

# On the Construction of Pseudo-Random Permutations: Luby-Rackoff Revisited

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## Abstract

Luby and Rackoff [27] showed a method for constructing a pseudo-random permutation from a pseudo-random function. The method is based on composing four (or three for weakened security) so called Feistel permutations, each of which requires the evaluation of a pseudo-random function. We reduce somewhat the complexity of the construction and simplify its proof of security by showing that two Feistel permutations are sufficient together with initial and final pair-wise independent permutations. The revised construction and proof provide a framework in which similar constructions may be brought up and their security can be easily proved. We demonstrate this by presenting some additional adjustments of the construction that achieve the following:

- Reduce the success probability of the adversary.
- Provide a construction of pseudo-random permutations with *large* input size using pseudo-random functions with *small* input size.

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# 1 Introduction

Pseudo-random permutations, which were introduced by Luby and Rackoff [27], formalize the well established cryptographic notion of block ciphers. Block ciphers are private-key encryption schemes such that the encryption of every plaintext-block is a single ciphertext-block *of the same length*. Therefore we can think of the private key as determining a permutation on strings of the length of the block. A highly influential example of a block cipher is the Data Encryption Standard (DES) [33].

The advantage of block ciphers (compared to using pseudo-random functions for private-key encryption) is that the plaintext and ciphertext are of the same length. This property saves on memory and prevents wasting communication bandwidth. Furthermore, it enables the easy incorporation of the encryption scheme into existing protocols or hardware components.

Luby and Rackoff defined the security of pseudo-random permutations in analogy to the different attacks considered in the context of block ciphers:

- Pseudo-random permutations can be interpreted as block ciphers that are secure against an adaptive *chosen-plaintext attack*. Informally, this means that an (efficient) adversary, with access to the encryptions of messages of its choice, cannot tell apart those encryptions from the values of a truly random permutation.
- Strong pseudo-random permutations can be interpreted as block ciphers that are secure against an adaptive *chosen plaintext and ciphertext attack*. Here, the adversary has the additional power to ask for the decryption of ciphertexts of its choice.

Pseudo-random permutations are closely related (both in definition and in their construction) to the earlier concept of pseudo-random functions which was defined by Goldreich, Goldwasser and Micali [17]. These are efficiently samplable and computable functions that are indistinguishable from random functions under all (efficient) black-box attacks (see Section 2 for a formal definition). Pseudo-random functions play a major role in private-key cryptography and have many additional applications (for some of these applications, see [11, 18, 26]).

Luby and Rackoff [27] provided a construction of strong pseudo-random permutations, (**LR-Construction**) which was motivated by the structure of DES. The basic building block is the so called Feistel permutation<sup>1</sup> based on a pseudo-random function defined by the key. Their construction consists of four rounds of Feistel permutations (or three rounds, for pseudo-random permutations) each round involves an application of a (different) pseudo-random function (see Figure 1.a for an illustration). The LR-Construction's main source of attraction is, most probably, its elegance.

Goldreich, Goldwasser and Micali [17] showed a construction of pseudo-random functions from pseudo-random generators [10, 51]. Thus, the construction of pseudo-random permutations reduces to the construction of pseudo-random generators. Recently a different construction of pseudo-random functions was introduced by Naor and Reingold [32]; this is a parallel construction based on a new primitive called a pseudo-random *synthesizer* that in particular can be constructed from any trapdoor permutation. This implies a parallel construction of pseudo-random permutations. Nevertheless, all known constructions of pseudo-random functions involve non-trivial (though of course polynomial time) computation, so it makes sense to attempt to minimize the number of invocations of pseudo-random functions.

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<sup>1</sup>A Feistel permutation for a function  $f : \{0, 1\}^n \mapsto \{0, 1\}^n$  is a permutation on  $\{0, 1\}^{2n}$  defined by  $D_f(L, R) \stackrel{\text{def}}{=} (R, L \oplus f(R))$ , where  $|L| = |R| = n$ . Each of the 16 rounds of DES involves a Feistel permutation of a function determined by the 56 key bits.

Alongside cryptographic pseudo-randomness the last two decades saw the development of the notion of limited independence in various setting and formulations [3, 4, 13, 14, 25, 31, 50]. For a family of functions  $\mathcal{F}$  to have some sort of (limited) independence means that if we consider the value of a function  $f$ , chosen uniformly at random from  $\mathcal{F}$ , at each point as a random variable (in the probability space defined by choosing  $f$ ) then these random variables possess the promised independence property. Thus, a family of permutations on  $\{0,1\}^n$  is pair-wise independent if for all  $x \neq y$  the values of  $f(x)$  and  $f(y)$  are uniformly distributed over strings  $(a,b) \in \{0,1\}^{2n}$  such that  $a \neq b$ . Functions of limited independence are typically much simpler to construct and easier to compute than (cryptographic) pseudo-random functions.

## 1.1 New Results and Organization

The goal of this paper is to provide a better understanding of the LR-Construction and as a result improve the construction in several respects. Our main observation is that the different rounds of the LR-Construction serve significantly different roles. We show that the first and last rounds can be replaced by pair-wise independent permutations and use this in order to :

1. Achieve an improvement in the computational complexity of the pseudo-random permutations – two applications of a pseudo-random function on  $n$  bits suffice for computing the value of a pseudo-random permutation on  $2n$  bits at a given point (vs. four applications in the original LR-Construction).
2. Simplify the proof of security of the construction (especially in the case of strong pseudo-random permutations) and provide a framework for proving the security of similar constructions.
3. Derive generalizations of the construction that are of practical and theoretical interest. The proof of security for each one of the constructions is practically “free of charge” given the proof of security of the main construction.

As discussed in Section 5.2, the new construction is in fact a generalization of the original LR-Construction. Thus, the proof of security (Theorem 3.2) also applies to the original construction. The following is a brief and informal description of the paper’s main results and organization:

**Section 2** Reviews the notations and definitions regarding pseudo-randomness and  $k$ -wise independence.

**Section 3** Presents the main construction and proves its security: pair-wise independent permutations can replace the first and fourth rounds of the LR-Construction (see Figure 1.b for an illustration).

**Section 4** Highlights the high-level structure of the proof of security which provides a framework that enables us to relax and generalize the main construction.

**Section 5** Shows how the main construction can be relaxed by:

- 5.1** Using a single pseudo-random function (instead of two) and
- 5.2** Using weaker and more efficient permutations (or functions) instead of the pair-wise independent permutations.

**Section 6** Provides a simple generalization of the main construction: using  $t$  rounds of (generalized) Feistel permutations (instead of two) the success probability of the distinguisher is reduced from approximately  $\frac{m^2}{2^{l/2}}$  to approximately  $\frac{t}{2} \cdot \frac{m^2}{2^{(1-1/t)l}}$ , where the permutation is on  $l$  bits and the distinguisher makes at most  $m$  queries (see Figure 3 for an illustration).

**Section 7** Provides a second generalization of the main construction. Instead of applying Feistel permutations on the entire outputs of the first and second rounds, Feistel permutations can be separately applied on each one of their sub-blocks. This is a construction of a strong pseudo-random permutation on *large* blocks using pseudo-random functions on *small* blocks (see Figure 4 for an illustration).

**Section 8** Analyzes the different constructions of the paper as constructions of  $k$ -wise  $\delta$ -dependent permutations.

**Section 9** Suggests directions for further research.

## 1.2 Related Work

The LR-Construction inspired a considerable amount of research. We try to refer to the more relevant (to this paper) part of these directions.

Several alternative proofs of the LR-Construction were presented over the years. Maurer [29] gives a proof of the three-round construction. His proof concentrates on the non-adaptive case, i.e., when the distinguisher has to specify all its queries in advance. A point worth noticing is that indistinguishability under non-adaptive attacks does not necessarily imply indistinguishability under adaptive attacks. For example, a random involution (an involution is a permutation which is the inverse of itself) and a random permutation are indistinguishable under non-adaptive attacks and can be distinguished using a very simple adaptive attack.<sup>2</sup> A different approach toward the proof was described by Patarin [35] (this is the only published proof, we are aware of, for the LR-Construction of *strong* pseudo-random permutations; another proof was given by Koren [23]).

Other papers consider the security of possible variants of the construction. A significant portion of this research deals with the construction of pseudo-random permutations and strong pseudo-random permutations from a *single* pseudo-random function. This line of work is described in Section 5.1.

Lucks [28] shows that a hash function can replace the pseudo-random function in the first round of the three-round LR-Construction. His proof is based on [29] and is motivated by his suggestion to use the LR-Construction when the input is divided into two *unequal* parts. Lucks left open the question of the construction of strong pseudo-random permutations.

Somewhat different questions were considered by Even and Mansour [15] and by Kilian and Rogaway [22]. Loosely speaking, the former construct several pseudo-random permutations from a single one, while the latter show how to make exhaustive key-search attacks more difficult. The construction itself amounts, in both cases, to XORing the input of the pseudo-random permutation with a random key and XORing the output of the permutation with a second random key.

The background and related work concerning other relevant issues are discussed in the appropriate sections: Definitions and constructions of efficient hash functions in Section 5.2, reducing the distinguishing probability in Section 6 and the construction of pseudo-random permutations (or functions) with large input size from pseudo-random permutations (or functions) with small input size in Section 7.

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<sup>2</sup>An even more striking example is obtained by comparing a random permutation  $P$  that satisfies  $P(P(0)) = 0$  with a truly random permutation.

## 2 Preliminaries

In this section, the concepts of pseudo-random functions and pseudo-random permutations are briefly reviewed. A more thorough and formal treatment can be found in [16, 26]. In addition, some basic notations and definitions are introduced.

### 2.1 Notations

- $I_n$  denotes the set of all  $n$ -bit strings,  $\{0, 1\}^n$ .
- $F_n$  denotes the set of all  $I_n \mapsto I_n$  functions and  $P_n$  denotes the set of all such permutations ( $P_n \subset F_n$ ).
- Let  $x$  and  $y$  be two bit strings of equal length, then  $x \oplus y$  denotes their bit-by-bit exclusive-or.
- For any  $f, g \in F_n$  denote by  $f \circ g$  their composition (i.e.,  $f \circ g(x) = f(g(x))$ ).
- For  $x \in I_{2n}$ , denote by  $x|_L$  the first (left)  $n$  bits of  $x$  and by  $x|_R$  the last (right)  $n$  bits of  $x$ .

**Definition 2.1 (Feistel Permutations)** For any function  $f \in F_n$ , let  $D_f \in P_{2n}$  be the permutation defined by  $D_f(L, R) \stackrel{\text{def}}{=} (R, L \oplus f(R))$ , where  $|L| = |R| = n$ .<sup>3</sup>

Notice that Feistel permutations are as easy to invert as they are to compute (since the inverse permutation satisfies  $D_f^{-1}(L, R) = (R \oplus f(L), L)$ ; that is,  $D_f^{-1}(L, R) \equiv \rho \circ D_f \circ \rho$  for  $\rho(L, R) \stackrel{\text{def}}{=} (R, L)$ ). Therefore, the LR-Construction (and its different variants which are introduced in Sections 6 & 7) are easy to invert.

### 2.2 Pseudo-Randomness

Pseudo-randomness is fundamental to cryptography and, indeed, essential in order to perform such tasks as encryption, authentication and identification. Loosely speaking, pseudo-random distributions cannot be efficiently distinguished from the truly random distributions (usually, random here means uniform). However, the pseudo-random distributions have substantially smaller entropy than the truly random distributions and are efficiently samplable.

#### 2.2.1 Overview of Pseudo-Random Primitives

In the case of **pseudo-random (bit) generators**, which were introduced by Blum and Micali and Yao [10, 51], the pseudo-random distribution is of bit-sequences. The distribution is efficiently sampled using a, relatively small, truly random bit-sequence (the seed). Hastad, Impagliazzo, Levin and Luby [21] showed how to construct a pseudo-random generator from any one-way function (informally, a function is one-way if it is easy to compute its value but hard to invert it).

**Pseudo-random function ensembles (PFE)**, which were introduced by Goldreich, Goldwasser and Micali [17], are distributions of functions. These distributions are indistinguishable from the uniform distribution under all (polynomially-bounded) black-box attacks (i.e. the distinguisher can only access the function by specifying inputs and getting the value of the function on these inputs). Goldreich, Goldwasser and Micali provided a construction of such functions based on the existence of pseudo-random generators.

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<sup>3</sup> $D$  stands for DES-like, another common term for these permutations.

Luby and Rackoff [27] define **pseudo-random permutation ensembles (PPE)** to be distributions of permutations that are indistinguishable from the uniform distribution to an efficient observer (that, again, has access to the value of the permutation at points of its choice). In addition, they consider a stronger notion of pseudo-randomness which they call *super pseudo-random permutation generators*. Here the distinguisher can also access the inverse permutation at points of its choice. Following [16] we use the term **strong pseudo-random permutation ensembles (SPPE)** instead.

Luby and Rackoff provided a simple construction of PPE and SPPE (*LR-Construction*) which is the focus of this work. Their construction is based on a basic compound of the structure of DES [33], namely, the compositions of several Feistel-permutations. Their definition of the PPE (resp. SPPE) is  $D_{f_3} \circ D_{f_2} \circ D_{f_1}$  (resp.  $D_{f_4} \circ D_{f_3} \circ D_{f_2} \circ D_{f_1}$ ) where all  $f_i$ s are independent pseudo-random functions and  $D_{f_i}$  as in Definition 2.1 (see Figure 1.a for an illustration).

### 2.2.2 Definitions

A *function ensemble* is a sequence  $H = \{H_n\}_{n \in \mathbb{N}}$  such that  $H_n$  is a distribution over  $F_n$ ,  $H$  is the *uniform function ensemble* if  $H_n$  is uniformly distributed over  $F_n$ . A *permutation ensemble* is a sequence  $H = \{H_n\}_{n \in \mathbb{N}}$  such that  $H_n$  is a distribution over  $P_n$ ,  $H$  is the *uniform permutation ensemble* if  $H_n$  is uniformly distributed over  $P_n$ .

