Adaptive Bayesian estimation and its self-informative limit in an indirect sequence space model

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The Gaussian sequence space model

Consider an indirect Gaussian sequence space model consisting of:

- an unknown parameter of interest $\left(\theta_j^\circ\right)_{j\in\mathbb{N}} = \theta^\circ$,
- a decreasing multiplicative sequence $\left(\lambda_j\right)_{j\in\mathbb{N}} = \lambda$ converging to 0,
- observations $\left(Y_j\right)_{j\in\mathbb{N}} = Y$, contaminated by an additive independent centered Gaussian noise with variance n^{-1},
$$Y = \left(\theta_j^\circ \cdot \lambda_j + \sqrt{n}^{-1}\cdot \xi_j\right)_{j\in\mathbb{N}}, \quad (\xi)_{j\in\mathbb{N}} \sim_{iid} \mathcal{N}(0,1).$$

The goal is to recover θ° and derive an upper bound.

The frequentist model selection

For any index j, an unbiased estimator of θ_j° is Y_j/λ_j. Hence, an intuitive class of estimators are the projection estimators: $\tilde{\theta}^m = \left(Y_j/\lambda_j \mathbb{1}_{\{j\leq m\}}\right)_{j\in\mathbb{N}}$ with m in $\mathbb{N}$. The model selection method offers a data driven way to select m in this context:

$$G_n := \max\left\{1\leq j\leq n : n^{-1}\lambda_j^{-2} \leq \lambda_1^{-2}\right\},$$

$$\widehat{m} := \underset{m\in[\![1,G_n]\!]}{\arg\min}\left\{3m - \sum_{j=1}^{m} Y_j^2\right\}, \qquad \widehat{\theta} := \left(\tilde{\theta}_j^{\widehat{m}}\right)_{j\in\mathbb{N}}.$$

It is shown in Massart [2003], in the direct case, that this estimator is consistent, converges in probability and $\mathbb{L}^2$-norm, noted $\|\cdot\|$, with minimax optimal rate over some Sobolev ellipsoid:

$$\Theta^\circ := \Theta^\circ(\mathfrak{a}, L^\circ)\left\{\theta : \sum_{j=1}^{\infty}\frac{1}{\mathfrak{a}_j}\theta_j^2 < L^\circ\right\}.$$

Bayesian paradigm, iterated posterior distribution and self informative limit

We adopt a Bayesian point of view:

- the parameter $\boldsymbol{\theta}$ is a random variable with prior $\mathbb{P}_{\boldsymbol{\theta}}$,
- given $\boldsymbol{\theta}$, the likelihood of Y is $\mathbb{P}^n_{Y|\boldsymbol{\theta}} = \mathcal{N}\left(\boldsymbol{\theta}\lambda, n^{-1}\mathbb{I}\right)$,
- we are interested in the posterior distribution $\mathbb{P}_{\boldsymbol{\theta}^n|Y} \propto \mathbb{P}^n_{Y|\boldsymbol{\theta}}\cdot\mathbb{P}_{\boldsymbol{\theta}}$.

In the spirit of Bunke and Johannes [2005], we then generate a posterior family by introducing an iteration parameter η:

- for $\eta = 1$, the prior distribution is $\mathbb{P}_{\boldsymbol{\theta}^1} = \mathbb{P}_{\boldsymbol{\theta}}$, the likelihood $\mathbb{P}^n_{Y^1|\boldsymbol{\theta}^1} = \mathbb{P}^n_{Y|\boldsymbol{\theta}}$ and the posterior distribution is $\mathbb{P}^n_{\boldsymbol{\theta}^1|Y^1} = \mathbb{P}^n_{\boldsymbol{\theta}|Y}$,
- for $\eta = 2$, we take the posterior for $\eta = 1$ as prior, hence, the prior distribution is $\mathbb{P}^n_{\boldsymbol{\theta}^2} = \mathbb{P}^n_{\boldsymbol{\theta}^1|Y^1}$, the likelihood is kept the same $\mathbb{P}^n_{Y^2|\boldsymbol{\theta}^2} = \mathbb{P}^n_{Y|\boldsymbol{\theta}}$ and we compute the posterior distribution with the same observations Y, which we note $\mathbb{P}^n_{\boldsymbol{\theta}^2|Y^2}$,
- ...
- for any value of $\eta > 1$, the prior is $\mathbb{P}^n_{\boldsymbol{\theta}^\eta} = \mathbb{P}^n_{\boldsymbol{\theta}^{\eta-1}|Y^{\eta-1}}$ and we compute the posterior with the same likelihood $\mathbb{P}_{Y^\eta|\boldsymbol{\theta}^\eta} = \mathbb{P}^n_{Y|\boldsymbol{\theta}}$ and same observation Y which gives $\mathbb{P}^n_{\boldsymbol{\theta}^\eta|Y^\eta}$.

