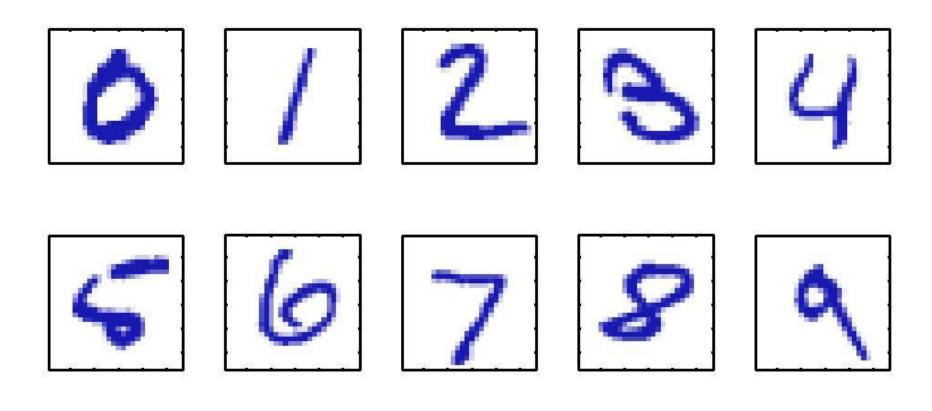
PATTERN RECOGNITION AND MACHINE LEARNING

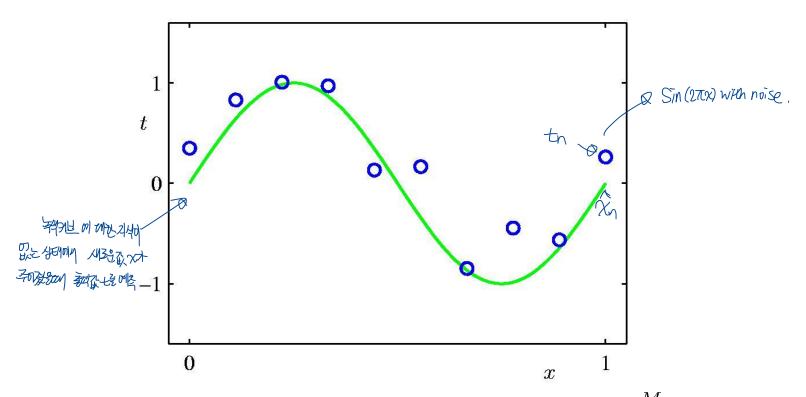
CHAPTER 1: INTRODUCTION

Example

Handwritten Digit Recognition

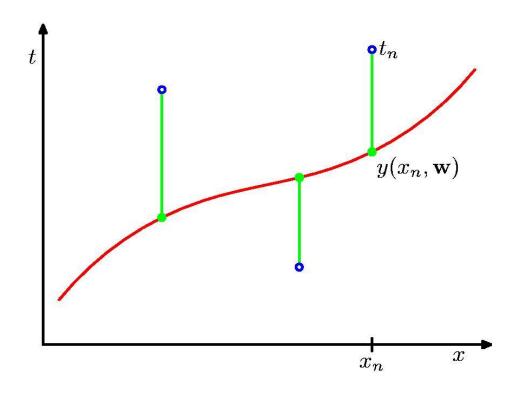


Polynomial Curve Fitting



$$y(x, \mathbf{w}) = w_0 + w_1 x + w_2 x^2 + \ldots + w_M x^M = \sum_{j=0}^{M} w_j x^j$$

