

Math CS 121 HW 2

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Question 1. 2.3.11 (pg. 51):

Six fair dice are tossed. What is the probability that at least two of them show the same face?

Proof. Here it'll be easier to consider the complement. Let A denote event with at least two of the dice show the same face, while A^c represents all outcomes with all dice showing distinct face.

If do consider the order of the dice, there are total of 6^6 total outcomes (since each dice has 6 possible faces, each face are equally probable to occur, while there are 6 dices). Then, for an event of A^c to happen, the dices need to show 6 distinct numbers, while there are total of 6 numbers the dices can show. Hence, They need to pick (or permute) 6 numbers out of total of 6 numbers, hence providing $6P6 = 6!$ ways for the dices to show distinct faces (if considering order).

Hence, $\mathbb{P}(A^c) = \frac{6!}{6^6} = \frac{5}{324}$, showing that $\mathbb{P}(A) = 1 - \mathbb{P}(A^c) = \frac{319}{324}$. So, the probability that at least two of the dice show the same face, is $\frac{319}{324}$. \square

Question 2. 2.3.19 (pg. 52):

One of the five elevators in a building starts with seven passengers and stops at nine floors. Assuming that it is equally likely that a passenger gets off at any of these nine floors, find the probability that at least two of these passengers will get off at the same floor.

Proof. Again it'll be easier to consider the complement. Let B denotes the event with at least two passengers getting off at the same floor, while B^c represents all outcomes with all passengers get off at distinct floors.

If do consider the order of them dropping off, there are total of 9^7 total outcomes (since each person has 9 possible floors to get dropped off, while there are 7 people). Then, for an event of B^c to happen, the people need to be dropped off at 7 distinct floors, while there are total of 9 floors. Since it's equally probable that a passenger gets off at any of the floor, then it suffices to count total number of outcomes where all people get off at distinct floors. Which, it essentially picks (or permutes) 7 floors out of the 9 floors, hence providing $9P7 = \frac{9!}{2!}$ ways for the passengers to get off at distinct floors (of consider order).

Hence, $\mathbb{P}(B^c) = \frac{9!/2!}{9^7} = \frac{2240}{59049}$, so $\mathbb{P}(B) = 1 - \mathbb{P}(B^c) = \frac{56809}{59049}$. This is the probability that at least two passengers get off at the same floor. \square

Question 3. 2.3.24 (pg. 52):

Let S and T be finite sets with n and m elements, respectively.

- (a) How many functions $f : S \rightarrow T$ can be defined?
- (b) If $m \geq n$, how many injective (one-to-one) functions $f : S \rightarrow T$ can be defined?
- (c) If $m = n$, how many surjective (onto) functions $f : S \rightarrow T$ can be defined?

Proof.

- (a) Since each element $s \in S$ can be sent to any element in T , hence there are total of $|T| = m$ choices where s can be sent to. And, there are total of $|S| = n$ elements, each endow with such m choices. Hence, there are total of m^n distinct ways of defining functions $f : S \rightarrow T$.
- (b) Given $m \geq n$, for a function to be injective, all $s \in S$ must be mapped to a distinct elements. Hence, since $|S| = n$, one needs to pick out n distinct elements in T for the elements in S to get mapped to. Hence, there are total of $mPn = \frac{m!}{(m-n)!}$ ways of mapping each element in S to a distinct element in T , which generates $\frac{m!}{(n-m)!}$ distinct injective functions $f : S \rightarrow T$.
- (c) Given $m = n$, since $|S| = |T|$, then $f : S \rightarrow T$ is surjective \iff it is injective. Hence, apply part (b), there are total of $\frac{m!}{(n-m)!} = \frac{m!}{0!} = m!$ surjective functions $f : S \rightarrow T$.

□

Question 4. 2.4.31 (pg. 67):

According to the 1998 edition of Encyclopedia Britannica, "there are at least 15,000 to as many as 35,000 species of orchids." These species have been found naturally and are distinct from each other. Suppose that hybrids can be created by crossing any two existing species. Furthermore, suppose that hybrids themselves can be continued to be hybridized with each other or with an original species. Orchid lovers develop thousands and thousands of hybrids for use as garden or greenhouse ornamental and for the commercial flower trade. Suppose that all species are crossed, two at a time, to create the first generations of hybrids. Then the first generation of hybrids are crossed with each other and with the original species, two at a time, to develop the second generation of hybrids. The second generation of hybrids are crossed with each other, with the first generation hybrids, and with the original species, two at a time, to generate the third generation of hybrids, and so on. Let n be the total number of original species of orchids. Let n_k be the number of hybrids in the k^{th} generation.

