Introduction to modelling sequence evolution

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Who am I?

- CNRS researcher at LBBE in Lyon, France
- Interested in methods for sequence and genome evolution, and in their application
- Keywords: gene tree-species tree, phylogenetic reconstruction (neural networks recently), siteand branch-heterogeneous models of sequence evolution, genome-phenotype associations

Cave oratorem

 There is a reason why the day off is in the middle of the "Phylogeny reconstruction methods" session



https://en.wikipedia.org/wiki/Beware_of_the_dog

This course

A lot of the good stuff was borrowed from Brian

Moore's slides



(http://phylolab.org/)

The bad stuff is mine

Generic statistical paradigm

- Question about some part of the world
- Model of how this part of the world works
- Collect data
- Estimate
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 question

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Example

Is my coin fair?

- Repeated throws=independent identically distributed Bernoulli draws
- Throw coin N times
- Estimate probability of heads

Generic statistical paradigm

Phylogeny example

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- Sites of alignment=independent identically distributed Markov chains running along a phylogeny
- Sequence rodents
- Estimate transition/transversion ratio

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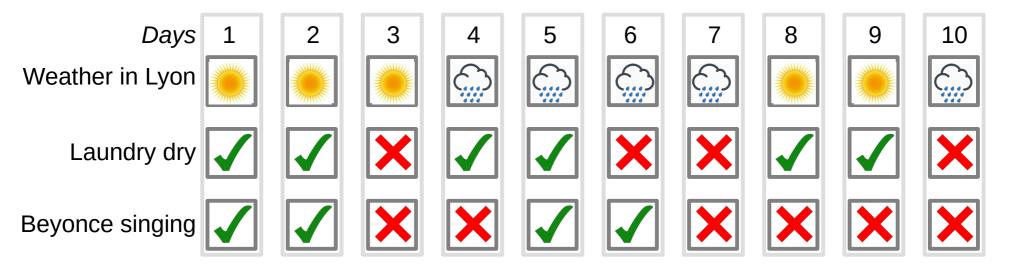
Aims and outline

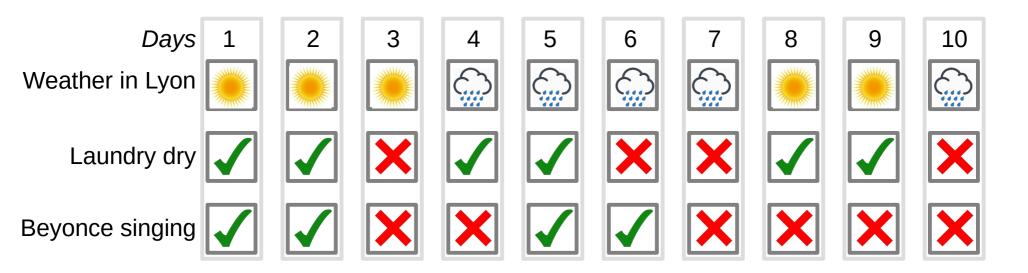
Understand the main ideas underlying models of sequence evolution

- To do so, we will:
 - Introduce important probability notions
 - Play with models of character evolution through simulations
- Briefly present some of the main models of nucleotide evolution

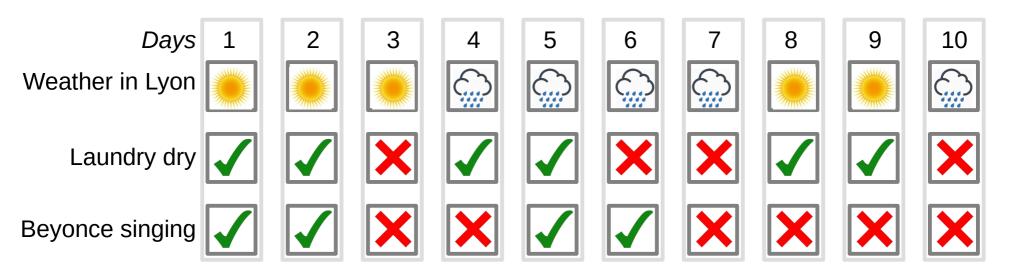
Useful probability concepts

- Conditional probabilities
- Independence/intersection
- Union
- Bayes theorem
- Common distributions that will be useful in this talk:
 - Bernoulli
 - Binomial
 - Poisson
 - Exponential

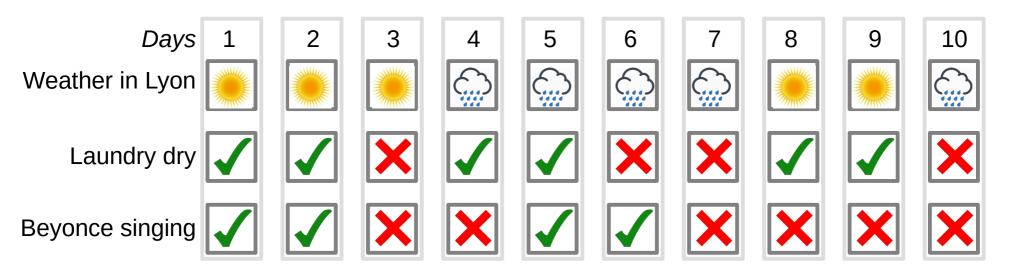




$$P(rainy) = ?$$



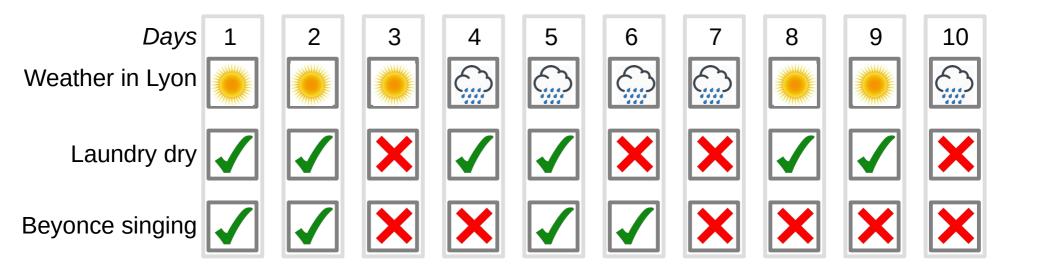
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 $P(sunny)=1-P(rainy)=0.5$



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$$P(dry laundry) = 0.6$$

Record of various events during 10 days

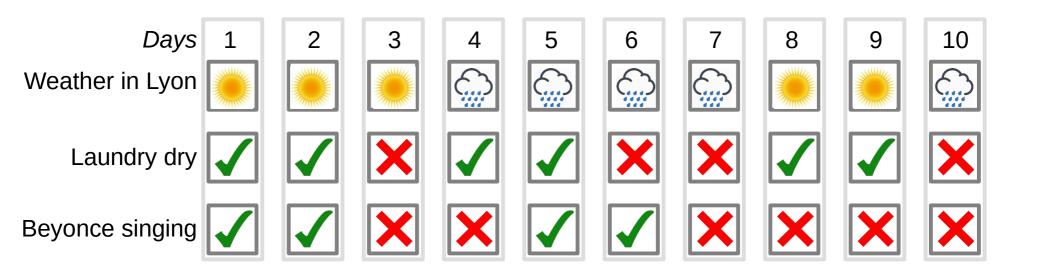


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P(dry | laundry | sunny) = ?

P(dry | laundry | rainy) = ?



$$P(rainy)=0.5$$
 $P(sunny)=1-P(rainy)=0.5$ $P(dry laundry)=0.6$ $P(dry laundry | sunny)=0.8$ Conditional probability: $P(A|B)$

$$P(dry | aundry | sum y) = 0.0$$

 $P(dry | aundry | rainy) = 0.4$

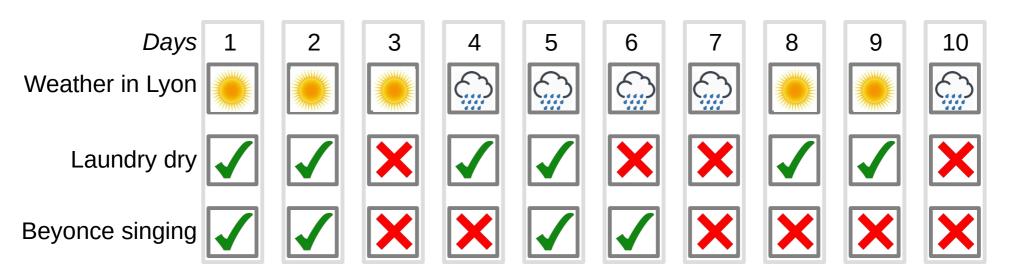
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P(Beyonce singing) = 0.4P(Beyonce singing) = P(Beyonce singing | rainy) = P(Beyonce singing | sunny) = 0.4

Record of various events during 10 days



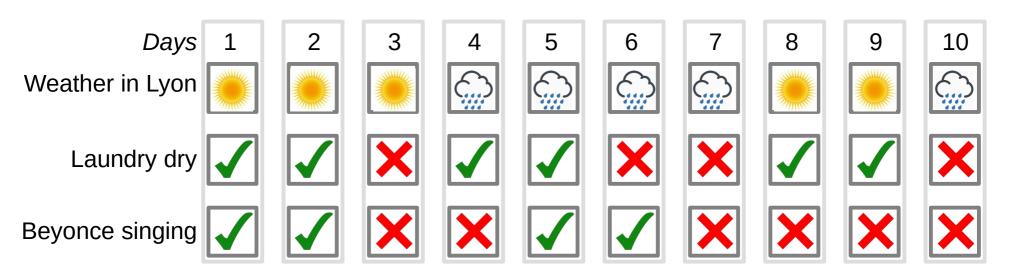
$$P(rainy)=0.5$$
 $P(sunny)=1-P(rainy)=0.5$

The events "Beyonce singing" and "sunny" are independent

P(Beyonce singing) = 0.4

 $P(Beyonce\ singing) = P(Beyonce\ singing|rainy) = P(Beyonce\ singing|sunny)^{20} = 0.4$

Record of various events during 10 days



$$P(rainy)=0.5$$
 $P(sunny)=1-P(rainy)=0.5$

$$P(dry laundry) = 0.6$$

$$P(dry | laundry | sunny) = 0.8$$

$$P(dry | laundry | rainy) = 0.4$$

The events "dry laundry" and "sunny" are NOT independent

$$P(rainy)=0.5$$
 $P(sunny)=1-P(rainy)=0.5$ $P(dry laundry)=0.6$ $P(dry laundry | sunny)=0.8$ $P(dry laundry | rainy)=0.4$

$$P(dry | laundry) = P(dry | laundry | sunny) \times P(sunny)$$

+ $P(dry | laundry | rainy) \times P(rainy)$
= $0.8 \times 0.5 + 0.4 \times 0.5 = 0.6$

Bayes formula

$$P(A|B) = \frac{P(A \land B)}{P(B)} = \frac{P(B \land A)}{P(B)}$$
$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

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$$P(sunny|dry | laundry) = \frac{P(sunny \land dry | laundry)}{P(dry | laundry)} = \frac{P(dry | laundry \land sunny)}{P(dry | laundry)}$$

$$P(sunny | dry | laundry) = \frac{P(dry | laundry | sunny) P(sunny)}{P(dry | laundry)}$$

- Discrete distributions (values in {0,1}, {0,1,2...}):
 - Bernoulli: coin flip: P(X=1)=p; P(X=0)=1-p

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$$Pr(k;n,p) = \Pr(X=k) = \binom{n}{k} p^k (1-p)^{n-k}$$

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 Poisson: how many events of a type over a continuous time: how many meteorites with diameter > 1m in a year:

$$P(k \text{ events in interval}) = e^{-\lambda} \frac{\lambda^{\kappa}}{k!}$$

- Continuous distributions (values in \mathbb{R} , [0,1]...):
 - Exponential: Time between events in a Poisson process: how much time between two meteorites with diameter

 $f(x;\lambda) = \begin{cases} \lambda e^{-\lambda x} & x \ge 0, \\ 0 & x < 0. \end{cases}$

Waiting times in a Poisson process

Some rare, discrete event that occurs at a constant rate in continuous time is described by a Poisson process



Siméon Poisson (1821)

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Waiting times in a Poisson process

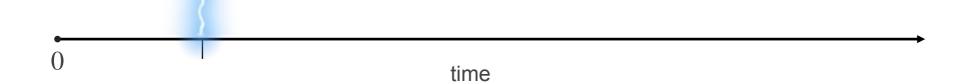
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0 time

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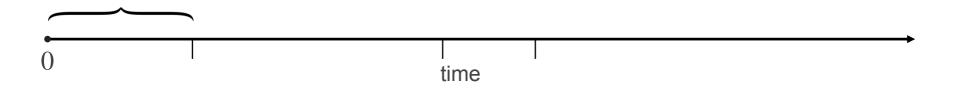


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The waiting (sojourn) time for the first event

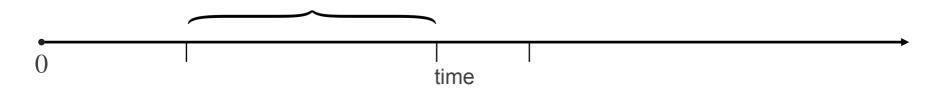


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The waiting (sojourn) time for the second event

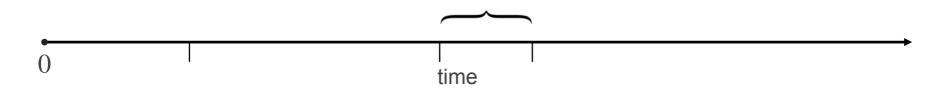


Waiting times in a Poisson process

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The waiting (sojourn) time for the third event

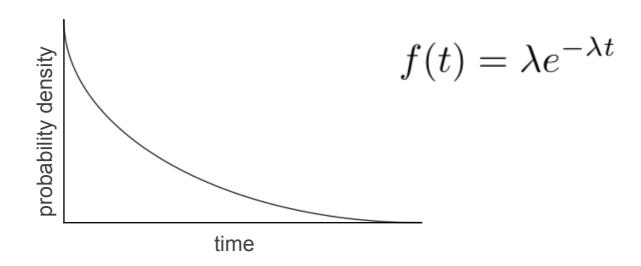


Waiting times in a Poisson process

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The waiting (sojourn) times are exponentially distributed random variables



Aims and outline

Understand the main ideas underlying models of sequence evolution

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Why are we interested in simulations?

