Global Spatial Autocorrelation

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global spatial autocorrelation

Moran scatter plot

correlogram

variogram

variogram models





Global Spatial Autocorrelation





Global Spatial Autocorrelation Measures

combination attribute similarity and locational similarity

one statistic for the whole pattern

test for clustering not for clusters (locations)





Moran's I





Moran's I

the most commonly used of many spatial autocorrelation statistics

• $I = [\Sigma_i \Sigma_j w_{ij} z_i.z_j/S_0]/[\Sigma_i z_i^2 / N]$

with $z_i = y_i - m_x$: deviations from mean

cross product statistic $(z_i.z_j)$ similar to a correlation coefficient

value depends on weights (wij)





Moran's I examined more closely

scaling factors in numerator and denominator

in numerator: $S_0 = \sum_i \sum_j w_{ij}$

the number of non-zero elements in the weights matrix, or the number of neighbor pairs (x2)

in denominator: N

the total number of observations





Inference

how to assess whether computed value of Moran's I is significantly different from a value for a spatially random distribution

compute analytically (assume normal distribution, etc.)

computationally, compare value to a reference distribution obtained from a series of randomly permuted patterns





Standardized z-value

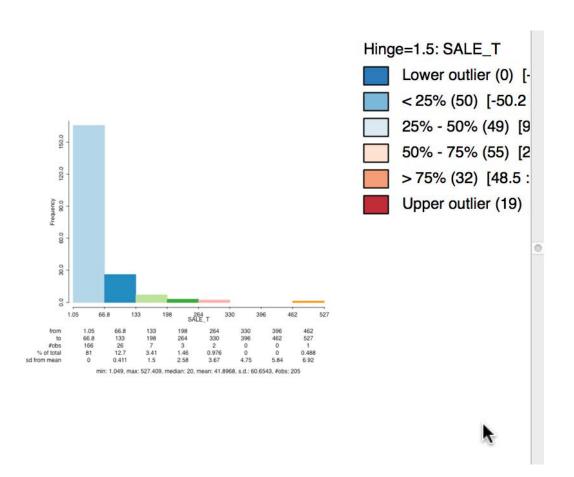
standardize by subtracting mean and dividing by standard deviation, computed from the reference distribution

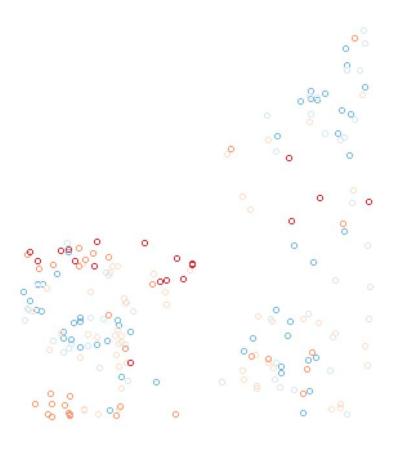
z = [Observed I - Mean(I)] / StandardDeviation(I)

z-values are comparable across variables and across spatial weights









Cleveland 2015 q4 house sales prices (in \$1,000)





	Normal	Randomization
MI	0.282 0.282	
E[MI]	-0.0049 -0.0049	
Var[MI]	0.00178 0.00158	
z-value	6.81 7.22	
p-value	<< 0.000001 << 0.000001	

normal vs randomization inference for Moran's I queen weights





Permutation Approach

as such, even a high Moran's I does not indicate significance

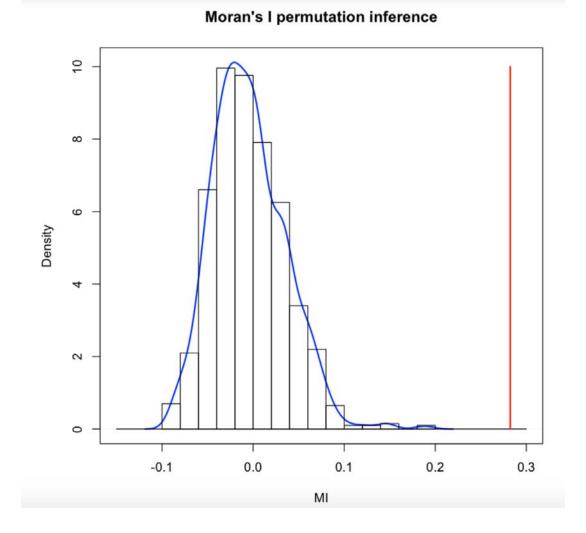
need to construct a reference distribution

randomly reshuffle observations and recompute Moran's I each time

compare observed value to reference distribution







999 permutations and reference distribution (queen weights)

$$MI = 0.282$$
 Mean = -0.0045 s.d. = 0.0401 z-value = 7.15





Interpretation of Moran's I





Sign of Moran's I

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theoretical mean is - I/(N - I), essentially zero for large N
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positive and significant = clustering of like value

