

Generalization of risk ratio using observational data

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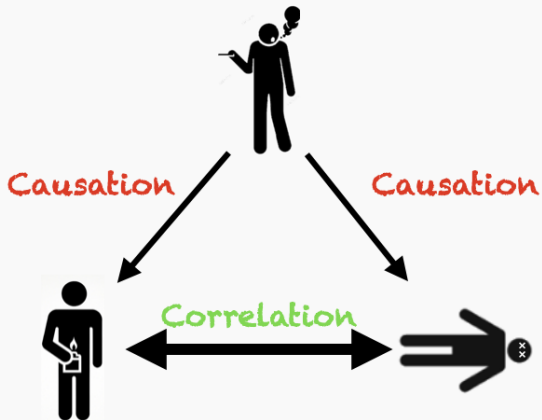
May 30, 2024

⇒ **Effect of a policy/intervention/treatment T on an outcome Y**

- What is the effect of smoking on COVID-19 mortality rate ?
- What is the effect of online classes on student performance ?
- How does 4 days work week affect the economy ?

Causal Inference

Effect of a policy/intervention/treatment T on an outcome Y



"People who have a lighter tend to have a smaller life expectancy"

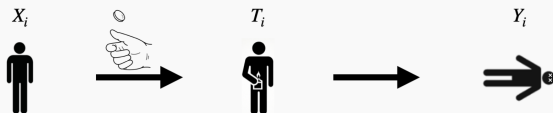
Causal Inference

"People who have a lighter tend to have a smaller life expectancy"



Causal Inference

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$$\Rightarrow N \text{ i.i.d. } (\underbrace{X_i}_{\text{covariates}}, \underbrace{T_i}_{\text{treatment}}, \underbrace{Y_i}_{\text{outcome}}) \in \mathbb{R}^d \times \{0, 1\} \times \mathbb{R}$$

$$Y = TY^{(1)} + (1 - T)Y^{(0)}.$$

Covariates			Treatment	Outcome	Potential outcomes	
X_1	X_2	X_3	T	Y	$Y^{(0)}$	$Y^{(1)}$
1.1	20	F	1	67	?	67
6	45	F	0	83	83	?
0	15	M	1	57	?	57
...
12	52	M	0	100	100	?

Ways to measure the causal effect

$$\mathbb{E} [Y^{(1)}] \quad \mathbb{E} [Y^{(0)}]$$

Expected outcome if treated (1) or control (0)

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Expected outcome if treated (1) or control (0)

Count the dead

$$\tau_{RR} = \frac{\mathbb{E}[Y^{(1)}]}{\mathbb{E}[Y^{(0)}]}$$

Count the living

$$\tau_{SR} = \frac{1 - \mathbb{E}[Y^{(1)}]}{1 - \mathbb{E}[Y^{(0)}]}$$

$$\tau_{RD} = \mathbb{E}[Y^{(1)}] - \mathbb{E}[Y^{(0)}]$$

Risk Difference

$$\tau_{NNT} = \tau_{RD}^{-1}$$

Number Needed to Treat

Odds Ratio

$$\tau_{OR} = \frac{\mathbb{E}[Y^{(1)}]}{1 - \mathbb{E}[Y^{(1)}]} \left(\frac{\mathbb{E}[Y^{(0)}]}{1 - \mathbb{E}[Y^{(0)}]} \right)^{-1}$$

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Expected outcome if treated (1) or control (0)

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<p>Risk Difference</p> $\tau_{RD} = \mathbb{E}[Y^{(1)}] - \mathbb{E}[Y^{(0)}]$	<p>Number Needed to Treat</p> $\tau_{NNT} = \tau_{RD}^{-1}$

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RESEARCH METHODS & REPORTING

CONSORT 2010 Explanation and Elaboration: updated guidelines for reporting parallel group randomised trials

David Moher,¹ Sally Hopewell,² Kenneth F Schulz,³ Victor Montori,⁴ Peter C Gøtzsche,⁵ P J Devereaux,⁶ Diana Elbourne,⁷ Matthias Egger,⁸ Douglas G Altman⁹

"[...] both the relative effect (risk ratio (relative risk) or odds ratio) and the absolute effect (risk difference) should be reported (with confidence intervals), as neither the relative measure nor the absolute measure alone gives a complete picture of the effect and its implications."

