Statistics

Introduction to R for Public Health Researchers

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Statistics

Now we are going to cover how to perform a variety of basic statistical tests in R.

- Correlation
- T-tests/Rank-sum tests
- Linear Regression
- Logistic Regression
- Proportion tests
- Chi-squared
- Fisher's Exact Test

Note: We will be glossing over the statistical theory and "formulas" for these tests. There are plenty of resources online for learning more about these tests, as well as dedicated Biostatistics series at the School of Public Health

cor() performs correlation in R

```
cor(x, y = NULL, use = "everything",
  method = c("pearson", "kendall", "spearman"))
```

Like other functions, if there are NAs, you get NA as the result. But if you specify use only the complete observations, then it will give you correlation on the non-missing data.

```
library(readr)
circ = read_csv("http://johnmuschelli.com/intro_to_r/data/Charm_City_Circulate
cor(circ$orangeAverage, circ$purpleAverage, use="complete.obs")
```

You can also get the correlation between matrix columns

You can also get the correlation between matrix columns

Or between columns of two matrices/dfs, column by column.

```
op = avgs %>% select(orangeAverage, purpleAverage)
gb = avgs %>% select(greenAverage, bannerAverage)
signif(cor(op, gb, use = "complete.obs"), 3)
```

orangeAverage greenAverage bannerAverage 0.840 0.545 purpleAverage 0.867 0.521

You can also use cor.test() to test for whether correlation is significant (ie non-zero). Note that linear regression may be better, especially if you want to regress out other confounders.

For many of these testing result objects, you can extract specific slots/results as numbers, as the ct object is just a list.

```
# str(ct)
names(ct)

[1] "statistic" "parameter" "p.value" "estimate" "null.value"
[6] "alternative" "method" "data.name" "conf.int"

ct$statistic

t
73.65553

ct$p.value

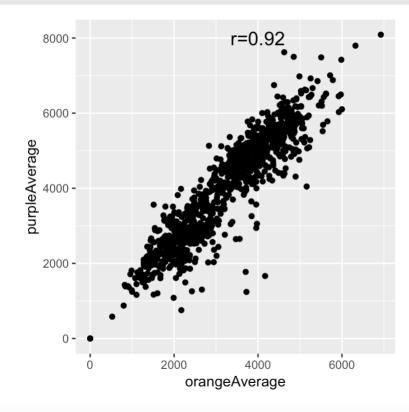
[1] 0
```

Broom package

The broom package has a tidy function that puts most objects into data.frames so that they are easily manipulated:

Note that you can add the correlation to a plot, via the annotate

```
library(ggplot2)
txt = paste0("r=", signif(ct$estimate,3))
q = qplot(data = circ, x = orangeAverage, y = purpleAverage)
q + annotate("text", x = 4000, y = 8000, label = txt, size = 5)
```



The T-test is performed using the t.test() function, which essentially tests for the difference in means of a variable between two groups.

In this syntax, x and y are the column of data for each group.

```
tt = t.test(circ$orangeAverage, circ$purpleAverage)
tt

Welch Two Sample t-test

data: circ$orangeAverage and circ$purpleAverage
t = -17.076, df = 1984, p-value < 2.2e-16
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
    -1096.7602   -870.7867
sample estimates:
mean of x mean of y
    3033.161   4016.935</pre>
```

Using t.test treats the data as independent. Realistically, this data should be treated as a paired t-test. The paired = TRUE argument to do a paired test

```
Paired t-test

data: circ$orangeAverage and circ$purpleAverage
t = -42.075, df = 992, p-value < 2.2e-16
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
-799.783 -728.505
sample estimates:
mean of the differences
-764.144
```

t.test(circ\$orangeAverage, circ\$purpleAverage, paired = TRUE)

t.test saves a lot of information: the difference in means estimate, confidence interval for the difference conf.int, the p-value p.value, etc.

```
names(tt)

[1] "statistic" "parameter" "p.value" "conf.int" "estimate"
[6] "null.value" "alternative" "method" "data.name"
```

```
tidy(tt)
```

You can also use the 'formula' notation. In this syntax, it is $y \sim x$, where x is a factor with 2 levels or a binary variable and y is a vector of the same length.

