

Applied Econometrics
 Prof. Leo Feler
 Quiz 6: Instrumental Variables in Practice

4 points

Name: Key

For this quiz, you'll be analyzing the attached Stata output. You have 30 minutes.

Research question: how do changes in local lending lead to changes in local employment?

Experiment: State-owned banks are privatized (by the state government) in Brazil. Cities with state bank branches see their branches converted to private ownership and possibly lose lending.

1. The Stata output shows the OLS estimates of the effects of changes in lending on changes in employment of low and high-skilled workers. What are the OLS estimates? How might these estimates be biased compared to the "true" effect that changes in lending have on employment? Specifically, under what conditions are the OLS estimates overstating or understating the "true" effect of lending on employment?

- a) 10% lending \rightarrow .79% low skilled employment
 b) 10% lending \rightarrow .76% high skilled employment

Estimates are biased, if we have an omitted variable

True model: $y_i = \alpha + \gamma s_i + \theta A_i + \varepsilon_i$

$E(\varepsilon_i | s_i, A_i) = 0$ unbiased

estimated: $y_i = \tilde{\alpha} + \tilde{\gamma} s_i + \tilde{\varepsilon}_i$

$E(\tilde{\varepsilon}_i | s_i) \neq 0$ biased

$\rightarrow E(\tilde{\gamma}) \neq \gamma$

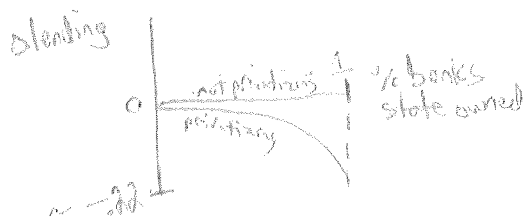
Direction of bias depends on relationship between A & Y

and between S & A :

$$E(\tilde{\gamma}) = \gamma + \theta \cdot \frac{\text{cov}(s, A)}{\text{var}(s)}$$

2. As an instrument for changes in lending, we use whether a city is "privatizing", i.e., whether it is in a state that privatizes, and its initial share (prior to privatization) of state-owned bank branches relative to all bank branches. For example, if a city has one state bank branch and one private bank branch, its initial share of state-owned bank branches is 0.50. If it only has one bank branch and this branch is state-owned, its initial share of state-owned bank branches is 1. We also use the $\text{Privatizing} \times (\text{InitialShareofStateBanks})^2$. What is the relationship between these instruments and lending? What is this estimation called? What is the relationship between these instruments and employment? What is this estimation called? At this point, can you calculate by hand what is the causal effect of changes in lending on changes in employment for a city that has 100% of its bank branches initially state-owned?

The relationship between the instruments and lending is that, for cities that are privatizing, there is a more rapid decline in lending the greater the initial share of banks that are state-owned. This is the first stage regression.



The relationship between the instruments and employment is similarly shaped. This is the reduced form.

When 100% of banks are initially state-owned, we can add the coefficients on the instruments and divide the reduced form over the first stage estimate to estimate the causal effect of changes in lending on changes in employment.

- .18

low skill:

+ - .09

- .22

$\frac{-0.15}{-0.22} = +0.68$

$$\frac{-0.22}{-0.22} = +1.00$$

↑ 1% lending → ↑ 1% low skilled employment

high skill:

$\frac{-0.15}{-0.22} = +0.68$

$\frac{-0.034}{-0.10} = +0.34$

$$\frac{-0.10}{-0.22} = +0.45$$

↑ 10% lending → ↑ 4.5% high skilled employment

3. What conditions must instruments satisfy? Do our instruments satisfy these conditions? How do you know if our instruments satisfy these conditions (if/when applicable, what are the null hypotheses that we're testing—do we reject them)?

