

1 Convexity

- A function is convex iff a line segment between two points on the function's graph always lies above the function.
- A function $h(u)$, $u \in \mathbb{R}^2$ is convex if $\forall u, v \in \mathbb{R}^2, 0 \leq \lambda \leq 1$: $h(\lambda u + (1-\lambda)v) \leq \lambda h(u) + (1-\lambda)h(v)$ (Strictly convex if $\leq \Rightarrow <$)
- A strictly convex function has a unique global minimum w^* . For convex functions, every local minimum is a global minimum.
- Sums of cvx fcn's \Rightarrow cvx (MSE+Lin model = cvx)

2 Optimization

- Local min $w^* \Rightarrow \exists \epsilon > 0$ s.t. $\mathcal{L}(w^*) \leq \mathcal{L}(w) \forall w / \|w - w^*\| < \epsilon$
- Global minimum w^* , $\mathcal{L}(w^*) \leq \mathcal{L}(w) \forall w \in \mathbb{R}^D$
- **Gradient Descent (GD)**: $w^{(t+1)} := w^{(t)} - \gamma \nabla \mathcal{L}(w^{(t)})$, $\gamma > 0$
- **Gradient Descent for Linear MSE**, $\mathcal{O}(N \times D)$
- $\mathcal{L}(w) := \frac{1}{2N} \sum_{n=1}^N (y_n - x_n^T w)^2 = \frac{1}{2N} e^T e$, $\nabla \mathcal{L}(w) = -\frac{1}{N} X^T e$
- $e = y - Xw$

Stochastic Gradient Descent (SGD), $\mathcal{O}(D)$

- $\mathcal{L}(w) := \frac{1}{N} \sum_{n=1}^N \mathcal{L}_n(w)$, $w^{(t+1)} := w^{(t)} - \gamma \nabla \mathcal{L}_n(w^{(t)})$
- Cheap and unbiased estimator of the gradient $E[\nabla \mathcal{L}_n(w)] = \nabla \mathcal{L}(w)$
- $\nabla \mathcal{L}_n(w) := \nabla (\frac{1}{2} (y - x^T w)^2) = \nabla \mathcal{L}(n)$
- **Mini-batch SGD**: $w^{(t+1)} := w^{(t)} - \gamma g$
- $g := \frac{1}{N_{\text{batch}}} \sum_{n \in B} \nabla \mathcal{L}_n(w^{(t)})$, subset $B \subseteq [N]$ of training samples

Non-Smooth Optimization

- Convexity, for differentiable functions: $\mathcal{L}(w) \geq \mathcal{L}(w) + \nabla \mathcal{L}(w)^T (u - w) \forall u, w$
- A vector $g \in \mathbb{R}^D$ s.t. $\mathcal{L}(u) \geq \mathcal{L}(w) + g^T (u - w) \forall u$ is called a subgradient to the function \mathcal{L} at w .
- This definition makes sense for objectives \mathcal{L} which are not necessarily differentiable (and not even necessarily convex).
- \mathcal{L} convex and differentiable at $w \rightarrow$ only subgradient $g = \nabla \mathcal{L}(w)$.

Subgradient Descent

- Intersections of convex sets are convex
- Projections onto convex sets are unique.
- Formal definition: $P_C(w) := \arg \min_{u \in C} \|w - u\|$.
- Projected Gradient Descent, SGD with Projection step
- Idea: add a projection onto C after every step: $P_C(w^{(t)}) := \arg \min_{u \in C} \|w^{(t)} - u\|$, $w^{(t+1)} := P_C(w^{(t)} - \gamma \nabla \mathcal{L}(w^{(t)}))$
- Constrained \rightarrow Unconstrained Problems**
- $\min_{u \in C} \mathcal{L}(w) \sim \min_w \mathcal{L}(w) + P(w)$