- (a) Find n_k in terms of n, n_1, \dots, n_{k-1} .
- (b) For $n = 25,000$, find the largest possible total number of all hybrids in the first four generations.

Proof.

- (a) For n_1 , since they're produced by pairing up 2 distinct species from the zeroth generation (total of n generations), then every distinct hybrid in the first generation is through choosing 2 distinct species from total of n species, which there are $n_1 = nC2 = \binom{n}{2}$ first generation hybrids.

Now, inductively, for $k > 1$, since each k^{th} generation hybrid is produced by crossing the $(k-1)^{\text{th}}$ generation with one of the $l \leq (k-1)$ generation, two at a time. Since crossing with each distinct generation is disjoint from each other (each generation is disjoint from each other), the amount of k^{th} generation can be computed by summing up the total number of crossings between the k^{th} generation, with any $l \leq (k-1)$ generation (where $l = 0$ is assumed to be the original generation, say $n_0 = n$).

For the crossings within the $(k-1)^{\text{th}}$ generation is $n_{k-1}C2 = \binom{n_{k-1}}{2}$, and for the crossings of the $(k-1)^{\text{th}}$ generation with any $l < (k-1)$ generation, is given by $n_l \cdot n_{k-1}$ (for each l^{th} generation species, there are n_{k-1} distinct ways of crossing with a $(k-1)^{\text{th}}$ generation species, and there are total of n_l distinct l^{th} generation species). Hence, the total is given by:

$$n_k = \binom{n_{k-1}}{2} + \sum_{l=0}^{k-2} n_l \cdot n_{k-1} \quad (1)$$

(b) Given $n = 25,000$, the first four generations are given as follow:

$$\begin{aligned} - n_1 &= \binom{25,000}{2} \\ - n_2 &= \binom{n_1}{2} + n_1 \cdot n. \\ - n_3 &= \binom{n_2}{2} + n_2 \cdot n_1 + n_2 \cdot n. \\ - n_4 &= \binom{n_3}{2} + n_3 \cdot n_2 + n_3 \cdot n_1 + n_3 \cdot n. \end{aligned}$$

Which, the largest possible number is $n_1 + n_2 + n_3 + n_4$, where the number is too large for manual calculation.

□

Question 5. 2.4.34 (pg. 67):

In a closet there are 10 pairs of shoes. If six shoes are selected at random, what is the probability of (a) no complete pairs; (b) exactly one complete pair; (c) exactly two complete pairs; (d) exactly three complete pairs?

Proof. Recall that there are 10 pairs of shoes, hence 20 distinct shoes. Choosing 6 shoes at random, there are total of $20C6 = \binom{20}{6}$ ways of doing so.

(a) Here, we'll view each pair of shoes as a whole first. To ensure there's no complete pairs, each of the 6 shoe must be coming from a distinct pair, hence providing $\binom{10}{6}$ ways of choosing distinct pairs. And, after distinguishing the 6 distinct pairs, since one can freely choose one out of the two shoes in

a pair, there are total of 2^6 ways of doing so. Hence, there are $2^6 \cdot \binom{10}{6}$ ways of having 6 shoes with no complete pair, which induces the following probability:

$$\mathbb{P}(\text{No complete pairs}) = \frac{2^6 \cdot \binom{10}{6}}{\binom{20}{6}} \quad (2)$$

- (b) For exactly one complete pairs, one must first fix the complete pair, where there are 10 distinct ways of doing so. Then, after distinguishing the complete pair (2 shoes), the rest of the 4 shoes must all be from distinct pairs of shoes (out of the remaining 9 pairs). Hence, there are $\binom{9}{4}$ ways of choosing distinct pairs out of the remaining pairs. Again, after distinguishing teh 4 distinct pairs, one can freely choose one of the two shoes in each pair, creating total of 2^4 ways of doing so. Hence, there are $10 \cdot 2^6 \cdot \binom{9}{4}$ ways of having 6 shoes with exactly one complete pair, this induces the following probability:

$$\mathbb{P}(\text{Exactly one complete pair}) = \frac{10 \cdot 2^6 \cdot \binom{9}{4}}{\binom{20}{6}} \quad (3)$$