- Simulating data forces us to think in terms of a generating process
- By comparing true to simulated data, we can get a sense of how realistic our model is
- Simulations are also central to a lot of inferential problems:
 - Validation of inference methods
 - Posterior predictive tests
 - Approximate Bayesian Computation (ABC)

– ...

Why are we interested in simulations?

- Simulating data forces us to think in terms of a generating process
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Assumption:

If I can simulate my data set, I understand my data set.

Approximate Bayesian Computation (ABC)

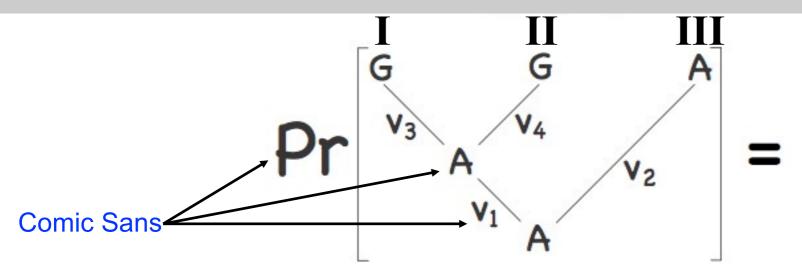
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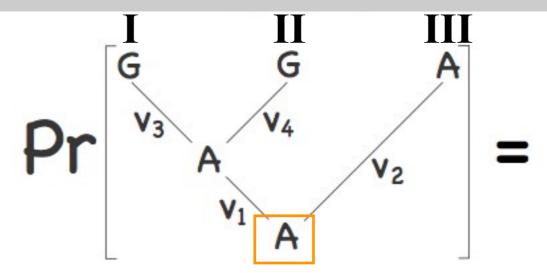
```
SpeciesSequence dataSpecies IGCG--CACCGGCGCAGTCA....Species IIGCGTTCA--GGCG--GTCA....Species IIIACGTTCACCGGCGCAGTCA....
```

How do we compute the likelihood of a site pattern?

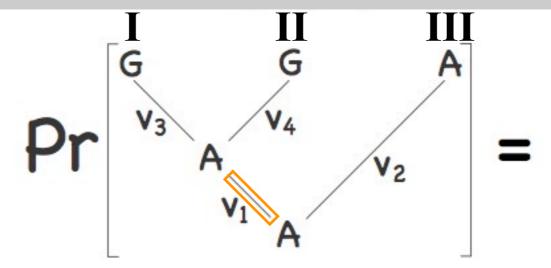
First, we'll see how we compute the likelihood of a site history.

To do that, we'll use simulations.

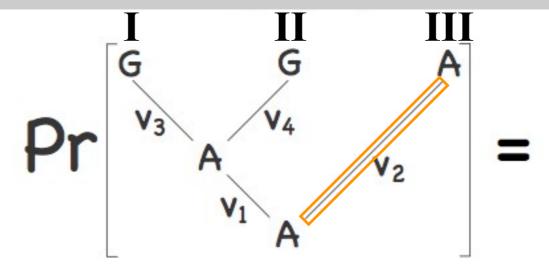




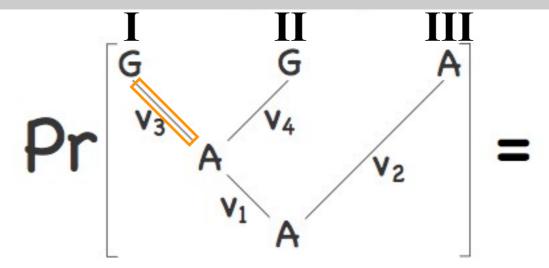
 π_A



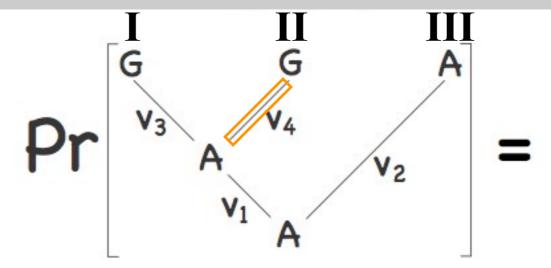
$$\pi_A \times p_{AA}(v_1)$$



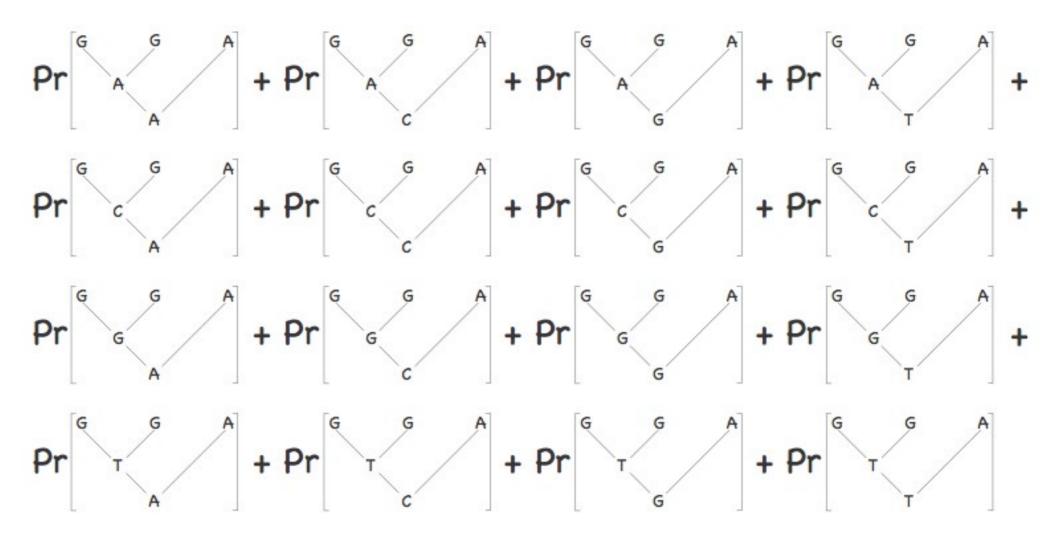
$$\pi_A \times p_{AA}(v_1) \times p_{AA}(v_2)$$

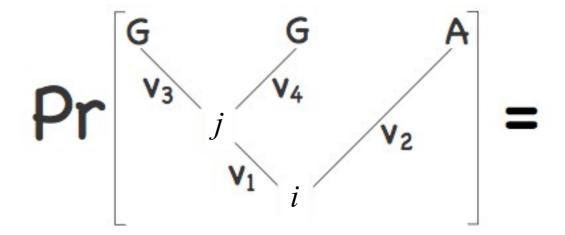


$$\pi_A \times p_{AA}(v_1) \times p_{AA}(v_2) \times p_{AG}(v_3)$$

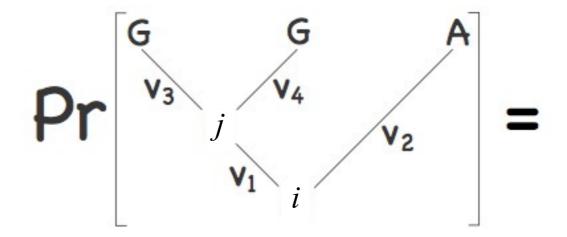


$$\pi_A \times p_{AA}(v_1) \times p_{AA}(v_2) \times p_{AG}(v_3) \times p_{AG}(v_4)$$





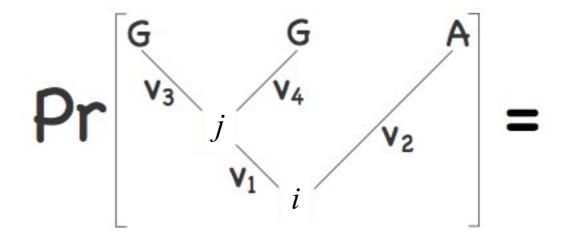
$$\pi_i \times p_{ij}(v_1) \times p_{iA}(v_2) \times p_{jG}(v_3) \times p_{jG}(v_4)$$



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 π_i Stationary frequencies

 $p_{ij}(v)$ Transition probabilities



$$\pi_i \times p_{ij}(v_1) \times p_{iA}(v_2) \times p_{jG}(v_3) \times p_{jG}(v_4)$$

 π_i Stationary frequencies $p_{ij}(v)$ Transition probabilities

Continuous-time Markov Chains (CTMC)

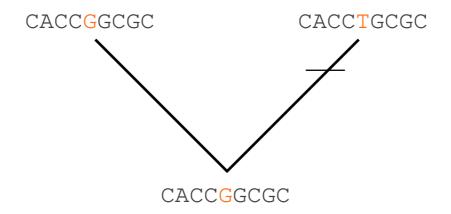
Evolution of discrete traits (*e.g.*, substitution models, morphological models)

We will introduce:

- Substitution rates
- Substitution probabilities
- Stationary frequencies

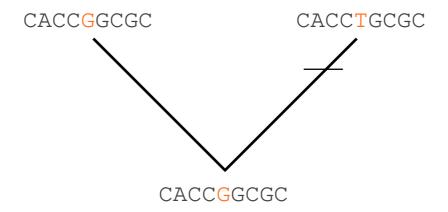
Models describe changes in the nucleotide sites at the species level

We are (generally) not trying to model the dynamics of allele frequencies in populations



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We are modeling the process of **nucleotide substitution**, which describes the outcome of the mutation and fixation processes within populations

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Stochastic model in which the next state of the chain depends only on the current state

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Stochastic model in which the next state of the chain depends only on the current state

The model is central to model-based inference

Even if the parameters of the substitution model are not of direct interest, they are nevertheless critical to estimation of the focal model parameters

 $p_{ij}(v)$

The Instantaneous-Rate Matrix

A Continuous-time Markov model is defined by a matrix of substitution rates

A table that specifies the rates of all possible changes between states.

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A hypothetical instantaneous-rate matrix

		To				
		A	C	G	T	
From	A	-1.916	0.541	0.787	0.588	
	C	0.148	-1.069	0.415	0.506	
	G	0.286	0.170	-0.591	0.135	
	T	0.525	0.236	0.594	-1.355	

This table of rates specifies the instantaneous rate of change between sates.

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The rows of the table must sum to zero.

A hypothetical instantaneous-rate matrix

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

If the current state of the Markov chain is i, the next substitution will occur after an exponentially distributed waiting time with rate parameter $-q_{ii}$

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$$P(C->A) = q_{cA}/-q_{cC} = 0.148 | 1.069 = 0.138$$

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$$P(C->A) = q_{cA}/-q_{cc} = 0.148 | 1.069 = 0.138$$

 $P(C->G) = q_{cG}/-q_{cc} = 0.415 | 1.069 = 0.388$

A hypothetical instantaneous-rate matrix

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$$P(C->A) = q_{cA}/-q_{cc} = 0.148 | 1.069 = 0.138$$

 $P(C->G) = q_{cG}/-q_{cc} = 0.415 | 1.069 = 0.388$
 $P(C->T) = q_{cG}/-q_{cc} = 0.506 | 1.069 = 0.474$

A hypothetical instantaneous-rate matrix

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

If the current state of the Markov chain is i, the next substitution will occur after an exponentially distributed waiting time with rate parameter $-q_{ii}$

We can read this rate parameter directly from the instantaneous-rate matrix: e.g., if the current state is C, the rate parameter is $-q_{cc} = -(-1.069) = 1.069$ or, equivalently: $q_{c4} + q_{c6} + q_{c7} = 0.148 + 0.415 + 0.506 = 1.069$

$$P(C->A) = q_{cA}/-q_{cc} = 0.148 | 1.069 = 0.138$$

$$P(C->G) = q_{cG}/-q_{cc} = 0.415 | 1.069 = 0.388$$

$$P(C->T) = q_{cT}/-q_{cc} = 0.506 | 1.069 = \underline{0.474}$$

$$\Sigma P_{ij} = 1.0$$

Developing Intuition for CTMCs: A Monte Carlo Simulation Experiment

What the heck is Monte Carlo Simulation?

We generate a number of **replicate outcomes** (we will perform multiple trials)

Developing Intuition for CTMCs: A Monte Carlo Simulation Experiment

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We generate a number of **replicate outcomes** (we will perform multiple trials) of a **stochastic process** (our replicate experiments involve randomness)

Developing Intuition for CTMCs: A Monte Carlo Simulation Experiment

What the heck is Monte Carlo Simulation?

We generate a number of **replicate outcomes** (we will perform multiple trials) of a **stochastic process** (our replicate experiments involve randomness) under a **fully specified model** (with specific values for all model parameters)

A simple Monte Carlo Simulation

We will assume that we have a **fully specified phylogenetic model**:

A simple Monte Carlo Simulation

V	/e will assume that we have a fully specified phylogenetic model:
	there is a single branch (the tree 'topology')

A simple Monte Carlo Simulation

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A simple Monte Carlo Simulation

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there is a single branch (the tree 'topology') the branch has a length of 0.5 (expected substitutions/site) the instantaneous-rate matrix is known

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

()

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We will assume that we have a **fully specified phylogenetic model**:

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What is the matrix of substitution probabilities over our branch of length 0.5?

