

Econometrics of Evaluation

Matching Methods

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You said matching ?

- You have data on treated and untreated individuals but treatment comes out of a choice
- **Endogenous selection**
- Very basic idea of matching : find a “match” for each treated individual in the untreated group
- **Good match ?** A “twin” with similar observable characteristics (apart from being treated)
- ⇒ Conditional on observables, treatment is *as-good-as-random*
- ⇒ Matched “twins” mimic the **unobserved counterfactual**
- Matching is a set of evaluation methods that rely on :
 - The assumption that **selection is only based on unobservables**
 - Conditionnal Independance Assumption (CIA) or “Treatment ignorability”
 - **Non-parametric** estimation
 - No assumptions on functional forms
- Much progress during the 1990's but very less since... Why ?
- Property (2) is attractive but property (1) much less so...
- ⇒ **What can we learn from matching methods ?**

Plan

Estimation under the CIA

Matching methods

Propensity score matching

Principle of matching

- Matching corrects (at least partly) for selection bias by **controlling for observable differences** between treated and control groups
- (Very strong) assumption that observable differences between treated and controls capture **all the determinants of selection**
- Conditional on observable variables X , treatment assignment is independent of potential outcomes (as-good-as random) :

$$(Y_0, Y_1) \perp T | X$$

⇒ **Conditional Independence Assumption (CIA)**

⇔ “Treatment ignorability”

⇔ “Uncounfoundedness assumption”

⇔ “Selection on observables”

What is the CIA ?

- Let's take 2 individuals with **similar observable characteristics**, but one is treated and the other is not.
 - The CIA assumption : if their characteristics are similar, being treated or not is not due to differences in potential outcomes.
 - The outcome of the untreated individual is a **good counterfactual** for the treated individual *in the absence of treatment* (and vice versa).
- ⇒ Comparing the “twins” provides an **unbiased estimator of the average treatment effect** (conditional on these observables).
- Note : For the ATT, un (little) less strong assumption is needed :

$$Y_0 \perp T | X$$

What do we need for matching to work ?

- CIA : there are no other characteristics than observables that influence both potential outcomes and treatment selection
→ **No selection on unobservables.**
- Matching methods require the existence of a **common support** (i.e. for all values of the observables there are both treated and untreated individuals) :

$$0 < P(T_i = 1|X_i) < 1$$

⇒ You need the “twin” to exist !

Regression as matching

- Let's assume that $E(Y_{i0}|X_i)$ is **linear** with respect to observable characteristics X_i :

$$E(Y_{i0}|X_i) = \alpha + \beta X_i$$

- Let's also assume **constant treatment effect** (Δ).
 The observed outcome writes :

$$Y_i = \alpha + \Delta T_i + \beta X_i + \epsilon_i$$

⇒ We can estimate treatment effect Δ with **OLS**, "**controlling**" for **observables**.

Note – *The assumption of a constant treatment effect can be relaxed by adding interaction terms between the treatment and the observables(observable heterogeneity).*

The pros and cons for linear regression

Pros

- ▶ Simplicity
- ▶ **Well-know theoretical foundations** both for estimation and for statistical inference (cf. simple assumptions for being BLUE)
- ▶ Even if the distribution of outcome is not exactly a linear function of the observables, linearity often provides a **(very) good approximation**.

Cons

- ▶ Simplicity can also be a limit...
- ▶ If the conditional distribution deviates too much from a linear function, even the best linear approximation may yield **biased estimates**.
- ▶ Particular issue when treatment and control groups have **very different observable characteristics** (no common support restriction)

The limits of regression

- Under CIA, the best linear estimate of the counterfactual outcome (\hat{Y}_0) is the mean outcome of the control group, **"corrected" for differences in observables** between the 2 groups :

$$\hat{E}(Y_{i0}|T_i = 1) = \bar{Y}_0 + (\bar{X}_1 - \bar{X}_0)\hat{\beta}$$

where \bar{X} and \bar{Y} denote empirical means of X and Y

- The estimator of the **ATT** is :

$$\begin{aligned}\hat{\Delta} &= E(Y_{i1}|T_i = 1) - \hat{E}(Y_{i0}|T_i = 1) \\ &= \bar{Y}_1 - \bar{Y}_0 - (\bar{X}_1 - \bar{X}_0)\hat{\beta}\end{aligned}$$

⇒ If the difference $\bar{X}_1 - \bar{X}_0$ is too large, so will be the correction (very sensitive to the specification).