A function ensemble (or a permutation ensemble),  $H = \{H_n\}_{n \in \mathbb{N}}$ , is *efficiently computable* if the distribution  $H_n$  can be sampled efficiently (i.e., there exists a probabilistic polynomial-time Turing-machine,  $I$ , and a mapping from strings to functions,  $\phi$ , such that  $\phi(I(1^n))$  and  $H_n$  are identically distributed) and the functions in  $H_n$  can be computed efficiently (i.e., there exists a probabilistic polynomial-time Turing-machine,  $V$ , such that  $V(i, x) = (\phi(i))(x)$ ).

We would like to consider efficiently computable function (or permutation) ensembles that cannot be efficiently distinguished from the uniform ensemble. In our setting, the distinguisher is an oracle machine that can make queries to a length preserving function (or functions) and outputs a single bit. We assume that on input  $1^n$  the oracle machine makes only  $n$ -bit long queries,  $n$  also serves as the security parameter. An oracle machine has an interpretation both under the uniform complexity model and under the non-uniform model. In the former it is interpreted as a Turing-machine with a special oracle-tape (in this case efficient means probabilistic polynomial-time) and in the latter as a circuit-family with special oracle-gates (in this case efficient means polynomial-size). The discussion of this paper is independent of the chosen interpretation.

Let  $M$  be an oracle machine, let  $f$  be a function in  $F_n$  and  $H_n$  a distribution over  $F_n$ . Denote by  $M^f(1^n)$  the distribution of  $M$ 's output when its queries are answered by  $f$  and denote by  $M^{H_n}(1^n)$  the distribution  $M^f(1^n)$ , where  $f$  is distributed according to  $H_n$ . We would also like to consider oracle machines with access both to a permutation and to its inverse. Let  $M$  be such a machine, let  $f$  be a permutation in  $P_n$  and  $H_n$  a distribution over  $P_n$ . Denote by  $M^{f, f^{-1}}(1^n)$  the distribution of  $M$ 's output when its queries are answered by  $f$  and  $f^{-1}$  and denote by  $M^{H_n, H_n^{-1}}(1^n)$  the distribution  $M^{f, f^{-1}}(1^n)$ , where  $f$  is distributed according to  $H_n$ .

**Definition 2.2 (advantage)** Let  $M$  be an oracle machine and let  $H = \{H_n\}_{n \in \mathbb{N}}$  and  $\tilde{H} = \{\tilde{H}_n\}_{n \in \mathbb{N}}$  be two function (or permutation) ensembles. We call the function

$$|\text{Prob}[M^{H_n}(1^n) = 1] - \text{Prob}[M^{\tilde{H}_n}(1^n) = 1]|$$

the advantage  $M$  achieves in distinguishing between  $H$  and  $\tilde{H}$ .

Let  $M$  be an oracle machine and let  $H = \{H_n\}_{n \in \mathbb{N}}$  and  $\tilde{H} = \{\tilde{H}_n\}_{n \in \mathbb{N}}$  be two permutation ensembles. We call the function

$$|\text{Prob}[M^{H_n, H_n^{-1}}(1^n) = 1] - \text{Prob}[M^{\tilde{H}_n, \tilde{H}_n^{-1}}(1^n) = 1]|$$

the advantage  $M$  achieves in distinguishing between  $\langle H, H^{-1} \rangle$  and  $\langle \tilde{H}, \tilde{H}^{-1} \rangle$ .

We say that  $M$  distinguishes between  $H$  and  $\tilde{H}$  (resp.  $\langle H, H^{-1} \rangle$  and  $\langle \tilde{H}, \tilde{H}^{-1} \rangle$ ) with advantage  $\epsilon = \epsilon(n)$  if for infinitely many  $n$ 's the advantage  $M$  achieves in distinguishing between  $H$  and  $\tilde{H}$  (resp.  $\langle H, H^{-1} \rangle$  and  $\langle \tilde{H}, \tilde{H}^{-1} \rangle$ ) is at least  $\epsilon(n)$ .

**Definition 2.3 (negligible functions)** A function  $h : \mathbb{N} \mapsto \mathbb{R}$  is negligible if for every constant  $c > 0$  and all sufficiently large  $n$ 's

$$h(n) < \frac{1}{n^c}$$

**Definition 2.4 (PFE)** Let  $H = \{H_n\}_{n \in \mathbb{N}}$  be an efficiently computable function ensemble and let  $R = \{R_n\}_{n \in \mathbb{N}}$  be the uniform function ensemble.  $H$  is a pseudo-random function ensemble if for every efficient oracle-machine  $M$ , the advantage  $M$  has in distinguishing between  $H$  and  $R$  is negligible.

**Definition 2.5 (PPE)** Let  $H = \{H_n\}_{n \in \mathbb{N}}$  be an efficiently computable permutation ensemble and let  $R = \{R_n\}_{n \in \mathbb{N}}$  be the uniform permutation ensemble.  $H$  is a pseudo-random permutation ensemble if for every efficient oracle-machine  $M$ , the advantage  $M$  has in distinguishing between  $H$  and  $R$  is negligible.

**Definition 2.6 (SPPE)** Let  $H = \{H_n\}_{n \in \mathbb{N}}$  be an efficiently computable permutation ensemble and let  $R = \{R_n\}_{n \in \mathbb{N}}$  be the uniform permutation ensemble.  $H$  is a strong pseudo-random permutation ensemble if for every efficient oracle-machine  $M$ , the advantage  $M$  has in distinguishing between  $\langle H, H^{-1} \rangle$  and  $\langle R, R^{-1} \rangle$  is negligible.

**Remark 2.1** We use the phrase “ $f$  is a pseudo-random function” as an abbreviation for “ $f$  is distributed according to a pseudo-random function ensemble” and similarly for “ $f$  is a pseudo-random permutation” and “ $f$  is a strong pseudo-random permutation”

## 2.3 k-Wise Independent Functions and Permutations

The notions of  $k$ -wise independent functions and  $k$ -wise “almost” independent functions [3, 4, 13, 14, 25, 31, 50] (under several different formulations) play a major role in contemporary computer science. These are distributions of functions such that their value on any given  $k$  inputs is uniformly or “almost” uniformly distributed. Several constructions of such functions and a large variety of applications were suggested over the years.

We briefly review the definitions of  $k$ -wise independence (and  $k$ -wise  $\delta$ -dependence). The definitions of pair-wise independence (and pair-wise  $\delta$ -dependence) can be derived by taking  $k = 2$ .

**Definition 2.7** Let  $D_1$  and  $D_2$  be two distributions defined over  $\Omega$ , the variation distance between  $D_1$  and  $D_2$  is

$$\|D_1 - D_2\| = \frac{1}{2} \sum_{\omega \in \Omega} |D_1(\omega) - D_2(\omega)|$$

**Definition 2.8** Let  $A$  and  $B$  be two sets,  $0 \leq \delta \leq 1$ ,  $k$  an integer ( $2 \leq k \leq |A|$ ) and  $F$  a distribution of  $A \rightarrow B$  functions.

Let  $x_1, x_2, \dots, x_k$  be  $k$  different members of  $A$ , consider the following two distributions:

1.  $\langle f(x_1), f(x_2), \dots, f(x_k) \rangle$  where  $f$  is distributed according to  $F$ .
2. The uniform distribution over  $B^k$ .

$F$  is  $k$ -wise independent if for all  $x_1, x_2, \dots, x_k$  the two distributions are identical.  $F$  is  $k$ -wise  $\delta$ -dependent if for all  $x_1, x_2, \dots, x_k$  the two distributions are of variation distance at most  $\delta$ .

These definitions are naturally extended to permutations:

**Definition 2.9** Let  $A$  be a set,  $0 \leq \delta \leq 1$ ,  $k$  an integer ( $2 \leq k \leq |A|$ ) and  $F$  a distribution of permutations over  $A$ .

Let  $x_1, x_2, \dots, x_k$  be  $k$  different members of  $A$ , consider the following two distributions:

1.  $\langle f(x_1), f(x_2), \dots, f(x_k) \rangle$  where  $f$  is distributed according to  $F$ .
2. The uniform distribution over sequences of  $k$  different elements of  $A$ .

$F$  is  $k$ -wise independent if for all  $x_1, x_2, \dots, x_k$  the two distributions are identical.  $F$  is  $k$ -wise  $\delta$ -dependent if for all  $x_1, x_2, \dots, x_k$  the two distributions are of variation distance at most  $\delta$ .

The connection of this paper to  $k$ -wise independence is bidirectional as described in the following two paragraphs.

As shown in Section 3, pair-wise independent permutations can replace the first and fourth rounds of the LR-Construction. Let  $A$  be a finite field then the permutation  $f_{a,b}(x) \stackrel{\text{def}}{=} a \cdot x + b$ , where  $a \neq 0, b \in A$  are uniformly distributed, is pair-wise independent. Thus, there are pair-wise independent permutations over  $I_n$  (the permutations  $f_{a,b}$  with operations over  $GF(2^n)$ ). In Section 5.2, it is shown that we can use even more efficient functions and permutations in our construction. In particular, we consider the concept of  $\epsilon$ -AXU<sub>2</sub> functions.

In contrast with the case of pair-wise independent permutations, we are not aware of any “good” constructions of  $k$ -wise  $\delta$ -dependent permutations for general  $k$  and  $\delta$ . The different variants of the LR-Construction offer a partial solution to this problem (“partial” because of the bounded values of  $\delta$  that can be achieved). For example, using  $k$ -wise  $\delta'$ -dependent functions on  $n$  bits instead of pseudo-random functions in the original LR-Construction yields a  $k$ -wise  $\delta$ -dependent permutation on  $2n$  bits (for  $\delta = O(k^2/2^n + \delta')$ ). In Section 8 we analyze the different constructions of this paper as constructions of  $k$ -wise  $\delta$ -dependent permutations.

### 3 Construction of PPE and SPPE

#### 3.1 Intuition

As mentioned in the introduction, a principle observation of this paper is that the different rounds of the LR-Construction serve significantly different roles. To illustrate this point, consider two rounds of the construction. Namely,  $E = D_{f_2} \circ D_{f_1}$ , where  $f_1, f_2 \in F_n$  are two independently chosen pseudo-random functions. It is not hard to verify that  $E$  is computationally indistinguishable from a random permutation to any efficient algorithm that has access to pairs  $\{\langle x_i, E(x_i) \rangle\}_{i=1}^m$ , where the sequence  $\{x_i\}_{i=1}^m$  is uniformly distributed. The intuition is as follows: First we should notice



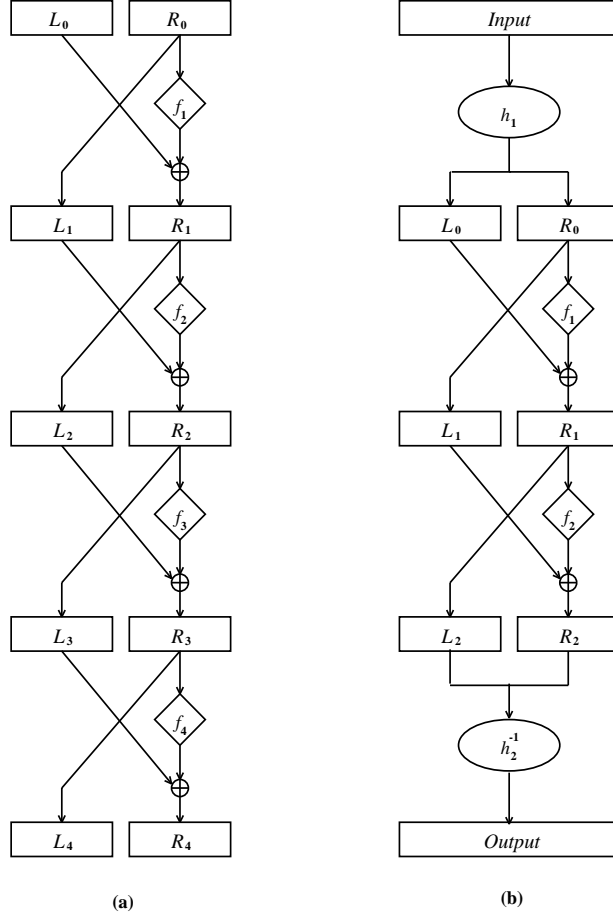


Figure 1: Constructions of SPPE: (a) The original LR-Construction (b) The revised Construction. In (a) and (b):  $\forall i \geq 1$ ,  $L_i = R_{i-1}$  and  $R_i = L_{i-1} \oplus f_i(R_{i-1})$ . In (b):  $\langle L_0, R_0 \rangle = h_1(\text{Input})$  and  $\text{Output} = h_2^{-1}(\langle L_2, R_2 \rangle)$ .

that it is enough to prove the pseudo-randomness of  $E$  when  $f_1$  and  $f_2$  are *truly random functions* (instead of pseudo-random). Let  $(L_i^0, R_i^0) = x_i$  and  $(L_i^2, R_i^2) = E(x_i)$ , by the definition of  $E$  we get that  $L_i^2 = L_i^0 \oplus f_1(R_i^0)$  and  $R_i^2 = R_i^0 \oplus f_2(L_i^2)$ . Since the sequence  $\{x_i\}_{i=1}^m$  is uniformly distributed, we have that with good probability (better than  $(1 - \frac{m^2}{2^{n+1}})$ )  $R_i^0 \neq R_j^0$  for all  $i \neq j$ . Conditioned on this event, the sequence  $\{L_i^2\}_{i=1}^m$  is uniformly distributed (since  $f_1$  is random). We now have that with good probability  $L_i^2 \neq L_j^2$  for all  $i \neq j$ . Conditioned on this event, the sequence  $\{R_i^2\}_{i=1}^m$  is uniformly distributed (and independent of the sequence  $\{L_i^2\}_{i=1}^m$ ). Notice that this argument still works if the sequence  $\{x_i\}_{i=1}^m$  is only pair-wise independent.

