



407 Assignment 2

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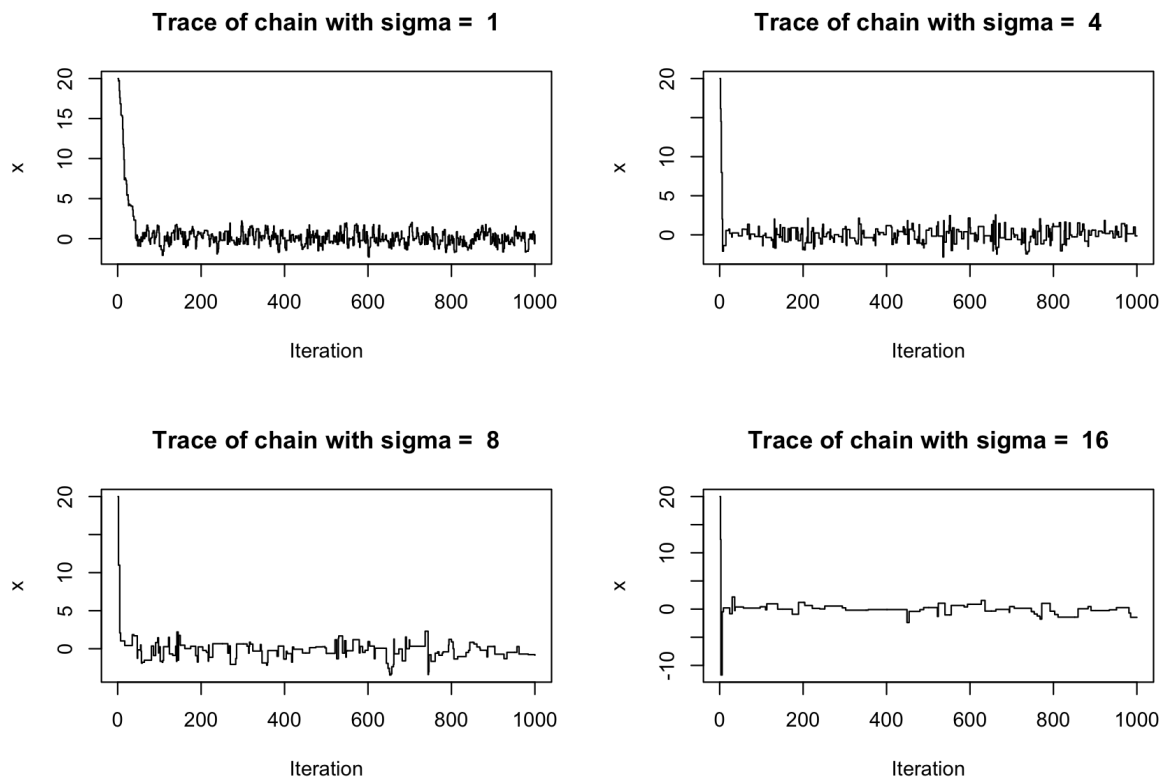
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question 1