Randomized Controlled Trial (RCT)

$$(1) \ T_i \perp\!\!\!\perp \{Y_i(0), Y_i(1), X_i\} \implies e = \mathbb{P}[T_i = 1]$$

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$$\stackrel{(1)}{=} \mathbb{E}[Y|T=1]/\mathbb{E}[Y|T=0]$$

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Simple mean approach:


$$\hat{\tau}_{RR,N,n} = \frac{\frac{1}{n_1} \sum_{i=1}^n T_i Y_i}{\frac{1}{n-n_1} \sum_{i=1}^n (1-T_i) Y_i} \quad \text{with} \quad n_1 = \sum_{i=1}^n T_i$$

$$\sqrt{n}(\hat{\tau}_{RR,N,n} - \tau_{RR}) \xrightarrow{d} \mathcal{N}(0, V_{RR,N})$$

where

$$V_{RR,N} = \tau_{RR}^2 \left(\frac{\text{Var}(Y^{(1)})}{e\mathbb{E}[Y^{(1)}]^2} + \frac{\text{Var}(Y^{(0)})}{(1-e)\mathbb{E}[Y^{(0)}]^2} \right)$$


Randomized Controlled Trial (RCT)

- **Gold standard** (allocation )
- $T_i \perp\!\!\!\perp \{Y_i(0), Y_i(1), X_i\}$ (1)
- Small sample size: expensive, long, ethical limitations
- Not representative
sample: restrictive inclusion criteria

Observational Data

- Low quality Data: contains **confounding bias**
- $\{Y_i(0), Y_i(1)\} \perp\!\!\!\perp T_i \mid X_i$ (2)
- Large sample size (registries, biobanks, EHR, claims)
- Representative sample

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$$\begin{aligned}\tau_{RR} &= \mathbb{E}[Y(1)]/\mathbb{E}[Y(0)] \\ &\stackrel{(2)}{=} \mathbb{E}[\mathbb{E}[Y|X, T=1]]/\mathbb{E}[\mathbb{E}[Y|X, T=0]]\end{aligned}$$

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- Inverse weighted approach

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- Regression based approach

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- Double robust based approach

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$$\hat{\tau}_{\text{RR,IPW}} = \frac{\sum_{i=1}^n \frac{T_i Y_i}{\hat{e}(X_i)}}{\sum_{i=1}^n \frac{(1-T_i) Y_i}{1-\hat{e}(X_i)}}$$

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$$\sqrt{n} (\tau_{\text{RR,IPW}}^{\star} - \tau_{\text{RR}}) \xrightarrow{d} \mathcal{N}(0, V_{\text{RR,IPW}})$$

where

$$V_{\text{RR,IPW}} = \tau_{\text{RR}}^2 \left(\frac{\mathbb{E} \left[\frac{(Y^{(1)})^2}{e(X)} \right]}{\mathbb{E} [Y^{(1)}]^2} + \frac{\mathbb{E} \left[\frac{(Y^{(0)})^2}{1-e(X)} \right]}{\mathbb{E} [Y^{(0)}]^2} \right).$$

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$$\sqrt{n} (\tau_{\text{RR,IPW}}^* - \tau_{\text{RR}}) \xrightarrow{d} \mathcal{N}(0, V_{\text{RR,IPW}})$$

where

$$V_{\text{RR,IPW}} = \tau_{\text{RR}}^2 \left(\frac{\mathbb{E} \left[\frac{(Y^{(1)})^2}{e(X)} \right]}{\mathbb{E} [Y^{(1)}]^2} + \frac{\mathbb{E} \left[\frac{(Y^{(0)})^2}{1-e(X)} \right]}{\mathbb{E} [Y^{(0)}]^2} \right).$$

