

Exercise 1. Consider the following discrete time deterministic dynamical system

$$z^{n+1} = \nu z^n (1 - z^n) \quad (1)$$

with $\nu = 4$ and $z^n \in (0, 1) \forall n \in \mathbb{N}$. Its invariant density is given by

$$\frac{1}{\pi \sqrt{z(1-z)}}. \quad (2)$$

a) Show analytically that (2) is invariant under the mapping (1).

Hint: obtain an integral equation for the evolution of the density π^n by starting from $\mathbb{P}(z^{n+1} \leq z)$.

We want to obtain the marginal density π^{n+1} and show that it is invariant, i.e. $\pi^{n+1} = \pi^n$.

$$F_{z^{n+1}}(z) = \mathbb{P}(4z^n(1-z^n) < z)$$

$$\text{CDF of } z^{n+1} = \mathbb{P}(z^n - (z^n)^2 < \frac{z}{4})$$

$$= \mathbb{P}((z^n)^2 - z^n + \frac{z}{4} > 0)$$

$$= \mathbb{P}\left(z^n < \frac{1-\sqrt{1-z}}{2}\right) + \mathbb{P}\left(z^n > \frac{1+\sqrt{1-z}}{2}\right)$$

$$= \mathbb{P}\left(z^n < \frac{1-\sqrt{1-z}}{2}\right) + \left(1 - \mathbb{P}\left(z^n < \frac{1+\sqrt{1-z}}{2}\right)\right)$$

$$= \left[\frac{2}{\pi} \arcsin\left(\sqrt{\frac{1-\sqrt{1-z}}{2}}\right) + 1 - \frac{2}{\pi} \arcsin\left(\sqrt{\frac{1+\sqrt{1-z}}{2}}\right) \right]$$

μ_1, μ_2 and you want to show

$\mu_1 = \mu_2$
then you do that by showing

$$\mu_1(A) = \mu_2(A)$$

for all subsets A

$$\begin{aligned} \text{change of variables } x &= \sqrt{z} \\ \frac{dx}{dz} &= \frac{1}{2\sqrt{z}} \\ dx &= \frac{1}{2\sqrt{z}} dz \end{aligned}$$

In 1-d it is enough to show that
 $\mu_1((-\infty, x]) = \mu_2((-\infty, x])$

$$F_1(x)$$

CDF

$$F_2(x)$$

CDF

$$\frac{d}{dz} \int_0^{1-\sqrt{1-z}} \dots dz$$

$$\text{Since } \int \frac{1}{\pi \sqrt{z(1-z)}} dz$$

$$= \frac{1}{\pi} \int \frac{1}{\sqrt{1-z} \sqrt{z}} dz$$

$$= \frac{2}{\pi} \int \frac{1}{\sqrt{1-x^2}} dx$$

$$\text{(standard integral)} = \frac{2}{\pi} \cdot \arcsin(x)$$

$$= \frac{2}{\pi} \cdot \arcsin(\sqrt{z})$$

$$\gamma^{n+1}(z) = F_{z^{n+1}}'(z) = (\circledast)'$$

↑
PDF of z^{n+1}

$$\begin{aligned} &= \frac{2}{\pi} \left(\frac{1}{\sqrt{1 - \left(\frac{1}{2} - \frac{\sqrt{1-z}}{2}\right)}} \cdot \frac{1}{2} \cdot \frac{1}{\sqrt{\frac{1}{2} - \frac{\sqrt{1-z}}{2}}} \cdot \left(-\frac{1}{2} \cdot \frac{1}{2\sqrt{1-z}} \cdot (-1)\right) \right. \\ &\quad \left. - \frac{1}{\sqrt{1 - \left(\frac{1}{2} + \frac{\sqrt{1-z}}{2}\right)}} \cdot \frac{1}{2} \cdot \frac{1}{\sqrt{\frac{1}{2} + \frac{\sqrt{1-z}}{2}}} \cdot \left(\frac{1}{2} \cdot \frac{1}{2\sqrt{1-z}} \cdot (-1)\right) \right) \\ &= \frac{1}{4\pi} \left(\frac{1}{\sqrt{\frac{1}{2} + \frac{\sqrt{1-z}}{2}} \cdot \sqrt{\frac{1}{2} - \frac{\sqrt{1-z}}{2}} \cdot \sqrt{1-z}} + \frac{1}{\sqrt{\frac{1}{2} - \frac{\sqrt{1-z}}{2}} \cdot \sqrt{\frac{1}{2} + \frac{\sqrt{1-z}}{2}} \cdot \sqrt{1-z}} \right) \\ &= \frac{1}{4\pi} \left(\frac{2}{\sqrt{\frac{1}{4} - \left(\frac{1-z}{4}\right)} \cdot \sqrt{1-z}} \right) \\ &= \frac{1}{4\pi} \left(\frac{2}{\frac{\sqrt{z}}{2} \cdot \sqrt{1-z}} \right) \\ &= \frac{1}{\pi \sqrt{z(1-z)}} = \gamma^n(z) \end{aligned}$$