Implementation Issues

Optimality: 1st order: if \mathcal{L} is cvx and $\nabla \mathcal{L}(w^*) = 0 \rightarrow$ global 2nd order: \mathcal{L} potentially nonconv, if $\nabla \mathcal{L}(w^*) = 0$, $\nabla^2 \mathcal{L}(w^*) \geq 0 \rightarrow w$ local minimum ($\nabla^2 \mathcal{L}(w) = \frac{\partial^2 \mathcal{L}}{\partial w \partial w^T}(w)$ expensive)

3 Least Squares (Linear regression + MSE)

- $\mathcal{L}(w) := \frac{1}{2N} \sum_{n=1}^N (y_n - x_n^T w)^2 = \frac{1}{2N} (y - Xw)^T (y - Xw)$.
- **Proof of convexity**:
- 1) Composition of a linear function with a convex function (the square function).
- 2) Verify definition, for any $\lambda \in [0, 1]$ and w, w' : $\mathcal{L}(\lambda w + (1 - \lambda)w') - (\lambda \mathcal{L}(w) + (1 - \lambda)\mathcal{L}(w')) \leq 0$
- $LHS = \frac{1}{2N} \lambda(1 - \lambda) (Xw - Xw')^T (Xw - Xw') \leq 0$
- 3) Hessian positive semidefinite (all its eigenvalues are non-negative)
- $H(w) = \frac{1}{N} \nabla^2 ((y - Xw)^T (y - Xw)) = \frac{1}{N} X^T X$

Singular value decomposition (SVD): $X = USV^T$ and V and S are orthogonal matrices, and S is a diagonal matrix with the singular values σ_i on the diagonal.

- Now find its minimum $\nabla \mathcal{L}(w) = -\frac{1}{N} X^T (y - Xw) = 0$

Closed form (if Gram matrix $X^T X \in \mathbb{R}^{D \times D}$ is invertible)

- $w^* = (X^T X)^{-1} X^T y$
- Invertibility and Uniqueness**
- $X^T X \in \mathbb{R}^{D \times D}$ invertible iff $\text{rank}(X) = D$.
- Proof: Assume $\text{rank}(X) < D \Rightarrow \exists u \neq 0$ s.t. $Xu = 0 \Rightarrow X^T Xu = 0 \Rightarrow \text{rank}(X^T X) < D \Rightarrow X^T X$ is not invertible.
- Rank deficiency and III-Conditioning
- In practice, X is often rank deficient.
- If $D > N$, we always have $\text{rank}(X) < D$ (row rank = col. rank)
- If $D \leq N$, but some of the columns x_n are (nearly) collinear, then the matrix is illconditioned \rightarrow numerical issues

4 Maximum Likelihood

Gaussian distribution and independence

$p(y | \mu, \Sigma) = \mathcal{N}(y | \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp(-\frac{(y-\mu)^2}{2\sigma^2})$, Σ psd.

$\mathcal{N}(y | \mu, \Sigma) = \frac{1}{\sqrt{(2\pi)^D \det(\Sigma)}} \exp(-\frac{1}{2} (y - \mu)^T \Sigma^{-1} (y - \mu))$

- Two RVs X, Y independent when $p(x, y) = p(x)p(y)$.

A probabilistic model for least-squares

- $y_n = x_n^T w + \epsilon_n$ where the ϵ_n zero-mean Gaussian RV
- $p(y | X, w) = \prod_{n=1}^N p(y_n | x_n, w) = \prod_{n=1}^N \mathcal{N}(y_n | x_n^T w, \sigma^2)$

Defining cost with log-likelihood

$\mathcal{L}_{LL}(w) := \log p(y | X, w) = -\frac{1}{2N} \sum_{n=1}^N (y_n - x_n^T w)^2 + \text{cnst.}$

Maximum-likelihood estimator (MLE)

$\arg \min_w \mathcal{L}_{MSE}(w) = \arg \max_w \mathcal{L}_{LL}(w)$.

