

College of Natural Sciences

Some notes on the Dirichlet

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Dirichlet distribution

Dirichlet prior over probability vector:

$$p(\theta) = \frac{\Gamma(\sum_{k=1}^{K} \alpha_i)}{\prod_{k=1}^{K} \Gamma(\alpha_i)} \prod_{k=1}^{K} \theta_i^{\alpha_i - 1}$$

Discrete prior on observation category:

$$p(Z_i = k|\theta) = \theta_i$$

Posterior over θ :

$$\begin{split} p(\theta|z_i) \propto & p(z_i|\theta) p(\theta) \propto \theta_{z_i} \prod_{k=1}^K \theta_i^{\alpha_i - 1} = \prod_{k=1}^K \theta_i^{\alpha_i + \mathbb{I}(z_i = k) - 1} \\ \text{so } (\theta|z_i) \sim & \text{Dirichlet} \left((\alpha_k + \mathbb{I}(z_i = k))_{k=1}^K \right) \end{split}$$

Dirichlet distribution

If we have multiple samples $z_i \sim \theta$, then

$$p(\theta|z_1,\ldots,z_n) = Dirichlet((\alpha_k + m_k)_{k=1}^K)$$

where
$$m_k = \sum_{i=1}^n \mathbb{I}(z_i = k)$$

Posterior predictive

Let's start with $p(z_{i+1}|z_{1:i})$:

$$p(z_{i+1} = k|z_{1:i}) = \int_{\mathcal{M}} p(z_{i+1}|\theta)p(\theta|z_{1:i})d\theta$$

$$= \int_{\mathcal{M}} \theta_k \text{Dirichlet}(\theta|(\alpha_k + m_k)_{k=1}^K))d\theta$$

$$= \frac{\alpha_k + m_k}{\sum_{j=1}^K \alpha_j + m_j}$$

Posterior predictive: multiple observations

How about $p(z_{i+1:i+j}|z_{1:i})$?

$$p(z_{i+1:i+j}|z_{1:i}) = p(z_{i+j}|z_{1:i+j-1})p(z_{i+j-1}|z_{1:i+j-2})\cdots p(z_{i+1}|z_{1:i})$$

$$= \frac{\alpha_{z_{i+j}} + m_{z_{i+j}}^{(1:i+j-1)}}{i+j-1+\sum_{k}\alpha_{k}} \times \frac{\alpha_{z_{i+j-1}} + m_{z_{i+j-1}}^{(1:i+j-2)}}{i+j-2+\sum_{k}\alpha_{k}}$$

$$\times \cdots \times \frac{\alpha_{z_{i+1}} + m_{z_{i+1}}^{(1:i)}}{i+\sum_{k}\alpha_{k}}$$

$$= \frac{\Gamma(i+\sum_{k}\alpha_{k})}{\Gamma(i+j+\sum_{k}\alpha_{k})} \prod_{k=1}^{K} \frac{\Gamma(m_{k}^{1:i+j} + \alpha_{k})}{\Gamma(m_{k}^{1:i} + \alpha_{k})}$$

where $m_k^{1:j} = \sum_{i=1}^j \mathbb{I}(z_i = k)$. (If you look up the Dirichlet-Multinomial distribution, you will find this with some constant combinatorics terms, and i = 0).

Mixture models

General format:

$$\pi \sim \mathsf{Dirichlet}(lpha)$$
 $heta_k \sim p(heta), \qquad k = 1, \ldots, K$
 $z_i \sim \pi, \qquad i = 1, \ldots, n$
 $x_i \sim f(heta_{z_i})$

Conditional distributions:

$$p(z_i = k|z_{-i}) \propto m_k^{-i} + \alpha_k$$

$$p(z_i = k|z_{-i}, x_i, \theta_k) \propto (m_k^{-i} + \alpha_k) f(x_i; \theta_k)$$

- ▶ We can sample $z_i|z_{-i}, x_i, \theta_k$ by calculating each unnormalized probability, normalizing, and using them to parametrize a multinomial.
- ▶ If f and θ are conjugate, we may be able to integrate out θ_k and instead use $p(x_i | \{x_i, j \neq i, z_i = k\})$

Concrete example: Mixture of multinomials

$$\pi \sim \mathsf{Dirichlet}(lpha)$$
 $\eta_k \sim \mathsf{Dirichlet}(eta), \qquad k = 1, \dots, K$
 $z_i \sim \pi, \qquad i = 1, \dots, n$
 $x_i \sim \mathsf{Multinomial}(M_i, \eta_{z_i})$

