ECOSPHERE



Analyzing animal movement patterns using potential functions

Haiganoush K. Preisler, ¹, † Alan A. Ager, ² and Michael J. Wisdom ²

Citation: Preisler, H. K., A. A. Ager, and M. J. Wisdom. 2013. Analyzing animal movement patterns using potential functions. Ecosphere 4(3):32. http://dx.doi.org/10.1890/ES12-00286.1

Abstract. The advent of GPS technology has made it possible to study human-wildlife interactions on large landscapes and quantify behavioral responses to recreation and other anthropogenic disturbances at increasingly fine scales. Of particular interest are the potential impacts on habitat use patterns, energetics, and cascading impacts on fecundity and other life history traits for key wildlife species that are exposed to human activities. Statistical models quantifying effects of human activity on animal movement on a heterogeneous landscape are essential for understanding these potential impacts. Here we present a statistical framework for analyzing movement data that is based on the concept of a potential surface. The potential surface is motivated by the assumption that animals are moving on a space-time surface with regions or points of attraction or of repulsion. We demonstrate the use of the technique by analyzing movement data from a long-term controlled experiment to evaluate the responses of free ranging Rocky Mountain elk (*Cervus elaphus*) to anthropogenic disturbances that vary in time and space. Our results demonstrated a strong avoidance of elk to all-terrain vehicles detected up to one km from the disturbance. Elk avoidance of mountain bikers was detected up to 500 m, and avoidance of hikers and horseback riders was detected to 200 m.

Key words: Cervus elaphus; gradient system; movement; potential function; R code; Rocky Mountain elk; Starkey Experimental Forest and Range; stochastic differential equation.

Received 11 September 2012; revised 10 January 2013; accepted 11 January 2013; **published** 04 March 2013. Corresponding Editor: R. Parmenter.

Copyright: © 2013 Preisler et al. This is an open-access article distributed under the terms of the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited. http://creativecommons.org/licenses/by/3.0/

† E-mail: hpreisler@fs.fed.us

INTRODUCTION

Fine scale data-logging technology, such as global positioning systems (GPS), has stimulated many new approaches for analyzing and modeling movement patterns of free-ranging animals. Models of animal movements provide a quantitative framework to analyze spatiotemporal effects of anthropogenic disturbances, predators, and conspecifics. Of particular interest are the connections between movements and habitat use, energetics, and life history traits for keystone wildlife species that are increasingly exposed to human activities (Sawyer and Kaufman 2011).

Dynamics of free-ranging animal movements are complex, even without consideration of anthropogenic impacts. A wide range of models have been formulated and applied to animal movement. Most methods involve a Markov process such as uncorrelated or simply correlated random walk where the location of an animal at each step depends on the location in the previous step (Turchin 1998, Okubo and Levin 2001, Morales et al. 2004, Codling and Hill 2005, Smouse et al. 2010). Other processes include the Brownian Bridges models where the continuous movement paths are estimated assuming conditional random walk between successive locations

1

¹Pacific Southwest Research Station, UDSA Forest Service, Albany, California 94710 USA ²Pacific Northwest Research Station, USDA Forest Service, La Grande, Oregon 97850 USA

(Horne et al. 2007, Sawyer et al. 2009). Statespace or hidden Markov models assume locations of animals are a function of the unobserved state of the system (location or velocity) at a previous general time (Johnson et al. 2008, Dowd and Joy 2011). Other models have been explored for modeling movement using mechanistic models that incorporate effects of explanatory variables, such as human disturbances and landscape features (Forester et al. 2007, McClintock et al. 2012). The approaches listed above model the stochastic differential equations of motion (SDE), or the velocity of the motion, as a function of explanatory variables. Our present work is based on the idea of potential functions that are used to motivate specific functional forms for the SDEs. This framework may be used in conjunction with other methods, such as state-space models, by presenting a mechanism for arriving at an equation for the expected step sizes of movement. Potential functions are motivated by methods from physics where the motion of a physical particle is modeled as being affected by its location on a potential surface and by its relative distances from other particles, regions, or barriers (Hirsch et al. 2004, Taylor 2004). An attractive feature of the potential function approach is the ease with which multiple factors affecting motion (e.g., foraging behavior, topography, human disturbance) can be incorporated into a mechanistic model and then parameterized via a regression routine. The estimation methods assume a Markovian/diffusion process which is then generalized to more realistic stochastic processes by introducing, for example, serial correlations in the error term. Our goals in the current paper are to (1) provide ecologists an overview of the potential function modeling approach, (2) demonstrate its practical implementation with readily-available regression routines (e.g., open access R programing software), (3) extend the techniques presented in our previous work (Brillinger et al. 2004, Brillinger 2010, Brillinger et al. 2011) to accommodate serial correlations, and (4) use the technique to study movements from a long-term controlled experiment for evaluating responses of free ranging Rocky Mountain elk (Cervus elaphus) to multiple anthropogenic disturbances, including all-terrain vehicle (ATV) riding, hiking, mountain biking, and horseback riding (Wisdom et al. 2004,

Naylor et al. 2009).

