

Homework 3

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- (1). Location parameters family, with Jeffrey prior $\pi(\theta) \propto 1$.
- (2). Scale parameters family, with Jeffrey prior $\pi(\beta) \propto 1/\beta$.
- (3). Either if one of $\{\mu, \sigma\}$ fixed, with Jeffrey prior $\pi(\mu, \sigma) \propto 1/\sigma$.
- (4). Scale parameters family, with Jeffrey prior $\pi(x_0) \propto 1/x_0$.

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- (1). By computing

$$\text{Sum}[\text{-D}[\text{Log}[\text{Exp}[-w] * w^i / i!], \{w, 2\}] * \text{Exp}[-w] * w^i / i!, \{i, 0, \text{Infinity}\}]$$

the Fisher information for Poisson distribution $P(\lambda)$ equals to

$$I(\lambda) = 1/\lambda$$

And the Jeffrey prior distribution is

$$\pi(\lambda) \propto \lambda^{-\frac{1}{2}}$$

- (2). By computing

$$\text{Sum}[\text{-D}[\text{Log}[\text{Binomial}[n+r-1, r-1] * w^r * (1-w)^n], \{w, 2\}] * \text{Binomial}[n+r-1, r-1] * w^r * (1-w)^n, \{n, 0, \text{Infinity}\}]$$

the Fisher information for distribution $NBN(r, \theta)$ equals to

$$I(\theta) = \frac{r}{(1-\theta) \cdot \theta^2}$$

And the Jeffrey prior distribution is

$$\pi(\theta) \propto \frac{1}{\theta \sqrt{1-\theta}}$$

- (3). By computing

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Integrate[
  -D[Log[Exp[-x/w]/w],{w,2}]*Exp[-x/w]/w,
  {x,0,Infinity}]
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the Fisher information for distribution $Exp(\frac{1}{\lambda})$ equals to

$$I(\lambda) = \frac{1}{\lambda^2}$$

And the Jeffrey prior distribution is

$$\pi(\lambda) \propto \frac{1}{\lambda}$$

(4). By computing

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Integrate[
  -D[Log[w^a*Exp[-w*x]*x^(a-1)/Gamma[a]],{w,2}]*w^a*Exp[-w*x]*x^(a-1)/Gamma[a],
  {x,0,Infinity}]
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the Fisher information for distribution $\Gamma(\alpha, \lambda)$ where α is already known equals to

$$I(\lambda) = \frac{\alpha}{\lambda^2}$$

And the Jeffrey prior distribution is

$$\pi(\lambda) \propto \frac{1}{\lambda}$$

(5). The likelihood function of the multinomial model is

$$p((x_1, \dots, x_{k-1})|(p_1, \dots, p_{k-1})) = A_{n,\mathbf{x}} \cdot \left(\prod_1^{k-1} p_i^{x_i}\right) \cdot \left(1 - \sum_1^{k-1} p_i\right)^{n - \sum_1^{k-1} x_i}$$

the log likelihood

$$\log p(\mathbf{x}|\mathbf{p}) = C(\mathbf{x}) + \sum_1^{k-1} x_i \log(p_i) + (n - \sum_1^{k-1} x_i) \log(1 - \sum_1^{k-1} p_i)$$

and for $i \neq j$,

$$-D_{ij}(l_{\mathbf{x}}\mathbf{p}) = \frac{n - \sum_1^{k-1} x_i}{(1 - \sum_1^{k-1} p_i)^2}$$

whereas for i ,

$$-D_{ii}(l_{\mathbf{x}}\mathbf{p}) = \frac{n - \sum_1^{k-1} x_i}{(1 - \sum_1^{k-1} p_i)^2} + \frac{x_i}{p_i^2}$$