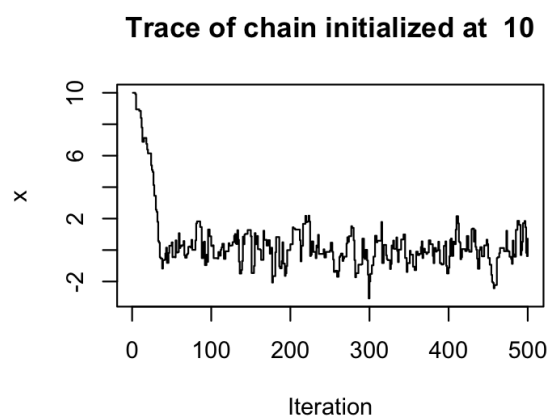
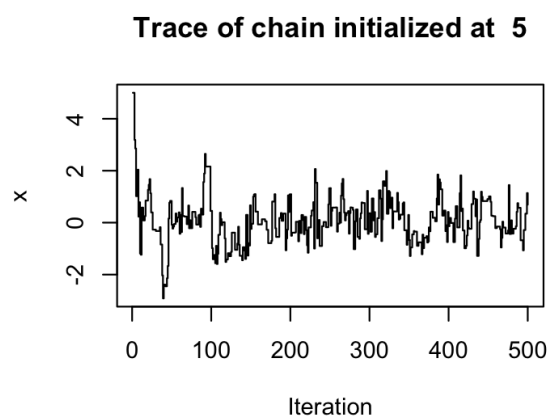
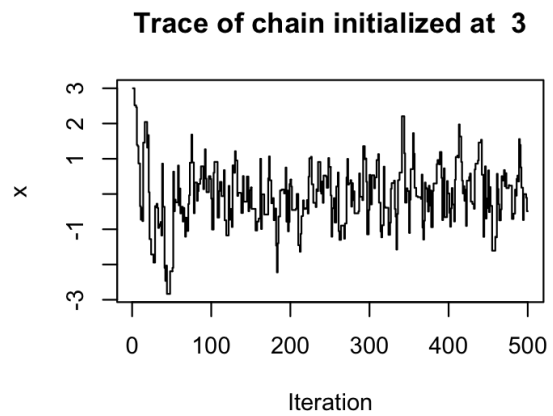
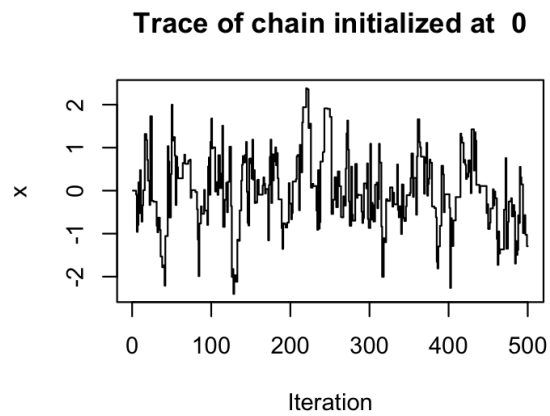
```
metrp = function(n=1000,Sigma = 1,miu = 0,x0=0,Beta = 1.5){  
  X = matrix(0,nrow = 1,ncol = n)  
  X[1] = x0  
  U = runif(n)  
  for (i in 2:n) {  
    Y = X[i-1]+rnorm(1,mean = miu,sd = Sigma)  
    # alpha = exp((abs((X[i-1]-miu)/Sigma))^Beta-(abs((Y-miu)/Sigma))^Beta)  
    alpha = exp(abs(X[i-1])^Beta-abs(Y)^Beta)  
    if(U[i]<alpha){  
      X[i] = Y  
    }  
    else{  
      X[i] = X[i-1]  
    }  
  }  
  X  
}
```

(a)

- i. the choice of Sigma have influence on the acceptance rate, I would like to choose Sigma=1

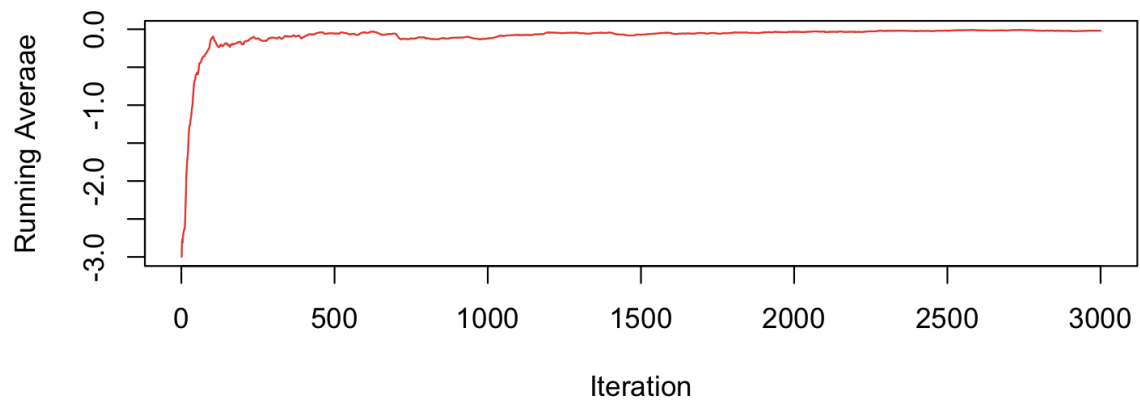


- ii. the initial X_0 would have influence on the *Burn-in* process, through these plots we can see that the sequence will finally fluctuate around 0, so we can decide by the stabilization of the chain

