







# Causal alternatives to metaanalysis

#### Clément Berenfeld (Premedical)

Joint work with A. Boughdiri<sup>1</sup>, B. Colnet, W. van Amsterdam<sup>2</sup>, A. Bellet<sup>1</sup>, R. Kellaf<sup>1</sup>, E. Scornet<sup>3</sup> and J. Josse<sup>1</sup>

<sup>1</sup>Inria-Inserm Premedical <sup>2</sup>UMC Utrecht <sup>3</sup>U. Sorbonne

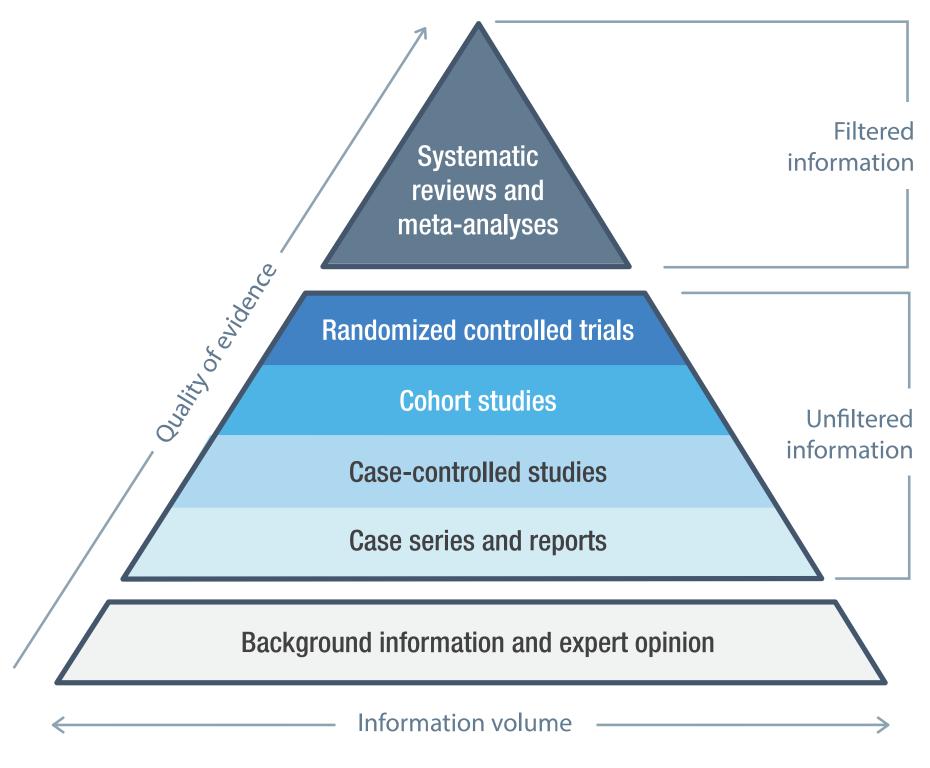
**CIRC Seminar - 13.10.2025** 

#### What is a meta-analysis?

A meta-analysis is a statistical method that combines the results from multiple studies to derive a more precise and reliable overall estimate

→ stands at the top of the pyramid of evidence in clinical research

→ used by regulatory bodies to formulate recommendations or fix drug prices



source: https://openmd.com/guide/levels-of-evidence

#### What is a meta-analysis?

Let us consider K different RCTs measuring the effect of the same binary treatment  $A \in \{0,1\}$  and the same binary outcome  $Y \in \{0,1\}$ .

The results are publicly available in a contingency table of the form

	Y = 1	Y = 0
A = 1	$n_{11}(k)$	$n_{10}(k)$
A = 0	$n_{01}(k)$	$n_{00}(k)$

#### What is a meta-analysis?

From these tables, one can compute a desired treatment effect  $\hat{\theta}_k$  (e.g. risk difference, risk ratio, odds ratio, etc), along with a standard error  $\hat{\sigma}_k$ .

Ex: for the log-risk ratio

$$\hat{\theta}_k = \log \frac{n_{11}(k)/n_1(k)}{n_{01}(k)/n_0(k)} \qquad \hat{\sigma}_k^2 = \frac{n_{10}(k)}{n_{11}(k)n_1(k)} + \frac{n_{00}(k)}{n_{01}(k)n_0(k)}$$

where 
$$n_a(k) := n_{a0}(k) + n_{a1}(k)$$

#### What is a meta-analysis?

#### Fixed-effect model: Gaussian model

$$\hat{\theta}_k \sim \mathcal{N}(\theta^*, \sigma_k^2)$$

- → no heterogeneity between studies
- → maximum likelihood estimator is given by

$$\hat{\theta} = \sum_{k=1}^{K} \omega_k \hat{\theta}_k$$
 where  $\omega_k \propto \frac{1}{\hat{\sigma}_k^2}$ 

 $\rightarrow$  **final variance** estimator is given by  $\hat{\sigma}^2 = \left(\sum_{k=1}^K \frac{1}{\hat{\sigma}_k^2}\right)^{-1}$ 

#### What is a meta-analysis?

Random-effects model: Hierarchical model

$$\hat{\theta}_k \mid \theta_k \sim \mathcal{N}(\theta_k, \sigma_k^2)$$

$$\theta_k \sim \mathcal{N}(\theta^*, \tau^2)$$

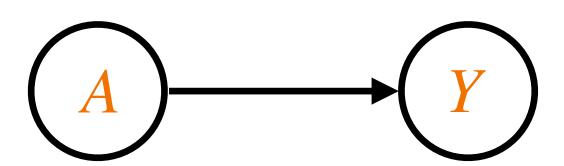
- → heterogeneity between studies
- $\rightarrow$  many **different methods** to estimate  $\tau$  (e.g, DerSimonian and Laird, Paule-Mandel, etc)
- → final estimator is given by

$$\hat{\theta} = \sum_{k=1}^K \omega_k \hat{\theta}_k \qquad \text{where} \qquad \omega_k \propto \frac{1}{\hat{\sigma}_k^2 + \hat{\tau}^2} \qquad \text{and} \qquad \hat{\sigma}^2 = \left(\sum_{k=1}^K \frac{1}{\hat{\sigma}_k^2 + \hat{\tau}^2}\right)^{-1}$$

#### What is causal inference?

Causal inference pertains to the process of understanding the relationships between a cause and its effects.