Instruments must be strong and valid. Our instruments satisfy both conditions. (don't not satisfy)
 significant relationship between instrument and instrumented variable
 only affects variable of interest through the instrumented variable

$$E[z_i \varepsilon_i] = 0$$

Check by looking at F-stat in 1st stage regression for joint significance of IVs

We want $F \geq 10$

$$F_{low} = \frac{39.84}{\cancel{10.10}}$$

$$F_{high} = \frac{\cancel{38.75}}{38.75}$$

Check with Hansen-Sargan test (if more instruments than endogenous variables)

H_0 : Instruments not invalid

H_a : Instruments are invalid

$$pvalue_{low} = .8155$$

$$pvalue_{high} = .7175$$

4. Look at the difference between the 2SLS and LIML estimates. What can we conclude? Were we ever concerned that our 2SLS estimates would have been biased toward the OLS estimates? Why or why not?

	2SLS	LIML
high	.422	.423
low	1.00	1.00

A comparison between the 2SLS and LIML estimates confirms our belief in the unbiasedness of the estimators.

We had little reason to be concerned about the 2SLS estimates being biased towards the OLS estimates since the instruments satisfied ~~both conditions in question~~ the strength condition ($F_{stat} > 10$).

LIML estimates are less likely to be biased but have higher standard errors.

5. What is Heckman selection? When is a correction for Heckman selection used? Suppose you are estimating the following equation: $\ln w_i = \alpha + \beta s_i + \varepsilon_i$, where w_i , the wage, is only observed for people for whom $L=1$, i.e., for people in the labor force. We want a measure of the return to schooling, β , for the total population, regardless of whether they are in the labor force. For all individuals, you can observe their characteristics (x 's) and whether or not they participate in the labor force, but you only observe wages for those who actually work. Using the Heckman selection correction, how do you estimate β , and how do you test for whether selection into employment exists?

Heckman selection allows us to make results obtained from a non-representative sample more generalizable to the population, i.e., it allows us to correct for selection into the sample.

Heckman selection is used when a sample is not representative of the overall population to which we would like to extend inference. To make results obtained from the selected sample generalizable, we need to correct for selection into the sample.

Based on the example, estimate

- ① $P(L=1|X) = \Phi(x\delta)$ using all N observations, via probit

$$L = \delta_1 x_1 + \delta_2 x_2 + \dots + \delta_k x_k$$

obtain predicted values, $\hat{\delta}_1, \dots, \hat{\delta}_k$.

- ② Calculate Inverse Mills Ratio, $\lambda(x\hat{\delta})$, for N_1 , which is a subsample (i.e., the selected sample) of N . The Inverse Mills Ratio is $\lambda(x\hat{\delta}) = \frac{\phi(x\hat{\delta})}{\Phi(x\hat{\delta})}$ is the probability density function over

[END OF QUIZ]. the cumulative distribution function.

- ③ with $\hat{\lambda}_i$ for each individual (observation), calculated from ②, estimate $\ln w_i = \alpha + \beta s_i + \gamma \hat{\lambda}_i + \varepsilon_i$.

β is now unbiased since we have controlled for what was previously an omitted variable (sample selection). [Note, however, that β might still be biased b/c of other omitted variables (like unobserved ability) but we are assuming the only reason for bias in this case was non-random sampling].

A simple t-test of $H_0: \gamma = 0$ versus $H_a: \gamma \neq 0$ will let us know whether there's selection into employment and whether this selection also affects wages.

```

name: <unnamed>
log: C:\Users\lfeler1\Documents\Banking\Data\Quiz6.log
log type: text
opened on: 20 Apr 2011, 14:35:46

```

```
. reg dlnlowskill dlnmeandefvalue1600 lnlowskill0, robust
```

Linear regression

```

Number of obs = 2664
F( 2, 2661) = 236.46
Prob > F = 0.0000
R-squared = 0.3101
Root MSE = .43117

```

dlnlowskill	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
dlnmean~1600	.0791818	.0177977	4.45	0.000	.044283	.1140806
lnlowskill0	-.1765377	.0083785	-21.07	0.000	-.1929667	-.1601087
_cons	1.55004	.0590426	26.25	0.000	1.434266	1.665815

```
. reg dlnhighskill dlnmeandefvalue1600 lnhighskill0, robust
```

