TEMPORAL EXPONENTIAL-FAMILY RANDOM GRAPH MODELING (TERGMS) WITH STATNET

Prof. Steven Goodreau

Prof. Martina Morris

Prof. Michal Bojanowski

Prof. Mark S. Handcock



Source for all things STERGM

Pavel N. Krivitsky and Mark S. Handcock (2014). <u>A Separable Model for Dynamic Networks</u>. *Journal of the Royal Statistical Society, Series B*, Volume 76, Issue 1, pages 29–46.

Terminology

- The phrase "temporal ERGMs," or TERGMs, refers to all ERGMs that are dynamic
- The specific class of TERGMs that have been implemented thus far are called "separable temporal ERGMs," or STERGMs
- In the relevant R package, we left open the possibility that we would develop more in the future
- Thus:

	Cross-sectional	Dynamic
Name of package	ergm	tergm
Name of function in package	ergm	stergm

ERGMs: Review

Probability of observing a graph (set of relationships) y on a fixed set of nodes:

$$P(Y = y \mid) = \frac{\exp(\boldsymbol{\theta}' \boldsymbol{g}(\boldsymbol{y}))}{k(\boldsymbol{\theta})}$$

Conditional log-odds of a tie

$$logit(P(Y_{ij} = 1 | rest of the graph)) = log(\frac{P(Y_{ij} = 1 | rest of the graph)}{P(Y_{ij} = 0 | rest of the graph)})$$

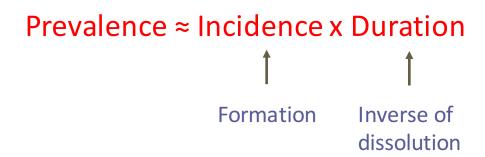
$$= \theta' \partial(g(y))$$

 $\mathbf{k}(\theta)$ = numerator summed over all possible networks on node set y $\partial(\mathbf{g}(\mathbf{y}))$ represents the change in $\mathbf{g}(\mathbf{y})$ when Y_{ij} is toggled between 0 and 1

- ERGMs are great for modeling cross-sectional network structure
- But they can only predict the presence of a tie; they are unable to separate the processes of tie formation and dissolution
- Why separate formation from dissolution?

- Intuition: The social forces that facilitate formation of ties are often different from those that facilitate their dissolution.
- Interpretation: Because of this, we would want model parameters to be interpreted in terms of ties formed and ties dissolved.
- Simulation: We want to be able to control cross-sectional network structure and relational durations separately in our disease simulations, matching both to data

- E.g. if a particular type of tie is rare in the cross-section, is that because:
 - They form infrequently?
 - They form frequently, but then dissolve frequently as well?
- The classic approximation formula from epidemiology helps us see the basic relationship among our concepts:



Core idea:

- The y_{ij} values (ties in the network) and Y (the set of all y_{ij} values) are now indexed by time
- Represent evolution from Y_t to Y_{t+1} as a product of two phases: one in which ties are formed and another in which they are dissolved, with each phase a draw from an ERGM.
- Thus, two formulas: a formation formula and a dissolution formula
- And, two corresponding sets of statistics

ERGM: Conditional log-odds of a tie existing

$$logit(P(Y_{ij} = 1 | rest of the graph)) = \theta' \partial(g(y))$$

STERGM: Conditional log-odds of a tie forming (formation model):

$$logit(P(Y_{ij,t+1} = 1 | Y_{ij,t} = 0, rest of the graph)) = \theta^{+\prime} \partial(g^+(y))$$

STERGM: Conditional log-odds of a tie *persisting* (dissolution model):

$$logit\left(P(Y_{ij,t+1}=1 | Y_{ij,t}=1, \text{ rest of the graph})\right) = \boldsymbol{\theta}^{-\prime} \boldsymbol{\partial} (\boldsymbol{g}^{-}(\boldsymbol{y}))$$

where: $g^+(y)$ = vector of network statistics in the formation model θ^+ = vector of parameters in the formation model

