

1. A Brief Summary of Expectation

For a random variable X ,

$$\mathbb{E}[X] \stackrel{\text{def}}{=} \begin{cases} \sum_{k \in S_X} k \mathbb{P}(X=k) & X: \text{discrete} \\ \int_{-\infty}^{\infty} x \underbrace{f(x)}_{\text{P.D.F}} dx & X: \text{continuous} \end{cases}$$

$\mathbb{E}[X]$: the expected value X .

More generally, let $g: \mathbb{R} \rightarrow \mathbb{R}$, then $g(X)$ is also a R.V.

with

$$\mathbb{E}[g(X)] \stackrel{\text{def}}{=} \begin{cases} \sum_{k \in S_X} g(k) \mathbb{P}(X=k) & X: \text{discrete} \\ \int_{-\infty}^{\infty} g(x) f(x) dx & X: \text{continuous} \end{cases}$$

Moreover, $\text{Var}(X) \stackrel{\text{def}}{=} \mathbb{E}[(X - \mathbb{E}[X])^2]$

$$= \underbrace{\mathbb{E}[X^2]}_{\text{easier way to calculate}} - (\mathbb{E}[X])^2$$

2. Tutorial on Multivariate Calculus

Problem 2. Let (X, Y) be a random vector with joint probability density function given by

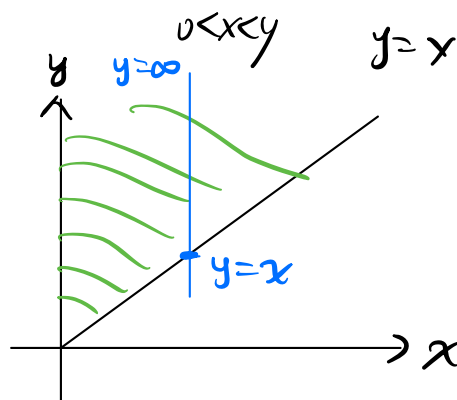
$$f_{X,Y}(x, y) = \begin{cases} ce^{-2(x+y)}, & 0 < x < y \\ 0, & \text{otherwise} \end{cases}$$

Find the value of c that makes $f_{X,Y}$ a valid joint probability density function.

Hint: $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx dy = 1.$

$$\begin{aligned} 1 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy dx \\ &= c \int_0^{\infty} \int_x^{\infty} e^{-2x} \cdot e^{-2y} dy dx \\ &= c \int_0^{\infty} e^{-2x} \cdot \left(-\frac{1}{2}\right) e^{-2y} \Big|_{y=x}^{y=\infty} dx \\ &= \frac{c}{2} \int_0^{\infty} e^{-4x} dx \\ &= \frac{c}{2} \cdot \left(-\frac{1}{4}\right) e^{-4x} \Big|_{x=0}^{x=\infty} \\ &= \frac{c}{2} \cdot \frac{1}{4} \\ &= \frac{c}{8} \end{aligned}$$

$$\text{So } c = 8$$



Problem 3. Let (X, Y, Z) be a random vector with joint probability density function given by

$$f_{X,Y,Z}(x, y, z) = \begin{cases} c, & 0 < x < y < z < 2 \\ 0, & \text{otherwise} \end{cases}$$

Find the value of c that makes $f_{X,Y,Z}$ a valid joint probability density function.

Hint: $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y,Z}(x, y, z) dx dy dz = 1.$

$$1 = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y,Z}(x, y, z) dx dy dz$$

$$\stackrel{(*)}{=} \int_0^2 \int_0^z \int_0^y c dx dy dz$$

To see (*):

① take the integration of x , which satisfies $0 < x < y$,
so the bound for x is 0 and y .

