

Assignment2 STAT 353

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Question 1 and 3 is on the paper attached.

Question2:

a)

```
gpa<-read.table("2.7.txt",header = T)

fit1<-lm(gpa$GPA~gpa$GMAT)
summary(fit1)
```

```
##
## Call:
## lm(formula = gpa$GPA ~ gpa$GMAT)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.98608 -0.25048 -0.04539  0.47659  0.64531
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  2.157611    2.014430   1.071   0.309
## gpa$GMAT      0.001931    0.003510   0.550   0.594
##
## Residual standard error: 0.5326 on 10 degrees of freedom
## Multiple R-squared:  0.02937,    Adjusted R-squared:  -0.06769
## F-statistic: 0.3026 on 1 and 10 DF,  p-value: 0.5943
```

So the R_{square} is 0.02937 the model is $\text{GPA} = 2.157611 + 0.001931(\text{GMAT}) + e$

b):

the second person fitted value is: $2.157611 + 0.001931(540) = 3.20034$

so the estimate GPA for seconde person is: 3.2

c):

```
anova(fit1)
```

```
## Analysis of Variance Table
##
## Response: gpa$GPA
##           Df Sum Sq Mean Sq F value Pr(>F)
## gpa$GMAT    1  0.08585  0.085847   0.3026  0.5943
## Residuals  10  2.83698  0.283698
```

The p-value is significantly larger than 0.05, which is 0.5943, so the significant evidence shows that do not reject H_0 : the β_1 is 0, so we can conclude that the GMAT is not an important predictor variable.

Question 4 a)

Assumptions:

** x is fixed

** $E(\epsilon_i) = 0$

** $V(\epsilon_i) = \sigma^2$ which is a constant

** $\text{cov}(\epsilon_i, \epsilon_j) = 0$ for $i \neq j$ (ϵ_i are independent)

```
data<-read.table("2.12.txt",header = T)
```

```
fit2<-lm(data$y~data$x)
summary(fit2)
```

```
##
## Call:
## lm(formula = data$y ~ data$x)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.16217 -0.10178 -0.07266  0.03979  0.49064
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.228090   0.137840  -1.655   0.137
## data$x       0.994757   0.005219 190.585 6.43e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.2067 on 8 degrees of freedom
## Multiple R-squared:  0.9998, Adjusted R-squared:  0.9998
## F-statistic: 3.632e+04 on 1 and 8 DF,  p-value: 6.429e-16
```

So the model is $y = -0.228090 + 0.994757x + e$

b):

```
confint(fit2, level = 0.95)
```

```
##              2.5 %      97.5 %
## (Intercept) -0.5459503 0.08977054
## data$x       0.9827204 1.00679271
```

The 95% confidence interval for the intercept(β_0) is : (-0.5459503,0.08977054)

c): The 95% confidence interval for the slope of the model(β_1) is : (0.9827204 ,1.00679271)

d): i): If the $x=0$, $y=0.228090$, which is not a no calcium present. also the CI for the intercept(β_0) contains zero, also shows there should be no corresponding result. ii): As 1 is included in the confidence interval for slope(β_1), so the slope could be 1.

e):

```
#we need beta0 is 0 so we use the zero intercept.
fit3<-lm(data$y~0+data$x)
summary(fit3)
```

```
##
## Call:
## lm(formula = data$y ~ 0 + data$x)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.24861 -0.19054 -0.09167  0.00104  0.49827
##
## Coefficients:
##      Estimate Std. Error t value Pr(>|t|)
## data$x 0.987153   0.002704   365.1  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.2258 on 9 degrees of freedom
## Multiple R-squared:  0.9999, Adjusted R-squared:  0.9999
## F-statistic: 1.333e+05 on 1 and 9 DF,  p-value: < 2.2e-16
```

```
confint(fit3, level = 0.95)
```

```
##           2.5 %    97.5 %
## data$x 0.9810362 0.9932693
```

Comment: The new model line is $y=0.987153x+e$

here the new 95% confidence interval for the slope of the model(beta1) is : (0.9810362, 0.9932693)

f): As we change the model, we remove the intercept term so the slope will change either. The new CI for the beta1 is not include 1 since the model slope is change.