Introduction to modelling sequence evolution

Bastien Boussau Bastien.boussau@univ-lyon1.fr @bastounette







Who am I?

- CNRS researcher at LBBE in Lyon, France
- Interested in methods for sequence and genome evolution, and in their application
- Keywords: gene tree-species tree, phylogenetic reconstruction site- and branch-heterogeneous models of sequence evolution, genomephenotype associations

This course

A lot of the good stuff was borrowed from Brian

Moore's slides



(http://phylolab.org/)

The bad stuff is mine

Aims

Understand the main ideas underlying models of sequence evolution

- To do so, we will:
 - Introduce important probability notions
 - Play with models of character evolution through simulations
- Briefly present some of the main models of nucleotide evolution

Generic statistical paradigm

- Question about some part of the world
- Model of how this part of the world works
- Collect data
- Estimate
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Example

Is my coin fair?

- Repeated throws=independent identically distributed Bernoulli draws
- Throw coin N times
- Estimate probability of heads

Generic statistical paradigm

Phylogeny example

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- Sites of alignment=independent identically distributed Markov chains running along a phylogeny
- Sequence rodents
- Estimate transition/transversion ratio

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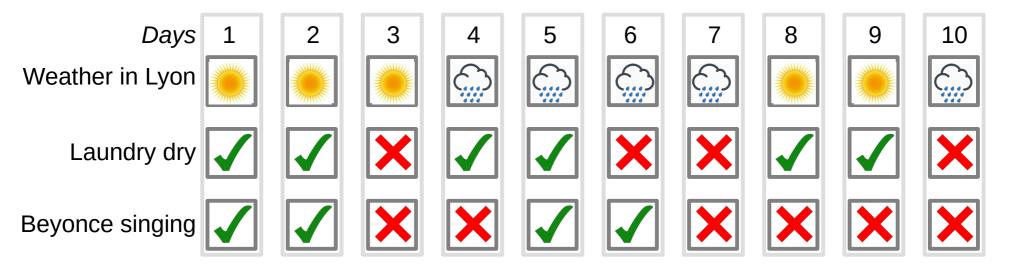
Why are we interested in simulations?

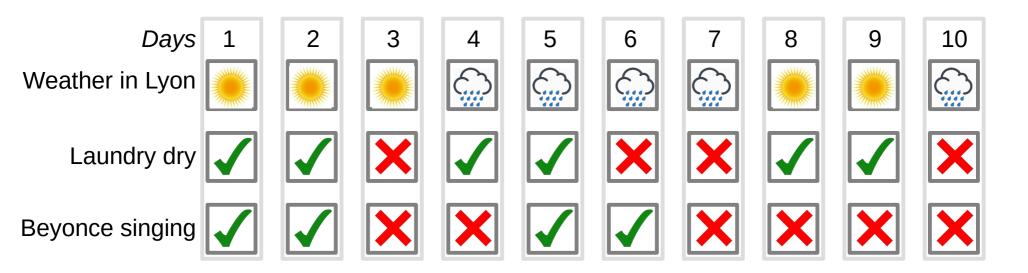
- Simulating data forces us to think in terms of a generating process
- By comparing true to simulated data, we can get a sense of how realistic our model is
- Simulations are also central to a lot of inferential problems:
 - Validation of inference methods
 - Posterior predictive tests
 - Approximate Bayesian Computation (ABC)

- ...

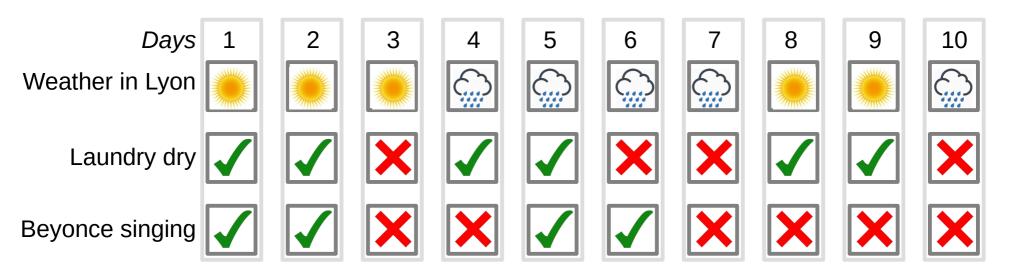
Useful probability concepts

- Conditional probabilities
- Independence/intersection
- Union
- Bayes theorem
- Common distributions that will be useful in this talk:
 - Bernoulli
 - Binomial
 - Poisson
 - Exponential

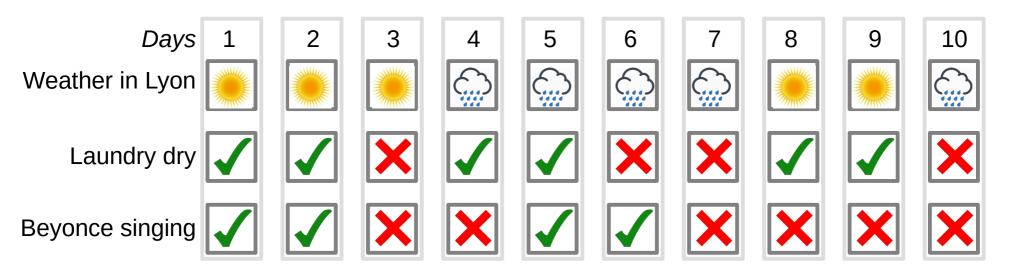




$$P(rainy) = ?$$



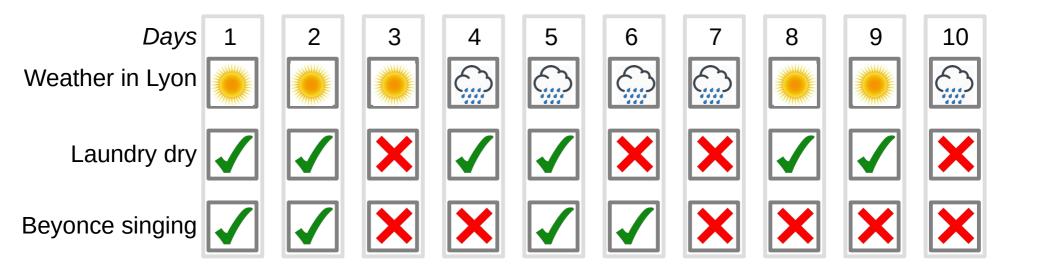
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 $P(sunny)=1-P(rainy)=0.5$



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$$P(dry laundry) = 0.6$$

Record of various events during 10 days

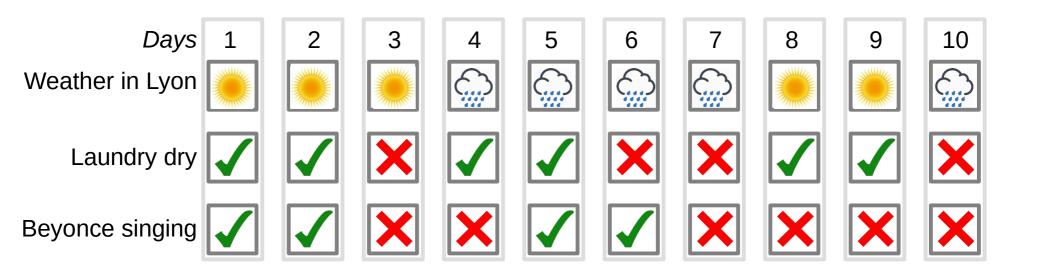


$$P(rainy)=0.5$$
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$$P(dry laundry) = 0.6$$

P(dry | laundry | sunny) = ?

P(dry | laundry | rainy) = ?



$$P(rainy)=0.5$$
 $P(sunny)=1-P(rainy)=0.5$ $P(dry laundry)=0.6$ $P(dry laundry | sunny)=0.8$ Conditional probability: $P(A|B)$

$$P(dry | aundry | sum y) = 0.0$$

 $P(dry | aundry | rainy) = 0.4$

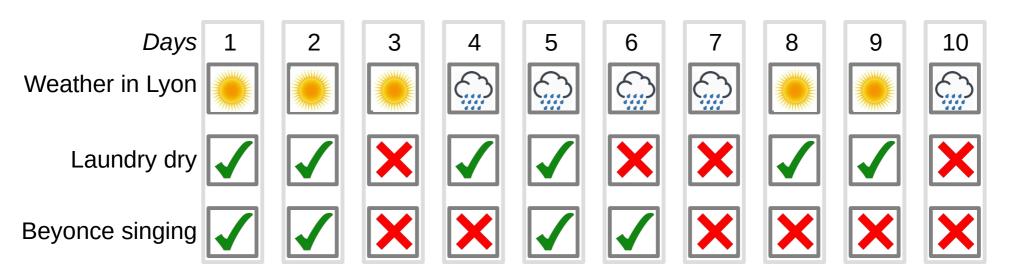
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 $P(dry laundry | rainy)=0.4$

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P(Beyonce singing) = 0.4P(Beyonce singing) = P(Beyonce singing | rainy) = P(Beyonce singing | sunny) = 0.4

Record of various events during 10 days



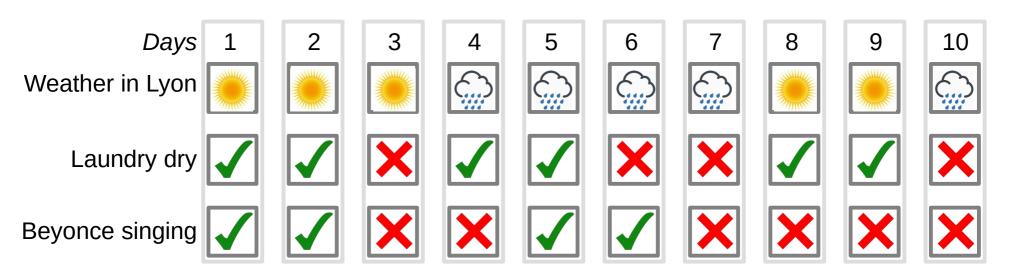
$$P(rainy)=0.5$$
 $P(sunny)=1-P(rainy)=0.5$

The events "Beyonce singing" and "sunny" are independent

P(Beyonce singing) = 0.4

 $P(Beyonce\ singing) = P(Beyonce\ singing|rainy) = P(Beyonce\ singing|sunny)^{20} = 0.4$

Record of various events during 10 days



$$P(rainy)=0.5$$
 $P(sunny)=1-P(rainy)=0.5$

$$P(dry laundry) = 0.6$$

$$P(dry | laundry | sunny) = 0.8$$

$$P(dry | laundry | rainy) = 0.4$$

The events "dry laundry" and "sunny" are NOT independent

$$P(rainy)=0.5$$
 $P(sunny)=1-P(rainy)=0.5$ $P(dry laundry)=0.6$ $P(dry laundry | sunny)=0.8$ $P(dry laundry | rainy)=0.4$

$$P(dry | laundry) = P(dry | laundry | sunny) \times P(sunny)$$

+ $P(dry | laundry | rainy) \times P(rainy)$
= $0.8 \times 0.5 + 0.4 \times 0.5 = 0.6$

Bayes formula

$$P(A|B) = \frac{P(A \land B)}{P(B)} = \frac{P(B \land A)}{P(B)}$$
$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

$$P(rainy)=0.5$$
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$$P(sunny|dry | laundry) = \frac{P(sunny \land dry | laundry)}{P(dry | laundry)} = \frac{P(dry | laundry \land sunny)}{P(dry | laundry)}$$

$$P(sunny | dry | laundry) = \frac{P(dry | laundry | sunny) P(sunny)}{P(dry | laundry)}$$

Useful distributions

- *Discrete distributions (values in {0,1}, {0,1,2...}):*
 - Bernoulli: coin flip: P(X=1)=p; P(X=0)=1-p
 - Binomial: how many heads in several coin flips:

$$Pr(k;n,p) = \Pr(X=k) = inom{n}{k} p^k (1-p)^{n-k}$$

 Poisson: how many events of a type over a continuous time: how many meteorites with diameter > 1m in a year:

$$P(k \text{ events in interval}) = e^{-\lambda} \frac{\lambda^{\kappa}}{k!}$$

- Continuous distributions (values in \mathbb{R} , [0,1]...):
 - Exponential: Time between events in a Poisson process: how much time between two meteorites with diameter

>1m:
$$f(x;\lambda) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0, \\ 0 & x < 0. \end{cases}$$

Waiting times in a Poisson process

Some rare, discrete event that occurs at a constant rate in continuous time is described by a Poisson process



Siméon Poisson (1821)

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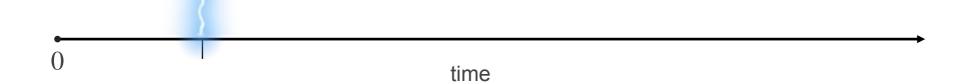
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These events occur with a rate λ

0 time

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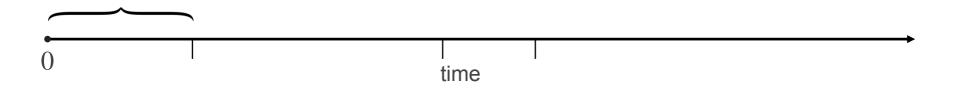


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The waiting (sojourn) time for the first event

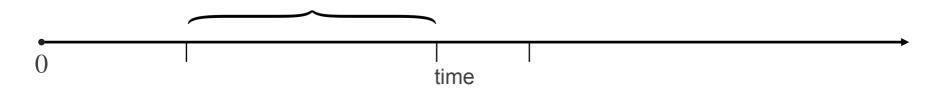


Waiting times in a Poisson process

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The waiting (sojourn) time for the second event

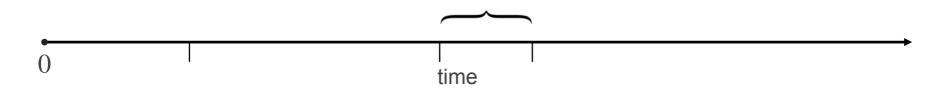


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The waiting (sojourn) time for the third event