THE POTENTIAL FUNCTION APPROACH

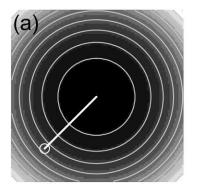
Our approach is motivated by the assumption that particles are moving on a space-time surface containing regions or points of attraction and repulsion. This potential field is similar to a topographic surface under the influence of gravity where objects are attracted to low points or hollows and repelled from high points or hills. A simple example with a point of attraction at (0, 0) is provided by the potential $H(x, y) = (x^2 + y^2)$ where a particle at (x, y) moves directly to (0, 0)on the surface (Fig. 1a). Next, if some random noise or perturbation is added to the potential, for example by shaking the surface, then the particle will still drift toward the point with lowest potential but with the path not straight but wiggling with an amplitude dependent on the level of the noise (see Fig. 1b, c). A second elementary example is provided by a process where a particle moves on a flat potential surface by means of statistically independent random steps, i.e., a random walk process with no drift.

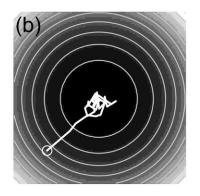
An analytic foundation for this modeling framework is provided by a formal relationship between the potential surface and the velocity of the particle at a given location (x, y) and time t. In Newtonian physics, the velocity of a particle at location (x, y) is given by the negative of the gradient of the potential function at (x, y). The gradient of a surface is defined by the slopes in the x- and y-directions, that is grad $H = (\partial H/\partial x, \partial H/\partial y)$. For example, for the potential surface in Fig. 1 with a point of attraction at (0, 0) the potential function $H(x, y) = (x^2 + y^2)$, and the velocity of the particle, as given by the speeds in the x- and y-directions are:

$$\frac{\partial x}{\partial t} = -\frac{\partial H(x, y)}{\partial x} = -2x;$$

$$\frac{\partial y}{\partial t} = -\frac{\partial H(x, y)}{\partial y} = -2y.$$
(1)

The minus signs in Eq. 1 are traditional. Eq. 1 provides one example of a gradient system (see Hirsch and Smale 1974:199–204). Other examples of parametric potential functions and the corresponding surfaces are provided in Fig. 2. The potential can be any differentiable function, including non-parametric smooth functions that





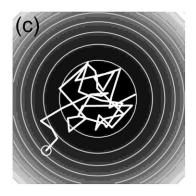


Fig. 1. Tracks of a particle moving in a potential with a point of attraction (lowest point) at the center of the surface. Tracks in the three panels were generated by adding various amounts of random noise to the same potential. (a) No noise; (b) some random noise (small variance); (c) larger amount of random noise (larger variance). In each panel the starting point is indicated by the open circle.

are estimated from the data.

Eq. 1 is important because while the potential itself may not be directly observable, the step sizes in the x- and y-directions, and thereby the approximate velocity of the animal, is observable. Consequently, if we assume the existence of a potential affecting the movements of an animal (up to random fluctuations) then it may be estimated, given observations on consecutive locations. This approach assumes that the potential function is differentiable. Complicated potential surfaces may be modeled and assessed

using this framework. For example, there may be multiple regions of attraction where animals prefer to forage, versus regions of repulsion where human disturbances are frequent. A potential surface that integrates multiple factors affecting movement is developed by adding separate surfaces describing each of the attractions and repulsions together (Fig. 3 and Eq. 6). Potential functions may also be used to model animal movement as they are attracted to, or repulsed from, a moving entity (e.g., predator, hunter, vehicle). Brillinger et al. (2011) considered

	Parametric Potential Functions			
Example	Orenstein-Uhlenbeck	Gaussian	Gravitational	Zohdi
Potential function	$H = \beta_o + \beta_1 (x - x_o)^2 + \beta_2 (y - y_o)^2$	$H = \exp\left\{\alpha \left(x - x_o\right)^2 + \beta \left(y - y_o\right)^2\right\}$	$H = \beta / D$	$H = \alpha D^{-\beta} - y D^{-\delta}$
Potential surface	x	x y	X V	x

Fig. 2. Four examples of parametric potential functions, with one point of attraction or repulsion at (x_0, y_0) , and plots of the corresponding surfaces for specific values of the parameters (α, β, δ) . $D = \sqrt{(x - x_0)^2 + (y - y_0)^2}$. An example of the use of the Zohdi function is found in Zohdi (2003, 2009) and Brillinger et al. (2011).

21508925, 2013, 3, Downloaded from https://esajournals.onlinelibrary.wiely.com/doi/10.1890ES12-00286.1 by Yale University, Wiley Online Library on [07/10/2024]. See the Terms and Conditions (https://onlinelibrary.wiely.com/erms-and-conditions) on Wiley Online Library for rules of use; OA articles are governed by the applicable Creative Commons Licenses

Fig. 3. An example of a potential surface with two repulsion regions and simulated movement tracks within this surface. Red squares are the starting points of the tracks. The highest background values (pink then orange) indicate ridges (high potential), followed by yellow then green and finally blue then purple indicating valleys (low potential regions).

the case of modeling attraction between a given animal and conspecifics. When the point of attraction or repulsion is moving (e.g., person or conspecific) then the potential surface that defines the regions of attraction/repulsion are changing as well.