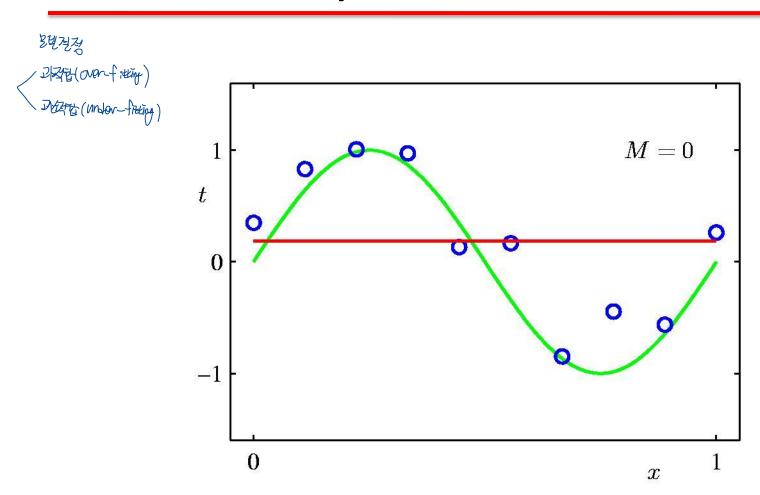
Sum-of-Squares Error Function



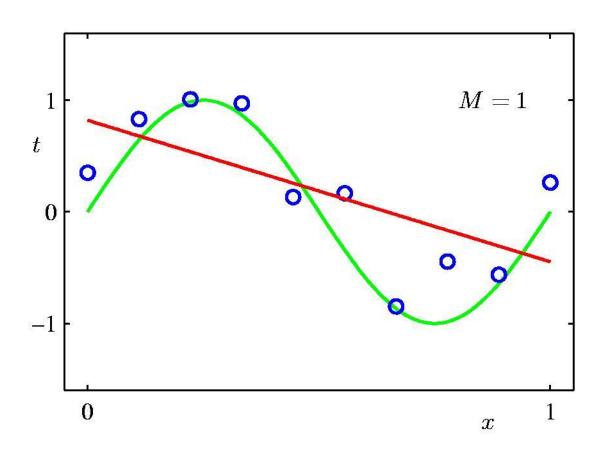
$$E(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} \{y(x_n, \mathbf{w}) - t_n\}^2$$