 $g^{-}(y)$ = vector of network statistics in the dissolution model

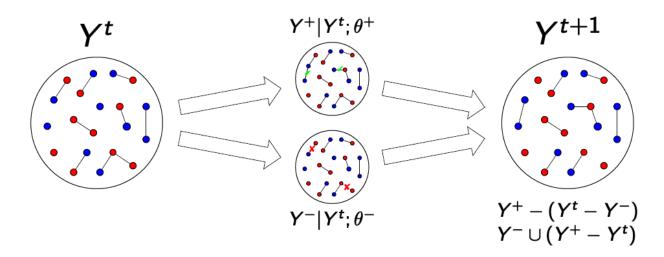
 θ^- = vector of parameters in the dissolution model

Dissolution? Or persistence?

$$logit\left(P(Y_{ij,t+1}=1 | Y_{ij,t}=1, \text{ rest of the graph})\right) = \boldsymbol{\theta}^{-\prime}\boldsymbol{\partial}(\boldsymbol{g}^{-}(\boldsymbol{y}))$$

- The model is expressed as log odds of tie equaling 1 given it equaled 1 at the last time step
- This is done to make it consistent with the formation model, so all the math works out nicely
- But it implies that the model, and thus the coefficients, should be interpreted in terms of effects on relational persistence
- That said, people tend to thing in terms of relational formation and dissolution, since relational dissolution is a more salient event than relational persistence
- Thus, we often use the language of dissolution

During simulation, two processes occur separately within a time step:



- Y⁺ = network in the formation process after evolution
- Y = network in the dissolution process after evolution
- This is the origin of the "S" in STERGM

- The statistical theory in Krivitsky and Handcock 2014:
 - demonstrates a given combination of formation and dissolution model will converge to a stable equilibrium, i.e.:

Prevalence ≈ Incidence x Duration

■ This and other work in press provide the statistical theory for methods for estimating the two models, given certain kinds of data

Term = ~edges

	$oldsymbol{ heta} otag $	$oldsymbol{ heta} \mathrel{\searrow}$
Formation model	more new ties created each time step	fewer new ties created each time step
Dissolution (persistence) model	more existing ties pre- served (fewer dissolved); longer average duration	fewer existing ties pre- served (more dissolved); shorter average duration

What combo do you think is most common in empirical networks?

Term = ~edges

	$oldsymbol{ heta} otag $	$oldsymbol{ heta} \mathrel{\searrow}$
Formation model	more new ties created each time step	fewer new ties created each time step
Dissolution (persistence) model	more existing ties pre- served (fewer dissolved); longer average duration	fewer existing ties pre- served (more dissolved); shorter average duration

What combo do you think is most common in empirical networks?

Term = ~concurrent (# of nodes with degree 2+)

	$oldsymbol{ heta}$ /	$oldsymbol{ heta} \mathrel{\searrow}$
Formation model	more ties added to actors with exactly 1 tie	fewer ties added to actors with 1 tie
Dissolution (persistence) model	actors with 2 ties more likely to have them be preserved	actors with 2 ties more likely to have them dissolve

What combo do you think is most common in empirical sexual networks?

Term = ~concurrent (# of nodes with degree 2+)

	$oldsymbol{ heta}$ $ heta$	$oldsymbol{ heta} \mathrel{\searrow}$
Formation model	more ties added to actors with exactly 1 tie	fewer ties added to actors with 1 tie
Dissolution (persistence) model	actors with 2 ties more likely to have them be preserved	actors with 2 ties more likely to have them dissolve

What combo do you think is most common in empirical sexual networks?

STERGMs: Data sources

- 1. Multiple cross-sections of complete network data
 - easy to work with
 - but rare-to-non-existent in some fields
- 2. One snapshot of a cross-sectional network (census, egocentric, or otherwise), plus information on relational durations
 - more common
 - but introduces some statistical issues in estimating relation lengths

STERGMs: nodal dynamics

- All of the statistical theory presented so far regards networks with
 - Dynamic relationships, but still
 - Static actors
- I.e. no births and deaths, no changing of nodal attributes
- The statistical theory of STERGM can handle nodal dynamics during simulation, with a few added tweaks
 - Most important is an offset term to deal with changing population size
 - Without it, density is preserved as population size changes
 - With it, mean degree is preserved as population size changes

STERGMs: nodal dynamics

For more info, see:

Pavel N. Krivitsky, Mark S. Handcock, and Martina Morris (January 2011). Adjusting for Network Size and Composition Effects in Exponential-Family Random Graph Models. Statistical Methodology, 8(4): 319–339

- And for more help with using STERGMs to simulate dynamic networks along with changing nodes and attributes:
 - Take our intensive summer workshop on network modeling for epidemic diffusion
 - Explore the online materials for the workshop (on the statnet webpage)
 - Try the EpiModel package

To the tutorial.....

