#### LECTURE 15: Linear models with normal noise

$$X_i = \sum_{j=1}^m a_{ij} \Theta_j + W_i$$
  $W_i$ ,  $\Theta_j$ : independent, normal

- Very common and convenient model
- Bayes' rule: normal posteriors
- MAP and LMS estimates coincide
- simple formulas
   (linear in the observations)
- Many nice properties
- Trajectory estimation example

## Recognizing normal PDFs

$$X \sim N(\mu, \sigma^2) \qquad f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-(x-\mu)^2/2\sigma^2} \qquad \frac{2 \langle x + \beta \rangle}{\sigma^2} = 0$$

$$c \cdot e^{-8(x-3)^2}$$
  $\mu = 3$   $\frac{1}{2\sigma^2} = 8 \implies \sigma^2 = \frac{1}{16}$   $C = \frac{1}{4}\sqrt{2\pi}$ 

$$f_X(x) = c \cdot e^{-(\alpha x^2 + \beta x + \gamma)}$$
  $\alpha > 0$  Normal with mean  $-\beta/2\alpha$  and variance  $1/2\alpha$ 

## Estimating a normal random variable in the presence of additive normal noise

$$X = \Theta + W$$

$$X = \Theta + W$$
  $\Theta, W : N(0,1),$  independent

$$f_{\Theta|X}(\theta \mid x) = \frac{f_{\Theta}(\theta) f_{X|\Theta}(x \mid \theta)}{f_{X}(x)}$$

$$f_X(x) = \int f_{\Theta}(\theta) f_{X|\Theta}(x \mid \theta) d\theta$$

$$f_{X|\Theta}(x|\theta): X = \theta + W \qquad \mathcal{N}(\theta, 1)$$

$$f_{\Theta|X}(\theta|x) = \frac{1}{f_{X}(x)} \quad c \quad e^{-\frac{1}{2}\theta^{2}} \quad c \quad e^{-\frac{1}{2}(x-\theta)^{2}} = c(x)e^{-9u \cdot \alpha d \cdot \alpha k \cdot c(\theta)}$$

Fix 
$$\alpha$$
 min  $\left[\frac{1}{2}\theta^2 + \frac{1}{2}(\alpha - \theta)^2\right]$   $\theta + (\theta - \alpha) = 0$ 

$$\hat{\theta}_{MAP} = \hat{\theta}_{LMS} = \mathbf{E}[\Theta | X = x] = 2/2$$

$$\theta + (\theta - x) = 0$$

$$\widehat{\Theta}_{MAP} = \mathbf{E}[\Theta \mid X] = \frac{\mathsf{X}}{2}$$

# Estimating a normal parameter in the presence of additive normal noise

$$X = \Theta + W$$
  $\Theta, W : N(0,1)$ , independent

$$\widehat{\Theta}_{\mathsf{MAP}} = \widehat{\Theta}_{\mathsf{LMS}} = \mathbf{E}[\Theta \,|\, X] = \frac{X}{2}$$

$$f_{\Theta|X}(\theta \mid x) = \frac{f_{\Theta}(\theta) f_{X|\Theta}(x \mid \theta)}{f_{X}(x)}$$
$$f_{X}(x) = \int f_{\Theta}(\theta) f_{X|\Theta}(x \mid \theta) d\theta$$

- Even with general means and variances:
  - posterior is normal
  - LMS and MAP estimators coincide
  - these estimators are "linear," of the form  $\widehat{\Theta} = aX + b$

## The case of multiple observations

$$X_1 = \Theta + W_1$$
  $\Theta \sim N(x_0, \sigma_0^2)$   $W_i \sim N(0, \sigma_i^2)$   
 $\vdots$   
 $X_n = \Theta + W_n$   $\Theta, W_1, \dots, W_n$  independent

$$f_{X_i|\Theta}(x_i|\theta) = \frac{c_i}{\epsilon} e^{-\left(\infty_i - \theta\right)^2/2\sigma_i^2}$$

$$f_{X|\Theta}(x|\theta) = f_{X_1,\ldots,X_n|\Theta}(x_1,\ldots,x_n|\theta) = \prod_{i=1}^n f_{X_i|\Theta}(x_i|\theta)$$

given 0=0: Wi independent => Xi independent

$$f_{\Theta|X}(\theta|x) = \frac{1}{f_{X}(x)} \cdot c_{o} e^{-(\theta-x_{o})^{2}/2\sigma_{o}^{2}} \prod_{i=1}^{n} c_{i} e^{-(x_{i}-\theta)^{2}/2\sigma_{o}^{2}}$$
Normal!