(P) 多 3分科 V 代码 / PB 例 WONANT 在路侧 E(M) 多好社 以是 要给现在

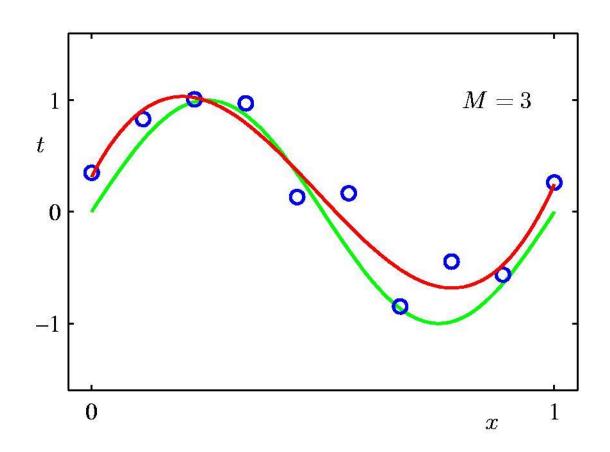
Oth Order Polynomial



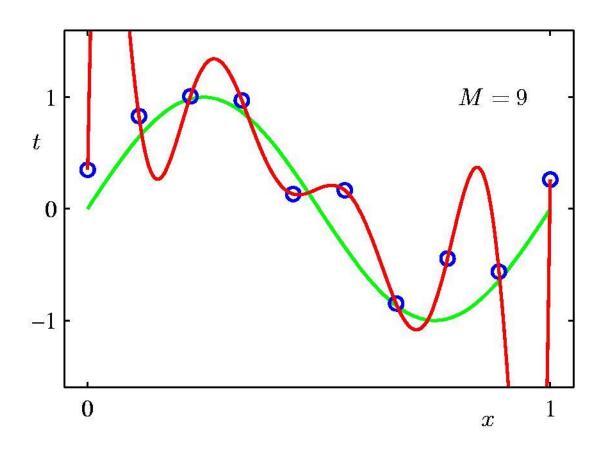
1st Order Polynomial



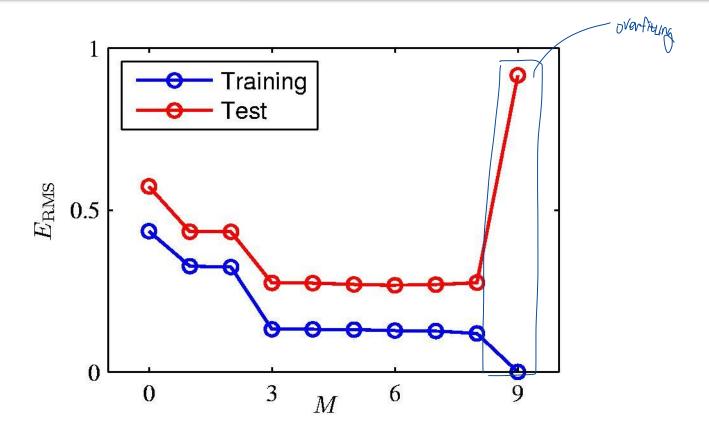
3rd Order Polynomial



9th Order Polynomial



Over-fitting



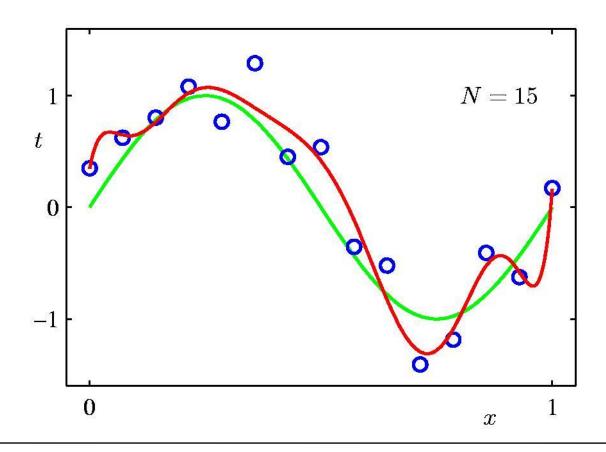
Root-Mean-Square (RMS) Error: $E_{\rm RMS} = \sqrt{2E(\mathbf{w}^\star)/N}$

Polynomial Coefficients

	M=0	M = 1	M = 3	M = 9
$\overline{w_0^{\star}}$	0.19	0.82	0.31	0.35
w_1^{\star}		-1.27	7.99	232.37
w_2^{\star}			-25.43	-5321.83
w_3^{\star}			17.37	48568.31
w_4^{\star}				-231639.30
w_5^{\star}				640042.26
w_6^{\star}				-1061800.52
w_7^{\star}				1042400.18
w_8^{\star}				-557682.99
w_9^{\star}				125201.43

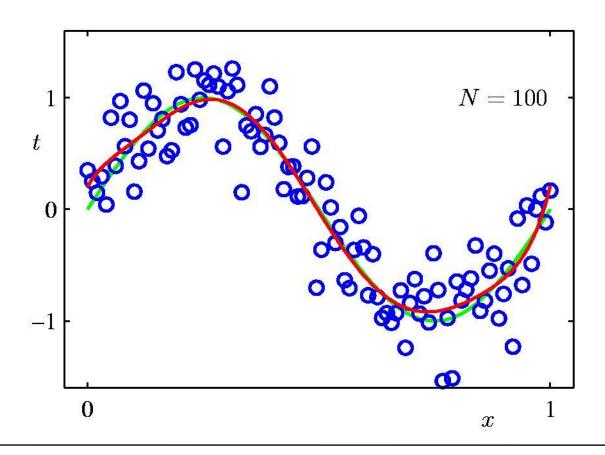
Data Set Size: N = 15

9th Order Polynomial



Data Set Size: N = 100

9th Order Polynomial with salety.



Regularization 344

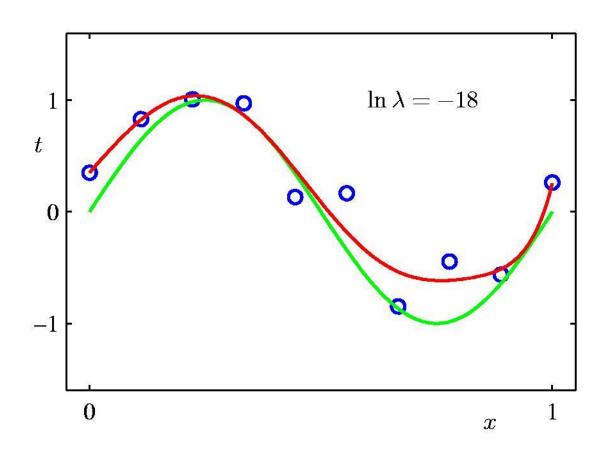
Penalize large coefficient values

$$\widetilde{E}(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} \left\{ y(x_n, \mathbf{w}) - t_n \right\}^2 + \frac{\lambda}{2} ||\mathbf{w}||^2$$