This iteration procedure corresponds to giving more and more weight to the observations and make the prior knowledge vanish.
Within this framework we define the family of estimators:

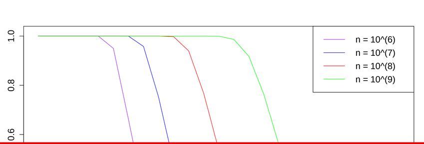
$$\widehat{\theta}^{(\eta)} := \mathbb{E}^n_{\boldsymbol{\theta}^\eta|Y^\eta}[\boldsymbol{\theta}],$$

and call self-informative limit the limit of the estimate with $\eta\to\infty$.
We are interested in the behavior of the family $\left(\mathbb{P}^n_{\boldsymbol{\theta}^\eta|Y^\eta}\right)_{\eta\in\mathbb{N}^\star}$ as n and/or η tend to infinite.
In particular, the question of oracle and minimax concentration (resp. convergence) is answered for any element of the family of posterior distributions (resp. posterior means), including when η tends to infinite.

Hierarchical prior

- Consider a random hyper-parameter M, with values in a subset of $\mathbb{N}$, acting like a threshold:
$$\forall j > m, \quad \mathbb{P}_{\boldsymbol{\theta}_j|M=m} = \delta_0,$$
$$\forall j \leq m, \quad \mathbb{P}_{\boldsymbol{\theta}_j|M=m} = \mathcal{N}(0,1).$$
- if we denote $\mathbb{P}_M$ the distribution of M (to be specified later), then
$$\mathbb{P}^n_{\boldsymbol{\theta}|Y} = \sum_{m\in\mathbb{N}}\mathbb{P}^n_{\boldsymbol{\theta}|M=m,Y}\cdot\mathbb{P}_{M=m|Y}{}^n.$$
- Hence, given M, the posterior is
$$\forall j > m, \quad \boldsymbol{\theta}_j|M=m, Y \sim \delta_0,$$
$$\forall j \leq m, \quad \boldsymbol{\theta}_j|M=m, Y \sim \mathcal{N}\left(\frac{Y_j\cdot n\cdot\lambda_j}{1+n\cdot\lambda_j^2}, \frac{1}{1+n\cdot\lambda_j^2}\right).$$

Remark: the family of hierarchical priors with deterministic threshold M is called family of sieve priors.



Existing results

In Johannes et al. [2016], under a pragmatic Bayesian point of view; that is, the existence of a true parameter θ° is accepted; it is shown that, by choosing $\mathbb{P}_M$ suitably:

- the estimator $\widehat{\theta}^{(1)}$ converges with,
  - oracle optimal rate for the quadratic risk which means, $\forall\theta^\circ\in\Theta^\circ, \exists C^\circ\in[1,\infty[: \forall n\in\mathbb{N}, \exists\Phi_n^\circ\in\mathbb{R}:$
  $$\inf_{m\in\mathbb{N}}\mathbb{E}^n_{\theta^\circ}\left[\left\|\tilde{\theta}^m - \theta^\circ\right\|^2\right] \geq \Phi_n^\circ,$$
  $$\mathbb{E}^n_{\theta^\circ}\left[\left\|\widehat{\theta}^{(1)} - \theta^\circ\right\|^2\right] \leq C^\circ\Phi_n^\circ;$$
  - minimax optimal rate for the maximal risk over Θ°, that is to say, $\exists C^\star\in[1,\infty[: \forall n\in\mathbb{N}, \exists\Phi_n^\star\in\mathbb{R}:$
  $$\inf_{\tilde{\theta}}\sup_{\theta^\circ\in\Theta^\circ}\mathbb{E}^n_{\theta^\circ}\left[\left\|\tilde{\theta} - \theta^\circ\right\|^2\right] \geq \Phi_n^\star,$$
  $$\sup_{\theta^\circ\in\Theta^\circ}\mathbb{E}^n_{\theta^\circ}\left[\left\|\widehat{\theta}^{(1)} - \theta^\circ\right\|^2\right] \leq C^\star\Phi_n^\star,$$
  where $\inf_{\tilde{\theta}}$ is taken over all possible estimators of θ°;
- the posterior distribution concentrates with,
  - oracle optimal rate for the quadratic loss which means, $\forall\theta^\circ\in\Theta^\circ, \exists K^\circ\in[1,\infty[:$
  $$\lim_{n\to\infty}\mathbb{E}^n_{\theta^\circ}\left[\mathbb{P}^n_{\boldsymbol{\theta}^1|Y^1}\left(\|\boldsymbol{\theta}-\theta^\circ\|^2 \leq K^\circ\Phi_n^\circ\right)\right] = 1;$$
  - minimax optimal rate Θ°, that is to say, for any unbounded sequence $K_n\in\mathbb{R}^{\mathbb{N}}$:
  $$\lim_{n\to\infty}\sup_{\theta^\circ\in\Theta^\circ}\mathbb{E}^n_{\theta^\circ}\left[\mathbb{P}^n_{\boldsymbol{\theta}^1|Y^1}\left(\|\boldsymbol{\theta}-\theta^\circ\|^2 \leq K_n\Phi_n^\star\right)\right] = 1.$$