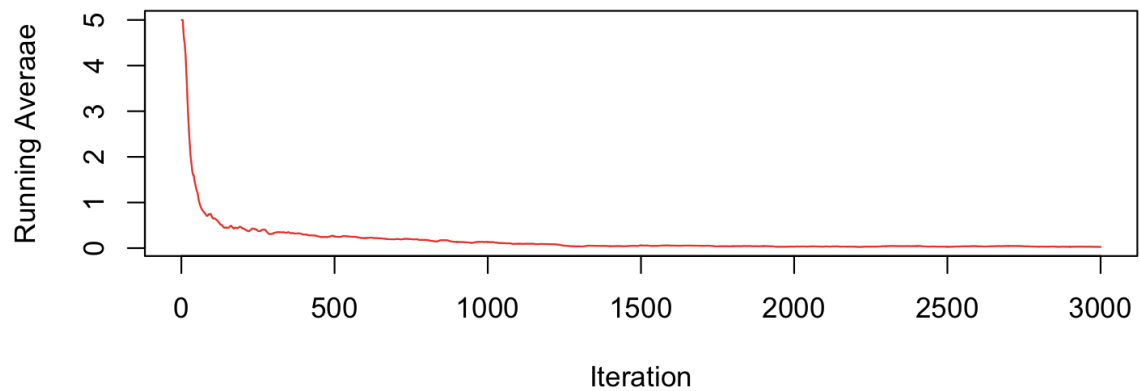


iii. I plot the empirical means of a chain through iteration, it become so steady around 0.

Trace of empirical means starting at -3



Trace of empirical means starting at 5



(b)

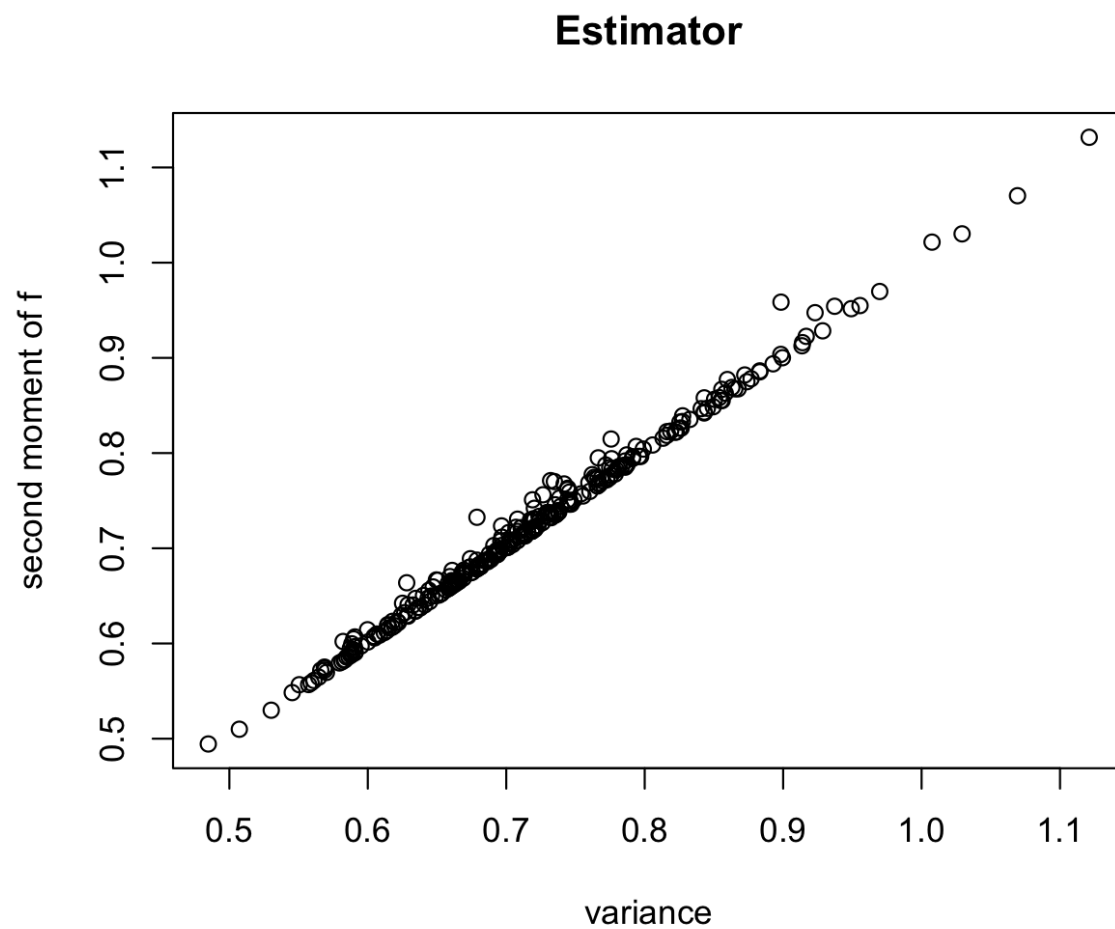
```
est = c()
for (i in 1:300) {
  xb = metrp(x0=0,Sigma=1 ,n = 1000)
  est[i] = mean(xb^2)
}
var(est)
```