$$f_{\Theta|X}(\theta \mid x) = \frac{f_{\Theta}(\theta) f_{X|\Theta}(x \mid \theta)}{f_{X}(x)}$$

$$f_X(x) = \int f_{\Theta}(\theta) f_{X|\Theta}(x \mid \theta) d\theta$$

## The case of multiple observations

$$f_{\Theta|X}(\theta \,|\, x) = c \cdot \exp\left\{-\operatorname{quad}(\theta)\right\} \qquad \operatorname{quad}(\theta) = \frac{(\theta - x_0)^2}{2\sigma_0^2} + \frac{(\theta - x_1)^2}{2\sigma_1^2} + \dots + \frac{(\theta - x_n)^2}{2\sigma_n^2}$$

$$\frac{d}{d\theta} quad(\theta) = 0: \quad \sum_{i=0}^{n} \frac{(\theta - x_i)}{\sigma_i^2} = 0 \Rightarrow \theta \stackrel{\sum}{\geq} \frac{1}{\sigma_i^2} = \sum_{i=0}^{n} \frac{x_i}{\sigma_i^2}$$

$$\widehat{\theta}_{\mathsf{MAP}} = \widehat{\theta}_{\mathsf{LMS}} = \mathbf{E}[\Theta \mid X = x] = \frac{\sum_{i=0}^{n} \frac{x_i}{\sigma_i^2}}{\sum_{i=0}^{n} \frac{1}{\sigma_i^2}}$$

## The case of multiple observations

- Key conclusions:
  - posterior is normal
  - LMS and MAP estimates coincide
  - these **estimates** are "linear," of the form  $\hat{\theta} = a_0 + a_1x_1 + \cdots + a_nx_n$
- Interpretations:
  - estimate  $\hat{\theta}$ : weighted average of  $x_0$  (prior mean) and  $x_i$  (observations)
  - weights determined by variances

$$\widehat{\theta}_{\mathsf{MAP}} = \widehat{\theta}_{\mathsf{LMS}} = \mathbf{E}[\Theta \mid X = x] = \frac{\sum_{i=0}^{n} \frac{x_i}{\sigma_i^2}}{\sum_{i=0}^{n} \frac{1}{\sigma_i^2}} \bullet$$

## The mean squared error

$$f_{\Theta|X}(\theta \,|\, x) = c \cdot \exp\left\{-\operatorname{quad}(\theta)\right\}$$

$$\operatorname{quad}(\theta) = \frac{(\theta - x_0)^2}{2\sigma_0^2} + \frac{(\theta - x_1)^2}{2\sigma_1^2} + \dots + \frac{(\theta - x_n)^2}{2\sigma_n^2}$$

$$\widehat{\theta} = \frac{\sum_{i=0}^{n} \frac{x_i}{\sigma_i^2}}{\sum_{i=0}^{n} \frac{1}{\sigma_i^2}}$$

Performance measures:

$$\mathbf{E}\big[(\Theta - \widehat{\Theta})^2 \mid X = x\big] = \mathbf{E}\big[(\Theta - \widehat{\theta})^2 \mid X = x\big] = \text{var}(\Theta \mid X = x) = 1 / \sum_{i=0}^n \frac{1}{\sigma_i^2}$$

$$\mathbf{E}[(\Theta - \widehat{\Theta})^2] = \int E[(\Theta - \widehat{\Theta})^2 / X = \infty] \int_X (\infty) dx$$

$$f_X(x) = c \cdot e^{-(\alpha x^2 + \beta x + \gamma)}$$
  $\alpha > 0$ 

Normal with mean  $-\beta/2\alpha$  and variance  $1/2\alpha$ 

$$\alpha = \frac{1}{200^2} + 0.00 + \frac{1}{200^2}$$

## The mean squared error

$$\mathbf{E}\left[(\Theta - \widehat{\Theta})^2 \mid X = x\right] \widehat{\Theta} \mathbf{E}\left[(\Theta - \widehat{\Theta})^2\right] = 1 / \sum_{i=0}^n \frac{1}{\sigma_i^2}$$