Note – $ATT = ATE$ due to constant treatment effect.

Why is matching so different ?

- Matching methods get rid of the linearity assumption
- **Non-parametric estimation**
- You just need to find in the data the **best possible counterfactual “twin”** for treated units based on observed characteristics X
Then compare their outcomes.
 - **But how do we find this “twin”?**
- ⇒ Several matching methods for building this counterfactual !

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Nearest-neighbor matching

- The simplest method is to **find a “twin”** for each treated individual.
 - We then estimate the ATT by comparing the outcome Y_i of each treated individual with an untreated individual with exactly the same observable characteristics X_i .
 - However, it might be hard to find an exactly identical individual in the control group (especially if X_i are continuous variables)
→ **The “nearest-neighbor” is chosen.**
- ⇒ We need to define a **metric for the distance** between individuals

Metric for distance

Two main metrics are widely used :

- **Euclidean distance** : the distance between two individuals is the sum of the distance between all covariates.

$$d(l_i, l_j) = \sqrt{\sum_{k=1}^P (x_i^k - x_j^k)^2}$$

- **The Mahalanobis distance** is sometimes preferred, because it weights the distance by the variance-covariance matrix of the covariates X_i :

$$d(x_i, x_j) = (x_i - x_j)' \Sigma^{-1} (x_i - x_j)$$

⇒ You match each individual in the treated sample with his
“nearest-neighbor” in the untreated sample

Estimating causal effects

- **Out of matching :**

$$\hat{Y}_{i0} = \begin{cases} Y_{i(j)0} & \text{if } T_i = 1 \\ Y_{i0} & \text{if } T_i = 0 \end{cases}$$

$$\hat{Y}_{i1} = \begin{cases} Y_{i1} & \text{if } T_i = 1 \\ Y_{i(j)1} & \text{if } T_i = 0 \end{cases}$$

- **The average treatment effect on the treated** $E(Y_{i1} - Y_{i0} | T_i = 1)$ is estimated by the average of matched differences in the treated sample :

$$\widehat{ATT} = \frac{1}{N_1} \sum_{E_1} (Y_{i1} - \hat{Y}_{i0})$$

- **The average treatment effect** $E(Y_{i1} - Y_{i0})$ is estimated by the average of matched differences in the whole sample :

$$\widehat{ATE} = \frac{1}{N} \sum_i (\hat{Y}_{i1} - \hat{Y}_{i0})$$

where N is the total sample size

N_1 is the size of the treated sample E_1 ,

Y is the observed outcome of individual i

\hat{Y} is the outcome of i 's nearest-neighbor

Types of nearest-neighbor matching

- Matching can be done :
 - ▶ **Without replacement** : an individual in the control group can only be matched once with an individual in the treatment group
 - ▶ **With replacement** : the whole sample is used each time, which allows several matches with the same individual.
- But matching without replacement has some limitations :
 - ▶ It is **very demanding** in terms of data (large sample)
 - ▶ Estimates are **sensitive to the order** in matching

Limitations of nearest-neighbor matching

- Nearest-neighbor matching is one of the most widely used matching estimators.
→ Intuitive and does not require any parametric choice.
 - **Two limitations :**
 - ▶ No control over the **quality of the match** : the concept of nearest-neighbor is relative by nature.
 - ▶ Uses very **few information** : each individual's counterfactual uses only 1 observation (loses the information brought by other individuals)
- ⇒ **Variants** : estimate \hat{Y}_{i0} using more individuals from the control group
- ⇒ **Efficiency gains** if counterfactual for individual i is based on some average over several “similar” individual j
- ⇒ **Tradeoff between bias and variance**

M-closest matching

- Matching with a **fixed number M of nearest-neighbors**.
- The counterfactual outcome of individual i is the **average outcome** of his M nearest-neighbors.

$$\hat{Y}_{i0} = \begin{cases} \frac{1}{M} \sum_M Y_{i(j)0} & \text{if } T_i = 1 \\ Y_{i0} & \text{if } T_i = 0 \end{cases}$$

$$\hat{Y}_{i1} = \begin{cases} Y_{i1} & \text{if } T_i = 1 \\ \frac{1}{M} \sum_M Y_{i(j)1} & \text{if } T_i = 0 \end{cases}$$