Properties of MLE

- MLE is a sample approximation to the expected log-likelihood: $\mathcal{L}_{LL}(w) \approx \mathbb{E}_{p(y_n | x_n, w)} [\log p(y_n | x_n, w)]$
- MLE is consistent, i.e., it will give us the correct model assuming that we have a sufficient amount of data. $w_{MLE} \xrightarrow{P} w_{true}$
- The MLE is asymptotically normal, i.e., $(w_{MLE} - w_{true}) \xrightarrow{d} \frac{1}{\sqrt{N}} \mathcal{N}(w_{MLE} | 0, F^{-1}(w_{true}))$

where $F(w) = -\mathbb{E}_{p(y)} [\frac{\partial^2 \mathcal{L}}{\partial w \partial w^T}]$ is the Fisher information.

- MLE is efficient, i.e. it achieves the Cramer-Rao lower bound. Covariance ($w_{MLE} = F^{-1}(w_{true})$)

Laplace distribution $p(y_n | x_n, w) = \frac{1}{2b} e^{-\frac{1}{2b} |y_n - x_n^T w|}$

5 Regularization: $\min_w \mathcal{L}(w) + \Omega(w)$

Ridge Regression: L_2 -Regularization + \mathcal{L}_{MSE}

- Euclidean norm (L_2 norm) $\Omega(w) = \|w\|_2^2, \|w\|_2^2 = \sum_i w_i^2$.
- $\min_w \frac{1}{2N} \sum_{n=1}^N |y_n - x_n^T w|^2 + \lambda \|w\|_2^2 = \min_w \frac{1}{2N} \|y - Xw\|_2^2 + \lambda \|w\|_2^2$

Explicit solution for w :

$\nabla = \nabla \mathcal{L} + \nabla \Omega = -\frac{1}{N} X^T (y - Xw) + \lambda 2w = 0$

$\Rightarrow w_{ridge} = (X^T X + \lambda I)^{-1} X^T y$, ($\frac{\lambda}{2N} = \lambda$)

Ridge Regression Fights III-Conditioning

- Lifting the eigenvalues: The eigenvalues of $(X^T X + \lambda I)$ are all at least λ' and so the inverse always exists.
- Proof: Write the Eigenvalue decomposition of $X^T X$ as USU^T , $X^T X + \lambda I = USU^T + \lambda U I U^T = U[S + \lambda I]U^T$
- L_1 -Regularization: The Lasso (L_1 -norm + \mathcal{L}_{MSE})**
- $\min_w \frac{1}{2N} \sum_{n=1}^N |y_n - x_n^T w|^2 + \lambda \|w\|_1$ where $\|w\|_1 := \sum_i |w_i|$

6 Model Selection

Generalization Error

- Expected error over all samples drawn from distribution \mathcal{D} : $L_{\mathcal{D}}(f) = \mathbb{E}_{(x,y) \sim \mathcal{D}} \ell(y, f(x))$
- approximate the true error by averaging the loss function over the dataset: $L_S(f) = \frac{1}{|S|} \sum_{(x_n, y_n) \in S} \ell(y_n, f(x_n))$, (RV)
- unbiased estimator of the true error
- Law of large number: $L_S(f) \xrightarrow{|S| \rightarrow \infty} L_{\mathcal{D}}(f)$ but fluctuations!

Generalization gap $|L_{\mathcal{D}}(f) - L_S(f)|$

- **Claim**: given a model f and a test set $S_{test} \sim \mathcal{D}$ i.i.d. (not used to learn f) and a loss $\ell(\cdot, \cdot) \in [a, b]$,
- $\mathbb{P}[|L_{\mathcal{D}}(f) - L_{S_{test}}(f)| \geq \sqrt{\frac{(b-a)^2 \ln(2/|\mathcal{D}|)}{2|S_{test}|}}] \leq \delta$
- Error \downarrow as $\mathcal{O}(1/\sqrt{|S_{test}|})$ with nb. of test points
- High probability bound: δ is only in the ln