NOT clustering of high or low

could be either or a combination

negative and significant = alternating values

presence of spatial outliers

spatial heterogeneity (checkerboard pattern)





Comparing Moran's I

Moran's I depends on spatial weights

relative magnitude for same weights and different variables is meaningful

but NOT for different spatial weights

instead, use standardized z-value to compare





	Queen	K6	Distance
MI	0.282	0.343	0.335
E[MI]	-0.0049	-0.0049	-0.0049
Var[MI]	0.00158	0.00124	0.00110
z-value	7.22	9.88	10.3
p-value	<< 0.000001	<< 0.000001	<< 0.000001

Moran's I, different spatial weights MI is largest for K6, but z is largest for Distance





Clustering vs Clusters

Moran's I is a global statistic, i.e., a single value for the whole spatial pattern

Moran's I does NOT provide the location of clusters

cluster detection requires a local statistic





True vs Apparent Contagion

the indication of clustering does not provide an explanation for why the clustering occurs

different processes can result in the same pattern

true contagion: evidence of clustering due to spatial interaction (peer effects, mimicking, etc.)

apparent contagion: evidence of clustering due to spatial heterogeneity (different spatial structures create local similarity)





Geary's c





Focus on Dissimilarity

squared difference as measure of dissimilarity

similar to notion of variogram (geostatistics)

values between 0 and 2





Geary's c Statistic

$$c = (N-1) \sum_{i} \sum_{j} w_{ij} (x_i - x_j)^2 / 2S_0 \sum_{i} z_i^2$$

alternatively

c =
$$[\Sigma_i \Sigma_j w_{ij} (x_i - x_j)^2 / 2S_0] / [\Sigma_{z_i}^2 / (N-I)]$$

with z_i in deviations from the mean

 $S_0 = \sum_i \sum_j w_{ij}$ sum of all the weights





Interpretation

positive spatial autocorrelation c < l or z < 0

negative spatial autocorrelation c > 1 or z > 0

opposite sign of Moran's I





Inference

same approach as for Moran's I

analytical

approximate (normal, randomization)

permutation





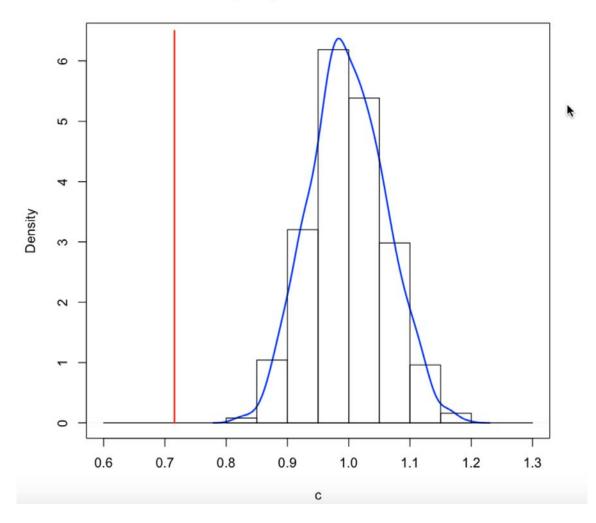
	Queen	K6	Distance
С	0.716	0.496	0.544
E[c]	1.0	1.0	1.0
Var[c]	0.00381	0.00545	0.00285
z-value	-4.61	-6.82	-8.55
p-value	< 0.000004	<< 0.00000 I	<< 0.00000 I

geary's c inference under randomization, different spatial weights





Geary's c permutation inference



999 permutations and reference distribution (queen weights)

$$c = 0.716$$
 Mean = 0.998 s.d. = 0.062 z-value = -4.58





Moran Scatter Plot





Recap: Moran's I

• $I = [\Sigma_i \Sigma_j w_{ij} z_i.z_j/S_0]/[\Sigma_i z_i^2 / N]$

with $z_i = y_i - m_x$: deviations from mean

for row-standardized weights $S_0 = N$

$$I = \sum_{i} \sum_{j} w_{ij} z_{i}.z_{j} / \sum_{i} z_{i}^{2} = \sum_{i} z_{i} (\sum_{j} w_{ij}.z_{j}) / \sum_{i} z_{i}^{2}$$

Moran's I is slope in a regression of Σ_i $w_{ij}.z_j$ on z_i





Moran Scatter Plot

scatter plot of [z_i, Σ_j w_{ij}.z_j]

