanced Joint Convexity of DP

**Joint Convexity:**

$$\begin{aligned}\nu &= p_1\nu_1 + p_2\nu_2 \\ \mu &= p_1\mu_1 + p_2\mu_2\end{aligned} \Rightarrow D_{e^\epsilon}(\mu | \nu) \leq p_1D_{e^\epsilon}(\mu_1 | \nu_1) + p_2D_{e^\epsilon}(\mu_2 | \nu_2)$$

$$\mu_1 = \nu_1 \Rightarrow D_{e^\epsilon}(\mu | \nu) \leq p_2D_{e^\epsilon}(\mu_2 | \nu_2)$$

**Advanced Joint Convexity:**  $\tilde{\epsilon} = \log(1 + p_2(e^\epsilon - 1))$

$$\mu_1 = \nu_1 \Rightarrow D_{e^{\tilde{\epsilon}}}(\mu | \nu) \leq p_2 \left( \left( 1 - \frac{e^{\tilde{\epsilon}}}{e^\epsilon} \right) \underline{D_{e^\epsilon}(\mu_2 | \mu_1)} + \frac{e^{\tilde{\epsilon}}}{e^\epsilon} \underline{D_{e^\epsilon}(\mu_2 | \nu_2)} \right)$$

# Overlapping Mixtures in Subsampling

*Let  $M$  be  $(\varepsilon, \delta)$ -DP and  $S$  be subsampling w.o.r. of a  $\gamma$  fraction of the dataset:*

$$\Pr[M \circ S(x) \in E] = \Pr_{y \sim S(x)} [M(y) \in E]$$

# Overlapping Mixtures in Subsampling

*Let  $M$  be  $(\varepsilon, \delta)$ -DP and  $S$  be subsampling w.o.r. of a  $y$  fraction of the dataset:*

$$\Pr[M \circ S(x) \in E] = \Pr_{y \sim S(x)} [M(y) \in E]$$

*Let  $x$  and  $x'$  differ in one record:*

$$M \circ S(x) : \quad \Pr[x_n \notin y] \Pr[M(y) \in E \mid x_n \notin y] + \Pr[x_n \in y] \Pr[M(y) \in E \mid x_n \in y]$$

$$M \circ S'(x) : \quad \Pr[x'_n \notin y] \Pr[M(y) \in E \mid x'_n \notin y] + \Pr[x'_n \in y] \Pr[M(y) \in E \mid x'_n \in y]$$

# Overlapping Mixtures in Subsampling

Let  $M$  be  $(\varepsilon, \delta)$ -DP and  $S$  be subsampling w.o.r. of a  $\gamma$  fraction of the dataset:

$$\Pr[M \circ S(x) \in E] = \Pr_{y \sim S(x)} [M(y) \in E]$$

Let  $x$  and  $x'$  differ in one record:

$$M \circ S(x) : \quad \Pr[x_n \notin y] \Pr[M(y) \in E \mid x_n \notin y] + \Pr[x_n \in y] \Pr[M(y) \in E \mid x_n \in y]$$
$$p_1 = 1 - \gamma \qquad \qquad \qquad p_2 = \gamma$$
$$M \circ S'(x) : \quad \Pr[x'_n \notin y] \Pr[M(y) \in E \mid x'_n \notin y] + \Pr[x'_n \in y] \Pr[M(y) \in E \mid x'_n \in y]$$
$$\nu_1 \qquad \qquad \qquad \nu_2$$

# Overlapping Mixtures in Subsampling

Let  $M$  be  $(\varepsilon, \delta)$ -DP and  $S$  be subsampling w.o.r. of a  $\gamma$  fraction of the dataset:

$$\Pr[M \circ S(x) \in E] = \Pr_{y \sim S(x)} [M(y) \in E]$$

Let  $x$  and  $x'$  differ in one record:

$$M \circ S(x) : \quad \Pr[x_n \notin y] \Pr[M(y) \in E | x_n \notin y] + \Pr[x_n \in y] \Pr[M(y) \in E | x_n \in y]$$
$$p_1 = 1 - \gamma \qquad \qquad \qquad \mu_1 \qquad \qquad \qquad p_2 = \gamma \qquad \qquad \qquad \mu_2$$
$$M \circ S'(x') : \quad \Pr[x'_n \notin y] \Pr[M(y) \in E | x'_n \notin y] + \Pr[x'_n \in y] \Pr[M(y) \in E | x'_n \in y]$$
$$\nu_1 \qquad \qquad \qquad \qquad \qquad \qquad \qquad \nu_2$$

# Overlapping Mixtures in Subsampling

Let  $M$  be  $(\varepsilon, \delta)$ -DP and  $S$  be subsampling w.o.r. of a  $\gamma$  fraction of the dataset:

$$\Pr[M \circ S(x) \in E] = \Pr_{y \sim S(x)} [M(y) \in E]$$

Let  $x$  and  $x'$  differ in one record:

$$M \circ S(x) : \quad \Pr[x_n \notin y] \Pr[M(y) \in E | x_n \notin y] + \Pr[x_n \in y] \Pr[M(y) \in E | x_n \in y]$$

$$p_1 = 1 - \gamma \qquad \qquad \qquad \parallel \qquad \qquad \qquad p_2 = \gamma$$

$$M \circ S'(x') : \quad \Pr[x'_n \notin y] \Pr[M(y) \in E | x'_n \notin y] + \Pr[x'_n \in y] \Pr[M(y) \in E | x'_n \in y]$$

$$\nu_1 \qquad \qquad \qquad \qquad \qquad \qquad \qquad \nu_2$$

Bound the divergences:

$$y_2 \sim S_{\gamma n-1}(x \setminus x_n) \cup \{x_n\}$$

$$y_1 \sim S_{\gamma n}(x \setminus x_n)$$

Coupling

$$y_2 \simeq y_1$$

Joint Convexity

$$\Rightarrow D_{e^\epsilon}(\mu_2 \mid \mu_1) \leq \delta$$

# Overlapping Mixtures in Subsampling

Let  $M$  be  $(\varepsilon, \delta)$ -DP and  $S$  be subsampling w.o.r. of a  $\gamma$  fraction of the dataset:

$$\Pr[M \circ S(x) \in E] = \Pr_{y \sim S(x)} [M(y) \in E]$$

Let  $x$  and  $x'$  differ in one record:

$$M \circ S(x) : \quad \Pr[x_n \notin y] \Pr[M(y) \in E | x_n \notin y] + \Pr[x_n \in y] \Pr[M(y) \in E | x_n \in y]$$

$$p_1 = 1 - \gamma \qquad \qquad \qquad \parallel \qquad \qquad \qquad p_2 = \gamma$$

$$M \circ S'(x') : \quad \Pr[x'_n \notin y] \Pr[M(y) \in E | x'_n \notin y] + \Pr[x'_n \in y] \Pr[M(y) \in E | x'_n \in y]$$

$$\nu_1 \qquad \qquad \qquad \qquad \qquad \qquad \qquad \nu_2$$

Bound the divergences:

$$y_2 \sim S_{\gamma n-1}(x \setminus x_n) \cup \{x_n\}$$

$$y_1 \sim S_{\gamma n}(x \setminus x_n)$$

Coupling

$$y_2 \simeq y_1 \quad \Rightarrow \quad D_{e^\epsilon}(\mu_2 | \mu_1) \leq \delta$$

$$y_2 \sim S_{\gamma n-1}(x \setminus x_n) \cup \{x_n\}$$

$$y'_2 \sim S_{\gamma n-1}(x' \setminus x'_n) \cup \{x'_n\}$$

Coupling

$$y_2 \simeq y'_2 \quad \Rightarrow \quad D_{e^\epsilon}(\mu_2 | \nu_2) \leq \delta$$

Joint Convexity

# Overlapping Mixtures in Postprocessing

Let  $K$  be  $\log(1/(1-\gamma))$ -LDP. For any  $\omega$  in the image of  $K$ : [DLM'03]

$$\Pr[K \circ M(x) \in E] = (1 - \gamma)\omega(E) + \gamma \Pr[\tilde{K} \circ M(x) \in E] \quad \text{with} \quad \tilde{K} = \frac{K - (1 - \gamma)\omega}{\gamma}$$

# Overlapping Mixtures in Postprocessing

Let  $K$  be  $\log(1/(1-\gamma))$ -LDP. For any  $\omega$  in the image of  $K$ : [DLM'03]

$$\Pr[K \circ M(x) \in E] = (1 - \gamma)\omega(E) + \gamma \Pr[\tilde{K} \circ M(x) \in E] \quad \text{with} \quad \tilde{K} = \frac{K - (1 - \gamma)\omega}{\gamma}$$

Applying advanced joint convexity:

$$D_{e^{\tilde{\epsilon}}}(K \circ M(x) \mid K \circ M(x')) \leq \gamma \left( \left( 1 - \frac{e^{\tilde{\epsilon}}}{e^\epsilon} \right) D_{e^\epsilon}(\tilde{K} \circ M(x) \mid \omega) + \frac{e^{\tilde{\epsilon}}}{e^\epsilon} D_{e^\epsilon}(\tilde{K} \circ M(x) \mid \tilde{K} \circ M(x')) \right)$$

# Overlapping Mixtures in Postprocessing

Let  $K$  be  $\log(1/(1-\gamma))$ -LDP. For any  $\omega$  in the image of  $K$ : [DLM'03]

$$\Pr[K \circ M(x) \in E] = (1 - \gamma)\omega(E) + \gamma \Pr[\tilde{K} \circ M(x) \in E] \quad \text{with} \quad \tilde{K} = \frac{K - (1 - \gamma)\omega}{\gamma}$$

Applying advanced joint convexity:

$$D_{e^{\tilde{\epsilon}}}(K \circ M(x) \mid K \circ M(x')) \leq \gamma \left( \left( 1 - \frac{e^{\tilde{\epsilon}}}{e^\epsilon} \right) D_{e^\epsilon}(\tilde{K} \circ M(x) \mid \omega) + \frac{e^{\tilde{\epsilon}}}{e^\epsilon} \cancel{D_{e^\epsilon}(\tilde{K} \circ M(x) \mid \tilde{K} \circ M(x'))} \right) \leq \delta$$

# Overlapping Mixtures in Postprocessing

Let  $K$  be  $\log(1/(1-\gamma))$ -LDP. For any  $\omega$  in the image of  $K$ : [DLM'03]

$$\Pr[K \circ M(x) \in E] = (1 - \gamma)\omega(E) + \gamma \Pr[\tilde{K} \circ M(x) \in E] \quad \text{with} \quad \tilde{K} = \frac{K - (1 - \gamma)\omega}{\gamma}$$

Applying advanced joint convexity:

$$D_{e^{\tilde{\epsilon}}}(K \circ M(x) \mid K \circ M(x')) \leq \gamma \left( \left( 1 - \frac{e^{\tilde{\epsilon}}}{e^\epsilon} \right) D_{e^\epsilon}(\tilde{K} \circ M(x) \mid \omega) + \frac{e^{\tilde{\epsilon}}}{e^\epsilon} D_{e^\epsilon}(\tilde{K} \circ M(x) \mid \tilde{K} \circ M(x')) \right) \leq \delta$$

Taking  $\omega = K \circ M(x)$ :

$$\tilde{K} \circ M(x) = \frac{K \circ M(x) - (1 - \gamma)K \circ M(x)}{\gamma} = K \circ M(x)$$

[BBGG'19]

# Overlapping Mixtures in Postprocessing

Let  $K$  be  $\log(1/(1-\gamma))$ -LDP. For any  $\omega$  in the image of  $K$ : [DLM'03]

$$\Pr[K \circ M(x) \in E] = (1 - \gamma)\omega(E) + \gamma \Pr[\tilde{K} \circ M(x) \in E] \quad \text{with} \quad \tilde{K} = \frac{K - (1 - \gamma)\omega}{\gamma}$$

Applying advanced joint convexity:

$$D_{e^{\tilde{\epsilon}}}(K \circ M(x) \mid K \circ M(x')) \leq \gamma \left( \left(1 - \frac{e^{\tilde{\epsilon}}}{e^\epsilon}\right) D_{e^\epsilon}(\tilde{K} \circ M(x) \mid \omega) + \frac{e^{\tilde{\epsilon}}}{e^\epsilon} D_{e^\epsilon}(\tilde{K} \circ M(x) \mid \tilde{K} \circ M(x')) \right)$$
$$= 0 \leq \delta$$

Taking  $\omega = K \circ M(x)$ :

$$\tilde{K} \circ M(x) = \frac{K \circ M(x) - (1 - \gamma)K \circ M(x)}{\gamma} = K \circ M(x)$$

# Further Results

Subsampling	$\simeq_Y$	$\simeq_X$	$\eta$	$\delta'$	Theorem
Poisson( $\gamma$ )	R	R	$\gamma$	$\gamma\delta$	13
WOR( $n,m$ )	S	S	$\frac{m}{n}$	$\frac{m}{n}\delta$	14
WR( $n,m$ )	S	S	$1 - \left(1 - \frac{1}{n}\right)^m$	$\sum_{k=1}^m \binom{m}{k} \left(\frac{1}{n}\right)^k \left(1 - \frac{1}{n}\right)^{m-k} \delta_k$	15
WR( $n,m$ )	S	R	$1 - \left(1 - \frac{1}{n}\right)^m$	$\sum_{k=1}^m \binom{m}{k} \left(\frac{1}{n}\right)^k \left(1 - \frac{1}{n}\right)^{m-k} \delta_k$	16

**Amplification by  
subsampling in Rényi DP**

[WBK'19; ZW'19]

TABLE 1. Summary of privacy amplification bounds. Amplification parameter  $\eta$ :  $e^{\varepsilon'} = 1 + \eta(e^\varepsilon - 1)$ . Types of subsampling: without replacement (WOR) and with replacement (WR). Neighbouring relations: remove/add-one (R) and substitute one (S).

