

# Solutions for Practice Final

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## Final Exams

See exercise: 3.37, 3.67, 5.10, 5.22, 6.6, 6.10, 6.32, 8.6, 8.13, 10.3

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**Note:** This study guide is used for Botao Jin's sections only. Comments, bug reports: [b\\_jin@ucsb.edu](mailto:b_jin@ucsb.edu)

(3.37) Suppose  $X$  has a c.d.f.

$$F(x) = \begin{cases} \frac{x}{1+x} & x \geq 0 \\ 0 & x < 0 \end{cases}$$

(a) P.D.F  $f(x) = F'(x)$

$$x < 0: f(x) = 0$$

$$x > 0: f(x) = \left(1 - \frac{1}{1+x}\right)' = \frac{1}{(1+x)^2}$$

$$\text{so } f(x) = \begin{cases} 1/(1+x)^2 & x > 0 \\ 0 & x < 0 \end{cases}$$

(b)  $\mathbb{P}(2 < X < 3)$

$$= F(3) - F(2) = \frac{3}{4} - \frac{2}{3} = \frac{1}{12}$$

(c)  $\mathbb{E}[(1+X)^2 e^{-2X}]$

$$= \int_{-\infty}^{\infty} f(x) (1+x)^2 e^{-2x} dx$$

$$= \int_0^{\infty} e^{-2x} dx = 1/2$$

$$(3.67) \quad Z \sim \mathcal{N}(0, 1), \quad X \sim \mathcal{N}(\mu, \sigma^2)$$

$$(a) \quad \mathbb{E}[Z^3] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \underbrace{x^3}_{\text{odd}} \underbrace{\exp\{-\frac{x^2}{2}\}}_{\text{even}} dx = 0$$

$$(b) \quad X = \sigma Z + \mu$$

$$\begin{aligned} \mathbb{E}[X^3] &= \mathbb{E}[(\sigma Z + \mu)^3] \\ &= \mathbb{E}[\sigma^3 Z^3 + 3\sigma^2 Z^2 \mu + 3\sigma Z \mu^2 + \mu^3] \\ &= \sigma^3 \cancel{\mathbb{E}[Z^3]} + 3\sigma^2 \mu \overset{1}{\mathbb{E}[Z^2]} + 3\sigma \mu^2 \cancel{\mathbb{E}[Z]} + \mu^3 \\ &= 3\sigma^2 \mu + \mu^3 \end{aligned}$$

$$(5.10) \quad M(t) = \left(\frac{1}{5} + \frac{4}{5} e^t\right)^{30}$$

by checking distribution Table,  $X \sim \text{Bin}(30, 4/5)$

$$(5.22) \quad M_Y(t) = \mathbb{E}[e^{tY}] = \mathbb{E}[e^{t(3X-2)}]$$

$$= e^{-2t} \mathbb{E}[e^{(3t)X}] = e^{-2t} \varphi_X(3t)$$

$$\text{where } \varphi_X(t) = \mathbb{E}[e^{tX}]$$

$$= \frac{\lambda}{\lambda - t} \quad (\text{by checking Distribution Table})$$

$$\text{So } M_Y(t) = \begin{cases} e^{-2t} \frac{\lambda}{\lambda - 3t} & t < \lambda/3 \\ \infty & \text{o/w.} \end{cases}$$

be careful about  
the domain of  $X$

$$(6.6) \quad f(x,y) = \begin{cases} xe^{-x(1+y)} & x > 0, y > 0 \\ 0 & \text{o/w} \end{cases}$$

$$\begin{aligned} (a) \quad f_X(x) &= \int_0^\infty xe^{-x(1+y)} dy \\ &= xe^{-x} \int_0^\infty e^{-xy} dy \\ &= xe^{-x} \left(-\frac{1}{x}\right) e^{-xy} \Big|_{y=0}^{y=\infty} = e^{-x} \quad x > 0 \end{aligned}$$

$$\begin{aligned} f_Y(y) &= \int_0^\infty xe^{-x(1+y)} dx \\ &= \frac{1}{(1+y)^2} \quad y > 0 \end{aligned}$$