STOCHASTIC DIFFERENTIAL EQUATIONS

The potential function allows one to model the deterministic component (expected velocity) of movement in terms of biologically meaningful parameters. The stochastic component of the movement is modeled by adding a random term to the differential equations defining the velocity of the particle at time *t*. Specifically, the stochastic differential equation (SDE) is:

$$dr(t) = \mu(r, t)dt + \Sigma(r, t)d\mathbf{V}(t)$$
 (2)

where $d\mathbf{r}(t) = {\partial x, \partial y}$; $\mathbf{\mu}(t)$ is minus the gradient of the potential surface of the expected drift or movement direction of the particle in the next increment of time; $\Sigma(r, t)$ is a real valued matrix specifying the correlation between the steps in the x and y-directions; and dV(t) is a continuoustime stochastic process with expected value of zero. The SDE formulation with Brownian driver, i.e., dV(t) = dB(t) leads to a continuous time random walk under assumptions. Eq. 2 is the basis for the discrete time formulation used in this paper for estimation purposes. The Brownian driver can take on many forms (Table 1), including auto-correlated time series.

STATIONARY DISTRIBUTIONS AND POTENTIAL **FUNCTIONS**

Animal movement studies are sometimes concerned with the estimation of the home range, i.e., spatial extent of an animal's movement, or the stationary distribution of animal locations, sometimes referred to as the utilization distribution (Millspaugh et al. 2006). In special cases, in particular when the potential function *H* does not depend on time t and $\Sigma = \sigma^2 I$, the stationary distributions may be shown to be

Table 1. Examples of stochastic processes.

Example	$\mu(x(t), y(t))$	$d\mathbf{V}(t) = \{dV_x, dV_y\}$
Random walk Biased random walk O-U process with point of attraction	0 Constant (not zero) vector A $A r(t) - a $	V_x , V_y independent Brownian processes V_x , V_y independent Brownian processes V_x , V_y independent Brownian processes
at a Correlated random walk (consecutive moves are correlated)†	0	V_{x} , V_{y} positively correlated processes
Levy process (Brownian process with jumps)‡	Any parametric or non-parametric smooth function	V_{x} , V_{y} Brownian plus marked Poisson process

[†] Note that in both O-U process and correlated random walk (CRW) animals are moving with persistence in a given direction. That is, consecutive moves are correlated. However, in an O-U process the persistence is towards a point of attraction/repulsion.

‡ Jumps occur at random time points as characterized by a Poisson process.

$$\pi(r) = k\exp\{-2H(r)/\sigma^2\}. \tag{3}$$

The constant k is to make $\pi(r)$ integrate to 1. The potential function in this case is the logarithm of the stationary distribution (up to a constant). For example, if the potential function leads to the Ornstein-Uhlenbeck equations then the stationary distribution is the two-dimensional Gaussian density function with the mode of the distribution located at the point of attraction of the Ornstein-Uhlenbeck process.

ESTIMATION METHODS

There is substantial literature devoted to the topic of inference from stochastic differential equations (Sorensen 1997, Prakasa Rao 1999). Given a set of discrete observations on animal locations, the potential function and its gradients in continuous time (velocity) are used as a framework to build the discrete difference equation used in a regression model. Estimation is set in terms of the discrete approximation of the velocity in Eq. 2 as given by

$$\frac{[r(t_{i+1}) - r(t_i)]}{[t_{i+1} - t_i]} \approx \mu \Big(r(t_i) \Big) + \Sigma \mathbf{Z}_{i+1} / \sqrt{t_{i+1} - t_i} \quad (4)$$

with \mathbf{Z}_i a standard bivariate random vector with variance-covariance matrix given by Σ . The term $\sqrt{t_{i+1}-t_i}$ is needed because the variance-covariance matrix of the real valued continuous process, $d\mathbf{V}_i$, is Σdt .

Given observed locations (x_i, y_i) for what may be unequally spaced times $t_1 < t_2 < ... < t_i < ... < t_m$, parameters, Θ , of a differentiable function, $H(x, y|\Theta)$, may be estimated by the regression

model

$$\frac{\Delta x_i}{\Delta t_i} = H_x(x_i, y_i | \Theta) + \varepsilon_i$$

$$\frac{\Delta y_i}{\Delta t_i} = H_y(x_i, y_i | \Theta) + \gamma_i$$
(5)

where; $\Delta x_i = (x_{i+1} - x_i)$; $\Delta y_i = (y_{i+1} - y_i)$; $\Delta t_i = (t_{i+1} - y_i)$ $-t_i$); H_x , H_y are the partial derivatives of $H(x, t_i)$ $y|\Theta$) with respect to x and y and where ε and γ are random noises. In the case where the error terms ϵ_i and γ_i are independent, ordinary least squares regression routines may be used to estimate the parameters. The error terms may not be independent if the step sizes, Δx_i , Δy_i , are correlated, $cov(\varepsilon_i, \gamma_i) \neq 0$, or if the sample includes observations on more than one subject/ animal resulting in within subject correlated errors. Observation error in our study was assumed to be negligible relative to the stochastic error, although this assumption needs to be carefully examined on a study by study basis (see The Data section). Finally, serial correlation may necessitate the use of an autoregressive model.