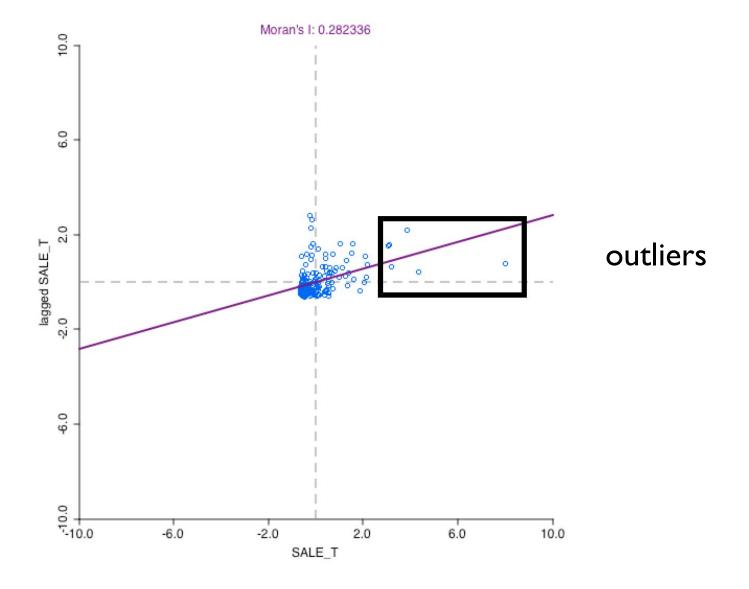
the value at i on the x-axis, its spatial lag (weighted average of neighboring values) on the y-axis