(reference slides follow)

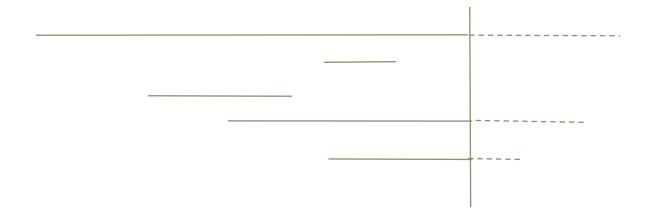
In some domains, often takes the form of

- asking respondents about individual relationships (either with or without identifiers).
- Often this is the n most recent, or all over some time period, or some combination (e.g. up to 3 in the last year)
- asking whether the relationship is currently ongoing
- if it's ongoing: asking how long it has been going on (or when it started)
- if it's over: asking how long it lasted (or when it started and when it ended)

From this we want to estimate

- the mean duration of relationships
- perhaps additional information about the variation in those durations (overall, across categories of respondents, etc.)

Issues?



- 1. Ongoing durations are right-censored
 - can use Kaplan-Meyer or other techniques to deal with

Issues?

- 2. Relationships are subject to length bias in their probability of being observed
 - This can also be adjusted for statistically
 - However, complex hybrid inclusion rules (e.g. most recent 3, as long as ongoing at some point in the last year) can make this complicated

- In practice (and for examples in this course), we sometimes rely on an elegant approximation:
 - If relation lengths are approximately exponential/geometric (a big if!),
 then the effects of length bias and right-censoring cancel out
 - The mean amount of time that the **ongoing** relationships have lasted until the day of interview (relationship age) is an unbiased estimator of the mean duration of relationships
 - Why?!?

- Exponential/geometric durations suggests a memoryless processes one in which the future does not depend on the past
- Imagine a fair, 6-sided die:
 - What is the probability I will get a 1 on my next toss?
 - What is the probability I will get a 1 on my next toss given that my previous 1 was five tosses ago?
 - On average, how many tosses will I need before I get my first 1?
 - On average, how many more tosses will I need before I get my next 1, given that my previous 1 was 8 tosses ago?

Geometric		
Parameters	$0 success probability (real)$	
Support	$k \in \{1, 2, 3, \ldots\}$	
Probability mass function (pmf)	$(1-p)^{k-1}p$	
Cumulative distribution function (CDF)	$1 - (1 - p)^k$	
Mean	$\frac{1}{p}$ 25	

SUNBELT 2015 - 23 JUNE 2015

- Now, let's imagine this fairly bizarre scenario:
 - You arrive in a room where there are 100 people who have each been flipping one die; they pause when you arrive.
 - You don't know how many sides those dice have, but you know they all have the same number.
 - You are not allowed to ask any information about what they've flipped in the past.
 - The only information people will give you is: how many flips after your arrival does it take until they get their first 1?
 - You are allowed to stay until all of the 100 people get their first 1, and they can inform you of the result.
- Given the information provided you, how will you estimate the number of sizes on the die?

- Simple: when everyone tells you how many flips it takes from your arrival until their first 1, just take the mean of those numbers. Call it *m*.
- Your best guess for the probability of getting a 1 per flip is 1/m.
- And your best guess for the number of sides is the reciprocal of the probability of any one outcome per flip, which is 1/1/m, which just equals m again.
- Voila!

Retrospective relationship surveys are like this, but in reverse:

Dice:

Relationships:

- If you have something approximating a memoryless process for relational duration, then an unbiased estimator for relationship length is to:
 - ask people about how long their ongoing relationships have lasted up until the present
 - take the mean of that number across respondents.

- In practice, we find that the geometric distribution doesn't often capture the distribution of relational durations overall.
- But, if you divide the relationships into 2+ types, it can do a reasonable job within type
- Especially if you remove any 1-time contacts and model them separately (for populations where they are common)
- Remember: DCMs model pretty much everything as a memoryless process, so approximating one aspect of our model that way is well within common practice