A simple Monte Carlo Simulation

We will use a random-number generator to mimic the **stochastic process**

A simple Monte Carlo Simulation

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

A simple Monte Carlo Simulation

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Rate of leaving the current state, T = 1.355

A simple Monte Carlo Simulation

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ \hline 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Probability of changing to A:

$$P(T->A) = q_{TA}/-q_{TT} = 0.525 | 1.355 = 0.387$$

A simple Monte Carlo Simulation

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Probability of changing to C:

$$P(T->A) = q_{TA}/-q_{TT} = 0.525 | 1.355 = 0.387$$

$$P(T->C) = q_{TC}/-q_{TT} = 0.236 | 1.355 = 0.174$$

A simple Monte Carlo Simulation

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Probability of changing to G:

$$P(T->A) = q_{TA}/-q_{TT} = 0.525 | 1.355 = 0.387$$

 $P(T->C) = q_{TC}/-q_{TT} = 0.236 | 1.355 = 0.174$
 $P(T->G) = q_{TC}/-q_{TT} = 0.594 | 1.355 = 0.438$

A simple Monte Carlo Simulation

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

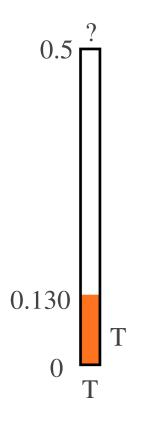
Generate an exponentially distributed waiting time, *x*:

A simple Monte Carlo Simulation

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate an exponentially distributed waiting time, x: rate when process is in T: $\lambda = 1.355$

A simple Monte Carlo Simulation



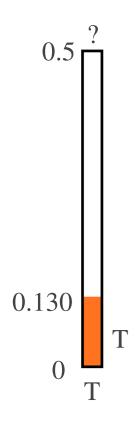
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Generate an exponentially distributed waiting time, *x*:

rate when process is in T: $\lambda = 1.355$

Draw x: $x \sim dnExponential(1.355) : x = 0.130$

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

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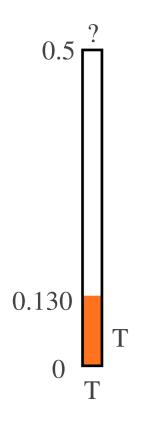
Probabilities of substitution events in state T:

$$P(T->A) = q_{TA}/-q_{TT} = 0.525 | 1.355 = 0.387$$

 $P(T->C) = q_{TC}/-q_{TT} = 0.236 | 1.355 = 0.174$

$$P(T->G) = q_{rr}/-q_{rr} = 0.594 | 1.355 = 0.438$$

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate an exponentially distributed waiting time, *x*:

rate when process is in T: $\lambda = 1.355$

Draw x: $x \sim dnExponential(1.355)$: x = 0.130

Specify a set of intervals: intervals

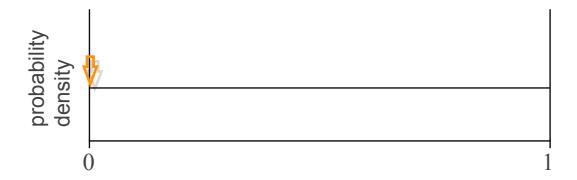
$$P(T->A) = 0.387$$
 $0 - 0.387$ (choose A)

$$P(T->C) = 0.174$$
 0.387 - 0.561 (choose C)

$$P(T->G) = 0.438$$
 0.561 – 1 (choose G)

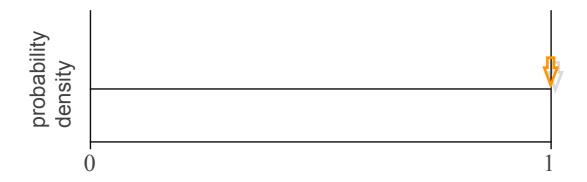
Only the width of the bins matters, not their order

Our uniform(0,1) random number will take any value between 0 and 1 with equal probability (by definition)



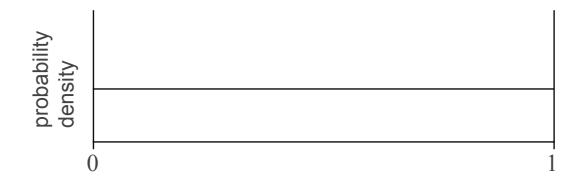
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Imagine that there are two possible outcomes, A and B, which occur with probabilities:

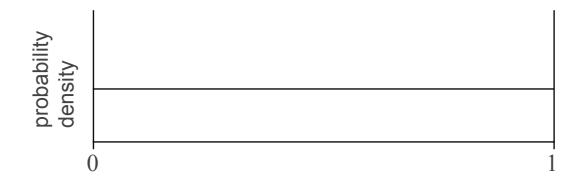
Probability of outcomes:

$$P(\text{option A}) = 0.6$$

$$P(\text{option B}) = 0.4$$

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We can choose randomly (according to their probabilities) by specifying these intervals...

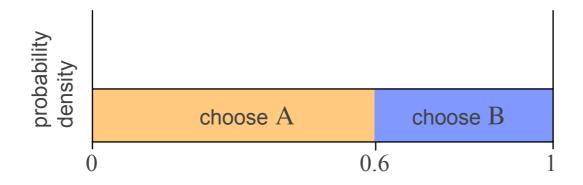
Probability of outcomes: intervals

$$P(\text{option } A) = 0.6$$
 $0.0 - 0.6$ (choose A)

$$P(\text{option B}) = 0.4$$
 $0.6 - 1.0$ (choose B)

Only the width of the bins matters, not their order

Our uniform(0,1) random number will take any value between 0 and 1 with equal probability (by definition)



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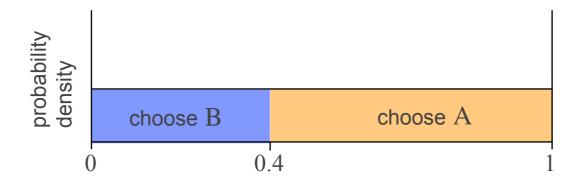
Probability of outcomes: intervals

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 $0.0 - 0.6$ (choose A)

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Imagine that there are two possible outcomes, A and B, which occur with probabilities:

Probability of outcomes:

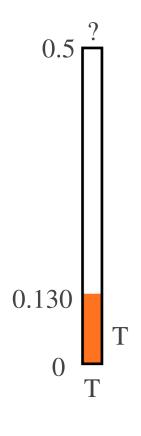
$$P(\text{option } A) = 0.6$$

$$P(\text{option B}) = 0.4$$

Or equivalently by specifying these intervals...

Probability of outcomes: intervals $P(\text{option }A) = 0.6 \qquad 0.4 - 1.0 \quad (\text{choose }A)$ $P(\text{option }B) = 0.4 \qquad 0.0 - 0.6 \quad (\text{choose }B)$

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate an exponentially distributed waiting time, *x*:

rate when process is in T: $\lambda = 1.355$

Draw x: $x \sim dnExponential(1.355)$: x = 0.130

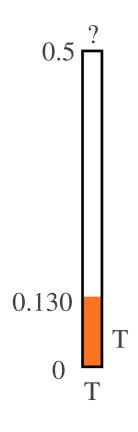
Specify a set of intervals: intervals

$$P(T->A) = 0.387$$
 $0 - 0.387$ (choose A)

$$P(T->C) = 0.174$$
 0.387 - 0.561 (choose C)

$$P(T->G) = 0.438$$
 0.561 – 1 (choose G)

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

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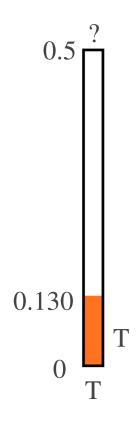
$$P(T->A) = 0.387$$
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 $0.561-1$ (choose G)

Draw a uniformly distributed number, u, to select substitution event: $u \sim dnUniform(0,1)$: u=0.446

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate an exponentially distributed waiting time, *x*:

rate when process is in T: $\lambda = 1.355$

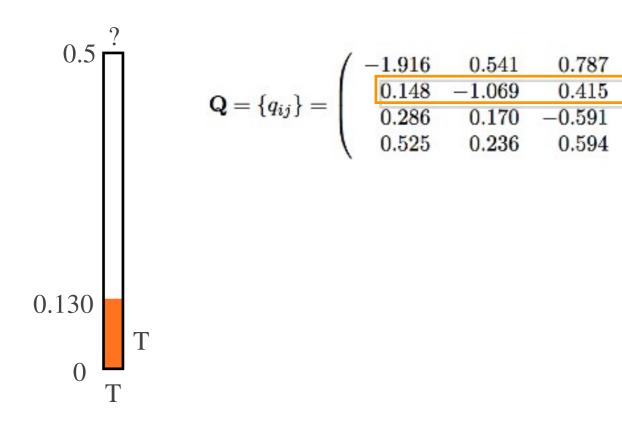
Draw x: $x \sim dnExponential(1.355)$: x = 0.130

Specify a set of intervals: intervals
$$P(T->A) = 0.387$$
 $0 - 0.387$ (choose A) $P(T->C) = 0.174$ $0.387 - 0.561$ (choose C) $P(T->G) = 0.438$ $0.561 - 1$ (choose G)

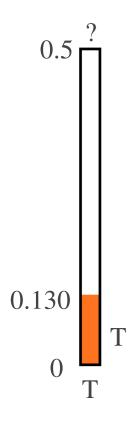
Draw a uniformly distributed number, u, to select substitution event: $u \sim dnUniform(0,1)$: u=0.446

0.588 0.506 0.135

A simple Monte Carlo Simulation



A simple Monte Carlo Simulation

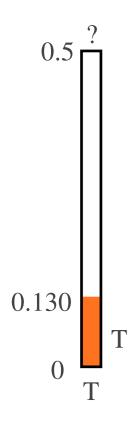


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Generate waiting time to next event:

Rate when process in state C: $-q_{cc} = \lambda = 1.069$

A simple Monte Carlo Simulation



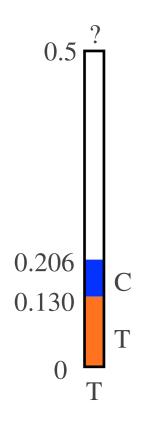
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Generate waiting time to next event:

Rate when process in state C: $-q_{cc} = \lambda = 1.069$

Draw x: $x \sim dnExponential(1.069)$: x = 0.076

A simple Monte Carlo Simulation



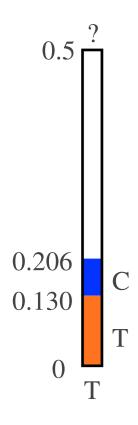
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A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ \hline 0.148 & -1.069 & 0.415 & 0.506 \\ \hline 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate waiting time to next event:

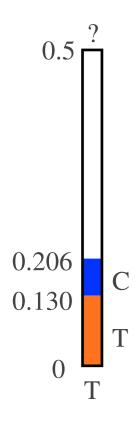
Rate when process in state C: $-q_{cc} = \lambda = 1.069$

Draw x: $x \sim dnExponential(1.069)$: x = 0.076

Substitution probabilities in state C:

$$P(C->A) = q_{cA}/-q_{cC} = 0.148 \mid 1.069 = 0.138$$

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

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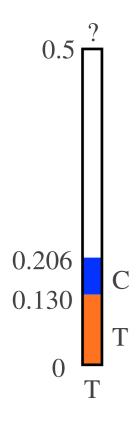
Draw x: $x \sim dnExponential(1.069)$: x = 0.076

Substitution probabilities in state C:

$$P(C->A) = q_{cA}/-q_{cC} = 0.148 \mid 1.069 = 0.138$$

$$P(C->G) = q_{cc}/-q_{cc} = 0.415 \mid 1.069 = 0.388$$

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate waiting time to next event:

Rate when process in state C: $-q_{cc} = \lambda = 1.069$

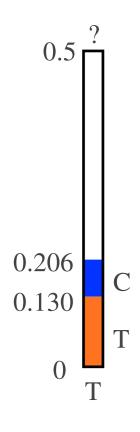
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Substitution probabilities in state C:

$$P(C->A) = q_{ca}/-q_{cc} = 0.148 | 1.069 = 0.138$$

 $P(C->G) = q_{cc}/-q_{cc} = 0.415 | 1.069 = 0.388$
 $P(C->T) = q_{cc}/-q_{cc} = 0.506 | 1.069 = 0.474$

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

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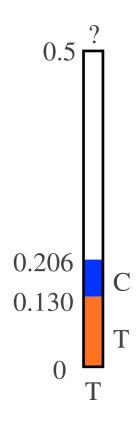
Specify a set of intervals: intervals

$$P(C->A) = 0.138$$
 $0 - 0.138$ (choose A)

$$P(C->G) = 0.388$$
 $0.138 - 0.526$ (choose G)

$$P(C->T) = 0.474$$
 0.526 - 1 (choose T)

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate waiting time to next event:

Rate when process in state C: $-q_{cc} = \lambda = 1.069$

Draw x: $x \sim dnExponential(1.069)$: x = 0.076

Specify a set of intervals: intervals

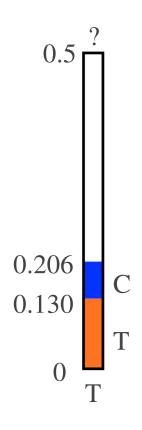
$$P(C->A) = 0.138$$
 $0 - 0.138$ (choose A)

$$P(C->G) = 0.388$$
 $0.138 - 0.526$ (choose G)

$$P(C->T) = 0.474$$
 $0.526 - 1$ (choose T)

Draw a uniformly distributed number, u, to select substitution event: $u \sim dnUniform(0,1)$: u=0.317

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate waiting time to next event:

Rate when process in state C: $-q_{cc} = \lambda = 1.069$

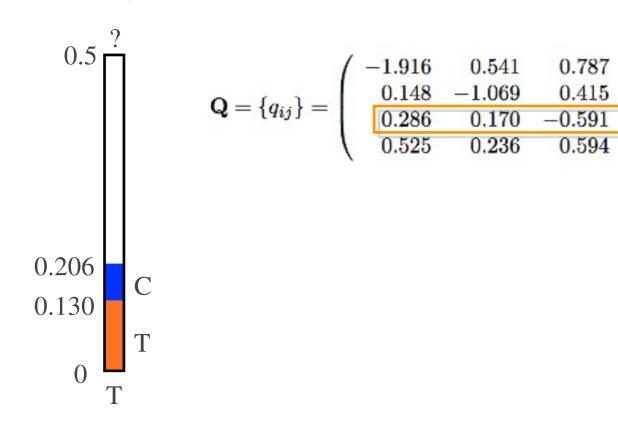
Draw x: $x \sim dnExponential(1.069)$: x = 0.076