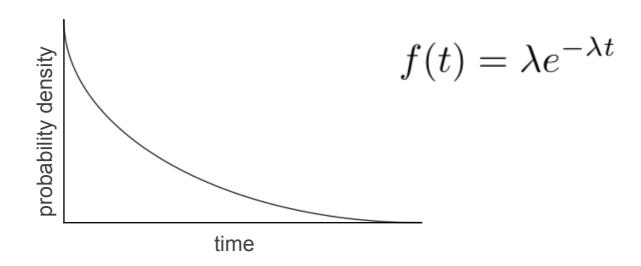


Waiting times in a Poisson process

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The waiting (sojourn) times are exponentially distributed random variables



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Understand the main ideas underlying models of sequence evolution

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Stochastic Models of Nucleotide Substitution

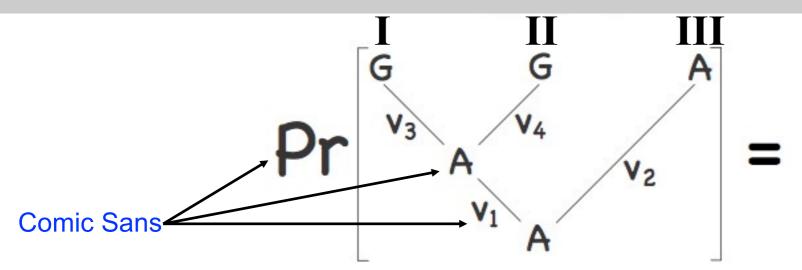
Species Sequ

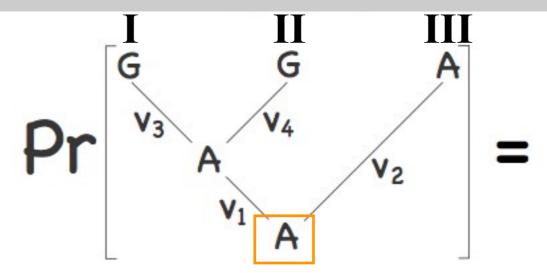
Species II Species III

Sequence data

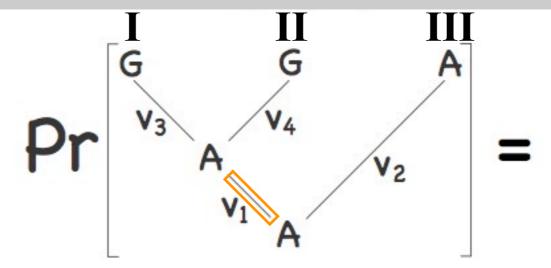
```
GCG--CACCGGCGCAGTCA....
GCGTTCA--GGCG--GTCA....
ACGTTCACCGGCGCAGTCA....
```

Stochastic Models of Nucleotide Substitution

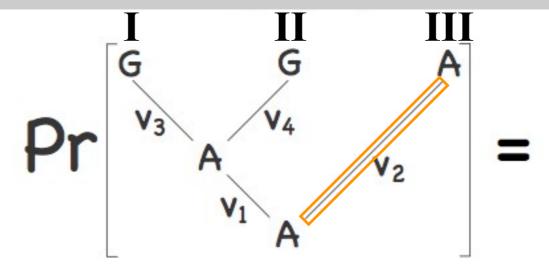




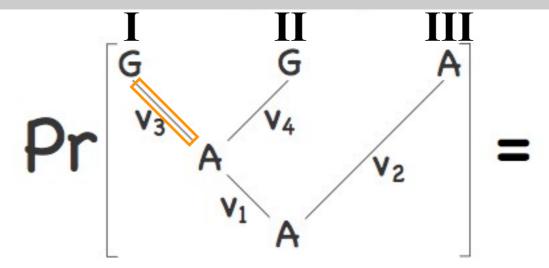
 π_A



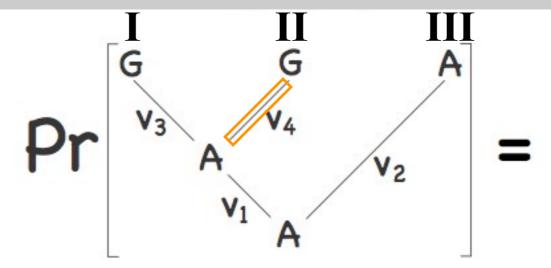
$$\pi_A \times p_{AA}(v_1)$$



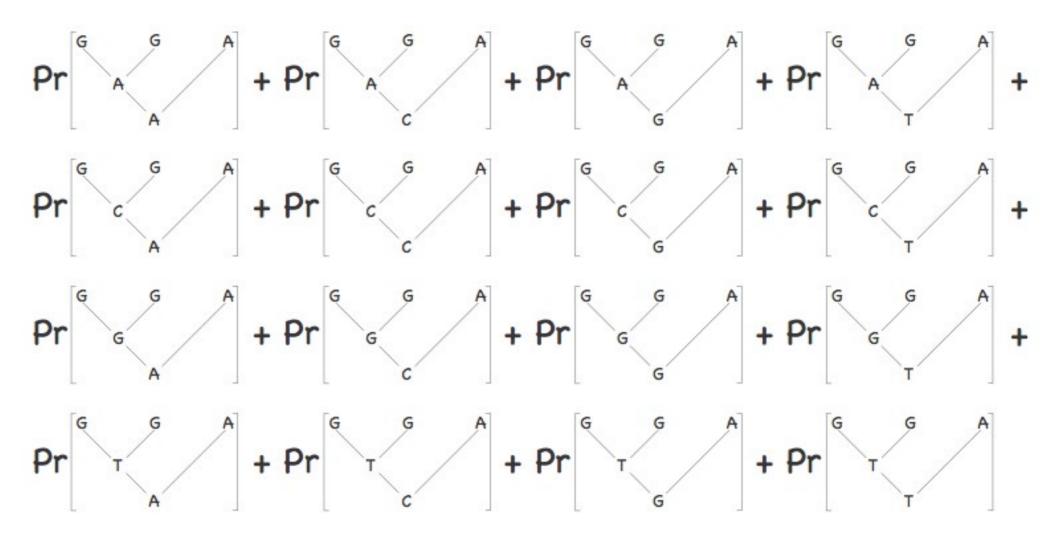
$$\pi_A \times p_{AA}(v_1) \times p_{AA}(v_2)$$

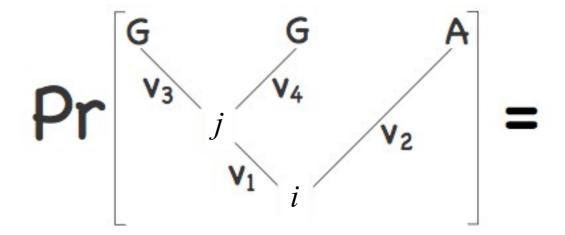


$$\pi_A \times p_{AA}(v_1) \times p_{AA}(v_2) \times p_{AG}(v_3)$$

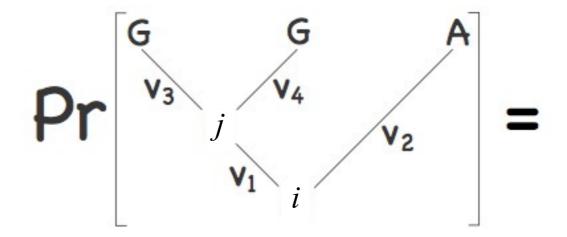


$$\pi_A \times p_{AA}(v_1) \times p_{AA}(v_2) \times p_{AG}(v_3) \times p_{AG}(v_4)$$





$$\pi_i \times p_{ij}(v_1) \times p_{iA}(v_2) \times p_{jG}(v_3) \times p_{jG}(v_4)$$



$$\pi_i \times p_{ij}(v_1) \times p_{iA}(v_2) \times p_{jG}(v_3) \times p_{jG}(v_4)$$

 π_i Stationary frequencies

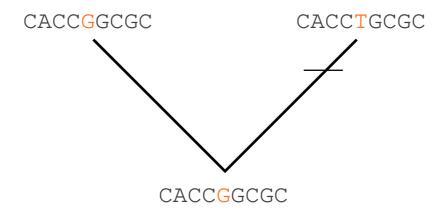
 $p_{ij}(v)$ Transition probabilities

Continuous-time Markov Chains (CTMC)

Evolution of discrete traits (e.g., substitution models, morphological models)

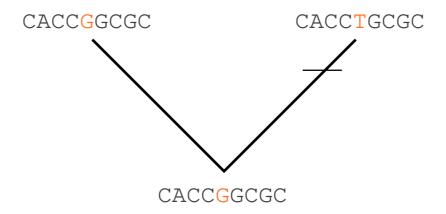
Models describe changes in the nucleotide sites at the species level

We are (generally) not trying to model the dynamics of allele frequencies in populations



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We are modeling the process of **nucleotide substitution**, which describes the outcome of the mutation and fixation processes within populations

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Character change (nucleotide substitution) is modeled as a continuous-time Markov chain (CTMC)

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Stochastic model in which the next state of the chain depends only on the current state

The model is central to model-based inference

Even if the parameters of the substitution model are not of direct interest, they are nevertheless critical to estimation of the focal model parameters

A Continuous-time Markov model is defined by a matrix of substitution rates

A table that specifies the rates of all possible changes between states.

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A hypothetical instantaneous-rate matrix

		To			
		A	C	G	T
From	A	-1.916	0.541	0.787	0.588
	C	0.148	-1.069	0.415	0.506
	G	0.286	0.170	-0.591	0.135
	T	0.525	0.236	0.594	-1.355

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The rows of the table must sum to zero.

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$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

If the current state of the Markov chain is i, the next substitution will occur after an exponentially distributed waiting time with rate parameter $-q_{ii}$

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$$P(C->A) = q_{cA}/-q_{cC} = 0.148 | 1.069 = 0.138$$

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$$P(C->A) = q_{cA}/-q_{cc} = 0.148 | 1.069 = 0.138$$

 $P(C->G) = q_{cG}/-q_{cc} = 0.415 | 1.069 = 0.388$

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$$P(C->A) = q_{ca}/-q_{cc} = 0.148 | 1.069 = 0.138$$

 $P(C->G) = q_{cg}/-q_{cc} = 0.415 | 1.069 = 0.388$
 $P(C->T) = q_{cg}/-q_{cc} = 0.506 | 1.069 = 0.474$

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$$P(C->T) = q_{cT}/-q_{cc} = 0.506 | 1.069 = \underline{0.474}$$

$$\Sigma P_{ij} = 1.0$$

Developing Intuition for CTMCs: A Monte Carlo Simulation Experiment

What the heck is Monte Carlo Simulation?

We generate a number of **replicate outcomes** (we will perform multiple trials)

Developing Intuition for CTMCs: A Monte Carlo Simulation Experiment

What the heck is Monte Carlo Simulation?

We generate a number of **replicate outcomes** (we will perform multiple trials) of a **stochastic process** (our replicate experiments involve randomness)

Developing Intuition for CTMCs: A Monte Carlo Simulation Experiment

What the heck is Monte Carlo Simulation?

We generate a number of **replicate outcomes** (we will perform multiple trials) of a **stochastic process** (our replicate experiments involve randomness) under a **fully specified model** (with specific values for all model parameters)

A simple Monte Carlo Simulation

We will assume that we have a **fully specified phylogenetic model**:

A simple Monte Carlo Simulation

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$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

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A simple Monte Carlo Simulation

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What is the matrix of substitution probabilities over our branch of length 0.5?

A simple Monte Carlo Simulation

We will use a random-number generator to mimic the **stochastic process**

A simple Monte Carlo Simulation

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Rate of leaving the current state, T = 1.355

A simple Monte Carlo Simulation

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Probability of changing to A:

$$P(T->A) = q_{TA}/-q_{TT} = 0.525 | 1.355 = 0.387$$

A simple Monte Carlo Simulation

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Probability of changing to C:

$$P(T->A) = q_{TA}/-q_{TT} = 0.525 | 1.355 = 0.387$$

$$P(T->C) = q_{TC}/-q_{TT} = 0.236 | 1.355 = 0.174$$

A simple Monte Carlo Simulation

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Probability of changing to G:

$$P(T->A) = q_{TA}/-q_{TT} = 0.525 | 1.355 = 0.387$$

 $P(T->C) = q_{TC}/-q_{TT} = 0.236 | 1.355 = 0.174$
 $P(T->G) = q_{TC}/-q_{TT} = 0.594 | 1.355 = 0.438$

A simple Monte Carlo Simulation

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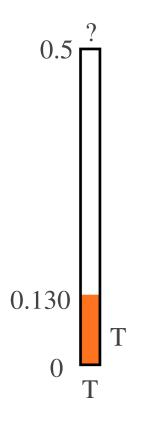
Generate an exponentially distributed waiting time, *x*:

A simple Monte Carlo Simulation

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Generate an exponentially distributed waiting time, x: rate when process is in T: $\lambda = 1.355$

A simple Monte Carlo Simulation



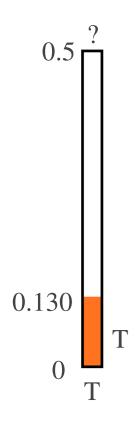
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Generate an exponentially distributed waiting time, *x*:

rate when process is in T: $\lambda = 1.355$

Draw x: $x \sim dnExponential(1.355) : x = 0.130$

A simple Monte Carlo Simulation



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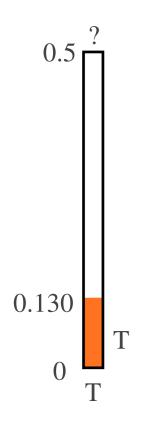
Probabilities of substitution events in state T:

$$P(T->A) = q_{TA}/-q_{TT} = 0.525 | 1.355 = 0.387$$

 $P(T->C) = q_{TC}/-q_{TT} = 0.236 | 1.355 = 0.174$

$$P(T->G) = q_{rr}/-q_{rr} = 0.594 | 1.355 = 0.438$$

A simple Monte Carlo Simulation



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Specify a set of intervals: intervals

