6 Sampling Distributions, Central Limit Theorem, and Estimation

Chapter Preview

This set of notes will examine statistics and their distributions along with the Central Limit Theorem. The Central Limit Theorem is arguably the most important theorem in Statistics and will bridge the gap between our study of Probability and Statistics. We will also examine estimators and their properties.

6.1: Statistics and Their Distributions

<u>Def:</u> A <u>statistic</u> is any quantity whose value can be calculated from sample data.

<u>Def:</u> The random variables X_1, X_2, \ldots, X_n are said to form a <u>random sample</u> of size n if X_i are independent and identically distributed (IID).

In chapter 1, we came across a few statistics. We can measure the center of a sample with the sample mean, \bar{x} and the sample median, \tilde{x} . If we want to describe the spread of a sample, there are the sample variance (s^2) and sample standard deviation (s).

In fact, since each random sample could potentially be different, sample statistics have a probability distribution associated with them.

Example: Let's simulate five exponential $(\lambda = 3)$ random variables a couple times and look at the different sample means.

```
sample1 = rexp(n=5,rate=3); sample1
## [1] 0.5274761 0.2818285 0.1090117 0.4920306 0.2566445

mean(sample1)
## [1] 0.3333983

sample2 = rexp(n=5,rate=3); sample2
## [1] 0.07052601 1.36882351 0.14881661 0.04802431 0.28632793

mean(sample2)
## [1] 0.3845037

sample3 = rexp(n=5,rate=3); sample3
## [1] 0.58110200 0.09693622 0.45714040 0.39613499 0.16153774

mean(sample3)
## [1] 0.3385703
```

Example: Let's repeat the previous experiment 1000 times and plot the results.

```
# initialize a vector of length 1000 for the sample means
means = rep(0, 1000)

# collect a sample of five RVs, calculate sample mean, and repeat
for (i in 1:1000) {
    randomsample = rexp(n = 5, rate = 3)
        means[i] = mean(randomsample)
}
mean(means)

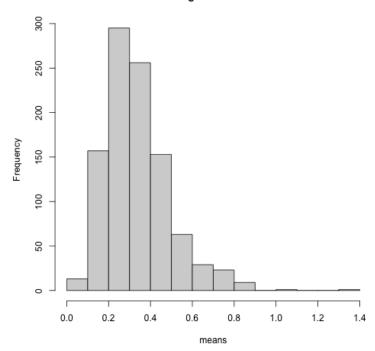
## [1] 0.3380338

var(means)

## [1] 0.02367919

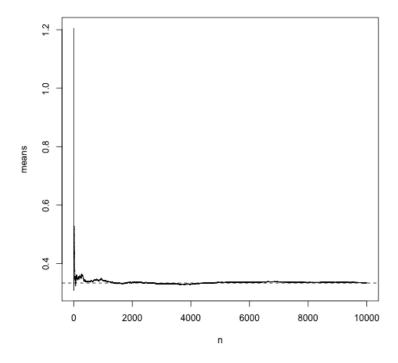
hist(means)
```

Histogram of means



Example: Now, let's look at the sample mean of a random sample as a function of sample size, n.

```
# generate a sample of 10000 exp(3) RVs
data = rexp(n = 10000, rate = 3)
means = cumsum(data)/seq_along(data)
plot(x = 1:10000, y = means, type = "1", xlab = "n")
abline(h = 1/3, lty = 2)
```



Theorem: The tendency of \bar{x} to approach μ as $n \to \infty$ is the result of the Law of Large Numbers. Specifically, if $|E[X]| = \mu < \infty$, then $\lim_{n \to \infty} 1/n \sum_{n=1}^{\infty} x_i = \lim_{n \to \infty} \bar{x} = \mu$.

6.2: The Distribution of the Sample Mean

Motivation: Suppose it is claimed that the average lifespan of Galapagos tortoises is 100 years old. To test this claim, a study of 20 Galapagos tortoises finds an average life span of $\bar{x} = 95.9$ years. Is this enough information to reject the claim?

To answer this question, we need to know the distribution of \bar{X} .

Theorem: Let X_1, X_2, \ldots, X_n be a random sample from a distribution with mean value μ and standard deviation σ and let $\bar{X} = \frac{1}{n} \sum X_i$. Then:

- $\begin{array}{l} \text{(i) } E[\bar{X}] = \mu \\ \text{(ii) } Var(\bar{X}) = \sigma_{\bar{X}}^2 = \frac{\sigma^2}{n} \\ \text{(iii) } \sigma_{\bar{X}} = \frac{\sigma}{\sqrt{n}} \end{array}$

Example: Show that $Var(\bar{X}) = \frac{\sigma^2}{n}$.