slope of linear fit is Moran's I

use local regression (Lowess) to identify possible structural breaks







Moran scatter plot - Cleveland house prices - queen weights





Categories of Local Spatial Autocorrelation

four quadrants of the scatter plot

upper right and lower left

positive spatial autocorrelation

clusters of like values

locations are similar to their neighbors

lower right and upper left

negative spatial autocorrelation

spatial outliers

locations are different from their neighbors

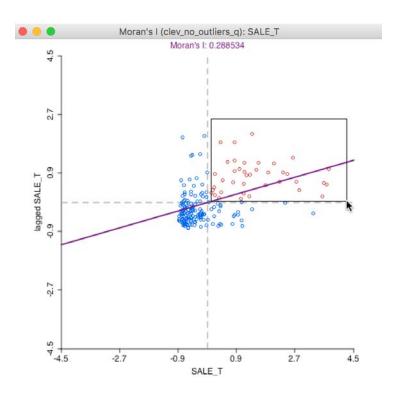




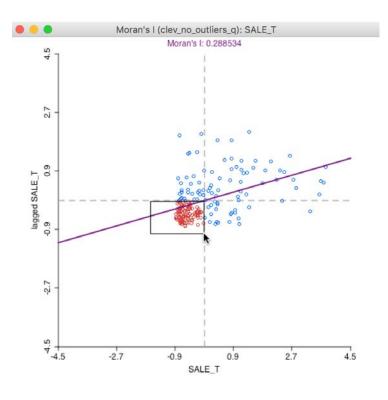
Positive Spatial Autocorrelation

all comparisons relative to the mean

not absolute high or low



high-high

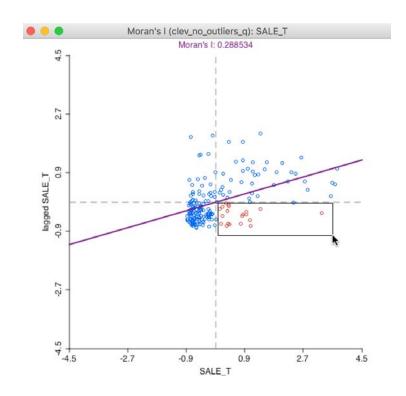


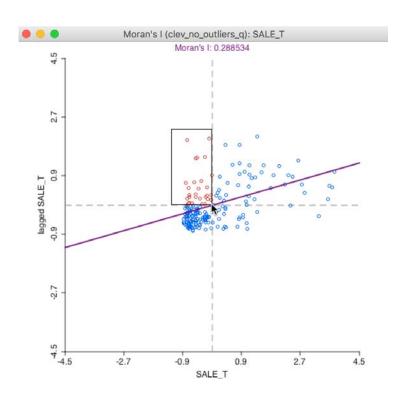
low-low





Negative Spatial Autocorrelation spatial outliers



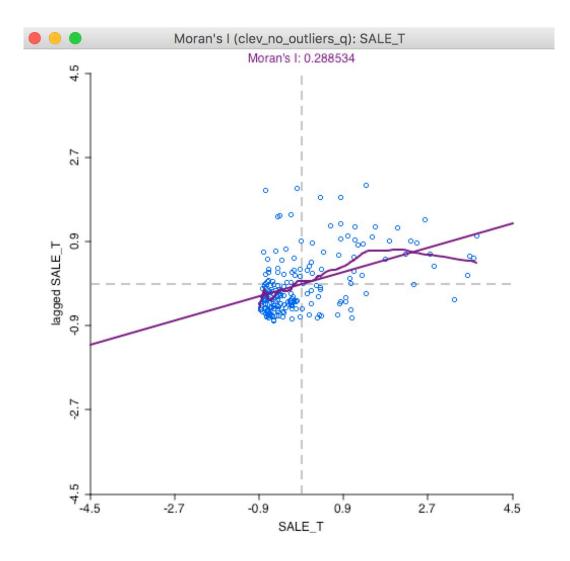


high-low

low-high







Moran scatter plot with Lowess smoother





Correlogram





Alternative Perspective - Nonparametric

non-parametric approach

no model for covariance

let the data suggest the functional form

based on sample autocorrelation

covariance function must be positive definite





Sample Autocorrelation

computed for each pair i, j

$$\rho_{ij} = \rho(z_i, z_j) = z_i z_j^* / (I/n) \Sigma_h (z_h - z_m)^2$$

 $z_i^* = z_i - z_m$ deviations from mean

in practice, easier to use standardized zi

incidental parameter problem

one parameter for each pair i-j

n(n-1)/2 individual values of ρ_{ij}





Non-Parametric Principle

spatial autocorrelation as an unspecified function of distance

$$\rho_{ij} = g(d_{ij})$$

how to fit the function? use kernel estimator





Kernel Regression

$$z_i z_j = g(d_{ij})$$

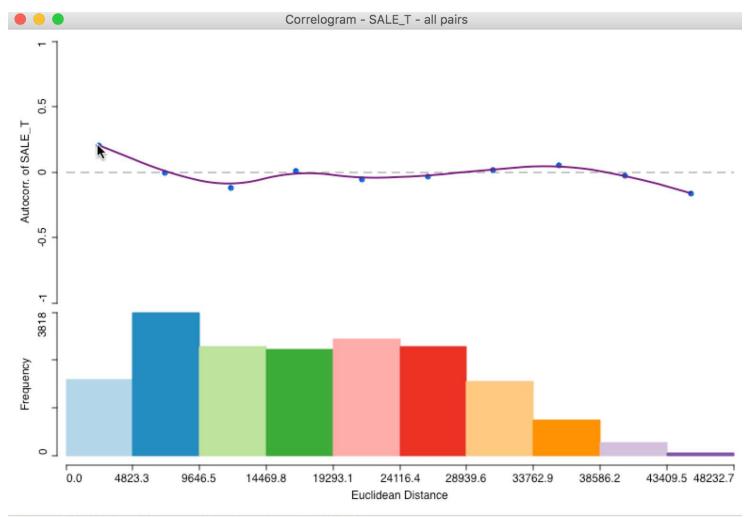
local regression

depends on choice of kernel function and bandwidth

values of the estimated $g(d_{ij})$ do not necessarily result in a valid (positive semi-definite) variance-covariance matrix





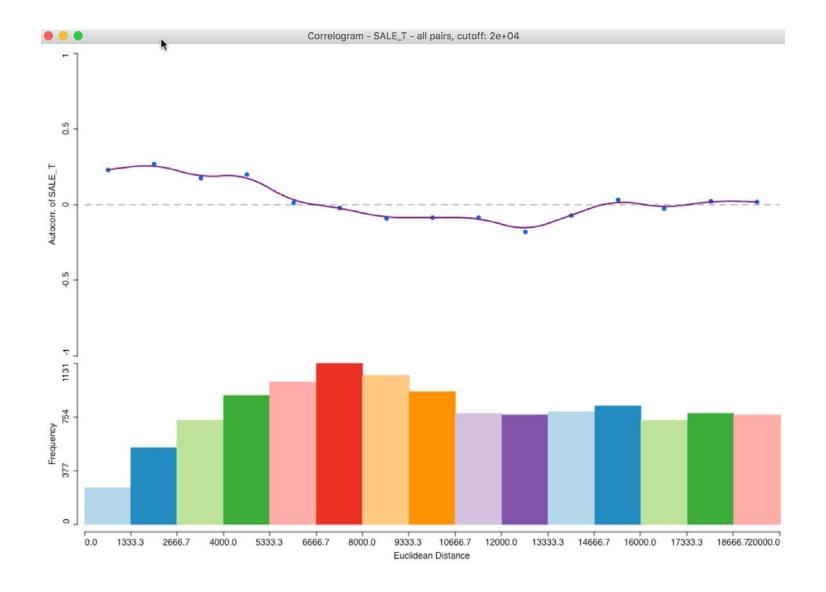


autocorrelation is 0.202048 for obs within distance band 0 to 4823.27

correlogram - full distance range







correlogram - 20,000 ft max distance





Interpretation

range of spatial autocorrelation

alternative to specifying spatial weights

sensitive to kernel fit

may violate Tobler's law





Variogram





Semi-Variogram





Variogram Function

magnitude of the variance of the difference as a function of displacement (h)