Nevertheless, as Luby and Rackoff showed,  $E$  can be easily distinguished from a random permutation by an algorithm that gets to see the value of  $E$  or  $E^{-1}$  on inputs of its choice. The reason is that for any values  $L_1, L_2$  and  $R$  such that  $L_1 \neq L_2$  we have that  $E(L_1, R)|_L \oplus E(L_2, R)|_L = L_1 \oplus L_2$ . In contrast, for a truly random permutation, the probability of this event is  $2^{-n}$ . This is the reason that the LR-Construction includes three or four rounds.

We think of the second and third rounds of the LR-Construction as the two-round construction  $E$ , described above and show that the role of the first and fourth rounds is to prevent the distinguisher from directly choosing the inputs of  $E$  and  $E^{-1}$ . As we shall see, this goal can also be achieved with “combinatorial” constructions (e.g., pair-wise independent permutations), rather than “cryptographic” (i.e., pseudo-random functions). In particular, the LR-Construction remains secure when the first and fourth rounds are replaced with pair-wise independent permutations (see Figure 1 for an illustration).

### 3.2 Construction and Main Result

**Definition 3.1** For any  $f_1, f_2 \in F_n$  and  $h_1, h_2 \in P_{2n}$ , define

$$W(h_1, f_1, f_2) \stackrel{\text{def}}{=} D_{f_2} \circ D_{f_1} \circ h_1$$

and

$$S(h_1, f_1, f_2, h_2) \stackrel{\text{def}}{=} h_2^{-1} \circ D_{f_2} \circ D_{f_1} \circ h_1$$

**Theorem 3.1** Let  $h_1, h_2 \in P_{2n}$  be pair-wise independent permutations (similarly to Remark 2.1 this is an abbreviation for “distributed according to a pair-wise independent permutation ensemble”) and let  $f_1, f_2 \in F_n$  be pseudo-random functions;  $h_1, h_2, f_1$  and  $f_2$  are independently chosen. Then,  $W = W(h_1, f_1, f_2)$  is a pseudo-random permutation and  $S = S(h_1, f_1, f_2, h_2)$  is a strong pseudo-random permutation ( $W$  and  $S$  as in Definition 3.1).

Furthermore, assume that no efficient oracle-machine that makes at most  $m = m(n)$  queries, distinguishes between the pseudo-random functions and random functions with advantage  $\epsilon = \epsilon(n)$  (see Definition 2.2). Then, no efficient oracle-machine that makes at most  $m$  queries to  $W$  (resp.  $S$  and  $S^{-1}$ ) distinguishes  $W$  (resp.  $S$ ) from a random permutation with advantage  $2\epsilon + \frac{m^2}{2^n} + \frac{m^2}{2^{2n}}$ .

**Remark 3.1** The conditions of Theorem 3.1 are meant to simplify the exposition of the theorem and of its proof. These conditions can be relaxed, as discussed in Section 5. The main points are the following:

1. A single pseudo-random function  $f$  can replace both  $f_1$  and  $f_2$
2.  $h_1$  and  $h_2$  may obey weaker requirements than pair-wise independence. For example, it is enough that for every  $x \neq y$ :

$$\text{Prob}[h_1(x)|_R = h_1(y)|_R] \leq 2^{-n} \text{ and } \text{Prob}[h_2(x)|_L = h_2(y)|_L] \leq 2^{-n}$$

### 3.3 Proof of Security

We now prove the security of the SPPE-construction; the proof of security for the PPE-construction is very similar (and, in fact, a bit simpler). As with the original LR-Construction, the main task is to prove that the permutations are pseudo-random when  $f_1$  and  $f_2$  are truly random (instead of pseudo-random).

**Theorem 3.2** *Let  $h_1, h_2 \in P_{2n}$  be pair-wise independent permutations and let  $f_1, f_2 \in F_n$  be random functions. Define  $S = S(h_1, f_1, f_2, h_2)$  (as in Definition 3.1) and let  $R \in P_{2n}$  be a random permutation. Then, for any oracle machine  $M$  (not necessarily an efficient one) that makes at most  $m$  queries:*

$$|\text{Prob}[M^{S, S^{-1}}(1^{2n}) = 1] - \text{Prob}[M^{R, R^{-1}}(1^{2n}) = 1]| \leq \frac{m^2}{2^n} + \frac{m^2}{2^{2n}}$$

Theorem 3.1 follows easily from Theorem 3.2 (see a proof-sketch in the subsequent). In order to prove Theorem 3.2, we introduce some additional notations.

Let  $G$  denote the permutation that is accessible to the machine  $M$  ( $G$  is either  $S$  or  $R$ ). There are two types of queries  $M$  can make: either  $(+, x)$  which denotes the query “what is  $G(x)$ ?” or  $(-, y)$  which denotes the query “what is  $G^{-1}(y)$ ?”. For the  $i$ th query  $M$  makes, define the query-answer pair  $\langle x_i, y_i \rangle \in I_{2n} \times I_{2n}$ , where either  $M$ ’s query was  $(+, x_i)$  and the answer it got was  $y_i$  or  $M$ ’s query was  $(-, y_i)$  and the answer it got was  $x_i$ . We assume that  $M$  makes exactly  $m$  queries and refer to the sequence  $\{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$  of all these pairs as the *transcript* (of  $M$ ’s computation).

Notice that no limitations were imposed on the computational power of  $M$ . Therefore,  $M$  can be assumed to be deterministic (we can always fix the random tape that maximizes the advantage  $M$  achieves). This assumption implies that for every  $1 \leq i \leq m$  the  $i$ th query of  $M$  is fully determined by the first  $i - 1$  query-answer pairs. Thus, for every  $i$  it can be determined from the transcript whether the  $i$ th query was  $(+, x_i)$  or  $(-, y_i)$ . We also get that  $M$ ’s output is a (deterministic) function of its transcript. Denote by  $C_M[\{\langle x_1, y_1 \rangle, \dots, \langle x_{i-1}, y_{i-1} \rangle\}]$  the  $i$ th query of  $M$  as a function of the previous query-answer pairs and denote by  $C_M[\{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}]$  the output of  $M$  as a function of its transcript.

**Definition 3.2** *Let  $\sigma$  be a sequence  $\{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$ , where for  $1 \leq i \leq m$  we have that  $\langle x_i, y_i \rangle \in I_{2n} \times I_{2n}$ . Then,  $\sigma$  is a possible  $M$ -transcript if for every  $1 \leq i \leq m$*

$$C_M[\{\langle x_1, y_1 \rangle, \dots, \langle x_{i-1}, y_{i-1} \rangle\}] \in \{(+, x_i), (-, y_i)\}$$

Let us consider yet another distribution on the answers to  $M$ ’s queries (which, in turn, induces another distribution on the possible  $M$ -transcripts). Consider a random process  $\tilde{R}$  that on the  $i$ th query of  $M$  answers as follows:

1. If  $M$ ’s query is  $(+, x)$  and for some  $1 \leq j < i$  the  $j$ th query-answer pair is  $\langle x, y \rangle$ , then  $\tilde{R}$ ’s answer is  $y$  (for an arbitrary such query-answer pair,  $\langle x, y \rangle$ ).
2. If  $M$ ’s query is  $(-, y)$  and for some  $1 \leq j < i$  the  $j$ th query-answer pair is  $\langle x, y \rangle$ , then  $\tilde{R}$ ’s answer is  $x$  (for an arbitrary such query-answer pair,  $\langle x, y \rangle$ ).
3. If neither 1 nor 2 holds, then  $\tilde{R}$ ’s answer is a uniformly chosen  $2n$ -bit string.

It is possible that  $\tilde{R}$  provides answers that are not consistent with *any* permutation; that is, we can have two query-answer pairs of the form  $\langle x_1, y \rangle$  and  $\langle x_2, y \rangle$  for  $x_1 \neq x_2$  or  $\langle x, y_1 \rangle$  and  $\langle x, y_2 \rangle$  for  $y_1 \neq y_2$ . In this case call the transcript *inconsistent*, otherwise, the transcript is called *consistent*.

We first show (in Proposition 3.3) that the advantage  $M$  might have in distinguishing between the process  $\tilde{R}$  and the random permutation  $R$  is small. The reason is that as long as  $\tilde{R}$  answers consistently (which happens with good probability) it “behaves” exactly as a random permutation. In order to formalize this, we consider the different distributions on the transcript of  $M$  (induced by the different distributions on the answers it gets).

**Definition 3.3** Let  $T_S$ ,  $T_R$  and  $T_{\tilde{R}}$  be the random variables such that  $T_S$  is the transcript of  $M$  when its queries are answered by  $S$ ,  $T_R$  is the transcript of  $M$  when its queries are answered by  $R$  and  $T_{\tilde{R}}$  is the transcript of  $M$  when its queries are answered by  $\tilde{R}$ .

Notice that by these definitions (and by our assumptions)  $M^{S, S^{-1}}(1^{2n}) = C_M(T_S)$  (are the same random variables) and  $M^{R, R^{-1}}(1^{2n}) = C_M(T_R)$ .

**Proposition 3.3**

$$|Prob_{\tilde{R}}[C_M(T_{\tilde{R}}) = 1] - Prob_R[C_M(T_R) = 1]| \leq \frac{m^2}{2^{2n+1}}$$

*Proof:* For any possible and consistent  $M$ -transcript  $\sigma$  we have that

$$Prob_R[T_R = \sigma] = \frac{2^{2n}!}{(2^{2n} - m)!} = Prob_{\tilde{R}}[T_{\tilde{R}} = \sigma \mid T_{\tilde{R}} \text{ is consistent}].$$

Therefore, the distribution of  $T_{\tilde{R}}$  conditioned on  $T_{\tilde{R}}$  being consistent is exactly the distribution of  $T_R$ . Furthermore, the probability that  $T_{\tilde{R}}$  is inconsistent is small:  $T_{\tilde{R}}$  is inconsistent if for some  $1 \leq j < i \leq m$  the corresponding query-answer pairs satisfy  $x_i = x_j$  and  $y_i \neq y_j$  or  $y_i = y_j$  and  $x_i \neq x_j$ . For a given  $i$  and  $j$  this event happens with probability at most  $2^{-2n}$ . Hence,

$$Prob_{\tilde{R}}[T_{\tilde{R}} \text{ is inconsistent}] \leq \binom{m}{2} \cdot 2^{-2n} < \frac{m^2}{2^{2n+1}}$$

The proposition follows:

$$\begin{aligned} & |Prob_{\tilde{R}}[C_M(T_{\tilde{R}}) = 1] - Prob_R[C_M(T_R) = 1]| \\ \leq & |Prob_{\tilde{R}}[C_M(T_{\tilde{R}}) = 1 \mid T_{\tilde{R}} \text{ is consistent}] - Prob_R[C_M(T_R) = 1]| \cdot Prob_{\tilde{R}}[T_{\tilde{R}} \text{ is consistent}] \\ + & |Prob_{\tilde{R}}[C_M(T_{\tilde{R}}) = 1 \mid T_{\tilde{R}} \text{ is inconsistent}] - Prob_R[C_M(T_R) = 1]| \cdot Prob_{\tilde{R}}[T_{\tilde{R}} \text{ is inconsistent}] \\ \leq & Prob_{\tilde{R}}[T_{\tilde{R}} \text{ is inconsistent}] \\ < & \frac{m^2}{2^{2n+1}} \end{aligned}$$

□

It remains to bound the advantage  $M$  might have in distinguishing between  $T_{\tilde{R}}$  and  $T_S$ . The intuition is that for every possible and consistent  $M$ -transcript  $\sigma$  unless some “bad” and “rare” event on the choice of  $h_1$  and  $h_2$  (as in the definition of  $S$ ) happens, the probability that  $T_S = \sigma$  is exactly the same as the probability that  $T_{\tilde{R}} = \sigma$ . We now formally define this event (Definition 3.4) and bound its probability (Proposition 3.4).

For any possible  $M$ -transcript  $\sigma = \{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$  we can assume that if  $\sigma$  is consistent then for  $i \neq j$  both  $x_i \neq x_j$  and  $y_i \neq y_j$  (this means that  $M$  never asks a query if its answer is determined by a previous query-answer pair).