- (c) To have exactly two complete pairs, one must first fix the two pairs, where there are $\binom{10}{2}$ ways of doing so. Then, after distinguishing the two complete pairs (4 shoes), the rest of the 2 shoes must be from distinct pairs, out of the remaining 8 pairs. Hence, there are $\binom{8}{2}$ ways of choosing 2 distinct pairs out of the remaing 8 pairs. Again, after distinguishing the 2 pairs, one can freely choose one of the two shoes in each pair, creating a total of 2^2 ways of doing so. Hence, there are $\binom{10}{2} \cdot 2^2 \cdot \binom{8}{2}$ ways of having 6 shoes with exactly two complete pairs, inducign the following probability:

$$\mathbb{P}(\text{Exactly two complete pairs}) = \frac{\binom{10}{2} \cdot 2^2 \cdot \binom{8}{2}}{\binom{20}{6}} \quad (4)$$

- (d) To have exactly three complete pairs, one must fix 3 pairs of shoes out of the 10 pairs. Hence, providing a total of $\binom{10}{3}$ ways of doing so. Which, it induces the following probability:

$$\mathbb{P}(\text{Exactly three complete pairs}) = \frac{\binom{10}{3}}{\binom{20}{6}} \quad (5)$$

□

6 Not done

Question 6. 2.4.44 (pg. 68):

A lake has N trout, and t of them are caught at random, tagged, and returned. We catch n trout at a later time randomly and observe that m of them are tagged.

- (a) Find P_N , the probability of what we observed actually happen.
- (b) To estimate the number of trout in the lake, statisticians find the value of N that maximizes P_N . Such a value is called the **maximum likelihood estimator** of N . Show that the maximum of P_N is $[nt/m]$, where by $[nt/m]$ we mean the greatest integer is less than or equal to nt/m . That is, prove that the maximum likelihood estimator of the number of trout in the lake is $[nt/m]$.

(Hint: Investigate for what values of N the probability P_N is increasing and what values it is decreasing).

Proof.

- (a) Overall, to catch n trouts when there are total of N trouts, it's given by $\binom{N}{n}$. Then, if out of the n trouts being caught, there are m of them that are previously labeled, it implies one must have chosen m out of t trouts that have been labeled previously, which has a total of $\binom{t}{m}$ cases; and, since the rest $(n - m)$ non-labeled trouts must be chosen from the $(N - t)$ trouts that are not labeled, there are $\binom{N - t}{n - m}$ cases. Hence, in general we have the following probability:

$$P_N = \frac{\binom{t}{m} \binom{N - t}{n - m}}{\binom{N}{n}} = \frac{n! \cdot (N - n)! \cdot t! \cdot (N - t)!}{N! \cdot m! \cdot (t - m)! \cdot (n - m)! \cdot (N - t - n + m)!} \quad (6)$$

- (b) To investigate when the probability is increasing, we'll consider the ratio $\frac{P_{N+1}}{P_N}$ for all valid $N \in \mathbb{N}$ (which $N \geq t, n, m$ at least). Using the formula derived in part (a), the ratio is expressed as follow:

$$\frac{P_{N+1}}{P_N} = \frac{n! \cdot (N + 1 - n)! \cdot t! \cdot (N + 1 - t)!}{(N + 1)! \cdot m! \cdot (t - m)! \cdot (n - m)! \cdot (N + 1 - t - n + m)!} \cdot \frac{N! \cdot m! \cdot (t - m)! \cdot (n - m)! \cdot (N - t - n + m)!}{n! \cdot (N - n)! \cdot t! \cdot (N - t)!} \quad (7)$$

$$= \frac{(N + 1 - t)(N + 1 - n)}{(N + 1)(N + 1 - t - n + m)} \quad (8)$$

Set $\frac{P_{N+1}}{P_N} > 1$, one yields the follow:

$$N^2 + N - nN + N + 1 - n - tN - t + tn > N^2 + N - tN - nN + mN + N + 1 - t - n + m \quad (9)$$

$$\implies tn > mN + m \implies (N + 1) < \frac{tn}{m} \implies N < \frac{tn}{m} - 1 \quad (10)$$

Which, for $N' = N + 1 = \lceil tn/m \rceil$, it satisfies $\frac{P_{N'}}{P_N} = \frac{P_{N+1}}{P_N} > 1$ since N satisfies the above inequality, while $\frac{P_{N'+1}}{P_{N'}} \leq 1$ since $N' = N + 1$ doesn't satisfy the above inequality. Hence, we deduce that $P_{N'} \geq P_N, P_{N'+1}$, or $P_{N+1} \geq P_N, P_{N+2}$.