[BBG'18]

**Theorem 1.** Let  $M$  be an  $(\varepsilon, \delta)$ -DP mechanism. For a given Markov operator  $K$ , the post-processed mechanism  $K \circ M$  satisfies:

- (1)  $(\varepsilon, \delta')$ -DP with  $\delta' = \gamma\delta$  if  $K$  is  $\gamma$ -Dobrushin,
- (2)  $(\varepsilon, \delta')$ -DP with  $\delta' = \gamma\delta$  if  $K$  is  $(\gamma, \tilde{\varepsilon})$ -Dobrushin with<sup>3</sup>  $\tilde{\varepsilon} = \log(1 + \frac{e^\varepsilon - 1}{\delta})$ ,
- (3)  $(\varepsilon', \delta')$ -DP with  $\varepsilon' = \log(1 + \gamma(e^\varepsilon - 1))$  and  $\delta' = \gamma(1 - e^{\varepsilon' - \varepsilon}(1 - \delta))$  if  $K$  is  $\gamma$ -Doeblin,
- (4)  $(\varepsilon', \delta')$ -DP with  $\varepsilon' = \log(1 + \gamma(e^\varepsilon - 1))$  and  $\delta' = \gamma\delta e^{\varepsilon' - \varepsilon}$  if  $K$  is  $\gamma$ -ultra-mixing.

Mixing Condition	Local DP Condition
$\gamma$ -Dobrushin	$(0, \gamma)$ -LDP
$(\gamma, \varepsilon)$ -Dobrushin	$(\varepsilon, \gamma)$ -LDP
$\gamma$ -Doeblin	Blanket condition <sup>4</sup>
$\gamma$ -ultra-mixing	$(\log \frac{1}{1-\gamma}, 0)$ -LDP

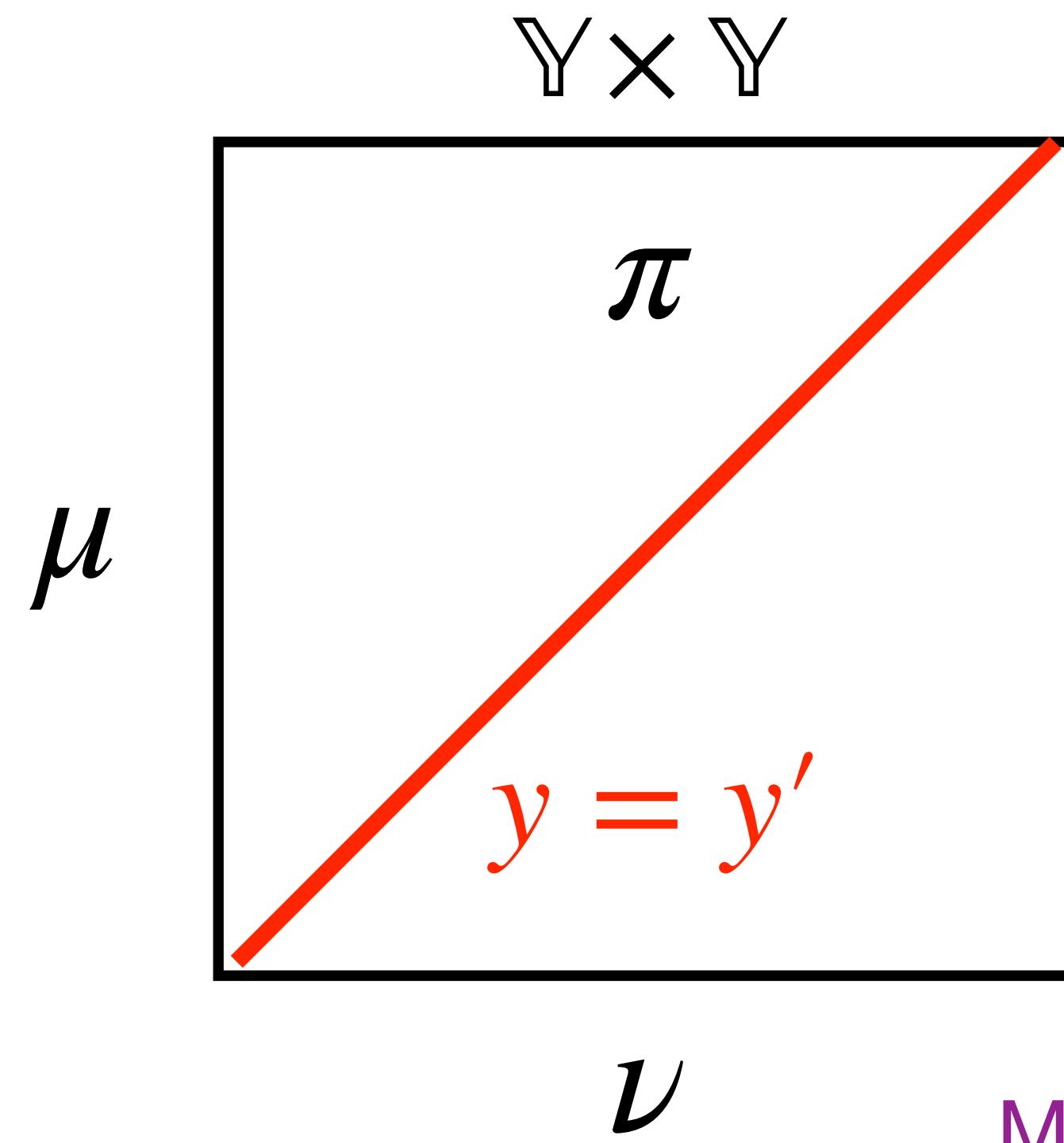
[BBGG'19]

# Overlapping Mixtures as Couplings

$$TV(\mu, \nu) = \inf_{\pi \in C(\mu, \nu)} \Pr_{(Y, Y') \sim \pi} [Y \neq Y'] = 1 - \int \min\{p_\mu(y), p_\nu(y)\} dy = \gamma$$

# Overlapping Mixtures as Couplings

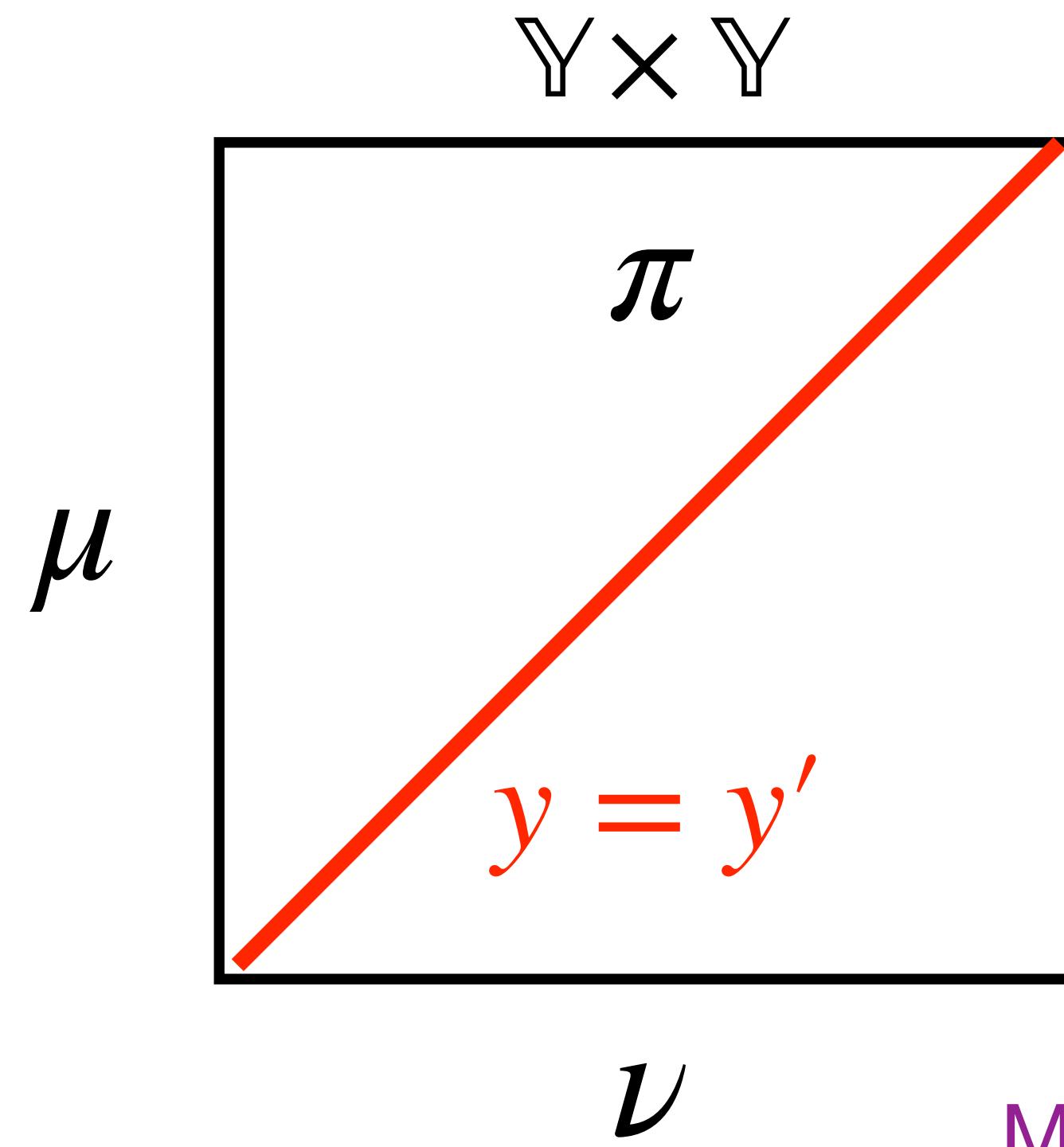
$$TV(\mu, \nu) = \inf_{\pi \in C(\mu, \nu)} \Pr_{(Y, Y') \sim \pi} [Y \neq Y'] = 1 - \int \min\{p_\mu(y), p_\nu(y)\} dy = \gamma$$



Maximal coupling: put as much mass as possible on the diagonal!

# Overlapping Mixtures as Couplings

$$TV(\mu, \nu) = \inf_{\pi \in C(\mu, \nu)} \Pr_{(Y, Y') \sim \pi} [Y \neq Y'] = 1 - \int \min\{p_\mu(y), p_\nu(y)\} dy = \gamma$$



$$\omega = \pi|_{y=y'}$$

$$\mu = (1 - \gamma)\omega + \gamma\mu_2$$

$$\nu = (1 - \gamma)\omega + \gamma\nu_2$$

Maximal coupling: put as much mass as possible on the diagonal!

# Couplings Beyond Overlapping Mixtures

# Other Coupling Strategies

- The maximal coupling leads to two components:
  - One component does not depend on the “data that changed”
  - The remaining components have disjoint support:  $\mu_2 \perp v_2$
- More than two components might be necessary
  - Amplification by shuffling: as many components as individuals
  - In some applications disjoint support is not the right condition
    - Amplification by iteration: keep all mass close to the diagonal