Hence, $\gamma^{n+1}(z) = \gamma^n(z)$, $\frac{1}{\pi \sqrt{z(1-z)}}$ is invariant under the mapping $z^{n+1} = 4z(1-z)$.

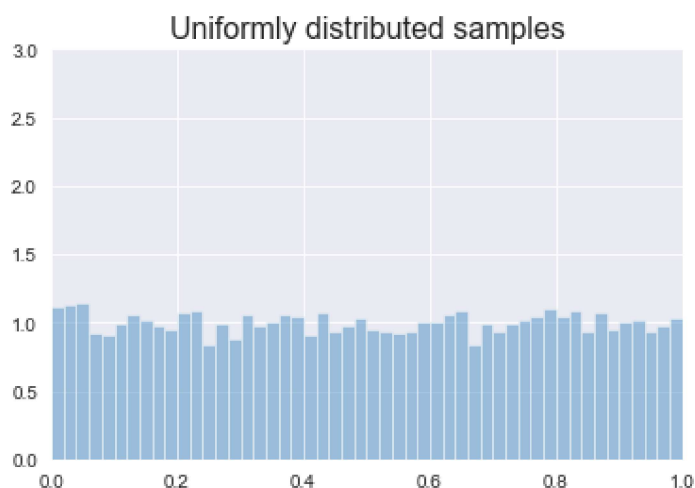
```
In [1]: 1 import numpy as np
2 from matplotlib import pyplot as plt
3 %matplotlib inline
4 import seaborn as sns
5 sns.set_style('darkgrid')
6 #sns.set(rc={'figure.figsize':(4,4)})
```

b) *Simulation exercise:* Consider z^0 as being uniformly distributed on $(0,1)$. Write a code to determine how this uniform distribution evolves under (1). Produce histograms after 5, 10, 100 iterations and overlay the invariant density (2).

We start from $U[0,1]$ and see how this density evolves.

```
In [2]: 1 samples = np.random.uniform(low=0, high=1, size=10000)
2
3 fig, ax = plt.subplots()
4 ax.set(xlim=(0,1), ylim=(0,3))
5 sns.distplot(samples, bins=50, ax=ax)
6 ax.get_lines()[0].remove()
7 plt.title("Uniformly distributed samples", fontsize=16)
```

Out[2]: Text(0.5, 1.0, 'Uniformly distributed samples')



Deterministic mapping function

$$z^{n+1} = \nu z^n (1 - z^n) \quad (1)$$

```
In [3]: 1 def DeterministicDynamicalSystem (samples_mapped, iterations):
2     for i in range(iterations):
3         samples_mapped = 4*samples_mapped*(1-samples_mapped)
4     return samples_mapped
```