The gradient of the potential surface, when displayed as a vector field of arrows, provides an estimate of the average velocity of an animal at a point. Sometimes the average velocity of animals is of interest in itself. One may use a nonparametric smooth function of location to estimate the discrete velocities without relating them to a specific parametric potential surface. Estimation in this case can be carried out using the two-dimensional tensor spline functions (Wood 2006) or a two-dimensional locally weighted regression

routine (Hastie and Tibshirani 1990).

AN EXAMPLE ON ELK MOVEMENT

The experiment and the data

Here we demonstrate potential functions that describe the strength of repulsion between elk and four different human activities (riding an All-terrain vehicle (ATV), biking, hiking, and equestrian riding), using data from a controlled landscape experiment at the Starkey Experimental Forest and Range in eastern Oregon (Wisdom et al. 2004, Preisler et al. 2006, Naylor et al. 2009). In this experiment, 25 elk were followed (using Loran-C technology and GPS collars) during summer for four years (2002-2005). The purpose of the study was to understand responses of elk to four distinct recreational disturbances (henceforth treatments). During 2002-2004, each treatment was implemented for five days followed by a nine-day control period where all human activities were excluded. For each 5-d treatment period, two recreationists, moving independently, implemented one type of disturbance (ATV riding, mountain biking, hiking, or horseback riding). Each type of disturbance was implemented on 32 km of trails twice daily, once in the morning (approximately 08:00 to 12:00 local time) and once in the afternoon (approximately 1300 to 1700 local time) (Wisdom et al. 2004, Naylor et al. 2009). Each pair of recreationists was equipped with a GPS unit that tracked their locations on a continual basis (Wisdom et al. 2004). Data used in our analysis were limited to cases with elapsed time between consecutive observations less than 15 min. The median time between observations was 5 min with 85% less than 5.5 min. The mean error associated with the telemetry locations was <20 m for the GPS data (92% of the data used) and <50 m for the Loran-C data. Consequently, the assumption of negligible measurement error in the model was not unreasonable. For each elk observation, the distance to the disturbance associated with the step size between time t and $t + \Delta$ was calculated by using the nearest GPS location of the disturbance within 5 min of the observed elk location at time t. In 2005, treatments were excluded from the study area and elk location data were obtained to analyze movement

patterns in the absence of human activities.

Estimation

We assumed that elk are moving on a hypothetical potential surface as they react to two outside forces: (1) attraction towards unknown foraging and resting areas, and (2) reaction to a human disturbance. The potential function for this model is given by

$$H[x(t), y(t)] = H_1[x(t), y(t)] + H_2[d(x(t), y(t))]$$
 (6)

where $H_1[x(t), y(t)]$ is assumed to be a nonparametric smooth function of the elk location $\{x(t), y(t)\}$ describing the regions of attraction for elk during different periods of the day (movement towards foraging or resting grounds). $H_2[x(t), y(t)]$ is assumed to be a parametric function of the distance, d(x(t), y(t)), between the animal and the disturbance at time t. Specifically, we used a fourth-degree polynomial

$$H_2\Big(d(x(t), y(t)|\Theta\Big)$$

$$= \beta_o + \beta_1 d + \beta_2 d^2 + \beta_3 d^3 + \beta_4 d^4$$
for $d < \infty$

$$H_2(\infty|\Theta) = \alpha \qquad \text{for } d = \infty$$
(7)

where

$$d = \sqrt{(x - z_x)^2 + (y - z_y)^2}$$

is the distance between the location (x, y) of the elk and the location (z_x, z_y) of the disturbance at a given time t. The distance between an elk and a human disturbance is set to ∞ when there is no human activity in the region. The degree of the polynomial was arbitrary, and the parameters, β_0 , \dots , β_4 , individually have no physical interpretation. We simply required an equation for the potential surface that could accommodate most non-linear shapes. The potential surface as a whole, rather than the individual parameters, is what provides the interpretation. Note that, given a distance between an elk and a disturbance, the potential function term $H_2[x(t), y(t)]$ in Eq. 7 is assumed to be independent of time of day. On the other hand we assumed that $H_1[x(t)]$, y(t) is independent of time only within each of four day-time periods discussed below. These time periods corresponded to two distinct daily periods of maximum time spent foraging (04:00-

07:00 h and 17:00–20:00 h) and two distinct periods of maximum time spent resting (08:00–16:00 h and 21:00–03:00 h) (Ager et al. 2003). These distinct time periods were based on results of activity monitoring by Naylor et al. (2009) for elk in the northeast pasture. Treatments were implemented only during one of the four periods (08:00–16:00 h).