Wilcoxon Rank-Sum Tests

Nonparametric analog to t-test (testing medians):

Lab Part 1

Website

Now we will briefly cover linear regression. I will use a little notation here so some of the commands are easier to put in the proper context. $y_i = \alpha + \beta x_i + \epsilon_i$ where:

- y_i is the outcome for person i
- \cdot α is the intercept
- \cdot β is the slope
- \cdot x_i is the predictor for person i
- \cdot ϵ_i is the residual variation for person i

The R version of the regression model is:

y ~ x

where:

- · y is your outcome
- x is/are your predictor(s)

For a linear regression, when the predictor is binary this is the same as a t-test:

The summary command gets all the additional information (p-values, t-statistics, r-square) that you usually want from a regression.

```
sfit = summary(fit)
print(sfit)
Call:
lm(formula = avg ~ line, data = long)
Residuals:
   Min 1Q Median 3Q Max
-4016.9 -1121.2 64.3 1060.8 4072.6
Coefficients:
                Estimate Std. Error t value Pr(>|t|)
(Intercept) 3033.16 38.99 77.79 <2e-16 ***
linepurpleAverage 983.77 57.09 17.23 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 1314 on 2127 degrees of freedom
  (163 observations deleted due to missingness)
Multiple R-squared: 0.1225, Adjusted R-squared: 0.1221
F-statistic: 296.9 on 1 and 2127 DF, p-value: < 2.2e-16
```

The coefficients from a summary are the coefficients, standard errors, t-statistcs, and p-values for all the estimates.

We can tidy linear models as well and it gives us all of this in one::

```
tidy(fit)
```

Using Cars Data

```
http data dir = "http://johnmuschelli.com/intro to r/data/"
   cars = read csv(
     paste0(http data dir, "kaggleCarAuction.csv"),
     col types = cols(VehBCost = col double()))
   head(cars)
   # A tibble: 6 \times 34
     RefId IsBadBuy PurchDate Auction VehYear VehicleAge Make Model Trim
     <dbl> <dbl> <chr> <dbl> <dbl> <chr> <dbl> <dbl> <chr> <dbl> 
                  0 12/7/2009 ADESA 2006
                                                     3 MAZDA MAZD... i
                 0 12/7/2009 ADESA 2004 5 DODGE 1500... ST
   3
                 0 12/7/2009 ADESA 2005 4 DODGE STRA... SXT
                 0 12/7/2009 ADESA 2004
   4
                                                     5 DODGE NEON SXT
   5
                 0 12/7/2009 ADESA 2005
                                                  4 FORD FOCUS ZX3
                  0 12/7/2009 ADESA 2004
                                                     5 MITS... GALA... ES
     ... with 25 more variables: SubModel <chr>, Color <chr>,
   #
       Transmission <chr>, WheelTypeID <chr>, WheelType <chr>, VehOdo <dbl>,
       Nationality <chr>, Size <chr>, TopThreeAmericanName <chr>,
       MMRAcquisitionAuctionAveragePrice <chr>,
       MMRAcquisitionAuctionCleanPrice <chr>,
   #
       MMRAcquisitionRetailAveragePrice <chr>,
       MMRAcquisitonRetailCleanPrice <chr>,
       MMRCurrentAuctionAveragePrice <chr>,
       MMRCurrentAuctionCleanPrice <chr>, MMRCurrentRetailAveragePrice <chr>,
       MMRCurrentRetailCleanPrice <chr>, PRIMEUNIT <chr>, AUCGUART <chr>,
       BYRNO <dbl>, VNZIP1 <dbl>, VNST <chr>, VehBCost <dbl>,
       IsOnlineSale <dbl>, WarrantyCost <dbl>
                                                                       23/38
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```

We'll look at vehicle odometer value by vehicle age:

Note that you can have more than 1 predictor in regression models. The interpretation for each slope is change in the predictor corresponding to a one-unit change in the outcome, holding all other predictors constant.