**Definition 3.4** For every specific choice of pair-wise independent permutations  $h_1, h_2 \in P_{2n}$  (in the definition of  $S$ ) define  $BAD(h_1, h_2)$  to be the set of all possible and consistent  $M$ -transcripts,  $\sigma = \{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$ , satisfying:

$$\exists 1 \leq i < j \leq m \text{ such that } h_1(x_i)|_R = h_1(x_j)|_R \text{ or } h_2(y_i)|_L = h_2(y_j)|_L$$

**Proposition 3.4** Let  $h_1, h_2 \in P_{2n}$  be pair-wise independent permutations then for any possible and consistent  $M$ -transcript  $\sigma = \{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$  we have that:

$$Prob_{h_1, h_2}[\sigma \in BAD(h_1, h_2)] < \frac{m^2}{2^n}$$

*Proof:* By definition,  $\sigma \in BAD(h_1, h_2)$  if there exist  $1 \leq i < j \leq m$  such that either  $h_1(x_i)|_R = h_1(x_j)|_R$  or  $h_2(y_i)|_L = h_2(y_j)|_L$ . For any given  $i$  and  $j$  both  $Prob_{h_1}[h_1(x_i)|_R = h_1(x_j)|_R]$  and  $Prob_{h_2}[h_2(y_i)|_L = h_2(y_j)|_L]$  are smaller than  $2^{-n}$  (since  $h_1$  and  $h_2$  are pair-wise independent). Therefore,

$$Prob_{h_1, h_2}[\sigma \in BAD(h_1, h_2)] < \binom{m}{2} \cdot 2 \cdot 2^{-n} < \frac{m^2}{2^n}$$

□

The key lemma for proving Theorem 3.2 is:

**Lemma 3.5** Let  $\sigma = \{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$  be any possible and consistent  $M$ -transcript, then

$$Prob_S[T_S = \sigma \mid \sigma \notin BAD(h_1, h_2)] = Prob_{\tilde{R}}[T_{\tilde{R}} = \sigma]$$

*Proof:* Since  $\sigma$  is a possible  $M$ -transcript we have that for all  $1 \leq i \leq m$ ,

$$C_M[\{\langle x_1, y_1 \rangle, \dots, \langle x_{i-1}, y_{i-1} \rangle\}] \in \{(+, x_i), (-, y_i)\}$$

Therefore,  $T_{\tilde{R}} = \sigma$  iff for all  $1 \leq i \leq m$ , the  $i$ th answer  $\tilde{R}$  gives is  $y_i$  in the case that  $C_M[\{\langle x_1, y_1 \rangle, \dots, \langle x_{i-1}, y_{i-1} \rangle\}] = (+, x_i)$  and otherwise its  $i$ th answer is  $x_i$ . By our assumptions and the definition of  $\tilde{R}$ , given that  $\tilde{R}$  answered “correctly” on each one of the first  $i - 1$  queries its  $i$ th answer is an independent and uniform  $2n$ -bit string. Therefore,

$$Prob_{\tilde{R}}[T_{\tilde{R}} = \sigma] = 2^{-2nm}$$

Since  $\sigma$  is a possible  $M$ -transcript we have that  $T_S = \sigma$  iff for all  $1 \leq i \leq m$ ,  $y_i = S(x_i)$ . Consider any specific choice of permutations  $h_1$  and  $h_2$  (for which  $S = S(h_1, f_1, f_2, h_2)$ ) such that  $\sigma \notin BAD(h_1, h_2)$ . Let  $(L_i^0, R_i^0) = h_1(x_i)$  and  $(L_i^2, R_i^2) = h_2(y_i)$ . By the definition of  $S$ , we get that:

$$y_i = S(x_i) \iff f_1(R_i^0) = L_i^0 \oplus L_i^2 \text{ and } f_2(L_i^2) = R_i^0 \oplus R_i^2$$

For every  $1 \leq i < j \leq m$  both  $R_i^0 \neq R_j^0$  and  $L_i^2 \neq L_j^2$  (otherwise  $\sigma \in BAD(h_1, h_2)$ ). Therefore, since  $f_1$  and  $f_2$  are random, we have that for every choice of  $h_1$  and  $h_2$  such that  $\sigma \notin BAD(h_1, h_2)$  the probability that  $T_S = \sigma$  is exactly  $2^{-2nm}$ . We can conclude:

$$Prob_S[T_S = \sigma \mid \sigma \notin BAD(h_1, h_2)] = 2^{-2nm}$$

which complete the proof of the lemma. □

*Proof: (of Theorem 3.2)* Let  $\Gamma$  be the set of all possible and consistent  $M$ -transcripts  $\sigma$  such that  $M(\sigma) = 1$ .

$$\begin{aligned}
& |Prob_S[C_M(T_S) = 1] - Prob_{\tilde{R}}[C_M(T_{\tilde{R}}) = 1]| \\
& \leq \left| \sum_{\sigma \in \Gamma} (Prob_S[T_S = \sigma] - Prob_{\tilde{R}}[T_{\tilde{R}} = \sigma]) \right| + Prob_{\tilde{R}}[T_{\tilde{R}} \text{ is inconsistent}] \\
& \leq \sum_{\sigma \in \Gamma} |Prob_S[T_S = \sigma | \sigma \notin \text{BAD}(h_1, h_2)] - Prob_{\tilde{R}}[T_{\tilde{R}} = \sigma]| \cdot Prob_{h_1, h_2}[\sigma \notin \text{BAD}(h_1, h_2)] \quad (1) \\
& + \left| \sum_{\sigma \in \Gamma} (Prob_S[T_S = \sigma | \sigma \in \text{BAD}(h_1, h_2)] - Prob_{\tilde{R}}[T_{\tilde{R}} = \sigma]) \cdot Prob_{h_1, h_2}[\sigma \in \text{BAD}(h_1, h_2)] \right| \quad (2) \\
& + Prob_{\tilde{R}}[T_{\tilde{R}} \text{ is inconsistent}] \quad (3)
\end{aligned}$$

We already showed in the proof of Proposition 3.3 that  $(3) = Prob_{\tilde{R}}[T_{\tilde{R}} \text{ is inconsistent}] < \frac{m^2}{2^{2n+1}}$ , by Lemma 3.5 we get that  $(1) = 0$ . Therefore, it remains to bound  $(2)$ : Assume without loss of generality that

$$\begin{aligned}
& \sum_{\sigma \in \Gamma} Prob_S[T_S = \sigma | \sigma \in \text{BAD}(h_1, h_2)] \cdot Prob_{h_1, h_2}[\sigma \in \text{BAD}(h_1, h_2)] \\
& \leq \sum_{\sigma \in \Gamma} Prob_{\tilde{R}}[T_{\tilde{R}} = \sigma] \cdot Prob_{h_1, h_2}[\sigma \in \text{BAD}(h_1, h_2)]
\end{aligned}$$

then using Proposition 3.4 we get that

$$\begin{aligned}
& \left| \sum_{\sigma \in \Gamma} (Prob_S[T_S = \sigma | \sigma \in \text{BAD}(h_1, h_2)] - Prob_{\tilde{R}}[T_{\tilde{R}} = \sigma]) \cdot Prob_{h_1, h_2}[\sigma \in \text{BAD}(h_1, h_2)] \right| \\
& \leq \sum_{\sigma \in \Gamma} Prob_{\tilde{R}}[T_{\tilde{R}} = \sigma] \cdot Prob_{h_1, h_2}[\sigma \in \text{BAD}(h_1, h_2)] \\
& \leq \max_{\sigma \in \Gamma} Prob_{h_1, h_2}[\sigma \in \text{BAD}(h_1, h_2)] \\
& < \frac{m^2}{2^n}
\end{aligned}$$

Thus, we can conclude that:

$$|Prob_S[C_M(T_S) = 1] - Prob_{\tilde{R}}[C_M(T_{\tilde{R}}) = 1]| < \frac{m^2}{2^n} + \frac{m^2}{2^{2n+1}}$$

Using Proposition 3.3 we complete the proof:

$$\begin{aligned}
& \left| Prob_S[M^{S, S^{-1}}(1^{2n}) = 1] - Prob_R[M^{R, R^{-1}}(1^{2n}) = 1] \right| \\
& = |Prob_S[C_M(T_S) = 1] - Prob_R[C_M(T_R) = 1]| \\
& \leq |Prob_S[C_M(T_S) = 1] - Prob_{\tilde{R}}[C_M(T_{\tilde{R}}) = 1]| + |Prob_{\tilde{R}}[C_M(T_{\tilde{R}}) = 1] - Prob_R[C_M(T_R) = 1]| \\
& < \frac{m^2}{2^n} + \frac{m^2}{2^{2n}}
\end{aligned}$$

□

Given Theorem 3.2, the proof of Theorem 3.1 is essentially the same as the corresponding proof of the original LR-Construction (the proof of Theorem 1 of [27], given their main Lemma). The proof-idea is the following: Define three distributions:

- $S_1 = S(h_1, f_1, f_2, h_2)$ , where  $h_1, h_2 \in P_{2n}$  are pair-wise independent and  $f_1, f_2 \in F_n$  are pseudo-random functions.
- $S_2 = S(h_1, g_1, f_2, h_2)$ , where  $h_1, h_2 \in P_{2n}$  are pair-wise independent,  $f_2 \in F_n$  a pseudo-random function and  $g_1 \in F_n$  a random function.
- $S_3 = S(h_1, g_1, g_2, h_2)$ , where  $h_1, h_2 \in P_{2n}$  are pair-wise independent and  $g_1, g_2 \in F_n$  are random functions.

It is enough to show that for every oracle machine, for all but finite number of  $n$ 's:

1.  $|Prob[M^{S_1, S_1^{-1}}(1^{2n}) = 1] - Prob[M^{S_2, S_2^{-1}}(1^{2n}) = 1]| \leq \epsilon(n)$
2.  $|Prob[M^{S_2, S_2^{-1}}(1^{2n}) = 1] - Prob[M^{S_3, S_3^{-1}}(1^{2n}) = 1]| \leq \epsilon(n)$

If (1) or (2) do not hold, then we can construct an efficient oracle-machine  $M'$  that distinguishes a pseudo-random function from a random one with advantage greater than  $\epsilon$ , in contradiction to the assumption. Assume for example that for infinitely many  $n$ 's:

$$|Prob[M^{S_1, S_1^{-1}}(1^{2n}) = 1] - Prob[M^{S_2, S_2^{-1}}(1^{2n}) = 1]| > \epsilon(n)$$

The oracle-machine  $M'$  on input  $1^n$  and with access to a function  $O \in F_n$  first samples pair-wise independent permutations,  $h_1, h_2 \in P_{2n}$ , and a pseudo-random function  $f_2 \in F_n$ .  $M'$  then invokes  $M$  with input  $1^{2n}$  and answers its queries with the values of  $S$  and  $S^{-1}$ , for  $S = S(h_1, O, f_2, h_2)$ . When  $M$  halts so does  $M'$  and it outputs whatever was the output of  $M$ . Notice that if  $O$  is a pseudo-random function then the distribution of  $S$  is  $S_1$  whereas if  $O$  is a truly random function then the distribution of  $S$  is  $S_2$ . This is the reason that  $M'$  distinguishes a pseudo-random function from a random one with advantage greater than  $\epsilon$ . Similar hybrid-arguments apply to all other constructions of this paper.

## 4 The Framework

As we shall see in Sections 5–7, the construction of Section 3 can be relaxed and generalized in several ways. The different pseudo-random permutations obtained share a similar structure and almost identical proof of security. In this section we examine the proof of Theorem 3.2 in a more abstract manner. Our goal is to establish a framework for proving (almost) all the constructions of this paper and to suggest a way for designing and proving additional constructions.

Our framework deals with constructions of a pseudo-random permutation  $S$  on  $\ell$  bits which is the composition of three permutations:  $S \equiv h_2^{-1} \circ E \circ h_1$ . (see Figure 2 for an illustration). In general,  $h_1$  and  $h_2^{-1}$  are “lightweight” and  $E$  is where most of the work is done.  $E$  is constructed from pseudo-random functions and for the purpose of the analysis we assume (as in Theorem 3.2) that these functions are truly random. In Section 3, for example,  $\ell = 2n$ ,  $h_1$  and  $h_2$  are chosen as pair-wise independent permutations and  $E \equiv D_{f_2} \circ D_{f_1}$  for random  $f_1, f_2 \in F_n$ .

The framework starts with  $E$  which may be easily distinguished from a truly random permutation and transforms it via  $h_1$  and  $h_2$  into a pseudo-random permutation. The property  $E$  should have is that for almost every sequence,  $\{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$ , the probability that  $\forall i, y_i = E(x_i)$  is “close” to what we have for a truly random permutation: Call a sequence,  $\{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$ ,  $E$ -Good if  $Prob_E[\forall i, y_i = E(x_i)] = 2^{-l \cdot m}$ . We assume that apart from some “rare” sequences all others are  $E$ -Good. Loosely speaking, the role of  $h_1$  and  $h_2$  is to ensure that under any (adaptive

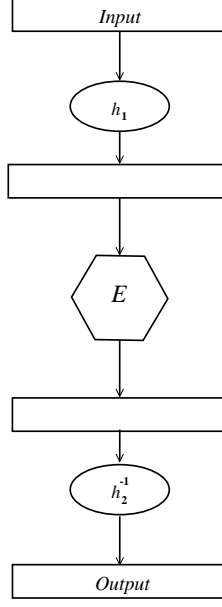


Figure 2: The high-level structure of the different constructions of SPPE.

chosen plaintext and ciphertext) attack on  $S$  the inputs and outputs of  $E$  form an  $E$ -Good sequence with a very high probability.