The regression equations used for estimation purposes were motivated by Eqs. 6 and 7. They are the discrete approximations (difference equations) of their partial derivatives with respect to x and y. For $d < \infty$, i.e., for days with human disturbance, the difference equations are given by

$$\frac{\Delta x_i}{\Delta t_i} = -H_{1x}(x, y) - (\beta_1 + 2\beta_2 d_i + 3\beta_3 d_i^2 + 4\beta_4 d_i^3)$$

$$\times (x_i - z_{x,i})/d_i + \varepsilon_i$$

$$\frac{\Delta y_i}{\Delta t_i} = -H_{1y}(x, y) - (\beta_1 + 2\beta_2 d_i + 3\beta_3 d_i^2 + 4\beta_4 d_i^3)$$

$$\times (y_i - z_{v,i})/d_i + \gamma_i$$
 (8)

and for control days $(d = \infty)$ are given by

$$\frac{\Delta x_i}{\Delta t_i} = -H_{1x}(x, y) + \varepsilon_i$$

$$\frac{\Delta y_i}{\Delta t_i} = -H_{1y}(x, y) + \gamma_i$$
(9)

 H_{1x} , H_{1y} are nonparametric smooth functions such as a locally weighted regression function or a two dimensional spline. The estimation process involved, first, using the control data collected in 2005 and the difference equations in Eq. 9 to obtain separate estimates of the diurnal movement patterns of elk for each of the four time periods described above. Next, Eq. 8 is used, with H_{1x} , H_{1y} replaced by their estimated values, $\hat{H}_{1x}, \hat{H}_{1y}$, from the first stage, to estimate the parameters $\{\beta_1, \ldots, \beta_4\}$. The R-code for estimating the regression parameters $\{\beta_1, \ldots, \beta_4\}$, is given in the Supplement. Parameter estimates were then used in Eq. 7 to develop an estimate of the potential surface up to the constant parameter, β_0 . Standard errors for the potential surface were evaluated from the estimated SEs of the individual parameters, $\{\beta_1, \ldots, \beta_4\}$, and those for \hat{H}_{1x} , \hat{H}_{1y} , using the delta-method.

RESULTS

Stochastic terms

Correlations between the step sizes in the xand y-directions were negligible for both control days and treatment days with $\rho(\varepsilon_i, \gamma_i)$ equal to 0.035 for control and -0.003 for treatment. Therefore, the matrix Σ in Eq. 2 was set to the diagonal matrix σ^2 I. The serial correlations $\rho(\varepsilon_{i\nu})$ ε_{i-1}), $\rho(\gamma_i, \gamma_{i-1})$ ranged between 0.21 and 0.36 for the four time periods during control days and between 0.32 and 0.52 for the four disturbances during treatment days. The larger values tended to be for serial correlations in the y-direction. The estimated serial correlations are an indication that the underlying continuous process is not Markovian because there is some evidence that an elk's location at time t depends not only on its location at t - 1 (Markov process) but also on its location two periods ago.

Diurnal movement patterns during control days

Plots of the estimated gradient of the potential function (movement arrows) evaluated using the control data (Fig. 4) demonstrate significant directional movement patterns. Arrows in Fig. 4 were plotted as a random sample of all locations where elk were detected during control days. Only arrows that were significantly greater than zero (arrow length >2 SE) were plotted. The estimated movement arrows revealed daily cycles of spatial movement patterns from resting to foraging areas and back. For example, during the early morning hours (04:00-07:00), elk appeared to be moving to preferred foraging areas, as depicted by movements to the northeast and away from the southern and eastern areas. Significant movement patterns were not evident during the night (21:00-03:00) nor daytime (08:00-16:00) periods when elk typically rest and ruminate. At dusk (16:00-20:00), elk again appeared to move from foraging areas to rest areas along the border fence in the west and south.

Movements with respect to disturbance

As noted above, treatments were introduced into the experimental region only during the day time hours 08:00–16:00. Undisturbed elk exhibited seemingly random movement and regions of attraction or repulsion were not apparent, con-

onlinelibrary.wiley.com/doi/10.1890/ES12-00286.1 by Yale University, Wiley Online Library on [07/10/2024]. See the Terms

: (https://onlinelibrary.wiley.com/terms-and-conditions) on Wiley Online Library for rules of use; OA articles are governed by the applicable Creative Commons

Fig. 4. Estimated (smoothed) movement vectors for four periods of the day in the absence of human disturbances. Speed of movement is proportional to the length of the movement arrows. The two time intervals on the left are associated with periods of maximum crepuscular movements to and from foraging areas. The two time intervals on the right are associated with periods of rest and rumination.

trary to the pattern seen during the early morning (04:00–07:00) or evening hours (Fig. 4). Accordingly, it seemed reasonable to use a model with the non-parametric term $H_1[x(t), y(t)]$ in Eq. 6 set to a constant, and consequently, its partial derivatives, H_{1x} , H_{1y} , in Eq. 8 are zero. The exact model fitted to the data on treatment days is shown in the R-code given in the Supplement. The estimated parameter values for $\{\beta_1, ..., \beta_4\}$

were next used in Eq. 7 to produce Fig. 5 depicting the height of the potential surface as a function of the distance between elk and the four disturbances. The estimated curves in Fig. 5 seem to indicate that elk within a few hundred meters of any of the four disturbances were, on average, moving away (repelled) from the disturbance. The repulsion appeared to be strongest in terms of both magnitude and