The proof relies only on concentration inequalities

- $(\epsilon_n, y_n) \in S_{test}$ are chosen independently, the associated losses $\ell(y_n, f(x_n)) \in [a, b]$ given a fixed model f , are also i.i.d. random variables.
- Empirical loss: $\frac{1}{N} \sum_{n=1}^N \Theta_n = \frac{1}{N} \sum_{n=1}^N \ell(y_n, f(x_n)) = L_{S_{test}}(f)$
- True loss: $\mathbb{E}[\Theta_n] = \mathbb{E}[\ell(y_n, f(x_n))] = L_{\mathcal{D}}(f)$
- Hoeffding inequality: a simple concentration bound**
- **Claim**: Let $\Theta_1, \dots, \Theta_N$ be a sequence of i.i.d. random variables with mean $\mathbb{E}[\Theta]$ and range $[a, b]$. Then, for any $\epsilon > 0$
- $\mathbb{P}[\frac{1}{N} \sum_{n=1}^N \Theta_n - \mathbb{E}[\Theta] \geq \epsilon] \leq 2e^{-2N\epsilon^2/(b-a)^2}$
- Concentration bound: the empirical mean is concentrated around its mean
- A. Use it with $\Theta_n = \ell(y_n, f(x_n))$ B. Equating $\delta = 2e^{-2|S_{test}| \epsilon^2/(b-a)^2}$

we get $\epsilon = \sqrt{\frac{(b-a)^2 \ln(2/|\mathcal{D}|)}{2|S_{test}|}}$

How far is each of the K test series $L_{S_{test}}(f_k)$ from the true $L_{\mathcal{D}}(f_k)$?

- **Claim**: we can bound the maximum deviation for all K candidates, by $\mathbb{P}[\max_k |L_{\mathcal{D}}(f_k) - L_{S_{test}}(f_k)| \geq \sqrt{\frac{(b-a)^2 \ln(2K/\delta)}{2|S_{test}|}}] \leq \delta$

- The error decreases as $\mathcal{O}(1/\sqrt{|S_{test}|})$ with the number test points
- test K hyper-parameters, the error only goes up by $\sqrt{\ln(K)}$
- can test many hyperparameters without incurring a large penalty
- It can be extended to infinitely many models
- Proof: A simple union bound**
- Special case $K = 1$ $\mathbb{P}[\max_k |L_{\mathcal{D}}(f_k) - L_{S_{test}}(f_k)| \geq \epsilon] = \mathbb{P}[|L_{\mathcal{D}}(f) - L_{S_{test}}(f)| \geq \epsilon] \leq \sum_k \mathbb{P}[|L_{\mathcal{D}}(f_k) - L_{S_{test}}(f_k)| \geq \epsilon] \leq 2Ke^{-2N\epsilon^2/(b-a)^2}$
- $\delta = 2Ke^{-2N\epsilon^2/(b-a)^2} \Rightarrow \epsilon = \sqrt{\frac{(b-a)^2 \ln(2K/\delta)}{2N}}$
- Let $k^* = \arg \min_k L_{\mathcal{D}}(f_k)$ (smallest true risk) and $\hat{k} = \arg \min_k L_{S_{test}}(f_k)$ (smallest empirical risk) then:
- $\mathbb{P}[L_{\mathcal{D}}(f_{\hat{k}}) \geq L_{\mathcal{D}}(f_{k^*}) + 2\sqrt{\frac{(b-a)^2 \ln(2K/\delta)}{2|S_{test}|}}] \leq \delta$