# The Shuffle Model

$$\mathcal{R} : \mathbb{X} \rightarrow \mathbb{Y}$$

$$x_1 \rightarrow \mathcal{R} \rightarrow y_1$$

$$x_2 \rightarrow \mathcal{R} \rightarrow y_2$$

⋮  
⋮  
⋮

$$x_n \rightarrow \mathcal{R} \rightarrow y_n$$

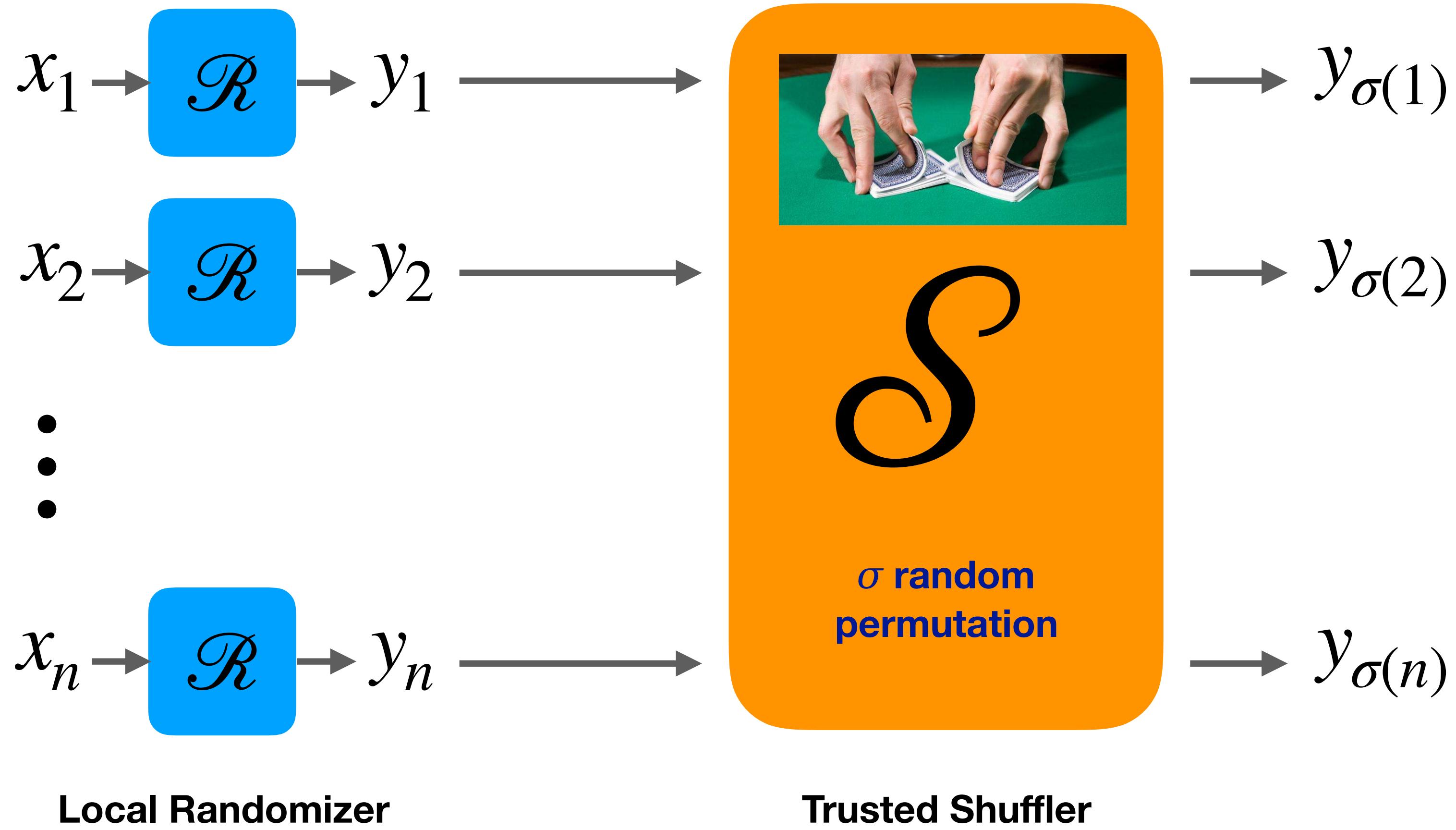
**Local Randomizer**

[BEMMRLRKTS17; EFMRTT19; CSUZZ19]

# The Shuffle Model

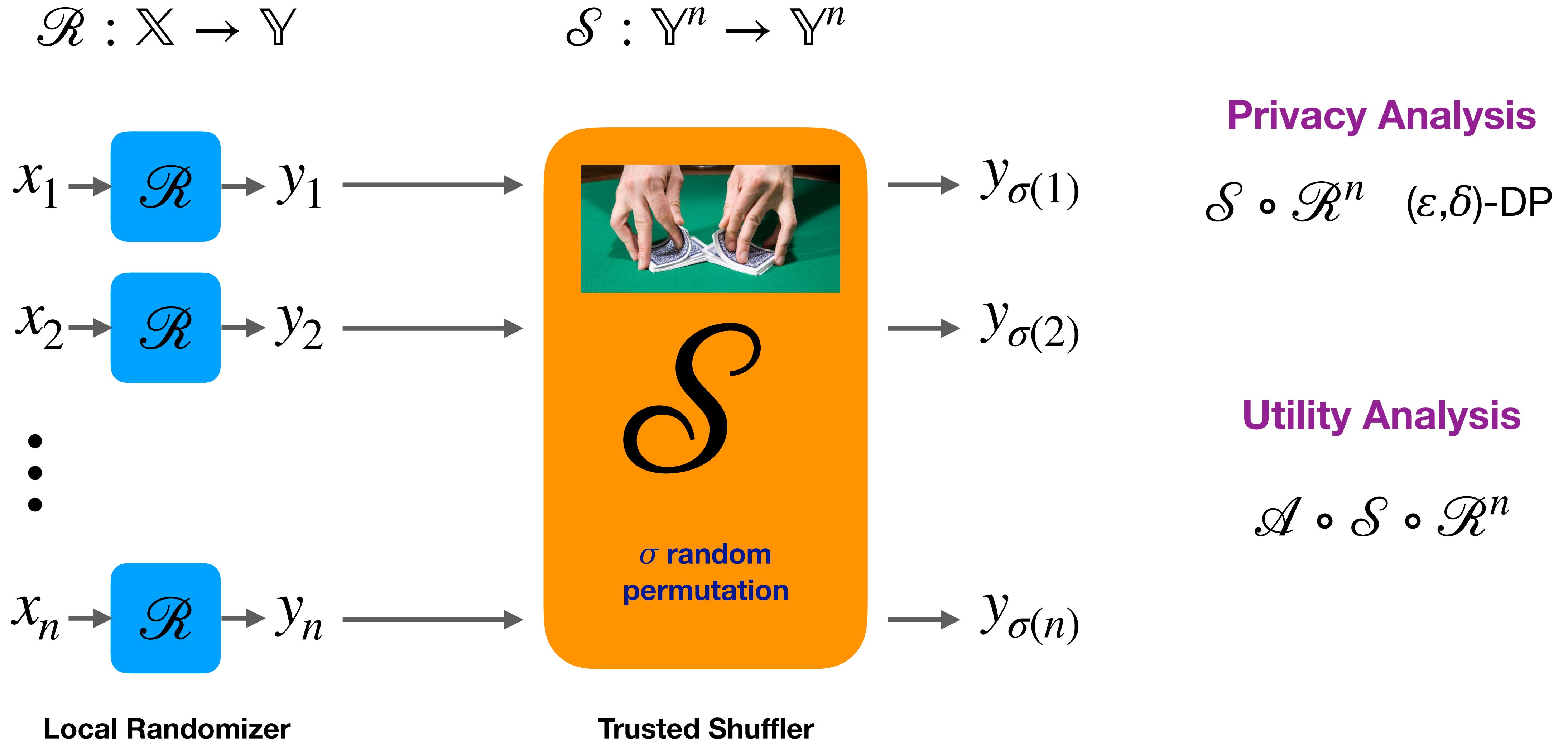
$$\mathcal{R} : \mathbb{X} \rightarrow \mathbb{Y}$$

$$\mathcal{S} : \mathbb{Y}^n \rightarrow \mathbb{Y}^n$$



[BEMMRLRKTS17; EFMRTT19; CSUZZ19]

# The Shuffle Model

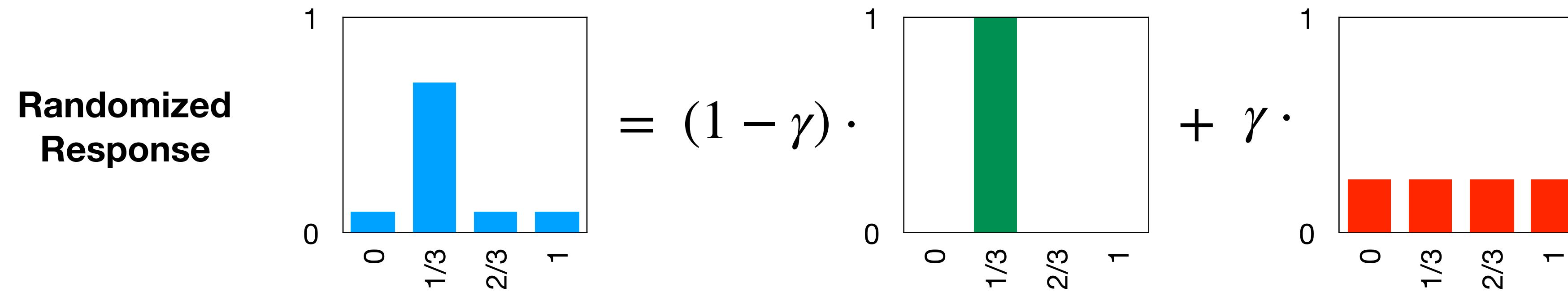


[BEMMRLRKTS17; EFMRTT19; CSUZZ19]

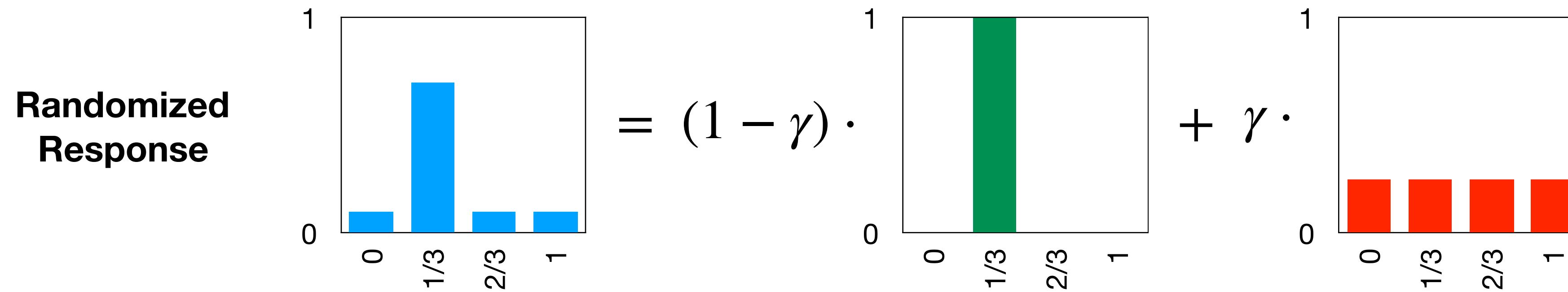
# Privacy Amplification by Shuffling

- **Problem Statement**
  - Characterize the privacy of shuffled mechanisms in terms of the privacy of its local randomizers
- **Previous Work [EFMRTT, SODA 2019]**
  - Shuffle-then-randomize (with adaptativity):
$$\varepsilon = O\left(\varepsilon_0 \sqrt{\log(1/\delta)/n}\right)$$
for  $\varepsilon_0 = O(1)$
- **Our Result**
  - Randomize-then-shuffle (one randomizer):
$$\varepsilon = O\left((\varepsilon_0 \wedge 1)e^{\varepsilon_0} \sqrt{\log(1/\delta)/n}\right)$$
for  $\varepsilon_0 \leq 0.5 \log(n) + O(1)$

# Blanket of a Local Randomizer

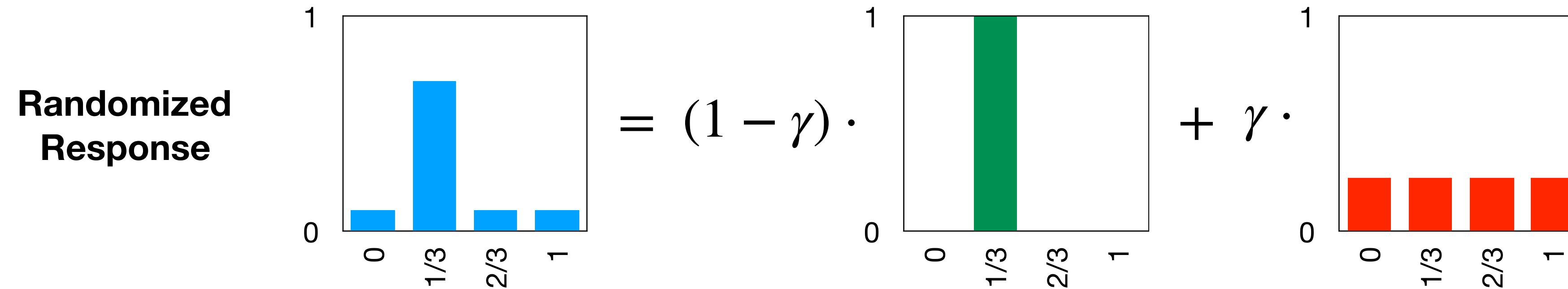


# Blanket of a Local Randomizer



- **Theorem (Blanket Decomposition):** Every  $\epsilon_0$ -LDP randomizer admits a (unique maximal) mixture decomposition where one of the components is independent of the input

# Blanket of a Local Randomizer

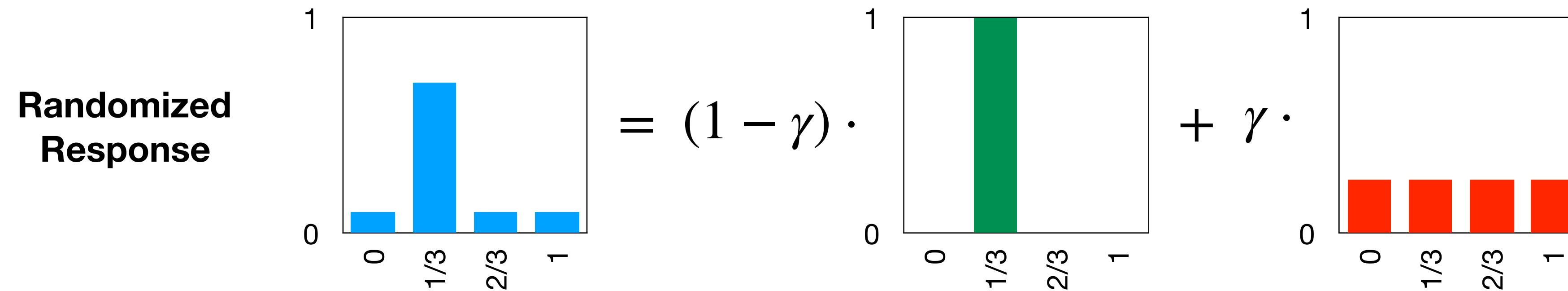


- **Theorem (Blanket Decomposition):** Every  $\varepsilon_0$ -LDP randomizer admits a (unique maximal) mixture decomposition where one of the components is independent of the input

$$\mathcal{R}(x) = (1 - \gamma)\mathcal{R}'(x) + \gamma\omega$$

$$\begin{aligned}\mathcal{R} : \mathbb{X} &\rightarrow \mathbb{Y} \\ \mathcal{R}' : \mathbb{X} &\rightarrow \mathbb{Y} \quad e^{-\varepsilon_0} \leq \gamma \leq 1 \\ \omega &\in \text{Dist}(\mathbb{Y})\end{aligned}$$