```
fit2 = lm(VehOdo ~ IsBadBuy + VehicleAge, data = cars)
summary(fit2)
Call:
lm(formula = VehOdo ~ IsBadBuy + VehicleAge, data = cars)
Residuals:
  Min 1Q Median 3Q Max
-70856 -9490 1390 10311 41193
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 60141.77 134.75 446.33 <2e-16 ***
IsBadBuy 1329.00 157.84 8.42 <2e-16 ***
VehicleAge 2680.33 30.27 88.53 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 13810 on 72980 degrees of freedom
Multiple R-squared: 0.1031, Adjusted R-squared: 0.1031
F-statistic: 4196 on 2 and 72980 DF, p-value: < 2.2e-16
```

Linear Regression: Interactions

The * does interactions:

```
fit3 = lm(VehOdo ~ IsBadBuy * VehicleAge, data = cars)
summary(fit3)
Call:
lm(formula = VehOdo ~ IsBadBuy * VehicleAge, data = cars)
Residuals:
  Min 1Q Median 3Q Max
-70857 -9490 1391 10311 41193
Coefficients:
               Estimate Std. Error t value Pr(>|t|)
(Intercept) 60139.70 143.23 419.872 < 2e-16 ***
IsBadBuy
                1347.28 456.23 2.953 0.00315 **
VehicleAge 2680.84 32.54 82.380 < 2e-16 ***
IsBadBuy: VehicleAge -3.79 88.74 -0.043 0.96594
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 13810 on 72979 degrees of freedom
Multiple R-squared: 0.1031, Adjusted R-squared: 0.1031
F-statistic: 2798 on 3 and 72979 DF, p-value: < 2.2e-16
```

Linear Regression: Interactions

You can take out main effects with minus

```
fit4 = lm(VehOdo ~ IsBadBuy * VehicleAge -IsBadBuy , data = cars)
summary(fit4)
Call:
lm(formula = VehOdo ~ IsBadBuy * VehicleAge - IsBadBuy, data = cars)
Residuals:
  Min 1Q Median 3Q Max
-70822 -9493 1389 10311 41172
Coefficients:
                Estimate Std. Error t value Pr(>|t|)
(Intercept) 60272.49 136.00 443.184 < 2e-16 ***
VehicleAge 2652.94 31.14 85.186 < 2e-16 ***
IsBadBuy:VehicleAge 242.08 30.70 7.885 3.19e-15 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 13810 on 72980 degrees of freedom
Multiple R-squared: 0.103, Adjusted R-squared: 0.103
F-statistic: 4192 on 2 and 72980 DF, p-value: < 2.2e-16
```

Factors get special treatment in regression models - lowest level of the factor is the comparison group, and all other factors are relative to its values.

```
fit3 = lm(VehOdo ~ factor(TopThreeAmericanName), data = cars)
summary(fit3)
Call:
lm(formula = VehOdo ~ factor(TopThreeAmericanName), data = cars)
Residuals:
  Min
        10 Median
                       30
                             Max
-71947 -9634 1532 10472 45936
Coefficients:
                                Estimate Std. Error t value Pr(>|t|)
(Intercept)
                                68248.48 92.98 733.984 < 2e-16 ***
factor(TopThreeAmericanName)FORD 8523.49 158.35 53.828 < 2e-16 ***
factor(TopThreeAmericanName)GM 4952.18 128.99 38.393 < 2e-16 ***
factor(TopThreeAmericanName)NULL -2004.68 6361.60 -0.315 0.752670
factor(TopThreeAmericanName)OTHER 584.87 159.92 3.657 0.000255 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 14220 on 72978 degrees of freedom
Multiple R-squared: 0.04822, Adjusted R-squared: 0.04817
F-statistic: 924.3 on 4 and 72978 DF, p-value: < 2.2e-16
```

Logistic Regression and GLMs

Generalized Linear Models (GLMs) allow for fitting regressions for non-continous/normal outcomes. The glm has similar syntax to the lm command. Logistic regression is one example.