- **Refinement** : exclude individual for whom you cannot find a “twin” (or M “twin”) within a given distance d (to be fixed).
- ⇒ **But how do we select the M -neighbors ?**

Radius or Caliper matching

- You select all individuals in the control group who are **located in a fixed neighborhood** of individual i , for a given **neighborhood radius** h , i.e. such that :

$$\|X_i - X_j\| < h$$

- Need to define a metric (Euclidian or Mahalanobis)

Kernel matching

- The counterfactual of individual i is computed with a **kernel estimation**.
- All individuals in the control group are used, but **weighted by their distance** from the treated group :

$$\hat{Y}_{i0} = \frac{\sum_{E_0}^k K\left(\frac{\|X_i - X_k\|}{h}\right) Y_k}{\sum_{E_0}^k K\left(\frac{\|X_i - X_k\|}{h}\right)}$$

where $K(.)$ is the kernel function used (most often the gaussian density)

i is a treated individual of the treatment group E_1

k is an untreated individual of the control group E_0 .

h is the window (*bandwidth*) of the kernel.

More on Kernel matching

- The bandwidth h gives the size of the neighborhood outside of which **weights are very small**.
- **The smaller the bandwidth**, the more likely the counterfactual will be estimated only for individuals in the control group whose observable characteristics are **very close**.
- **No set rule** for choosing the bandwidth (in practice, *ad hoc* choice or classic "rules of thumb".)
- Need to **check the sensitivity** of the results to various bandwidth h .

Note – *Have a look at the data to identify a “natural” threshold, but also test the sensitivity of the results to the chosen threshold, to finally choose the threshold that best arbitrates between precision and bias.*

How do we choose between matching methods?

- Each of matching methods has **pros and cons**.
 - The opposition between the simplest (nearest-neighbor) and the most complex (Kernel) method illustrates the standard **tradeoff between bias and precision**.
 - Nearest-neighbor matching does not use all available information and thus reduces precision.
 - Kernel function estimates are always more accurate but might generate mismatches, and thus bias.
- ⇒ **Check the sensitivity of your results to the method used !**

Can we be more simple ?

- For the CIA to hold, we want to use as much information as possible. You want to match on as many variables as possible
 - Difficult to find (a) close neighbor(s)
 - **High-dimensionality comparisons**
 - Imagine that you have to cut the population into boxes according to the whole set of observable characteristics you have, and in each of these boxes you have to find a treatment and its control...
 - It is shown that at finite distance, the estimators are **all the more biased** when the number of conditioning variables X is high (and even more when the conditioning variables are continuous).
- ⇒ Matching on the **propensity score**
- ⇒ **But what is the propensity score ?**

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A revised CIA

- An important property shown by Rosenbaum and Rubin (1983).
- If the CIA assumption holds for the X variables, then the potential outcomes are also independent of the treatment **conditional on any function of X**
- **The propensity score** is such a function : it gives the probability of being treated conditional on the X observable characteristics :

$$p(X) = P(T = 1|X)$$

- We can then **revise the CIA assumption** :

$$(Y_0, Y_1) \perp T | X \implies (Y_0, Y_1) \perp T | p(X)$$

Why is the propensity score useful ?

- This property helps to **reduce the dimensionality** of the comparisons.
- But the propensity score is unknown
→ **Need to be estimated**
- We often use **logit or probit specifications** to estimate the propensity score
- **Previous matching methods** (nearest neighbor matching, radius, or kernel) can all be applied on the estimated propensity score $\hat{p}(X)$ to measure the distance between two observations.

Estimating the propensity score

- To take into account the fact that the score is bounded between $[0, 1]$, we often use a logistic form (or a probit) :

$$\hat{p}(X) = \frac{\exp(f(X))}{1 + \exp(f(X))}$$

where $f(X)$ is a function of observable characteristics X . The simplest function has a linear form $f(X) = X\beta$

Note – *It is however recommended to use a polynomial approximation to get closer to the true distribution (see Hirano, Imbens & Ridder, 2003)*

Restriction to common support

- Remember that the CIA assumption is inherently based on the **existence of a common support**
→ We can find “twins”, i.e. individuals with similar values of observable characteristics

$$0 < P(T_i = 1|X_i) < 1$$

- If not, impossible to find comparable treated and controls.