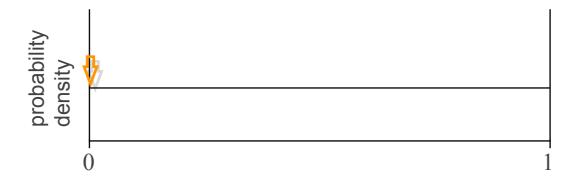
$$P(T->A) = 0.387$$
 $0 - 0.387$ (choose A)

$$P(T->C) = 0.174$$
 0.387 - 0.561 (choose C)

$$P(T->G) = 0.438$$
 $0.561 - 1$ (choose G)

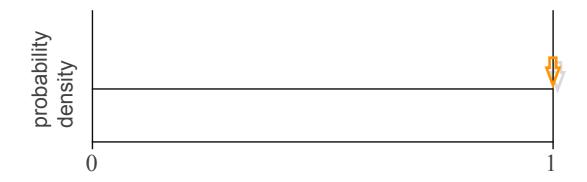
Only the width of the bins matters, not their order

Our uniform(0,1) random number will take any value between 0 and 1 with equal probability (by definition)



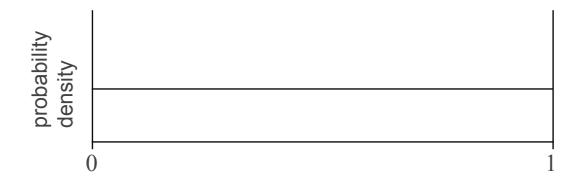
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Imagine that there are two possible outcomes, A and B, which occur with probabilities:

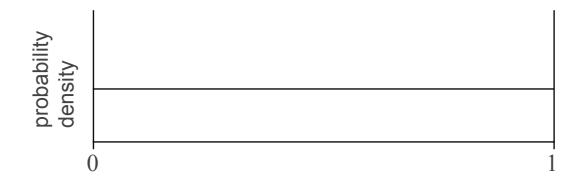
Probability of outcomes:

$$P(\text{option A}) = 0.6$$

$$P(\text{option B}) = 0.4$$

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We can choose randomly (according to their probabilities) by specifying these intervals...

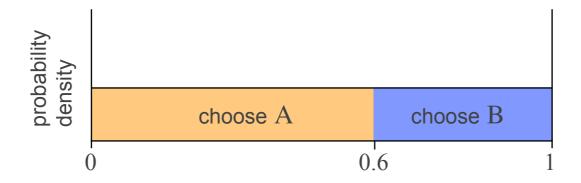
Probability of outcomes: intervals

$$P(\text{option } A) = 0.6$$
 $0.0 - 0.6$ (choose A)

$$P(\text{option B}) = 0.4$$
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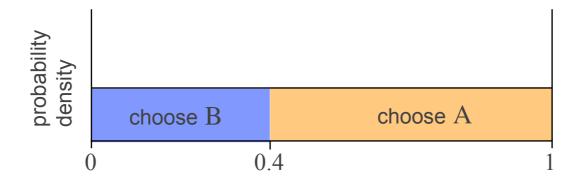
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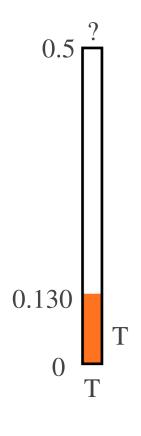
$$P(\text{option } A) = 0.6$$

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Or equivalently by specifying these intervals...

Probability of outcomes: intervals $P(\text{option }A) = 0.6 \qquad 0.4 - 1.0 \quad (\text{choose }A)$ $P(\text{option }B) = 0.4 \qquad 0.0 - 0.6 \quad (\text{choose }B)$

A simple Monte Carlo Simulation



$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

Generate an exponentially distributed waiting time, *x*:

rate when process is in T: $\lambda = 1.355$

Draw x: $x \sim dnExponential(1.355) : x = 0.130$

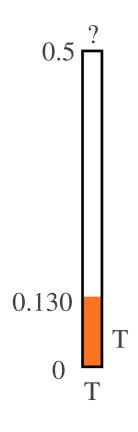
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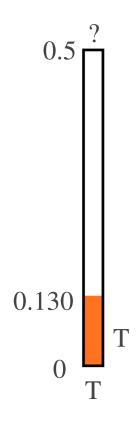
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Draw a uniformly distributed number, u, to select substitution event: $u \sim dnUniform(0,1)$: u=0.446

A simple Monte Carlo Simulation



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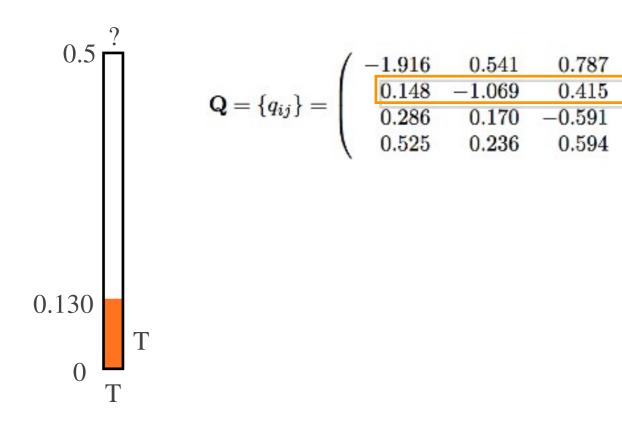
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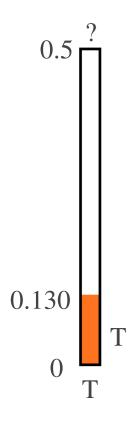
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0.588 0.506 0.135

A simple Monte Carlo Simulation



A simple Monte Carlo Simulation

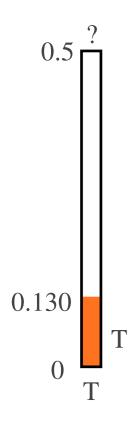


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Generate waiting time to next event:

Rate when process in state C: $-q_{cc} = \lambda = 1.069$

A simple Monte Carlo Simulation



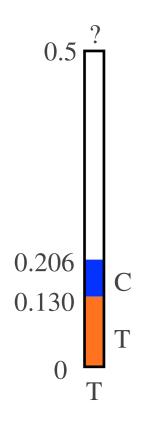
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Generate waiting time to next event:

Rate when process in state C: $-q_{cc} = \lambda = 1.069$

Draw x: $x \sim dnExponential(1.069)$: x = 0.076

A simple Monte Carlo Simulation



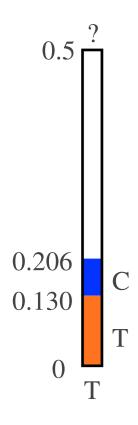
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A simple Monte Carlo Simulation



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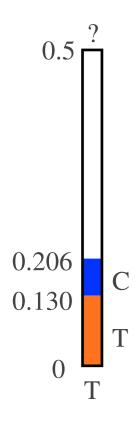
Rate when process in state C: $-q_{cc} = \lambda = 1.069$

Draw x: $x \sim dnExponential(1.069)$: x = 0.076

Substitution probabilities in state C:

$$P(C->A) = q_{cA}/-q_{cC} = 0.148 \mid 1.069 = 0.138$$

A simple Monte Carlo Simulation



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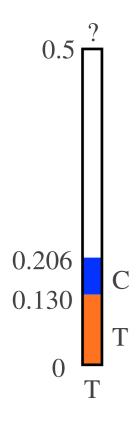
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Substitution probabilities in state C:

$$P(C->A) = q_{cA}/-q_{cC} = 0.148 \mid 1.069 = 0.138$$

$$P(C->G) = q_{cc}/-q_{cc} = 0.415 \mid 1.069 = 0.388$$

A simple Monte Carlo Simulation



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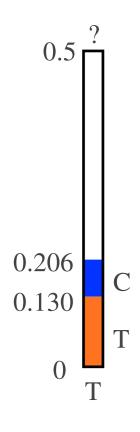
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 $P(C->G) = q_{cc}/-q_{cc} = 0.415 | 1.069 = 0.388$
 $P(C->T) = q_{cc}/-q_{cc} = 0.506 | 1.069 = 0.474$

A simple Monte Carlo Simulation



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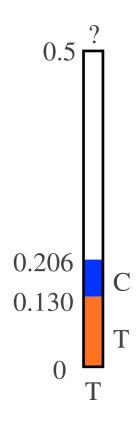
Specify a set of intervals: intervals

$$P(C->A) = 0.138$$
 $0 - 0.138$ (choose A)

$$P(C->G) = 0.388$$
 $0.138 - 0.526$ (choose G)

$$P(C->T) = 0.474$$
 0.526 - 1 (choose T)

A simple Monte Carlo Simulation



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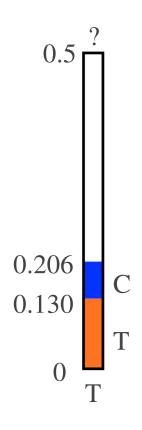
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 $0.526 - 1$ (choose T)

Draw a uniformly distributed number, u, to select substitution event: $u \sim dnUniform(0,1)$: u=0.317

A simple Monte Carlo Simulation



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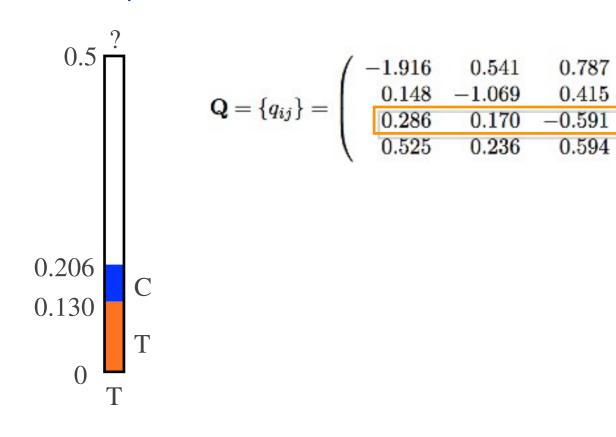
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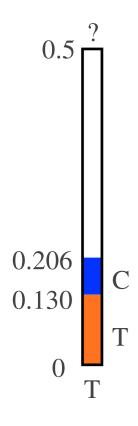
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0.588 0.506 0.135

A simple Monte Carlo Simulation



A simple Monte Carlo Simulation

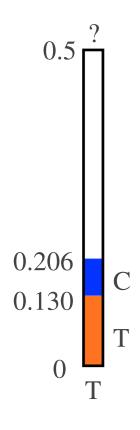


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Generate waiting time to next event:

Rate when process in state G: $-q_{cc} = \lambda = 0.591$

A simple Monte Carlo Simulation



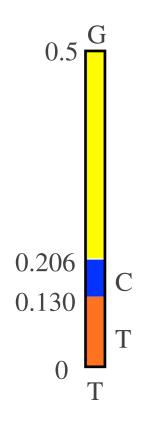
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Generate waiting time to next event:

Rate when process in state G: $-q_{cc} = \lambda = 0.591$

Draw x: $x \sim dnExponential(1.069)$: x = 1.820

A simple Monte Carlo Simulation



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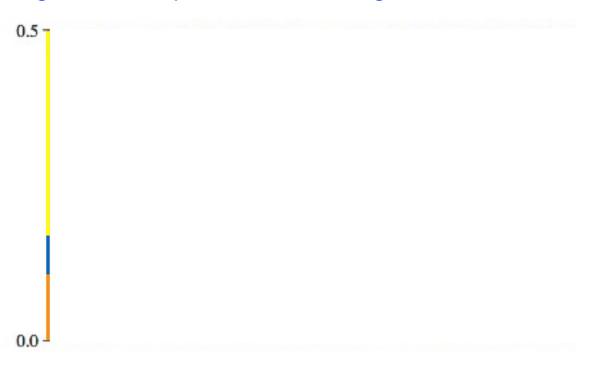
Rate when process in state G: $-q_{cc} = \lambda = 0.591$

Draw x: $x \sim dnExponential(1.069)$: x = 1.820

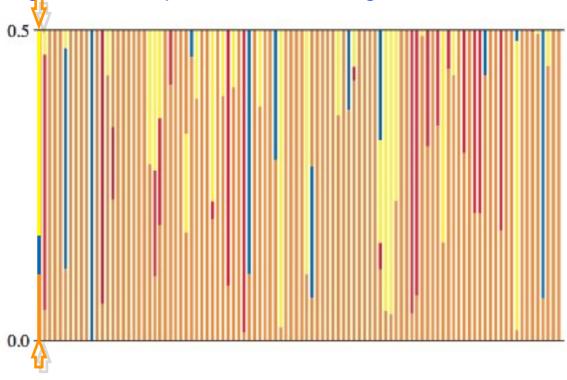
 $1.820 > 0.5-0.206 \rightarrow Terminate simulation (in state G)$