Ex: What is the effect of a given treatment on a given outcome?



Counterfactual variables: Y(0) and Y(1) are the outcome if the patient has, possibly contrary to the fact, taken treatment A=0 or A=1.

#### What is causal inference?

A causal effect is a measure of how Y(1) and Y(0) differ in a given population of interest

Ex: the risk difference among

- $\rightarrow$  the study population (ATE)  $\theta = \mathbb{E}[Y(1)] \mathbb{E}[Y(0)]$
- $\rightarrow$  the treated population (ATT)  $\theta = \mathbb{E}[Y(1)|A=1] \mathbb{E}[Y(0)|A=1]$
- $\rightarrow$  the control population (ATC)  $\theta = \mathbb{E}[Y(1)|A=0] \mathbb{E}[Y(0)|A=0]$

In a RCT, thanks to randomization, all these quantities coincide.

#### The big question

Do usual meta-analysis methods target an estimand which is a causal effect, in the sense that it pertains to the effect of the treatment in a specific population?

→ would have big implications in term of interpretability of meta-analysis results!

### 1. Causal meta-analysis with aggregated data (AD)

### 2. Causal meta-analysis with individual data (ID)

B. C., Boughdiri, A., Colnet, B., van Amsterdam, W. A., Bellet, A., Khellaf, R., Scornet, E., & Josse, J. (2025). Causal meta-analysis: rethinking the foundations of evidence-based medicine. arXiv preprint arXiv:2505.20168.

Boughdiri, A., B. C., Josse, J., & Scornet, E. (2025). A unified framework for the transportability of population-level causal measures. *NeuRIPS 2025*.

**Notations:** A patient's data is typically of the form (A, Y, X, H) where

- A is the treatment variable
- Y = AY(1) + (1 A)Y(0) is the individual outcome
- $X \in \mathcal{X}$  is the patient covariate
- $H \in [K]$  is the study membership

In the aggregated data setting, we have access to the aggregated values

$$n_{av}(k) := \#\{i \mid H_i = k, A_i = a, Y_i = y\}, \quad a \in \{0,1\}, y \in \{0,1\}$$

for all studies  $k \in [K]$ .

We can also have access to summary information about the covariate distribution in each study:

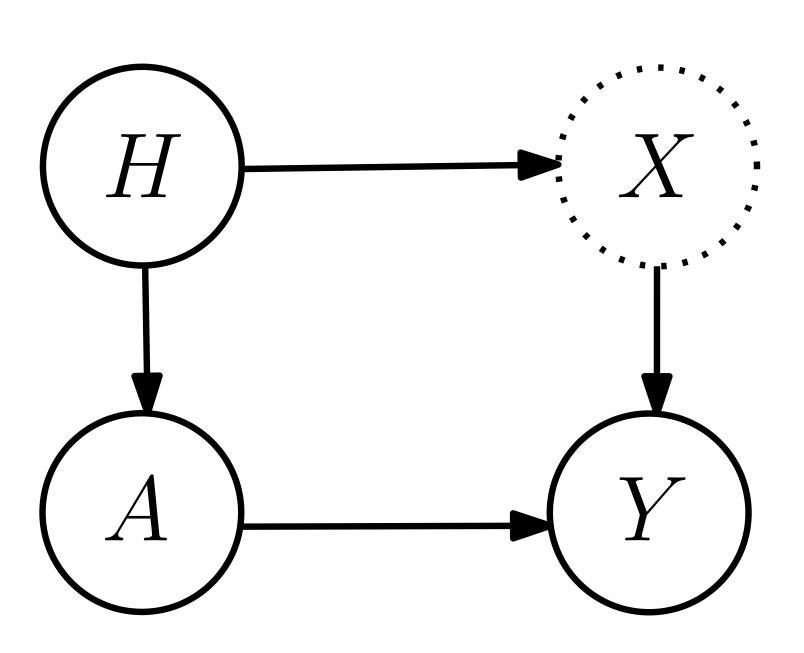
$$S_k := \mathcal{S}_k(\hat{P}_k)$$

where  $\hat{P}_k$  is the **empirical distribution** of X in study K, and  $\mathcal{S}_k: \mathcal{P}(\mathcal{X}) \to \mathbb{R}^{d_k}$  is a **summary map** (e.g., means, standard deviations, quantiles, etc)

#### We assume that

- each study is a RCT (no arrow from X to A)
- there is no center effect (no arrow from H to Y)

$$\forall k, \ell \in [K],$$
 
$$\mathbb{E}[Y(a) \mid X, H = k] = \mathbb{E}[Y(a) \mid X, H = \ell]$$
 
$$=: \mu_a(X)$$



We also assume that  $\theta_k$  only depends on  $\mathbb{E}[Y(a) | H = k]$  through