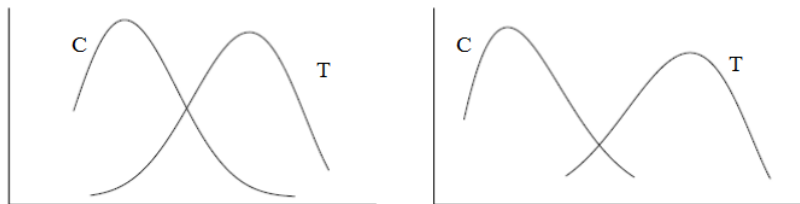
⇒ **What is a common support ?**

→ Area of $\widehat{p(X)}$ distribution over which this condition is verified
(you can find both treated and controls)

How do we find the common support ?

- Important to check that this area is **large enough**.
- **Graphical analysis** : you can plot the distribution of the score over the two subsamples (treated and controls)
 - **Histograms** of the estimated probability or **density functions** of being treated for both treated and controls
- **Check that the overlap is large** : for each value of the score, there must be a sufficient number of individuals in both subsamples (treatment and controls)
 - If no overlap for some values of the observables, incorrect to use these individuals for estimation.
 - If you do not restrict to the common support, estimates may be biased !

Graphical analysis



Distributions du score de propension

How do we restrict to the common support ?

- Several methods for restricting to the common support.
- **Warning !** It changes the scope the estimation : the impact is now estimated on part of the population (i.e. whose observables are such that an overlap is observed for the two subsamples)
→ **Local treatment effect estimator**

Min/max method

- ▶ For the *ATT*, you can drop individuals from the control group whose propensity score is below the minimum observed in the treatment group.
- ▶ For the *ATE*, also do the reverse : drop individuals from the treatment group whose propensity score is higher than the maximum observed in the control group.

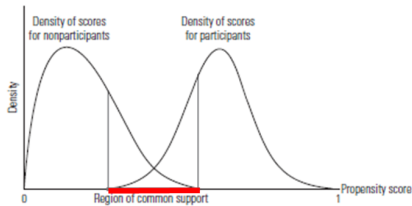
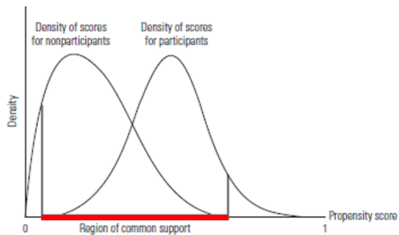
Trimming method

- ▶ Drop individuals whose propensity is too high or too low :

$$\alpha \leq p(X) \leq 1 - \alpha$$

- ▶ Imbens and Wooldridge (2008) suggest as a rule of thumb $\alpha = 0.01$ for the estimator to be efficient.

Graphical analysis



How do we choose the conditioning variables?

- The CIA assumption requires a **sufficient number of observable characteristics**
- The choice of these variables is crucial
 - They must have an impact on the variable of interest **AND** on the probability of being treated
- However, **no precise rule** for selecting "good" variables
 - ▶ Do not use variables that are measured **"after" treatment** (they may also be affected by the treatment)
 - **Endogeneity issues**
 - ▶ For the common support to be large enough, conditioning variables must not explain the probability of being treated **"too much"**

To sum up

Steps for implementing matching methods :

- 1 Identify the **control group**.
- 2 Select a set of **conditioning variables**.
- 3 Choose an **estimation method** (a linear specification can be used first).
- 4 For propensity score matching, estimate the **propensity score**.
- 5 Check that the **common support** is large enough (and that matching does reduce differences between treated and controls).
- 6 Estimate the **average treatment effects** of interest
- 7 Check the **sensitivity** of your results to alternative matching methods

What can we learn from matching ?

- The robustness of matching methods relies on the validity of the **Conditional Independence Assumption (CIA)**
 - This hypothesis is **very strong**.
 - It is **very demanding** in terms of data (too little information won't eliminate the selection bias)
 - **Standard errors** are hard to compute (bootstrap)
- ⇒ **Matching methods are often matched with other evaluation methods !**