<u>Def:</u> If $\hat{\sigma}$ is an estimator for σ , then the <u>standard error</u> of $\hat{\sigma}$ is: $SE(\hat{\sigma}) = \sqrt{Var(\hat{\sigma})}$.

Example: If X_1, X_2, \ldots, X_{50} are iid Uniform[10,20] random variables, then determine $E[\bar{X}]$ and $SE(\bar{X})$.

<u>Theorem:</u> (Central Limit Theorem) Let X_1, X_2, \ldots, X_n be a random sample from a distribution with mean μ and variance σ^2 . If n is sufficiently large, \bar{X} has approximately a normal distribution with mean $\mu_{\bar{X}} = \mu$ and $\sigma_{\bar{X}}^2 = \sigma^2/n$.

This result is nontrivial and extremely useful! It allows us to determine how likely or unlikely a sample mean is if we are given a hypothesized population mean.

Illustration:

Example: Suppose that X_1, X_2, \dots, X_{10} is a random sample from a discrete uniform (1,6) random variable. Determine the approximate sampling distribution of \bar{X} .

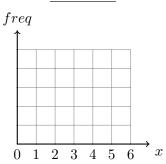
 $\underline{\text{Example:}}$ Class exercise. We will collect data based on the following experiments. Plot a histogram for each instance.

(a) Roll your die once and record your number: _____

Sample data

X	1	2	3	4	5	6
freq						

Histogram

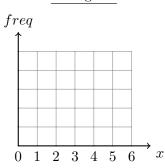


(b) Roll your die twice and record the average of the two die rolls:

Sample data

X	(0.5, 1]	(1,1.5]	(1.5,2]	(2,2.5]	(2.5,3]	(3,3.5]
freq						
x	(3.5, 4]	(4,4.5]	(4.5,5]	(5,5.5]	(5.5,6]	
freq						

Histogram

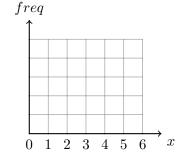


(c) Roll your die ten times and record the average of the ten die rolls:

Sample data

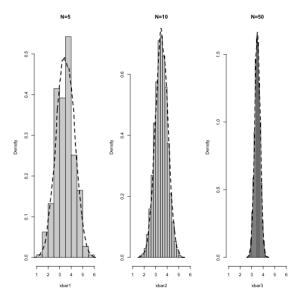
X	(0.5, 1]	(1,1.5]	(1.5,2]	(2,2.5]	(2.5,3]	(3,3.5]
freq						
Х	(3.5, 4]	(4,4.5]	(4.5,5]	(5,5.5]	(5.5,6]	
freq						