$$\theta_k = \phi(\mathbb{E}[Y(1) | H = k], \mathbb{E}[Y(0) | H = k])$$

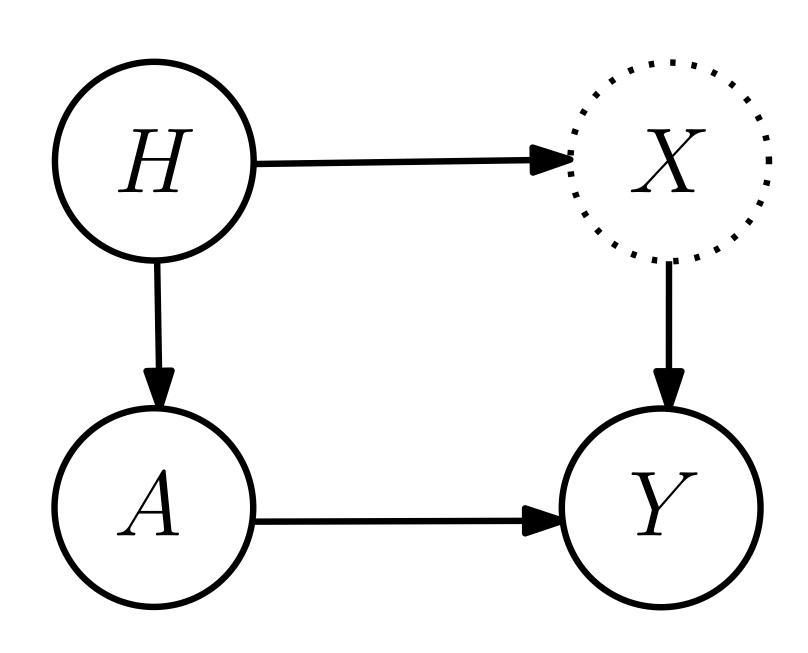
#### Ex:

- risk difference:  $\phi(a,b) = a - b$ 

- risk ratio:  $\phi(a,b) = a/b$ 

odds ratio:  $\phi(a,b) = \frac{a}{1-a} \frac{1-b}{b}$ 

- etc



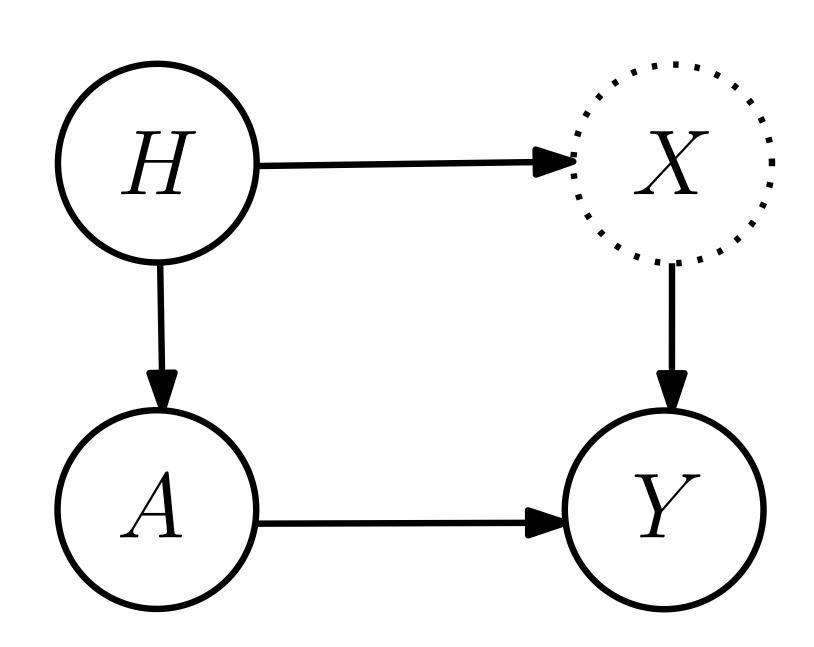
Under these assumptions,  $\theta_k$  only depends on  $P_k = \mathcal{L}(X \mid H = k)$  through

$$\mathbb{E}[Y(a) | H = k] = \mathbb{E}[\mathbb{E}[Y(a) | X, H = k] | H = k] = \int \mu_a(x) dP_k(x)$$

Defining

$$\theta(P) := \phi\left(\int \mu_1(x) dP(x), \int \mu_0(x) dP(x)\right),$$

we find that  $\theta_k = \theta(P_k)$ 



Given a meta-analysis aggregate  $\hat{\theta}$ , we denote by  $\theta_{\infty}$  the values towards which it converges as  $n \to \infty$  (if it exists)

**Ex**: For the random effect model,  $\hat{\theta}_k \to \theta_k$ ,  $\hat{\sigma}_k \to 0$  and  $\hat{\tau} \to \tau$  so that, when  $\tau \neq 0$ , it holds that  $\theta_\infty = \frac{1}{K} \sum^K \theta_k$ 

A meta-analysis effect  $\hat{\theta}$  is *causal* if, for all covariate distributions  $P_1,\ldots,P_K$ , there exists  $P^*$ , independent from  $\mu_1$  and  $\mu_0$ , such that  $\theta_\infty=\theta(P^*)$ 

 $\rightarrow$  we say that  $P^*$  is the target population

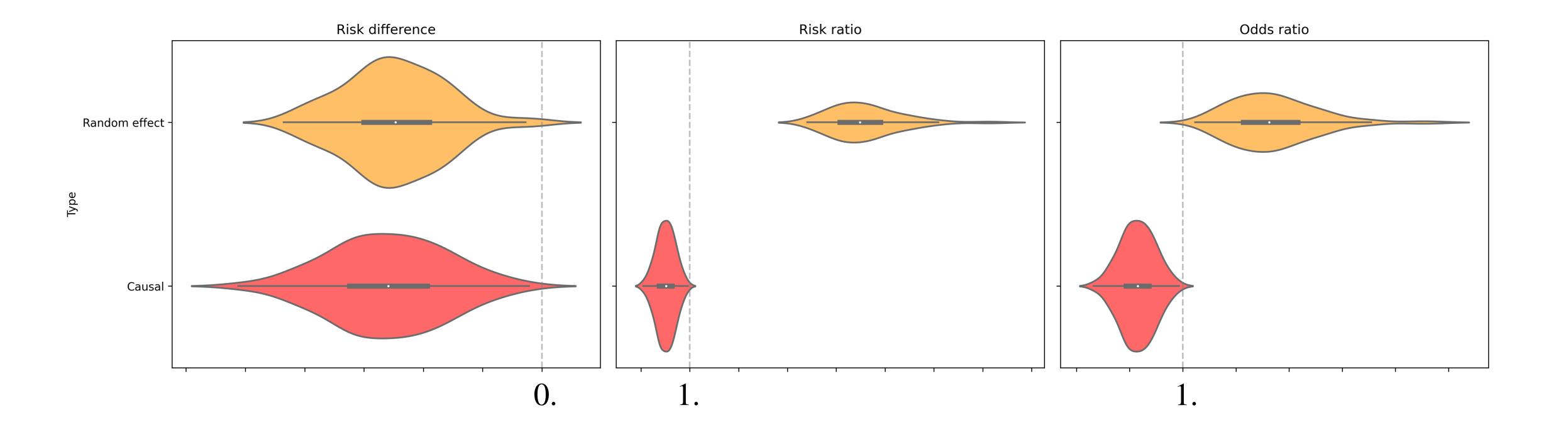
#### Theorem.