7 Bias Variance Decomposition

- Data model: output perturbed by some noise $y = f(x) + \epsilon$, $\epsilon \sim \mathcal{D}$, \mathcal{D} s.i.d. independent of x $\mathbb{E}[\epsilon] = 0$
- Error Decomposition** $\mathbb{E}_{x,y \sim \mathcal{D}} [(y - f_S(x))^2]$
- Consider the expected error of f_S for a fixed element x_0 : $L(f_S) = \mathbb{E}_{x,y \sim \mathcal{D}} [(f(x_0) + \epsilon - f_S(x_0))^2]$, (RV from train set S)
- Run experiment many times \rightarrow average and the variance of the predictions $\{f_S, \dots, f_S\}$ over these multiple runs
- A decomposition in three terms**
- Rank deficiency and III-Conditioning
- $\mathbb{E}_{x,y \sim \mathcal{D}} [(f(x_0) - \mathbb{E}_{S \sim \mathcal{D}} [f_S(x_0)])^2] = \mathbb{E}_{x,y \sim \mathcal{D}} [(f(x_0) - \mathbb{E}_{S \sim \mathcal{D}} [f_S(x_0)])^2]$
- $\mathbb{E}_{x,y \sim \mathcal{D}} [(f(x_0) - \mathbb{E}_{S \sim \mathcal{D}} [f_S(x_0)])^2] = \mathbb{E}_{x,y \sim \mathcal{D}} [e^2] + 2\mathbb{E}_{x,y \sim \mathcal{D}} [e(f(x_0) - f_S(x_0))] + \mathbb{E}_{S \sim \mathcal{D}} [(f(x_0) - f_S(x_0))^2]$
- Using that $\mathbb{E}_{x,y \sim \mathcal{D}} [e] = 0$ and $\mathbb{E} \perp \perp S$:

- $\bullet \mathbb{E}_{x,y \sim \mathcal{D}} [e^2] = \text{Var}_{x,y \sim \mathcal{D}} [e]$ $\bullet \mathbb{E}_{x,y \sim \mathcal{D}} [e(f(x_0) - f_S(x_0))] = \mathbb{E}_{x,y \sim \mathcal{D}} [e] \mathbb{E}_{x,y \sim \mathcal{D}} [f(x_0) - f_S(x_0)] = 0$
- $\bullet \mathbb{E}_{x,y \sim \mathcal{D}} [e^2] = \text{Var}_{x,y \sim \mathcal{D}} [e]$ $\bullet \mathbb{E}_{x,y \sim \mathcal{D}} [e(f(x_0) - f_S(x_0))] = 0$
- $\bullet \mathbb{E}_{x,y \sim \mathcal{D}} [e^2] = \text{Var}_{x,y \sim \mathcal{D}} [e]$ $\bullet \mathbb{E}_{x,y \sim \mathcal{D}} [e(f(x_0) - f_S(x_0))] = 0$
- Trick: we add and subtract the constant term $\mathbb{E}_{S \sim \mathcal{D}} [f_S(x_0)]$, where S' is a second training set independent from S
- $\mathbb{E}_{S \sim \mathcal{D}} [(f(x_0) - f_S(x_0))^2] = \mathbb{E}_{S \sim \mathcal{D}} [(f(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)] + \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)] - f_S(x_0))^2]$
- $= \mathbb{E}_{S \sim \mathcal{D}} [(f(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)])^2] + \mathbb{E}_{S \sim \mathcal{D}} [f_S(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)] - f_S(x_0)]^2$
- $+ 2(f(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)]) \cdot \mathbb{E}_{S \sim \mathcal{D}} [f_S(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)] - f_S(x_0)]$
- Cross-term:
- $\mathbb{E}_{S \sim \mathcal{D}} [(f(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)]) \cdot \mathbb{E}_{S \sim \mathcal{D}} [f_S(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)] - f_S(x_0)] = \mathbb{E}_{S \sim \mathcal{D}} [(f(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)]) \cdot \mathbb{E}_{S \sim \mathcal{D}} [\mathbb{E}_{S' \sim \mathcal{D}} [f_S(x_0)] - f_S(x_0)] = (f(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)]) \cdot \mathbb{E}_{S \sim \mathcal{D}} [\mathbb{E}_{S' \sim \mathcal{D}} [f_S(x_0)] - f_S(x_0)] = 0$
- $\Rightarrow \mathbb{E}_{S \sim \mathcal{D}} [(f(x_0) - f_S(x_0))^2] = (f(x_0) - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)])^2 + \mathbb{E}_{S \sim \mathcal{D}} [\mathbb{E}_{S' \sim \mathcal{D}} [f_S(x_0)] - f_S(x_0)]^2$

