Note for DS2023s (Machine Learning Section)

Topic 1: Linear Regression 一般线性回归

CASE: $n\gg p$

$$\hat{\beta}_{OLS} = (X'X)^{-1}X'y$$

- When $n\gg p$, $X'X\in\mathbb{R}^{p\times p}$, $X'y\in\mathbb{R}^{p\times 1}$, thus it is appliable to calculate in the memory
- The key is to **calculate the two matrix multiplication above (重点即为求解两矩阵乘法**,最后再求解 一个线性方程组即可)
- Note that, to maximize the calculation performance, we can combine the calculation into one(为方便 计算,常用如下方式整合矩阵): X'[X,y]
- After obtaining the two matrix, at last we have to solve the linear equation to get $\hat{\beta}$ (Remember that we should always try to avoid calcualte the inverse of a matrix)
- Moreover, if we have to consider the intercept, we can transform the data matrix (若考虑截距项): $X^* = [1,X]$

Note: TRY TO REALIZE THE PROCESS USING PYTHON WITH RDD

Topic 2: Regularization (Shrinkage Methods) 正则化(数据缩减技术)

CASE: n < p

- ullet When n < p, X'X is not reversible. In this case, OLS **DOES NOT HAVE UNIQUE SOLUTION**
 - o Note that it does not mean that there is no solution! By contrast, there are infinite sets of solutions that can hold $\min S_c = 0$ (you can think of it as there are more variables than equations, so the variables might have many solutions)(此时OLS并非无解,而是有无数解)
- ullet Such case acutally will also accures when npprox p
- In this case, OLS is too **flexible**, and the solution though perfectly fits the modle, will perform badly in test sets. It is called **OVERFIT** (过拟合、模型在测试集的表现欠佳)

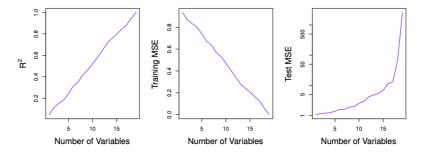


FIGURE 6.23. On a simulated example with n=20 training observations, features that are completely unrelated to the outcome are added to the model. Left: The R^2 increases to 1 as more features are included. Center: The training set MSE decreases to 0 as more features are included. Right: The test set MSE increases as more features are included.

Method 1: Ridge Regression 岭回归

ESSENCE: ℓ_2 Regularization

• Recall in OLS, our loss function:

$$\text{Loss} = ||Y - X\beta||_2 = \sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2$$

• In Ridge Regression, we add an extra penalty:

$$\operatorname{Loss}^* = ||Y - X\beta||_2 + \lambda ||\beta||_2 = \sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2 + \sum_{j=1}^p \beta_j^2$$

- We are going to find the optimum β that minimize that loss function
- Here $\lambda \geq 0$ is called the **Tunning Parameter**, $\lambda ||\beta||_2$ is called **Shrinkage Penalty**
 - \circ When $\lambda=0$, ridge regression will degenereate into OLS (此时退化为普通最小二乘); when $\lambda\to\infty$, the penalty term will squeeze out the β and approach to zero(此时各系数将趋近于0)
- The parameter estimation will relies on the selection of λ , i.e. $\hat{\beta}^R_{\lambda}$, and to choose a good λ is important. We can use methods such as Cross Valediction, etc. (可用交叉验证等手段选取一个最优的 λ 以得到较优的岭回归系数)
- ullet Note that the ridge regression **DOES NOT INCLUDE** eta_0 (在岭回归中不会对截距项进行缩减)

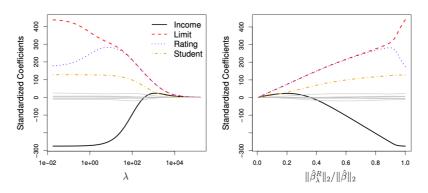


FIGURE 6.4. The standardized ridge regression coefficients are displayed for the Credit data set, as a function of λ and $\|\hat{\beta}_{\lambda}^{R}\|_{2}/\|\hat{\beta}\|_{2}$.

As is shown in FIGURE 6.4, **Income, Limit, Rating & Student** are all variables' parameter estimation. From the lest graph, as λ increases, all estimation shrinks to zero. From the right graph, the ratio shows how different the ridge estimation and OLS estimation are (1 means no difference and 0 means utmost different) (上图即展示了lambda对于参数估计的影响,其取值越大,参数约趋近于0,压缩越明显)