- 1. If the link function  $\phi$  is non-linear, then both the random-effects and the fixed-effect estimator are not causal
- 2. If the link function  $\phi$  is linear then the random-effects estimator is causal
- 3. If the link function  $\phi$  is linear, and if the ratios  $\hat{\sigma}_k^2/\hat{\sigma}_{k'}^2$  all converges towards a value in  $[0,\infty]$ , then the fixed-effects estimator is **causal**

#### random effects on risk ratios → not causal

#### Ex:

- random effects on risk differences → causal



A violin plot

#### So how do we construct causal meta-analysis estimands?

- $\rightarrow$  First, define a target population  $P^*$
- ightarrow Try to realize  $P^*$  as a convex combination of  $P_1, \ldots, P_K$

$$P^* pprox \sum_{k=1}^K \hat{\alpha}_k P_k$$
 where  $\sum_{k=1}^K \hat{\alpha}_k = 1$ 

 $\rightarrow$  Estimate  $\theta(P^*)$  with

$$\hat{\theta} = \phi \left( \sum_{k=1}^{K} \hat{\alpha}_k \frac{n_{11}(k)}{n_1(k)}, \sum_{k=1}^{K} \hat{\alpha}_k \frac{n_{01}(k)}{n_0(k)} \right)$$

Depending on the choice of  $P^*$ , and on the summary informations on the  $P_k$ 's, the computation of  $\hat{\alpha}_k$  can range to very simple to very complicated Covariate-free targets:

- Pooled target: 
$$P^* = \sum_{k=1}^K \mathbb{P}(H=k)P_k$$
 and  $\hat{\alpha}_k = n_k/n$ 

o  $\theta^*$  corresponds to the ATE on the population of all studies pooled together

- Uniform target: 
$$P^* = \sum_{k=1}^K \frac{1}{K} P_k$$
 and  $\hat{\alpha}_k = 1/K$ 

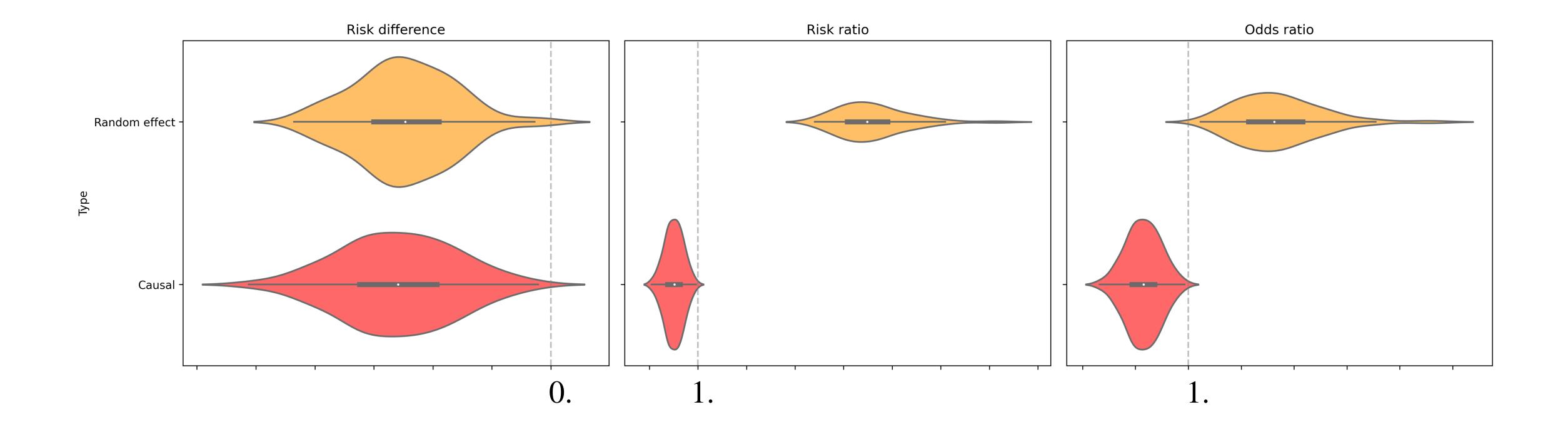
Are these effects really different from classical approaches, e.g. random-effects model? Yes

In the large scale limit  $n \to \infty$ , the random effects estimate converge to, in the case of the **risk ratio** 