Transition Probabilities of a CTMC

Estimating transition probabilities using Monte Carlo simulation

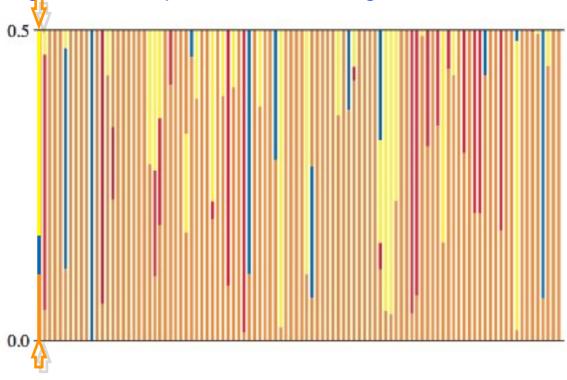


Estimating transition probabilities using Monte Carlo simulation



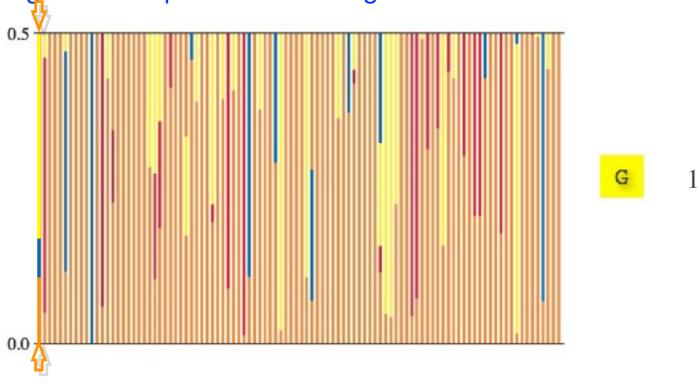
 $T \rightarrow C \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



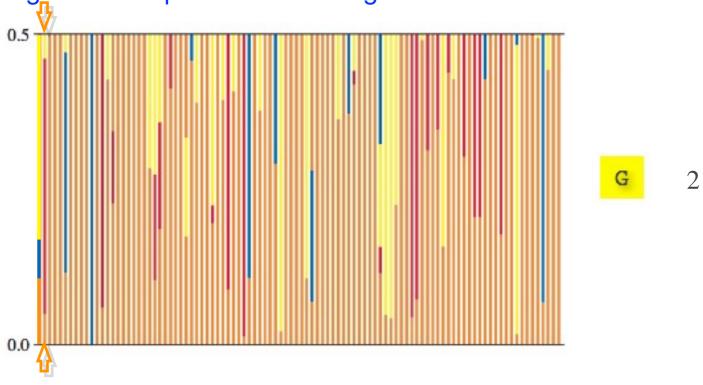
 $T \rightarrow C \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



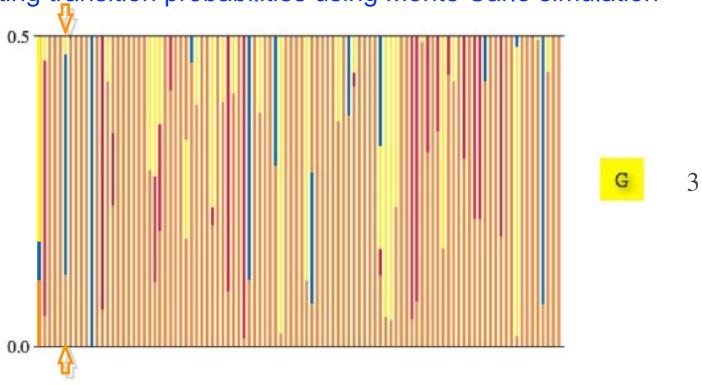
 $T \rightarrow C \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



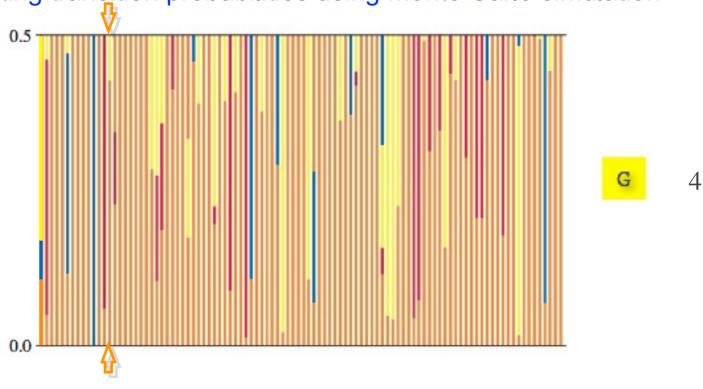
 $T \rightarrow A \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



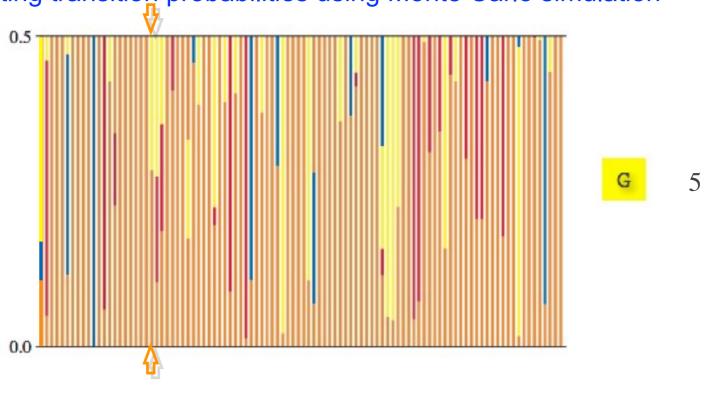
$$T \rightarrow C \rightarrow G$$

Estimating transition probabilities using Monte Carlo simulation



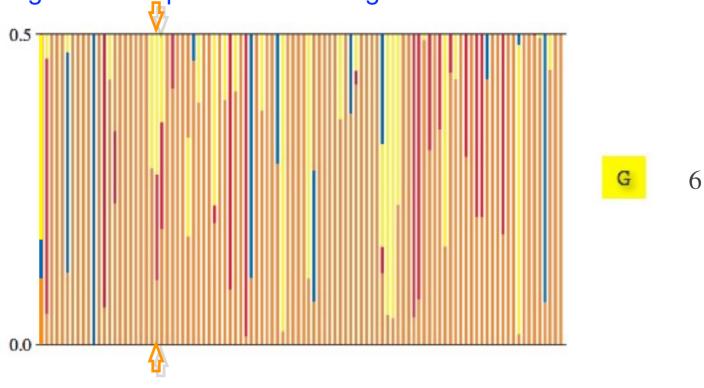
 $T \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation



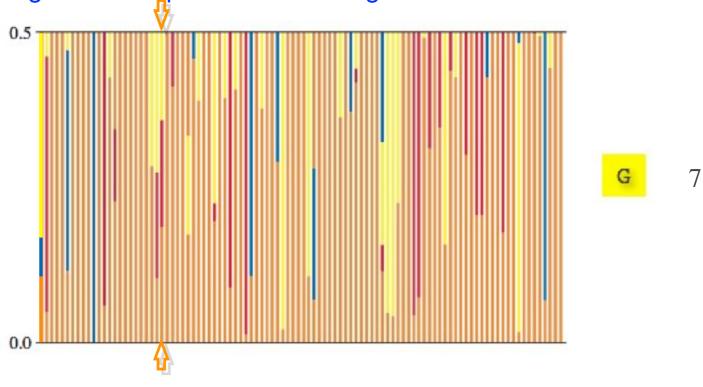
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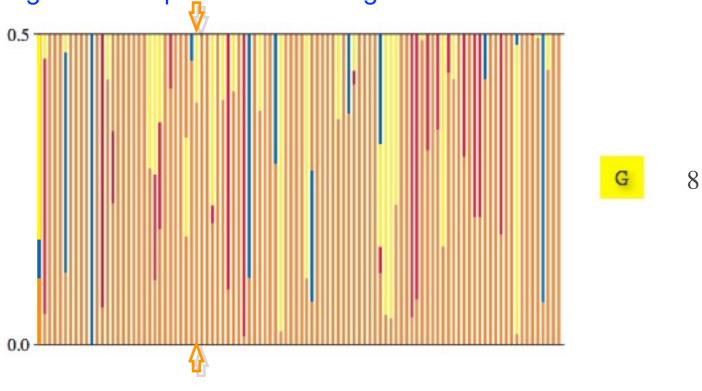
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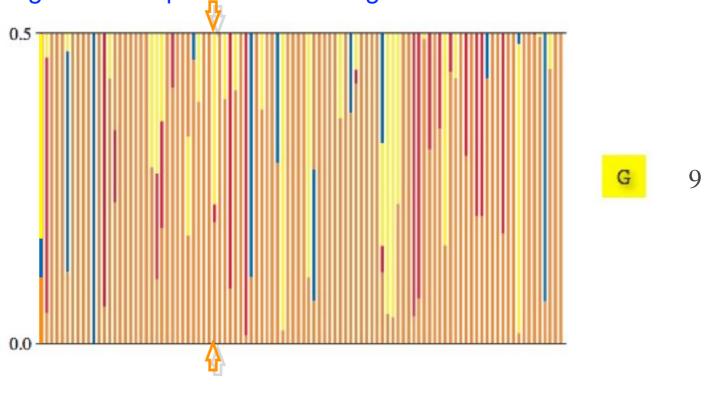
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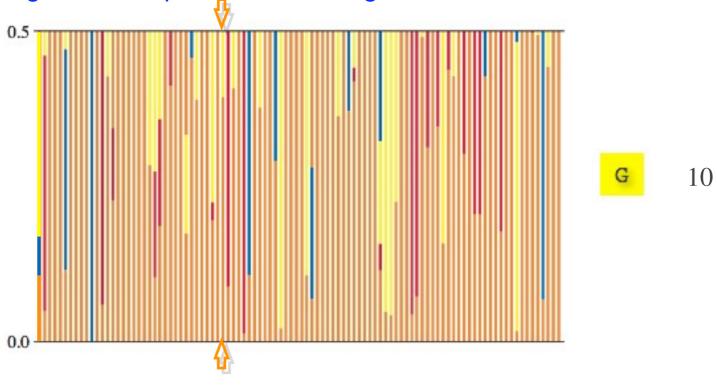
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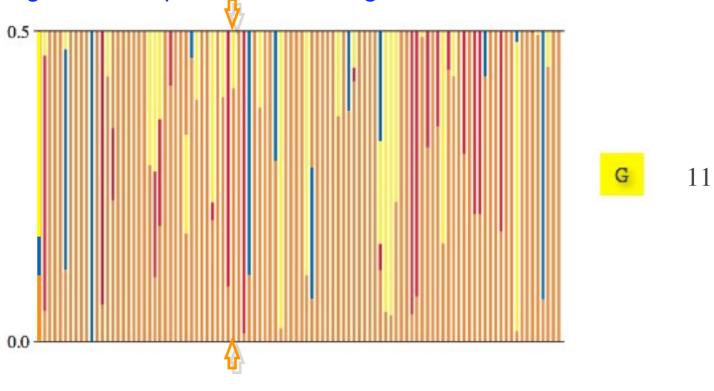
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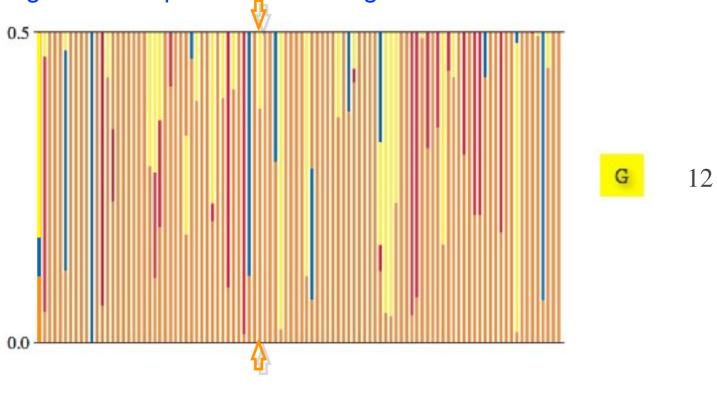
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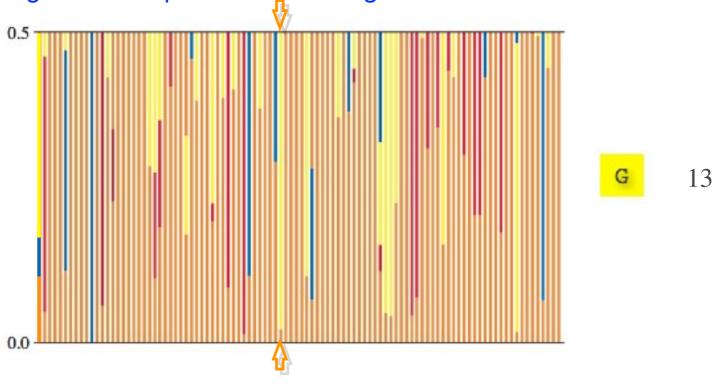
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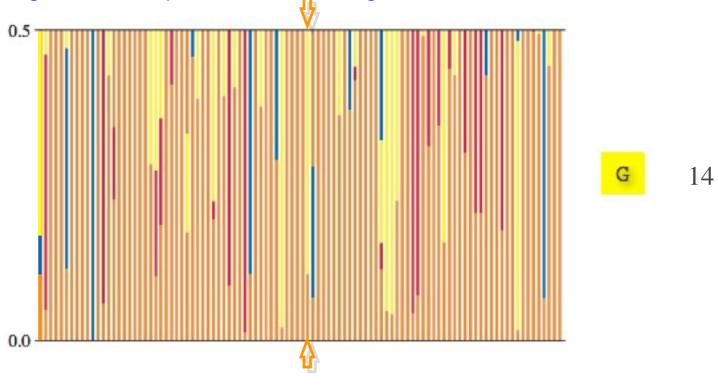
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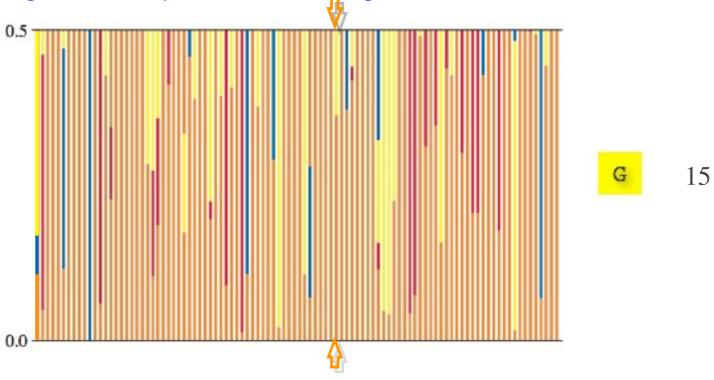
 $T \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation

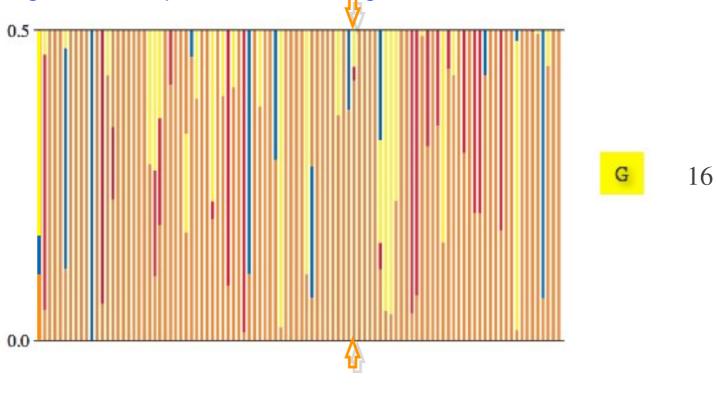


 $T \rightarrow G$

Estimating transition probabilities using Monte Carlo simulation

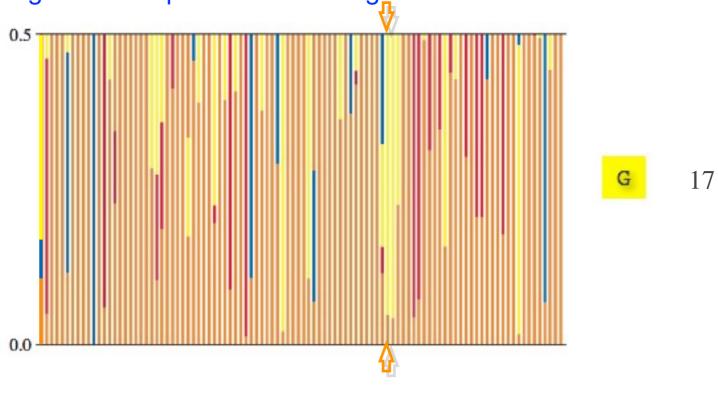


Estimating transition probabilities using Monte Carlo simulation



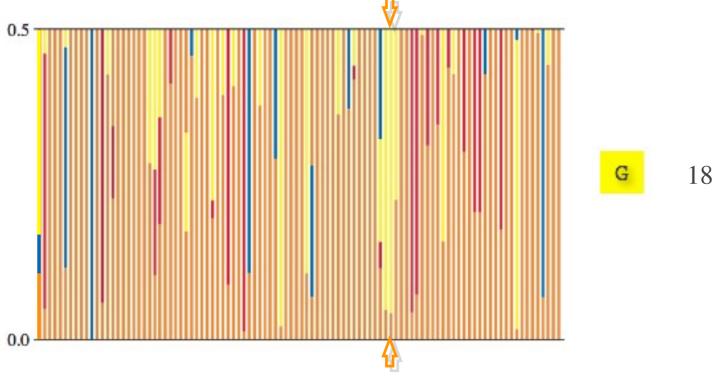
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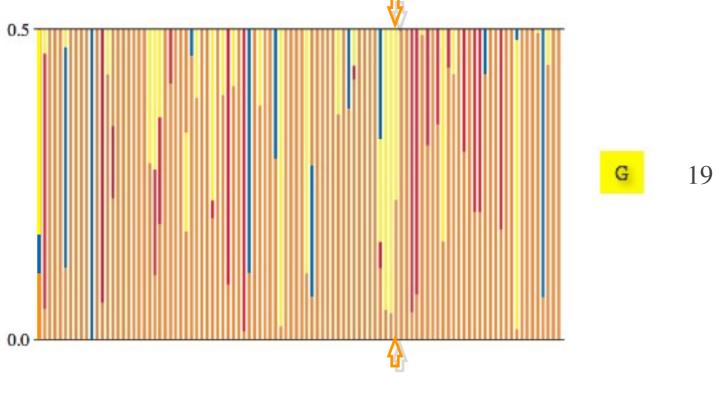
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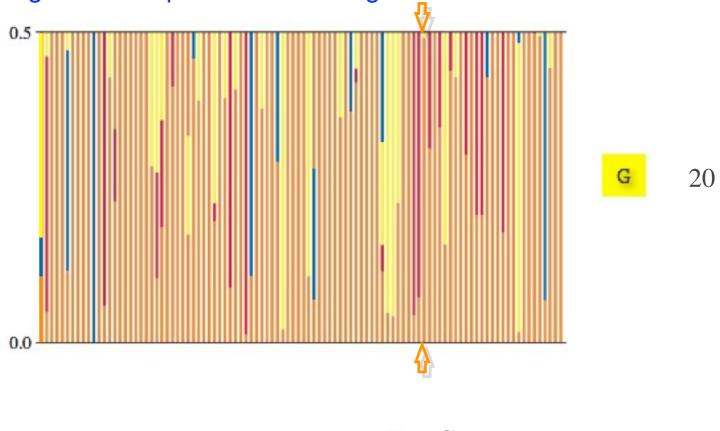
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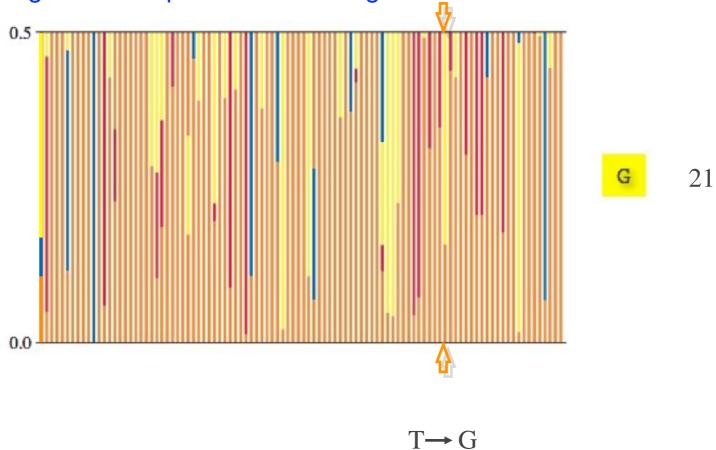
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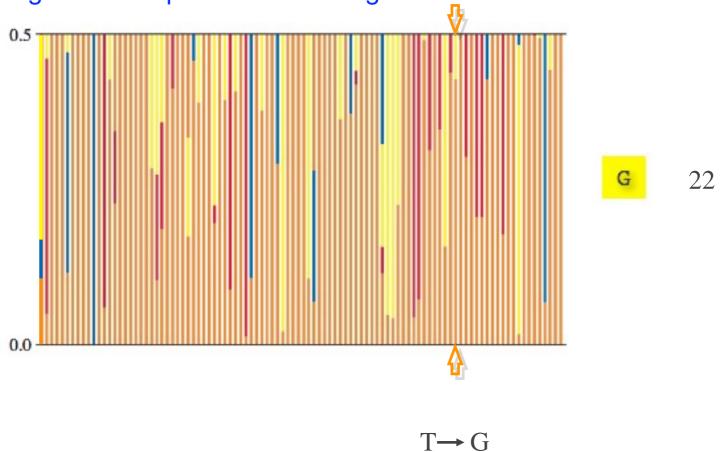


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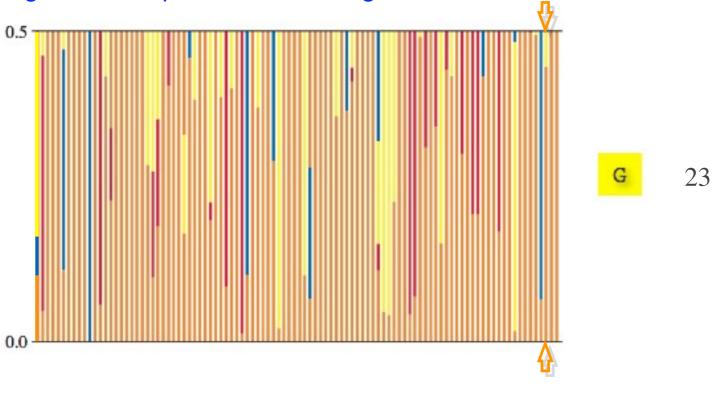
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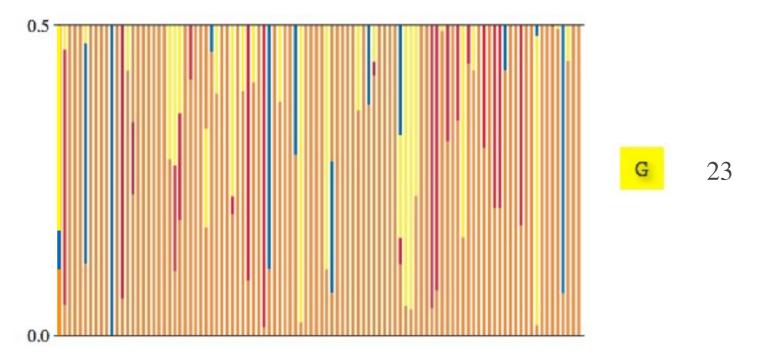


Estimating transition probabilities using Monte Carlo simulation

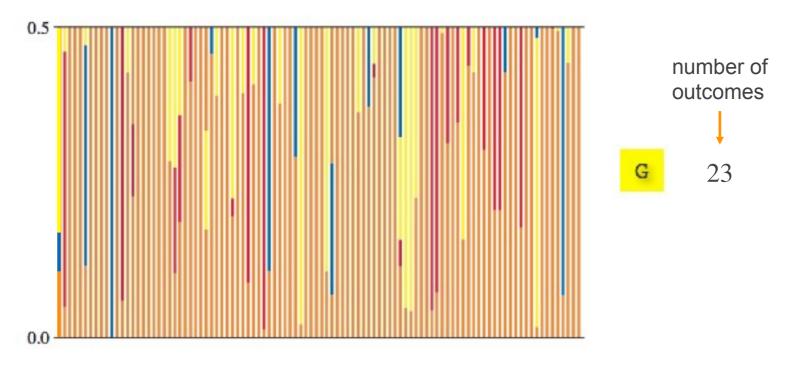


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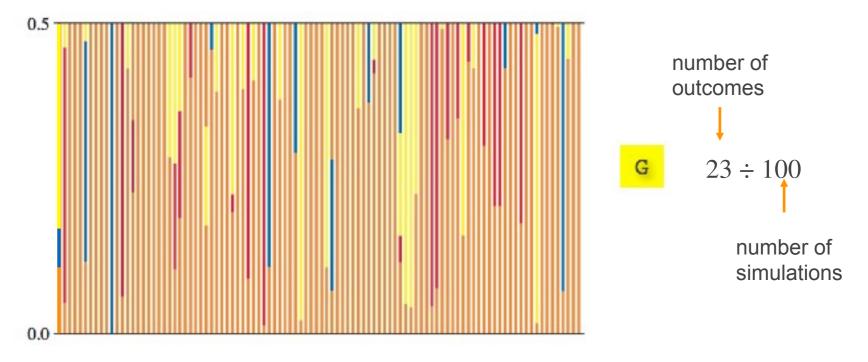
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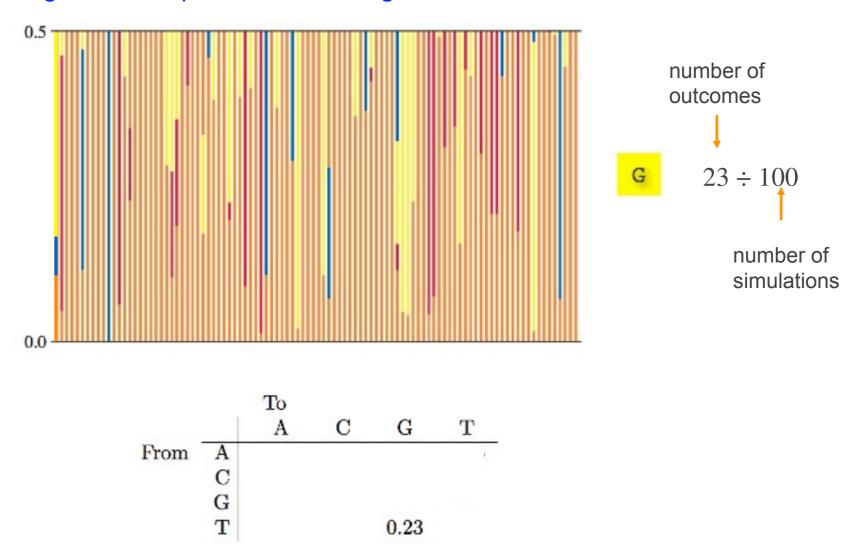
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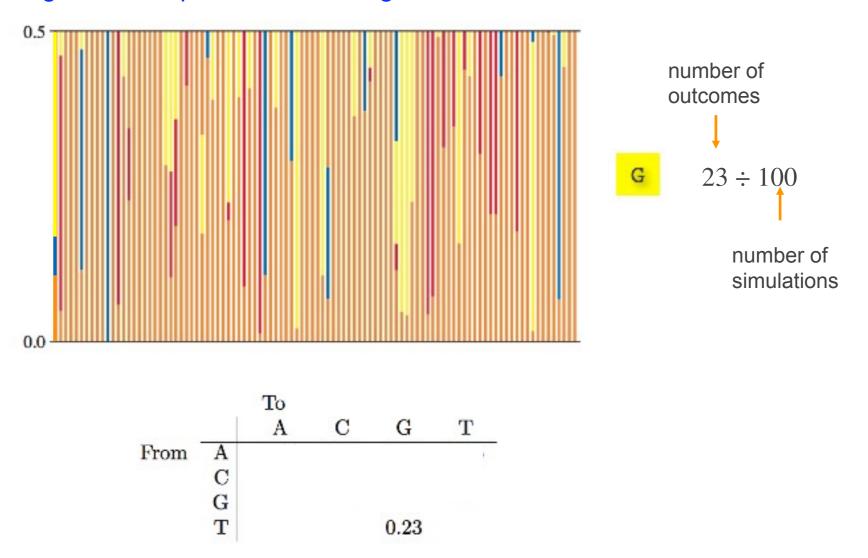
Estimating transition probabilities using Monte Carlo simulation