Invariant density function

$$\frac{1}{\pi \sqrt{z(1-z)}}. \quad (2)$$

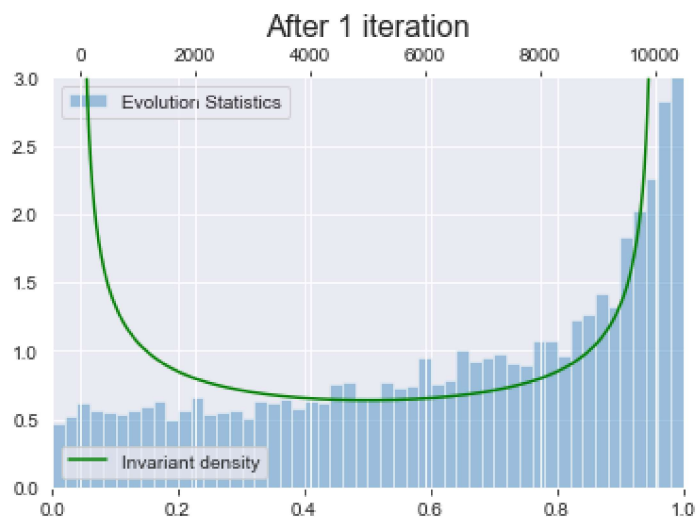
```
In [4]: 1 def InvariantDensity(z):
2         z = 1/(np.pi*np.sqrt(z*(1-z)))
3         return z
4 invariant_density = InvariantDensity(np.linspace(0.001, 0.999, 10000))
```

function for plotting histogram

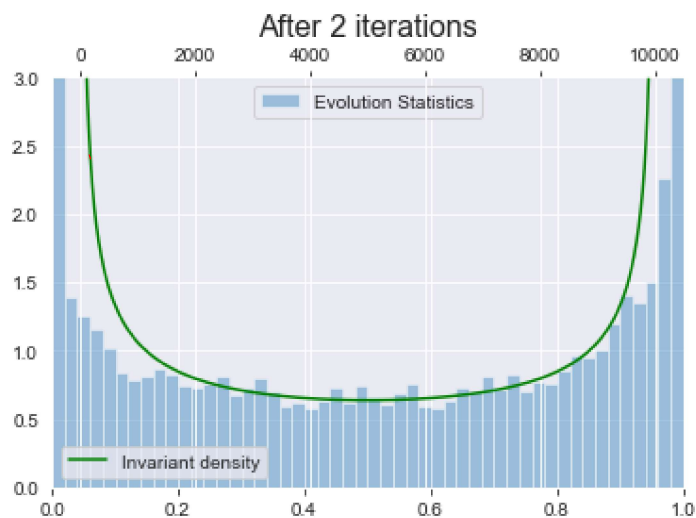
```
In [5]: 1 def PlotHistogram(samples_iter, title):
2         fig, ax = plt.subplots()
3         ax.set(xlim=(0,1), ylim=(0,3))
4         sns.distplot(samples_iter, bins=50, ax=ax)
5         ax.get_lines()[0].remove()
6         ax2 = ax.twinx()
7         ax2.plot(invariant_density,color='green')
8         ax2.legend(labels=['Invariant density'])
9         ax.legend(labels=['Evolution Statistics'])
10        plt.title(title, fontsize=16)
```

Now we will see the evolution of the initial density.

```
In [6]: 1 samples_iter1 = DeterministicDynamicalSystem (samples, 1)
2 PlotHistogram(samples_iter1, "After 1 iteration")
```



```
In [7]: 1 samples_iter2 = DeterministicDynamicalSystem (samples, 2)
2 PlotHistogram(samples_iter2, "After 2 iterations")
```



Unif Distr.