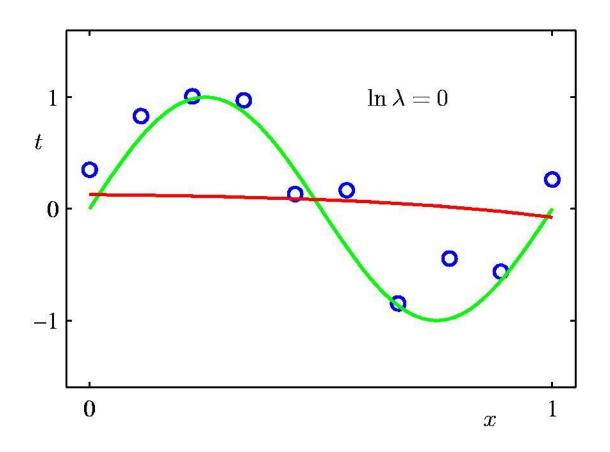
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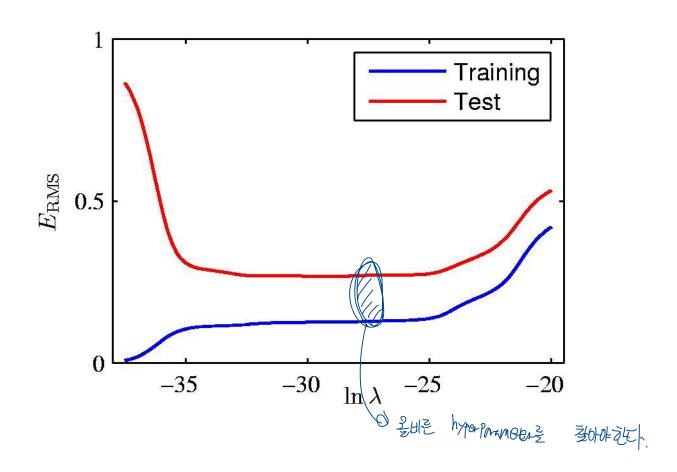
Regularization: $\ln \lambda = -18$



Regularization: $\ln \lambda = 0$



Regularization: $E_{\rm RMS}$ vs. $\ln \lambda$

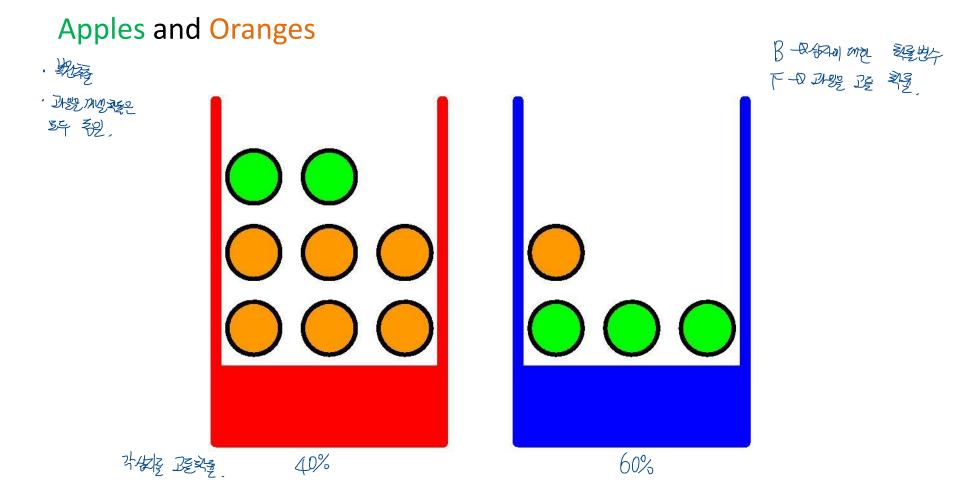


Polynomial Coefficients

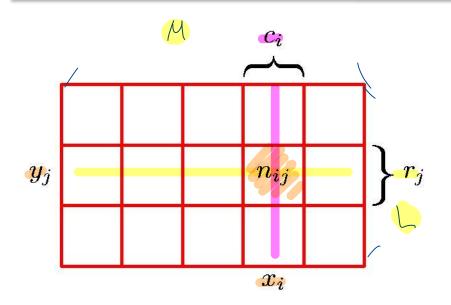
	$\ln \lambda = -\infty$	$\ln \lambda = -18$	$\ln \lambda = 0$
w_0^{\star}	0.35	0.35	0.13
w_1^{\star}	232.37	4.74	-0.05
w_2^{\star}	-5321.83	-0.77	-0.06
w_3^{\star}	48568.31	-31.97	-0.05
w_4^{\star}	-231639.30	-3.89	-0.03
w_5^{\star}	640042.26	55.28	-0.02
w_6^{\star}	-1061800.52	41.32	-0.01
w_7^{\star}	1042400.18	-45.95	-0.00
w_8^{\star}	-557682.99	-91.53	0.00
w_9^{\star}	125201.43	72.68	0.01