Remark: You can consider  $(1+y)e^{-(1+y)x}$  as the density of  $\exp(\lambda=1+y)$ , then use the fact that

$$\int_0^\infty x(1+y)e^{-(1+y)x} dx = \frac{1}{1+y} \quad (\text{Expected value is } \frac{1}{1+y})$$

$$\text{so } \int_0^\infty xe^{-(1+y)x} dx = \frac{1}{(1+y)^2}.$$

$$\begin{aligned} (b) \quad \mathbb{E}[XY] &= \int_0^\infty \int_0^\infty xy f(x,y) dx dy \\ &= \int_0^\infty \int_0^\infty x^2 y e^{-x(1+y)} dx dy \\ &= \int_0^\infty x^2 e^{-x} \underbrace{\left[ \int_0^\infty y e^{-xy} dy \right]}_{\frac{1}{x^2} \text{ (using same method as part a)}} dx = \int_0^\infty e^{-x} dx = 1 \end{aligned}$$

$$(c) \quad \mathbb{E}\left[\frac{X}{1+Y}\right] = \int_0^\infty \int_0^\infty \frac{x}{1+y} x e^{-x(1+y)} dx dy$$

$$= \int_0^\infty \frac{1}{1+y} \left\{ \int_0^\infty x^2 e^{-x(1+y)} dx \right\} dy$$

$\frac{(1+y)^3}{\Gamma(3)} x^2 e^{-x(1+y)}$  : density of  $\text{Gamma}(3, 1+y)$

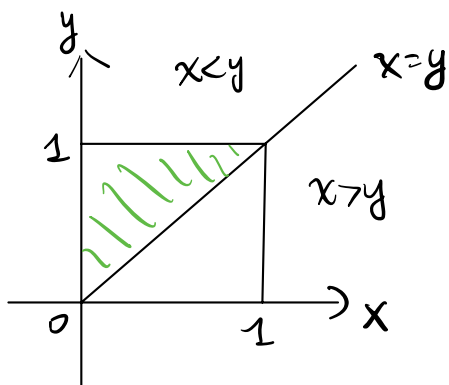
$$\text{so } \int_0^\infty x^2 e^{-x(1+y)} dx = \frac{2}{(1+y)^3}$$

$$= 2 \int_0^\infty \frac{1}{(1+y)^4} dy$$

$$= -\frac{2}{3} \cdot \frac{1}{(1+y)^3} \Big|_0^\infty = \frac{2}{3}$$

(6.10)  $X, Y \sim \text{Unif}(0, 1) \quad X \perp Y$

$$f(x, y) = \begin{cases} 1 & 0 < x < 1, 0 < y < 1 \\ 0 & \text{o.w.} \end{cases}$$



$$\mathbb{P}(X < Y)$$

$$= \int_0^1 \int_0^y dx dy$$

$$= \int_0^1 y dy = \frac{1}{2}$$

(6.32) Support of  $N = \mathbb{N}$

Support of  $Y = \{1, 2\}$

$$\mathbb{P}(N=k, Y=1)$$

$$= \mathbb{P}(\text{White in the first } (k-1)^{\text{th}} \text{ time, Green at the } k^{\text{th}} \text{ time})$$

$$= \left(\frac{2}{9}\right)^{k-1} \cdot \frac{4}{9}$$

$$\text{Similarly, } \mathbb{P}(N=k, Y=2) = \left(\frac{2}{9}\right)^{k-1} \cdot \frac{1}{3}$$

$$\text{Note: } \mathbb{P}(N=k) = \left(\frac{2}{9}\right)^{k-1} \cdot \frac{7}{9} \quad \text{b/c } N \sim \text{Geo}(7/9)$$

$$\mathbb{P}(Y=1) = \sum_{k=1}^{\infty} \mathbb{P}(N=k, Y=1)$$

$$= \frac{4}{9} \sum_{k=1}^{\infty} \left(\frac{2}{9}\right)^{k-1} = \frac{4}{9} \cdot \frac{1}{1-2/9} = \frac{4}{7}$$

$$\mathbb{P}(Y=2) = \sum_{k=1}^{\infty} \mathbb{P}(N=k, Y=2)$$

$$= \frac{1}{3} \sum_{k=1}^{\infty} \left(\frac{2}{9}\right)^{k-1} = \frac{1}{3} \cdot \frac{1}{1-2/9} = \frac{3}{7}$$