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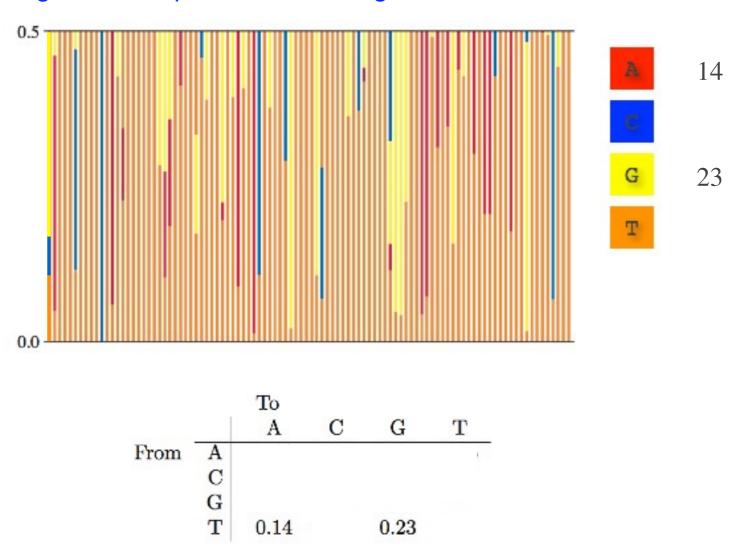


Estimating transition probabilities using Monte Carlo simulation



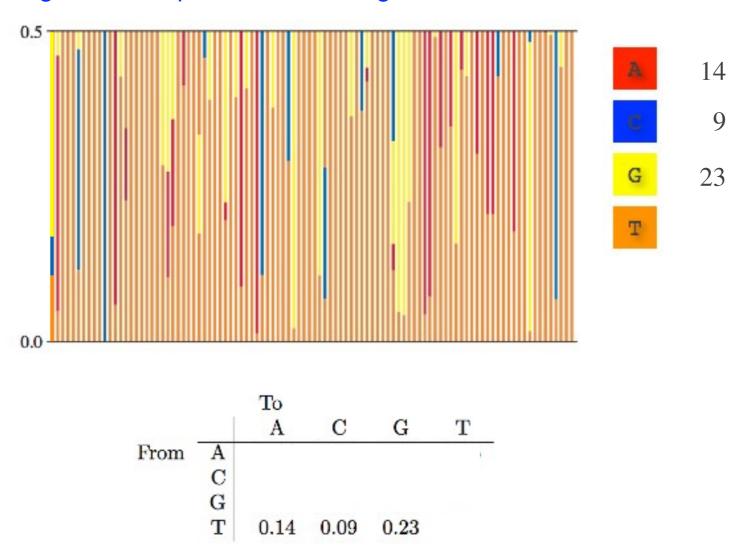
This 'transition probability' reflects all possible histories that start in T and end in G (*i.e.*, histories with different numbers and/or positions of changes)

Estimating transition probabilities using Monte Carlo simulation



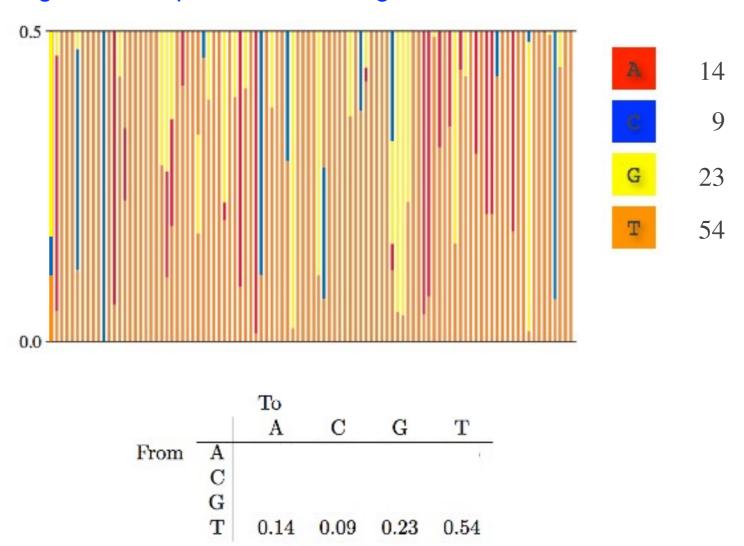
We can repeat this process for the other end states...

Estimating transition probabilities using Monte Carlo simulation



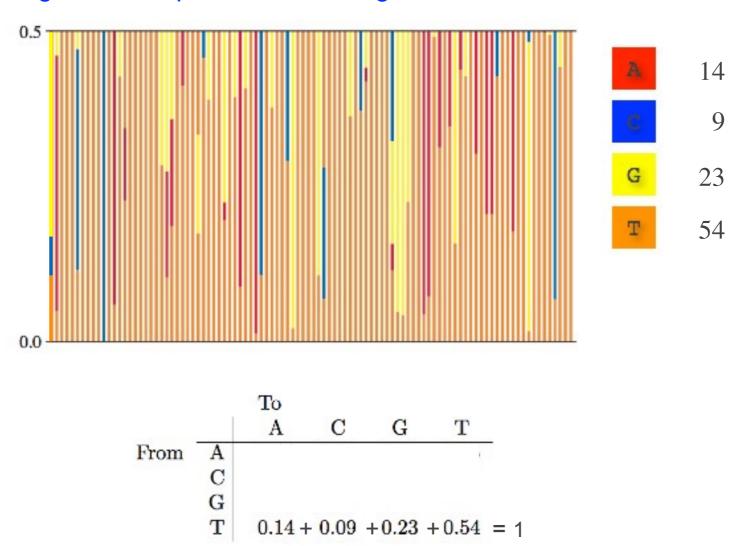
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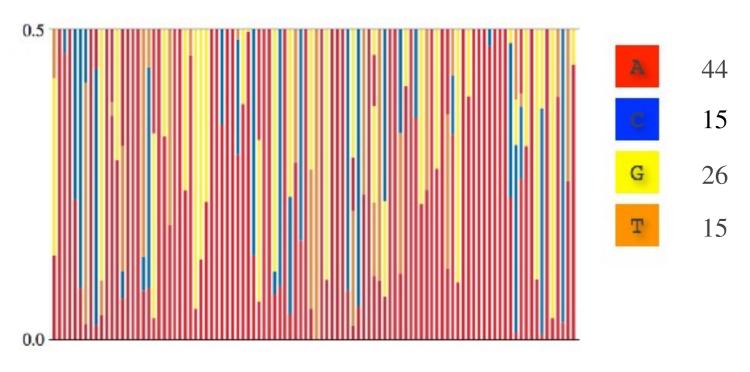
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Estimating transition probabilities using Monte Carlo simulation



Note that each row of the transition-probability matrix sums to 1 (*c.f.*, the Law of Total Probability).

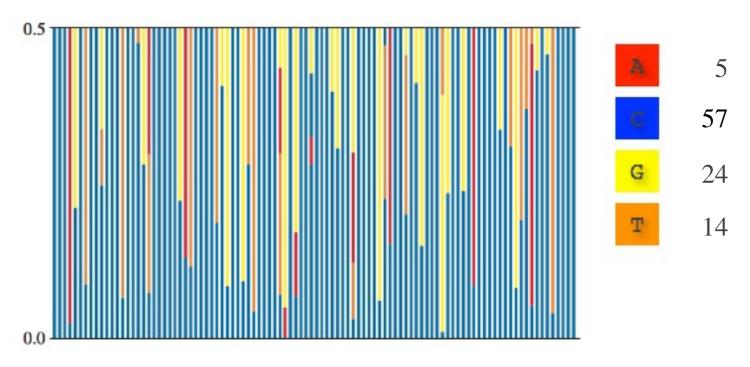
Realizations of 100 replicate simulations starting in state A



		To			
		A	C	\mathbf{G}	\mathbf{T}
From	A	0.44	0.15	0.26	0.15
	C				
	G				
	T	0.14	0.09	0.23	0.54

We can perform new Monte Carlo simulations that start in A to fill out the corresponding row of the transition-probability matrix.

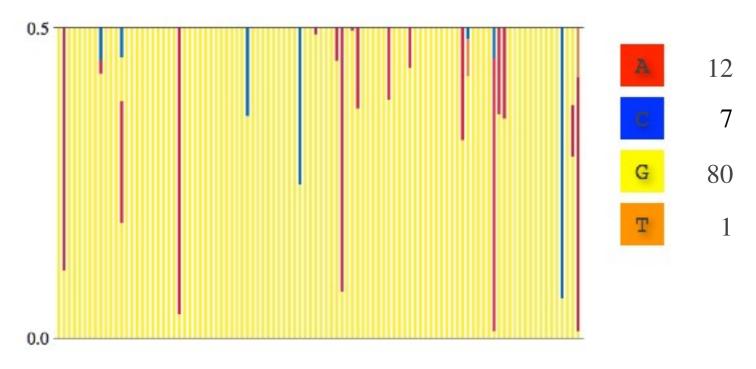
Realizations of 100 replicate simulations starting in state C



		To				
		A	C	\mathbf{G}	\mathbf{T}	
From	A	0.44	0.15	0.26	0.15	
	C	0.05	0.57	0.24	0.14	
	G					
	T	0.14	0.09	0.23	0.54	

And then for simulations that start in C ...

Realizations of 100 replicate simulations starting in state G



		To			
		A	C	\mathbf{G}	\mathbf{T}
From	A	0.44	0.15	0.26	0.15
	C	0.05	0.57	0.24	0.14
	G	0.12	0.07	0.80	0.01
	T	0.14	0.09	0.23	0.54

And finally for simulations that start in G.

Accuracy of Monte Carlo approximation depends on the number of replicates

100 replicates		To			
		A	C	G	\mathbf{T}
From	A	0.44	0.15	0.26	0.15
	C	0.05	0.57	0.24	0.14
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100,000 replica	ites	To			
		A	C	\mathbf{G}	\mathbf{T}
From	A	0.42119	0.15365	0.26361	0.16155
	C	0.06209	0.60811	0.17602	0.15378
	G	0.08834	0.07241	0.77796	0.06129
	\mathbf{T}	0.13534	0.09411	0.22724	0.54331

Analytical solutions for the transition probabilities: matrix exponentiation

Monte Carlo simulation is computationally expensive and unnecessary, as the transition probabilities can be solved 'analytically'

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The exact solution for the transition probability matrix for our instantaneous-rate matrix and branch length (0.5) is:

$$\mathbf{P}(\nu) = \{p_{ij}(\nu)\} = \begin{pmatrix} 0.422927 & 0.153118 & 0.263330 & 0.160625 \\ 0.062896 & 0.609068 & 0.175153 & 0.152883 \\ 0.087566 & 0.071950 & 0.778271 & 0.062212 \\ 0.134967 & 0.093601 & 0.226962 & 0.544470 \end{pmatrix}$$

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Compare with approximate solution (based on 100,000 replicates)

		To			
		A	C	\mathbf{G}	\mathbf{T}
From	A	0.42119	0.15365	0.26361	0.16155
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SIAM REVIEW Vol. 45, No. 1, pp. 3-000 © 2003 Society for Industrial and Applied Mathematics

Nineteen Dubious Ways to Compute the Exponential of a Matrix, Twenty-Five Years Later*

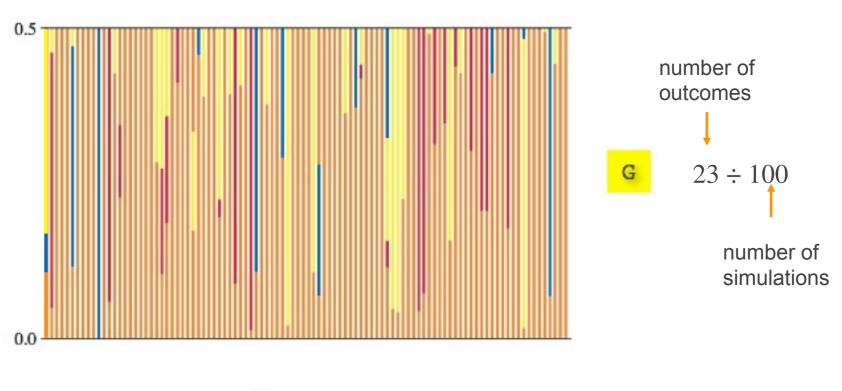
Cleve Moler[†] Charles Van Loan[‡]

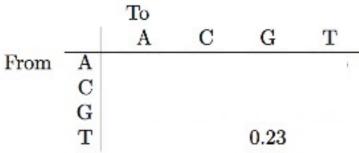
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An aside about transition probabilities

Transition probabilities account for all possible histories that a CTMC can end in a particular state, given a particular starting state (and fully specified model)

Reminder:





An aside about transition probabilities

Transition probabilities account for all possible histories that a CTMC can end in a particular state, given a particular state (and fully specified model)

Transition probabilities play a key role in computing the likelihood, as they avoid the need to condition on a particular history of character change (nucleotide substitution)

Transition probabilities

The probability of observing state j conditioned on starting in state i and running the process over a branch of length v; i.e., $p_{ij}(v)$

Can be estimated by Monte Carlo simulation or matrix exponentiation, $P(v) = e^{Qv}$