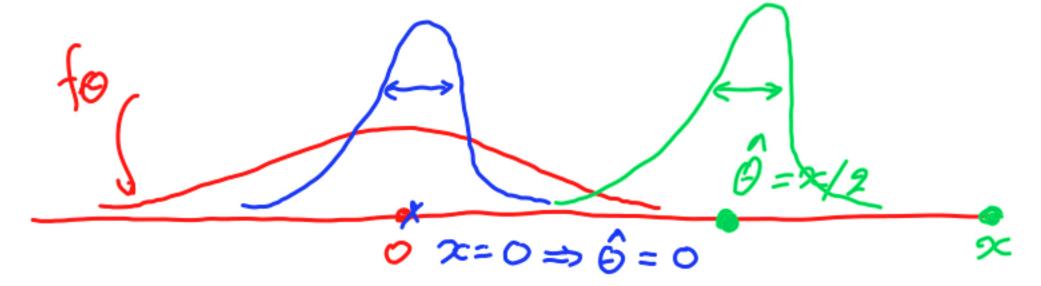
- Example:  $\sigma_0^2 = \sigma_1^2 = \dots = \sigma_n^2 = \sigma^2$
- conditional mean squared error same for all x
- Example:  $X = \Theta + W \quad \Theta \sim N(0, 1), \quad W \sim N(0, 1)$ independent  $\Theta, W$

$$\Theta \sim N(0, 1),$$

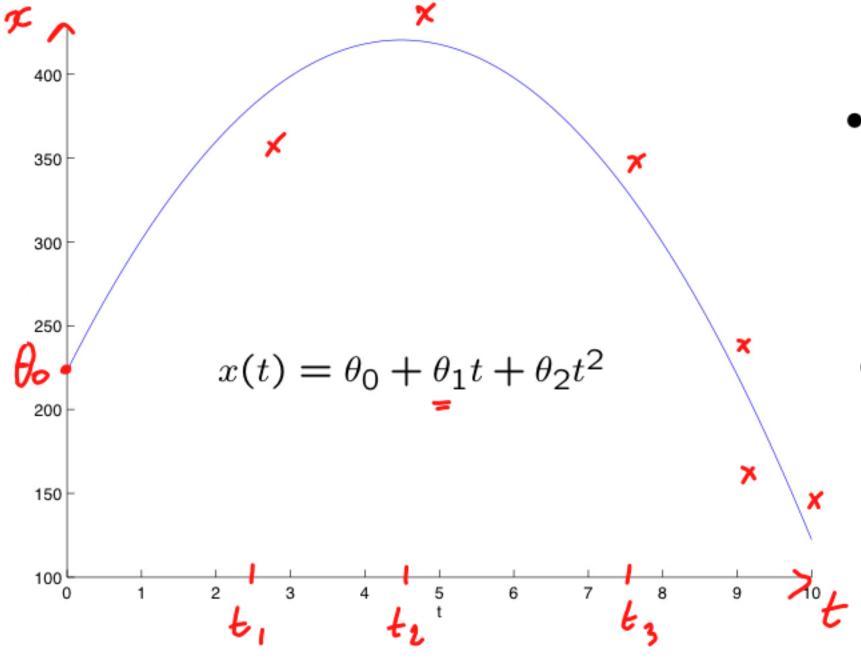
$$\widehat{\Theta} = X/2$$

$$W \sim N(0, 1)$$

$$\widehat{\Theta} = X/2 \qquad \mathbf{E} \left[ (\Theta - \widehat{\Theta})^2 \mid X = \underline{x} \right] = \frac{1/2}{2}$$