Furthermore, if for all i , $0 < m \leq Y_i \leq M$ we have:

$$|Bias(\tau_{\text{RR,IPW}}^*)| \leq \frac{2M^3(1-\eta)^3}{nm^3\eta^3} \quad |Var(\tau_{\text{RR,IPW}}^*)| \leq \frac{4M^4(1-\eta)^6}{nm^6\eta^4}$$

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Regression based approach: We estimate $\hat{\mu}_t(X) = \mathbb{E}[Y|T = t, X]$

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$$\sqrt{n}(\tau_{\text{RR}, G, n}^* - \tau_{\text{RR}}) \xrightarrow{d} \mathcal{N}(0, V_{\text{RR}, G})$$

where

$$V_{\text{RR}, G} = \tau_{\text{RR}}^2 \text{Var} \left(\frac{\mu_1^*(X)}{\mathbb{E}[Y(1)]} - \frac{\mu_0^*(X)}{\mathbb{E}[Y(0)]} \right)$$

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Assume for all i , $M_0 \geq \mu_0^*(x_i) \geq m_0 > 0$ and $M_1 \geq \mu_1^*(x_i)$. Then, we have:

$$|\text{Bias}(\hat{\tau}_{\text{RR}, G, n})| \leq \frac{2M_1M_0^2}{nm_0^3} \quad |\text{Var}(\hat{\tau}_{\text{RR}, G, n})| \leq \frac{2M_0^2M_1(M_1 + M_0)}{nm_0^6}$$

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Double robust based approach: We estimate $\hat{\mu}_t(X) = \mathbb{E}[Y|T = t, X]$ and $\hat{e}(X) = \mathbb{P}[T = 1|X]$:

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If $\mathbb{E}[(\hat{\mu}_{(t)}(X_i) - \mu_{(t)}(X_i))^2] \mathbb{E}[(\hat{e}(X_i) - e(X_i))^2] = o(\frac{1}{n})$ then:

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$$\sqrt{n}(\hat{\tau}_{\text{RR,AIPW}} - \tau_{\text{RR}}) \xrightarrow{d} \mathcal{N}(0, V_{\text{RR,AIPW}})$$

where

$$\frac{V_{\text{RR,AIPW}}}{\tau_{\text{RR}}^2} = \text{Var}\left(\frac{\mu_1(X)}{\mathbb{E}[Y^{(1)}]} - \frac{\mu_0(X)}{\mathbb{E}[Y^{(0)}]}\right) + \mathbb{E}\left[\frac{\text{Var}(Y^{(1)}|X)}{e(X)\mathbb{E}[Y^{(1)}]^2}\right] + \mathbb{E}\left[\frac{\text{Var}(Y^{(0)}|X)}{(1-e(X))\mathbb{E}[Y^{(0)}]^2}\right].$$

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$$\hat{\tau}_{\text{RR, OS}} = \frac{\sum_{i=1}^n \hat{\mu}_1(X_i)}{\sum_{i=1}^n \hat{\mu}_0(X_i)} \left(1 - \frac{\sum_{i=1}^n \hat{\mu}_1(X_i) + \frac{T_i(Y_i - \hat{\mu}_1(X_i))}{\hat{e}(X_i)}}{\sum_{i=1}^n \hat{\mu}_0(X_i)} \right) + \frac{\sum_{i=1}^n \hat{\mu}_0(X_i) + \frac{(1 - T_i)(Y_i - \hat{\mu}_0(X_i))}{\hat{e}(X_i)}}{\sum_{i=1}^n \hat{\mu}_0(X_i)}$$

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If $\mathbb{E}\left[(\hat{\mu}_{(t)}(X_i) - \mu_{(t)}(X_i))^2\right] \mathbb{E}[(\hat{e}(X_i) - e(X_i))^2] = o\left(\frac{1}{n}\right)$ and [...] then:

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$$\sqrt{n}(\hat{\tau}_{\text{RR,OS}} - \tau_{\text{RR}}) \xrightarrow{d} \mathcal{N}(0, V_{\text{RR,AIPW}})$$

where

$$\frac{V_{\text{RR,AIPW}}}{\tau_{\text{RR}}^2} = \text{Var}\left(\frac{\mu_1(X)}{\mathbb{E}[Y^{(1)}]} - \frac{\mu_0(X)}{\mathbb{E}[Y^{(0)}]}\right) + \mathbb{E}\left[\frac{\text{Var}(Y^{(1)}|X)}{e(X)\mathbb{E}[Y^{(1)}]^2}\right] + \mathbb{E}\left[\frac{\text{Var}(Y^{(0)}|X)}{(1-e(X))\mathbb{E}[Y^{(0)}]^2}\right].$$

Observational Data

Efficiency Theory:

Suppose we want to estimate a function f in a value a :

$$a \approx \hat{a} \implies f(a) \approx f(\hat{a}) \quad f(a) = f(\hat{a}) + \int_{\hat{a}}^a f'(t) dt$$

Observational Data

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$$f(a) \implies \Psi(\mathcal{P}_{X,T,Y}) = \frac{\mathbb{E}[\mathbb{E}[Y|T=1, X]]}{\mathbb{E}[\mathbb{E}[Y|T=0, X]]}$$

$$\mathcal{P}_{X,T,Y} \approx \hat{\mathcal{P}}_{X,T,Y} \implies \Psi(\mathcal{P}) \approx \Psi(\hat{\mathcal{P}})$$

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$$\mathcal{P}_{X,T,Y} \approx \hat{\mathcal{P}}_{X,T,Y} \implies \Psi(\mathcal{P}) \approx \Psi(\hat{\mathcal{P}})$$

$$\begin{aligned} \Psi(\mathcal{P}) &= \Psi(\hat{\mathcal{P}}) + \int \varphi(z; \hat{\mathcal{P}}) d\mathcal{P}(z) + R_2(\hat{\mathcal{P}}, \mathcal{P}) \\ &\simeq \Psi(\mathbb{P}_n) + \frac{1}{n} \sum_{i=1}^n \varphi(Z_i; \mathbb{P}_n) := \hat{\tau}_{\text{RR,OS}} \end{aligned}$$

φ can be seen as the derivative of the Ψ according to a distribution \mathcal{P} .

Observational Data

We generate datasets $(X_i, T_i, Y_i^{(0)}, Y_i^{(1)})$ according to the model:

$$\begin{aligned} Y_i^{(1)} &= m(X_i) + b(X_i) + \varepsilon_i^{(1)} & e(X_i) &= \mathbb{P}[T_i = 1|X_i] \\ Y_i^{(0)} &= b(X_i) + \varepsilon_i^{(0)} & \varepsilon^{(t)} &\sim \mathcal{N}(0, \sigma^2) \end{aligned}$$

$$\hat{\tau}_{\text{RR,IPW}} \quad \hat{\tau}_{\text{RR,G, n}} \quad \hat{\tau}_{\text{RR,AIPW}} \quad \hat{\tau}_{\text{RR,OS}}$$

We need to estimate $\hat{\mu}_t(X) = \mathbb{E}[Y|T = t, X]$ and $\hat{e}(X) = \mathbb{P}[T = 1|X]$:

Random Forest:

- `RandomForestClassifier()`
- `RandomForestRegressor()`

Regressions:

- `LogisticRegression()`
- `LinearRegression()`

Simulations

Linear and logistic setting:

$$m(X, V) = 2,$$

$$b(X, V) = \beta_0^\top [X, V], \quad \text{where } \beta_0 = (-1, 1, -1, -1, 1, 1)$$

$$e(X) = (1 + \exp(-\beta_e X))^{-1}, \quad \text{where } \beta_e = (-0.6, 0.6, -0.6).$$