$$2\gamma(h) = Var [Z_{s+h} - Z_s]$$

factor 2 is by convention, so half of this function is the semi-variogram

$$\gamma(h) = (1/2) \text{ Var } [Z_{s+h} - Z_s]$$





Operational Semi-Variogram

constant mean assumption

$$E[Z_{s+h} - Z_s] = E[Z_{s+h}] - E[Z_s] = 0$$

such that
$$Var[Z_{s+h} - Z_s] = E[Z_{s+h} - Z_s]^2 - 0$$

semi-variogram

$$\gamma(h) = (1/2) E [Z_{s+h} - Z_s]^2$$

estimate as average of the squared differences





Variogram Cloud Plot

plot of all squared differences against the distance separating the pair of observations

exploring the variogram cloud plot

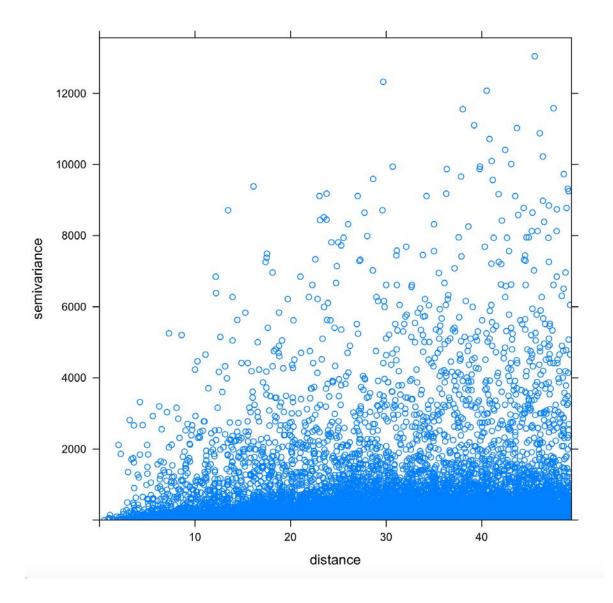
identify outliers

large difference for small distance

negative spatial autocorrelation



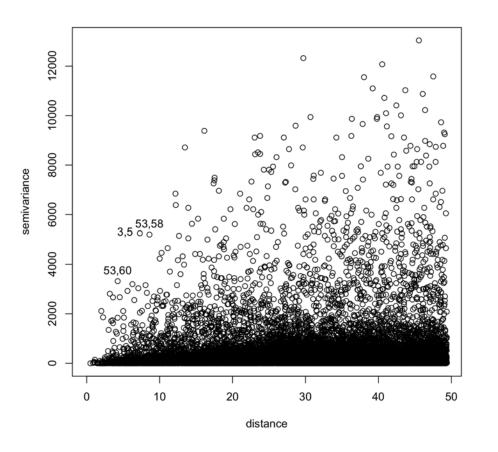




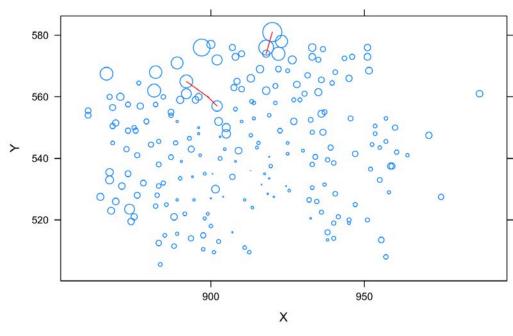
Variogram Cloud Plot Baltimore House Prices







selected point pairs



Checking for outliers





The Empirical Variogram





Estimating the Empirical Variogram

method of moments

bin the pairs by distance bands

average squared difference within bin

• $2\gamma(h) = (1 / |N(h)|) \Sigma_h [Z_{s+h} - Z_s]^2$

N(h) number of pairs in distance bin h





Practical Issues

maximum distance = "distance of reliability"

different rules of thumb

h < D/2 (D is maximum distance)

h < d/3 (d is diagonal of bounding box)