Thus for any  $k \in \mathbb{N}$ ,

$$\mathbb{P}(N=k, Y=1) = \mathbb{P}(N=k) \mathbb{P}(Y=1)$$

$$\mathbb{P}(N=k, Y=2) = \mathbb{P}(N=k) \mathbb{P}(Y=2)$$

Hence,  $N \perp Y$

□

$$(8.6) \quad X \sim \text{Geo}(p), \quad Y \sim \text{Bin}(n, r)$$

$$\mathbb{E}[X] = \frac{1}{p} \quad \text{Var}(X) = \frac{1-p}{p^2} \quad \text{so } \mathbb{E}X^2 = (\mathbb{E}X)^2 + \text{Var}(X) = \frac{2-p}{p^2}$$

$$\mathbb{E}[Y] = nr \quad \text{Var}(Y) = nr(1-r) \quad \text{so } \mathbb{E}Y^2 = (\mathbb{E}Y)^2 + \text{Var}(Y) = n(n-1)r^2 + nr$$

using the distribution sheet

$$(a) \quad \mathbb{E}[X+Y] = \frac{1}{p} + nr$$

$$(b) \quad \mathbb{E}[XY] = \mathbb{E}X \cdot \mathbb{E}Y = \frac{nr}{p}$$

$$(c) \quad \mathbb{E}[X^2+Y^2] = \frac{2-p}{p^2} + n(n-1)r^2 + nr$$

$$(d) \quad \mathbb{E}[(X+Y)^2] = \mathbb{E}[X^2+Y^2+2XY] = \frac{2-p}{p^2} + n(n-1)r^2 + nr + \frac{2nr}{p}$$

$$(8.13) \quad \text{Let } Z = X_1 + X_2 + \dots + X_{36}$$

with  $X_1, \dots, X_{36}$  are i.i.d. Random Variables

independent and identically distributed

$$\text{and } M_{X_i}(t) = \frac{1}{2}e^{-t} + \frac{2}{5} + \frac{1}{10}e^{t/2}$$

then P.M.F of  $X_i$  is

|              |               |               |                |
|--------------|---------------|---------------|----------------|
| $X_i$        | -1            | 0             | $\frac{1}{2}$  |
| $\mathbb{P}$ | $\frac{1}{2}$ | $\frac{2}{5}$ | $\frac{1}{10}$ |

for  $i=1, 2, \dots, 36$

$$(10.3) \quad \mathbb{P}(Y=k) = \frac{1}{6} \quad k=1, 2, \dots, 6$$

$$X|Y=k \sim \text{Bin}(k, \frac{1}{2})$$

flip the coins for  $k$  times

for  $n=0, 1, 2, \dots, 6$

$$\begin{aligned} \mathbb{P}(X=n) &= \sum_{k=n}^6 \mathbb{P}(X=n|Y=k) \mathbb{P}(Y=k) \quad \swarrow \text{Law of Total Probability} \\ &= \sum_{k=n}^6 \binom{k}{n} \left(\frac{1}{2}\right)^k \frac{1}{6} \end{aligned}$$

$$\mathbb{E}[X] = \mathbb{E}[\mathbb{E}[X|Y]]$$

$$= \sum_{k=1}^6 \mathbb{E}[X|Y=k] \mathbb{P}(Y=k)$$

$$= \sum_{k=1}^6 \frac{k}{2} \cdot \frac{1}{6} = \frac{7}{4}$$