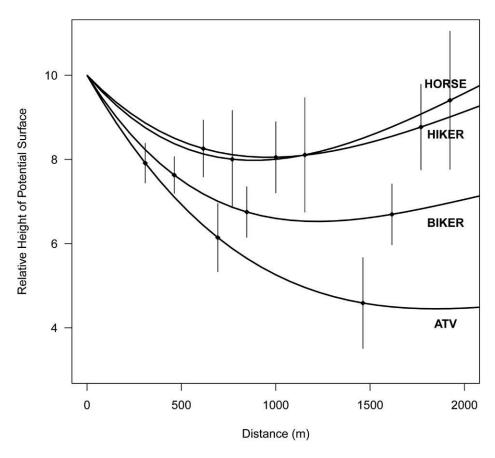


Fig. 5. Estimated height of the potential surfaces as a function of the distance to each of the four disturbances, ATV riding (ATV), bike riding (BIK), hiking (HIK) and horseback riding (HRS). Vertical lines are ± 2 SE bounds.

distance for the ATV treatment, with some repulsion observed up to 1.0 km. Repulsion from bikers was detected up to about 500 m, after which the function is not significantly different from a horizontal line. The smallest estimated repulsion effect was observed between elk and hikers and elk and horseback riders, with significant repulsion observed only up to about 200 m.

Potential surfaces were next evaluated for treatments at selected locations along the treatment routes. This type of analysis can be used to predict the potential impact of human activities at specific locations. A map of the study area and the treatment routes is shown in Fig. 6. Overlaid on the maps are the estimated potential functions assuming the disturbance is at a particular location on the route. When an ATV was at the indicated location, the estimated potential function decreased from a value of 10.0 to 4.0. These

values are significantly higher than the minimum level (grey, flat area) was; according to our potential function model, elk appear to be feeling no force of attraction or repulsion. When the disturbance was a hiker or a horseback rider, the repulsion appeared much less and the potential surface is mostly flat (Figs. 4 and 5) with no regions of attraction or repulsion, i.e., similar to days when no human disturbance was allowed in the study area.

DISCUSSION

We have presented a framework for studying the movement of animals (in particular attraction and repulsion) that couples a conceptual model of behavior with statistical methods and estimation procedures. In this framework, motion of an animal is assumed to be affected by its location relative to surrounding biophysical factors, in-

21508252, 2013, 3, Downadaed from https://esajournals.onlinelibrary.wiely.com/ab/10.1899ES12-00286.1 by Yale University, Wiley Online Library on [07/10/2024]. See the Terms and Conditions (https://onlinelibrary.wiely.com/terms-and-conditions) on Wiley Online Library for rules of use; OA articles are governed by the applicable Creative Commons Licensea

Fig. 6. Estimated potential surfaces when the human disturbance was located at a particular point (red triangle) on the route within the study area. The grey areas indicate regions where the potential surface is approximately flat.

cluding conspecifics, predators, humans, or sources of food. We also provided example SDEs and accompanying R scripts (Supplemental Material) that can be modified to implement the framework as part of other movement studies concerned with detecting and quantifying land-scape patterns of movements and their timing with respect to different behavioral states (e.g., foraging, migration, avoidance).

Our methods were initially developed to

explain animal movement in terms of human and ecological disturbances, and the example we presented concerned quantifying the strength of repulsion of elk from four different human disturbances. This is in contrast to a number of other recent studies that characterized temporal changes in movement patterns associated with different behavioral states (Forester et al. 2007, Dowd and Joy 2011). Potential surfaces can be applied to studies used to model animal interac-

tions with patch boundaries and core area by including locational variables like distance to patch edge or centroid. State changes can be then modeled as a bivariate function that includes both distance metrics and time as the explanatory variables. Time steps can be hour of day or season in year (or both) depending on the temporal scale of the data and the movement behavior in question. An example of fitting such a potential function is given in Preisler et al. (2004) where the effect of habitat features, such as refugia from roads, food patches, streams and canyons on movement were studied by time of day and season. In movement studies on marine mammals, for example, where sea surface temperature can be an important explanatory variable (e.g., Jonsen et al. 2003), a univariate function of the changing temperature, H(T(t)), can be used as a potential function, with points of attraction (valleys in the potential surface) being regions with higher temperatures. One may use a polynomial function with temperature, T, as an explanatory variable. Note that in this example temperature is a function of time, thus it is possible to represent time in the model (i.e., a potential function varying in time) without having to include it explicitly in the equation. In Brillinger et al. (2008) the authors study the movement of Hawaiian monk seals (Monachus schauinslandi) using a time varying potential function with two points of attraction to describe the migration of the seal as it moves between a foraging region in the sea and resting areas on land.

The complexity of the stochastic component of a movement model is determined by the data at hand. In our elk example, serial correlation was detected beyond that of a Markov process and an autoregressive model was employed. A Lévy process (diffusion process with jumps) can be useful for studying reactions of animals to outside stimuli (e.g., cars, or other animals). However, it was not used in the present study because locations and times of the disturbances were known. More complicated error structures, such as when measurement errors are too large to ignore, may necessitate the use of a linear or non-linear state-space model. However, one may still base the form of the difference equation on a function derived from a hypothesized potential function.

Brownian-bridges are another technique that has been used to analyze movements. In particular, when the animal's location is known at a sequence of locations, one can estimate positions in between using Brownian bridges (Horne et al. 2007), and then estimate the utilization distribution as a function of location-specific habitat variables. Eq. 3 of our model shows the relationship between the utilization distribution and a potential function independent of time. Thus Brownian-bridges models may be linked to a particular form of a potential function.