- In ridge regression, the final estimation of eta is $\hat{eta}_{\lambda}=(X^TX+\lambda I)^{-1}X^TY$
- ullet It can be proved that $(X^TX+\lambda I)^{-1}$ must be inversible
 - o First, it can be proved that for column full rank matrix X, X^TX is **positive semi-definite** (PSD) (首先证 X^TX 是半正定的)
 - ullet $\forall a \in \mathbb{R}^{p imes 1}, y := Xa
 eq 0$ (This is held by column full rank 由列满秩)
 - $a^T X^T X a := y^T y = \sum y_i^2 \ge 0$
 - \circ Second it can be proved that $X^TX + \lambda I$ is PD.
 - For a PSD, the eignalvalues $\alpha_i \geq 0$.
 - By adding λI whose eignvalues must be positive, the overall eignvalues $\Lambda>0$, thus the matrix is PD (其次证 $X^TX+\lambda I$ 是正定的)
 - \circ Moreover, it can be proved that $X^TX + \lambda I$ is symmetric (再可证矩阵是对称的)

- $(X^TX + \lambda I)^T = \lambda I + X^TX$
- o Last, for a symmetric PD matrix, it is inversible (对称正定阵是可逆的)
 - Recall that the determinant of a matrix is equal to the product of the eignvalues
 - Since all eignvalues are greater than zero thus the determinant is greater than zero
- Also note that $(X^TX + \lambda I)^T$ though invertible, but in this case is a rahter large matrix which is hard to actually calculate the analytical value. In this case, we choose to use **Conjugate Gradient Method** (由于矩阵很大,求解逆矩阵的解析解是困难的,常常用共轭梯度法等求解数值解)

Method 2: LASSO

ESSENCE: ℓ_1 Regularization

- **Problem** of *Ridge Regression*:
 - o Though ridge regression can squeeze the parameters to **CLOSE** to zero, it cannot set any to exact zero(岭回归永远无法真正将参数压缩至0)
 - It might be **acceptable for prediction accuracy**, but might cause **inconvenience in interpretation**, especially when *p* is large and we only need few of them
- In LASSO, the loss function is: $\operatorname{Loss} = ||Y X\beta||_2 + \lambda ||\beta||_1$
- We can see that the only difference is the **penalty term**, Ridge regression uses a ℓ_2 norm penalty regularization, while LASSO uses a ℓ_1 norm penalty regularization (即通过1-范数和2-范数进行正则化)

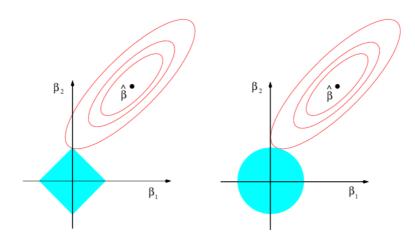


FIGURE 6.7. Contours of the error and constraint functions for the lasso (left) and ridge regression (right). The solid blue areas are the constraint regions, $|\beta_1| + |\beta_2| \le s$ and $\beta_1^2 + \beta_2^2 \le s$, while the red ellipses are the contours of the RSS.

Topic 3: Logistic Regression

Essence: Modeling for the distribution of $Y \sim \mathrm{Bernoulli}(\cdot)$

• In Logistic Regression, the final purpose is to answer a bunch of "YES or NO" "BELONGS TO or NOT BELONGS TO" problems. Such problems can be abstracted into a 0-1 questions, and there we have **Bernoulli Distribution**, which is extactly the one trying to demonstrate "how likely something will happen or not"(最终的目的是对Y的Bernoulli分布进行建模)

- For Y_1, \ldots, Y_n , they all follow some patterns of Bernoulli Distribution. But the parameters of Bernoulli distribution (i.e. p) are different for different Y's, which depends on the given information (variables X 's) (Y服从Bernoulli分布是确定的,但并不是独立的;其分布的概率参数p依赖于给定的其他解释变量X的信息)
- We can describe such dependency as $p \propto f(X)$. To be more specific, here we assume that Y depends on the linear combinition of X, thus $p \propto X\beta$ (这种相依关系这里假设通过解释变量的线性组合进行刻画)
- Another problem is that $0 \le p \le 1$, but $X\beta$ literally can be any value, thus we need another bridging function to constrain the value to [0,1] (需要另外引入一个函数使得分布在区间范围内)
 - o One possible solution: CDF
 - o Another suggested solution: **Sigmoid** (it is the most commonly used, but not the only one) $ho(x)=rac{1}{1+e^{-x}}$
- Thus we have:

$$Y|x \sim \mathrm{Bernoulli}(
ho(eta^T x))$$

- $\circ \ \
 ho(eta^T x)$ describes the possibility of Y=1
- ullet Here, the combinition patterns are still unknown. We want to have an optimum estimation to best fit the probabilities for all given Y's and X's. And here we have **Maximum Likelihood Estimation (通过** 极大似然函数以估计eta)
 - $P(Y_i = y) = p_i^y (1 p_i)^y, y = 0 \text{ or } 1$
 - $\circ \ \ l = \sum \log P(Y_i = y_i) = \sum [y_i \log p_i + (1-y_i) \log (1-p_i)]$, where $p_i =
 ho(x_i eta)$
 - \circ Usually, set loss function as $L(\beta) = -l$ for minimization
 - Unfortunately, it is also hard to get the analytical solution. We also have to use iterations to calculate the numerical solution

Topic 4: Optimization Methods