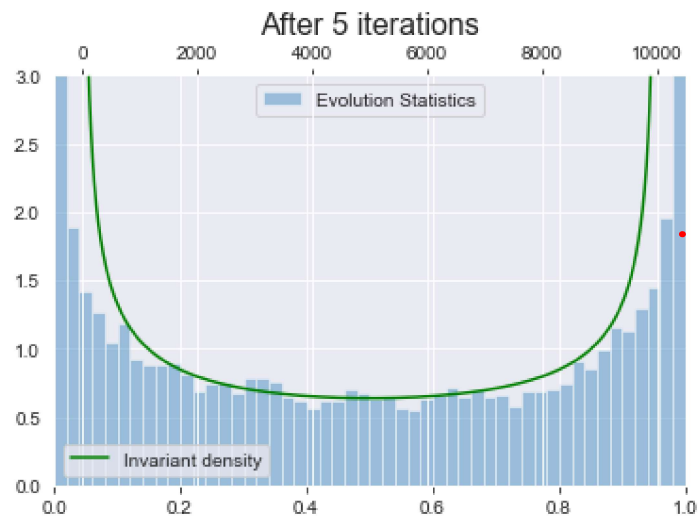
→

↪

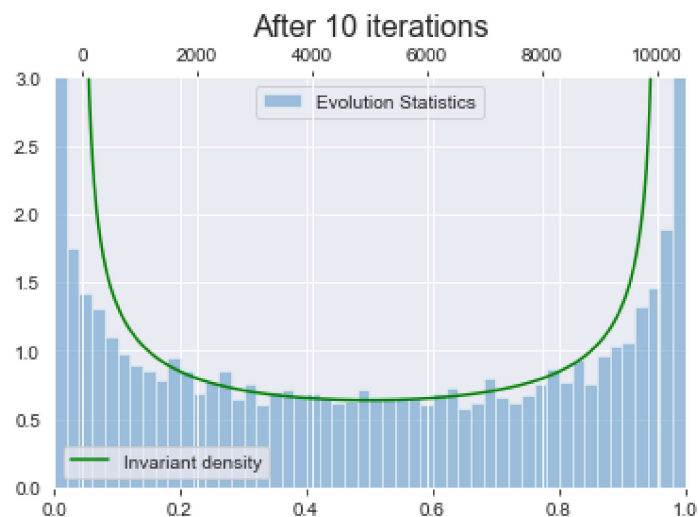
"geometric ergodicity"

Even after 2 iterations, it already approaches to invariant density!

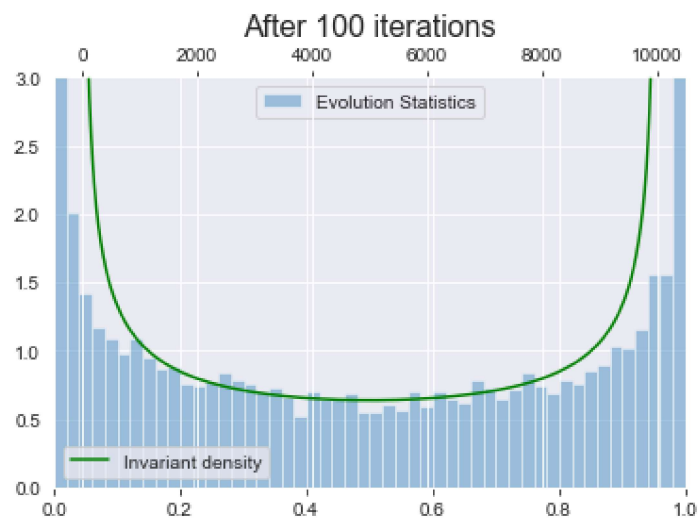
```
In [8]: ▶ 1 samples_iter5 = DeterministicDynamicalSystem (samples, 5)
        2 PlotHistogram(samples_iter5, "After 5 iterations")
```



```
In [9]: ▶ 1 samples_iter10 = DeterministicDynamicalSystem (samples, 10)
        2 PlotHistogram(samples_iter10, "After 10 iterations")
```



```
In [10]: ▶ 1 samples_iter100 = DeterministicDynamicalSystem (samples, 100)
        2 PlotHistogram(samples_iter100, "After 100 iterations")
```



Exercise 2. The weather in a simple earth system has 4 possible states: sun, rain, snow, hail. We want to use a discrete time Markov chain to model the daily weather of this system. The weather does not change from one day to the next with probability $1 - 3\alpha$. If the weather does change, then all other 3 states are equally likely. We have that $0 < \alpha < \frac{1}{3}$.

(a) Show that $P_{i,i}^n = \frac{1}{4} + \frac{3}{4}(1 - 4\alpha)^n \quad \forall i = \{1, 2, 3, 4\}$ where n corresponds to the n th time index. Then obtain an expression for $P_{i,j}^n$ for $i \neq j$.