Elapsed time between consecutive observations was small (~5 min) in the present study. However, the utility of potential function models is not limited to high frequency data; the approach can be applied to whatever spatiotemporal scale is of interest. In the present study short-term responses to human disturbances were of interest, whereas in the Hawaiian monk seal example (Brillinger et al. 2008) the data were daily observations. In an example on the transition of Sierra Nevada yellow-legged frogs (Rana sierrae) to breeding, feeding and overwintering locations the time steps were months (Matthews and Preisler 2010). One can also interpolate between consecutively observed locations to increase the temporal resolution of location data using Brownian bridges (Horne et al. 2007) or simple linear or spline interpolation (Brillinger et al. 2011).

Our work is motivated in part by the fact that models for analyzing movement trajectories can provide ecologists with valuable insights into the behavioral mechanisms that control movement parameters. The study of movement is fundamental to understanding individual and population responses to emergent anthropogenic and natural disturbance. Recent studies have followed three lines of investigation (Schick et al. 2008), namely: (1) modeling for realistic movement, (2) animal-environment interactions, and (3) inferring movement when the data are incomplete. We offer a stochastic model that can be applied to all three lines of investigation. SDEs can be used to statistically model the ecology of animal movement and associated behavioral states in continuous time, as seen in our maps of movement vectors. The approach also encapsulates basic mechanisms of movement, namely, internal state, navigation, and

external influences (Nathan et al. 2008) as timespace explanatories. Moreover, the potential function surfaces can have underlying biological meaning that can lead to insights about the factors that control movement patterns, such as energetic considerations and habitat quality. For example, conditions of high habitat quality and low energetic cost to an animal would be indicated by a potential surface and movement vector of animal use that is concentrated in small areas, centered on the best habitats. By contrast, conditions of low habitat quality and high energetic cost would be indicated by a potential surface and movement vectors where animal use is substantially less concentrated and more widely distributed across a larger area. Similarly, environmental conditions that cause animal repulsion (e.g., roads, human activities) can be accurately mapped and quantified, as shown by the patterns of animal avoidance in the present study (Figs. 4 and 5).

Deriving ecological inferences from animal movement data has been difficult, in part because it is a multistate, stochastic process (McClintock et al. 2012). Elk, like many ungulates, exhibit pronounced switching behavior between fine- and broad-scale movements, corresponding to crepuscular transitions in habitat preferences (Ager et al. 2003). Different landscape features and associated scales of perception may be responsible for decision making and navigation during these phases. State-space models are one approach to handle multiphasic movements where both time-dependent and time-independent factors must be considered (Forester et al. 2007). More recently, nested, discrete, multi-state movement models have been proposed as a simplified framework to facilitate the analyses of the growing body of movement data (McClintock et al. 2012). We submit that SDEs derived from potential functions enhance the above approaches by connecting a physical model to the discrete difference equations that may then be readily estimated with existing statistical packages.

ACKNOWLEDGMENTS

We thank Professor David Brillinger for his comments and help during the preparation of this manuscript. We thank Jennifer Hafer and Bridgett Naylor for data preparation and management. We are

also thankful to two referees and the associate editor for their valuable comments on the manuscript. Special thanks go to Michelle Day for her help with Fig. 2 and other copy editing issues. This research was supported with funding from the Oregon Department of Parks and Recreation and the USDA Forest Service Pacific Northwest and Pacific Southwest Research Stations.

LITERATURE CITED

- Ager, A. A., B. K. Johnson, J. W. Kern, and J. G. Kie. 2003. Daily and seasonal movements and habitat use by female rocky mountain elk and mule deer. Journal of Mammalogy 84:1076–1088.
- Brillinger, D. R. 2010. Modelling spatial trajectories. Pages 463–474 *in* A. Gelfand, P. Diggle, P. Guttorp, and M. Fuentes, editors. Handbook of spatial statistics. CRC Press, Boca Raton, Florida, USA.
- Brillinger, D. R., H. K. Preisler, A. A. Ager, and J. G. Kie. 2004. An exploratory data analysis (EDA) of the paths of moving animals. Journal of Statistical Planning and Inference 122:43–63.
- Brillinger, D. R., H. K. Preisler, and M. J. Wisdom. 2011. Modelling particles moving in a potential field with pairwise interactions and an application. Brazilian Journal of Probability and Statistics 25:421–436.
- Brillinger, D. R., B. S. Stewart, and C. L. Littnan. 2008. Three months journeying of a Hawaiian monk seal. Institute of Mathematical Statistics Collection. Probability and Statistics: Essays in Honor of David A. Freedman 2:246–264.
- Codling, E. A., and N. A. Hill. 2005. Sampling rate effects on measurements of correlated and biased random walks. Journal of Theoretical Biology 233:573–588.
- Dowd, M., and R. Joy. 2011. Estimating behavioral parameters in animal movement models using a state-augmented particle filter. Ecology 92:568–575.
- Forester, J. D., A. R. Ives, M. G. Turner, D. P. Anderson, D. Fortin, H. L. Beyer, D. W. Smith, and M. S. Boyce. 2007. State-space models link elk movement patterns to landscape characteristics in Yellowstone National Park. Ecological Monographs 77:285–299.
- Hastie, T. J., and R. J. Tibshirani. 1990. Generalized additive models. Chapman & Hall, New York, New York, USA.
- Hirsch, M. W., and S. Smale. 1974. Differential equations, dynamical systems, and linear algebra. Academic Press, New York, New York, USA.
- Hirsch, M. W., S. Smale, and R. L. Devaney. 2004. Differential equations, dynamical systems, and an introduction to chaos. Academic Press, Elsevier, Waltham, Massachusetts, USA.
- Horne, J. S., E. O. Garton, S. M. Krone, and J. S. Lewis. 2007. Analyzing animal movements using Brownian bridges. Ecology 88:2354–2363.