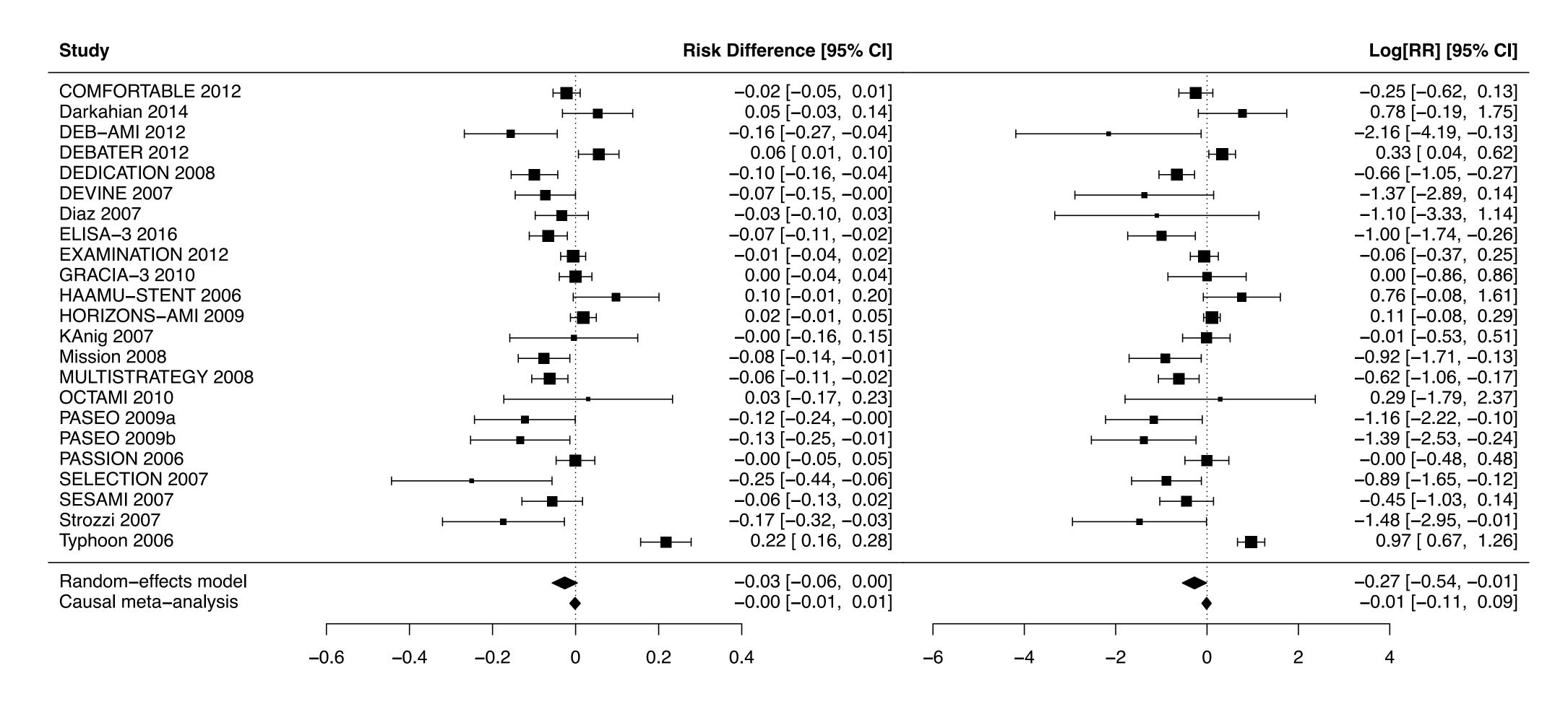
$$\frac{\prod_{k=1}^{K} P_k(\mu_1)^{1/K}}{\prod_{k=1}^{K} P_k(\mu_0)^{1/K}} \neq RR(P^*)$$

while a covariate-free causal approach will yield

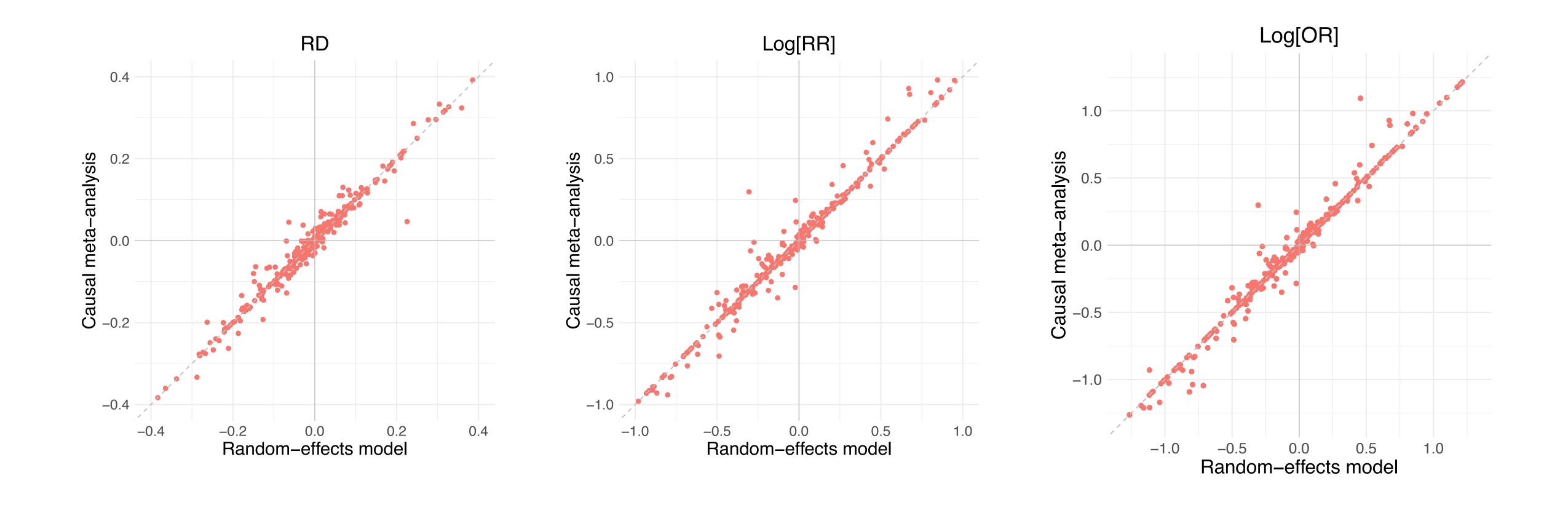
$$\frac{\sum_{k=1}^{K} \alpha_k P_k(\mu_1)}{\sum_{k=1}^{K} \alpha_k P_k(\mu_0)} = \frac{P^*(\mu_1)}{P^*(\mu_0)} = RR(P^*)$$



The same violin plot



Reanalysis of Feinberg et al, *Drug-eluting stents versus bare-metal stents for acute coronary syndrome*. Cochrane Database of Systematic Reviews, (8), 2017



Comparison of the random effect models and the pooled meta-analysis for 597 meta-analyses from the Cochrane Library

Covariate-dependent target: covariate information through  $S_k = S_k(\hat{P}_k)$ 

- Linear summaries: if all  $\mathcal{S}_k$  are identical and linear, one can solve

$$\hat{\alpha} \in \operatorname{argmin} ||S^* - S\alpha|| + \Omega(\alpha)$$

where 
$$S = (S_1, ..., S_K)$$
 and  $S^* = S(P^*)$ .