By taking expectation over \mathbf{X}_{k-1} , we have for $i \neq j$,

$$\begin{aligned}
& \mathbb{E}_{\mathbf{P}} \left[\frac{n - \sum_1^{k-1} x_i}{(1 - \sum_1^{k-1} p_i)^2} \right] \\
&= \sum_{x_1 + \dots + x_{k-1} \leq n} \left(A_{n, \mathbf{x}} \cdot \left(\prod_1^{k-1} p_i^{x_i} \right) \cdot (1 - \sum_1^{k-1} p_i)^{n - \sum_1^{k-1} x_i} \cdot \frac{n - \sum_1^{k-1} x_i}{(1 - \sum_1^{k-1} p_i)^2} \right) \\
&= \sum_{x_1 + \dots + x_{k-1} \leq n-1} \left(A_{n-1, \mathbf{x}} \cdot \left(\prod_1^{k-1} p_i^{x_i} \right) \cdot (1 - \sum_1^{k-1} p_i)^{n-1 - \sum_1^{k-1} x_i} \cdot \frac{n}{1 - \sum_1^{k-1} p_i} \right) \\
&= \frac{n}{1 - \sum_1^{k-1} p_i}
\end{aligned}$$

i.e. for $i \neq j$,

$$I_{ij}(\mathbf{p}) = \frac{n}{1 - \sum_1^{k-1} p_i}$$

and similarly, for each i ,

$$I_{ii}(\mathbf{p}) = \frac{n}{1 - \sum_1^{k-1} p_i} + \frac{n}{p_i}$$

Hence the Jeffrey prior distribution is

$$\pi(\mathbf{p}) \propto \sqrt{|\det(I(\mathbf{p}))|} \propto \prod_1^{k-1} p_i^{-\frac{1}{2}}$$

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For prior distribution

$$\pi(\theta) \propto e^{k_1 a(\theta) + k_2 c(\theta)}$$

with an exponential family of single parameter

$$p(x|\theta) = e^{a(\theta)b(x) + c(\theta)d(x)}$$

The posterior is

$$\pi(\theta|x) \propto e^{a(\theta)(b(x)+k_1) + c(\theta)(1+k_2) + d(x)}$$

i.e. for some k_3 ,

$$\pi(\theta|x) = e^{a(\theta)(b(x)+k_1) + c(\theta)(1+k_2) + (d(x)+k_3)}$$

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To maximize

$$\begin{aligned}
KL(\pi; \pi_0) &= \int_a^b \log \frac{d\Pi_0}{d\Pi} d\Pi \\
&= \int_a^b -\log(\theta f(\theta)) f(\theta) d\theta \\
&= \left(\int_a^z + \int_z^b \right) -\log(\theta f(\theta)) f(\theta) d\theta
\end{aligned}$$

given that $F(z) = \frac{1}{2}$.

If F is optimized in (a, z) , by variation inside $\mathcal{G}_1 = \{g \mid \int_a^z g d\theta = 0\}$

$$\begin{aligned}
&\left. \frac{\partial}{\partial \epsilon} \right|_{\epsilon=0} \int_a^z -\log(\theta(f + \epsilon g)) \cdot (f + \epsilon g) d\theta \\
&= - \int_a^z [\log(\theta) + 1 + \log(f)] \cdot g d\theta \\
&= 0, \quad \forall g \in \mathcal{G}_1
\end{aligned}$$

Hence $(\log(\theta) + 1 + \log(f))$ is constant, *i.e.* $f = \frac{A_1}{\theta}$. Recall that $F(z) = \frac{1}{2}$, we have

$$f(\theta) = \frac{1}{\theta} (2 \ln \frac{z}{a})^{-1}, \quad \theta \in (a, z)$$

Similarly,

$$f(\theta) = \frac{1}{\theta} (2 \ln \frac{b}{z})^{-1}, \quad \theta \in (z, b)$$

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Since the reference prior is independent with the choice of coordinates, we have for coordinate $\{\mu_1, \mu_2\}$ and $\{\phi_1, \phi_2\}$,

$$\begin{aligned}
\pi(\phi_1, \phi_2) &= \pi(\mu_1, \mu_2) \cdot |\det(T_{\mu \rightarrow \phi})|^{-1} \\
&\propto \pi(\mu_1, \mu_2) \cdot \frac{\mu_1}{\mu_2}
\end{aligned}$$

And then consider the reference prior $\pi(\mu_1, \mu_2)$, which is equivalent to the Jeffrey prior. (Bernardo, 1979)

$$\pi(\mu_1, \mu_2) \propto \frac{1}{\mu_1 \mu_2}$$

Hence

$$\pi(\phi_1, \phi_2) \propto \frac{1}{\mu_2^2} = (\phi_1 \phi_2)^{-1}$$