Specify a set of intervals: intervals
$$P(C->A) = 0.138$$
 $0 - 0.138$ (choose A) $P(C->G) = 0.388$ $0.138 - 0.526$ (choose G) $P(C->T) = 0.474$ $0.526 - 1$ (choose T)

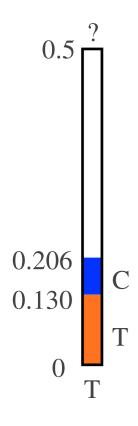
Draw a uniformly distributed number, u, to select substitution event: $u \sim dnUniform(0,1)$: u=0.317

0.588 0.506 0.135

A simple Monte Carlo Simulation



A simple Monte Carlo Simulation

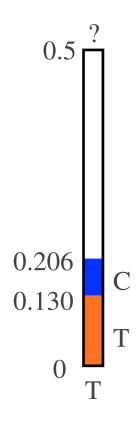


$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate waiting time to next event:

Rate when process in state G: $-q_{cc} = \lambda = 0.591$

A simple Monte Carlo Simulation



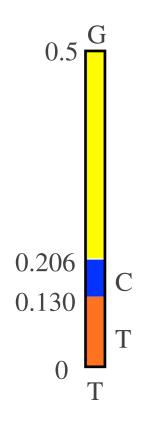
$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate waiting time to next event:

Rate when process in state G: $-q_{cc} = \lambda = 0.591$

Draw x: $x \sim dnExponential(1.069)$: x = 1.820

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

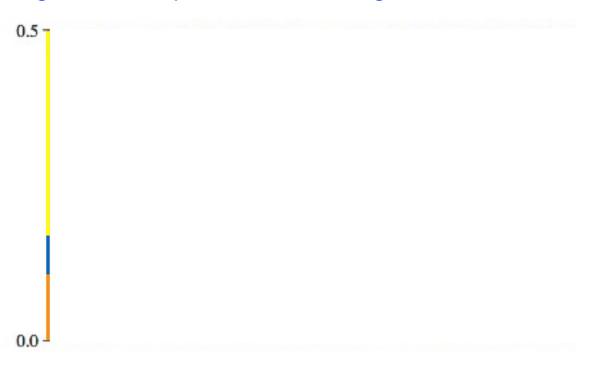
Generate waiting time to next event:

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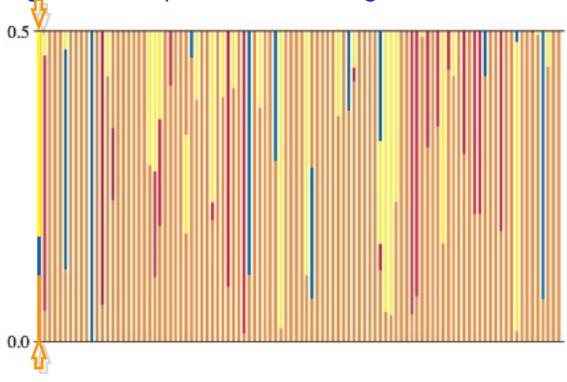
Draw x: $x \sim dnExponential(1.069)$: x = 1.820

 $1.820 > 0.5-0.206 \rightarrow Terminate simulation (in state G)$

Estimating transition probabilities using Monte Carlo simulation

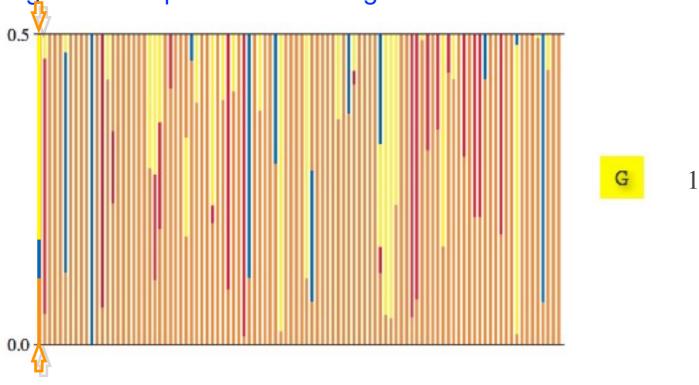


Estimating transition probabilities using Monte Carlo simulation



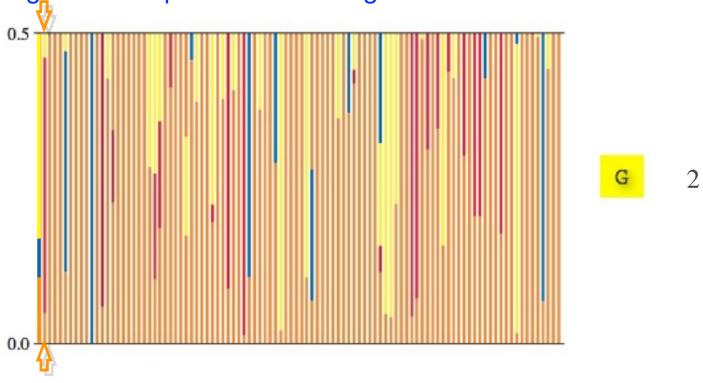
 $T \rightarrow C \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



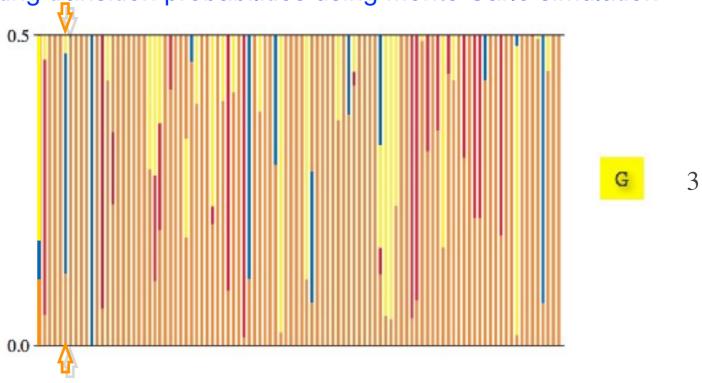
 $T \rightarrow C \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



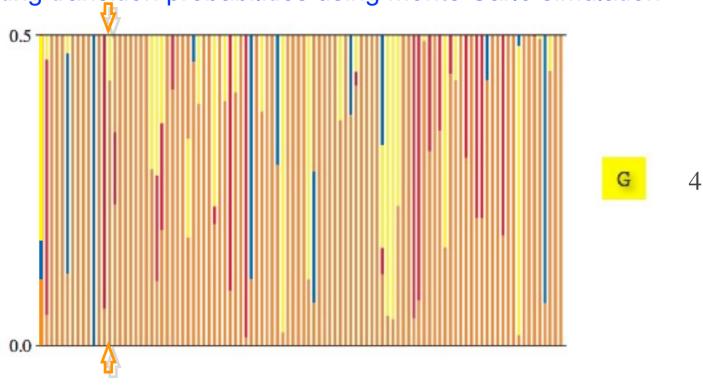
 $T \rightarrow A \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



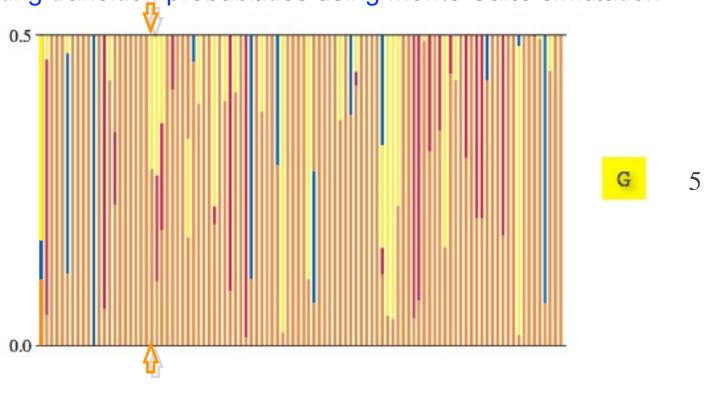
$$T \rightarrow C \rightarrow G$$

Estimating transition probabilities using Monte Carlo simulation



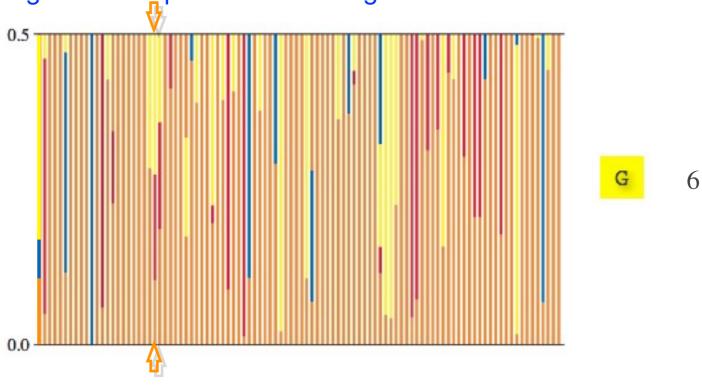
 $T \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



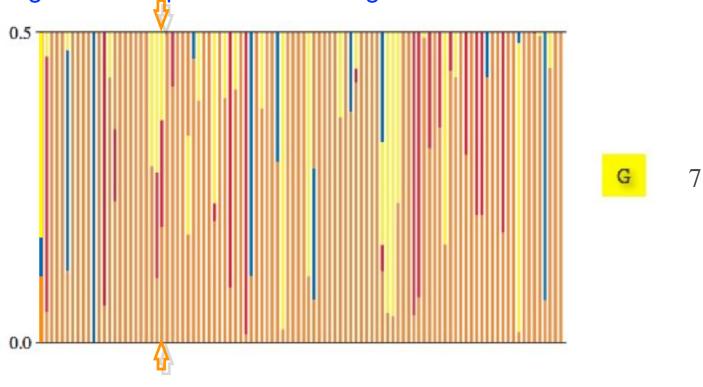
 $T \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



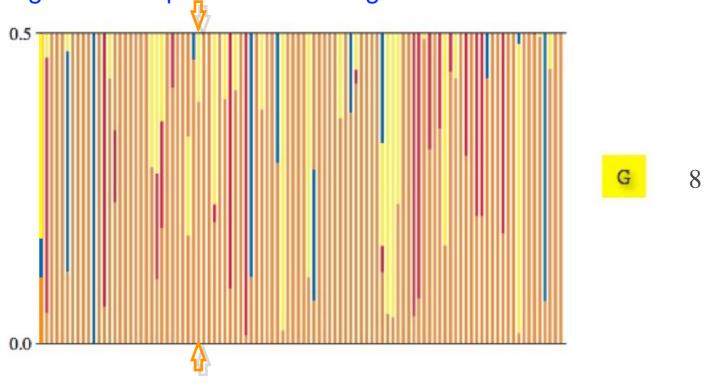
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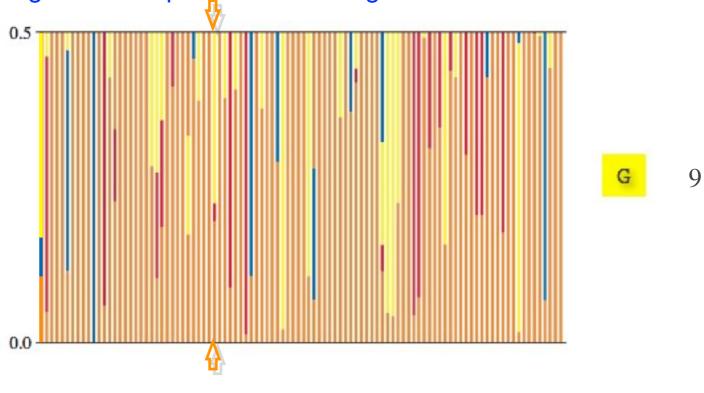
 $T \rightarrow A \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