- Parametric proxies: one can solve  $\mathcal{S}_k(Q_k) \approx S_k$  under a parametric model  $Q_k \in \Sigma_k$  and adjust

$$\hat{\alpha} \in \operatorname{argmin} \operatorname{dist} \left( P^*, \sum_{k=1}^K \alpha_k Q_k \right) + \Omega(\alpha)$$

(a.k.a. generalization)

Imagine now that we have access to all the individual data

$$(X_i, A_i, Y_i, H_i)$$
 for  $i \in [n]$ 

Given a target population  $P^*$ , we wish to estimate  $\theta(P^*)$ 

 $\to$  having access to all the covariates  $\{X_i\}_{i\in[n]}$  allows to adjust much more precisely to the covariate distribution

#### A slight change of setting:

- We consider a single source study (rather than  $\it K$ ) for which  $\it H=1$
- We have a sample from the target population, denoted by H=0
- Data is collapsed to the form

$$(X_i, H_iA_i, H_iY_i, H_i), i \in [N]$$

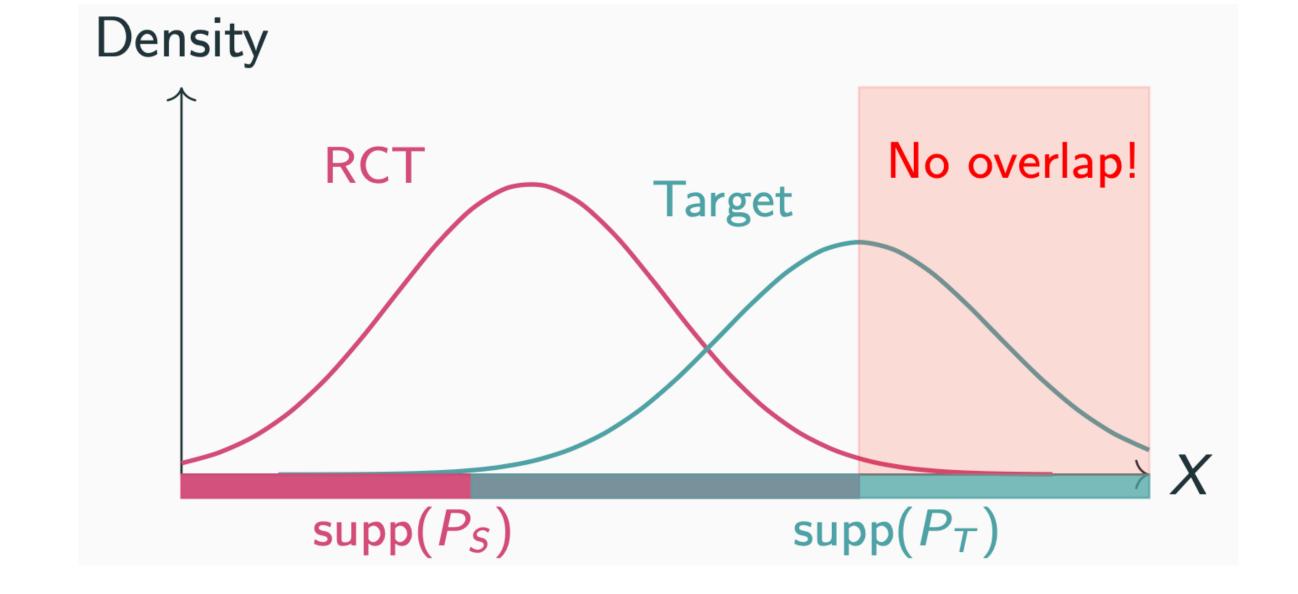
- We let  $n = \#\{i, H_i = 1\}$  and  $m = \#\{i, H_i = 0\}$ .

We let T (resp. S) denotes distribution conditional to H=0 (resp. H=1)

#### Overlap assumption:

$$supp P_T \subset supp P_S$$

#### **Exchangeability in mean:**



$$\mathbb{E}_{\mathsf{T}}[Y(a) \mid X] = \mathbb{E}_{\mathsf{S}}[Y(a) \mid X] \qquad (= \mu_a(X))$$

Goal: estimate the effect in the target population  $\theta_{\rm T} = \theta(P_{\rm T})$ 

Identifiability formulae: letting 
$$r(X) := \frac{\mathrm{d}P_{\mathrm{T}}}{\mathrm{d}P_{\mathrm{S}}}(X)$$
, it holds

$$\mathbb{E}_{T}[Y(a)] = \mathbb{E}_{T}[\mathbb{E}_{T}[Y(a)|X]]$$

$$= \mathbb{E}_{T}[\mu_{a}(X)]$$

$$= \mathbb{E}_{S}[r(X)\mu_{a}(X)]$$

$$= \mathbb{E}_{S}[r(X)Y(a)]$$

$$= \mathbb{E}_{S}[r(X)Y|A = a]$$
(2)

#### **G-formula:**

$$\mathbb{E}_{\mathsf{T}}[Y(a)] = \mathbb{E}_{\mathsf{T}}[\mu_a(X)] \tag{1}$$

1. Fit two (or one) models on the source data

$$Y_{A=a} = \mu_a(X) + \varepsilon$$

2. Predict the counterfactual outcomes in the target population

$$\hat{Y}_i(a) = \hat{\mu}_a(X_i)$$

3. Average over the target population

$$\hat{\theta} = \phi \left( \frac{1}{m} \sum_{H_i=0} \hat{\mu}_1(X_i), \frac{1}{m} \sum_{H_i=0} \hat{\mu}_0(X_i) \right)$$