$$\text{var}(\text{Estimator}) = 0.0114115$$

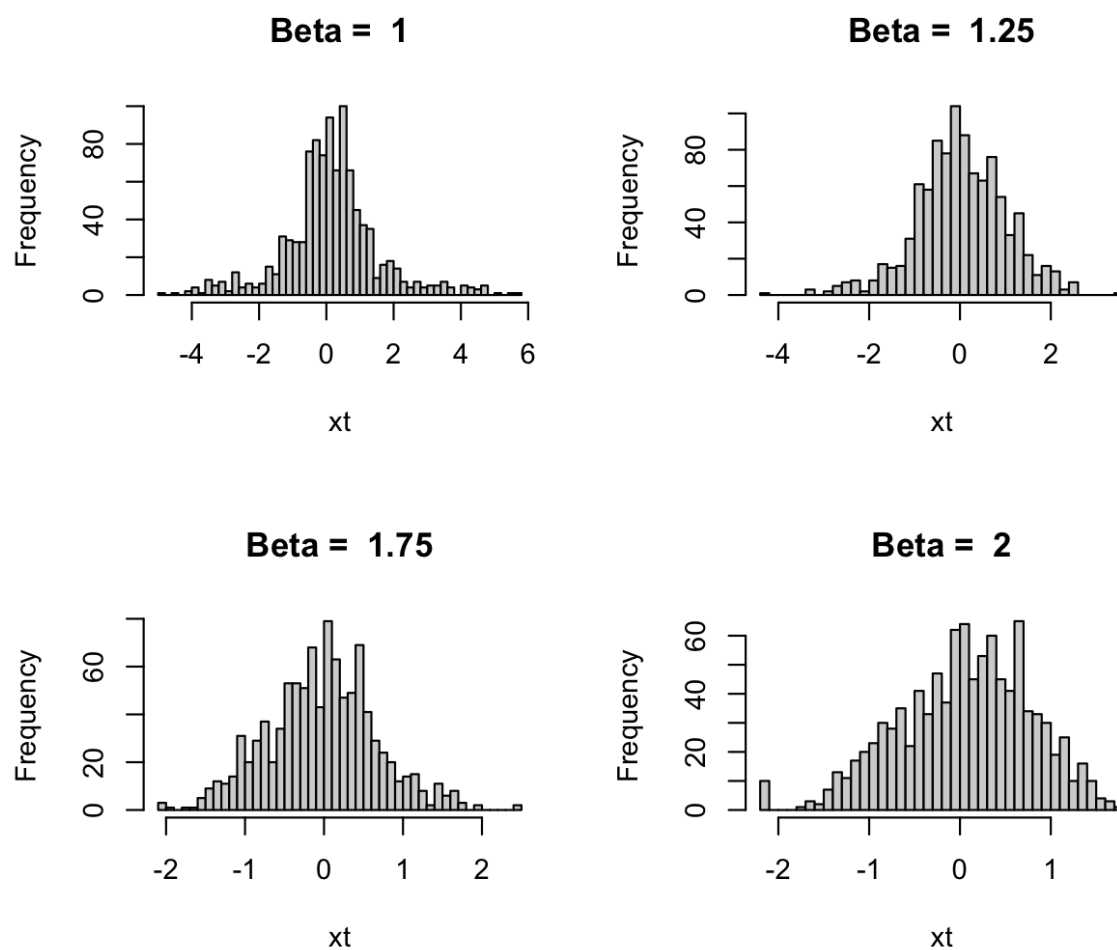
(c)