Histogram

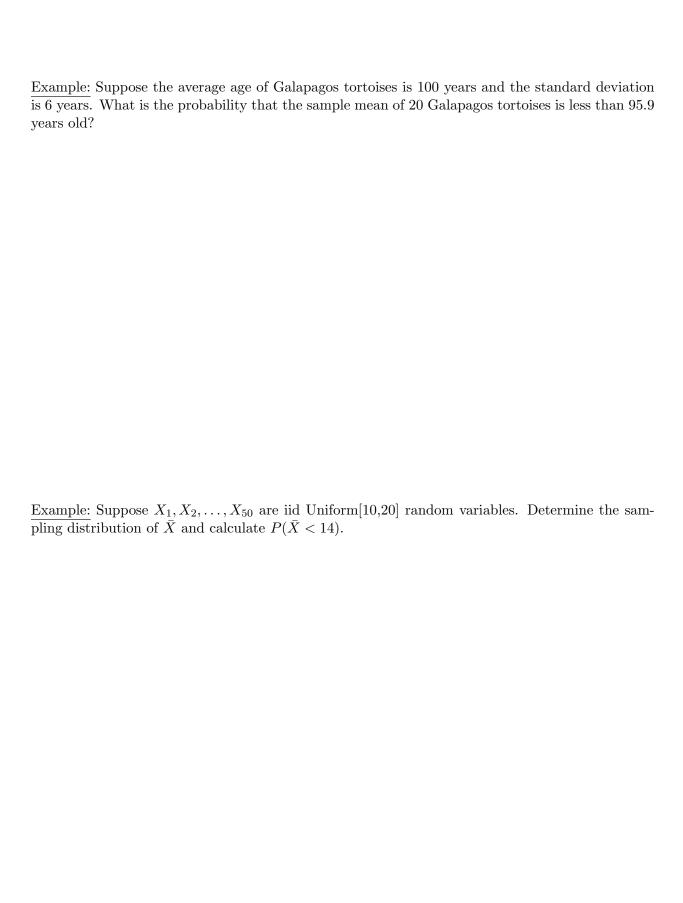


Example: Let's simulate N fair die rolls 10,000 times. Here we will do the simulation for N=5, $\overline{N=10}$, and N=50.

```
xbar1 = rep(0, 10000) # initialize this vector to zeros
xbar2 = rep(0, 10000) # initialize this vector to zeros
xbar3 = rep(0, 10000) # initialize this vector to zeros
\# 10,000 simulated die rolls for N=5,10,50
for (i in 1:10000) {
    x1 = sample(1:6, 5, replace = T)
    xbar1[i] = mean(x1)
    x2 = sample(1:6, 10, replace = T)
    xbar2[i] = mean(x2)
    x3 = sample(1:6, 50, replace = T)
    xbar3[i] = mean(x3)
par(mfrow = c(1, 3))
hist(xbar1, main = "N=5", xlim = c(1, 6), prob = TRUE)
lines(density(xbar1), lwd = 2, lty = 2)
hist(xbar2, main = "N=10", xlim = c(1, 6), prob = TRUE)
lines(density(xbar2), lwd = 2, lty = 2)
hist(xbar3, main = "N=50", xlim = c(1, 6), prob = TRUE)
lines(density(xbar3), lwd = 2, lty = 2)
```



Notice that the variance decreases when N increases. Also notice that the histograms appear to look more "normal" as N increases.



6.3: Several General Concepts of Point Estimation

<u>Def:</u> A <u>point estimate</u> of a parameter θ is a single number that can be regarded as a sensible value for θ and is often denoted with a hat (i.e., $\hat{\theta}$). A point estimate is obtained by selecting a suitable statistic and computing its value from the given sample data. The selected statistic, $\hat{\theta}$ is called the point estimator of θ .

Example:

<u>Def:</u> The bias of a point estimator $\hat{\theta}$ for a parameter θ is defined as $Bias[\hat{\theta}] = E[\hat{\theta}] - \theta$.

<u>Def:</u> A statistic $\hat{\theta}$ is called an <u>unbiased estimator</u> of the parameter θ if $E[\hat{\theta}] = \theta$.

<u>Note</u>: \bar{X} is an unbiased estimator for μ since $E[\bar{X}] = \mu$

<u>Note:</u> From the theorem above, note that as $n \to \infty$, $Var(\bar{X}) \to 0$. This means that we can estimate μ as precisely as we would like as long as we make n large.

Example: Suppose X_1, X_2, \ldots, X_n are exponential $(\lambda = 2)$ random variables. Is \bar{X} an unbiased estimator for λ ? How about if X_1, X_2, \ldots, X_n are Poisson $(\lambda = 2)$ random variables?

<u>Def:</u> The mean squared error (MSE) of an estimator $\hat{\theta}$ of a parameter θ is defined by $E[(\hat{\theta} - \theta)^2]$.

Typically, unbiased estimators are preferred though some advanced methods consider biased estimators that minimize MSE.

Theorem: Suppose X_1, X_2, \ldots, X_n is a random sample from a distribution with mean μ and variance σ^2 . Then:

- (i) $\hat{\theta} = \bar{X}$ is an unbiased estimator for μ . (ii) $\hat{\sigma}^2 = S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i \bar{X})^2$ is an unbiased estimator for σ^2 .

Example: Suppose X_1, X_2, \dots, X_n are a random sample of Binomial(n,p) random variables. Derive $E[\bar{X}]$ using the Binomial distribution and compare it to the result from the theorem above. Also, find Var(X).

<u>Def:</u> The <u>standard error</u> of an estimator $\hat{\theta}$ is its standard deviation $\sigma_{\hat{\theta}} = \sqrt{Var(\hat{\theta})}$. It is the magnitude of a typical or representative deviation between an estimate and the value of θ . If the standard error is a function of unknown parameters, then we can estimate it with the estimated standard error which is denoted $s_{\hat{\theta}}$.

Example: Find the standard error for \bar{X} in the example above.

6.4: Methods of Point Estimation

There are a variety of ways to obtain a point estimate. These methods include:

- (i) Method of Moments
- (ii) Maximum Likelihood Estimation
- (iii) Bayesian Estimation

While we won't cover how these estimation methods are completed, we will compare different estimators. Each of these methods could produce a different point estimate.

Example: Suppose X_1, X_2, X_3 are a random sample of $\mathcal{N}(\mu = 10, \sigma^2 = 4)$. Four estimators are given below. Determine if each estimator is unbiased and calculate its standard error. Which of the four estimators is preferred?

(a)
$$Y_1 = \frac{X_1 + X_2 + X_3}{3}$$

(b)
$$Y_2 = \frac{X_1}{2} + \frac{X_2}{4} + \frac{X_3}{4}$$

(c)
$$Y_3 = X_1$$

(d)
$$Y_4 = \frac{X_1 + X_2 + X_3}{2}$$