(Approach using induction)

Since the other $(P^n)_{ij}$ for $i \neq j$ take the same values in each column,

$$\begin{aligned} \text{three } (P^n)_{ij} &= (1 - (\frac{1}{4} + \frac{3}{4}(1 - 4\alpha)^n)) / 3 \\ &= \frac{1}{4} - \frac{1}{4}(1 - 4\alpha)^n \end{aligned}$$

claim: $(P^n)_{ii} = \frac{1}{4} (1 + 3(1 - 4\alpha)^n)$ and $(P^n)_{ij} = \frac{1}{4} (1 - (1 - 4\alpha)^n)$ for $i \neq j, i, j \in \{1, 2, 3, 4\}$

Proof by induction:

* Base case: $(P^1)_{ii} = \frac{1}{4} + \frac{3}{4} - 3\alpha = 1 - 3\alpha$
 $(P^1)_{ij} = \frac{1}{4} - \frac{1}{4} + \alpha = \alpha$

* Inductive step: Show that for any $n \geq 1$ with $C(n) \Rightarrow C(n+1)$

$$\begin{aligned} (P^{n+1})_{ii} &= \sum_{k=1}^4 (P)_{ik} (P^n)_{ki} \\ &\stackrel{C(n)}{=} P_{ii} (P^n)_{ii} + 3 \cdot P_{il} (P^n)_{li} \quad \text{for } l \neq i \\ &\stackrel{C(n)}{=} (1 - 3\alpha) \left(\frac{1}{4} + \frac{3}{4}(1 - 4\alpha)^n \right) + 3 \cdot \alpha \left(\frac{1}{4} - \frac{1}{4}(1 - 4\alpha)^n \right) \\ &= \frac{1}{4} (1 + 3(1 - 4\alpha)^n - 3\alpha - 9\alpha(1 - 4\alpha)^n + 3\alpha - 3\alpha(1 - 4\alpha)^n) \\ &= \frac{1}{4} (1 + (1 - 4\alpha)^n (3 - 12\alpha)) \\ &= \frac{1}{4} (1 + 3 \cdot (1 - 4\alpha)^{n+1}) \\ &= \frac{1}{4} + \frac{3}{4}(1 - 4\alpha)^{n+1} \quad \times \end{aligned}$$

$$\begin{aligned} (P^{n+1})_{ij} &= \sum_{k=1}^4 (P)_{ik} (P^n)_{kj} \\ &\stackrel{C(n)}{=} (P)_{ii} (P^n)_{ij} + (P)_{ij} (P^n)_{jj} + 2 \cdot (P)_{il} (P^n)_{lj} \quad \text{for } l \neq i, j \\ &\stackrel{C(n)}{=} (1 - 3\alpha) \cdot \frac{1}{4} (1 - (1 - 4\alpha)^n) + \alpha \cdot \frac{1}{4} (1 + 3(1 - 4\alpha)^n) + 2 \cdot \alpha \cdot \frac{1}{4} (1 - (1 - 4\alpha)^n) \\ &= \frac{1}{4} (1 - (1 - 4\alpha)^n - 3\alpha + 3\alpha(1 - 4\alpha)^n + \alpha + 3\alpha(1 - 4\alpha)^n + 2\alpha - 2\alpha(1 - 4\alpha)^n) \\ &= \frac{1}{4} (1 - (1 - 4\alpha)^n (1 - 4\alpha)) \\ &= \frac{1}{4} (1 - (1 - 4\alpha)^{n+1}) \quad \times \end{aligned}$$

By sym. it has to hold that all P_{ij}^{n+1} are equal if $i \neq j$

$$\Rightarrow P_{ij}^{n+1} = \frac{1}{3} (1 - P_{ii}^n)$$

(alternative approach using diagonalization)

$$P = \begin{bmatrix} 1-3\alpha & \alpha & \alpha & \alpha \\ \alpha & 1-3\alpha & \alpha & \alpha \\ \alpha & \alpha & 1-3\alpha & \alpha \\ \alpha & \alpha & \alpha & 1-3\alpha \end{bmatrix}$$

Since P is a real symmetric matrix, P can be diagonalized, i.e.

$$P = Q\Lambda Q^T \text{ where } \Lambda \text{ is } \text{diag}(\lambda_1, \lambda_2, \lambda_3, \lambda_4) \quad \Lambda \in \mathbb{R}^{4 \times 4}$$

and Q is orthogonal. $Q \in \mathbb{R}^{4 \times 4}$

Since P is irreducible, by Perron-Frobenius theorem,

We know that only one eigenvalue of P is 1. $\Rightarrow \lambda_1 = 1$

Since the sum of the all eigenvalue of P is equal to its trace,

$$1 + \lambda_2 + \lambda_3 + \lambda_4 = 4 - 12\alpha \Rightarrow \lambda_2 + \lambda_3 + \lambda_4 = 3 - 12\alpha$$

Since the converge rate P_{ij}^n for $i \neq j$ are the same, $\lambda_2 = \lambda_3 = \lambda_4 = 1 - 4\alpha$.