Hence, at $N' = N + 1 = \lceil tn/m \rceil$, it's when the probability goes from increasing to decreasing, which is the maximum likelihood estimator of $P_{N'}$.

□

Question 7. 2.4.51 (pg. 70):

A fair coin is tossed n times. Calculate the probability of getting no successive heads.

(Hint: Let x_i be the number of sequences of H's and T's of length i with no successive H's. Show that x_i satisfies $x_i = x_{i-1} + x_{i-2}$, $i \geq 2$, where $x_0 = 1$ and $x_1 = 2$. The answer is $x_n/2^n$. Note that $\{x_i\}_{i=1}^\infty$ is a Fibonacci-type sequence).

Proof. Based on the hint, if tossed $i = 0$ times of a fair coin, there is total of 1 outcome (namely no tosses), which vacuously satisfies the condition of no successive H's, hence $x_0 = 1$. For $i = 1$ toss of a fair coin, there is total of 2 outcomes (H or T), which again vacuously satisfies the condition of no successive H's. Hence $x_1 = 2$.

Now, for $i \geq 2$, we can classify such sequences into 2 types: One with the $(i - 1)^{\text{th}}$ toss getting H, and the other type getting T.

If the $(i - 1)^{\text{th}}$ toss yields H, in case to avoid successive H's, the i^{th} toss is limited to T (which only has 1 possible outcome). Else, if the $(i - 1)^{\text{th}}$ toss yields T, the i^{th} toss can be either H or T (which has 2 possible outcomes).

Let H_{i-1}, T_{i-1} denote numbers of $(i - 1)$ sequence of tosses with no successive H's, that end in H and T respectively. Then, $H_{i-1} + T_{i-1} = x_{i-1}$ (since this collects all cases of $(i - 1)$ tosses of sequences that satisfy the desired condition), and $H_{i-1} + 2T_{i-1} = x_i$ (since H_{i-1} collects all sequence of length i lands with H on the $(i - 1)^{\text{th}}$ position, which each has one outcome; while T_{i-1} collects all sequence of length i lands with T on the $(i - 1)^{\text{th}}$ position, which each has two distinct outcomes).

However, also notice that for every sequence of length $(i - 2)$, regardless if the $(i - 2)^{\text{th}}$ toss being H or T, the $(i - 1)^{\text{th}}$ toss is valid to get T. Hence, every sequence of tosses with length $(i - 2)$ after concatenating T as the $(i - 1)^{\text{th}}$ toss, becomes a sequence of toss with length $(i - 1)$, and ends at T. Hence, $T_{i-1} = x_{i-2}$.

As a result, we have $x_i = H_{i-1} + 2T_{i-1} = (H_{i-1} + T_{i-1}) + T_{i-1} = x_{i-1} + x_{i-2}$. Where, for n tosses, there are total of x_n sequences with such property, hence providing the probability of $\mathbb{P}_n = \frac{x_n}{2^n}$ (since there are 2^n distinct sequences of tosses with length n).

Notice that this forms a Fibonacci sequence starting with $x_0 = 1$ and $x_1 = 2$. For $n = 5, 10, 20$, since the Fibonacci Sequence has value $x_5 = 8$, $x_{10} = 89$, and $x_{20} = 10946$. Which, $\mathbb{P}_5 = 0.03125$, $\mathbb{P}_{10} \approx 0.0869$, and $\mathbb{P}_{20} \approx 0.0104$. □

Question 8. Review 2.R.10 (pg. 72):

In how many arrangements of the letters BERKELEY are all three E's adjacent?

Proof. Since there are 8 letters total, if view each E as distinct initially, there are total of $8!$ ways of arranging these letters. Yet, since we're identifying the E 's as identical, out of the 3 positions the E 's acquired, there are total of $3! = 6$ ways of arranging them (that are considered the same). Hence, there are total of $\frac{8!}{3!}$ ways of arranging the letters BERKELEY, without identifying each E as distinct.