bin width

avoid sparse bins

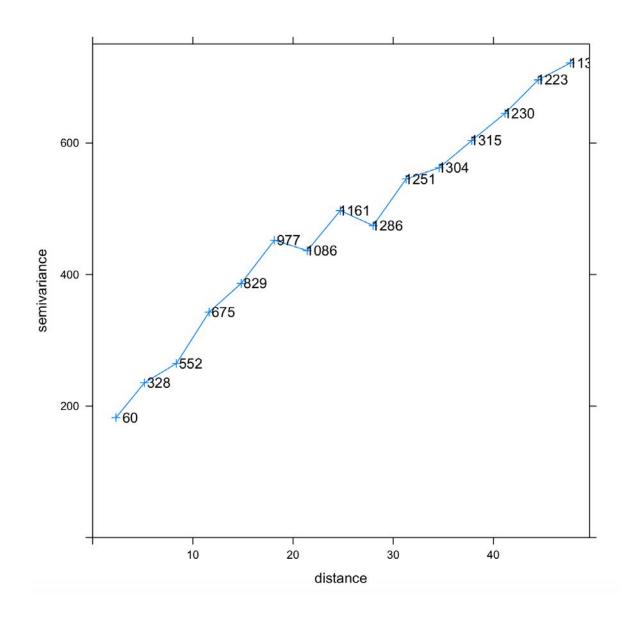
at least 30 pairs

precision of estimate will depend on number of pairs in the bin





```
dist
                     gamma
     np
         2.325018 182.4068
    328
         5.171430 235.3985
3
    552
         8.358653 264.7059
    675 11.626348 342.9250
    829 14.824826 386.5458
    977 18.136492 452.1512
   1086 21.419961 436.3443
   1161 24.718931 497.1898
   1286 27.991568 474.3808
10 1251 31.285467 545.2776
  1304 34.588754 562.3477
  1315 37.838726 603.7175
  1230 41.159131 645.1485
  1223 44.475928 696.0524
15 1138 47.681960 721.7097
```



Empirical Variogram Baltimore house prices





Issues

variogram should increase with distance, but level off at some point

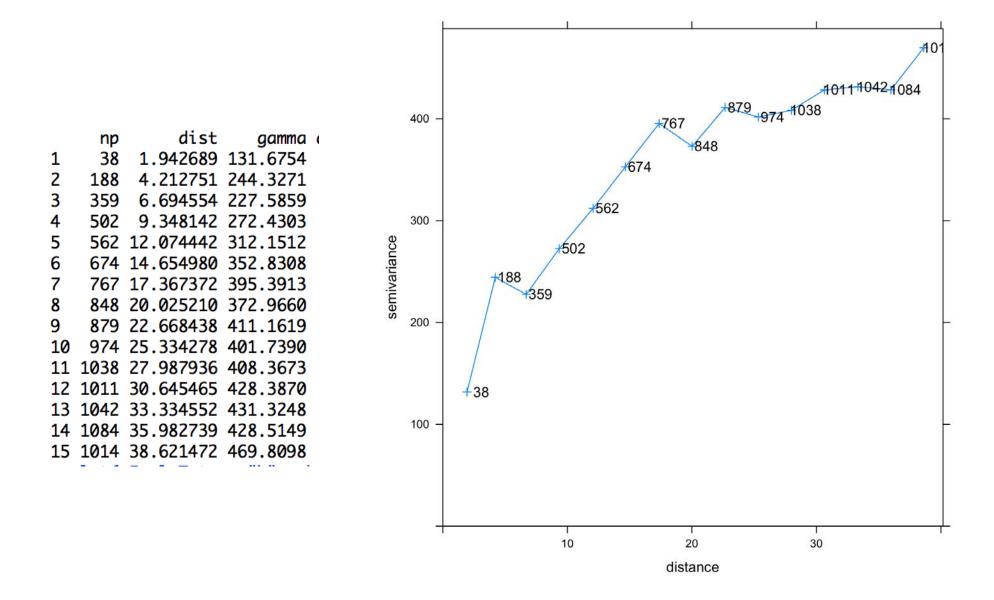
point beyond it levels off = no more spatial autocorrelation