#### Reweighted Neyman:

$$\mathbb{E}_{\mathsf{T}}[Y(a)] = \mathbb{E}_{\mathsf{S}}[r(X)Y|A = a] \tag{2}$$

1. Notice that

$$r(X) = \frac{\mathbb{P}(H=1 \mid X)\mathbb{P}(H=1)}{\mathbb{P}(H=0 \mid X)\mathbb{P}(H=0)} = \frac{\alpha \rho(X)}{1-\rho(X)} \quad \text{where} \quad \begin{cases} \rho(X) = \mathbb{P}(H=1 \mid X) \\ \alpha = \mathbb{P}(H=1)/\mathbb{P}(H=0) \end{cases}$$

2. Fit a model for  $\rho$  (using e.g. a logistic regression) on the whole data

$$H = \rho(X) + \varepsilon$$

3. Average over the source population

$$\hat{\theta} = \phi \left( \frac{1}{n_1} \sum_{H_i=1} A_i \hat{r}(X_i) Y_i, \frac{1}{n_0} \sum_{H_i=1} (1 - A_i) \hat{r}(X_i) Y_i \right) \qquad n_a := \#\{i, A_i = a, H_i = 1\}$$

Both approach require to fit a model (r or  $\mu_a$ )

→ one can combine both for a double-robust estimator

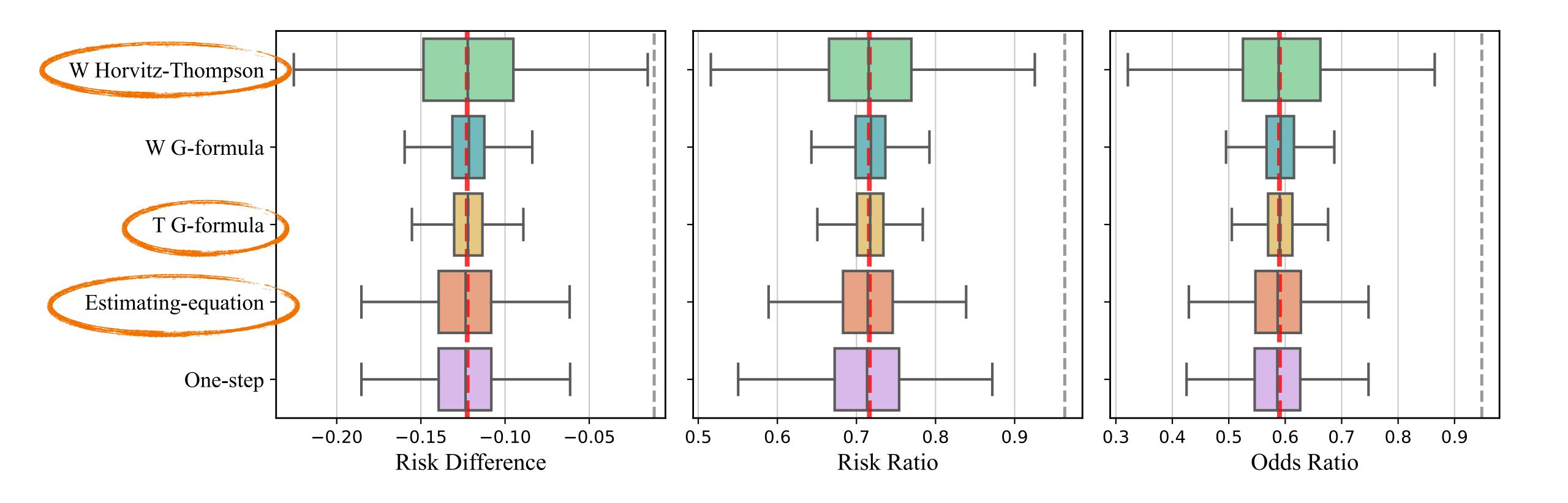
$$\frac{1}{m} \sum_{H_i=0} \hat{\mu}_a(X_i) + \frac{1}{n_a} \sum_{H_i=1} \mathbb{I}\{A_i = a\} \hat{r}(X_i) (Y_i - \hat{\mu}_a(X_i))$$

G-formula

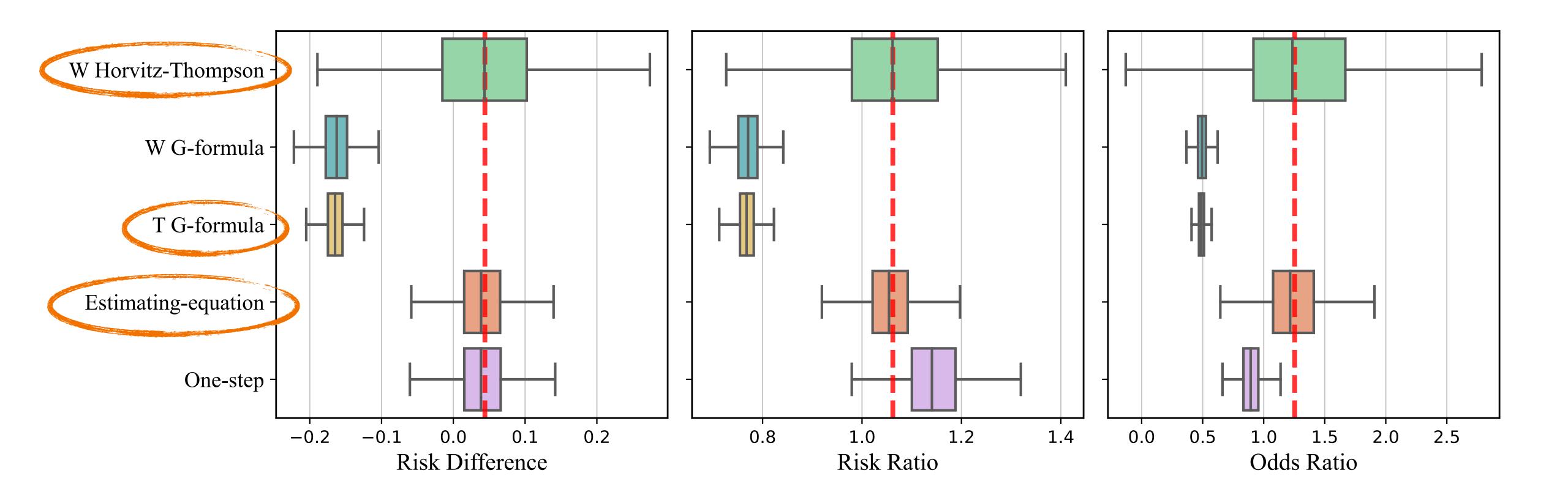
corrective term

ightarrow consistent estimation of  $heta^*$  as soon as  $\hat{\mu}_a$  or  $\hat{r}$  is well specified