as the mean of distribution is 0, so the variance of f is equal to
second moment of f

so we can estimate like part b. to calculate the variance.



invariant distribution of Markov chain for different β



Possibly, as we have known the combination distribution of X and β , we could calculate the conditional distribution and then apply *Gibbs sampling*

question 2

```
pdf = function(x){
  if(x > -5 && x < (-3)){
    return(3/8*(1-(x+4)^2))
  }
  else if(x > 3 && x < 5){
    return(3/8*(1-(x-4)^2))
  }
  else return(0)
}
```

(a)

(a) through the sample steps, it's obvious that

$$\textcircled{1} P(u_t | x_{t-1}) = U(0, f(x_{t-1}))$$

$$\textcircled{2} P(x_t | u_t) = U(x: f(x) > u_t)$$

$$P(x_t, u_t) = P(x_{t-1}, u_{t-1}) \cdot P(u_t | x_{t-1}) \cdot P(x_t | u_t)$$

it corresponds to Gibbs Sampling Steps
Starting at x_0 and through the 2 steps, the
sample area is limited into $\{(x, u) : u < f(x)\}$,

(b)

$$\begin{aligned} \text{(b). } p(x) &= \int_u p(x, u) du \\ &= \int_0^{f(x)} \frac{1}{Z} du \\ &= \frac{f(x)}{Z}, \quad Z = \int f(x) dx = 1 \end{aligned}$$

\therefore the invariant distribution is $f(x)$

(c)