Bias-Variance Decomposition

- We obtain the following decomposition into three positive terms
- $= \mathbb{E}_{S \sim \mathcal{D}} [e_{\text{bias}}^2] + \mathbb{E}_{S \sim \mathcal{D}} [e_{\text{var}}^2] + \mathbb{E}_{S \sim \mathcal{D}} [e_{\text{noise}}^2]$
- $= \mathbb{E}_{S \sim \mathcal{D}} [(f(x_0) + \epsilon - \mathbb{E}_{S' \sim \mathcal{D}} [f_{S'}(x_0)] - f_S(x_0))^2] + \mathbb{E}_{S \sim \mathcal{D}} [e_{\text{noise}}^2]$
- $= \text{Noise variance} + \text{Bias} + \text{Var.}$

Bias Variance tradeoff

8 Classification

- Loss function: (0-1 Loss) $\ell(y, y') = 1_{y \neq y'} = \begin{cases} 1 & \text{if } y \neq y' \\ 0 & \text{if } y = y' \end{cases}$
- True risk for the classification: $L_{\mathcal{D}}(f) = \mathbb{E}_{\mathcal{D}} [1_{y \neq f(x)}] = \mathbb{P}_{\mathcal{D}}[Y \neq f(X)]$
- minimize $L_{\mathcal{D}}(f)$
- Bayes classifier f , $f_S = \arg \min L_{\mathcal{D}}(f)$**
- **Claim**: $f_S(x) = \arg \max_{y \in \{1,-1\}} \mathbb{P}(y = y' | X = x)$
- Proof of the Bayes classifier**
- **Claim 1**: $\forall x \in \mathcal{X}$, $f_S(x) \in \arg \min_{y \in \mathcal{Y}} \mathbb{P}(Y \neq y | X = x) \Rightarrow f_S \in \arg \min_{f \in \mathcal{F}} L_{\mathcal{D}}(f)$
- **Claim 2**: $f_S(x) = \arg \min_{f \in \mathcal{F}} \mathbb{P}(Y \neq y | X = x)$

Classification by empirical risk minimization

$\min_{f \in \mathcal{F}} L_{\mathcal{D}}(f) = \min_{f \in \mathcal{F}} \frac{1}{N} \sum_{n=1}^N 1_{f(x_n) \neq y_n} = \frac{1}{N} \sum_{n=1}^N 1_{y_n f(x_n) < 0}$

- **PROBLEM**: L_{train} is not convex because \mathcal{Y} is discrete

The set of classifiers is not convex because \mathcal{Y} is discrete

The indicator function 1 is not convex because it is not continuous

9 Logistic Regression

Logistic function: $\sigma(\eta) := \frac{1}{1+e^{-\eta}}$, $1 - \sigma(\eta) = \frac{1}{1+e^{\eta}}$

$\sigma'(\eta) = \sigma(\eta)(1 - \sigma(\eta))$

Logistic Regression

$p(1 | x) := \mathbb{P}(Y = 1 | X = x) = \sigma(x^T w + w_0)$

$p(0 | x) := \mathbb{P}(Y = 0 | X = x) = 1 - \sigma(x^T w + w_0)$

If $p(1 | x) \geq 1/2$, you predict the class 1, else class 0

MLE for logistic regression

- Assumption: The inputs X do not depend on the parameter w we choose:
- $\mathcal{L}(w) := p(Y | X, w) = p(X | w)p(Y | X, w)$
- $\mathbb{X} = \begin{bmatrix} x_1^T \\ \vdots \\ x_N^T \end{bmatrix}$, $p(X | w) = \prod_{n=1}^N p(y_n | x_n, w) = \prod_{n=1}^N p(y_n | x_n, w)^{y_n} [1 - p(x_n^T w)]^{1-y_n}$
- $\mathcal{L}(w) \propto \prod_{n=1}^N \sigma(x_n^T w)^{y_n} [1 - \sigma(x_n^T w)]^{1-y_n}$