# Blanket of a Local Randomizer



- **Theorem (Blanket Decomposition):** Every  $\epsilon_0$ -LDP randomizer admits a (unique maximal) mixture decomposition where one of the components is independent of the input

$$\mathcal{R}(x) = (1 - \gamma)\mathcal{R}'(x) + \gamma\omega$$

$$\mathcal{R} : \mathbb{X} \rightarrow \mathbb{Y}$$

$$\mathcal{R}' : \mathbb{X} \rightarrow \mathbb{Y}$$

$$\omega \in \text{Dist}(\mathbb{Y})$$

$$e^{-\epsilon_0} \leq \gamma \leq 1$$

## Blanket Construction

$$\gamma = \int_{\mathbb{Y}} \min_{x \in \mathbb{X}} p_{\mathcal{R}(x)}(y) dy$$

$$p_{\omega}(y) = \frac{\min_{x \in \mathbb{X}} p_{\mathcal{R}(x)}(y)}{\gamma}$$



# Example Blanket Decompositions

$\varepsilon_0$ -LDP RR on  $[k]$

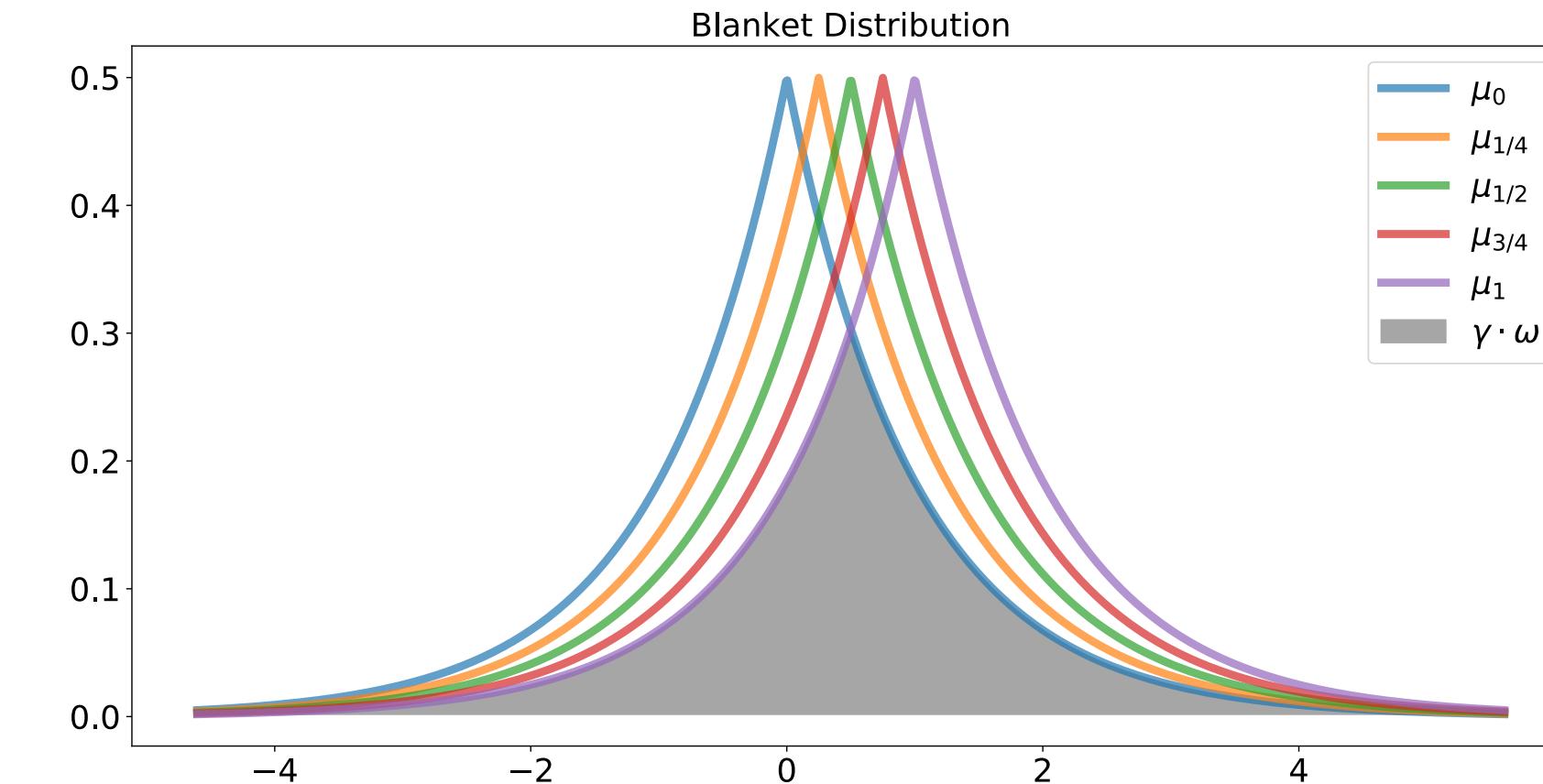
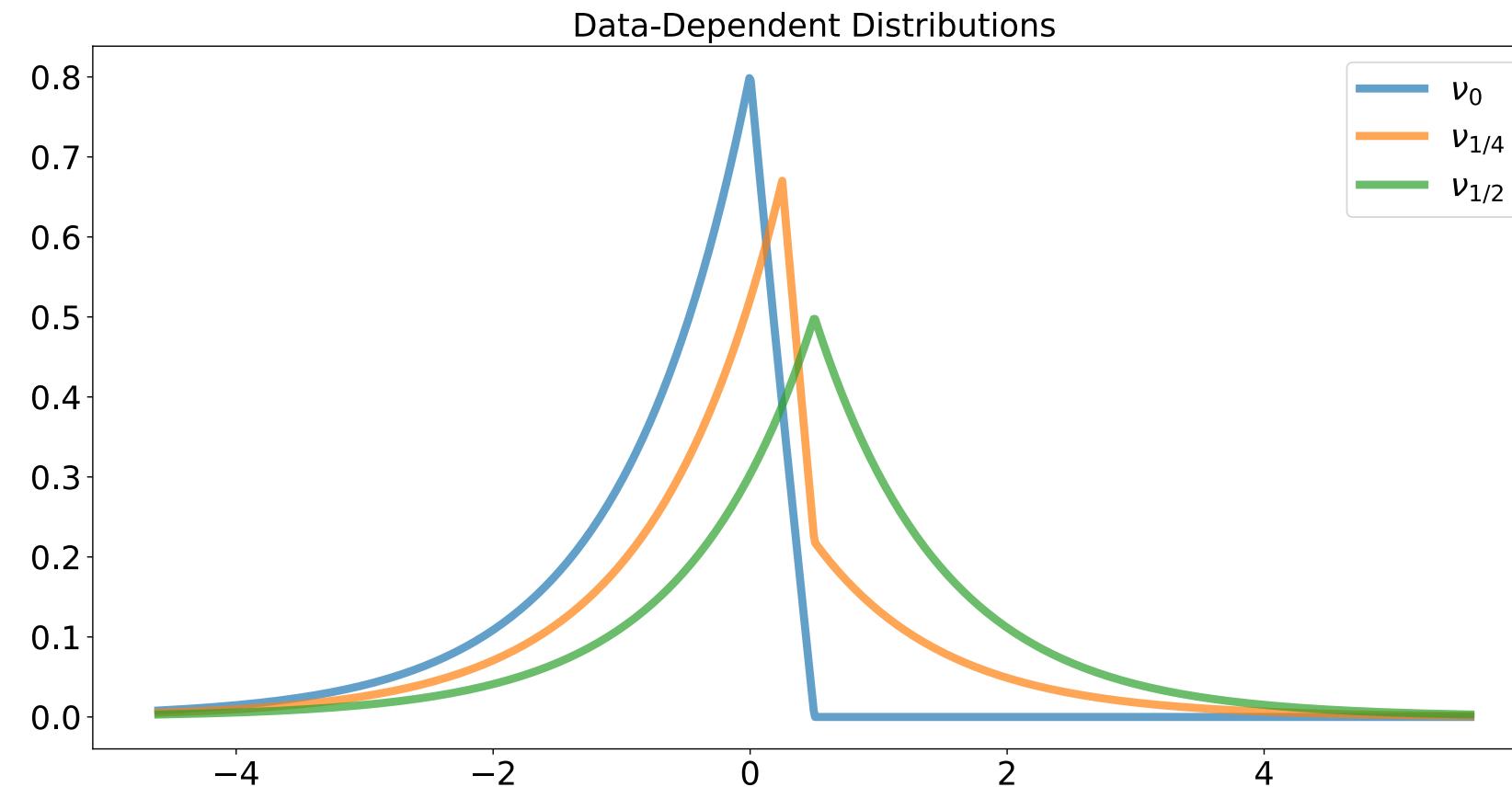
$$\gamma = \frac{k}{e^{\varepsilon_0} + k - 1}$$

$\varepsilon_0$ -LDP Laplace on  $[0,1]$

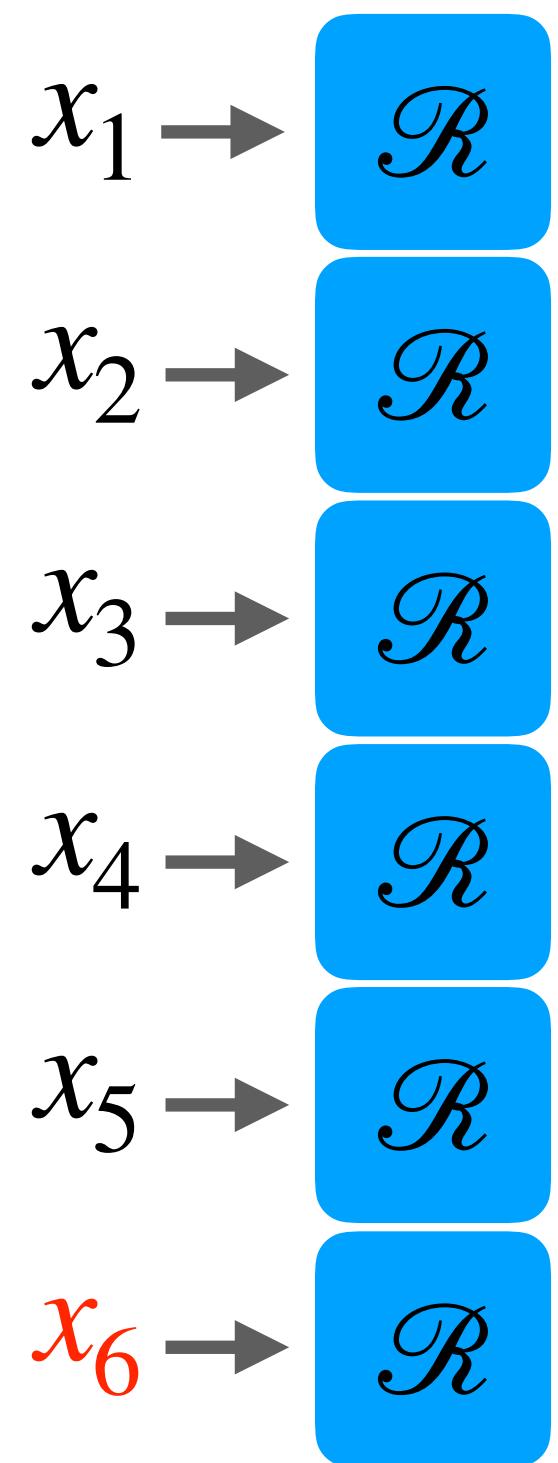
$$\gamma = e^{-\frac{\varepsilon_0}{2}}$$