Linear regression

```

Number of obs = 2664
F( 2, 2661) = 178.79
Prob > F = 0.0000
R-squared = 0.2078
Root MSE = .44065

```

dlnhighskill	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
dlnmean~1600	.0762369	.0173699	4.39	0.000	.0421772	.1102967
lnhighskill0	-.1415063	.007543	-18.76	0.000	-.156297	-.1267156
_cons	1.758588	.0455028	38.65	0.000	1.669364	1.847813

.

```
. reg dlnmeandefvalue1600 lnlowskill0 privxinitsharesB privxinitsharesB2, robust
```

Linear regression

```

Number of obs = 2664
F( 3, 2660) = 95.87
Prob > F = 0.0000
R-squared = 0.0869
Root MSE = .50161

```

dlnmean~1600	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lnlowskill0	.0728979	.0070515	10.34	0.000	.059071	.0867248
privxinit~B	-.1852632	.1005906	-1.84	0.066	-.3825069	.0119805
privxinit~2	-.0409581	.1035499	-0.40	0.692	-.2440045	.1620884
_cons	-.1276807	.0460681	-2.77	0.006	-.2180136	-.0373477

```
. reg dlnmeandefvalue1600 lnhighskill0 privxinitsharesB privxinitsharesB2, robust
```

Linear regression

Number of obs = 2664
 F(3, 2660) = 83.92
 Prob > F = 0.0000
 R-squared = 0.0780
 Root MSE = .50406

dlnmean~1600	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lnhighskill0	.0676812	.0076967	8.79	0.000	.052589	.0827733
privXinits~B	-.1549929	.1013876	-1.53	0.126	-.3537995	.0438136
privXinits~2	-.076597	.1051499	-0.73	0.466	-.2827808	.1295869
_cons	-.0181677	.042047	-0.43	0.666	-.1006158	.0642804

```
.
.
. reg dlnlowskill lnlowskill0 privXinitsharesB privXinitsharesB2, robust
```

Linear regression

Number of obs = 2664
 F(3, 2660) = 161.73
 Prob > F = 0.0000
 R-squared = 0.3215
 Root MSE = .42766

dlnlowskill	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lnlowskill0	-.1801152	.0084626	-21.28	0.000	-.1967091	-.1635213
privXinits~B	-.1576193	.0781498	-2.02	0.044	-.3108598	-.0043788
privXinits~2	-.0718823	.0855897	-0.84	0.401	-.2397113	.0959468
_cons	1.647099	.0640072	25.73	0.000	1.52159	1.772608

```
. reg dlnhighskill lnhighskill0 privXinitsharesB privXinitsharesB2, robust
```

Linear regression

Number of obs = 2664
 F(3, 2660) = 121.19
 Prob > F = 0.0000
 R-squared = 0.2051
 Root MSE = .44147

dlnhighskill	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
lnhighskill0	-.1408115	.0077016	-18.28	0.000	-.1559133	-.1257096
privXinits~B	-.0341719	.0806024	-0.42	0.672	-.1922217	.1238778
privXinits~2	-.0664065	.0877393	-0.76	0.449	-.2384506	.1056377
_cons	1.796883	.0505231	35.57	0.000	1.697815	1.895952

```
.
.
. ivreg2 dlnlowskill lnlowskill0 (dlnmeandefvalue1600=privXinitsharesB
privXinitsharesB2), ffirst robust
```

Summary results for first-stage regressions

dlnlowskill	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]
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dlnmean~1600	1.001681	.1749025	5.73	0.000	.6588782	1.344483
lnlowskill10	-.2526149	.0191228	-13.21	0.000	-.2900948	-.215135
_cons	1.773286	.0891614	19.89	0.000	1.598533	1.948039