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$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

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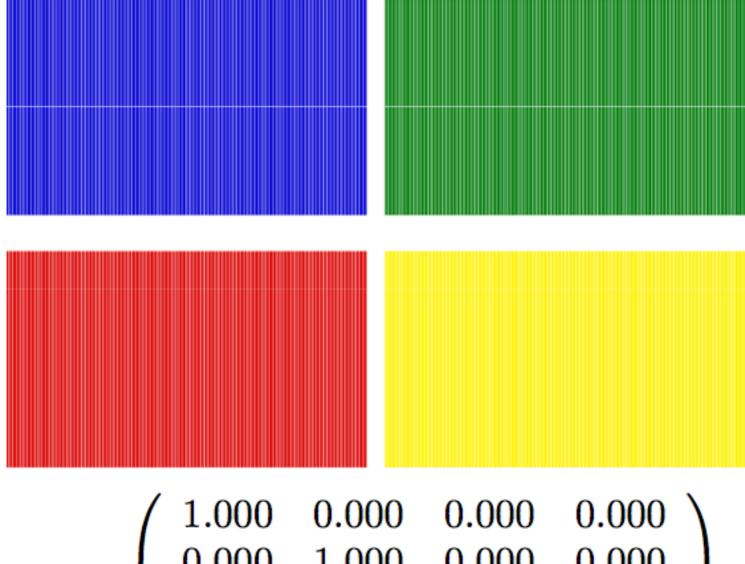
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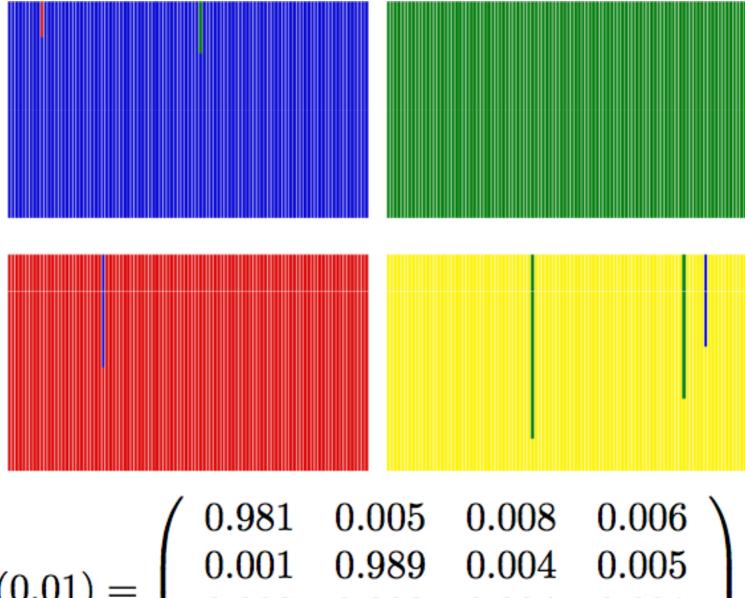
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Transition probabilities over a branch of length v = 0.0:

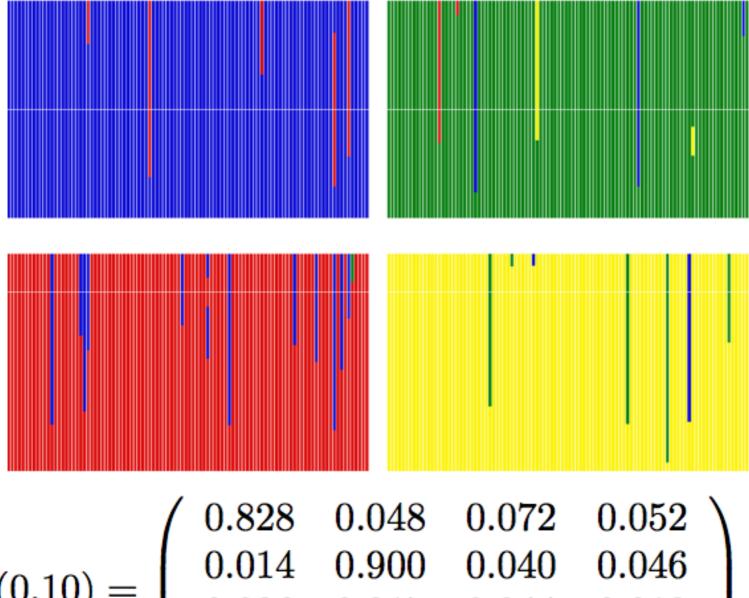
$$\mathbf{P}(0.0) = \begin{pmatrix} 1.000 & 0.000 & 0.000 & 0.000 \\ 0.000 & 1.000 & 0.000 & 0.000 \\ 0.000 & 0.000 & 1.000 & 0.000 \\ 0.000 & 0.000 & 0.000 & 1.000 \end{pmatrix}$$



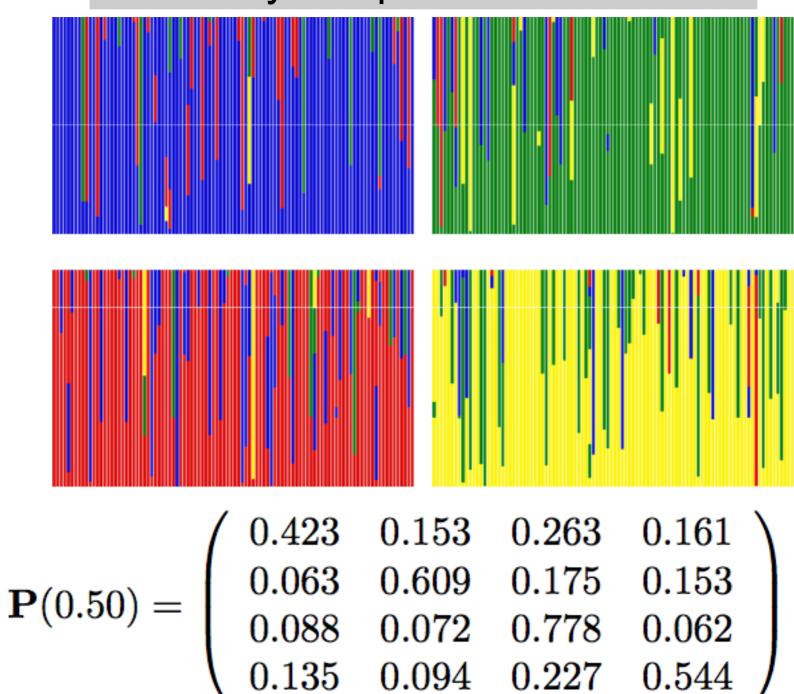
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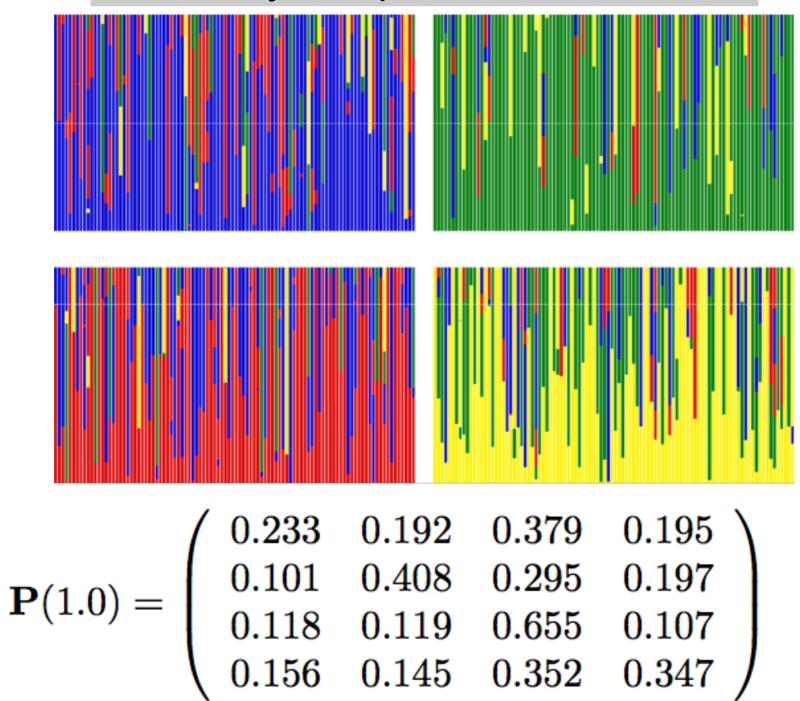


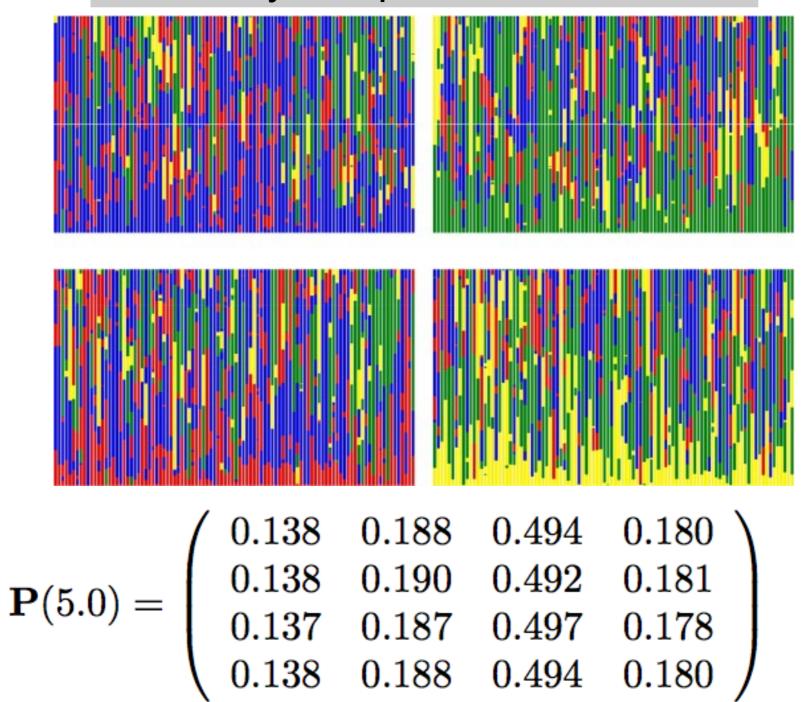
$$\mathbf{P}(0.01) = \begin{pmatrix} 0.001 & 0.003 & 0.003 & 0.003 \\ 0.001 & 0.989 & 0.004 & 0.005 \\ 0.003 & 0.002 & 0.994 & 0.001 \\ 0.005 & 0.002 & 0.006 & 0.986 \end{pmatrix}$$

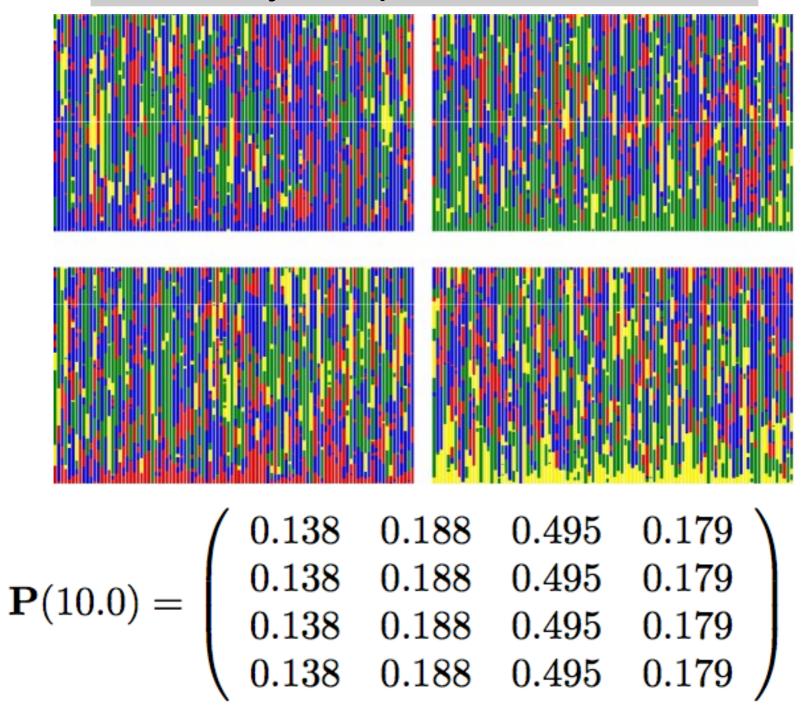


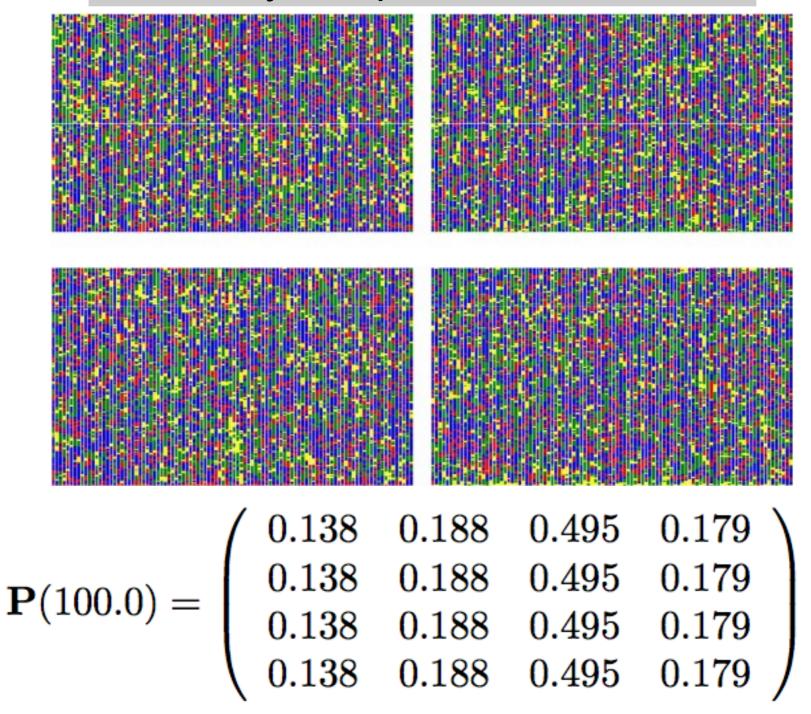
$$\mathbf{P}(0.10) = \begin{pmatrix} 0.014 & 0.900 & 0.040 & 0.046 \\ 0.026 & 0.017 & 0.944 & 0.013 \\ 0.046 & 0.023 & 0.056 & 0.876 \end{pmatrix}$$