Eigenvalues and Eigenvectors of P :

$$\begin{array}{l} \lambda^1 \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} \\ V^1 \end{array} \quad \begin{array}{l} (1-4\alpha) \\ \lambda^2 \text{ with} \\ \text{multiplicity} \\ 3 \end{array} \quad \begin{array}{l} \begin{bmatrix} 1 \\ -1 \\ 1 \\ -1 \end{bmatrix}, \begin{bmatrix} 1 \\ 1 \\ -1 \\ -1 \end{bmatrix}, \begin{bmatrix} 1 \\ -1 \\ -1 \\ 1 \end{bmatrix} \\ V^2 \quad V^3 \quad V^4 \end{array}$$

Since $\|V^1\| = \|V^2\| = \|V^3\| = \|V^4\| = \sqrt{4} = 2$,

$$Q = \begin{bmatrix} 1/2 & 1/2 & 1/2 & 1/2 \\ 1/2 & -1/2 & 1/2 & -1/2 \\ 1/2 & 1/2 & -1/2 & -1/2 \\ 1/2 & -1/2 & -1/2 & 1/2 \end{bmatrix}$$

$$P = Q\Lambda Q^T = \frac{1}{4} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 1 & -1 & 1 & -1 \\ 1 & 1 & -1 & -1 \\ 1 & -1 & -1 & 1 \end{bmatrix} \begin{bmatrix} 1 & & & \\ & 1-4\alpha & & \\ & & 1-4\alpha & \\ & & & 1-4\alpha \end{bmatrix} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 1 & -1 & 1 & -1 \\ 1 & 1 & -1 & -1 \\ 1 & -1 & -1 & 1 \end{bmatrix}$$

$$\begin{aligned} P^n &= Q\Lambda^n Q^T = \frac{1}{4} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 1 & -1 & 1 & -1 \\ 1 & 1 & -1 & -1 \\ 1 & -1 & -1 & 1 \end{bmatrix} \begin{bmatrix} 1^n & & & \\ & (1-4\alpha)^n & & \\ & & (1-4\alpha)^n & \\ & & & (1-4\alpha)^n \end{bmatrix} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 1 & -1 & 1 & -1 \\ 1 & 1 & -1 & -1 \\ 1 & -1 & -1 & 1 \end{bmatrix} \\ &= \frac{1}{4} \begin{bmatrix} 1 & (1-4\alpha)^n & (1-4\alpha)^n & (1-4\alpha)^n \\ 1 & -(1-4\alpha)^n & (1-4\alpha)^n & -(1-4\alpha)^n \\ 1 & (1-4\alpha)^n & -(1-4\alpha)^n & (1-4\alpha)^n \\ 1 & -(1-4\alpha)^n & -(1-4\alpha)^n & (1-4\alpha)^n \end{bmatrix} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 1 & -1 & 1 & -1 \\ 1 & 1 & -1 & -1 \\ 1 & -1 & -1 & 1 \end{bmatrix} \end{aligned}$$

$$P^n = \frac{1}{4} \begin{bmatrix} 1 + 3(1-4\alpha)^n & 1 - (1-4\alpha)^n & 1 - (1-4\alpha)^n & 1 - (1-4\alpha)^n \\ 1 - (1-4\alpha)^n & 1 + 3(1-4\alpha)^n & 1 - (1-4\alpha)^n & 1 - (1-4\alpha)^n \\ 1 - (1-4\alpha)^n & 1 - (1-4\alpha)^n & 1 + 3(1-4\alpha)^n & 1 - (1-4\alpha)^n \\ 1 - (1-4\alpha)^n & 1 - (1-4\alpha)^n & 1 - (1-4\alpha)^n & 1 + 3(1-4\alpha)^n \end{bmatrix}$$

Hence, $P_{ii}^n = \frac{1}{4} + \frac{3}{4}(1-4\alpha)^n$ for $\forall i = \{1, 2, 3, 4\}$

and $P_{ij}^n = \frac{1}{4} + \frac{1}{4}(1-4\alpha)^n$ for $i \neq j$ *

(b) What is the long-run proportion of time the chain is in each state?