```
glmfit = glm(IsBadBuy ~ VehOdo + VehicleAge, data=cars, family = binomial())
summary(glmfit)
Call:
glm(formula = IsBadBuy ~ VehOdo + VehicleAge, family = binomial(),
   data = cars)
Deviance Residuals:
   Min
            10 Median 30
                                       Max
-0.9943 \quad -0.5481 \quad -0.4534 \quad -0.3783 \quad 2.6318
Coefficients:
             Estimate Std. Error z value Pr(>|z|)
(Intercept) -3.778e+00 6.381e-02 -59.211 <2e-16 ***
VehOdo 8.341e-06 8.526e-07 9.783 <2e-16 ***
VehicleAge 2.681e-01 6.772e-03 39.589 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 54421 on 72982 degrees of freedom
Residual deviance: 52346 on 72980 degrees of freedom
AIC: 52352
Number of Fisher Scoring iterations: 5
```

Tidying GLMs

```
tidy(glmfit, conf.int = TRUE)
```

Tidying GLMs

Logistic Regression

Note the coefficients are on the original scale, we must exponentiate them for odds ratios:

```
exp(coef(glmfit))

(Intercept) VehOdo VehicleAge 0.02286316 1.00000834 1.30748911
```

Chi-squared tests

chisq.test() performs chi-squared contingency table tests and goodness-of-fit tests.

```
0 1
0 62375 1632
1 8763 213
```

Chi-squared tests

You can also pass in a table object (such as tab here)

```
cq = chisq.test(tab)
Cq
   Pearson's Chi-squared test with Yates' continuity correction
data: tab
X-squared = 0.92735, df = 1, p-value = 0.3356
names (cq)
[1] "statistic" "parameter" "p.value" "method" "data.name" "observed"
[7] "expected" "residuals" "stdres"
cq$p.value
[1] 0.3355516
```

Chi-squared tests

Note that does the same test as prop.test, for a 2x2 table (prop.test not relevant for greater than 2x2).

```
chisq.test(tab)
    Pearson's Chi-squared test with Yates' continuity correction
data: tab
X-squared = 0.92735, df = 1, p-value = 0.3356
prop.test(tab)
    2-sample test for equality of proportions with continuity
    correction
data: tab
X-squared = 0.92735, df = 1, p-value = 0.3356
alternative hypothesis: two.sided
95 percent confidence interval:
 -0.005208049 0.001673519
sample estimates:
  prop 1 prop 2
0.9745028 0.9762701
```

Fisher's Exact test

fisher.test() performs contingency table test using the hypogeometric distribution (used for small sample sizes).

control = list(), or = 1, alternative = "two.sided",

fisher.test(x, y = NULL, workspace = 200000, hybrid = FALSE,

conf.int = TRUE, conf.level = 0.95, simulate.p.value = FALSE, B = 2000)

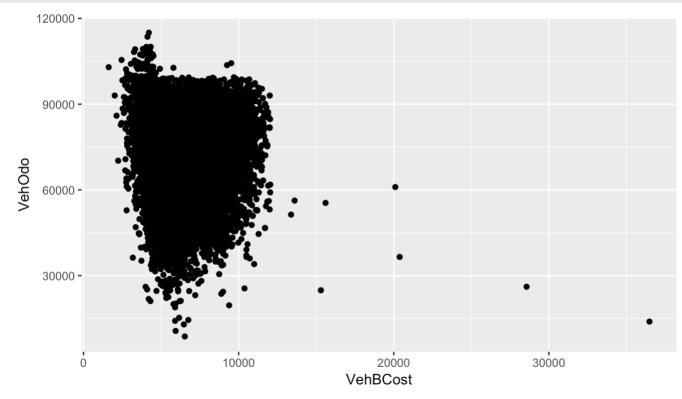
```
fisher.test(tab)

Fisher's Exact Test for Count Data

data: tab
p-value = 0.3324
alternative hypothesis: true odds ratio is not equal to 1
95 percent confidence interval:
    0.8001727 1.0742114
sample estimates:
odds ratio
    0.9289923
```

Sampling

Also, if you want to only plot a subset of the data (for speed/time or overplotting)



Lab Part 2

Website