Now, if all three E 's are adjacent, one can consider them as a single letter, hence remaining with total of 6 letters: B, EEE, R, K, L, Y. Which, there are total of $6!$ ways of arranging these 6 new letters. Hence, the probability of having all three E 's being adjacent is as follow:

$$\mathbb{P}(\text{All three E's are adjacent}) = \frac{6!}{8!/3!} = \frac{6}{56} \quad (11)$$

□

Question 9. *Review 2.R.16 (pg. 73):*

Suppose that four women and two men enter a restaurant and sit at random around a table that has four chairs on one side and another four on the other side. What is the probability that the men are not all sitting on one side?

Proof. Since there are 6 people and 8 chairs, there are total of $8P6 = \frac{8!}{(8-6)!}$ ways of picking seats for people to sit down (when viewing each person as distinct appearance). However here if people are identified only with the specified genders (namely men and women here), there are $4!$ ways of letting 4 women sitting in specified 4 chairs, while $2!$ ways of letting 2 men sitting in specified 2 chairs. Hence, the total ways of letting them sit down (without identifying people with information besides gender) is $\frac{8!}{(8-6)! \cdot 4! \cdot 2!}$.

Now, let A denote the event that the men are not sitting on one side, we'll first distinguish the position the men are sitting. Since each side of the table has 4 chairs, there are $4 \cdot 4 = 16$ ways of choosing one chair on each side for the two men to not sit on the same side.

For the remaining 6 chairs and 4 women, there are $\binom{6}{4}$ ways of identifying 4 distinct chairs for the women. This results in $16 \cdot \binom{6}{4} = \frac{16 \cdot 6!}{2! \cdot 4!}$ ways of having the men not sitting on the same side.

Finally, this provides the following probability:

$$\mathbb{P}(\text{Men are not on the same side}) = \frac{16 \cdot 6! / (2! \cdot 4!)}{8! / ((8-6)! \cdot 4! \cdot 2!)} = \frac{16 \cdot 6!}{8!} \cdot \frac{2! \cdot 4! \cdot 2!}{8!} = \frac{4}{7} \quad (12)$$

□

Question 10. *From the set of integers $\{1, 2, 3, \dots, 100000\}$ a number is selected at random. What is the probability that the sum of its digits is 8?*

Hint: Establish a one-to-one correspondence between the set of integers from $\{1, 2, \dots, 100000\}$ the sum of whose digits is 8, and the set of possible ways 8 identical objects can be placed into 5 distinguishable cells. Then use Example 2.22.

Proof. First, observe that the last number 100000 doesn't satisfy the given condition, hence we can ignore it (when calculating the numbers). Which, we're able to consider numbers with only at most 5 digits in the given range of numbers, and for simplicity, if the number has less than 5 digits, add 0 in the front until it "formally" has 5 digits, so WLOG one can assume the rest of the numbers all have 5 digits.

Now, for each number (in digits, not multiplication) $abcde \in \{1, 2, \dots, 100000\}$ that has digits sum $a + b + c + d + e = 8$, correspond this number with 5 distinguishable cells c_1 to c_5 , with a, b, c, d, e identical balls in them respectively.

Notice that such correspondance is injective, since for two of the ways of putting 8 identical objects in 5 distinguishable cells, each c_i must contain the same amount of balls. Hence, the number must have the same digits in each position (for each c_i to have matching numbers).

On the other hand, the correspondance is also surjective, since given a way of putting 8 balls in cells c_1 to c_5 , say each c_i has number a_i , then the number $a_1a_2a_3a_4a_5$ corresponds to the given way, has digit sum to 8 by definition, and it belongs to the set $\{1, 2, \dots, 100000\}$.

Therefore, the numbers with digits sum to 8 is in a one-to-one correspondance with the ways of partition 8 balls into 5 distinct cells. Since each ball has 5 choices, while every ball has an independent choice, then there are total of $5^8 = 32678$ different choices, showing that there are 32678 numbers in $\{1, \dots, 100000\}$, with digits sum to 8.

As a result, the probability of selecting a number with digits sum being 8, is $\mathbb{P} = \frac{32678}{100000}$. \square