$\sigma^2$  Gaussian on  $[0,1]$

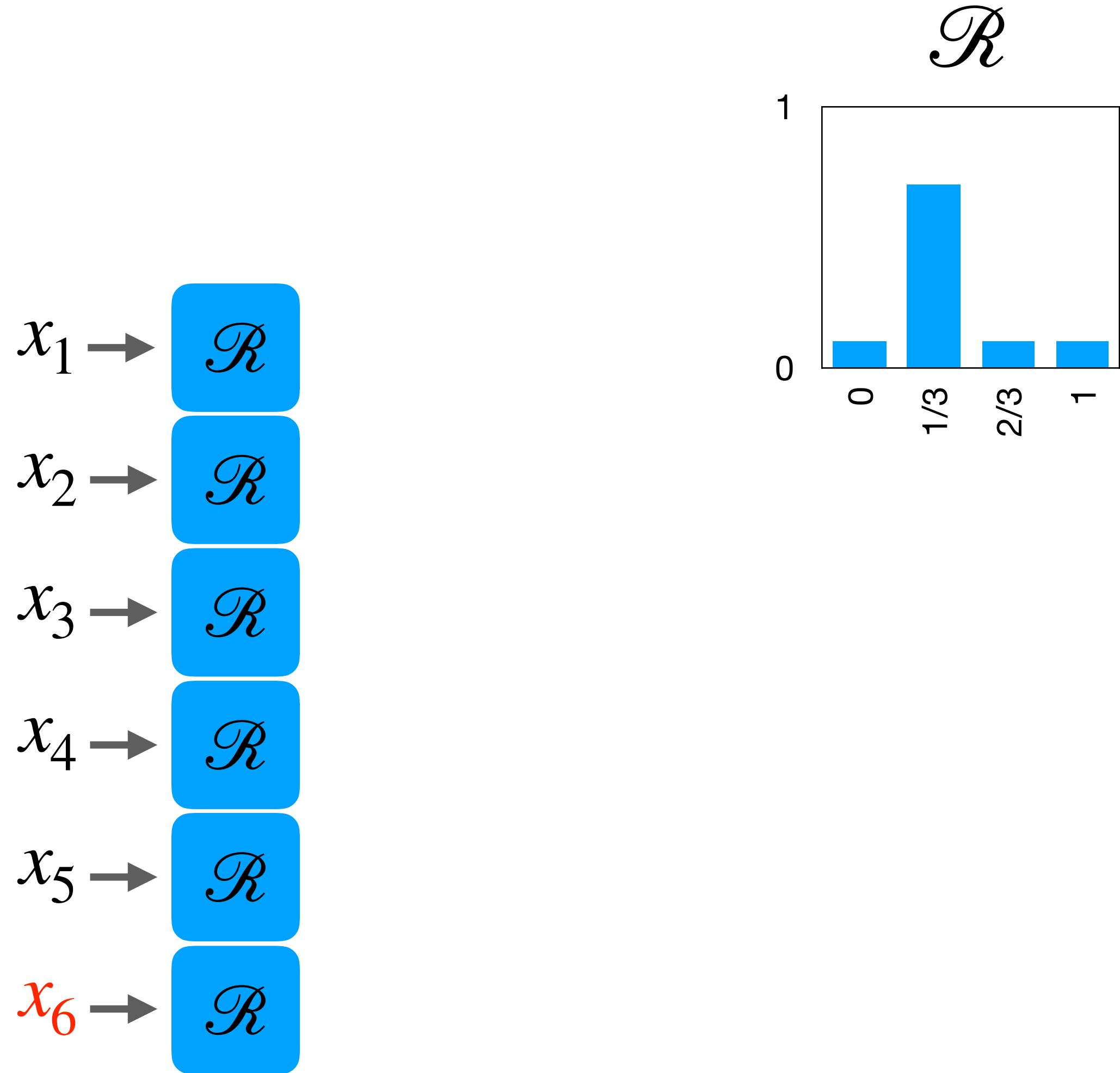
$$\gamma = 2\mathbb{P}[N(0,\sigma^2) \leq -1/2]$$



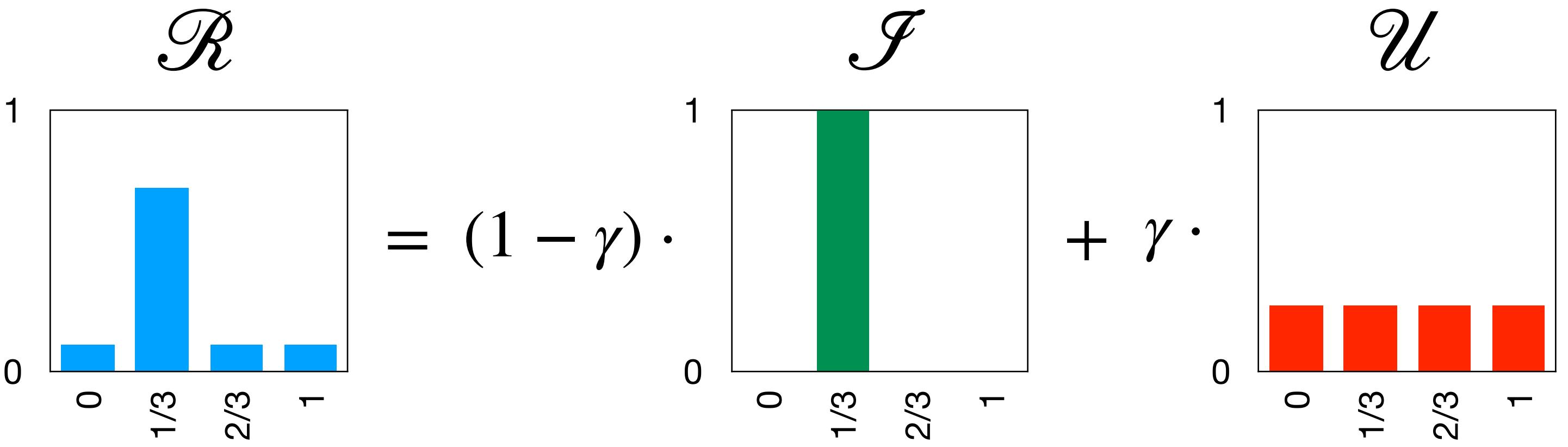
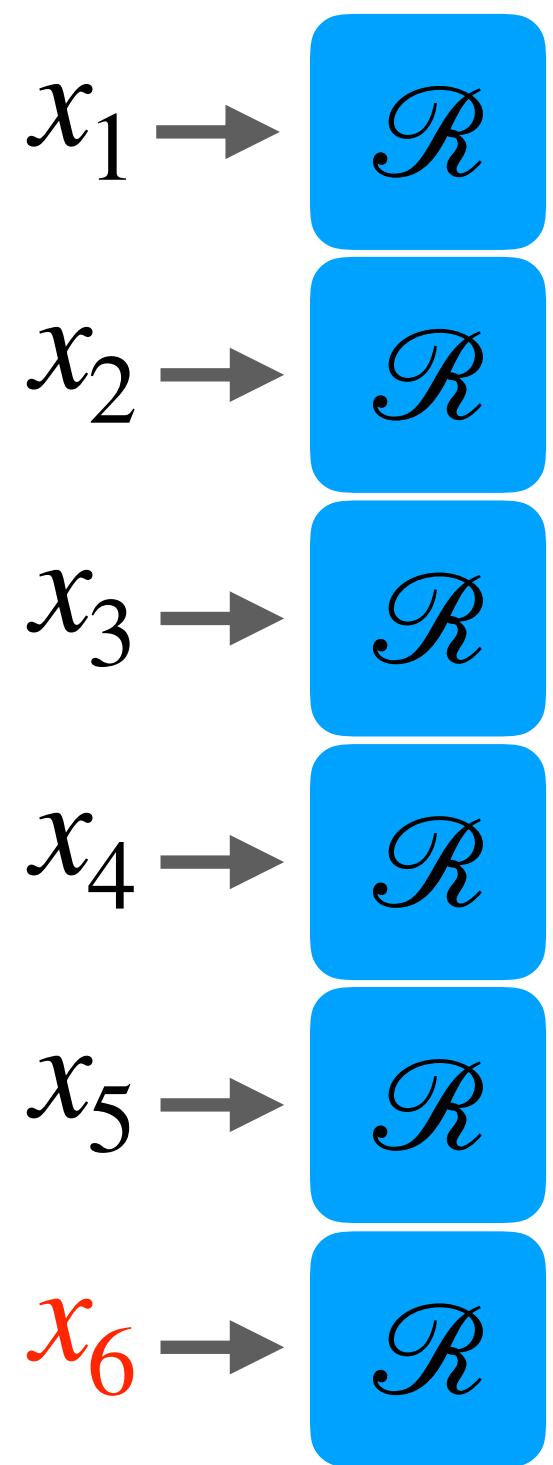
# Coupling the Blankets



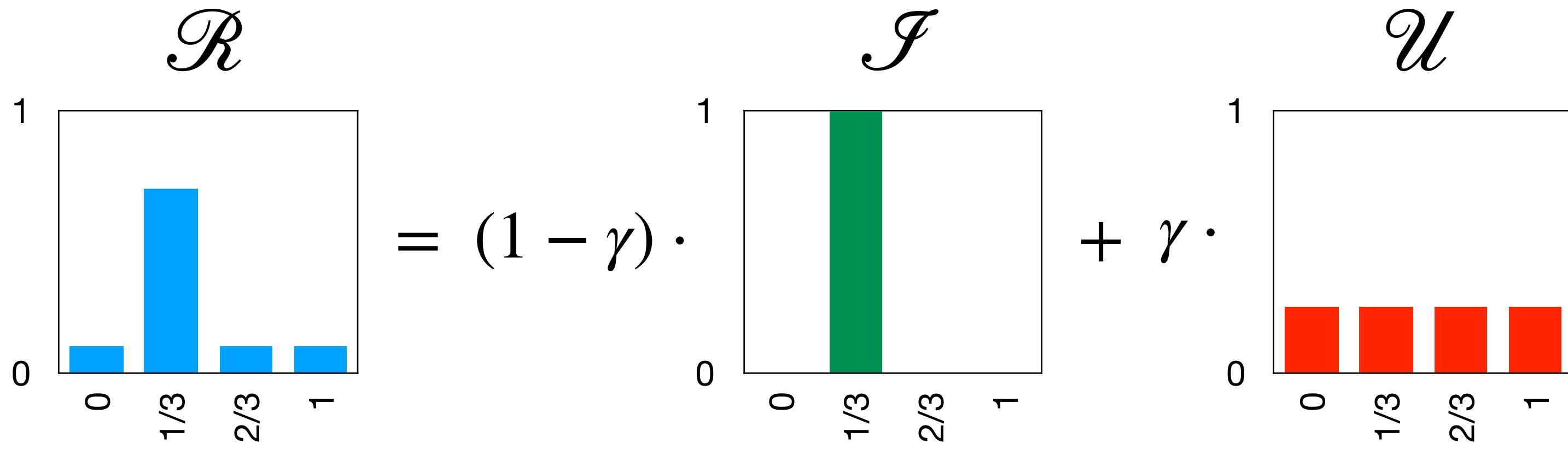
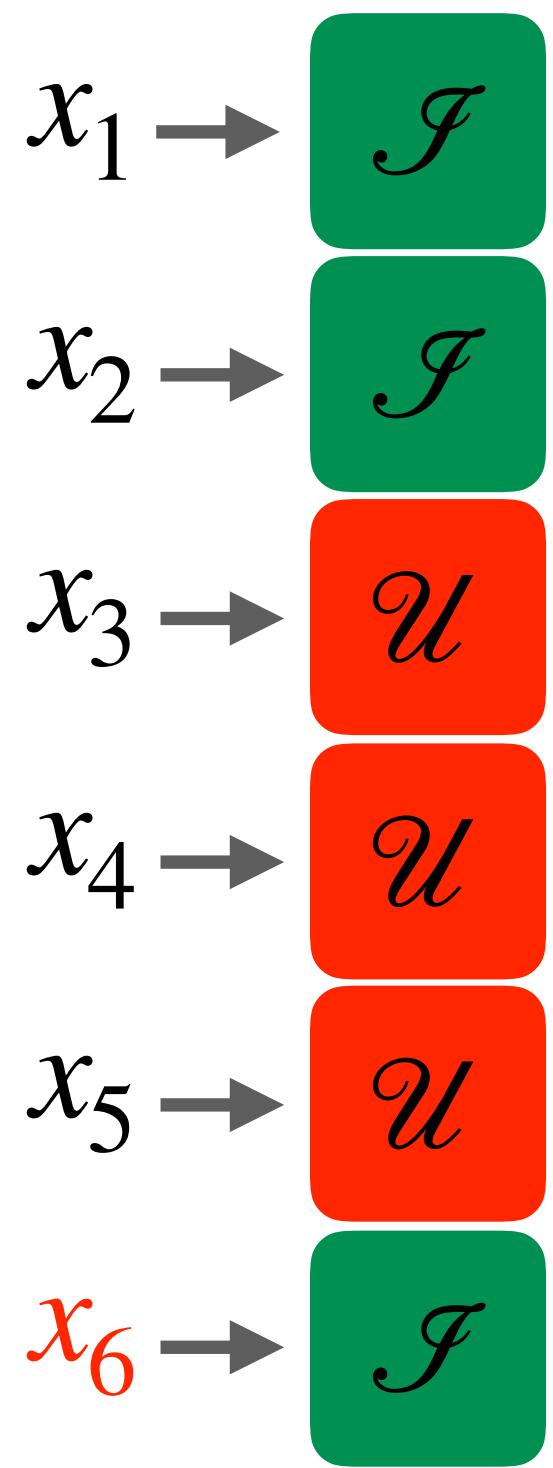
# Coupling the Blankets