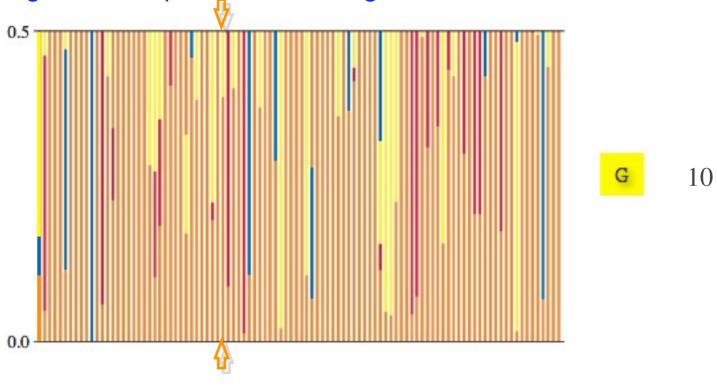
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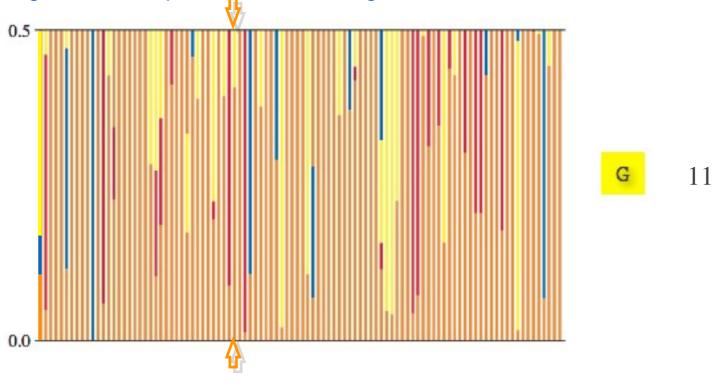
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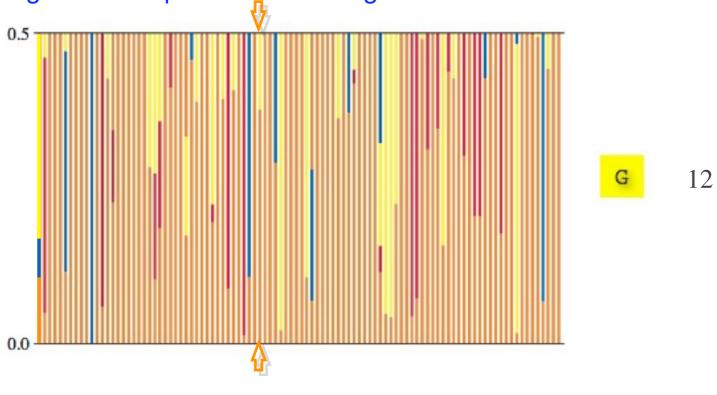
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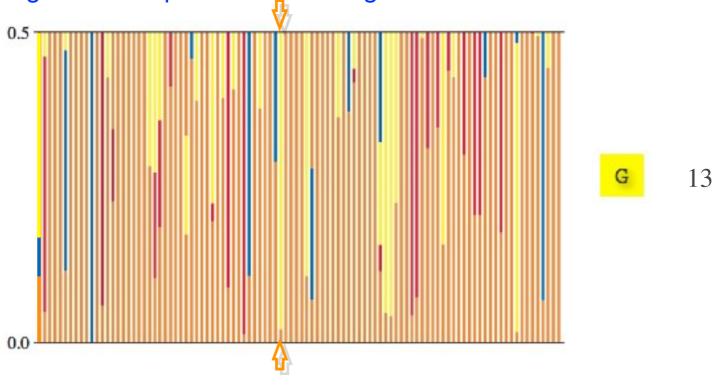
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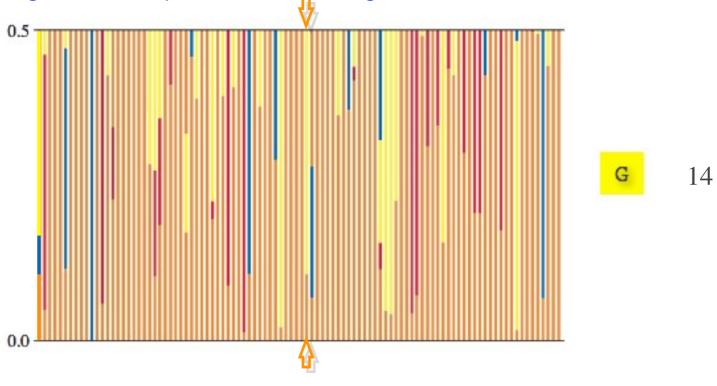
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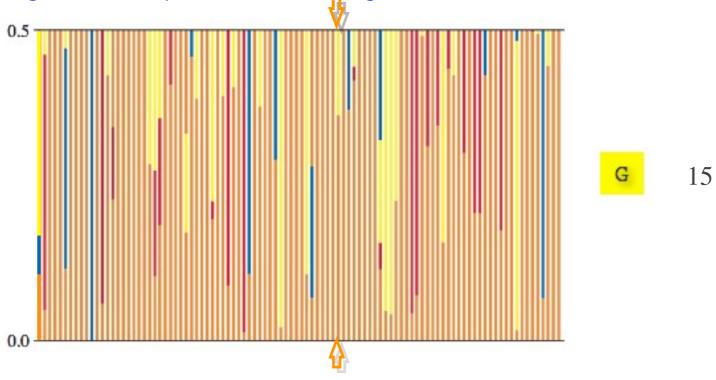
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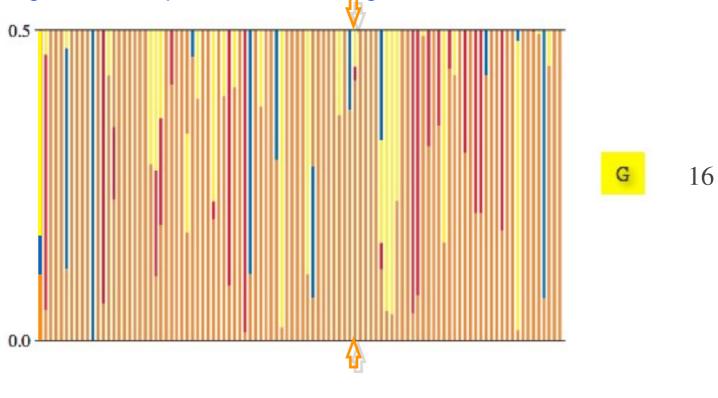


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Estimating transition probabilities using Monte Carlo simulation

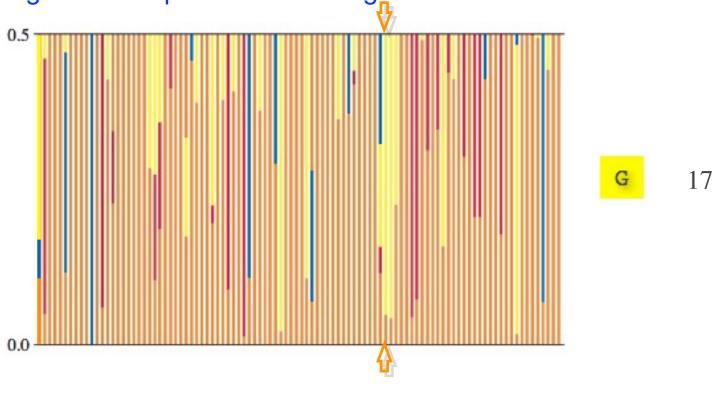


Estimating transition probabilities using Monte Carlo simulation



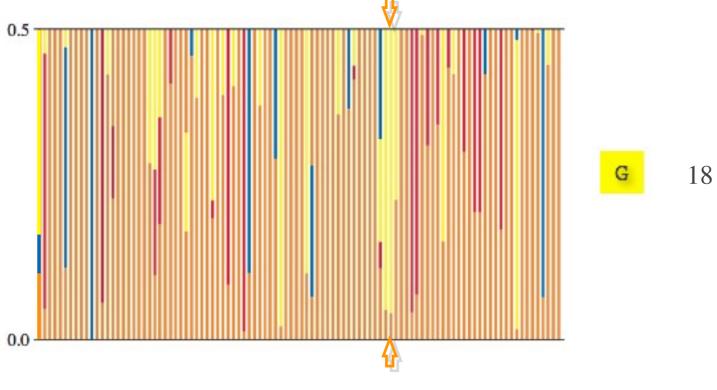
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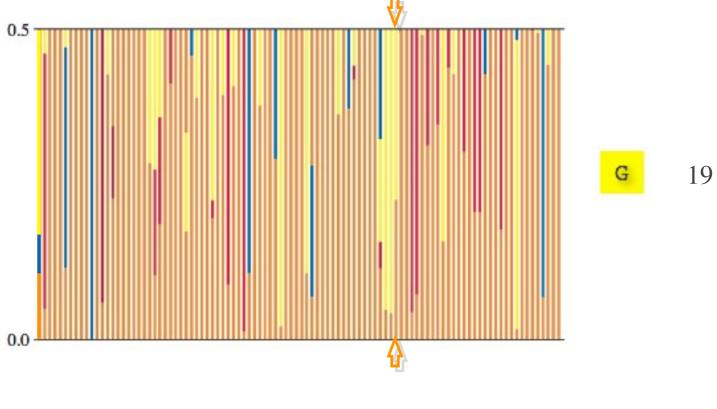
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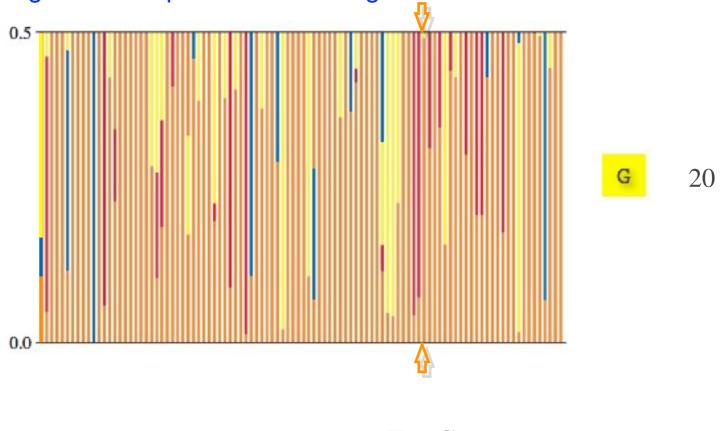
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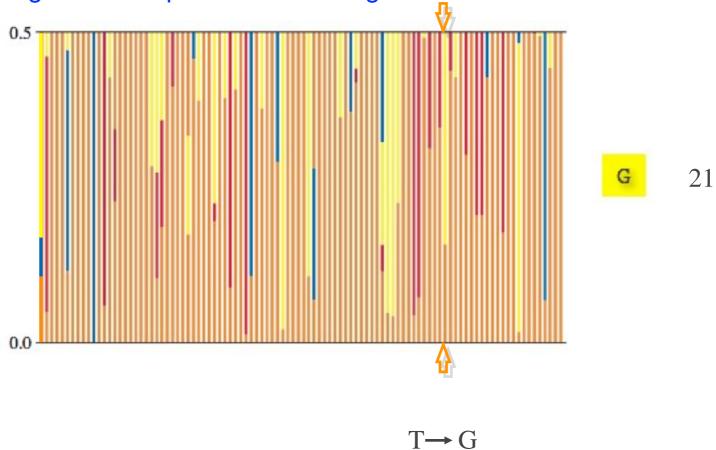
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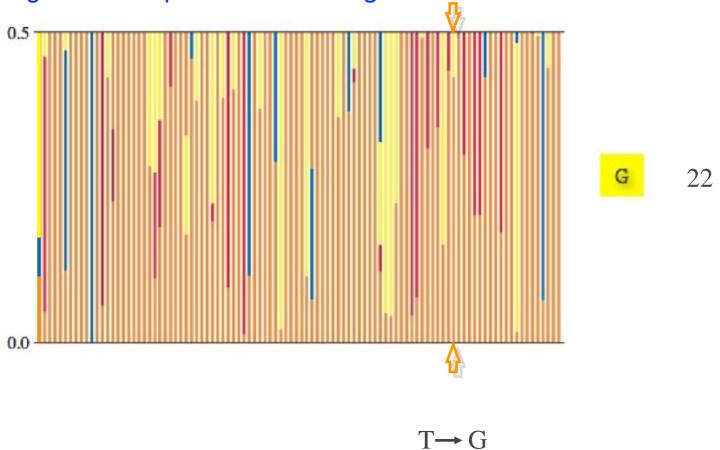


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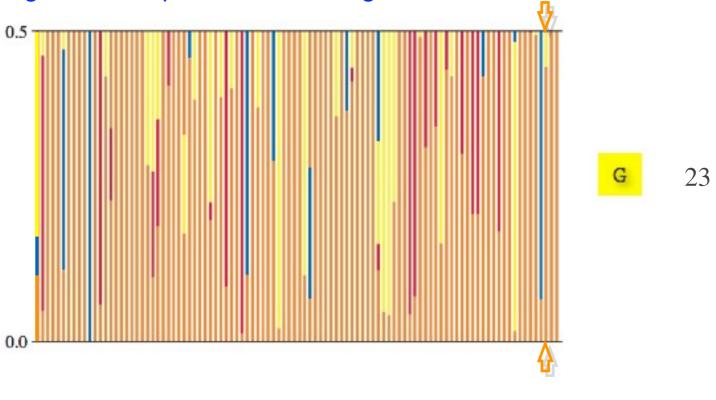
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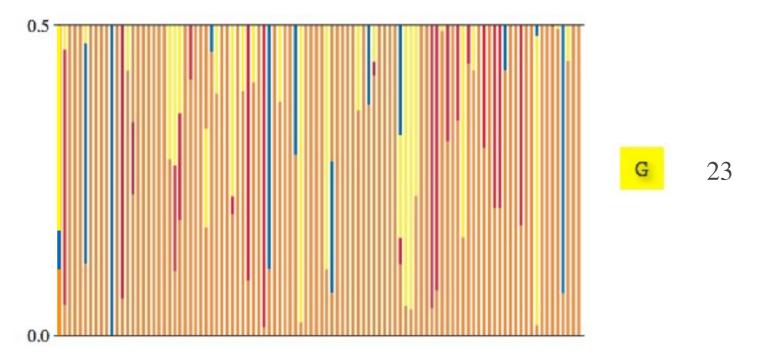


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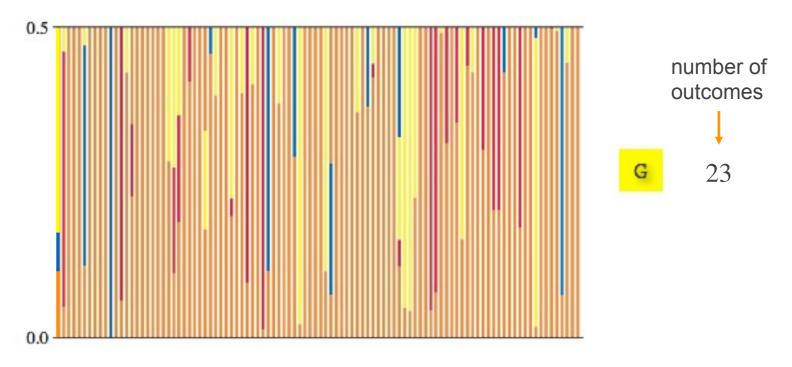