For the exact properties needed from the distributions on  $h_1, h_2$  and  $E$ , we shall try to follow the statement and proof of Theorem 3.2. The goal is to show that  $S$  is indistinguishable from a truly random permutation  $R$  on  $\ell$  bits. Specifically, that for some small  $\epsilon$  (whose choice will be explained hereafter), for any oracle machine  $M$  (not necessarily an efficient one) that makes at most  $m$  queries:

$$|\text{Prob}[M^{S,S^{-1}}(1^\ell) = 1] - \text{Prob}[M^{R,R^{-1}}(1^\ell) = 1]| \leq \epsilon + \frac{m^2}{2^\ell}.$$

Let the notions of query-answer pair, a transcript, the function  $C_M$ , a possible  $M$ -transcript, the random process  $\tilde{R}$ , a consistent transcript and the different random variables  $T_S$ ,  $T_R$  and  $T_{\tilde{R}}$  be as in the proof of Theorem 3.2. Proposition 3.3 (saying that the distance between  $T_R$  and  $T_{\tilde{R}}$  is bounded by the probability that  $T_{\tilde{R}}$  is inconsistent and that this probability is bounded by  $\frac{m^2}{2^{\ell+1}}$ ) still holds. The heart of applying the framework is in specifying the “bad”  $M$ -transcripts for given  $h_1$  and  $h_2$ . This set  $\text{BAD}_E(h_1, h_2)$  replaces  $\text{BAD}(h_1, h_2)$  in Definition 3.4 and in the rest of the proof. It contains possible and consistent  $M$ -transcripts and should have the property that any  $\{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$  not in  $\text{BAD}_E(h_1, h_2)$  satisfies that  $\{\langle h_1(x_1), h_2(y_1) \rangle, \dots, \langle h_1(x_m), h_2(y_m) \rangle\}$  is  $E$ -Good. Note that Definition 3.4 is indeed a special case of the above and also that, by this property,

$$\text{Prob}_S[T_S = \sigma \mid \sigma \notin \text{BAD}_E(h_1, h_2)] = 2^{-\ell \cdot m}$$

This implies that Lemma 3.5 where  $\text{BAD}(h_1, h_2)$  is replaced with  $\text{BAD}_E(h_1, h_2)$  is true:

**Lemma 4.1** *Let  $\sigma = \{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$  be any possible and consistent  $M$ -transcript, then*

$$\text{Prob}_S[T_S = \sigma \mid \sigma \notin \text{BAD}_E(h_1, h_2)] = \text{Prob}_{\tilde{R}}[T_{\tilde{R}} = \sigma].$$



For  $\text{BAD}_E(h_1, h_2)$  to be useful we must have that

$$\text{Prob}_{h_1, h_2}[\sigma \in \text{BAD}_E(h_1, h_2)] \leq \epsilon \quad (1)$$

and this substitutes Proposition 3.4. This is the only place in the proof where we use the definition of  $\epsilon$  and the definition of the distributions of  $h_1$  and  $h_2$ . As demonstrated in Sections 5.2 & 7.1, there is actually a tradeoff between reducing the requirements from  $h_1$  and  $h_2$  and having a somewhat larger value of  $\epsilon$ . Applying (1) and Lemma 4.1 as in the proof of Theorem 3.2 we conclude:

**Theorem 4.2** *Let  $h_1, h_2, E$  be distributed over permutations in  $P_\ell$ , let  $S \equiv h_2^{-1} \circ E \circ h_1$  and let  $R \in P_\ell$  be a random permutation. Suppose that  $\text{BAD}_E(h_1, h_2)$  is as above and  $\epsilon$  satisfies (1). Then, for any oracle machine  $M$  (not necessarily an efficient one) that makes at most  $m$  queries:*

$$|\text{Prob}[M^{S, S^{-1}}(1^\ell) = 1] - \text{Prob}[M^{R, R^{-1}}(1^\ell) = 1]| \leq \epsilon + \frac{m^2}{2^\ell}$$

To summarize, the major point in proving the security of the different constructions is to define the set  $\text{BAD}_E(h_1, h_2)$  such that for any possible and consistent  $M$ -transcript,  $\sigma$ , both  $\text{Prob}_S[T_S = \sigma \mid \sigma \notin \text{BAD}_E(h_1, h_2)] = 2^{-\ell \cdot m}$  and  $\text{Prob}_{h_1, h_2}[\sigma \in \text{BAD}_E(h_1, h_2)] \leq \epsilon$  (for the specific  $\epsilon$  in the claim we are proving). This suggests that the critical step for designing a pseudo-random permutation, using the framework described in this section, is to come up with a permutation  $E$  such that the set of  $E$ -Good sequences is “large enough” and “nice enough”. Note that to meet this end one can use different or more general definitions of an  $E$ -Good sequence with only minor changes to the proof (as is the case for the permutation  $\hat{S}$  in Section 7).

## 5 Relaxing the Construction

### 5.1 PPE and SPPE with a Single Pseudo-Random Function

Since Luby and Rackoff introduced their construction a considerable amount of research [34, 35, 36, 37, 39, 43, 44, 45, 47, 52] was focused on the following question: Can we obtain a similar construction of PPE or SPPE such that every permutation will be constructed from a *single* pseudo-random function?

Apparently, this line of research originated in the work of Schnorr [47]. Schnorr considered the LR-Construction, where the functions used are truly random, as a pseudo-random generator that is secure if not too many bits are accessible. The security of Schnorr’s generator does not depend on any unproven assumption. This notion of local-randomness is further treated in [29, 30]. Since the key of a random function is huge it makes sense to minimize the number of functions and, indeed, Schnorr suggested  $D_f \circ D_f \circ D_f$  as pseudo-random (the suggested permutation was later shown to be distinguishable from random [43]).

Following is an informal description of some of these results. Let  $f \in F_n$  be a random function, then:

- For all  $i, j, k \geq 1$  the permutation  $D_{f^i} \circ D_{f^j} \circ D_{f^k}$  is not pseudo-random [52].
- For all  $i, j, k, l \geq 1$  the permutation  $D_{f^i} \circ D_{f^j} \circ D_{f^k} \circ D_{f^l}$  is not strongly pseudo-random [44].
- $D_{f^2} \circ D_f \circ D_f \circ D_f$  is pseudo-random [39].
- $D_f \circ D_I \circ D_{f^2} \circ D_f \circ D_I \circ D_{f^2}$  is strongly pseudo-random, where  $I \in F_n$  is the identity function [45].

- $D_{f \circ \xi \circ f} \circ D_f \circ D_f$  is pseudo-random and  $D_{f \circ \xi \circ f} \circ D_f \circ D_f \circ D_f$  is strongly pseudo-random, where  $\xi$  is, for example, a rotation of one bit [37].

A critique which has been voiced often is that using only one pseudo-random function does not seem too significant: A pseudo-random function on  $n+2$  bits can replace 4 pseudo-random functions on  $n$  bits or, alternatively, a small key can be used to pseudo-randomly generate a larger key. It should also be noticed that the new constructions require additional invocations of the pseudo-random functions which imply an increase in the computation time. Furthermore, these results involve detailed and non-trivial proofs (to a point, where some papers claim to find inaccuracies in others).

The adjustment of the LR-Construction we suggest in Section 3 can easily be converted into a construction of PPE and SPPE from a single pseudo-random function. Simply replace both (pseudo-random) functions,  $f_1$  and  $f_2$ , with a single (pseudo-random) function  $f$ . This solution does not suffer from the drawbacks of the previous ones. The construction and the proof remain as simple as before and the pseudo-random function is only invoked twice at each computation of the permutation. The additional key-size for the pair-wise independent functions ( $h_1$  and  $h_2$ ) is not substantial (especially compared to the size of a truly random function). Consider, for example, the construction of *SPPE* when we use a truly random function  $f$ :

**Theorem 5.1** *Let  $h_1, h_2 \in P_{2n}$  be pair-wise independent permutations and let  $f \in F_n$  be a random function. Define  $S = S(h_1, f, f, h_2)$  (as in Definition 3.1) and let  $R \in P_{2n}$  be a random permutation. Then, for any oracle machine  $M$  (not necessarily an efficient one) that makes at most  $m$  queries:*

$$|Prob[M^{S, S^{-1}}(1^{2n}) = 1] - Prob[M^{R, R^{-1}}(1^{2n}) = 1]| \leq \frac{2m^2}{2^n} + \frac{m^2}{2^{2n}}$$

The proof follows the framework described in Section 4. The set  $BAD(h_1, h_2)$  (Definition 3.4) is replaced with the set  $BAD_1(h_1, h_2)$  defined to be:

*The set of all possible and consistent  $M$ -transcripts,  $\sigma = \{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$ , satisfying that there exist  $1 \leq i < j \leq m$  such that either  $h_1(x_i)|_R = h_1(x_j)|_R$  or  $h_2(y_i)|_L = h_2(y_j)|_L$  (as before), or there exist  $1 \leq i, j \leq m$  such that  $h_1(x_i)|_R = h_2(y_j)|_L$ .*

In order to apply Theorem 4.2, it is enough to note that by this definition we get that for any possible and consistent  $M$ -transcripts,  $\sigma$ , both  $Prob_S[T_S = \sigma \mid \sigma \notin BAD_1(h_1, h_2)] = 2^{-2nm}$  (hence, it is a proper definition according to the framework) and  $Prob_{h_1, h_2}[\sigma \in BAD_1(h_1, h_2)] < \frac{2m^2}{2^n}$ .

## 5.2 Relaxing the Pair-Wise Independence Requirement

One might interpret the construction of Section 3 in the following way: given the task of constructing *efficient* pseudo-random permutations it is enough to concentrate on the efficient construction of pseudo-random *functions*. The assumption that supports such a claim is that the computation of pseudo-random functions is much more expensive than the computation of pair-wise independent permutations. Therefore, computing the value of the pseudo-random permutation (that is constructed in Section 3) on any input of  $2n$  bits is essentially equivalent to two invocations of a pseudo-random function with  $n$ -bit inputs. In this section we show that we can use even weaker permutations instead of the pair-wise independent ones – resulting in an even more efficient construction of pseudo-random permutations.

As mentioned in Section 4, the only place in Section 3 we use the fact that  $h_1$  and  $h_2$  are pairwise independent permutations is in the proof of Proposition 3.4. In fact, the exact requirement on  $h_1$  and  $h_2$  we use is that for every  $x \neq y$ :

$$\text{Prob}_{h_1}[h_1(x)|_R = h_1(y)|_R] \leq 2^{-n} \text{ and } \text{Prob}_{h_2}[h_2(x)|_L = h_2(y)|_L] \leq 2^{-n}.$$

Furthermore, we can replace  $2^{-n}$  with any  $\epsilon \geq 2^{-n}$  and still get a construction of pseudo-random permutations (with somewhat larger distinguishing probability). Consider, for example, the revised statement of Theorem 3.2:

**Theorem 5.2** *Let  $H^1$  and  $H^2$  be distributions of permutations in  $P_{2n}$  such that for every  $2n$ -bit strings  $x \neq y$ :*

$$\text{Prob}_{h_1 \in H^1}[h_1(x)|_R = h_1(y)|_R] \leq \epsilon \text{ and } \text{Prob}_{h_2 \in H^2}[h_2(x)|_L = h_2(y)|_L] \leq \epsilon$$

*Let  $h_1$  be distributed according to  $H^1$ ,  $h_2$  distributed according to  $H^2$  and let  $f_1, f_2 \in F_n$  be random functions. Define  $S = S(h_1, f_1, f_2, h_2)$  (as in Definition 3.1) and let  $R \in P_{2n}$  be a random permutation. Then, for any oracle machine  $M$  (not necessarily an efficient one) that makes at most  $m$  queries:*

$$|\text{Prob}[M^{S, S^{-1}}(1^{2n}) = 1] - \text{Prob}[M^{R, R^{-1}}(1^{2n}) = 1]| < m^2 \cdot \epsilon + \frac{m^2}{2^{2n}}$$

The proof follows the framework described in Section 4. This time the definition of  $\text{BAD}(h_1, h_2)$  stays unchanged and, in order to apply Theorem 4.2, we only need to note that for any possible and consistent  $M$ -transcript  $\sigma$ ,  $\text{Prob}_{h_1, h_2}[\sigma \in \text{BAD}(h_1, h_2)] < m^2 \cdot \epsilon$ .

The conditions on  $H^1$  and  $H^2$  in Theorem 5.2 are somewhat nonstandard (since the requirements are on half the bits of the output). Nevertheless, these conditions are satisfied by more traditional requirements on function-families. In particular, we can take both  $H^1$  and  $H^2$  to be a distribution of permutations,  $H$ , such that for every  $x \neq y$  and every  $z$  ( $x, y, z \in I_{2n}$ ),

$$\text{Prob}_{h \in H}[h(x) \oplus h(y) = z] \leq (2^n - 1)^{-1} \cdot \epsilon$$

Such a distribution (for  $\epsilon = (2^n + 1)^{-1}$ ) is  $h_a(x) \stackrel{\text{def}}{=} a \cdot x$  where  $a$  is uniform in  $I_{2n} \setminus \{0\}$  and the multiplication is in  $GF(2^{2n})$ .

Another way to construct  $H^1$  and  $H^2$  is by using Feistel permutations with  $\epsilon$ -AXU<sub>2</sub> functions. A distribution on  $I_n \mapsto I_n$  function,  $H$ , is  $\epsilon$ -AXU<sub>2</sub> if for every  $x \neq y$  and every  $z$  ( $x, y, z \in I_n$ ),

$$\text{Prob}_{h \in H}[h(x) \oplus h(y) = z] \leq \epsilon$$

The concept of  $\epsilon$ -AXU<sub>2</sub> functions was originally defined by Carter and Wegman [13]; We use the terminology of Rogaway [41]. Let  $H$  be a distribution of  $\epsilon$ -AXU<sub>2</sub> functions on  $n$  bits strings then, we can define  $H^1$  to be  $\{D_h\}_{h \in H}$  and  $H^2$  to be  $\{D_h^{-1}\}_{h \in H}$ . The reason is that for every two *different*  $2n$ -bit strings  $x = (L^1, R^1)$  and  $y = (L^2, R^2)$  and every function  $h \in F_n$  we have by definition that:

$$D_h(x)|_R = D_h(y)|_R \iff h(R^1) \oplus h(R^2) = L^1 \oplus L^2$$

If  $R^1 = R^2$  then  $L^1 \neq L^2$  and therefore  $D_h(x)|_R \neq D_h(y)|_R$  otherwise, by the definition of  $\epsilon$ -AXU<sub>2</sub> functions:

$$\text{Prob}_{h \in H}[D_h(x)|_R = D_h(y)|_R] = \text{Prob}_{h \in H}[h(R^1) \oplus h(R^2) = L^1 \oplus L^2] \leq \epsilon$$

Thus,  $H^1$  satisfies its requirement and similarly for  $H^2$ .

By using Feistel permutations to construct  $H^1$  and  $H^2$  we get the original LR-Construction as a special case (since a random function is in particular  $2^{-n}$ -AXU<sub>2</sub>). Thus, the proof of security in Section 3 also holds for the original LR-Construction. The idea of using  $\epsilon$ -AXU<sub>2</sub> functions instead of pseudo-random functions for the first round of the LR-Construction was previously suggested by Lucks [28].

