

Moment Generating Function

Def: For a random variable X , the Moment Generating Function (M.G.F) is defined by

$$M_X(t) = \mathbb{E}[e^{tX}] = \begin{cases} \sum_{k \in S_X} e^{tk} \mathbb{P}(X=k) & X: \text{discrete} \\ \int_{-\infty}^{\infty} e^{tx} f(x) dx & X: \text{continuous} \end{cases}$$

Moreover, given the M.G.F of X , the n -th moment of X can be obtained by

$$\mathbb{E}[X^n] = M_X^{(n)}(t) \big|_{t=0}$$

Example 1: Suppose a R.V. X has density

$$f(x) = \frac{1}{2} e^{-|x|} = \begin{cases} \frac{1}{2} e^{-x} & x > 0 \\ \frac{1}{2} e^x & x < 0 \end{cases}$$

m.g.f: $M_X(t) = \mathbb{E}[e^{tX}]$

$$= \int_0^{\infty} \frac{1}{2} e^{(t-1)x} dx + \int_{-\infty}^0 \frac{1}{2} e^{(t+1)x} dx$$

$$= \frac{1}{2} \left\{ \underbrace{\int_0^{\infty} e^{(t-1)x} dx}_{\textcircled{1}} + \underbrace{\int_{-\infty}^0 e^{(t+1)x} dx}_{\textcircled{2}} \right\}$$

For ①:

- when $t-1 > 0$, $e^{(t-1)x} \uparrow \infty$ as $x \uparrow \infty$, ① = ∞
- when $t-1 = 0$, ① = $\int_0^{\infty} dx = \infty$
- when $t-1 < 0$, then

$$\textcircled{1} = \int_0^{\infty} e^{(t-1)x} dx = \frac{1}{1-t}$$

For ②:

- When $t+1 < 0$, $e^{(t+1)x} \uparrow \infty$ as $x \downarrow -\infty$, ② = ∞
- When $t+1 = 0$, ② = $\int_{-\infty}^0 dx = \infty$
- when $t+1 > 0$, then

$$\textcircled{2} = \int_{-\infty}^0 e^{(t+1)x} dx = \frac{1}{1+t}$$

Combining ① and ②: when $-1 < t < 1$,

$$M_X(t) = \frac{1}{2} \left(\frac{1}{1+t} + \frac{1}{1-t} \right) = \frac{1}{1-t^2}$$

$$\mathbb{E}[X] = M'_X(0) = \left. \frac{2t}{(1-t^2)^2} \right|_{t=0} = 0$$

Remark: $\mathbb{E}[X]$ can also be obtained by definition, which is

$$\mathbb{E}[X] = \int_{-\infty}^{\infty} \underbrace{x \frac{1}{2} e^{-|x|}}_{\text{odd}} dx = 0$$

x : odd function; $\frac{1}{2} e^{-|x|}$: even function

Example 2: Suppose X has density

$$f(x) = \begin{cases} x e^{-x} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

m.g.f $M_X(t) = \mathbb{E}[e^{tx}]$

$$= \int_0^{\infty} x e^{-x} e^{tx} dx = \int_0^{\infty} x e^{-(1-t)x} dx$$

(1) when $1-t \leq 0$, $x e^{-(1-t)x} \nearrow \infty$ as $x \nearrow \infty$

(2) when $1-t > 0$, use the fact that

$$\int_0^{\infty} x^{\alpha-1} e^{-\beta x} dx = \frac{\Gamma(\alpha)}{\beta^{\alpha}} \quad \text{for } \alpha, \beta > 0$$

$$\Rightarrow \int_0^{\infty} x e^{-(1-t)x} dx = \frac{\Gamma(2)}{(1-t)^2} = \frac{1}{(1-t)^2} \quad (\Gamma(2) = 1! = 1)$$

$$M_X(t) = \left(\frac{1}{1-t} \right)^2 \quad t < 1$$

Given m.g.f, use $\mathbb{E}[X^n] = M_X^{(n)}(t) \big|_{t=0}$

$$M_X'(t) = \frac{2}{(1-t)^3}, \quad M_X''(t) = \frac{2 \cdot 3}{(1-t)^4}$$

$$M_X'''(t) = \frac{2 \cdot 3 \cdot 4}{(1-t)^5}, \quad \dots, \quad M_X^{(n)}(t) = \frac{(n+1)!}{(1-t)^{n+2}}$$

Thus, $\mathbb{E}[X^n] = (n+1)!$

Remark 1: Given $f(x) = x e^{-x}$ $x > 0$, we know

$X \sim \text{Gamma}(r=2, \lambda=1)$, by distribution sheet:

$$M_X(t) = \left(\frac{\lambda}{\lambda - t} \right)^r = \frac{1}{(1-t)^2}$$

Remark 2: The n -th moment of X can also be obtained by

$$\begin{aligned} \mathbb{E}[X^n] &= \int_0^\infty x^n \cdot x e^{-x} dx = \int_0^\infty x^{n+1} e^{-x} dx \\ &= \Gamma(n+2) = (n+1)! \end{aligned}$$

Q: Given the m.g.f of a R.V. X , how can we recover its probability distribution?

MTD 1: If the m.g.f. has the format

$$M_X(t) = p_1 e^{a_1 t} + \dots + p_n e^{a_n t} = \sum_{k=1}^n p_k e^{a_k t}$$

with $p_1 + p_2 + \dots + p_n = 1$, then X has p.m.f:

X	a_1	a_2	\dots	a_n
\mathbb{P}	p_1	p_2	\dots	p_n

$\Rightarrow X$ is now a discrete R.V.

e.g. $M_X(t) = \underbrace{\frac{1}{2}}_{\frac{1}{2}e^{0 \cdot t}} + \frac{1}{3}e^{-4t} + \frac{1}{6}e^{5t}$

then p.m.f of X :
$$\begin{cases} \mathbb{P}(X = -4) = 1/3 \\ \mathbb{P}(X = 5) = 1/6 \\ \mathbb{P}(X = 0) = 1/2 \end{cases}$$

MTD 2: Use the Distribution Sheet:

Example 1: $M_X(t) = e^{\lambda(e^t - 1)} \Rightarrow X \sim \text{Poi}(\lambda)$

Example 2: $M_Y(t) = \{(0.2) + (0.8)e^t\}^{10} \Rightarrow Y \sim \text{Bin}(10, 0.8)$

Example 3: $M_Z(t) = \exp\left\{\mu t + \frac{\sigma^2 t^2}{2}\right\} \Rightarrow Z \sim \mathcal{N}(\mu, \sigma^2)$

Extra exercise: Suppose that $M_X(t) = e^{2t + 3(e^{5t} - 1)}$,

derive the p.m.f. of X .

Soln: let Y follows the m.g.f $M_Y(t) = e^{3(e^t - 1)}$, $Y \sim \text{Poi}(3)$

Use $M_{aY+b} = \mathbb{E}[e^{(aY+b)t}] = e^{bt} \mathbb{E}[e^{atY}] = e^{bt} M_Y(at)$,

we have $X = aY + b$ with $a = 5$, $b = 2$

$$\mathbb{P}(Y = k) = \mathbb{P}(5Y + 2 = k) = \mathbb{P}\left(Y = \frac{k-2}{5}\right) = \frac{e^{-3} 3^{\frac{k-2}{5}}}{\left(\frac{k-2}{5}\right)!}$$

where $k = 5\ell + 2$ with $\ell = 0, 1, 2, \dots$