Stationary frequencies

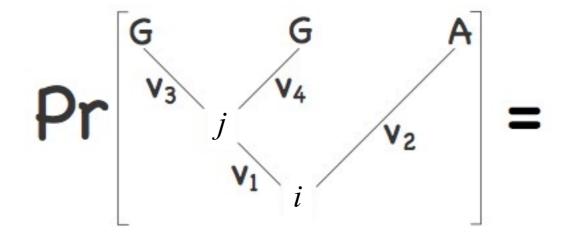
The probability of observing the process in a particular state *j* after a long (infinite) period of time

Stationary frequencies

The probability of observing the process in a particular state *j* after a long (infinite) period of time

When the continuous time Markov chain is at stationarity, the stochastic process has 'forgotten' the starting state: the process ends in a given state with the same probability, regardless of the starting state

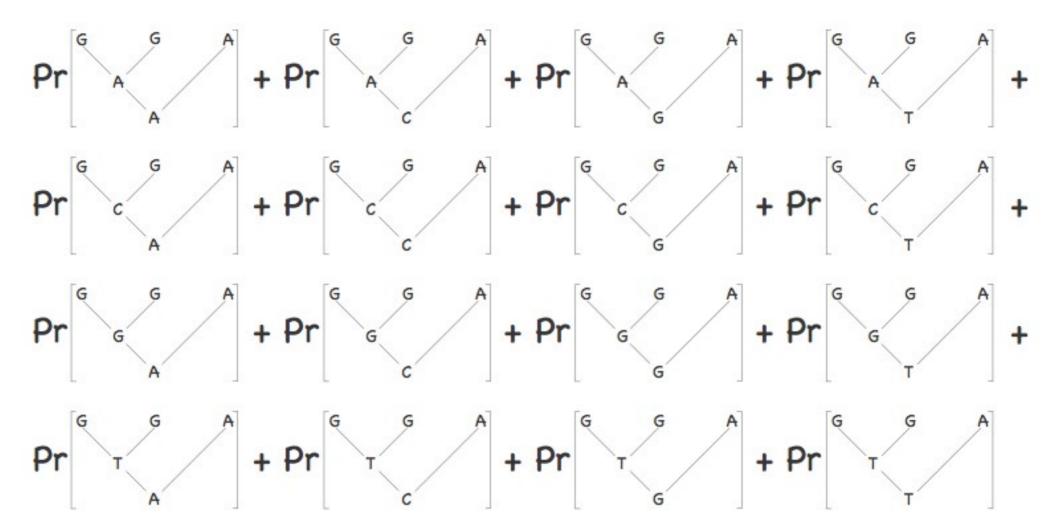
Now we can compute the likelihood of a site history



$$\pi_i \times p_{ij}(v_1) \times p_{iA}(v_2) \times p_{jG}(v_3) \times p_{jG}(v_4)$$

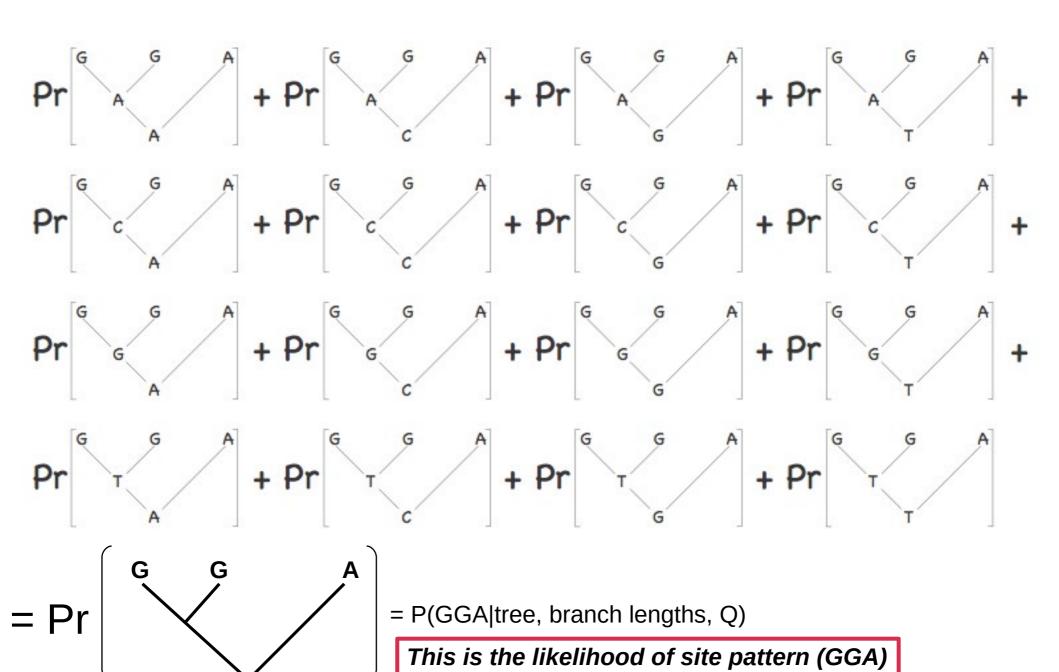
 π_i Stationary frequencies

 $p_{ij}(v)$ Transition probabilities



$$Pr\begin{bmatrix} G & G & A \\ A & A \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ A & C \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ G & A \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\$$

$$Pr\begin{bmatrix} G & G & A \\ A & A \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ A & C \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ A & C \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ A & C \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ A & C \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ G & C & A \end{bmatrix} + Pr\begin{bmatrix} G & G & A \\ G & C & C \\ G & C & C \\ G & C & C & C \\ G & C & C \\ G & C & C & C \\ G & C & C$$



Continuous-time Markov models

describe the stochastic process by which traits (nucleotides) evolve over the tree

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Stationary frequencies: the long-term probability of observing the chain in state *j*

Instantaneous-rate matrix, Q, and the transition probability matrix, P

The instantaneous rate matrix describes the probability of change between each state in an infinitesimal time interval, $q_{ij}(\partial t)$

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The transition probability matrix, $P(v) = \{p_{ij}(v)\}$, describes the probability of observing state j given that we started in state i and ran the process over a branch of length v

$$\mathbf{P}(\nu) = \{p_{ij}(\nu)\} = \begin{pmatrix} 0.422927 & 0.153118 & 0.263330 & 0.160625 \\ 0.062896 & 0.609068 & 0.175153 & 0.152883 \\ 0.087566 & 0.071950 & 0.778271 & 0.062212 \\ 0.134967 & 0.093601 & 0.226962 & 0.544470 \end{pmatrix}$$

Instantaneous-rate matrix, Q, and the transition probability matrix, P

The instantaneous rate matrix describes the probability of change between each state in an infinitesimal time interval, $q_{ij}(\partial t)$

$$\mathbf{Q} = \{q_{ij}\} = \begin{pmatrix} -1.916 & 0.541 & 0.787 & 0.588 \\ 0.148 & -1.069 & 0.415 & 0.506 \\ 0.286 & 0.170 & -0.591 & 0.135 \\ 0.525 & 0.236 & 0.594 & -1.355 \end{pmatrix}$$

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The relationship between **Q** and **P** is $P(v) = e^{\mathbf{Q}v}$

the transition probabilities integrate over all possible histories by which an initial state i can give rise to an end state j over branch length v

Aims

Understand the main ideas underlying models of sequence evolution

- To do so, we will:
 - Introduce important probability notions
 - Play with models of character evolution through simulations
- Briefly present some of the main models of nucleotide evolution

Substitution models

Rate matrix

$$Q = egin{pmatrix} -\mu_A & \mu_{GA} & \mu_{CA} & \mu_{TA} \ \mu_{AG} & -\mu_G & \mu_{CG} & \mu_{TG} \ \mu_{AC} & \mu_{GC} & -\mu_C & \mu_{TC} \ \mu_{AT} & \mu_{GT} & \mu_{CT} & -\mu_T \end{pmatrix}$$

Substitution models

Rate matrix

$$Q = egin{pmatrix} -\mu_A & \mu_{GA} & \mu_{CA} & \mu_{TA} \ \mu_{AG} & -\mu_G & \mu_{CG} & \mu_{TG} \ \mu_{AC} & \mu_{GC} & -\mu_C & \mu_{TC} \ \mu_{AT} & \mu_{GT} & \mu_{CT} & -\mu_T \end{pmatrix}$$

Jukes and Cantor 1969

Kimura 1980

Hasegawa, Kishino,

$$Q = \begin{pmatrix} * & \frac{\mu}{4} & \frac{\mu}{4} & \frac{\mu}{4} \\ \frac{\mu}{4} & * & \frac{\mu}{4} & \frac{\mu}{4} \\ \frac{\mu}{4} & \frac{\mu}{4} & * & \frac{\mu}{4} \\ \frac{\mu}{4} & \frac{\mu}{4} & \frac{\mu}{4} & * \end{pmatrix} Q = \begin{pmatrix} * & \kappa & 1 & 1 \\ \kappa & * & 1 & 1 \\ 1 & 1 & * & \kappa \\ 1 & 1 & \kappa & * \end{pmatrix} Q = \begin{pmatrix} * & \kappa \pi_C & \pi_A & \pi_G \\ \kappa \pi_T & * & \pi_A & \pi_G \\ \pi_T & \pi_C & * & \kappa \pi_G \\ \pi_T & \pi_C & \kappa \pi_A & * \end{pmatrix}$$

- 1 free parameter (0 if we impose one substitution per unit time)
- 1 transition/transversion ratio: 1 free parameter
- 1 transition/transversion ratio
- 4 equilibrium frequencies:
- 4 free parameters

Substitution models

Rate matrix

$$Q = egin{pmatrix} -\mu_A & \mu_{GA} & \mu_{CA} & \mu_{TA} \ \mu_{AG} & -\mu_G & \mu_{CG} & \mu_{TG} \ \mu_{AC} & \mu_{GC} & -\mu_C & \mu_{TC} \ \mu_{AT} & \mu_{GT} & \mu_{CT} & -\mu_T \end{pmatrix}$$

Jukes and Cantor 1969

Kimura 1980

$$Q = egin{pmatrix} * & \kappa & 1 & 1 \ \kappa & * & 1 & 1 \ 1 & 1 & * & \kappa \ 1 & 1 & \kappa & * \end{pmatrix}$$

Kishino,

$$Q = \begin{pmatrix} * & \frac{\mu}{4} & \frac{\mu}{4} & \frac{\mu}{4} \\ \frac{\mu}{4} & * & \frac{\mu}{4} & \frac{\mu}{4} \\ \frac{\mu}{4} & \frac{\mu}{4} & * & \frac{\mu}{4} \\ \frac{\mu}{4} & \frac{\mu}{4} & \frac{\mu}{4} & * \end{pmatrix} Q = \begin{pmatrix} * & \kappa & 1 & 1 \\ \kappa & * & 1 & 1 \\ 1 & 1 & * & \kappa \\ 1 & 1 & \kappa & * \end{pmatrix} Q = \begin{pmatrix} * & \kappa \pi_C & \pi_A & \pi_G \\ \kappa \pi_T & * & \pi_A & \pi_G \\ \pi_T & \pi_C & * & \kappa \pi_G \\ \pi_T & \pi_C & \kappa \pi_A & * \end{pmatrix}$$

Hasegawa,

1 free parameter (0 if we impose one substitution per unit time)

1 transition/transversion ratio: 1 free parameter

1 transition/transversion ratio

4 equilibrium frequencies:

4 free parameters

General Time Reversible model of substitution

Rate matrix

$$Q = egin{pmatrix} -\mu_A & \mu_{GA} & \mu_{CA} & \mu_{TA} \ \mu_{AG} & -\mu_G & \mu_{CG} & \mu_{TG} \ \mu_{AC} & \mu_{GC} & -\mu_C & \mu_{TC} \ \mu_{AT} & \mu_{GT} & \mu_{CT} & -\mu_T \end{pmatrix}$$

Lanave et al. 1984; Tavaré, 1986

$$Q = \begin{pmatrix} -(\alpha \pi_G + \beta \pi_C + \gamma \pi_T) & \alpha \pi_G & \beta \pi_C & \gamma \pi_T \\ \alpha \pi_A & -(\alpha \pi_A + \delta \pi_C + \epsilon \pi_T) & \delta \pi_C & \epsilon \pi_T \\ \beta \pi_A & \delta \pi_G & -(\beta \pi_A + \delta \pi_G + \eta \pi_T) & \eta \pi_T \\ \gamma \pi_A & \epsilon \pi_G & \eta \pi_C & -(\gamma \pi_A + \epsilon \pi_G + \eta \pi_C) \end{pmatrix}$$

4 **stationary frequencies**: 3 parameters

6 *exchangeability parameters*: 5 parameters (if we impose one substitution per unit time)

Summary on CTMCs

- We use CTMCs to model character evolution
- Given an instantaneous rate matrix, we can compute substitution probabilities on a branch with an arbitrary length
- We can combine these computations to compute the likelihood of a site history
- We can sum over site histories to get the likelihood of a site pattern
- The GTR family provides examples of instantaneous rate matrices