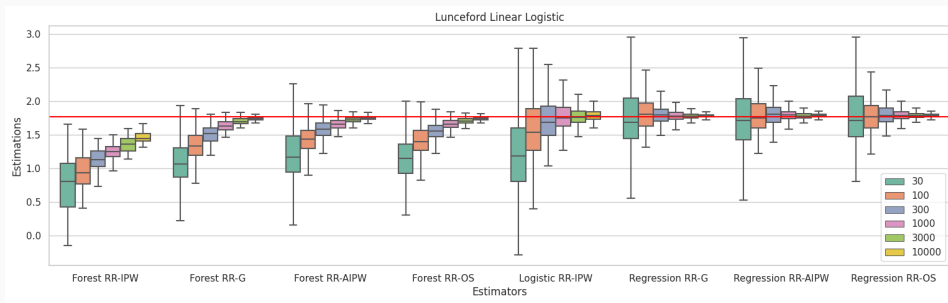


Figure 1: Estimation of the Risk Ratio in a linear and logistic setting

Simulations

Non-linear and non-logistic setting:

$$m(X) = \sin(\pi X_1 X_2) + 2(X_3 - 0.5)^2 + X_4 + 0.5X_5 - (X_1 + X_2)/4$$

$$b(X) = (X_1 + X_2)/2$$

$$e(X) = \max\{0.1, \min(\sin(\pi X_1), 0.9)\}, \quad \text{where } X \sim \text{Unif}(0, 1)^6.$$

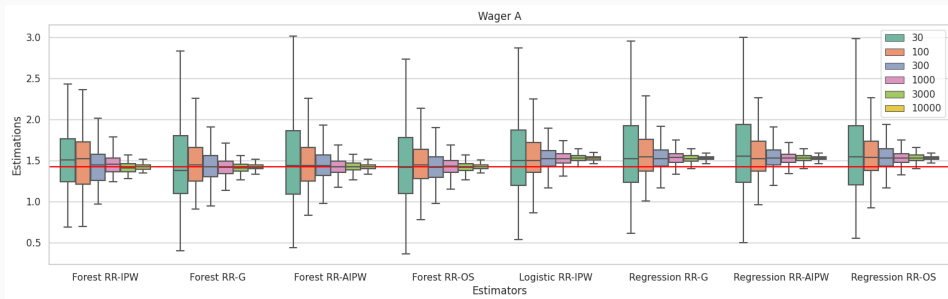


Figure 2: Estimation of the Risk Ratio in a Non-linear and non-logistic setting

Generalization of risk ratio using observational data

Generalization of risk ratio using observational data

- RCT

$$\tau_{RR,N} \quad \tau_{RR,HT}$$

Generalization of risk ratio using observational data

- **RCT**

$$\tau_{RR,N} \quad \tau_{RR,HT}$$

- **Observational Data**

$$\hat{\tau}_{RR,IPW} \quad \hat{\tau}_{RR,G, n} \quad \hat{\tau}_{RR,AIPW} \quad \hat{\tau}_{RR,OS}$$

Conclusion

Generalization of risk ratio using observational data

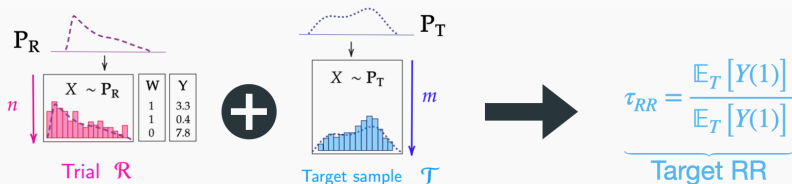
- **RCT**

$$\tau_{RR,N} \quad \tau_{RR,HT}$$

- **Observational Data**

$$\hat{\tau}_{RR,IPW} \quad \hat{\tau}_{RR,G, n} \quad \hat{\tau}_{RR,AIPW} \quad \hat{\tau}_{RR,OS}$$

Generalization:



Thanks for your attention!

Appendix

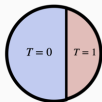
Causal Inference

$$Y(t) \text{ Vs } Y|T = t$$

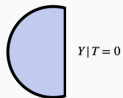
Population



Subpopulations



Conditioning



Intervening