Simulation study: under well-specification



Simulation study: under mis-specification of the treatment response



Case study: CRASH-3 results generalization to the Traumabase population

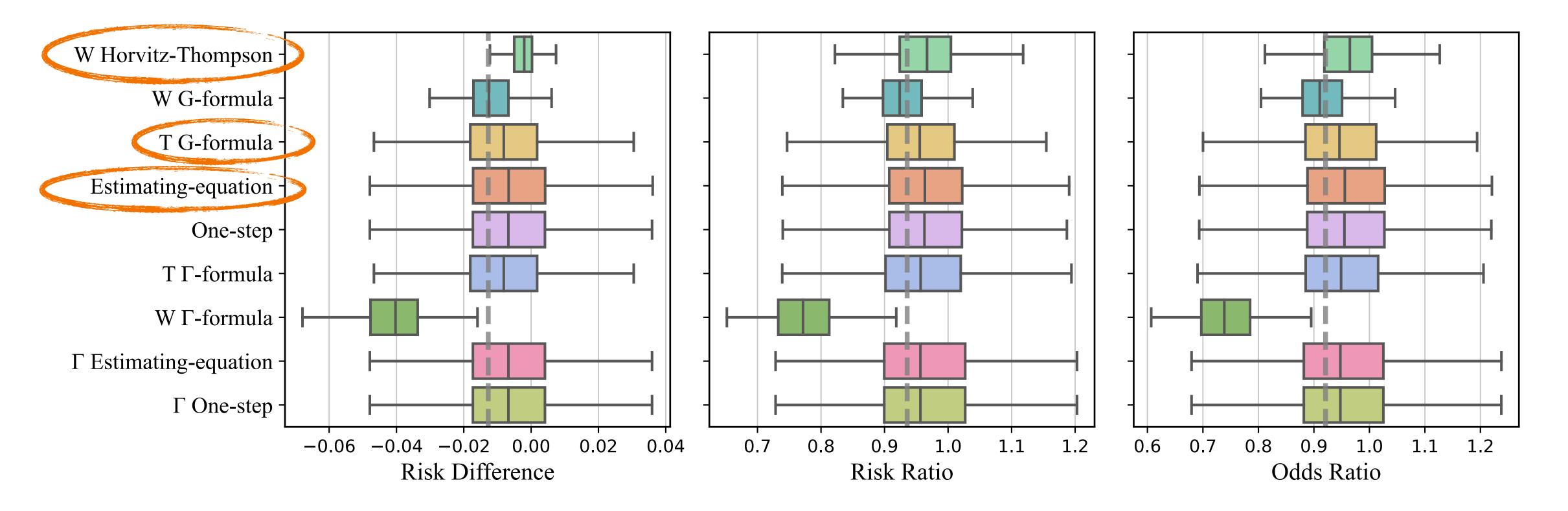
#### CRASH-3 trial

- Randomized trial ( $n \approx 12,000$ )
- Patients with TBI, GCS  $\leq$  12, within 3h
- Treatment: Tranexamic Acid (TXA)
- Outcome: Head injury-related death at 28 days

#### **Traumabase cohort**

- Observational registry  $(m \approx 9,000)$
- Severe trauma, real-world population
- Selected CRASH-3-eligible patients
- Aim: Apply TXA treatment effect to this cohort

Case study: CRASH-3 results generalization to Traumabase population



Grey line corresponds to the effect estimated from the RCT

### Conclusion

- Meta-analysis methods depends on the kind of data available (AD vs ID)
- Traditional methods with AD (e.g. random effects) do not target an estimand which can be interpreted as an average treatment effect
- One can alternatively aggregate the data using weighting strategies at the population level to specifically target a treatment effect
- Under the ID setting, one can reweight at the individual level or resort to double robust approaches

### What's next?

- Sensitivity analysis wrt no-study effect assumption
- Implementing and testing covariate-based weighting strategies for AD

#### We created a very simple R package you can play with:)

#### **CaMeA**

https://cran.r-project.org/package=camea

#### CaMeA: Causal Meta-Analysis for Aggregated Data

A tool for causal meta-analysis. This package implements the aggregation formulas and inference methods proposed in Berenfeld et al. (2025) < doi:10.48550/arXiv.2505.20168>. Users can input aggregated data across multiple studies and compute causally meaningful aggregated effects of their choice (risk difference, risk ratio, odds ratio, etc) under user-specified population weighting. The built-in function camea() allows to obtain precise variance estimates for these effects and to compare the latter to a classical meta-analysis aggregate, the random effect model, as implemented in the 'metafor' package < <a href="https://CRAN.R-project.org/package=metafor">https://CRAN.R-project.org/package=metafor</a>>.

### Thank you for your attention!

B. C., Boughdiri, A., Colnet, B., van Amsterdam, W. A., Bellet, A., Khellaf, R., Scornet, E., & Josse, J. (2025). Causal Meta-Analysis: Rethinking the Foundations of Evidence-Based Medicine. arXiv preprint arXiv:2505.20168.

Boughdiri, A., B. C., Josse, J., & Scornet, E. (2025). A unified framework for the transportability of population-level causal measures. *NeuRIPS 2025*.