## The case of multiple parameters: trajectory estimation



• Random variables  $\Theta_0, \Theta_1, \Theta_2$  independent; priors  $f_{\Theta_i}$ 

• Measurements at times  $t_1, \ldots, t_n$ 

$$X_i = \Theta_0 + \Theta_1 t_i + \Theta_2 t_i^2 + W_i$$

noise model:  $f_{W_i}$ 

independent  $W_i$ ; independent from  $\Theta_j$ 

## A model with normality assumptions

$$X_i = \Theta_0 + \Theta_1 t_i + \Theta_2 t_i^2 + W_i \qquad i = 1, \dots, \underline{n}$$

$$f_{\Theta|X}(\underline{\theta} \mid \underline{x}) = \frac{f_{\Theta}(\theta) f_{X|\Theta}(x \mid \theta)}{f_{X}(x)}$$
$$f_{X}(x) = \int f_{\Theta}(\theta) f_{X|\Theta}(x \mid \theta) d\theta$$

- assume  $\Theta_j \sim N(0, \sigma_j^2)$ ,  $W_i \sim N(0, \sigma^2)$ ; independent
- Given  $\Theta = \theta = (\theta_0, \theta_1, \theta_2), X_i$  is:  $N\left(\theta_0 + \theta_1 t_i + \theta_2 t_i^2, \sigma^2\right)$   $f_{X_i|\Theta}(x_i|\theta) = c \cdot \exp\left\{-(x_i \theta_0 \theta_1 t_i \theta_2 t_i^2)^2 / 2\sigma^2\right\}$  posterior:  $f_{\Theta|X}(\theta|x) = \frac{1}{f_{X_i}(x)} \int_{j=0}^{x} f_{\Theta_j}(\theta_j) \int_{i=1}^{m} f_{X_i}(\theta_j) d\theta_j$   $c(x) \exp\left\{-\frac{1}{2}(\frac{\theta_0^2}{\sigma_0^2} + \frac{\theta_1^2}{\sigma_1^2} + \frac{\theta_2^2}{\sigma_2^2}) \frac{1}{2\sigma^2} \sum_{i=1}^{n} (x_i \theta_0 \theta_1 t_i \theta_2 t_i^2)^2\right\}$

## A model with normality assumptions

$$f_{\Theta|X}(\theta \mid x) = c(x) \exp\left\{-\frac{1}{2}(\frac{\theta_0^2}{\sigma_0^2} + \frac{\theta_1^2}{\sigma_1^2} + \frac{\theta_2^2}{\sigma_2^2}) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \theta_0 - \theta_1 t_i - \theta_2 t_i^2)^2\right\}$$

MAP estimate: maximize over (θ<sub>0</sub>, θ<sub>1</sub>, θ<sub>2</sub>);
 (minimize quadratic function)

### Linear normal models .

- ullet  $\Theta_{i}$  and  $X_{i}$  and are linear functions of independent normal random variables
- $f_{\Theta|X}(\theta \mid x) = c(x) \exp \left\{ -\operatorname{quadratic}(\theta_1, \dots, \theta_m) \right\}$  in ear regression
- MAP estimate: maximize over  $(\theta_1, \dots \theta_m)$ ; (minimize quadratic function)
  - $\widehat{\Theta}_{\mathsf{MAP},j}$ : linear function of  $X=(X_1,\ldots,X_n)$
- Facts:
  - $\circ \ \widehat{\Theta}_{\mathsf{MAP},j} = \mathbf{E}[\Theta_j \,|\, X]$
  - o marginal posterior PDF of  $\Theta_j$ :  $f_{\Theta_j|X}(\theta_j|x)$ , is normal
  - MAP estimate based on the joint posterior PDF:
     same as MAP estimate based on the marginal posterior PDF
  - $\circ \mathbf{E}[(\widehat{\Theta}_{i,\mathsf{MAP}} \Theta_i)^2 \mid X = x]$ : same for all x