From $0 < \alpha < \frac{1}{3}$, it follows that $|1 - 4\alpha| < 1$.

As $n \rightarrow \infty$, $(1 - 4\alpha)^n \rightarrow 0$.

Hence, $P_{ii}^n = \frac{1}{4}$ and $P_{ij}^n = \frac{1}{4}$ for $i \neq j$ and $i, j \in \{1, 2, 3, 4\}$

and the limiting distribution is $(\frac{1}{4}, \frac{1}{4}, \frac{1}{4}, \frac{1}{4})$.

→ chain conv. to unif distr.

+ Unif. now is its stat.

(either by using that this is implied by
conv., or directly showing

$$P \begin{pmatrix} \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \end{pmatrix} = \begin{pmatrix} \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \end{pmatrix}$$

$$\alpha \neq 0, 1$$

Exercise 3. Consider the following 2-state discrete time Markov chain taking values $\{1, -1\}$ with $\pi_Z^0 = \pi_Z^*$ (invariant distribution) with transition probability matrix

$$P = \begin{bmatrix} \alpha & 1-\alpha \\ 1-\alpha & \alpha \end{bmatrix}$$

i) Show analytically that the (unnormalised) autocorrelation at lag τ , $\tau > 0$ defined as

$$\mathbb{E}[Z^n Z^{n+\tau}]$$

$$P \pi_Z^* = \begin{bmatrix} \alpha & 1-\alpha \\ 1-\alpha & \alpha \end{bmatrix} \begin{bmatrix} \pi_{z_1}^* \\ \pi_{z_2}^* \end{bmatrix} = \begin{bmatrix} \alpha \pi_{z_1}^* + (1-\alpha) \pi_{z_2}^* \\ (1-\alpha) \pi_{z_1}^* + \alpha \pi_{z_2}^* \end{bmatrix} = \begin{bmatrix} \pi_{z_1}^* \\ \pi_{z_2}^* \end{bmatrix} \Rightarrow \dots \text{Simple computation} \dots \Rightarrow \pi_{z_1}^* = \pi_{z_2}^* = 1/2$$

$$\text{claim: } p^n = \frac{1}{2} \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix} + \frac{(2\alpha-1)^n}{2} \begin{pmatrix} 1 & -1 \\ -1 & 1 \end{pmatrix} \Leftrightarrow C(n)$$

proof by induction: Show the claim for $n \geq 1$ with $C(n) \Rightarrow C(n+1)$

$$p^1 = \frac{1}{2} \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix} + \frac{(2\alpha-1)^1}{2} \begin{pmatrix} 1 & -1 \\ -1 & 1 \end{pmatrix} = \begin{pmatrix} 1/2 & 1/2 \\ 1/2 & 1/2 \end{pmatrix} + \begin{pmatrix} \alpha-1/2 & -\alpha+1/2 \\ -\alpha+1/2 & \alpha-1/2 \end{pmatrix} = \begin{pmatrix} \alpha & 1-\alpha \\ 1-\alpha & \alpha \end{pmatrix}$$