Minimum of the Negative Log Likelihood (NLL)

- $\log(p(Y | X, w)) = -\log(\prod_{n=1}^N \sigma(x_n^T w)^{y_n} [1 - \sigma(x_n^T w)]^{1-y_n})$
- $= \sum_{n=1}^N -y_n x_n^T w + \log(1 + e^{x_n^T w})$
- $w = \arg \min L(w) := \frac{1}{N} \sum_{n=1}^N -y_n x_n^T w + \log(1 + e^{x_n^T w})$
- NLL is equivalent to ERM for the logistic loss
- $y \in \{0, 1\} : \ell(y, g(x)) = -y g(x) + \log(1 + \exp(g(x)))$
- $y \in \{-1, 1\} : \ell(y, g(x)) = \log(1 + \exp(-y g(x)))$

Gradient of the negative log likelihood

$\nabla L(w) = \nabla \left[\frac{1}{N} \sum_{n=1}^N \log(1 + e^{x_n^T w}) - y_n x_n^T w \right] = \frac{1}{N} \sum_{n=1}^N \frac{e^{x_n^T w}}{1+e^{x_n^T w}} - y_n x_n = \frac{1}{N} \sum_{n=1}^N \sigma(x_n^T w) (x_n^T w - y_n) x_n$

$\nabla L(w) = \frac{1}{N} X^T \sigma(Xw) - y$

Convexity of the loss function L

- The Hessian $\nabla^2 L, \frac{\partial^2}{\partial w \partial w^T} L(w)$, is psd \Rightarrow convex
- $\nabla^2 L(w) = \frac{1}{N} \sum_{n=1}^N \sigma(x_n^T w) (1 - \sigma(x_n^T w)) x_n x_n^T$
- $\nabla^2 L(w) = \frac{1}{N} X^T S X$ where $S = \text{diag} \{ \sigma(x_1^T w) (1 - \sigma(x_1^T w)) \}$ $\triangleright 0 \Rightarrow L$ is convex since $\nabla^2 L(w) \geq 0$
- Newton's method** uses second order information
- Newton's method minimizes the quadratic approximation: $L(w) \approx L(w_0) + \nabla L(w_0)^T (w - w_0) + \frac{1}{2} (w - w_0)^T \nabla^2 L(w_0) (w - w_0) = \phi_1(w)$
- $\phi_1(w) = \arg \min \phi_1(w) \Rightarrow \nabla L(w_0) + \nabla^2 L(w_0) (\tilde{w} - w_0) = 0$
- Newton's method: $w_{t+1} = w_t - \gamma_t \nabla^2 L(w_t)^{-1} \nabla L(w_t)$
- Step-size needed to ensure convergence (damped Newton's method)
- Convergence faster than GD but comp. complex. higher.

Problem when the data are linearly separable

$\inf_w L(w) = 0 = \lim_{n \rightarrow \infty} \min_{w \in \mathbb{R}^D} \alpha(w) \Rightarrow$ the weights will go to ∞

- Solution: add a ℓ_2 -regularization (Ridge logistic regression)
- $\frac{1}{N} \sum_{n=1}^N -y_n x_n^T w + \log(1 + e^{x_n^T w}) + \frac{\lambda}{2} \|w\|_2^2$
- Optimization perspective: stabilize the optimization process
- Statistical perspective: avoid overfitting