Iterated posterior distributions

Note that in the framework of our hierarchical prior, we have:

$$\mathbb{P}^n_{\boldsymbol{\theta}^\eta|Y^\eta} = \sum_{m\in\mathbb{N}}\mathbb{P}^n_{\boldsymbol{\theta}^\eta|M^\eta=m,Y^\eta}\cdot\mathbb{P}^n_{M^\eta=m|Y^\eta},$$
$$\widehat{\theta}^{(\eta)} = \left(\mathbb{E}^n_{\boldsymbol{\theta}^\eta|M^\eta\geq j,Y^\eta}\left[\boldsymbol{\theta}_j\right]\cdot\mathbb{P}^n_{M^\eta|Y^\eta}(M^\eta\geq j)\right)_{j\in\mathbb{N}}.$$

Hence, we first compute $\boldsymbol{\theta}_j^\eta|M^\eta, Y^\eta$:

$$\forall j\in\mathbb{N}, \quad \boldsymbol{\theta}_j^\eta|M^\eta\geq j, Y^\eta \sim\mathcal{N}\left(\frac{\eta\cdot Y_j\cdot n\cdot\lambda_j}{1+\eta\cdot n\cdot\lambda_j^2}, \frac{1}{1+n\cdot\eta\cdot\lambda_j^2}\right),$$
$$\boldsymbol{\theta}_j^\eta|M^\eta < j, Y^\eta \sim\delta_0;$$

and then fix the distribution of M^1: $\forall m\in[\![1,G_n]\!]$,

$$\mathbb{P}_{M^1}(M=m)\propto\exp\left(-3\cdot\eta\cdot\frac{m}{2}\right)\cdot\prod_{j=1}^{m}\left(1+n\cdot\eta\cdot\lambda_j^2\right)^2.$$

Which gives the family of posterior distributions:

$$\mathbb{P}^n_{M^\eta|Y^\eta}(m)\propto\exp\left[-\frac{\eta}{2}\left(3m - \sum_{j=1}^{m}\frac{\eta\left(Y_j\cdot n\cdot\lambda_j^2\right)^2}{1+\eta\cdot n\cdot\lambda_j^2}\right)\right].$$

Self informative limit and model selection

Consider the limit of the family of posteriors as η tends to infinite:

$$\lim_{\eta\to\infty}\mathbb{P}^n_{\boldsymbol{\theta}^\eta|M^\eta=m,Y^\eta} = \delta_{\tilde{\theta}^m},$$

where $\tilde{\theta}^m$ is the projection estimator on the first m dimensions.
The distribution of M tends to a point mass:

$$\lim_{\eta\to\infty}\mathbb{P}^n_{M^\eta|Y^\eta} = \delta_{\widehat{m}},$$

where $\widehat{m}$ is the choice given by the frequentist model selection presented earlier.
The self-informative limit is equal to the model selection estimator, $\widehat{\theta}$, presented above.

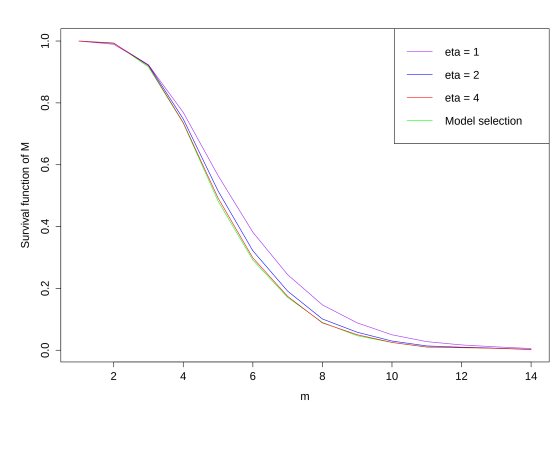


Figure: Survival function of M for different values of η

Notations

Define the following quantities:

$$\mathfrak{b}_m := \sum_{j=m+1}^{\infty}\left(\theta^\circ\right)^2, \quad \Lambda_j := \lambda_j^{-2}, \quad m\cdot\overline{\Lambda}_m := \sum_{j=1}^{m}\Lambda_j,$$

$$m_n^\circ := \underset{m\in[\![1,G_n]\!]}{\arg\min}\left[\mathfrak{b}_m\vee n^{-1}m\overline{\Lambda}_m\right], \quad \Phi_n^\circ := \left[\mathfrak{b}_{m_n^\circ}\vee n^{-1}m_n^\circ\overline{\Lambda}_{m_n^\circ}\right],$$

Set of assumptions

Define the following assumptions:

$$(\mathbb{H}_\lambda) : \exists a\in\mathbb{R}_+, c\geq 1: \quad \forall j\in\mathbb{N}, \quad \left(\frac{1}{c}j^{-a}\leq\lambda_j\leq cj^{-a}\right)$$

$$(\mathbb{H}_1) : 0 < \inf_{n\in\mathbb{N}}\left\{\frac{\left[\mathfrak{b}_{m_n^\circ}\wedge n^{-1}m_n^\circ\overline{\Lambda}_{m_n^\circ}\right]}{\left[\mathfrak{b}_{m_n^\circ}\vee n^{-1}m_n^\circ\overline{\Lambda}_{m_n^\circ}\right]}\right\}\leq 1$$

$$(\mathbb{H}_2) : 0 < \inf_{n\in\mathbb{N}}\left\{\frac{\left[\mathfrak{a}_{m_n^\star}\wedge n^{-1}m_n^\star\overline{\Lambda}_{m_n^\star}\right]}{\left[\mathfrak{a}_{m_n^\star}\vee n^{-1}m_n^\star\overline{\Lambda}_{m_n^\star}\right]}\right\}\leq 1$$

Note that under $(\mathbb{H}_\lambda)$, there exist a constant L such that,

$$\forall m\in\mathbb{N}, \quad \Lambda_m\leq L\overline{\Lambda}_m.$$

Concentration results for the threshold parameter M

For any η in $\overline{\mathbb{N}}$, we have the following results:

1. Under assumptions $(\mathbb{H}_1)$ and $(\mathbb{H}_\lambda)$, define
$$G_n^- := \min\left\{m\in[\![1,m_n^\circ]\!] : \mathfrak{b}_m\leq 9L\Phi_n^\circ\right\},$$
$$G_n^+ := \max\left\{m\in[\![m_n^\circ,G_n]\!] : (m-m_n^\circ)\,n^{-1}\leq 3\Lambda_{m_n^\circ}^{-1}\Phi_n^\circ\right\},$$
and we then have the following concentration for M,
$$\mathbb{P}^n_{M^\eta|Y^\eta}\left[M > G_n^+\right]\leq\exp\left[-\frac{5m_n^\circ}{9L}+\log(G_n)\right],$$
$$\mathbb{P}^n_{M^\eta|Y^\eta}\left[M < G_n^-\right]\leq\exp\left[-\frac{7m_n^\circ}{9}+\log(G_n)\right],$$
this means that M^η tends to select an oracle optimal threshold;
2. whereas under $(\mathbb{H}_2)$ and $(\mathbb{H}_\lambda)$, we define
$$G_n^{\star-} := \min\left\{m\in[\![1,m_n^\star]\!] : \quad \mathfrak{b}_m\leq 9\left(1\vee L^\circ\right)L\Phi_n^\star\right\},$$
$$G_n^{\star+} := \max\left\{m\in[\![m_n^\star,G_n]\!] : (m-m_n^\star)\,n^{-1}\leq 3\Lambda_{m_n^\star}^{-1}\left(1\vee L^\circ\right)\Phi_n^\star\right\},$$
and the following concentration stands,
$$\mathbb{P}^n_{M^\eta|Y^\eta}\left[M > G_n^{\star+}\right]\leq\exp\left[-\frac{5\left(1\vee L^\circ\right)m_n^\star}{9L}+\log(G_n)\right],$$
$$\mathbb{P}^n_{M^\eta|Y^\eta}\left[M < G_n^{\star-}\right]\leq\exp\left[-\frac{7\left(1\vee L^\circ\right)m_n^\star}{9}+\log(G_n)\right],$$
which means that M^η tends to select a minimax optimal threshold.