Another advantage of this approach is that it allows us to use many efficient constructions of function families. An example of efficient  $2^{-n}$ -AXU<sub>2</sub> functions are Vazirani’s “shift”-family [49]. A key of such a function is a uniformly chosen string  $a \in I_{2n-1}$  and the  $j$ ’s bit of  $f_a(x)$  ( $1 \leq j \leq n$ ) is defined to be  $\sum_{i=1}^n x_i a_{j+i-1} \bmod 2$ .

A substantial amount of research [13, 20, 24, 41, 48, 50] deals with the construction of *efficient* hash functions. This line of work contains constructions that obey weaker definitions on function families than pair-wise independence and in particular contains constructions of  $\epsilon$ -AXU<sub>2</sub> functions. Unfortunately, these functions were designed to be especially efficient when their output is substantially smaller than their input (since, they were mainly brought up in the context of authentication) which is not true in our case (but is relevant in Section 7). An additional objective is to reduce the size of the family of hash functions (e.g., [19, 24]). In our setting the purpose of this is to reduce the key-size of the pseudo-random permutations.

## 6 Reducing the Distinguishing Probability

There are various circumstances where it is desirable to have a pseudo-random permutation on relatively *few* bits (say 128). This is especially true when we want to minimize the size of the hardware-circuit that implements the permutation or the communication bandwidth with the (hardware or software) component that computes the permutation.

Let  $F$  be a pseudo-random permutation on  $\ell$  bits (note that  $n = \ell/2$  in Section 3) constructed from truly random functions (on  $\ell/2$  bits) using the LR-Construction. As shown by Patarin [36],  $F$  can be distinguished (with constant probability) from a random permutation using  $O(2^{\ell/4})$  queries (which means that the analysis of the LR-Construction, where the distinguishing probability for  $m$  queries is  $O(\frac{m^2}{2^{\ell/2}})$ , is tight). Therefore, the LR-Construction on  $\ell$  bits can only be used if  $2^{\ell/4}$  is large enough to bound the number of queries in the attack on the block cipher.

In this section, a simple generalization of the construction of Section 3 is presented. Using this construction, the adversary’s probability of distinguishing between the pseudo-random and random permutations can be reduced to roughly  $\frac{t}{2} \cdot \frac{m^2}{2^{(1-1/t)\ell}}$  for every integer  $2 \leq t \leq \ell$  (for  $t = 2$  we get the original construction). To achieve this security  $t + 2$  permutations are composed. The initial and final are pair-wise independent permutations, the rest are (generalized) Feistel permutations defined by  $I_{(1-1/t)\ell} \mapsto I_{\ell/t}$  random (or pseudo-random) functions (see Figure 3 for an illustration).

Patarin [38] shows that if we take six rounds of the LR-Construction (instead of three or four), then the resulting permutation cannot be distinguished from a random permutation with advantage better than  $\frac{5m^3}{2^\ell}$  (improving [36]). This means that distinguishing the six-round construction from a truly random permutation (with constant probability) requires at least  $\Omega(2^{\ell/3})$  queries. The bound we achieve in this section ( $\Omega(2^{(1-1/t)\ell/2})$ ) is better (for any  $t \geq 4$ ). Note that our construction uses pseudo-random functions with larger input size, which might be a disadvantage for some applications. Aiello and Venkatesan [1] show a construction of pseudo-random *functions* on  $\ell$  bits from pseudo-random functions on  $\ell/2$  bits. When using truly random functions in their construction, distinguishing the function they get from a truly random function (with constant probability) requires  $\Omega(2^{\ell/2})$  queries.

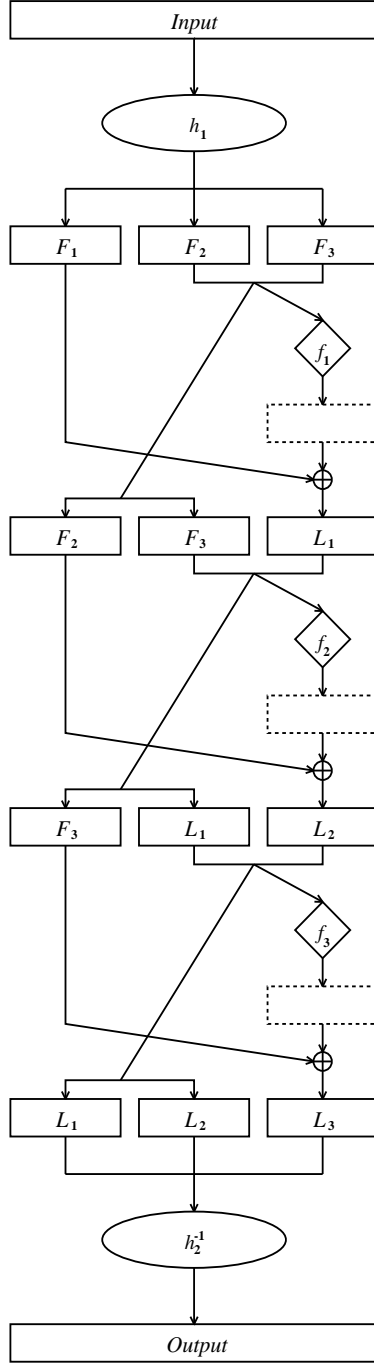


Figure 3: Construction of strong pseudo-random permutations with reduced distinguishing probability using  $t + 2$  rounds (here  $t = 3$ ). Recall,  $f_i : I_{(1-1/t)\ell} \mapsto I_{\ell/t}$  (here  $f_i : I_{2\ell/3} \mapsto I_{\ell/3}$ ).

In order to describe our generalized constructions we first extend Feistel permutations to deal with the case where the underlying functions have arbitrary input and output lengths (instead of length preserving functions as in Definition 2.1). We note that using such “unbalanced” Feistel permutations was previously suggested in [5, 28, 46].

**Definition 6.1 (Generalized Feistel Permutations)** For any integers  $0 < s < \ell$  and any function  $f : I_{\ell-s} \mapsto I_s$  let  $D_f \in P_\ell$  be the permutation defined by  $D_f(L, R) \stackrel{\text{def}}{=} (R, L \oplus f(R))$ , where  $|L| = s$  and  $|R| = \ell - s$  ( $D_f$  is defined for any function  $f$  since one can read  $\forall s, \forall(\ell - s)$ ).

We can now define the revised construction and consider its security. These are simple generalizations of the construction in Section 3 and of its proof of security.

**Definition 6.2 ( $t + 2$ -Round Construction)** For any integers  $2 \leq t \leq \ell$ , let  $s$  and  $r$  be integers such that  $\ell = s \cdot t + r$  (where  $r < t$ ). For any  $h_1, h_2 \in P_\ell$ ,  $f_1, f_2, \dots, f_r : I_{\ell-s-1} \mapsto I_{s+1}$  and  $f_{r+1}, \dots, f_t : I_{\ell-s} \mapsto I_s$  define

$$W(h_1, f_1, f_2, \dots, f_t) \stackrel{\text{def}}{=} D_{f_t} \circ D_{f_{t-1}} \circ \dots \circ D_{f_1} \circ h_1$$

and

$$S(h_1, f_1, f_2, \dots, f_t, h_2) \stackrel{\text{def}}{=} h_2^{-1} \circ D_{f_t} \circ D_{f_{t-1}} \circ \dots \circ D_{f_1} \circ h_1$$

(We get the constriction of Definition 3.1 by choosing  $t = 2$ ,  $s = \ell/2$  and  $r = 0$ .)

**Theorem 6.1** Let  $W$  and  $S$  be as in Definition 6.2, where  $h_1$  &  $h_2$  are pair-wise independent permutations and  $f_1, f_2, \dots, f_t$  are pseudo-random functions ( $t$  is allowed to be a function of  $\ell$ );  $h_1, h_2$  and  $f_1, f_2, \dots, f_t$  are independently chosen. Then,  $W$  is a pseudo-random permutation and  $S$  a strong pseudo-random permutation.

Furthermore, assume that no efficient oracle-machine that makes at most  $m = m(\ell)$  queries, distinguishes between the pseudo-random functions and random functions with advantage  $\epsilon = \epsilon(n)$ . Then, no efficient oracle-machine that makes at most  $m$  queries to  $W$  (resp.  $S$  and  $S^{-1}$ ) distinguishes  $W$  (resp.  $S$ ) from a random permutation with advantage  $t \cdot \epsilon + \frac{t}{2} \cdot \frac{m^2}{2^{\ell - \lceil \ell/t \rceil}} + \frac{m^2}{2^\ell}$

In case the middle functions are truly random this reduces to:

**Theorem 6.2** Let  $S$  be as in Definition 6.2, where  $h_1$  and  $h_2$  are pair-wise independent permutations and  $f_1, f_2, \dots, f_t$  are random functions and let  $R \in P_\ell$  be a random permutation. Then, for any oracle machine  $M$  (not necessarily an efficient one) that makes at most  $m$  queries:

$$|\text{Prob}[M^{S, S^{-1}}(1^\ell) = 1] - \text{Prob}[M^{R, R^{-1}}(1^\ell) = 1]| \leq \frac{t}{2} \cdot \frac{m^2}{2^{\ell - \lceil \ell/t \rceil}} + \frac{m^2}{2^\ell}$$

The proof of Theorem 6.2 follows the framework described in Section 4. Assume for simplicity that  $\ell = s \cdot t$ , the set  $\text{BAD}(h_1, h_2)$  (Definition 3.4) is replaced with the set  $\text{BAD}_2(h_1, h_2)$  defined to be:

The set of all possible and consistent  $M$ -transcripts,  $\sigma = \{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$ , satisfying that there exist  $1 \leq i < j \leq m$  and  $1 \leq k \leq t$  such that

$$(F_i^{k+1}, \dots, F_i^t, L_i^1, \dots, L_i^{k-1}) = (F_j^{k+1}, \dots, F_j^t, L_j^1, \dots, L_j^{k-1}),$$

where  $(F_i^1, F_i^2, \dots, F_i^t) = h_1(x_i)$  and  $(L_i^1, L_i^2, \dots, L_i^t) = h_2(y_i)$  ( $|F_i^1| = |F_i^2| = \dots = |F_i^t| = |L_i^1| = |L_i^2| = \dots = |L_i^t| = s$ ).

This guarantees that for any possible and consistent  $M$ -transcript  $\sigma$  we have that  $\text{Prob}_S[T_S = \sigma \mid \sigma \notin \text{BAD}_2(h_1, h_2)] = 2^{-lm}$  (and hence, it is a proper definition according to the framework). The reason is that, under the notations above,

$$\forall i, y_i = S(x_i) \iff \forall 1 \leq i \leq m, \forall 1 \leq k \leq t, f_k(F_i^{k+1}, \dots, F_i^t, L_i^1, \dots, L_i^{k-1}) = F_i^k \oplus L_i^k$$

Therefore, given any specific choice of  $h_1$  and  $h_2$  (in the definition of  $S$ ) such that  $\sigma \notin \text{BAD}_2(h_1, h_2)$  the event  $T_S = \sigma$  is composed of  $m \cdot t$  independent events each of which has probability  $2^{-s}$  to happen. In order to apply Theorem 4.2, it remains to note that for any such  $\sigma$  we have that

$$\text{Prob}_{h_1, h_2}[\sigma \in \text{BAD}_2(h_1, h_2)] < t \cdot \binom{m}{2} \cdot 2^{-(\ell - \lceil \ell/t \rceil)} < \frac{t}{2} \cdot \frac{m^2}{2^{\ell - \lceil \ell/t \rceil}}$$

**Remark 6.1** *The construction of this section achieves a substantial improvement in security over the construction in Section 3 even for a small constant  $t > 2$  (that is, with a few additional applications of the pseudo-random functions). Nevertheless, it might be useful for some applications to take a larger value of  $t$ . Choosing  $t = \ell$  reduces the advantage the distinguisher may achieve to roughly  $\frac{\ell \cdot m^2}{2^\ell}$ .*

## 7 SPPE on Large Blocks Using PFE or PPE on Small Blocks

Consider the application of pseudo-random permutations to encryption, i.e., using  $f(M)$  in order to encrypt a message  $M$ , where  $f$  is a pseudo-random permutation. Assume also that we want to use DES for this purpose. We now have the following problem: while DES works on fixed and relatively small length strings, we need a permutation on  $|M|$ -bit long strings, where the length of the message,  $|M|$ , may be large and may vary between different messages.

This problem is not restricted to the usage of DES (though the fact that DES was designed for hardware implementation contributes to it). Usually, a direct construction of pseudo-random permutations or pseudo-random functions (if we want to employ the LR-Construction) with large input size is expensive. Therefore, we would like a way to construct pseudo-random permutations (or functions) on *large blocks* from pseudo-random permutations (or functions) on *small blocks*.

Several such constructions were suggested in the context of DES (see e.g. [11] for the different modes of operation for DES). The simplest, known as the electronic codebook mode (ECB-mode), is to divide the input into sub-blocks and to apply the pseudo-random permutation on each sub-block separately. This solution suffers from the obvious drawback that every sub-block of output solely depends on a single sub-block of input (and, in particular, the permutation on the complete input is not pseudo-random). This may leak information about the message being encrypted (see further discussion in Section 7.2).

In this section we consider a generalization of the construction of Section 3 that uses pseudo-random functions (or permutations) on *small* blocks to construct strong pseudo-random permutations on *large* blocks. The idea is as follows: apply a pair-wise independent permutation on the entire input, divide the value you get into sub-blocks and apply two rounds of Feistel-permutations (or one round of a pseudo-random permutation) on each sub-block separately, finally, apply a second pair-wise independent permutation on the entire value you get (see Figure 4 for an illustration).

This solution resembles the electronic codebook mode and is almost as simple. But here, the security we achieve is relative to a random permutation applied on *the entire message* and not on each sub-block separately. As is the case with the electronic codebook mode, the construction is highly suitable for parallel implementation.