② take the integration of y , which satisfies $0 < y < z$

(since we've already taken the integration of x , so
the bound of y is independent of x)

③ take the integration of z , which satisfies $0 < z < 2$

so (*)

$$= \int_0^2 \int_0^z c y dy dz$$

$$= \frac{c}{2} \int_0^2 z^2 dz$$

$$= \frac{c}{2} \cdot \frac{8}{3} = \frac{4c}{3}$$

$$\text{so } c = \frac{3}{4}$$

3. Gamma Function

$$\text{Gamma Function } \Gamma(\alpha) \stackrel{\text{def}}{=} \int_0^{\infty} x^{\alpha-1} e^{-x} dx \quad \alpha > 0$$

Facts: ① $\Gamma(1) = \int_0^{\infty} e^{-x} dx = 1$
(Useful)

② $\Gamma(\alpha) = (\alpha-1)\Gamma(\alpha-1) \quad \alpha > 1$

(Proof using Integration by parts)

③ $\forall n \in \mathbb{N}, \Gamma(n) = (n-1)! \quad (\text{Mathematical Induction})$

④ $\alpha, \beta > 0 : \int_0^{\infty} x^{\alpha-1} e^{-\beta x} dx = \frac{\Gamma(\alpha)}{\beta^{\alpha}}$

HW5 P4 (b): Let $X \sim \text{exp}(1)$, $Y = X^{\beta}$, then calculate the n -th moment of Y , which is $\mathbb{E}[Y^n]$, when $\beta=3$?

Soln: $X \sim \text{exp}(1) \Rightarrow$ PDF $f_X(x) = \begin{cases} e^{-x} & x > 0 \\ 0 & \text{o.w.} \end{cases}$

$$\mathbb{E}[Y^n] = \mathbb{E}[X^{\beta n}]$$

$$= \int_0^{\infty} x^{\beta n} e^{-x} dx$$

$$= \int_0^{\infty} x^{(\beta n + 1) - 1} e^{-x} dx$$

$$= \Gamma(\beta n + 1)$$

when $\beta=3$, $3n+1 \in \mathbb{N}$, so by Facts ③: $\mathbb{E}[Y^n] = (3n)!$

□

4. Exponential Distribution

From Table of Distribution, we can see for

$$X \sim \exp(\lambda)$$

$$\text{PDF } f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$\begin{aligned} \text{CDF } F_X(x) &= \int_{-\infty}^x f_X(x) dx \\ &= \begin{cases} 1 - e^{-\lambda x} & x > 0 \\ 0 & x \leq 0 \end{cases} \end{aligned}$$

$$\text{Mean } \mathbb{E}[X] = \frac{1}{\lambda}$$

$$\text{Variance } \text{Var}(X) = \frac{1}{\lambda^2}$$

Q : How to apply what you learnt from prob. class

to verify $\int_0^{\infty} x e^{-\alpha x} dx = \frac{1}{\alpha^2}$ and $\int_0^{\infty} x^2 e^{-\alpha x} dx = \frac{2}{\alpha^3}$

without using integration by parts, for $\alpha > 0$?

$$\underline{A}: \int_0^{\infty} x e^{-\alpha x} dx$$

$$= \frac{1}{\alpha} \int_0^{\infty} x \alpha e^{-\alpha x} dx \quad (*)$$

Now, Define a Random variable $Y \sim \exp(\alpha)$, then
its density

$$f_Y(y) = \begin{cases} \alpha e^{-\alpha y} & y > 0 \\ 0 & \text{o.w.} \end{cases}$$

Distribution table

$$\text{and mean } \mathbb{E}[Y] = \int_0^{\infty} y \alpha e^{-\alpha y} dy = \frac{1}{\alpha}$$

So (X)

$$= \frac{1}{\alpha} \int_0^{\infty} x \alpha e^{-\alpha x} dx = \frac{1}{\alpha} \cdot \frac{1}{\alpha} = \frac{1}{\alpha^2}$$

$$\int_0^{\infty} x^2 e^{-\alpha x} dx$$

$$= \frac{1}{\alpha} \int_0^{\infty} x^2 \alpha e^{-\alpha x} dx$$

$$= \frac{1}{\alpha} \mathbb{E}[Y^2]$$

$$= \frac{1}{\alpha} (\mathbb{E}[Y]^2 + \text{Var}(Y))$$

$$= \frac{1}{\alpha} \left(\frac{1}{\alpha^2} + \frac{1}{\alpha^2} \right) = \frac{2}{\alpha^3}$$

We will see a lot of integration like this later in the future.