- Johnson, D. S., J. M. London, M. A. Lea, and J. W. Durban. 2008. Continuous-time correlated random walk model for animal telemetry data. Ecology 89:1208–1215.
- Jonsen, I. D., R. A. Myers, and J. M. Flemming. 2003. Metaanalysis of animal movement using statespace models. Ecology 84:3005–3063.
- Matthews, K. R., and H. K. Preisler. 2010. Site fidelity of the declining amphibian *Rana sierrae* (Sierra Nevada yellow-legged frog). Canadian Journal of Fisheries and Aquatic Sciences 67:243–255.
- McClintock, B. T., R. King, L. Thomas, J. Matthiopoulos, B. J. McConnell, and J. M. Morales. 2012. A general discrete-time modeling framework for animal movement using multistate random walks. Ecological Monographs 82:335–349.
- Millspaugh, J. J., R. M. Nielson, L. L. McDonald, J. M. Marzluff, R. A. Gitzen, C. Rittenhouse, M. W. Hubbard, and S. L. Sheriff. 2006. Analysis of resource selection using utilization distributions. Journal of Wildlife Management 70:384–395.
- Morales, J. M., T. Haydon, J. Frair, K. F. Holsinger, and J. M. Fryxell. 2004. Extracting more out of relocation data: Building movement models as mixtures of random walks. Ecology 85:2436–2445.
- Nathan, R., W. Getz, E. Revilla, M. Holyoak, R. Kadmon, D. Saltz, and P. Smouse. 2008. A movement ecology paradigm for unifying organismal movement research. Proceedings of the National Academy of Sciences USA 105:19052–19059.
- Naylor, L. M., M. J. Wisdom, and R. G. Anthony. 2009. Behavioral responses of North American elk to recreational activity. Journal of Wildlife Management 37:328–338.
- Okubo, A., and S. A. Levin. 2001. Diffusion and ecological problems: modern perspectives. Springer-Verlag, New York, New York, USA.
- Prakasa Rao, B. L. S. 1999. Statistical inference for diffusion type processes. Arnold, London, UK.
- Preisler, H. K., A. A. Ager, B. K. Johnson, and J. G. Kie. 2004. Modeling animal movements using stochastic differential equations. Environmetrics 15:643–657.
- Preisler, H. K., A. A. Ager, and M. J. Wisdom. 2006.

- Statistical methods for analyzing responses of wildlife to human disturbance. Journal of Applied Ecology 43:164–172.
- Sawyer, H., and M. J. Kaufman. 2011. Stopover ecology of a migratory ungulate. Journal of Animal Ecology 80:1078–1087.
- Sawyer, H., M. J. Kauffman, R. M. Nielson, and J. S. Horne. 2009. Identifying and prioritizing ungulate migration routes for landscape-level conservation. Ecological Applications 19:2016–2025.
- Schick, R. S., S. R. Loarie, F. Colchero, B. D. Best, A. Boustany, D. A. Conde, P. N. Halpin, L. N. Joppa, C. M. McClellan, and J. S. Clark. 2008. Understanding movement data and movement processes: current and emerging directions. Ecology Letters 11:1338–1350.
- Smouse, P. E., S. Focardi, P. R. Moorcroft, J. G. Kie, J. D. Forester, and J. M. Morales. 2010. Stochastic modelling of animal movement. Philosophical Transactions of the Royal Society 365:2201–2211.
- Sorensen, M. 1997. Estimating functions for discretely observed diffusions: A review. IMS Lecture Notes Monograph Series 32:305–325.
- Taylor, J. R. 2004. Classical mechanics. University Science Books, Sausalito, California, USA.
- Turchin, P. 1998. Quantitative analysis of movement: measuring and modeling population redistribution in animals and plants. Sinauer, Sunderland, Massachusetts. USA.
- Wisdom, M. J., A. A. Ager, H. K. Preisler, N. J. Cimon, and B. K. Johnson. 2004. Effects of off-road recreation on mule deer and elk. Transactions, North American Wildlife and Natural Resources Conference 69:531–550.
- Wood, S. N. 2006. Low rank scale invariant tensor product smooths for generalized additive mixed models. Biometrics 62:1025–1036.
- Zohdi, T. L. 2003. Computational designth of swarms. International Journal for Numerical Methods in Engineering 57:2205–2219.
- Zohdi, T. L. 2009. Mechanistic modeling of swarms. Computational Methods in Applied Mechanical Engineering 198:2039–2051.

SUPPLEMENTAL MATERIAL

SUPPLEMENT

R script and resulting output for estimating the potential surface described in Eq. 3 with data from elk during ATV treatment days (*Ecological Archives* C004-002-S1).