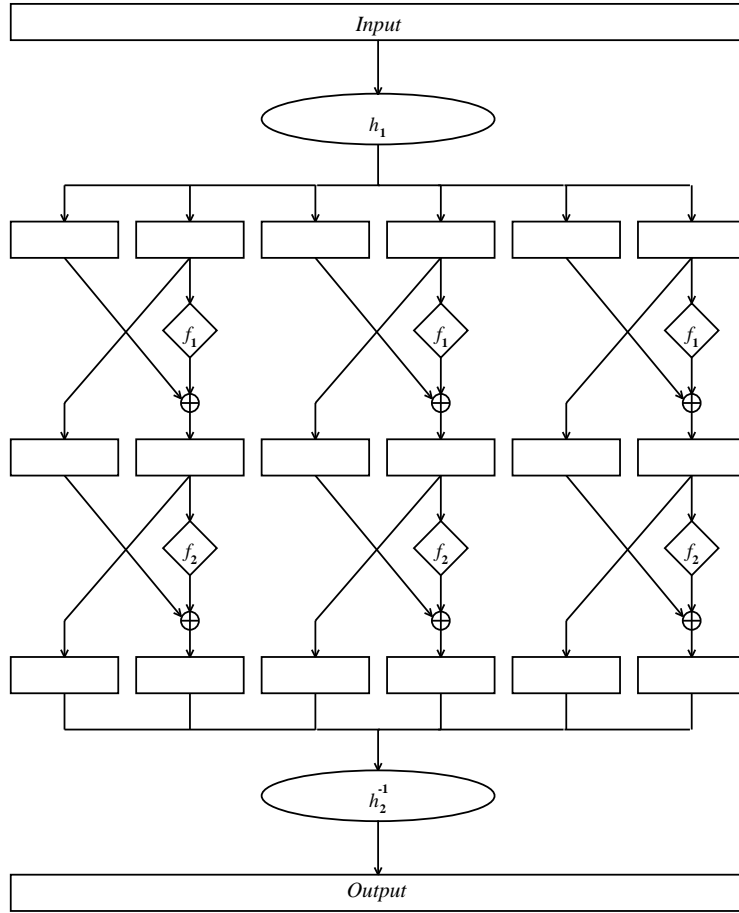


Figure 4: Construction of a strong pseudo-random permutation on many (six in this case) blocks from a pseudo-random function on a single block.



For simplicity, we only describe the construction using truly random functions (or a truly random permutation). The analysis of the construction when pseudo-random functions are used follows easily. In addition, we restrict our attention to the construction of *strong* pseudo-random permutations.

**Definition 7.1** For any two integers  $b$  and  $s$ , for any function  $g \in F_s$  let  $g^{\times b} \in F_{b \cdot s}$  be the function defined by:

$$g^{\times b}(x_1, x_2, \dots, x_b) \stackrel{\text{def}}{=} (g(x_1), g(x_2), \dots, g(x_b))$$

For any  $f_1, f_2 \in F_n$  and  $h_1, h_2 \in P_{2nb}$ , define:

$$S(h_1, f_1, f_2, h_2) \stackrel{\text{def}}{=} h_2^{-1} \circ D_{f_2}^{\times b} \circ D_{f_1}^{\times b} \circ h_1$$

For any  $p \in P_{2n}$  and  $h_1, h_2 \in P_{2nb}$ , define:

$$\hat{S}(h_1, p, h_2) \stackrel{\text{def}}{=} h_2^{-1} \circ p^{\times b} \circ h_1$$

**Theorem 7.1** Let  $h_1, h_2 \in P_{2nb}$  be pair-wise independent permutations, let  $f_1, f_2 \in F_n$  be random functions and  $p \in P_{2n}$  a random permutation. Define  $S = S(h_1, f_1, f_2, h_2)$  and  $\hat{S} = \hat{S}(h_1, p, h_2)$  (as in Definition 7.1) and let  $R \in P_{2nb}$  be a random permutation. Then, for any oracle machine  $M$  (not necessarily an efficient one) that makes at most  $m$  queries:

$$|\text{Prob}[M^{S, S^{-1}}(1^{2nb}) = 1] - \text{Prob}[M^{R, R^{-1}}(1^{2nb}) = 1]| \leq \frac{m^2 \cdot b^2}{2^n} + \frac{m^2}{2^{2nb}}$$

and

$$|\text{Prob}[M^{\hat{S}, \hat{S}^{-1}}(1^{2nb}) = 1] - \text{Prob}[M^{R, R^{-1}}(1^{2nb}) = 1]| \leq \frac{m^2 \cdot b^2}{2^{2n-1}}$$

The proof of Theorem 7.1 for  $S$  follows the framework described in Section 4. The set  $\text{BAD}(h_1, h_2)$  (Definition 3.4) is replaced with the set  $\text{BAD}_3(h_1, h_2)$  defined to be:

*The set of all possible and consistent  $M$ -transcripts,  $\sigma = \{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$ , such that either there are two equal values in  $\{F_i^{2j}\}_{1 \leq i \leq m, 1 \leq j \leq b}$  or there are two equal values in  $\{L_i^{2j-1}\}_{1 \leq i \leq m, 1 \leq j \leq b}$ , where  $(F_i^1, F_i^2, \dots, F_i^{2b}) = h_1(x_i)$  and  $(L_i^1, L_i^2, \dots, L_i^{2b}) = h_2(y_i)$  ( $|F_i^1| = |F_i^2| = \dots = |F_i^{2b}| = |L_i^1| = |L_i^2| = \dots = |L_i^{2b}| = n$ ).*

This guarantees that for any possible and consistent  $M$ -transcript  $\sigma$  we have that

$$\text{Prob}_S[T_S = \sigma \mid \sigma \notin \text{BAD}_3(h_1, h_2)] = 2^{-2n \cdot b \cdot m}$$

(and hence, it is a proper definition according to the framework). The reason is that, under the notations above,

$$\forall i, y_i = S(x_i) \iff \forall 1 \leq i \leq m, \forall 1 \leq j \leq b, f_1(F_i^{2j}) = F_i^{2j-1} \oplus L_i^{2j-1} \text{ and } f_2(L_i^{2j-1}) = F_i^{2j} \oplus L_i^{2j}$$

Therefore, given any specific choice of  $h_1$  and  $h_2$  (in the definition of  $S$ ) such that  $\sigma \notin \text{BAD}_3(h_1, h_2)$  the event  $T_S = \sigma$  is composed of  $2m \cdot b$  independent events each of which has probability  $2^{-n}$  to happen. In order to apply Theorem 4.2, it remains to note that for any such  $\sigma$  we have that

$$\text{Prob}_{h_1, h_2}[\sigma \in \text{BAD}_3(h_1, h_2)] \leq 2 \cdot \binom{m \cdot b}{2} \cdot 2^{-n} < \frac{m^2 \cdot b^2}{2^n}$$

The proof of Theorem 7.1 for  $\hat{S}$  slightly deviates from the framework described in Section 4 (providing yet another evidence to the claim that “nobody is perfect”). The set  $\text{BAD}(h_1, h_2)$  (Definition 3.4) is replaced with the set  $\text{BAD}_4(h_1, h_2)$  defined to be:

The set of all possible and consistent  $M$ -transcripts,  $\sigma = \{\langle x_1, y_1 \rangle, \dots, \langle x_m, y_m \rangle\}$ , such that either there are two equal values in  $\{F_i^j\}_{1 \leq i \leq m, 1 \leq j \leq b}$  or there are two equal values in  $\{L_i^j\}_{1 \leq i \leq m, 1 \leq j \leq b}$ , where  $(F_i^1, F_i^2, \dots, F_i^b) = h_1(x_i)$  and  $(L_i^1, L_i^2, \dots, L_i^b) = h_2(y_i)$  ( $|F_i^1| = |F_i^2| = \dots = |F_i^b| = |L_i^1| = |L_i^2| = \dots = |L_i^b| = 2n$ ).

Now we have that for any possible and consistent  $M$ -transcript  $\sigma$ ,

$$\text{Prob}_{h_1, h_2}[\sigma \in \text{BAD}_4(h_1, h_2)] \leq 2 \cdot \binom{m \cdot b}{2} \cdot 2^{-2n} < \frac{m^2 \cdot b^2}{2^{2n}}$$

but now for any such  $\sigma$ ,

$$\text{Prob}_{\hat{S}}[T_{\hat{S}} = \sigma \mid \sigma \notin \text{BAD}_4(h_1, h_2)] = \frac{2^{2n}!}{(2^{2n} - m \cdot b)!}$$

instead of  $2^{-2n \cdot b \cdot m}$  as “required” by the framework. However, the difference in probabilities is rather small which result in only a minor deviation from the proof of Theorem 3.2.

## 7.1 Relaxing the Construction

As in Section 5.2 we would like to reduce the requirements from  $h_1$  and  $h_2$  in Theorem 7.1. Our main motivation in doing so is to *decrease the key size* of the pseudo-random permutations. We would like the key-size to be of order  $n$  – the size of the small sub-blocks and *not* of order  $2nb$  – the size of the complete input (in some cases we may allow a small dependence on  $b$ ).

We sketch a way to redefine the distributions on  $h_1$  and  $h_2$  in the definition of  $\hat{S}$  (almost the same ideas apply to the definition of  $S$ ). The requirement these distributions have to obey is that for any possible and consistent  $M$ -transcript  $\sigma$  we have that  $\text{Prob}_{h_1, h_2}[\sigma \in \text{BAD}_4(h_1, h_2)]$  is “small”. We use the following notation: For any  $2n \cdot b$ -bit string  $z = (z_1, z_2, \dots, z_b)$  (such that  $\forall j, |z_j| = 2n$ ) and for all  $1 \leq i \leq b$ , denote by  $z_{|i}$  the substring  $z_i$  (the  $i$ th substring of  $z$ ). The requirement above can be achieved by sampling  $h_1$  and  $h_2$  according to a permutation distribution  $H$  such that for some small  $\epsilon \geq 2^{-2n}$  we have that:

1. For any  $2n \cdot b$ -bit string  $x$ ,  $\forall 1 \leq i < j \leq b$ ,  $\text{Prob}_{h \in H}[h(x)_{|i} = h(x)_{|j}] \leq \epsilon$  and
2. For any  $2n \cdot b$ -bit strings  $x \neq x'$ ,  $\forall 1 \leq i, j \leq b$ ,  $\text{Prob}_{h \in H}[h(x)_{|i} = h(x')_{|j}] \leq \epsilon$ .

We start by defining a permutation distribution  $H'$  that almost achieves this: A permutation  $h' = h'_{u_1, u_2}$  sampled from  $H'$  is defined by two  $\epsilon'$ -AXU<sub>2</sub> functions,  $u_1 : I_{2n} \mapsto I_{2n}$  and  $u_2 : I_{\lceil \log b \rceil} \mapsto I_{2n}$  (see definition of  $\epsilon$ -AXU<sub>2</sub> functions in Section 5.2). For any  $z = (z_1, z_2, \dots, z_b)$  (such that  $\forall j, |z_j| = 2n$ ),

$$h'_{u_1, u_2}(z) \stackrel{\text{def}}{=} (z_1 \oplus u_1(z_b) \oplus u_2(1), z_2 \oplus u_1(z_b) \oplus u_2(2), \dots, z_{b-1} \oplus u_1(z_b) \oplus u_2(b-1), z_b \oplus u_2(b))$$

It is not hard to verify that:

- 1' For any  $2n \cdot b$ -bit string  $x$ ,  $\forall 1 \leq i < j \leq b$ ,  $\text{Prob}_{h' \in H'}[h'(x)_{|i} = h'(x)_{|j}] \leq \epsilon'$  and
- 2' For any  $2n \cdot b$ -bit strings  $x \neq x'$  such that  $x_{|b} \neq x'_{|b}$  and for all  $1 \leq i, j \leq b$ ,  $\text{Prob}_{h' \in H'}[h'(x)_{|i} = h'(x')_{|j}] \leq \epsilon'$ .

In order to eliminate the additional requirement in (2') that  $x|_b \neq x'|_b$ , we define the permutation distribution  $H$  such that a permutation  $h$  sampled from  $H$  is defined to be  $h' \circ D_g$  (see Definition 6.1), where  $h'$  is sampled according to  $H'$  and  $g : I_{2n \cdot (b-1)} \mapsto I_{2n}$  is a  $\epsilon'$ -AXU<sub>2</sub> function. Using (1') and (2') and the fact that for any  $2n \cdot b$ -bit strings  $x \neq x'$ :

$$\text{Prob}_g[D_g(x)|_b = D_g(x')|_b] \leq \epsilon'$$

we get that  $H$  satisfies (1) and (2) for  $\epsilon = 2\epsilon'$ .

Notice that the computation of a function  $h \in H$  is essentially equivalent to one computation of an  $\epsilon$ -AXU<sub>2</sub> function,  $g : I_{2n \cdot (b-1)} \mapsto I_{2n}$ , and a few additional XOR operations per block. Using efficient constructions of  $\epsilon$ -AXU<sub>2</sub> functions [13, 20, 24, 41, 48, 50] we get an efficient function  $h$ . Krawczyk [24] shows a construction of  $\frac{m+\ell}{2^t-1}$ -AXU<sub>2</sub> functions from  $m$  bits to  $\ell$  bits with  $\ell$  key-bits. Using these functions we can achieve the desired goal of reducing the key-size of  $h$  to  $O(n)$  bits.