If we don't integrate out π and η :

$$\begin{split} \rho(z_i = k | \pi, x_i, \eta_k) &\propto \pi_i \prod_{v=1}^V \eta_{k,v}^{\sum_{j=1}^{M_i} x_{i,j} = v} \\ &\pi | z_{1:n} \sim \text{Dirichlet}\left(\left(\alpha_k + \sum_i z_i = k\right)_{k=1}^K\right) \\ &\eta_k | \{x_i : z_i = k\} \sim \text{Dirichlet}\left(\left(\beta_v + \sum_{i: z_i = k} \sum_{j=1}^{M_i} x_{i,j} = v\right)_{v=1}^V\right) \end{split}$$

Concrete example: Mixture of multinomials

$$p(z_i = k | \pi, x_i, \eta_k) \propto \pi_i \prod_{v=1}^V \eta_{k,v}^{\sum_{j=1}^{M_i} x_{i,j} = v}$$

$$\pi | z_{1:n} \sim \text{Dirichlet} \left(\left(\alpha_k + \sum_i z_i = k \right)_{k=1}^K \right)$$

$$\eta_k | \{ x_i : z_i = k \} \sim \text{Dirichlet} \left(\left(\beta_v + \sum_{i: z_i = k} \sum_{j=1}^{M_i} x_{i,j} = v \right)_{v=1}^V \right)$$

Integrating out π :

$$p(z_i = k | z_{-i} x_i, \eta_k) \propto (m_k^{-i} + \alpha_k) \prod_{v=1}^V \eta_{k,v}^{\sum_{j=1}^{M_i} \mathbb{I}(x_{i,j} = v)}$$

$$\eta_k | \{x_i : z_i = k\} \sim \text{Dirichlet} \left(\left(\beta_v + \sum_{i: z_i = k} \sum_{j=1}^{M_i} \mathbb{I}(x_{i,j} = v) \right)_{v=1}^V \right)$$

Concrete example: Mixture of multinomials

$$p(z_i = k | z_{-i} x_i, \eta_k) \propto (m_k^{-i} + \alpha_k) \prod_{v=1}^V \eta_{k,v}^{\sum_{j=1}^{M_i} \mathbb{I}(x_{i,j} = v)}$$

$$\eta_k | \{x_i : z_i = k\} \sim \text{Dirichlet} \left(\left(\beta_v + \sum_{i: z_i = k} \sum_{j=1}^{M_i} \mathbb{I}(x_{i,j} = v) \right)_{v=1}^V \right)$$

Integrating out both π and η_k :

where $\rho_{k,v} = \sum_{i:z_i=k} \sum_{j=1}^{M_i} \mathbb{I}(x_{i,j}=v)$ is the number of times we've seen token v in cluster k.

Latent Dirichlet allocation

Let's assume we use the the mixture model above to model documents. In such a model, a document is associated with a single cluster, or topic. It might be more reasonable to associate each document with a mixture over topics, so that

An uncollapsed Gibbs sampler

$$p(z_{i,j} = k | \theta_i, \eta_k, w_{i,j} = v) \propto \theta_{i,k} \eta_{k,v}$$

$$p(\theta_i | \{z_{i,j}\}_{j=1}^{M_i}) \sim \text{Dirichlet}\left((m_{i,k} + \alpha)_{k=1}^K\right)$$

$$p(\eta_k | \{w_{i,j} : z_{i,j} = k\}_{j=1}^{M_i}) \sim \text{Dirichlet}\left((\rho_{v,k} + \eta)_{k=1}^K\right)$$

where m_k is the number of times we've seen topic k in document i, and $\rho_{k,v}$ is the number of times we've seen word v in topic k.

A collapsed Gibbs sampler

Integrating out θ_i :

$$p(z_{i,j} = k | z_{i,-j}, \eta_k, w_{i,j} = v) \propto p(z_{i,j} = k | z_{i,-j}) p(w_{i,j} = v | \eta_k)$$

$$= \frac{m_{i,k}^{-j} + \alpha_k}{M_i - 1 + \sum_k \alpha_k} \eta_{k,v}$$

Integrating out η_k :

$$p(z_{i,j} = k | z_{i,-j}, \eta_k, w_{i,j} = v) \propto p(z_{i,j} = k | z_{i,-j}) p(w_{i,j} = v | \{w_{i,j} : z_{i,j} = k\})$$

$$= \frac{m_{i,k}^{-j} + \alpha_k}{M_i - 1 + \sum_k \alpha_k} \cdot \frac{\rho_{k,v}^{-w_{i,j}} + \beta_v}{\sum_{v'} (\rho_{k,v'}^{-w_{i,j}} + \beta_{v'})}$$

$$\propto (m_{i,k}^{-j} + \alpha_k) \cdot \frac{\rho_{k,v}^{-w_{i,j}} + \beta_v}{\sum_{v'} (\rho_{k,v'}^{-w_{i,j}} + \beta_{v'})}$$