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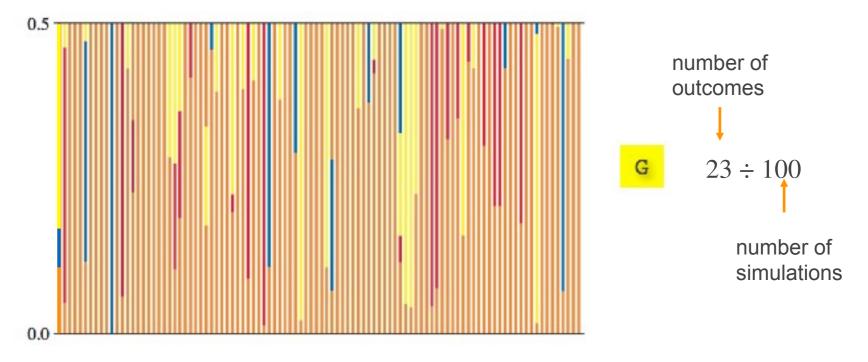
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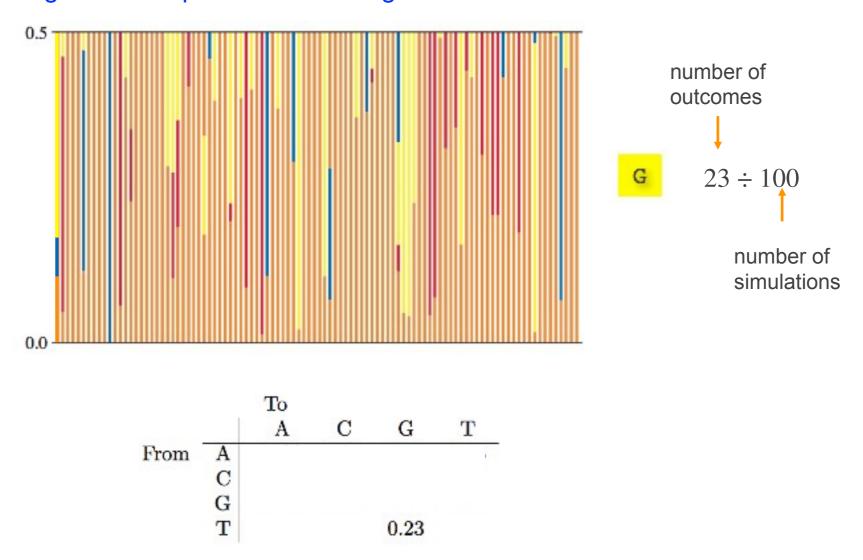
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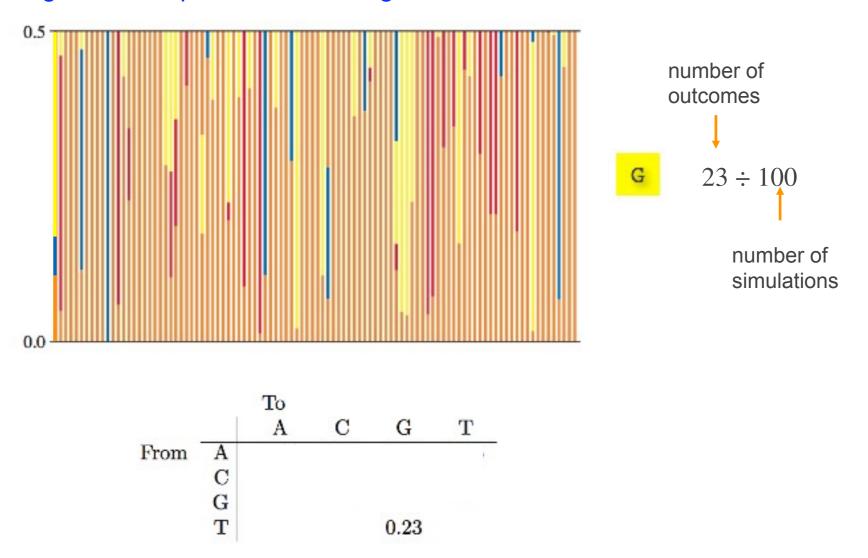
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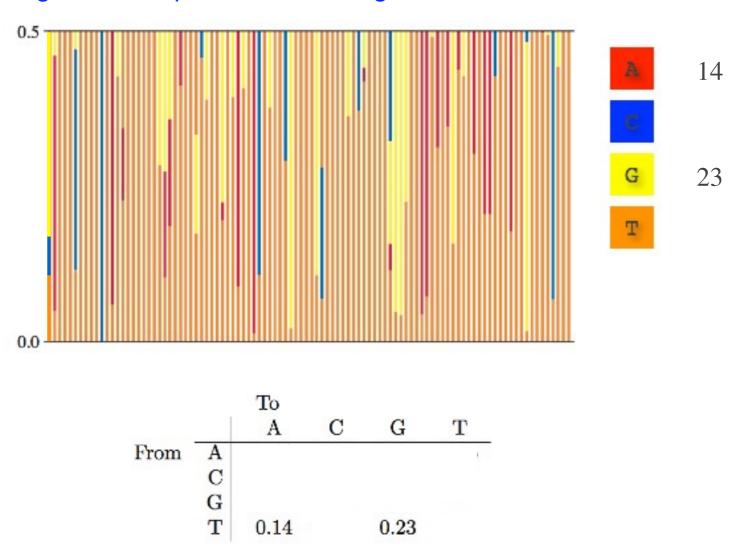


Estimating transition probabilities using Monte Carlo simulation



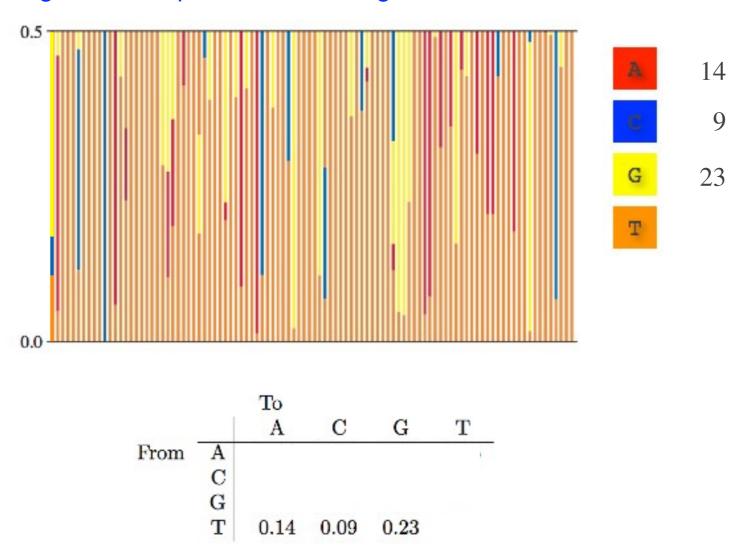
This 'transition probability' reflects all possible histories that start in T and end in G (*i.e.*, histories with different numbers and/or positions of changes)

Estimating transition probabilities using Monte Carlo simulation



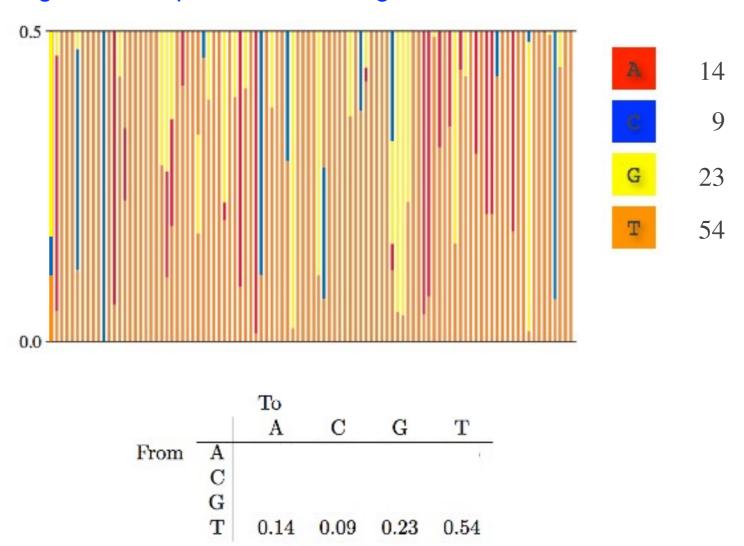
We can repeat this process for the other end states...

Estimating transition probabilities using Monte Carlo simulation



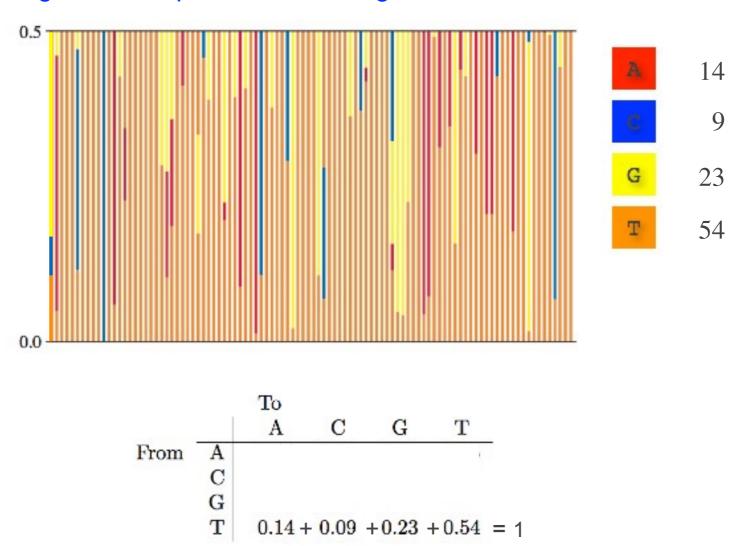
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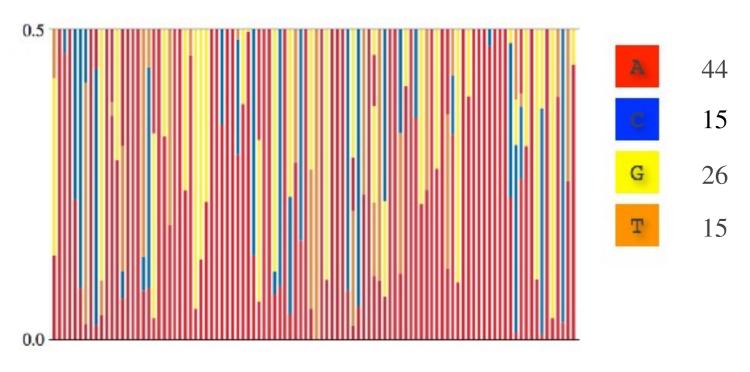
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Estimating transition probabilities using Monte Carlo simulation



Note that each row of the transition-probability matrix sums to 1 (*c.f.*, the Law of Total Probability).

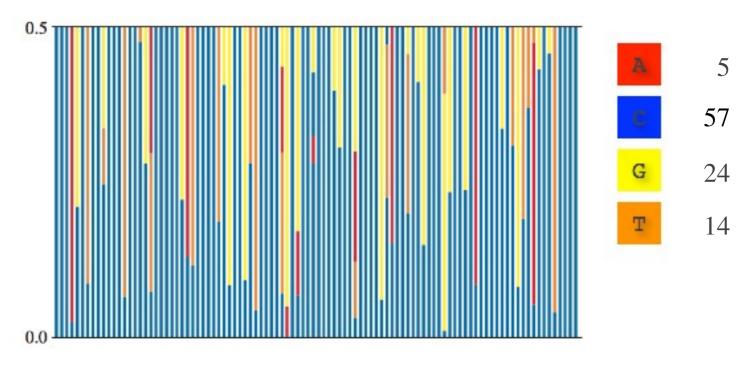
Realizations of 100 replicate simulations starting in state A



		To			
		A	C	\mathbf{G}	\mathbf{T}
From	A	0.44	0.15	0.26	0.15
	C				
	G				
	T	0.14	0.09	0.23	0.54

We can perform new Monte Carlo simulations that start in A to fill out the corresponding row of the transition-probability matrix.

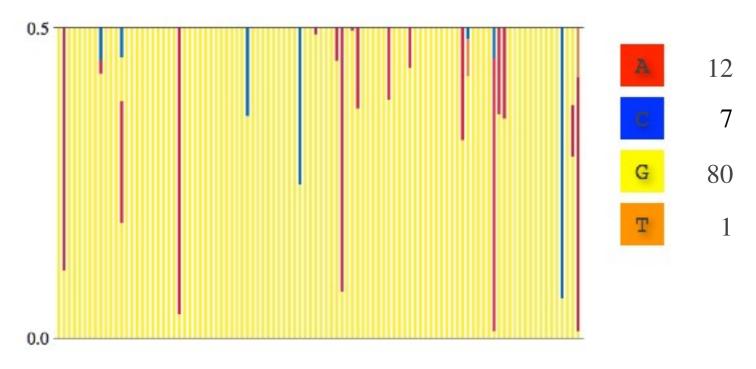
Realizations of 100 replicate simulations starting in state C



		To				
		A	C	\mathbf{G}	\mathbf{T}	
From	A	0.44	0.15	0.26	0.15	•
	C	0.05	0.57	0.24	0.14	
	G					
	T	0.14	0.09	0.23	0.54	

And then for simulations that start in C ...

Realizations of 100 replicate simulations starting in state G



		To			
		A	C	\mathbf{G}	\mathbf{T}
From	A	0.44	0.15	0.26	0.15
	C	0.05	0.57	0.24	0.14
	G	0.12	0.07	0.80	0.01
	T	0.14	0.09	0.23	0.54

And finally for simulations that start in G.