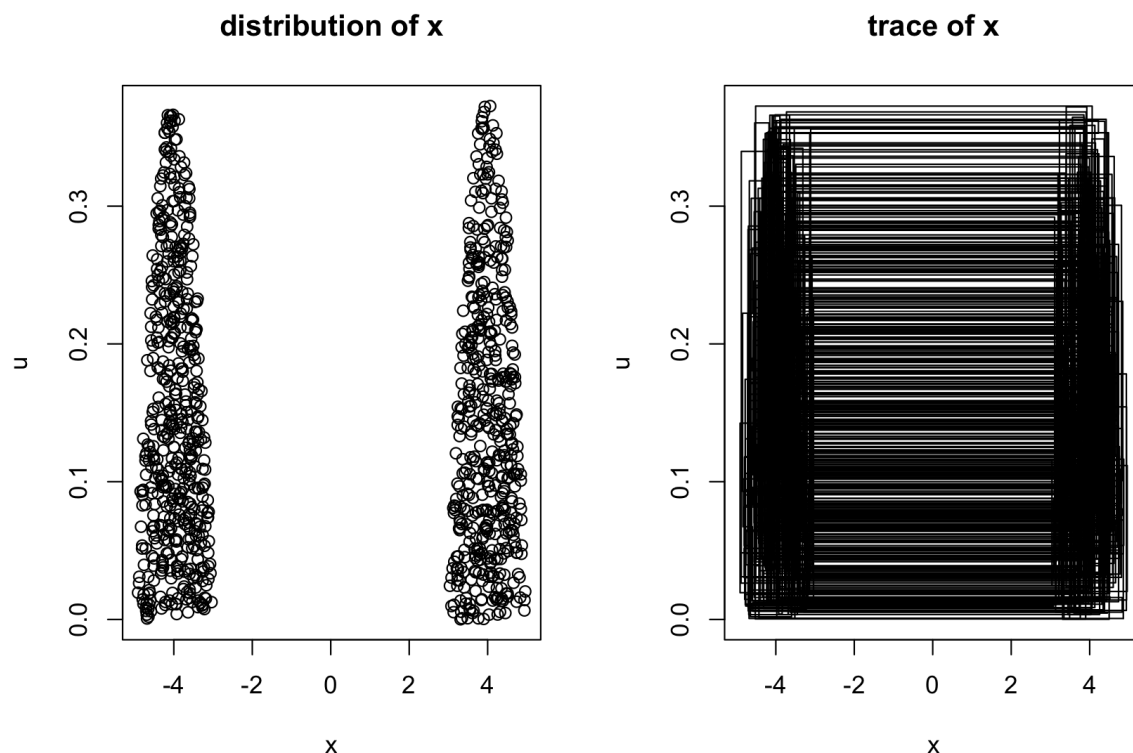
1. Draw $U^{(t)} \sim U(0, f(x_{(t-1)}))$

2. Draw $X^{(t)} \sim$

$$U \left\{ x: x \in \left[-4 - \sqrt{1 - \frac{8}{3}u_t}, -4 + \sqrt{1 - \frac{8}{3}u_t} \right] \cup \left[4 - \sqrt{1 - \frac{8}{3}u_t}, 4 + \sqrt{1 - \frac{8}{3}u_t} \right] \right\}$$

slice sampler:

```
slice = function(n=1000, x0=as.vector(c(0,0))){
  X = matrix(0, nrow = 2, ncol = n)
  X[,1] = x0
  for (i in 2:n) {
    ###U(t)
    X[1,i] = runif(1, 0, pdf(X[2,i-1]))
    ###X(t)
    X[2,i] = runif(1, -sqrt(1-X[1,i]*8/3), sqrt(1-X[1,i]*8/3)) + (runif(1, -1, 1) > 0) * 8 - 4
  }
  X
}
par(mfrow = c(1,2))
xss = slice(x0 = as.vector(c(0.1,4)))
plot(xss[2,], xss[1,], xlab = 'x', ylab = 'u', main = 'distribution of x')
plot(xss[2,], xss[1,], type = 's', xlab = 'x', ylab = 'u', main = 'trace of x')
```