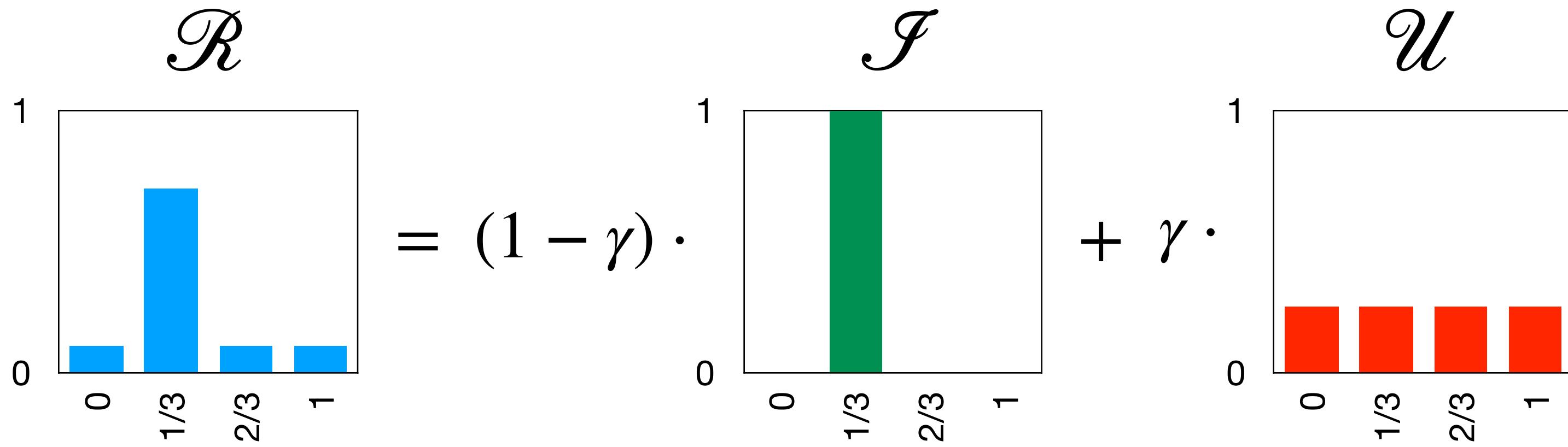
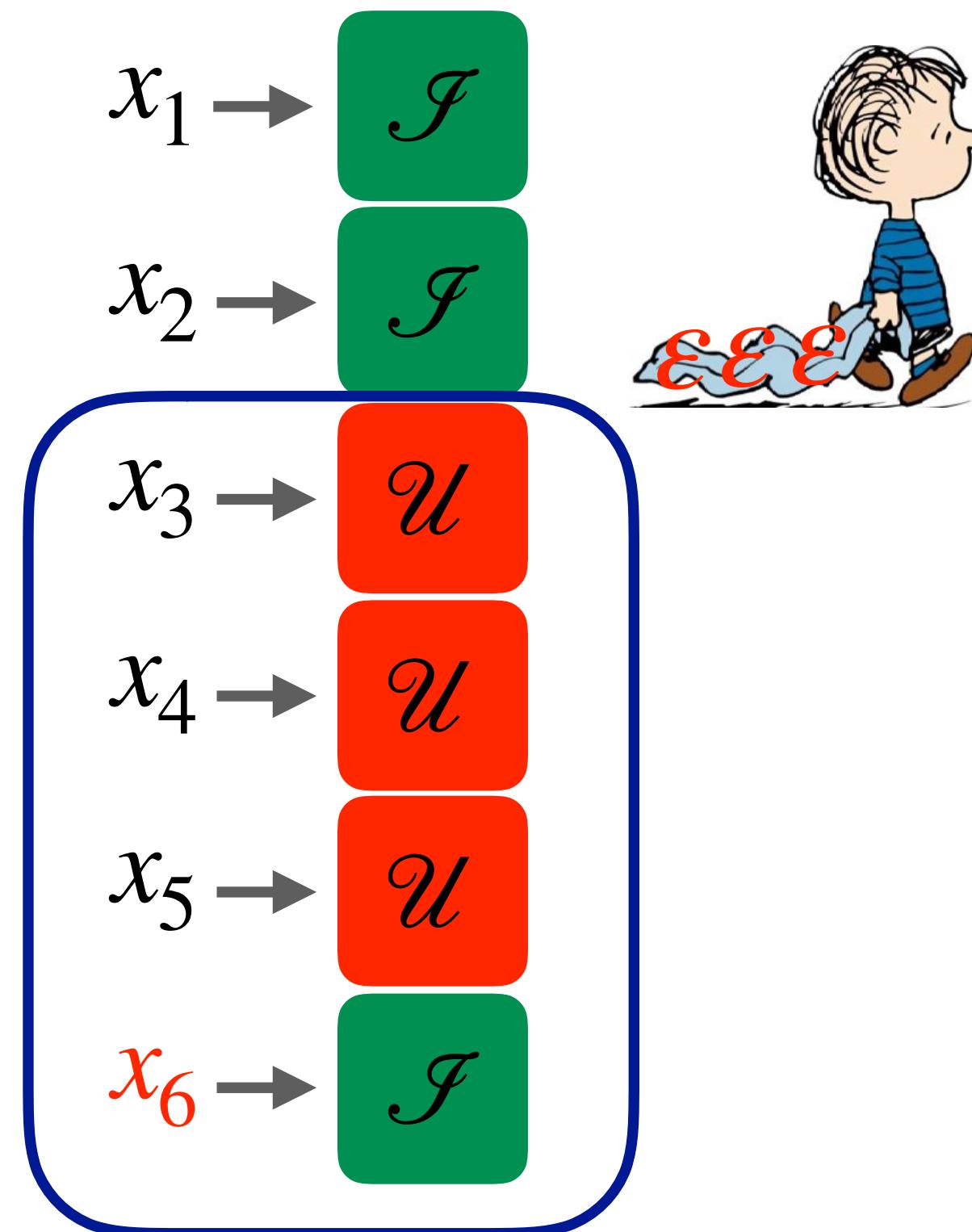
# Coupling the Blankets



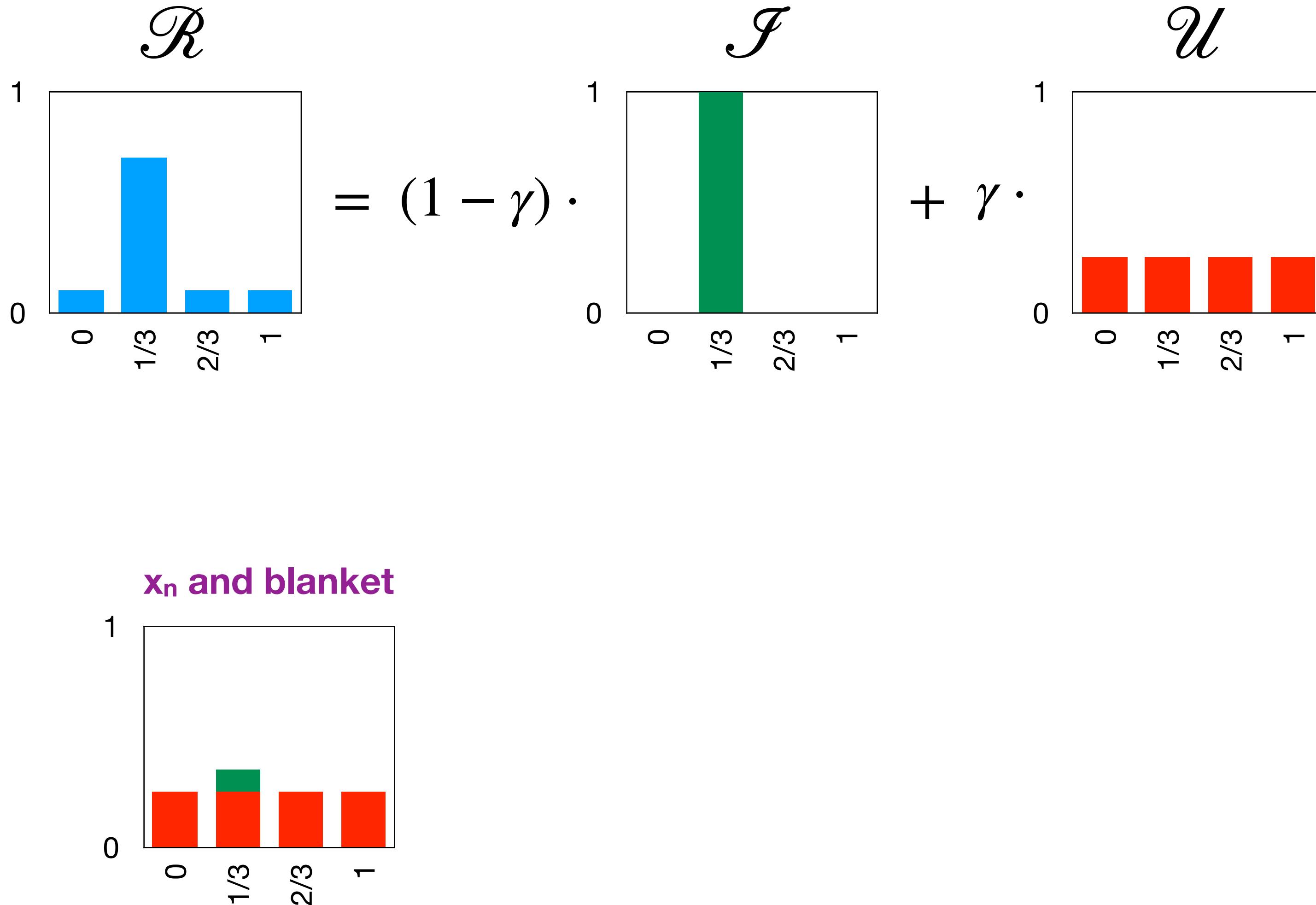
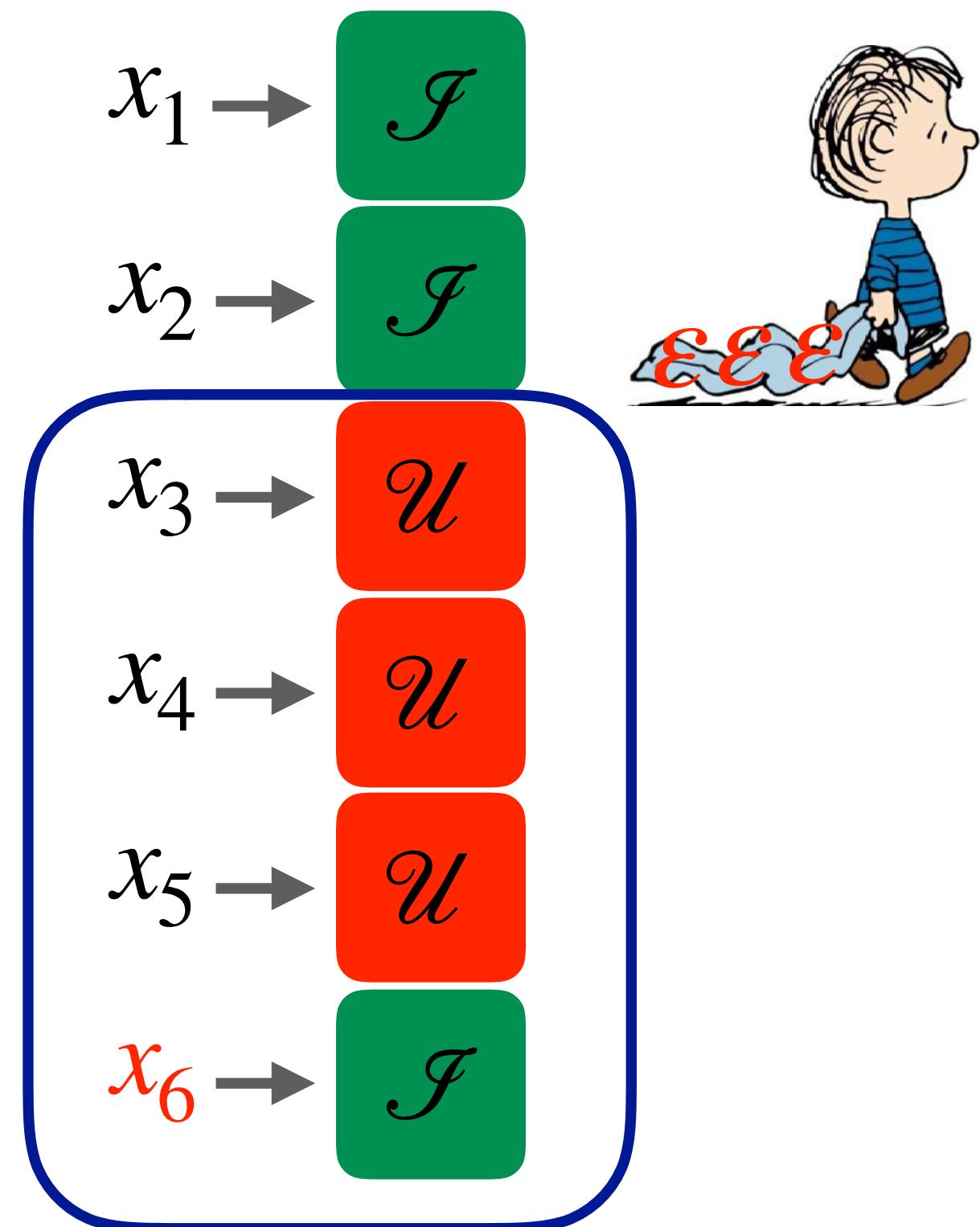
# Coupling the Blankets