Accuracy of Monte Carlo approximation depends on the number of replicates

100 replicates		To			
		A	C	G	\mathbf{T}
From	A	0.44	0.15	0.26	0.15
	C	0.05	0.57	0.24	0.14
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100,000 replica	ites	To			
		A	C	\mathbf{G}	\mathbf{T}
From	A	0.42119	0.15365	0.26361	0.16155
	C	0.06209	0.60811	0.17602	0.15378
	G	0.08834	0.07241	0.77796	0.06129
	\mathbf{T}	0.13534	0.09411	0.22724	0.54331

Analytical solutions for the transition probabilities: matrix exponentiation

Monte Carlo simulation is computationally expensive and unnecessary, as the transition probabilities can be solved 'analytically'

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The transition probability matrix, \mathbf{P} , can be solved by exponentiating the product of the instantaneous-rate matrix, \mathbf{Q} , and the branch length, v: $\mathbf{P}(v) = e^{\mathbf{Q}v}$

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The exact solution for the transition probability matrix for our instantaneous-rate matrix and branch length (0.5) is:

$$\mathbf{P}(\nu) = \{p_{ij}(\nu)\} = \begin{pmatrix} 0.422927 & 0.153118 & 0.263330 & 0.160625 \\ 0.062896 & 0.609068 & 0.175153 & 0.152883 \\ 0.087566 & 0.071950 & 0.778271 & 0.062212 \\ 0.134967 & 0.093601 & 0.226962 & 0.544470 \end{pmatrix}$$

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Compare with approximate solution (based on 100,000 replicates)

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		A	C	\mathbf{G}	\mathbf{T}
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SIAM REVIEW Vol. 45, No. 1, pp. 3-000 © 2003 Society for Industrial and Applied Mathematics

Nineteen Dubious Ways to Compute the Exponential of a Matrix, Twenty-Five Years Later*

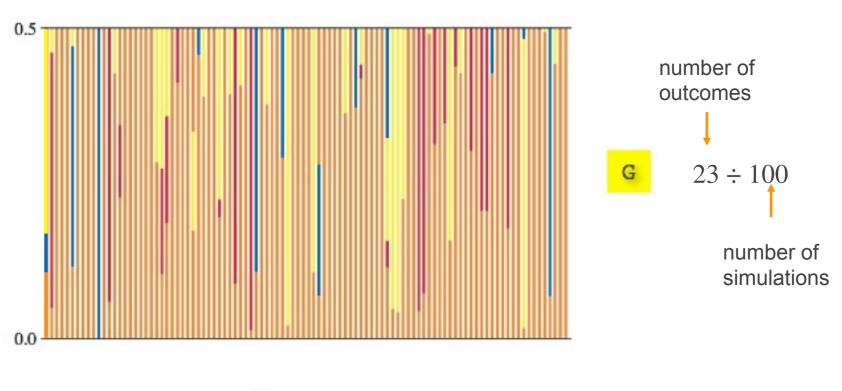
Cleve Moler[†] Charles Van Loan[‡]

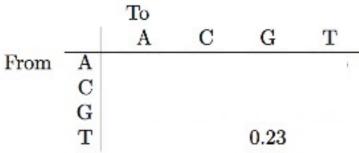
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An aside about transition probabilities

Transition probabilities account for all possible histories that a CTMC can end in a particular state, given a particular starting state (and fully specified model)

Reminder:

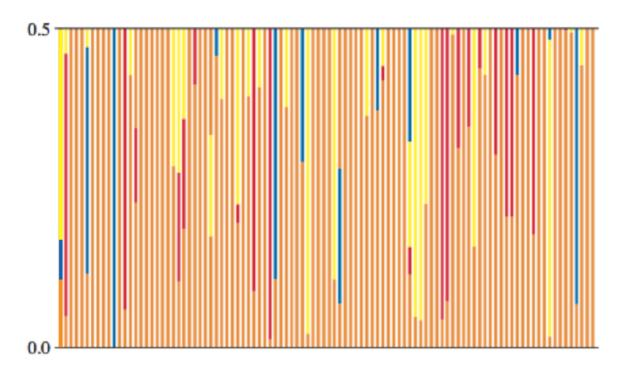




An aside about transition probabilities

Transition probabilities account for all possible histories that a CTMC can end in a particular state, given a particular starting state (and fully specified model)

Transition probabilities play a key role in computing the likelihood, as they avoid the need to condition on a particular history of character change (nucleotide substitution)



Transition probabilities

The probability of observing state j conditioned on starting in state i and running the process over a branch of length v; i.e., $p_{ij}(v)$

Can be estimated by Monte Carlo simulation or matrix exponentiation, $P(v) = e^{Qv}$

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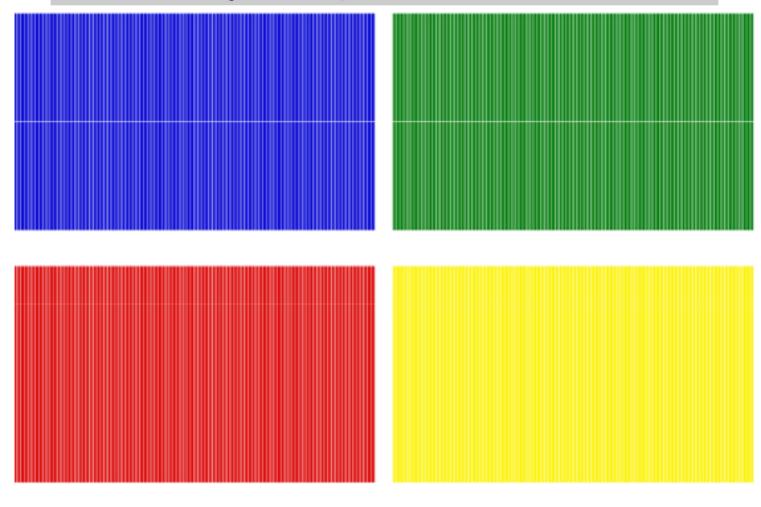
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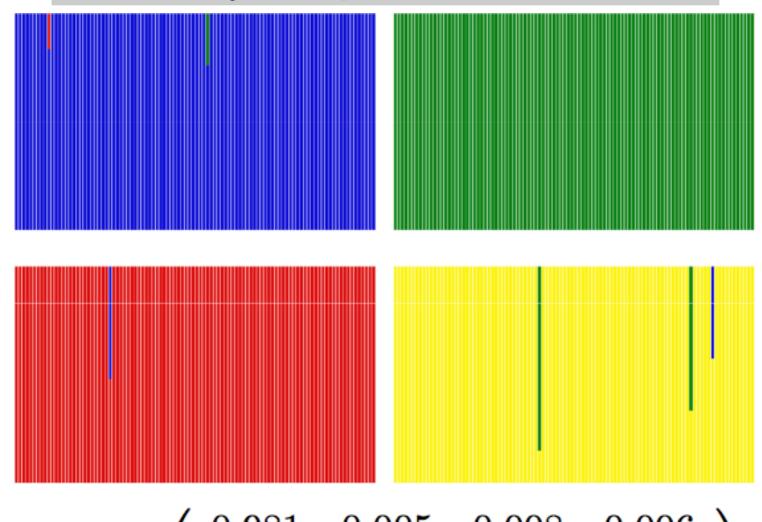
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Transition probabilities over a branch of length v = 0.0:

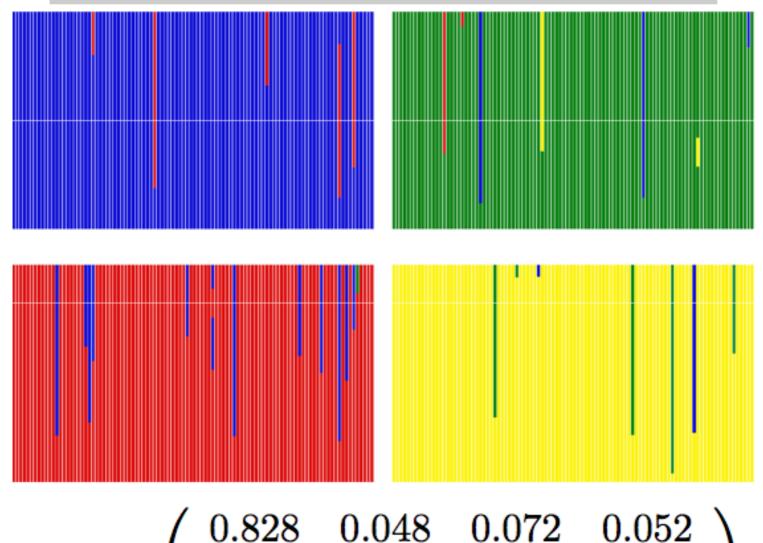
$$\mathbf{P}(0.0) = \begin{pmatrix} 1.000 & 0.000 & 0.000 & 0.000 \\ 0.000 & 1.000 & 0.000 & 0.000 \\ 0.000 & 0.000 & 1.000 & 0.000 \\ 0.000 & 0.000 & 0.000 & 1.000 \end{pmatrix}$$



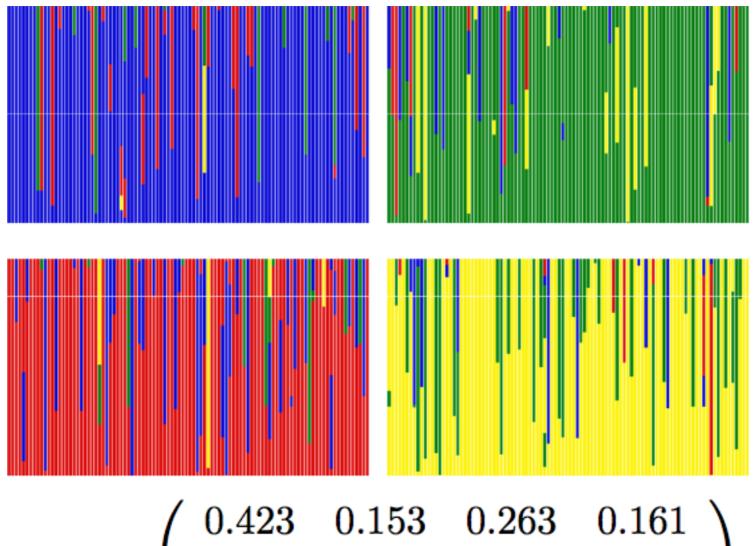
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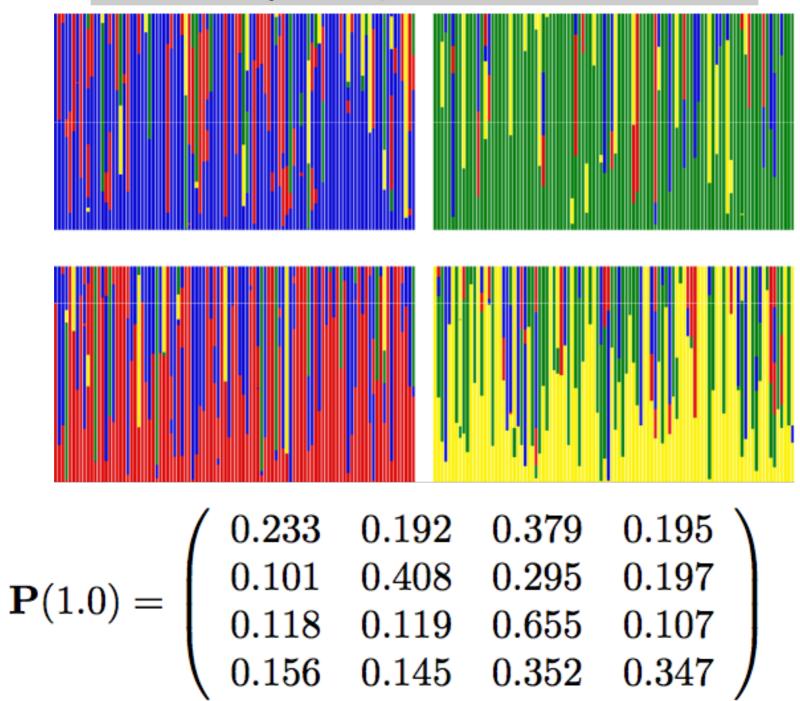
$$\mathbf{P}(0.01) = \begin{pmatrix} 0.981 & 0.005 & 0.008 & 0.006 \\ 0.001 & 0.989 & 0.004 & 0.005 \\ 0.003 & 0.002 & 0.994 & 0.001 \\ 0.005 & 0.002 & 0.006 & 0.986 \end{pmatrix}$$

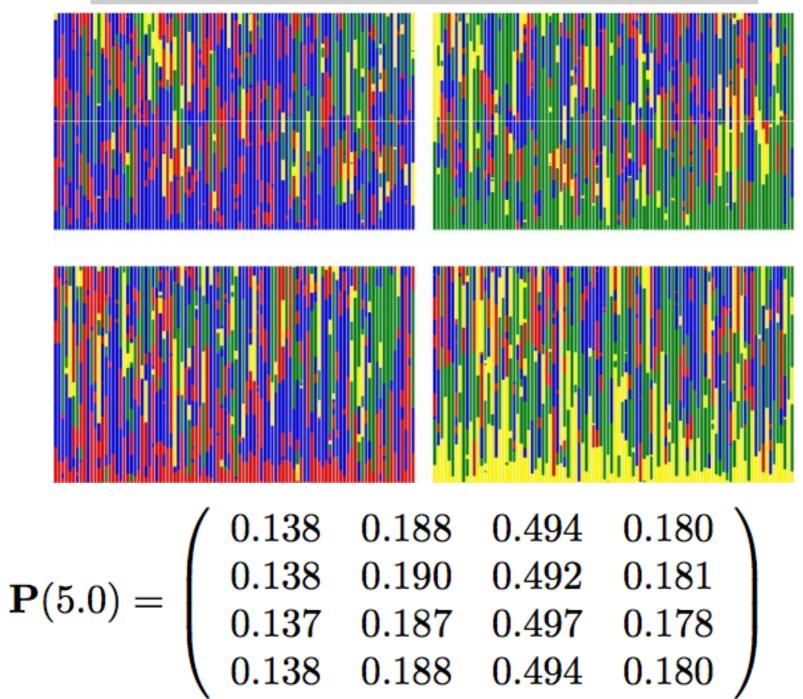


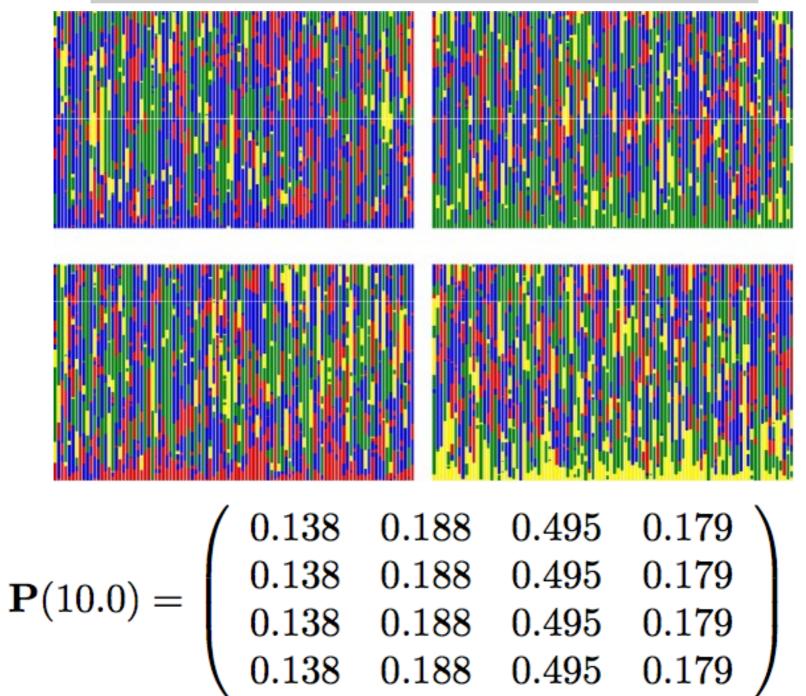
$$\mathbf{P}(0.10) = \begin{pmatrix} 0.028 & 0.048 & 0.042 & 0.032 \\ 0.014 & 0.900 & 0.040 & 0.046 \\ 0.026 & 0.017 & 0.944 & 0.013 \\ 0.046 & 0.023 & 0.056 & 0.876 \end{pmatrix}$$