when variogram keeps increasing = evidence of a trend in the data

solution = use regression residuals







Empirical Variogram Baltimore second order trend surface residuals





Adjusting the Empirical Variogram

adjust the cut-off distance

shorter distances are where the interest is

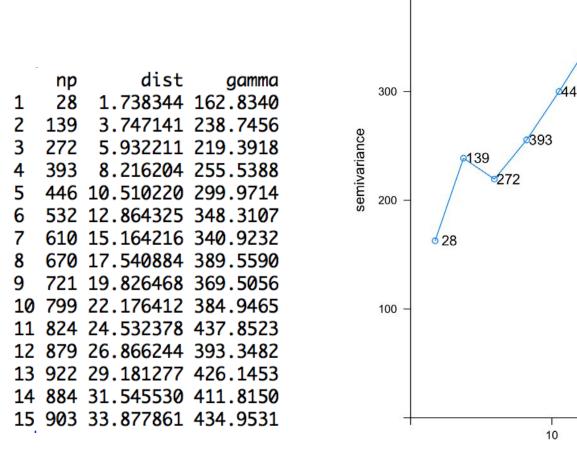
larger distances

tend not to be spatially correlated

fewer pairs to estimate variogram from







Empirical Variogram Baltimore trend surface residuals - cut off d=35





Directional Effects





Anisotropy - Directional Variogram

fundamental assumption = isotropy

only distance matters, direction doesn't matter

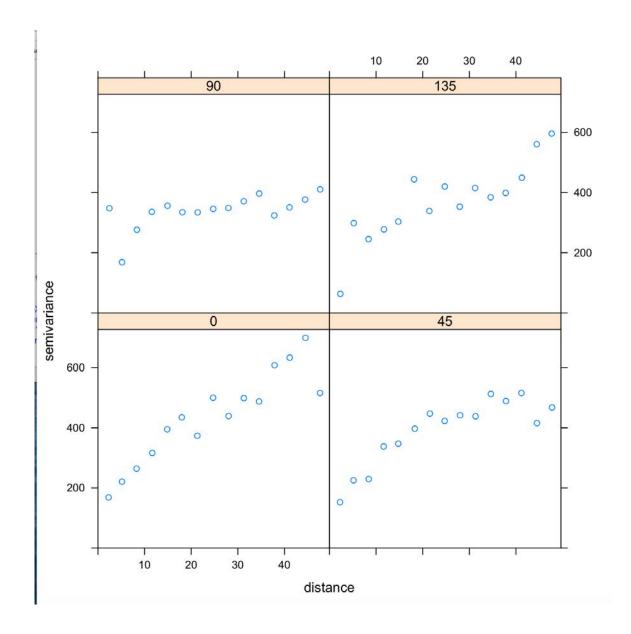
directional semi-variogram

different variogram for point pairs in angular sections

typically implemented by 45 degree sections







Directional variogram





Anisotropy - Variogram Map

alternative visualization of semi-variogram values centered on 0,0

rectangles for distance in E-W direction (dx) and N-S direction (dy)