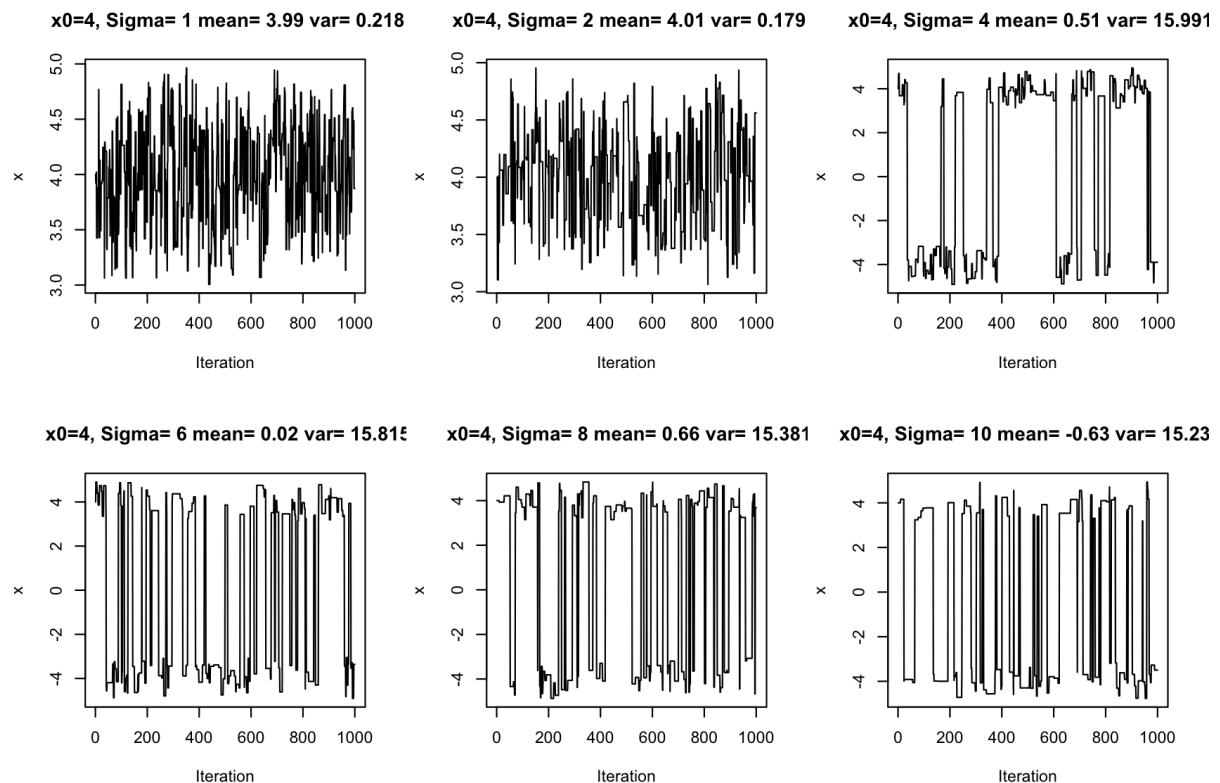
Metropolis algorithm

```
#####metropolis
mh = function(n=1000,Sigma = 1,miu = 0,x0=0,Beta = 1.5){
  X = matrix(0,nrow = 1,ncol = n)
  X[1] = x0
  U = runif(n)
  for (i in 2:n) {
    Y = X[i-1]+rnorm(1,mean = miu,sd = Sigma)
    alpha = pdf(Y)/pdf(X[i-1])
    if(U[i]<alpha){
      X[i] = Y
    }
    else{
      X[i] = X[i-1]
    }
  }
  X
}

par(mfrow = c(2,3))
xm = matrix(0,nrow = 8,ncol = 1000)
count = 1
for(i in c(1,2,4,6,8,10)){
  xm[count,] = mh(Sigma = i,x0 = 4)
  m = round(mean(xm[count,]),2)
  v = round(var(xm[count,]),3)
  name = paste(' x0=4, Sigma=',i,'mean=',m,'var=',v)
  plot(seq(1,1000,1),xm[count,],type = 'l',xlab = 'Iteration',ylab = 'x',main = name )
}
```

```
count = count + 1
}
```

we choose different Sigma and calculate the mean and var for every Sigma
plot the trace as following :



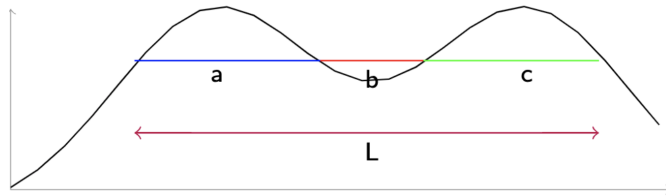
obviously there is an huge change of Mean and Var through the increase of Sigma, especially from Sigma = 2 to 4

Through the plots of two algorithms I think the quality of slice sampling is better as Metropolis algorithm could be influence by choice of Sigma, however using Slice algorithm request to choose a proper initial X_0 that $f(X_0) > 0$