$$\begin{aligned} p^{n+1} &= P \cdot p^n = \begin{pmatrix} \alpha & 1-\alpha \\ 1-\alpha & \alpha \end{pmatrix} \left(\frac{1}{2} \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix} + \frac{(2\alpha-1)^n}{2} \begin{pmatrix} 1 & -1 \\ -1 & 1 \end{pmatrix} \right) \\ &\stackrel{C(n)}{=} \frac{1}{2} \begin{pmatrix} \alpha+(1-\alpha) & \alpha+(1-\alpha) \\ (1-\alpha)+\alpha & (1-\alpha)+\alpha \end{pmatrix} + \frac{(2\alpha-1)^n}{2} \begin{pmatrix} \alpha-(1-\alpha) & -\alpha+(1-\alpha) \\ (1-\alpha)-\alpha & -(1-\alpha)+\alpha \end{pmatrix} \\ &= \frac{1}{2} \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix} + \frac{(2\alpha-1)^n}{2} \begin{pmatrix} 2\alpha-1 & -2\alpha+1 \\ -2\alpha+1 & 2\alpha-1 \end{pmatrix} \\ &= \frac{1}{2} \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix} + \frac{(2\alpha-1)^{n+1}}{2} \begin{pmatrix} 1 & -1 \\ -1 & 1 \end{pmatrix} \end{aligned}$$

$$\begin{aligned} \mathbb{E}[Z^n Z^{n+\tau}] &= 1 \cdot 1 \cdot P(Z^n=1, Z^{n+\tau}=1) + 1 \cdot (-1) \cdot P(Z^n=1, Z^{n+\tau}=-1) \\ &\quad + (-1) \cdot 1 \cdot P(Z^n=-1, Z^{n+\tau}=1) + (-1) \cdot (-1) \cdot P(Z^n=-1, Z^{n+\tau}=-1) \\ &= P(Z^{n+\tau}=1 | Z^n=1) P(Z^n=1) - P(Z^{n+\tau}=-1 | Z^n=1) P(Z^n=1) \\ &\quad - P(Z^{n+\tau}=1 | Z^n=-1) P(Z^n=-1) + P(Z^{n+\tau}=-1 | Z^n=-1) P(Z^n=-1) \\ &= (P^\tau)_{22} \cdot \frac{1}{2} - (P^\tau)_{12} \cdot \frac{1}{2} - (P^\tau)_{21} \cdot \frac{1}{2} + (P^\tau)_{11} \cdot \frac{1}{2} \\ &= \frac{1}{2} \left(\frac{1}{2} + \frac{(2\alpha-1)^\tau}{2} - \frac{1}{2} + \frac{(2\alpha-1)^\tau}{2} - \frac{1}{2} + \frac{(2\alpha-1)^\tau}{2} + \frac{1}{2} + \frac{(2\alpha-1)^\tau}{2} \right) \\ &= \frac{1}{2} \frac{4(2\alpha-1)^\tau}{2} \\ &= (2\alpha-1)^\tau \checkmark \end{aligned}$$

$$P = \begin{bmatrix} P_{11} & P_{12} \\ P_{21} & P_{22} \end{bmatrix} = \begin{bmatrix} \alpha & 1-\alpha \\ 1-\alpha & \alpha \end{bmatrix}$$

Def. of Exp. value

 p^n is in stat. distr. $\begin{pmatrix} 1/2 \\ 1/2 \end{pmatrix}$

ii) What happens to the autocorrelation function as the lag τ increases? Explain which property of the Markov chain leads to this behaviour.

$$\begin{aligned}\text{Cov}[Z^n, Z^{n+\tau}] &= \mathbb{E}[Z^n Z^{n+\tau}] - \mathbb{E}[Z^n] \mathbb{E}[Z^{n+\tau}] \\ &= \mathbb{E}[Z^n Z^{n+\tau}] \\ &= (2\alpha - 1)^\tau\end{aligned}$$

* Case $0 < \alpha < 1$, then $(2\alpha - 1)^\tau \rightarrow 0$ as $\alpha \rightarrow \infty$, which means that Z^n and $Z^{n+\tau}$ becomes uncorrelated.

The Markov chain with $0 < \alpha < 1$ is irreducible and aperiodic. \Rightarrow ergodic

* Case $\alpha = 0$, then the autocorrelation function oscillates between -1 and 1 .

The Markov chain with $\alpha = 0$, i.e. $P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$ has one communicating class (Irreducible and Periodic) with periodicity 2.

* Case $\alpha = 1$, then $(2\alpha - 1)^\tau$ stays 1,

which means that Z^n and $Z^{n+\tau}$ are completely positively correlated.

The Markov chain with $\alpha = 1$, i.e. $P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$ consists of

two self-transition communicating classes (reducible).

There are no unique stationary (invariant) distribution in this case.