Underidentification test (Kleibergen-Paap rk LM statistic): 73.805
Chi-sq(2) P-val = 0.0000

Weak identification test (Cragg-Donald Wald F statistic): 24.676
(Kleibergen-Paap rk Wald F statistic): 39.840
Stock-Yogo weak ID test critical values: 10% maximal IV size 19.93
15% maximal IV size 11.59
20% maximal IV size 8.75
25% maximal IV size 7.25

Source: Stock-Yogo (2005). Reproduced by permission.
NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Hansen J statistic (overidentification test of all instruments): 0.054
Chi-sq(1) P-val = 0.8155

Instrumented: dlnmeandefvalue1600
Included instruments: lnlowskill10
Excluded instruments: privXinitshareSB privXinitshareSB2

. ivreg2 dlnhighskill1 lnhighskill10 (dlnmeandefvalue1600=privXinitshareSB
privXinitshareSB2), ffirst robust

Summary results for first-stage regressions

Variable	F(2, 2660)	P-val	(Underid) AP Chi-sq(2)	P-val	(Weak id) AP F(2, 2660)
dlnmeandefva	38.75	0.0000	77.61	0.0000	38.75

NB: first-stage test statistics heteroskedasticity-robust

Stock-Yogo weak ID test critical values for single endogenous regressor:
10% maximal IV size 19.93
15% maximal IV size 11.59
20% maximal IV size 8.75
25% maximal IV size 7.25

Source: Stock-Yogo (2005). Reproduced by permission.
NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Underidentification test
Ho: matrix of reduced form coefficients has rank=K1-1 (underidentified)
Ha: matrix has rank=K1 (identified)
Kleibergen-Paap rk LM statistic Chi-sq(2)=71.42 P-val=0.0000

Weak identification test
Ho: equation is weakly identified
Cragg-Donald Wald F statistic 24.71
Kleibergen-Paap Wald rk F statistic 38.75

Stock-Yogo weak ID test critical values for K1=1 and L1=2:
10% maximal IV size 19.93
15% maximal IV size 11.59
20% maximal IV size 8.75
25% maximal IV size 7.25

Source: Stock-Yogo (2005). Reproduced by permission.
NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Weak-instrument-robust inference

Tests of joint significance of endogenous regressors B1 in main equation

Ho: B1=0 and orthogonality conditions are valid

Anderson-Rubin Wald test F(2,2660)= 4.29 P-val=0.0138

Anderson-Rubin Wald test Chi-sq(2)= 8.59 P-val=0.0136

Stock-Wright LM S statistic Chi-sq(2)= 8.32 P-val=0.0156

NB: Underidentification, weak identification and weak-identification-robust test statistics heteroskedasticity-robust

Number of observations N = 2664

Number of regressors K = 3

Number of endogenous regressors K1 = 1

Number of instruments L = 4

Number of excluded instruments L1 = 2

IV (2SLS) estimation

Estimates efficient for homoskedasticity only
Statistics robust to heteroskedasticity

		Number of obs =	2664
		F(2, 2661) =	149.37
		Prob > F =	0.0000
		Centered R2 =	0.0815
		Uncentered R2 =	0.8155
		Root MSE =	.4742
Total (centered) SS	=	652.1992465	
Total (uncentered) SS	=	3246.973899	
Residual SS	=	599.0156561	

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
dlnhighskill						
dlnmean~1600	.4220745	.1469603	2.87	0.004	.1340376	.7101115
lnhighskill0	-.1687982	.0150686	-11.20	0.000	-.1983322	-.1392643
_cons	1.803372	.0545017	33.09	0.000	1.696551	1.910194

Underidentification test (Kleibergen-Paap rk LM statistic): 71.420
Chi-sq(2) P-val = 0.0000

Weak identification test (Cragg-Donald Wald F statistic): 24.713

(Kleibergen-Paap rk Wald F statistic): 38.749

Stock-Yogo weak ID test critical values: 10% maximal IV size 19.93

15% maximal IV size 11.59

20% maximal IV size 8.75

25% maximal IV size 7.25

Source: Stock-Yogo (2005). Reproduced by permission.

NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Hansen J statistic (overidentification test of all instruments): 0.131
Chi-sq(1) P-val = 0.7175

Instrumented: dlnmeandefvalue1600

Included instruments: lnhighskill0

Excluded instruments: privXinitsharesB privXinitsharesB2

. ivreg2 dlnlowskill lnlowskill0 (dlnmeandefvalue1600=privXinitsharesB
privXinitsharesB2), ffirst robust liml

Summary results for first-stage regressions

			(Underid)		(weak id)
Variable	F(2, 2660)	P-val	AP Chi-sq(2)	P-val	AP F(2, 2660)
dlmneandefva	39.84	0.0000	79.80	0.0000	39.84

NB: first-stage test statistics heteroskedasticity-robust

Stock-Yogo weak ID test critical values for single endogenous regressor:

10% maximal LIML size	8.68
15% maximal LIML size	5.33
20% maximal LIML size	4.42
25% maximal LIML size	3.92

Source: Stock-Yogo (2005). Reproduced by permission.

NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Underidentification test

Ho: matrix of reduced form coefficients has rank=K1-1 (underidentified)

Ha: matrix has rank=K1 (identified)

Kleibergen-Paap rk LM statistic	Chi-sq(2)=73.80	P-val=0.0000
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Weak identification test

Ho: equation is weakly identified

Cragg-Donald Wald F statistic	24.68
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Kleibergen-Paap Wald rk F statistic	39.84
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Stock-Yogo weak ID test critical values for K1=1 and L1=2:

10% maximal LIML size	8.68
15% maximal LIML size	5.33
20% maximal LIML size	4.42
25% maximal LIML size	3.92

Source: Stock-Yogo (2005). Reproduced by permission.

NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Weak-instrument-robust inference

Tests of joint significance of endogenous regressors B1 in main equation

Ho: B1=0 and orthogonality conditions are valid

Anderson-Rubin Wald test	F(2,2660)=	25.08	P-val=0.0000
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Anderson-Rubin Wald test	Chi-sq(2)=	50.24	P-val=0.0000
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Stock-Wright LM S statistic	Chi-sq(2)=	46.05	P-val=0.0000
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NB: Underidentification, weak identification and weak-identification-robust test statistics heteroskedasticity-robust

Number of observations	N =	2664
Number of regressors	K =	3
Number of endogenous regressors	K1 =	1
Number of instruments	L =	4
Number of excluded instruments	L1 =	2

LIML estimation

k	=1.00002
lambda	=1.00002

Estimates efficient for homoskedasticity only
Statistics robust to heteroskedasticity

		Number of obs =	2664
		F(2, 2661) =	143.34
		Prob > F =	0.0000
Total (centered) SS	=	717.0313335	
Total (uncentered) SS	=	1110.381167	
Residual SS	=	1075.96272	
		Centered R2 =	-0.5006
		Uncentered R2 =	0.0310
		Root MSE =	.6355

| Robust

dlnlowskill	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
dlnmean~1600	1.002587	.1751292	5.72	0.000	.6593403	1.345834
lnlowskill0	-.2526897	.0191404	-13.20	0.000	-.2902042	-.2151751
_cons	1.773506	.0892126	19.88	0.000	1.598652	1.948359

Underidentification test (Kleibergen-Paap rk LM statistic): 73.805
Chi-sq(2) P-val = 0.0000

Weak identification test (Cragg-Donald Wald F statistic): 24.676
(Kleibergen-Paap rk Wald F statistic): 39.840

Stock-Yogo weak ID test critical values: 10% maximal LIML size 8.68
15% maximal LIML size 5.33
20% maximal LIML size 4.42
25% maximal LIML size 3.92

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NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Hansen J statistic (overidentification test of all instruments): 0.054
Chi-sq(1) P-val = 0.8156