(d)

i) Slice sampling is hard to implement for high dimensions because we when we choose X_t , we need to make sure that $X_t \sim U\{x_t : f(x_t) \geq U_t\}$, but in real problems the area of this is often divided into many different parts depending on U_t , and when it comes to high dimension it tends to be multi_dimension plates, which makes it harder to draw X_t uniformly.

ii) we can use Shrinkage Algorithm in order to deal with the gaps between the area of X_t ,



$$\Pr(X^{(t)} \in a \rightarrow X^{(t+1)} \in a) = \frac{a+b}{L}$$

$$\Pr(X^{(t)} \in a \rightarrow X^{(t+1)} \in c) = \frac{c}{L}$$

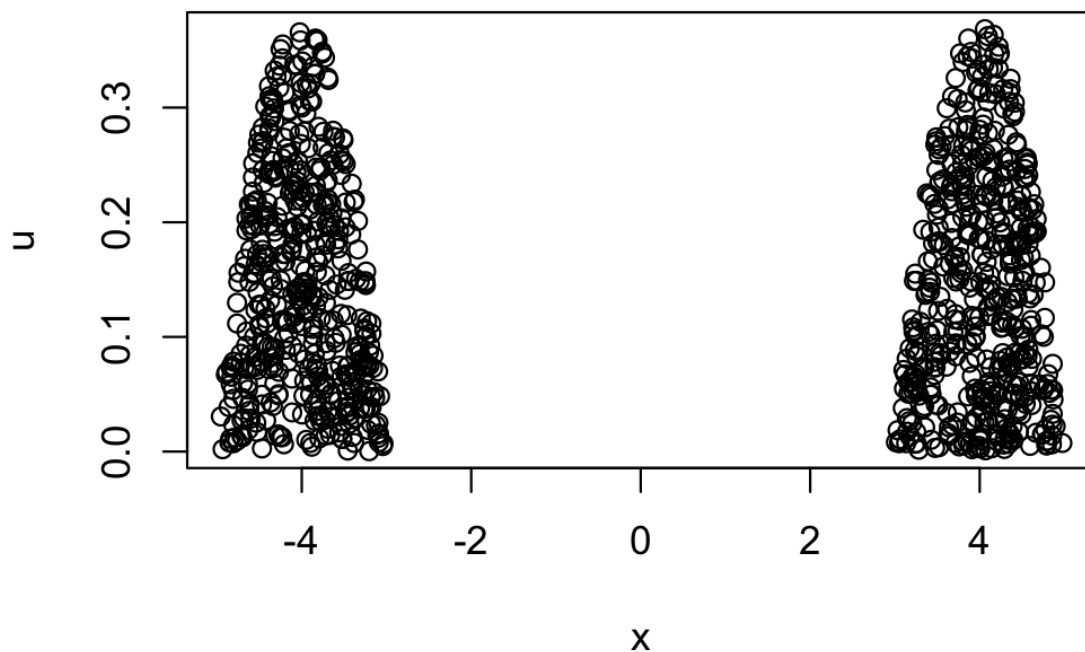
$$\Pr(X^{(t)} \in c \rightarrow X^{(t+1)} \in a) = \frac{a}{L}$$

$$\Pr(X^{(t)} \in c \rightarrow X^{(t+1)} \in c) = \frac{b+c}{L}$$

Therefore,

$$\begin{aligned} \Pr(X^{(t+1)} \in a) &= \Pr(X^{(t)} \in a \rightarrow X^{(t+1)} \in a) \Pr(X^{(t)} \in a) + \Pr(X^{(t)} \in c \rightarrow X^{(t+1)} \in a) \Pr(X^{(t)} \in c) \\ &= \frac{a+b}{L} \frac{a}{L} + \frac{a}{L} \frac{c}{L} = \frac{a^2 + ab + ac}{L^2} = \frac{a}{L} = \Pr(X^{(t)} \in a) \end{aligned}$$

distribution of x



iii)

I think the Shrinkage algorithm add a estimate to the Slice sampler which is like the way that Metropolis Hasting use: to change the probability of transforming depending on the different area of X_t in order to make sure the

Detailed Balance Function