semi-variogram computed for pairs in rectangle

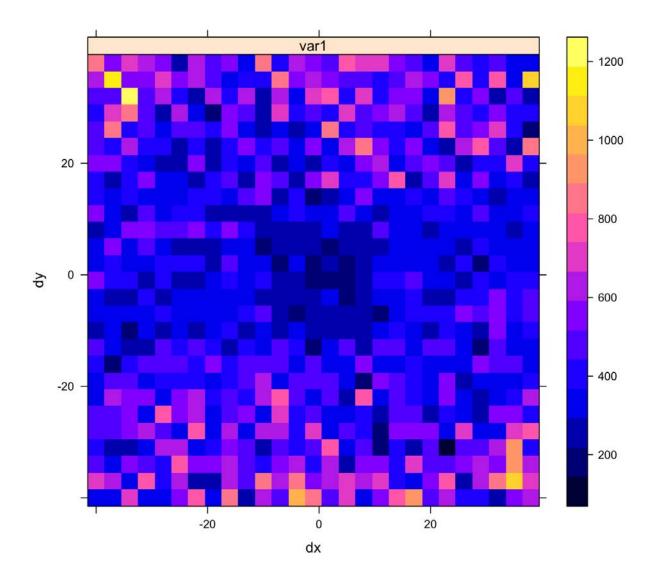
variogram map should be concentric

deviations point to anisotropy

eliminate cells with fewer than "n" pairs



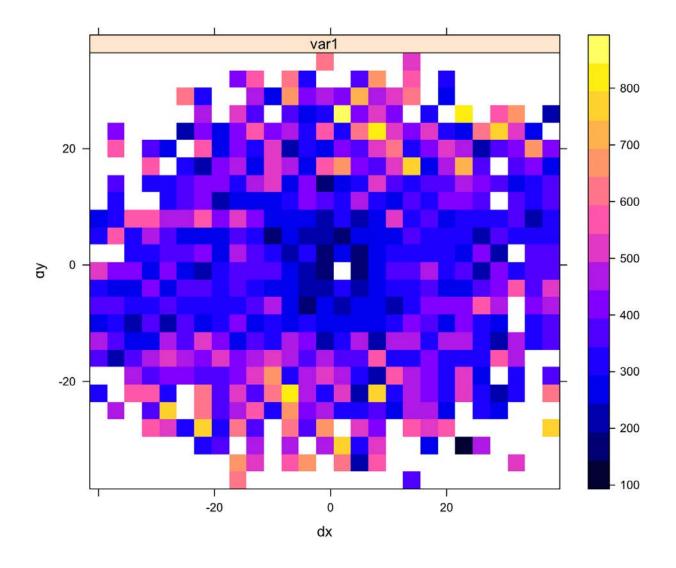




Variogram map - all pairs







Variogram map with more than 30 pairs





Variogram Models





Terminology

sill = process variance

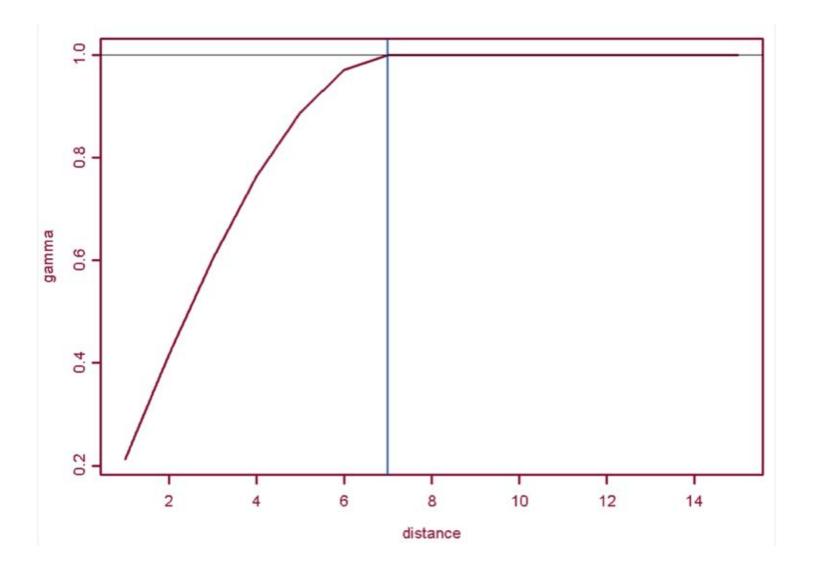
range = distance beyond which there is no more spatial autocorrelation

nugget = positive variance at distance 0 (variance at 0 should be 0)









Range = 7





Valid Variogram Models

need to ensure that variance-covariance matrix is positive definite

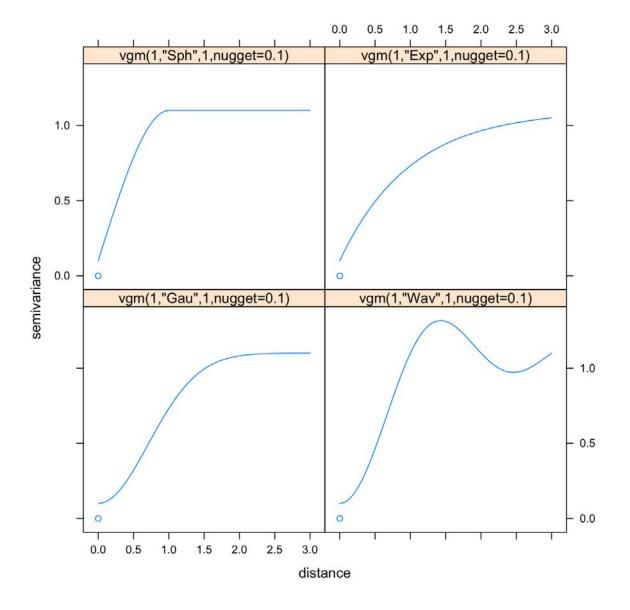
positive definite C(h) for all h

negative definite $\gamma(h)$ for all h

leads to a collection of valid parametric models (with limits on the parameter space)



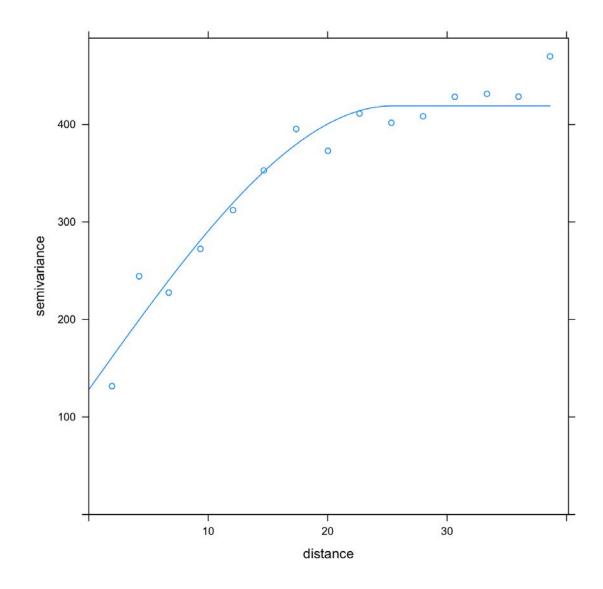




spherical, exponential, Gaussian and wave variogram models







best fit spherical model nugget = 128.04, partial sill = 291.01, range = 25.46 SSErr = 36852.6