## 7.2 Related Work

The construction presented in this section is certainly not the only solution to the problem at hand. We refer in brief to some additional solutions:

As mentioned above, DES modes of operation were suggested as a way of encrypting long messages. However, none of these modes constitutes a construction of a pseudo-random permutation. For instance, when using the cipher block chaining mode (CBC-mode), the encryptions of two messages with identical prefix will also have an identical prefix. Note that when the encryption of a message  $M$  is  $f(M)$ , for a pseudo-random permutation  $f$ , then the only information that is leaked on  $M$  is whether or not  $M$  is equal to previously encrypted messages. Bellare et. al. [8] show that the CBC-mode does define a construction of a pseudo-random function with small output length. They also provide a formal setting for the analysis of the security of pseudo-random functions with fixed input and output lengths. Bellare et. al. [6] consider the so called *cascade* construction of a pseudo-random function with small output length. Bellare and Rogaway [9] show how to use the CBC-mode in order to construct a pseudo-random *permutation* on large inputs (this is the only work we are aware of that explicitly refers to the problem). The work in their construction is comparable to two applications of the CBC-mode (approximately twice the work of our construction, assuming that  $h_1$  and  $h_2$  are relatively efficient). The security of all these constructions is of similar order to the security of our construction. In contrast to our construction, [6, 8, 9] are all sequential in nature.

A different approach that may be attributed in part to Carter and Wegman [50], is to define a length preserving pseudo-random function  $\tilde{F}$  as  $G \circ F \circ h$  where,  $h$  is a pair-wise independent hash function with short output,  $F$  is a length preserving pseudo-random function on short inputs and  $G$  a pseudo-random (bit) generator. It is now possible to use the LR-Construction in order to get a pseudo-random permutation on large inputs. Anderson and Biham [5] and Lucks [28] show how to directly apply similar ideas into the LR-Construction in order to get a pseudo-random permutation on large inputs. The comparison between these constructions and ours (both in terms of security and in terms of efficiency) is a bit more complicated because these constructions use an additional primitive – a pseudo-random generator. Therefore, the comparison relies on the specific parameters of the different primitives that are used (pseudo-random functions, pseudo-random generator and a hash function).

**Remark 7.1** *Our construction (as well as the other constructions described in this section) of a pseudo-random permutation on many blocks from a pseudo-random function (or permutation) on a*

single block is vulnerable to a birthday-attack on the size of a single block: If two different queries to the pseudo-random permutation on many blocks produce the same input to the pseudo-random function on a single block (at any stage in the computation) then our analysis fails. However, our construction (as well as the other constructions) reduces the problem of foiling birthday-attacks when constructing a pseudo-random permutation on many blocks to the problem of foiling birthday-attacks when constructing a pseudo-random function (or permutation) on two blocks. A solution to the latter is proposed by Aiello and Venkatesan [1].

## 8 Constructions of $k$ -Wise $\delta$ -Dependent Permutations

In this section, we summarize the connection between the various constructions of this paper and the task of obtaining  $k$ -wise  $\delta$ -dependent permutations. As mentioned in Section 5.1, Schnorr [47] suggested using the LR-Construction with truly random functions in order to get a pseudo-random generator that is secure as long as not too many bits of its output are accessible to the adversary. This idea is further treated by Maurer and Massey [30]. Maurer [29] suggested to replace the truly random functions with what he calls locally random (or almost random) functions. In the terminology of  $k$ -wise independence these ideas can be interpreted as a way of using the LR-Construction in order to obtain  $k$ -wise  $\delta$ -dependent permutations from  $k$ -wise  $\delta'$ -dependent functions (as long as  $k$  is not too large). Theorem 1 in [29] implies that

*when  $k$ -wise  $\delta'$ -dependent functions are used instead of pseudo-random functions in the LR-Construction the result is a  $k$ -wise  $\delta$ -dependent permutations for  $\delta = O(k^2/2^n + \delta')$ .*

Similar observations apply to the different constructions of our paper as discussed in this section.

**Corollary 8.1** (to Theorem 3.2) *Let  $h_1, h_2 \in P_{2n}$  be pair-wise independent permutations and let  $f_1, f_2 \in F_n$  be  $k$ -wise  $\delta'$ -dependent functions. Then,  $S = S(h_1, f_1, f_2, h_2)$  (as in Definition 3.1) is a  $k$ -wise  $\delta$ -dependent permutation for*

$$\delta \stackrel{\text{def}}{=} \frac{k^2}{2^n} + \frac{k^2}{2^{2n}} + 2\delta'$$

*Proof:* Let  $S_1, S_2 \in P_{2n}$  have the following distributions:

- $S_1 = S(h_1, g_1, f_2, h_2)$ , where  $h_1, h_2 \in P_{2n}$  are pair-wise independent,  $f_2 \in F_n$  a  $k$ -wise  $\delta'$ -dependent function and  $g_1 \in F_n$  a truly random function.
- $S_2 = S(h_1, g_1, g_2, h_2)$ , where  $h_1, h_2 \in P_{2n}$  are pair-wise independent and  $g_1, g_2 \in F_n$  are truly random functions.

and let  $R \in P_{2n}$  be a truly random permutation. It is enough to show that for every  $k$  strings of  $2n$ -bits,  $x_1, x_2, \dots, x_k$ , we have:

1.  $\| \langle S(x_1), S(x_2), \dots, S(x_k) \rangle - \langle S_1(x_1), S_1(x_2), \dots, S_1(x_k) \rangle \| \leq \delta'$
2.  $\| \langle S_1(x_1), S_1(x_2), \dots, S_1(x_k) \rangle - \langle S_2(x_1), S_2(x_2), \dots, S_2(x_k) \rangle \| \leq \delta'$
3.  $\| \langle S_2(x_1), S_2(x_2), \dots, S_2(x_k) \rangle - \langle R(x_1), R(x_2), \dots, R(x_k) \rangle \| \leq \frac{k^2}{2^n} + \frac{k^2}{2^{2n}}$

The reason (3) holds is that if we define an oracle machine  $M$  such that its  $i$ th query is always  $(+, x_i)$  and such that

$$\begin{aligned} C_M(\{\langle x_1, y_1 \rangle, \langle x_2, y_2 \rangle, \dots, \langle x_k, y_k \rangle\}) &= 1 \\ \iff \text{Prob}[\langle S_2(x_1), \dots, S_2(x_k) \rangle = \langle y_1, \dots, y_k \rangle] &< \text{Prob}[\langle R(x_1), \dots, R(x_k) \rangle = \langle y_1, \dots, y_k \rangle] \end{aligned}$$

we get by the definition of variation distance and from Theorem 3.2 that

$$\begin{aligned} &\| \langle S_2(x_1), S_2(x_2), \dots, S_2(x_k) \rangle - \langle R(x_1), R(x_2), \dots, R(x_k) \rangle \| \\ &= |\text{Prob}[M^{S_2, S_2^{-1}}(1^{2n}) = 1] - \text{Prob}[M^{R, R^{-1}}(1^{2n}) = 1]| \\ &\leq \frac{k^2}{2^n} + \frac{k^2}{2^{2n}} \end{aligned}$$

(1) and (2) hold by the definition of  $k$ -wise  $\delta'$ -dependent functions. For example, if

$$\| \langle S(x_1), S(x_2), \dots, S(x_k) \rangle - \langle S_1(x_1), S_1(x_2), \dots, S_1(x_k) \rangle \| > \delta'$$

then we can fix  $h_1, h_2 \in P_{2n}$  and  $f_2 \in F_n$  in the definition of both  $S$  and  $S_1$  such that the inequality still holds. This defines  $k$  strings of  $n$ -bits,  $z_1, z_2, \dots, z_k$ , (not necessarily all different) and a function  $V$  for which:

$$\begin{aligned} \langle S(x_1), S(x_2), \dots, S(x_k) \rangle &= V(\langle f_1(z_1), f_1(z_2), \dots, f_1(z_k) \rangle) \text{ and} \\ \langle S_1(x_1), S_1(x_2), \dots, S_1(x_k) \rangle &= V(\langle g_1(z_1), g_1(z_2), \dots, g_1(z_k) \rangle) \end{aligned}$$

We get a contradiction since for any function  $V$ :

$$\begin{aligned} &\| V(\langle f_1(z_1), f_1(z_2), \dots, f_1(z_k) \rangle) - V(\langle g_1(z_1), g_1(z_2), \dots, g_1(z_k) \rangle) \| \\ &\leq \| \langle f_1(z_1), f_1(z_2), \dots, f_1(z_k) \rangle - \langle g_1(z_1), g_1(z_2), \dots, g_1(z_k) \rangle \| \\ &\leq \delta' \end{aligned}$$

□

In a similar way we get the following two Corollaries from the constructions of Sections 6 & 7:

**Corollary 8.2** (to Theorem 6.2) *Let  $S$  be as in Definition 6.2, where  $h_1$  and  $h_2$  are pair-wise independent permutations and  $f_1, f_2, \dots, f_t$  are  $k$ -wise  $\delta'$ -dependent functions. Then,  $S$  is a  $k$ -wise  $\delta$ -dependent permutation for*

$$\delta \stackrel{\text{def}}{=} \frac{t}{2} \cdot \frac{k^2}{2^{\ell - \lceil \ell/t \rceil}} + \frac{k^2}{2^\ell} + t \cdot \delta'$$

**Corollary 8.3** (to Theorem 7.1) *Let  $h_1, h_2 \in P_{2nb}$  be pair-wise independent permutations, let  $f_1, f_2 \in F_n$  be  $b \cdot k$ -wise  $\delta'$ -dependent functions and let  $p \in P_{2n}$  be a  $b \cdot k$ -wise  $\delta'$ -dependent permutation. Define  $S = S(h_1, f_1, f_2, h_2)$  and  $\hat{S} = \hat{S}(h_1, p, h_2)$  (as in Definition 7.1). Then,  $S$  is a  $k$ -wise  $\delta$ -dependent permutation for*

$$\delta \stackrel{\text{def}}{=} \frac{k^2 \cdot b^2}{2^n} + \frac{k^2}{2^{2nb}} + 2\delta'$$

and  $\hat{S}$  is a  $k$ -wise  $\hat{\delta}$ -dependent permutation for

$$\hat{\delta} \stackrel{\text{def}}{=} \frac{k^2 \cdot b^2}{2^{2n-1}} + \delta'$$

By taking  $t = \ell$  in Corollary 8.2 we get a simple construction of a  $k$ -wise  $\delta$ -dependent permutation on  $\ell$  bits for  $\delta$  as close to  $\frac{(\ell+1) \cdot k^2}{2^\ell}$  as we wish. This construction requires  $\ell$  applications of  $k$ -wise  $\delta'$ -dependent functions from  $\ell - 1$  bits to a single bit. An interesting question is to find a simple construction of  $k$ -wise  $\delta$ -dependent permutations for an *arbitrarily small*  $\delta$  and an arbitrary  $k$ .

An “old” proposal by the first author (see [42, page 17]) is to use an “oblivious” card shuffling procedure that requires only few rounds. Oblivious here means that the location of a card after each round depends on a few random decisions. In [42] this idea is described for a specific card shuffling procedure that was suggested by Aldous and Diaconis [2]. Unfortunately, to the best of our knowledge, this procedure was never proven to give (with few rounds) an almost uniform ordering of the cards. However, we can still use this technique with different card shuffling procedures. For example, one can recursively construct such a procedure by permuting each half of the deck and merging the two halves in an almost uniform way (the merging is also performed recursively). This translates to an  $O(n^2)$  rounds procedure for a permutation on  $n$  bits (however, the construction is rather cumbersome). This proposal may become attractive given an efficient and simple “oblivious” card shuffling.

Leonard Schulman (private communication) suggested a way to generalize the algebraic construction of pair-wise independent permutations that yields 3-wise independent permutations. He’s suggestion is to use sharply 3-transitive permutation groups. A permutation group over the set  $[n] = \{1, 2, \dots, n\}$  is a subgroup of the symmetric group  $S_n$ . A permutation group  $G$  over  $[n]$  is  $k$ -transitive if for every two  $k$ -tuples  $\{a_1, \dots, a_k\}$  and  $\{b_1, \dots, b_k\}$  of distinct elements of  $[n]$  there exist a permutation  $\pi \in G$  such that  $\forall 1 \leq i \leq k, \pi(a_i) = b_i$ . A permutation group  $G$  over  $[n]$  is sharply  $k$ -transitive if for every two such tuples there exists exactly one permutation  $\pi \in G$  such that  $\forall 1 \leq i \leq k, \pi(a_i) = b_i$ . A sharply  $k$ -transitive permutation group is in particular  $k$ -wise independent and indeed the algebraic construction of pair-wise independent permutations use a sharply 2-transitive permutation group (containing all the linear permutations). Schulman suggested to use the fact that there are known constructions of sharply 3-transitive permutation groups. However, this approach cannot be generalized to larger values of  $k$ : from the classification of finite simple groups it follows that for  $k \geq 6$  there are no  $k$ -transitive groups over  $[n]$  other than the symmetric group  $S_n$  and the alternating group  $A_n$  and there are only few such groups for  $k = 4$  and  $k = 5$  (see [12, 40]). One should be careful not to interpret this as implying that for  $k \geq 4$  there are no efficient algebraic constructions of  $k$ -wise independent permutations. It is however justified to deduce that for  $k \geq 4$  any small family of  $k$ -wise independent permutations is not a permutation group (i.e. is not closed under composition and inverse).

## 9 Conclusion and Further Work

The constructions described in Sections 3 & 7 are optimal in their cryptographic work in the sense that the total number of bits on which the cryptographic functions are applied on is exactly the number of bits in the input. Therefore, it seems that in order to achieve the goal of constructing efficient block-ciphers it is sufficient to concentrate on the construction of efficient pseudo-random functions. The depth of the constructions, on the other hand, is twice the depth of the cryptographic functions. It is an interesting question whether there can be a construction of similar depth. The goal of reducing the depth is even more significant in the case of the  $t + 2$ -round construction in Section 6. A different question is finding a simple construction of  $k$ -wise  $\delta$ -dependent permutations for an *arbitrarily small*  $\delta$  and an arbitrary  $k$ . This question is discussed in Section 8.

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