10 Support Vector Machines

- Define a hyperplane as $\{x : w^T x = 0\}$ where $\|w\| = 1$
- Prediction: $f(x) = \text{sign}(x^T w)$
- **Claim**: distance x_0 -hyperplane defined by w is $|w^T x_0|$
- **Proof**: $\min_{u \perp w, \|u\| = \|w\|} \|x_0 - u\|$. Let $v = x_0 - w^T x_0 w$ then by the Pythagorean theorem for any u s.t. $w^T u = 0$, $\|x_0 - u\|^2 = \|x_0 - v\|^2 + \|v - u\|^2$
- **Hard-SVM rule**: max-margin separating hyperplane
- Margin of a hyperplane: $\min_{x \in \mathcal{X}} |w^T x_n|$
- Max-margin separating hyperplane: $\max_{x_n \in \mathcal{X}} \min_{|w| = 1} |w^T x_n| = \min_{|w| = 1} \max_{x_n \in \mathcal{X}} |w^T x_n| \geq 0$
- Equivalent to: $\max_{M \in \mathbb{R}^{N \times N}} \min_{|w| = 1} M$ s.t. $\forall n, y_n x_n^T w \geq M$
- Also equivalent to: $\min_{|w| = 1} \frac{1}{2} \|w\|^2$ such that $\forall n, y_n x_n^T w \geq 1$

Soft SVM: when training set is not linearly separable

- Maximize margin but allow some constraints to be violated
- Introduce positive slack variables ξ_1, \dots, ξ_N and replace the constraints with $y_n x_n^T w \geq 1 - \xi_n$
- $\min_{w, \xi} \frac{1}{2} \|w\|^2 + \frac{1}{N} \sum_{n=1}^N \xi_n$ s.t. $\forall n, y_n x_n^T w \geq 1 - \xi_n$ and $\xi_n \geq 0$
- Equivalent: $\min_{w, \xi} \frac{1}{2} \|w\|^2 + \frac{1}{N} \sum_{n=1}^N (1 - y_n x_n^T w)_+$ (a hinge loss)
- Proof**: 1) If $y_n x_n^T w \geq 1$, then $\xi_n = 0$, 2) If $y_n x_n^T w < 1$, $\xi_n = 1 - y_n x_n^T w$, therefore $\xi_n = (1 - y_n x_n^T w)_+$

ERM for the hinge loss with ridge regularization

- $\min_w \frac{1}{2} \|w\|^2 + \frac{1}{N} \sum_{n=1}^N (1 - y_n x_n^T w)_+$
- 1) Choose direction of w s.t. $w^T x_n$ acts as a separating hyperplane
- 2) Adjust scale of w to ensure that no point lies with the margin
- 3) Select the hyperplane with the largest margin
- Optimization: Convex (but non-smooth) objective which can be minimized with: 1) SubGD method, 2) StochSubGD method

Convex duality

- Define $G(w, \alpha) \text{ s.t. } \min_w L(w) = \min_w \max_{\alpha} G(w, \alpha)$
- Primal pb: $\min_w \max_{\alpha} G(w, \alpha)$, dual pb: $\max_{\alpha} \min_w G(w, \alpha)$
- $G(w, \alpha)$ for SVM
- $[z]_+ = \max(0, z) = \max_{\alpha \in [0, 1]} \alpha z$
- $(1 - y_n x_n^T w)_+ = \max_{\alpha_n \in [0, 1]} \alpha_n (1 - y_n x_n^T w)$
- $\min_w L(w) = \min_w \max_{\alpha \in [0, 1]} \frac{1}{N} \sum_{n=1}^N \alpha_n (1 - y_n x_n^T w) + \frac{\lambda}{2} \|w\|_2^2$
- The function G is convex in w and concave in α
- Min max interchangeable**
- $\max_w \min_{\alpha} G(w, \alpha) \leq \min_w \max_{\alpha} G(w, \alpha)$
- Equality if G is convex in w , concave in α and the domains of w and α are convex and compact
- **Proof**: $\min_w G(w, \alpha) \leq G(w, \alpha) \forall w, \alpha$
- $\max_w \min_{\alpha} G(w, \alpha) \leq \max_{\alpha} G(w, \alpha) \forall w, \alpha$
- $\max_w \min_{\alpha} G(w, \alpha) \leq \min_w \$