Concentration results for θ

For any η in $\mathbb{N}$, we have the following results:

1. under assumptions $(\mathbb{H}_1)$ and $(\mathbb{H}_\lambda)$, for all θ° in Θ°, there exist $K^\circ\geq 1$ and $C^\circ > 1$ such that we have
$$\lim_{n\to\infty}\inf_{\mathbb{Q}_{\boldsymbol{\theta}}}\mathbb{E}^n_{\theta^\circ}\left[\mathbb{Q}^n_{\boldsymbol{\theta}|Y}\left(\|\boldsymbol{\theta}-\theta^\circ\|^2\geq\Phi_n^\circ\right)\right] = 1,$$
$$\lim_{n\to\infty}\mathbb{E}^n_{\theta^\circ}\left[\mathbb{P}^n_{\boldsymbol{\theta}^\eta,M^\eta|Y^\eta}\left(\left(K^\circ\right)^{-1}\Phi_n^\circ\leq\|\boldsymbol{\theta}-\theta^\circ\|^2\leq K^\circ\Phi_n^\circ\right)\right] = 1,$$
$$\mathbb{E}^n_{\theta^\circ}\left[\|\widehat{\theta}^{(\eta)}-\theta^\circ\|^2\right]\leq C^\circ\Phi_n^\circ,$$
where $\inf_{\mathbb{Q}_{\boldsymbol{\theta}}}$ is taken over all possible sieve priors; establishing oracle optimal concentration and convergence of the posterior and Bayes estimate, respectively;
2. whereas under $(\mathbb{H}_2)$ and $(\mathbb{H}_\lambda)$, for a finite constant $C^\star\geq 1$ and any unbounded sequence K_n, we have
$$\lim_{n\to\infty}\inf_{\mathbb{Q}_{\boldsymbol{\theta}}}\sup_{\theta^\circ\in\Theta^\circ}\mathbb{E}^n_{\theta^\circ}\left[\mathbb{Q}^n_{\boldsymbol{\theta}|Y}\left(\|\boldsymbol{\theta}-\theta^\circ\|^2\geq\Phi_n^\star\right)\right] = 1,$$
$$\lim_{n\to\infty}\sup_{\theta^\circ\in\Theta^\circ}\mathbb{E}^n_{\theta^\circ}\left[\mathbb{P}^n_{\boldsymbol{\theta}^\eta,M^\eta|Y^\eta}\left(\|\boldsymbol{\theta}-\theta^\circ\|^2\leq K_n\Phi_n^\star\right)\right] = 1,$$
$$\sup_{\theta^\circ\in\Theta^\circ}\mathbb{E}^n_{\theta^\circ}\left[\|\widehat{\theta}^{(\eta)}-\theta^\circ\|^2\right]\leq C^\star\Phi_n^\star,$$
where $\inf_{\mathbb{Q}_{\boldsymbol{\theta}}}$ is taken over all possible sieve priors; establishing minimax optimal concentration and convergence of the posterior and Bayes estimate, respectively.

Note that in the case of $\eta\to\infty$, those results are still true and that the concentration corresponds to the convergence in probability as

$$\lim_{\eta\to\infty}\mathbb{E}^n_{\theta^\circ}\left[\mathbb{P}^n_{\boldsymbol{\theta}^\eta,M^\eta|Y^\eta}\left(\|\boldsymbol{\theta}-\theta^\circ\|^2\leq K_n\Phi_n\right)\right] = \mathbb{P}^n_{\theta^\circ}\left[\|\widehat{\theta}-\theta^\circ\|^2\leq K_n\Phi_n\right].$$

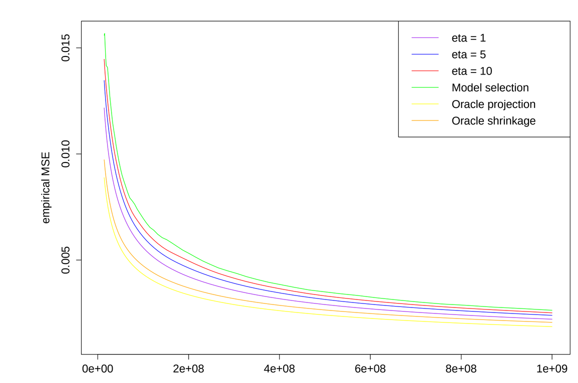


Figure: Estimated mean of the quadratic error of the Bayes estimate for θ° polynomial.

Bibliography

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