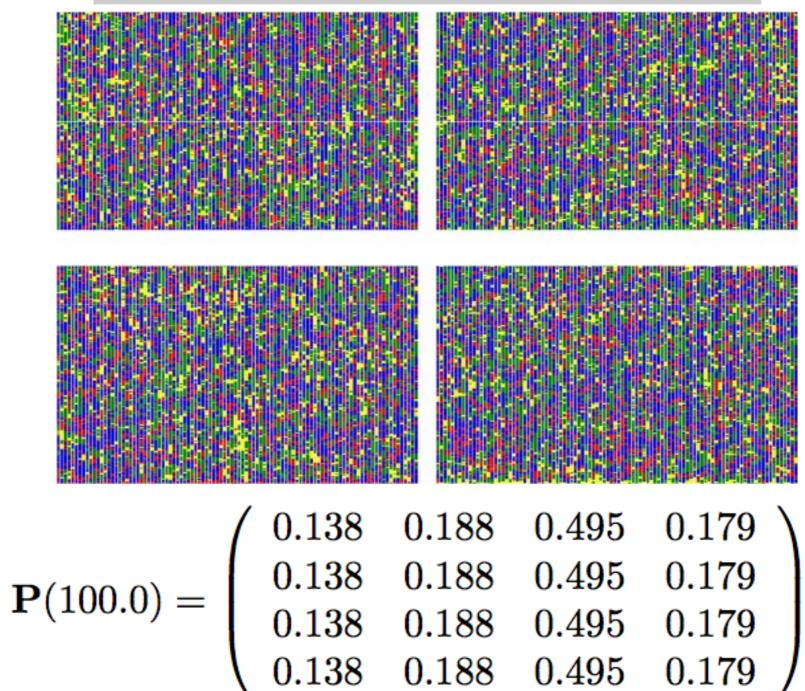


$$\mathbf{P}(0.50) = \begin{pmatrix} 0.425 & 0.165 & 0.265 & 0.161 \\ 0.063 & 0.609 & 0.175 & 0.153 \\ 0.088 & 0.072 & 0.778 & 0.062 \\ 0.135 & 0.094 & 0.227 & 0.544 \end{pmatrix}$$









Stationary Frequencies of a CTMC

Stationary frequencies

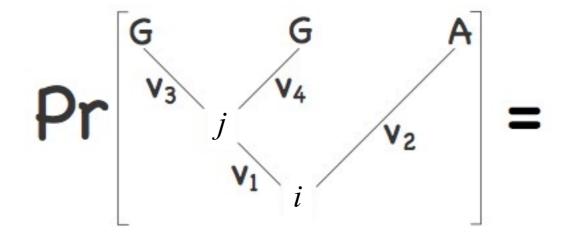
The probability of observing the process in a particular state *j* after a long (infinite) period of time

Stationary frequencies

The probability of observing the process in a particular state *j* after a long (infinite) period of time

When the continuous time Markov chain is at stationarity, the stochastic process has 'forgotten' the starting state: the process ends in a given state with the same probability, regardless of the starting state

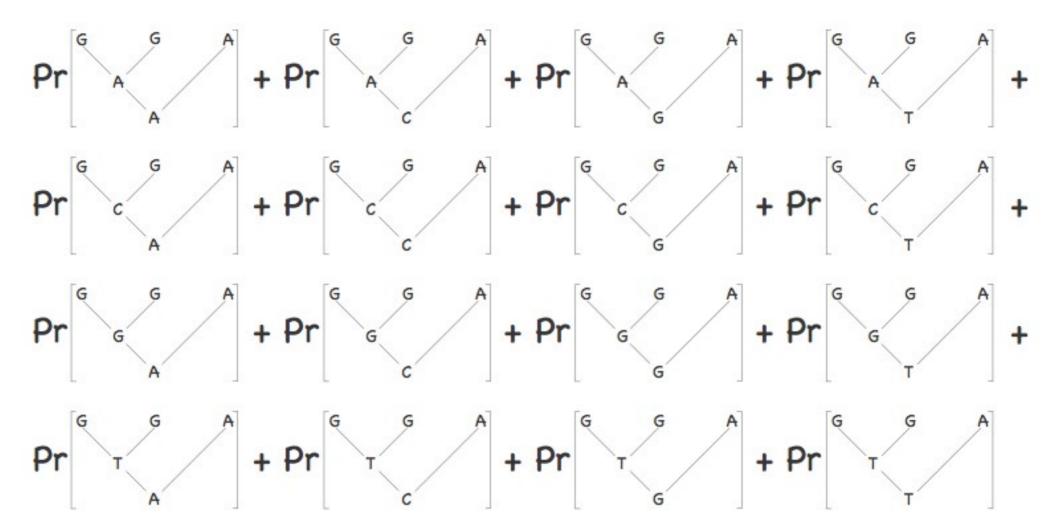
Now we can compute the likelihood of a site history



$$\pi_i \times p_{ij}(v_1) \times p_{iA}(v_2) \times p_{jG}(v_3) \times p_{jG}(v_4)$$

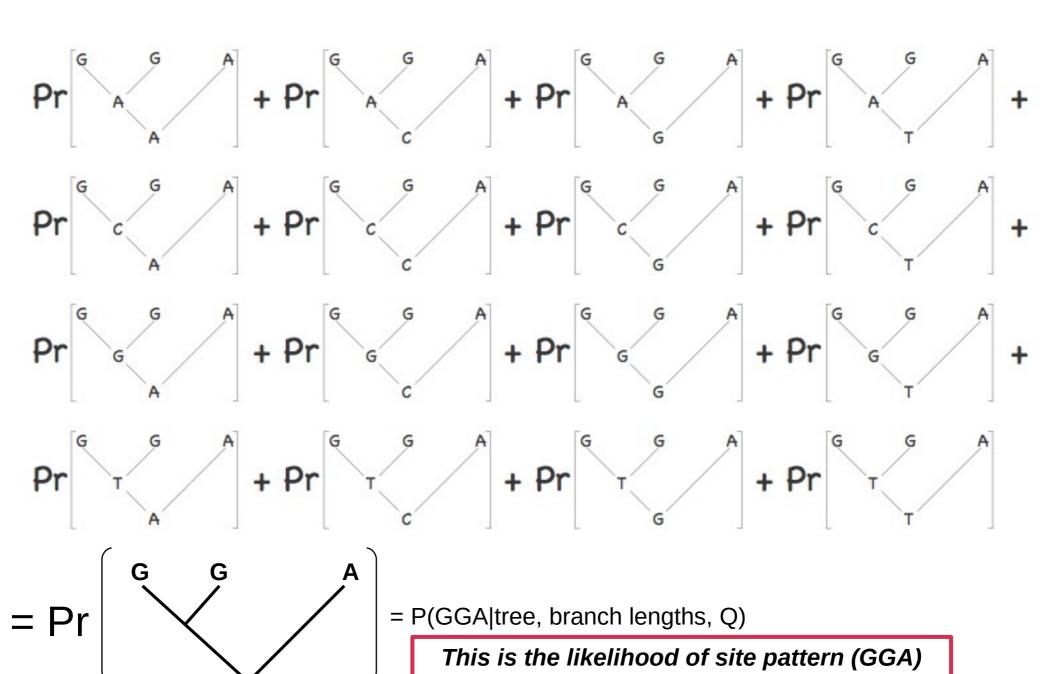
 π_i Stationary frequencies

 $p_{ij}(v)$ Transition probabilities



$$Pr\begin{bmatrix} G & G & A \\ A & A \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ A & C \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ G & A \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\$$

$$Pr \begin{bmatrix} G & G & A \\ A & A \end{bmatrix} + Pr \begin{bmatrix} G & G & A \\ A & C \end{bmatrix} + Pr \begin{bmatrix} G & G & A \\ A & C \end{bmatrix} + Pr \begin{bmatrix} G & G & A \\ A & C \end{bmatrix} + Pr \begin{bmatrix} G & G & A \\ G & C & A \end{bmatrix} + Pr \begin{bmatrix} G & G & A \\ G & C & C \\ G & C & C & C \end{bmatrix} + Pr \begin{bmatrix} G & G & A \\ G & C & C \\ G & C & C \\ G & C & C & C$$



Models of Character Change

Continuous-time Markov models

describe the stochastic process by which traits (nucleotides) evolve over the tree

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Instantaneous-rate matrix, Q

completely describes the stochastic process by specifying:

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Continuous-time Markov models

describe the stochastic process by which traits (nucleotides) evolve over the tree

Instantaneous-rate matrix, Q

completely describes the stochastic process by specifying:

Transition probabilities: $p_{ij}(v)$, the probability of observing state j conditioned on starting in state i and running the process over a branch of length v can be estimated by Monte Carlo simulation or matrix exponentiation, $P(v) = e^{\mathbf{Q}v}$

Stationary frequencies: the long-term probability of observing the chain in state *j*

Instantaneous-rate matrix, Q, and the transition probability matrix, P

The instantaneous rate matrix describes the probability of change between each state in an infinitesimal time interval, $q_{ij}(\partial t)$

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

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The transition probability matrix, $P(v) = \{p_{ij}(v)\}$, describes the probability of observing state j given that we started in state i and ran the process over a branch of length v

$$\mathbf{P}(\nu) = \{p_{ij}(\nu)\} = \begin{pmatrix} 0.422927 & 0.153118 & 0.263330 & 0.160625 \\ 0.062896 & 0.609068 & 0.175153 & 0.152883 \\ 0.087566 & 0.071950 & 0.778271 & 0.062212 \\ 0.134967 & 0.093601 & 0.226962 & 0.544470 \end{pmatrix}$$

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The transiti state *j* gi

Are we clear on what are:

- Substitution rates
- Substitution probabilities
- Stationary frequencies

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Aims and outline

Understand the main ideas underlying models of sequence evolution

- To do so, we will:
 - Introduce important probability notions
 - Play with models of character evolution through simulations
- Briefly present some of the main models of nucleotide evolution

Time-reversible substitution models

Rate matrix

$$\mathbf{Q} = q_{ij} = \begin{pmatrix} - & \mu a \pi_C & \mu b \pi_G & \mu c \pi_T \\ \mu a \pi_A & - & \mu d \pi_G & \mu e \pi_T \\ \mu b \pi_A & \mu d \pi_C & - & \mu f \pi_T \\ \mu c \pi_A & \mu e \pi_C & \mu f \pi_G & - \end{pmatrix}$$

Substitution models

Rate matrix

$$\mathbf{Q} = q_{ij} = \begin{pmatrix} - & \mu a \pi_C & \mu b \pi_G & \mu c \pi_T \\ \mu a \pi_A & - & \mu d \pi_G & \mu e \pi_T \\ \mu b \pi_A & \mu d \pi_C & - & \mu f \pi_T \\ \mu c \pi_A & \mu e \pi_C & \mu f \pi_G & - \end{pmatrix}$$

Jukes and Cantor 1969 Kimura 1980

Hasegawa, Kishino, Yano 1985

$$\mathbf{Q} = q_{ij} = \begin{pmatrix} - & \mu 1/4 & \mu 1/4 & \mu 1/4 \\ \mu 1/4 & - & \mu 1/4 & \mu 1/4 \\ \mu 1/4 & \mu 1/4 & - & \mu 1/4 \\ \mu 1/4 & \mu 1/4 & \mu 1/4 & - \end{pmatrix} \quad \mathbf{Q} = q_{ij} = \begin{pmatrix} - & \mu 1/4 & \mu \kappa 1/4 & \mu 1/4 \\ \mu 1/4 & - & \mu 1/4 & \mu \kappa 1/4 \\ \mu 1/4 & \mu 1/4 & - & \mu 1/4 \\ \mu 1/4 & \mu \kappa 1/4 & \mu 1/4 & - \end{pmatrix} \quad \mathbf{Q} = q_{ij} = \begin{pmatrix} - & \mu \pi_C & \mu \kappa \pi_G & \mu \pi_T \\ \mu \kappa 1/4 & \mu 1/4 & - & \mu 1/4 \\ \mu 1/4 & \mu \kappa 1/4 & \mu 1/4 & - \end{pmatrix} \quad \mathbf{Q} = q_{ij} = \begin{pmatrix} - & \mu \pi_C & \mu \kappa \pi_G & \mu \pi_T \\ \mu \pi_A & - & \mu \pi_G & \mu \kappa \pi_T \\ \mu \kappa \pi_A & \mu \pi_C & - & \mu \pi_T \\ \mu \pi_A & \mu \kappa \pi_G & \mu \pi_G & - \end{pmatrix}$$

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0 free parameter (1 if we do not impose one substitution per unit time)

1 transition/transversion ratio: 1 free parameter

1 transition/transversion ratio 4 equilibrium frequencies:

4 free parameters

Substitution models

Rate matrix

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General Time Reversible model of substitution

Rate matrix

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Lanave et al. 1984; Tavaré, 1986

4 **stationary frequencies**: 3 parameters

6 *exchangeability parameters*: 5 parameters (if we impose one substitution per unit time)

Summary on CTMCs

- We use CTMCs to model character evolution
- Given an instantaneous rate matrix, we can compute substitution probabilities on a branch with an arbitrary length
- We can combine these computations to compute the likelihood of a site history
- We can sum over site histories to get the likelihood of a site pattern
- The GTR family provides examples of reversible instantaneous rate matrices