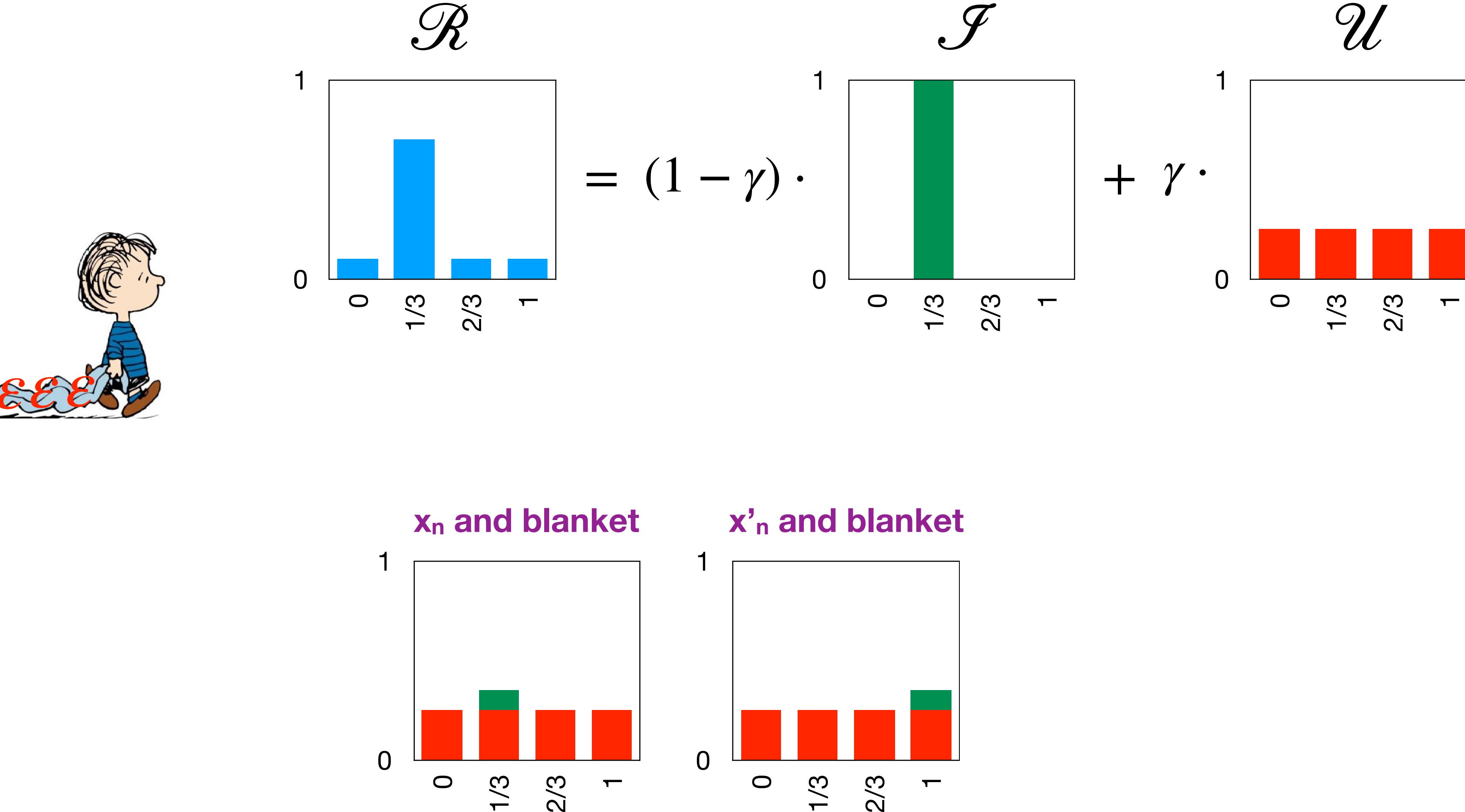
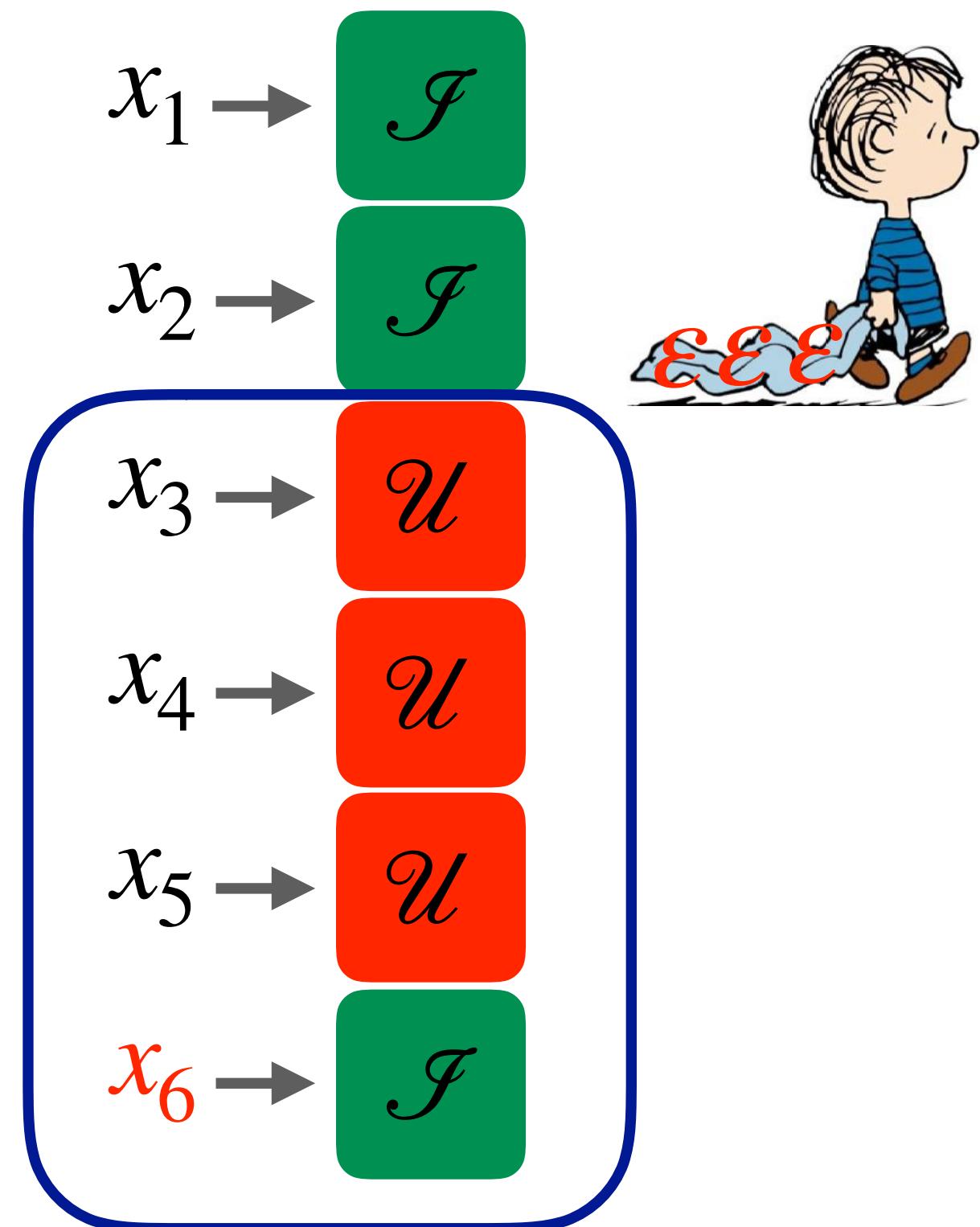
# Coupling the Blankets



# Coupling the Blankets



# Coupling the Blankets



# Amplification by Shuffling

- **Theorem:** Shuffling  $n$  copies of any  $\varepsilon_0$ -LDP randomizer with **blanket parameter  $\gamma$**  gives  $(\varepsilon, \delta)$ -DP with

$$\frac{\gamma(e^\varepsilon + 1)^2(e^{\varepsilon_0} - e^{-\varepsilon_0})^2}{4n(e^\varepsilon - 1)} \cdot \exp\left(-0.86n\left(\gamma \wedge \frac{(e^\varepsilon - 1)^2}{\gamma(e^\varepsilon + 1)^2(e^{\varepsilon_0} - e^{-\varepsilon_0})^2}\right)\right) \leq \delta$$

# Amplification by Shuffling

- **Theorem:** Shuffling  $n$  copies of any  $\varepsilon_0$ -LDP randomizer with **blanket parameter  $\gamma$**  gives  $(\varepsilon, \delta)$ -DP with

$$\frac{\gamma(e^\varepsilon + 1)^2(e^{\varepsilon_0} - e^{-\varepsilon_0})^2}{4n(e^\varepsilon - 1)} \cdot \exp\left(-0.86n\left(\gamma \wedge \frac{(e^\varepsilon - 1)^2}{\gamma(e^\varepsilon + 1)^2(e^{\varepsilon_0} - e^{-\varepsilon_0})^2}\right)\right) \leq \delta$$

- **Corollary:** Shuffling  $n$  copies of an  $\varepsilon_0$ -LDP randomizer gives  $(\varepsilon, \delta)$ -DP with

$$\varepsilon = O\left((\varepsilon_0 \wedge 1)e^{\varepsilon_0}\sqrt{\log(1/\delta)/n}\right) \quad \varepsilon_0 \leq \log(n/\log(1/\delta))/2$$

# Amplification: Proof Idea

- General idea
  - **Couple** who samples from the blanket in both executions
  - **Reveal** the identity of who samples from the blanket (**joint convexity**)
  - **Remove** the data from the users in  $1 \dots n-1$  who sampled from  $R'$  (**post-processing**)

- Define privacy amplification random variable  $\mathbb{E}[L] = 1 - e^\varepsilon < 0$

$$Y \sim \omega \quad L = L_{x,x'}^{\mathcal{R}} = \frac{p_{\mathcal{R}(x)}(Y) - e^\varepsilon p_{\mathcal{R}(x')}(Y)}{p_\omega(Y)} \quad \gamma(e^{-\varepsilon_0} - e^{\varepsilon+\varepsilon_0}) \leq L \leq \gamma(e^{\varepsilon_0} - e^{\varepsilon-\varepsilon_0})$$

- Reduce to bounding expectation, apply concentration for bounded r.v.'s

$$\sup_E (\mathbb{P}[\mathcal{S} \circ \mathcal{R}^n(\vec{x}) \in E] - e^\varepsilon \mathbb{P}[\mathcal{S} \circ \mathcal{R}^n(\vec{x}') \in E]) \leq \frac{1}{\gamma n} \mathbb{E} \left[ \sum_{i=1}^{Bin(n,\gamma)} L_i \right]_+$$

# Getting the Bound

- Applying Hoeffding's inequality we get

$$\frac{1}{\gamma n} \mathbb{E} \left[ \sum_{i=1}^{Bin(n,\gamma)} L_i \right]_+ \leq \frac{\gamma(e^\varepsilon + 1)^2(e^{\varepsilon_0} - e^{-\varepsilon_0})^2}{4n(e^\varepsilon - 1)} \cdot \exp \left( -0.86n \left( \gamma \wedge \frac{(e^\varepsilon - 1)^2}{\gamma(e^\varepsilon + 1)^2(e^{\varepsilon_0} - e^{-\varepsilon_0})^2} \right) \right)$$

- Refinements:
  - Use mechanism-specific bounds on  $L$  and  $\gamma$
  - Alternative concentration bounds, eg. Bennett's inequality

# Amplification by Iteration in NoisySGD

---

**Algorithm 1:** Noisy Projected Stochastic Gradient Descent — NoisyProjSGD( $D, \ell, \eta, \sigma, \xi_0$ )

---

**Input:** Dataset  $D = (z_1, \dots, z_n)$ , loss function  $\ell : \mathbb{K} \times \mathbb{D} \rightarrow \mathbb{R}$ , learning rate  $\eta$ , noise parameter  $\sigma$ , initial distribution  $\xi_0 \in \mathcal{P}(\mathbb{K})$

Sample  $x_0 \sim \xi_0$

**for**  $i \in [n]$  **do**

$x_i \leftarrow \Pi_{\mathbb{K}}(x_{i-1} - \eta(\nabla_x \ell(x_{i-1}, z_i) + Z))$  with  $Z \sim \mathcal{N}(0, \sigma^2 I)$