Instrumented: dlnmeandefvalue1600
Included instruments: lnlowskill0
Excluded instruments: privXinitsharesB privXinitsharesB2

. ivreg2 dlnhighskill lnhighskill0 (dlnmeandefvalue1600=privXinitsharesB
privXinitsharesB2), ffirst robust liml

Summary results for first-stage regressions

Variable	F(2, 2660)	P-val	(Underid) AP Chi-sq(2)	P-val	(weak id) AP F(2, 2660)
dlnmeandefva	38.75	0.0000	77.61	0.0000	38.75

NB: first-stage test statistics heteroskedasticity-robust

Stock-Yogo weak ID test critical values for single endogenous regressor:
10% maximal LIML size 8.68
15% maximal LIML size 5.33
20% maximal LIML size 4.42
25% maximal LIML size 3.92

Source: Stock-Yogo (2005). Reproduced by permission.
NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Underidentification test
Ho: matrix of reduced form coefficients has rank=K1-1 (underidentified)
Ha: matrix has rank=K1 (identified)
Kleibergen-Paap rk LM statistic Chi-sq(2)=71.42 P-val=0.0000

Weak identification test
Ho: equation is weakly identified
Cragg-Donald Wald F statistic 24.71
Kleibergen-Paap Wald rk F statistic 38.75

Stock-Yogo weak ID test critical values for K1=1 and L1=2:
10% maximal LIML size 8.68
15% maximal LIML size 5.33
20% maximal LIML size 4.42
25% maximal LIML size 3.92

Source: Stock-Yogo (2005). Reproduced by permission.
NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Weak-instrument-robust inference

Tests of joint significance of endogenous regressors B1 in main equation

H0: B1=0 and orthogonality conditions are valid

Anderson-Rubin Wald test F(2,2660)= 4.29 P-val=0.0138

Anderson-Rubin Wald test Chi-sq(2)= 8.59 P-val=0.0136

Stock-Wright LM S statistic Chi-sq(2)= 8.32 P-val=0.0156

NB: Underidentification, weak identification and weak-identification-robust test statistics heteroskedasticity-robust

Number of observations N = 2664

Number of regressors K = 3

Number of endogenous regressors K1 = 1

Number of instruments L = 4

Number of excluded instruments L1 = 2

LIML estimation

k =1.00004

lambda =1.00004

Estimates efficient for homoskedasticity only

Statistics robust to heteroskedasticity

Total (centered) SS	=	652.1992465	Number of obs	=	2664
Total (uncentered) SS	=	3246.973899	F(2, 2661)	=	149.29
Residual SS	=	599.3715161	Prob > F	=	0.0000
			Centered R2	=	0.0810
			Uncentered R2	=	0.8154
			Root MSE	=	.4743

dlnhighskill	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]
dlnmean~1600	.4228211	.147288	2.87	0.004	.1341419 .7115003
lnhighskill0	-.1688572	.0150919	-11.19	0.000	-.1984367 -.1392776
_cons	1.803469	.0545314	33.07	0.000	1.696589 1.910348

Underidentification test (Kleibergen-Paap rk LM statistic): 71.420
Chi-sq(2) P-val = 0.0000

Weak identification test (Cragg-Donald Wald F statistic): 24.713

(Kleibergen-Paap rk Wald F statistic): 38.749

Stock-Yogo weak ID test critical values: 10% maximal LIML size 8.68
15% maximal LIML size 5.33
20% maximal LIML size 4.42
25% maximal LIML size 3.92

Source: Stock-Yogo (2005). Reproduced by permission.

NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.

Hansen J statistic (overidentification test of all instruments): 0.131
Chi-sq(1) P-val = 0.7176

Instrumented: dlnmeandefvalue1600

Included instruments: lnhighskill0

Excluded instruments: privXinitshareSB privXinitshareSB2

. log close

name: <unnamed>

log: C:\Users\lfeler1\Documents\Banking\Data\Quiz6.log

log type: text

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