**return**  $x_n$

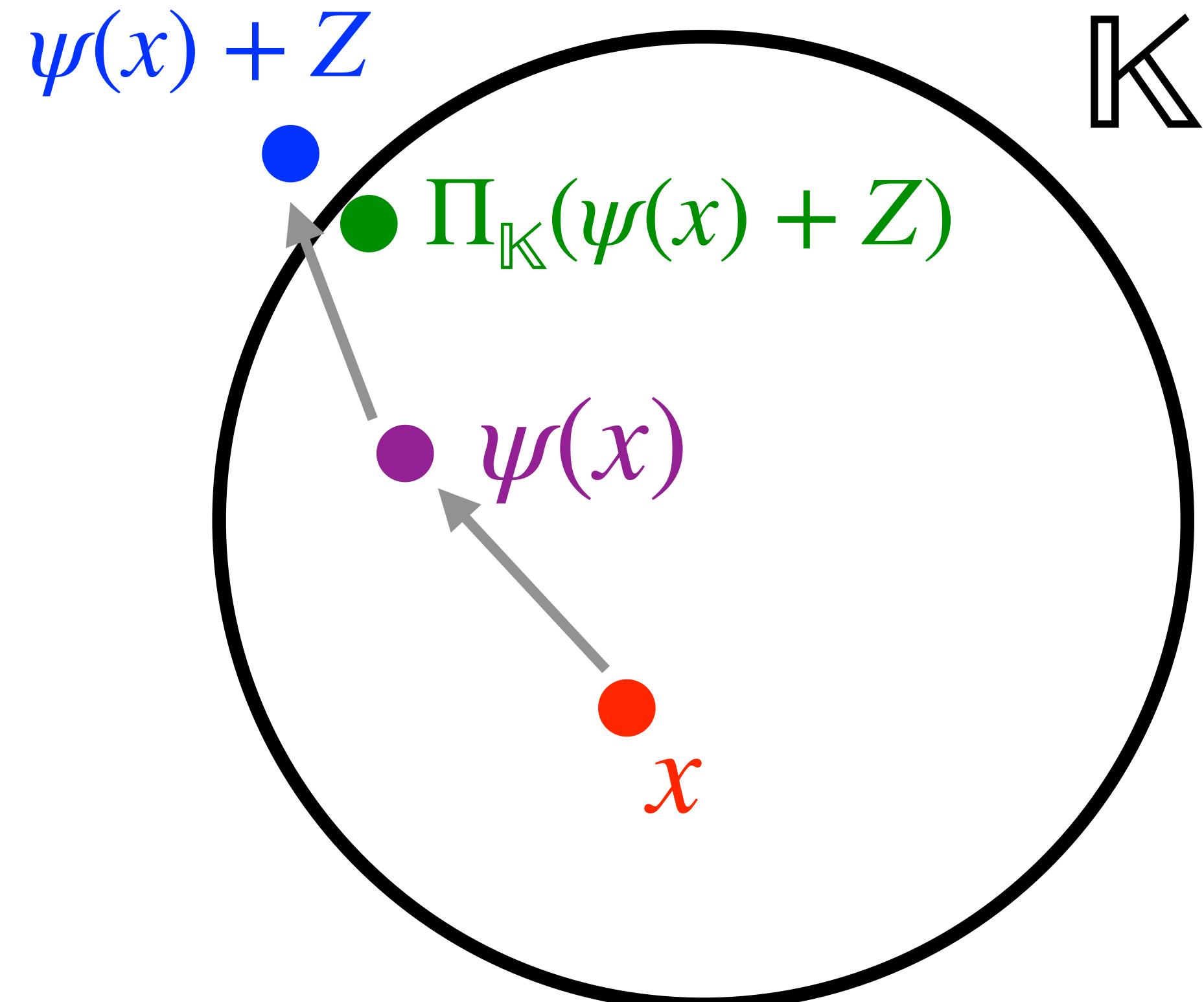
---

- If  $D$  and  $D'$  differ in position  $j$ , then the last  $n-j$  iterations are postprocessing
  - Can also use public data for the last  $r$  iterations
  - Start from a coupling between  $x_j$  and  $x'_j$  and propagate it through
    - Keep all the mass as close to the diagonal as possible

# Projected Generalized Gaussian Mechanism

$$K(x) = \Pi_{\mathbb{K}}(\mathcal{N}(\psi(x), \sigma^2 I))$$

$$\psi : \mathbb{R}^d \rightarrow \mathbb{R}^d$$



# Amplification by Coupling

Suppose  $\psi_1, \dots, \psi_r$  are  $L$ -Lipschitz

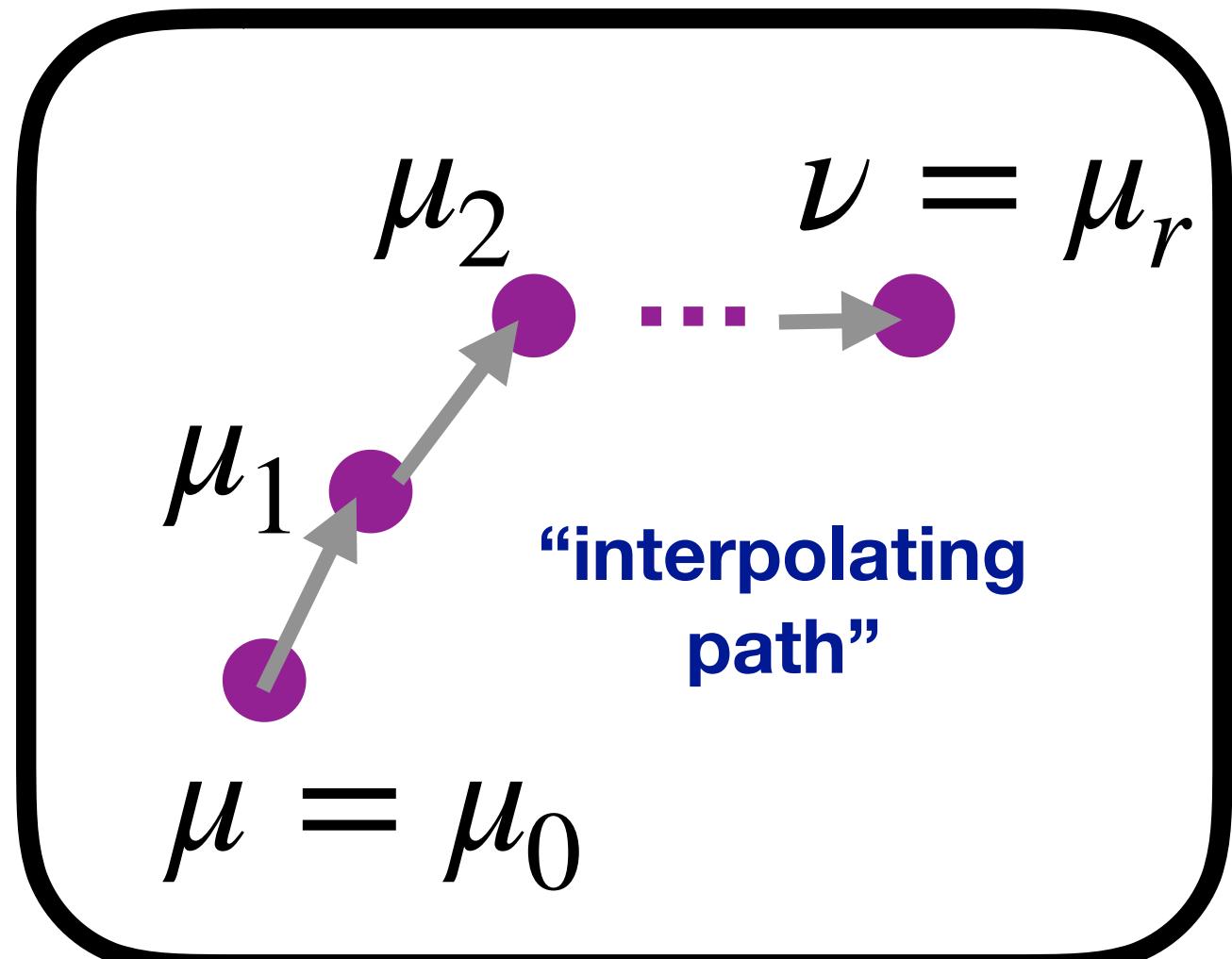
$$K_i(x) = \Pi_{\mathbb{K}}(\mathcal{N}(\psi_i(x), \sigma^2 I))$$

$$\text{R}_\alpha(\mu K_1 \cdots K_r \| \nu K_1 \cdots K_r) \leq \frac{\alpha L^2}{2\sigma^2} \sum_{i=1}^r L^{2(r-i)} W_\infty(\mu_i, \mu_{i-1})^2$$

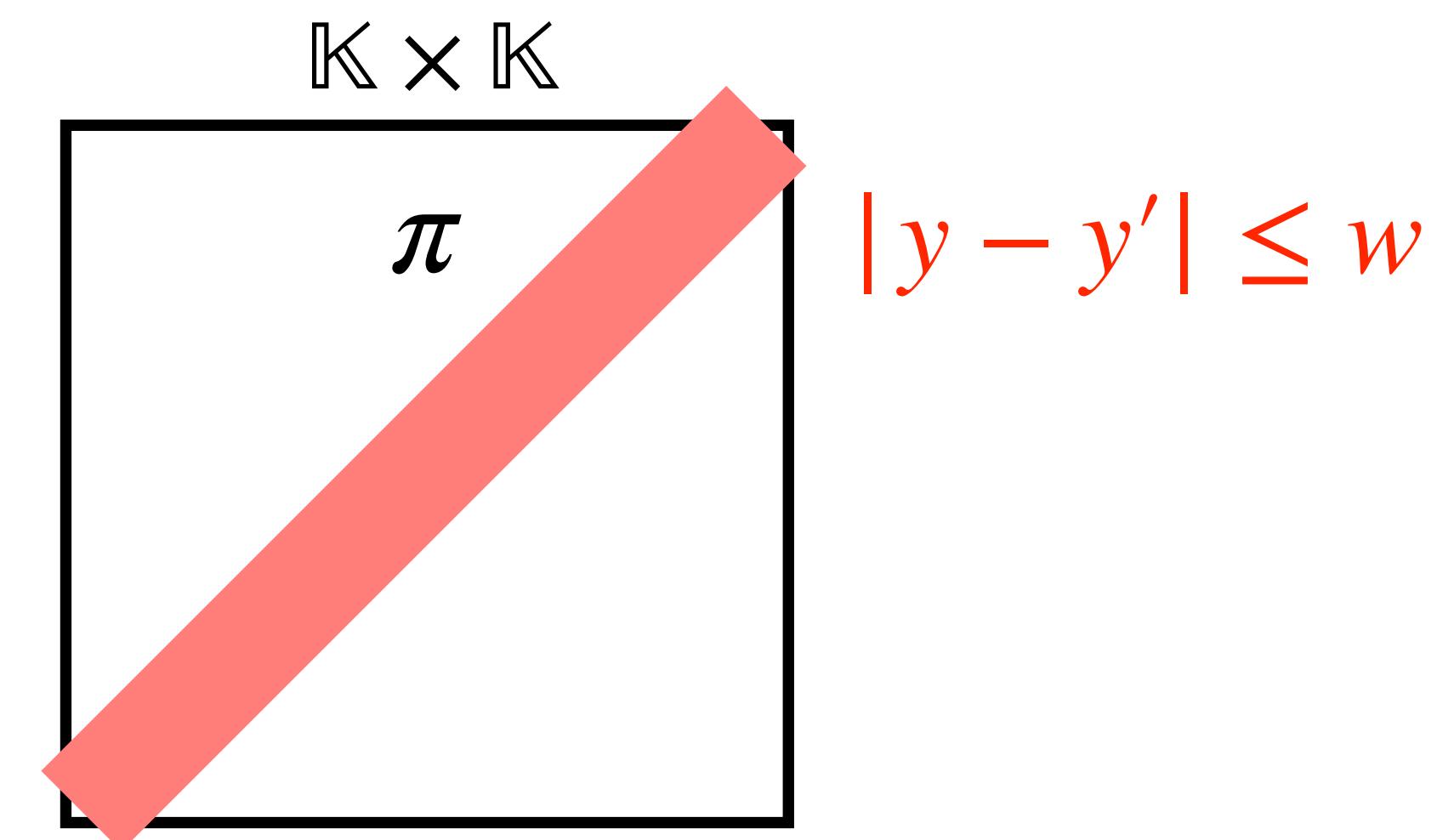
Rényi Divergence

Wasserstein Distance

$\mathcal{P}(\mathbb{K})$



- Applications:
- Bound  $L$
  - Optimize path



# Per-index RDP in NoisySGD

*Suppose the loss is Lipschitz and smooth*

*If loss is **convex** can take  $L=1$ . Then  $i$ -th person receives  $\epsilon_i(\alpha)$ -RDP with*

$$\epsilon_i(\alpha) = O\left(\frac{\alpha}{(n - i)\sigma^2}\right)$$

[FMTT'18]

*If loss is **strongly convex** can take  $L < 1$ . Then  $i$ -th person receives  $\epsilon_i(\alpha)$ -RDP with*

$$\epsilon_i(\alpha) = O\left(\frac{\alpha L^{(n-i)/2}}{(n - i)\sigma^2}\right)$$

[BBGG'19]

# Summary

- Couplings (including overlapping mixtures) provide a powerful methodology to study privacy amplification in many settings
  - Including: subsampling, postprocessing, shuffling and iteration
- Properties of divergences related to (R)DP (eg. advanced joint convexity) are “necessary” to get tight amplification bounds
- Different types of couplings are useful (eg. maximal and small distance)

# Conclusion

- **In DP constants (and exact expressions) matter**
  - Better noise calibration => better utility
  - See also the “analytic Gaussian mechanism” *[BW’18]*
- **Make every bit of randomness count!**
  - Does your postprocessing improve privacy?
  - **Other types of privacy amplification?**
    - Eg. is Report Noisy Max an instance of privacy amplification?