Probability Theory

是到给了平野的 处。



Probability Theory



Marginal Probability

$$p(X = x_i) = \frac{c_i}{N}.$$

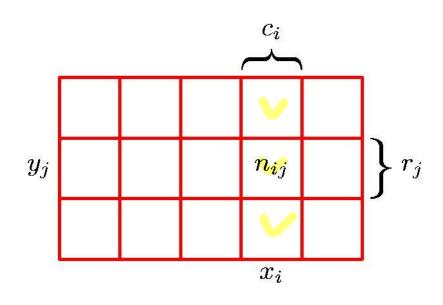
Joint Probability

$$p(X = x_i, Y = y_j) = \frac{n_{ij}}{N}$$

Conditional Probability

$$p(Y = y_j | X = x_i) = \frac{n_{ij}}{c_i}$$

Probability Theory



Sum Rule

$$r_j$$
 $p(X = x_i) = \frac{c_i}{N} = \frac{1}{N} \sum_{j=1}^{L} n_{ij}$
= $\sum_{j=1}^{L} p(X = x_i, Y = y_j)$

Product Rule

$$p(X = x_i, Y = y_j) = \frac{n_{ij}}{N} = \frac{n_{ij}}{c_i} \cdot \frac{c_i}{N}$$

$$= p(Y = y_j | X = x_i) p(X = x_i)$$

$$\frac{P(\forall i, 0, \forall i, j)}{P(\forall i, j, j)} \cdot P(\forall i, j, j)$$

The Rules of Probability

Sum Rule
$$p(X) = \sum_{Y} p(X,Y)$$

$$p(X) = \sum_{Y} p(X,Y)$$
 Product Rule
$$p(X,Y) = p(Y|X)p(X)$$

$$p(X,Y) = p(Y|X)p(X)$$

$$p(X,Y) = p(Y|X)p(X)$$

$$p(X,Y) = p(X|X)p(X)$$

Bayes' Theorem

$$p(Y|X) = \frac{p(X|Y)p(Y)}{p(X)}$$

$$p(X) = \sum_{Y} p(X|Y)p(Y)$$

$$p(X) = \underbrace{\sum_{Y} p(X|Y)p(Y)}_{p(Y|X)}$$

posterior ∝ likelihood × prior

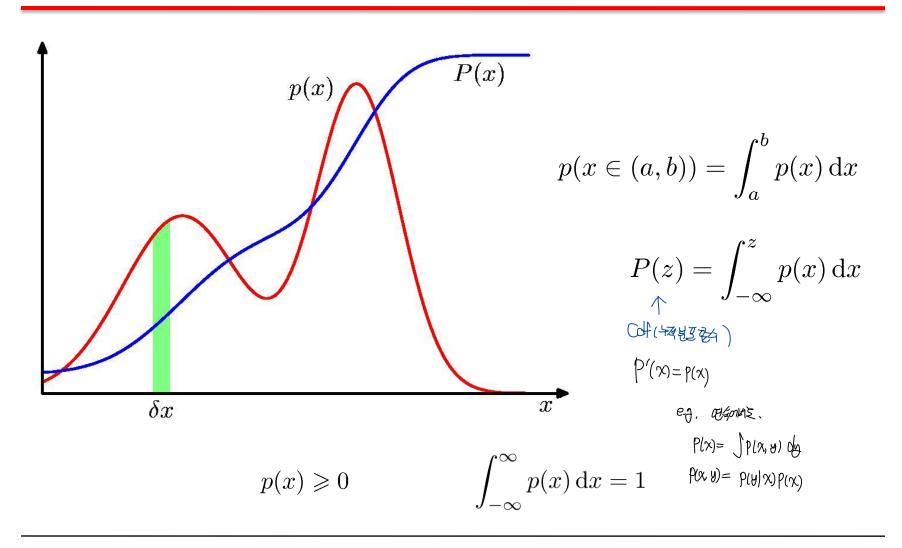
다. 독립

P(X,Y)=P(X)P(Y)

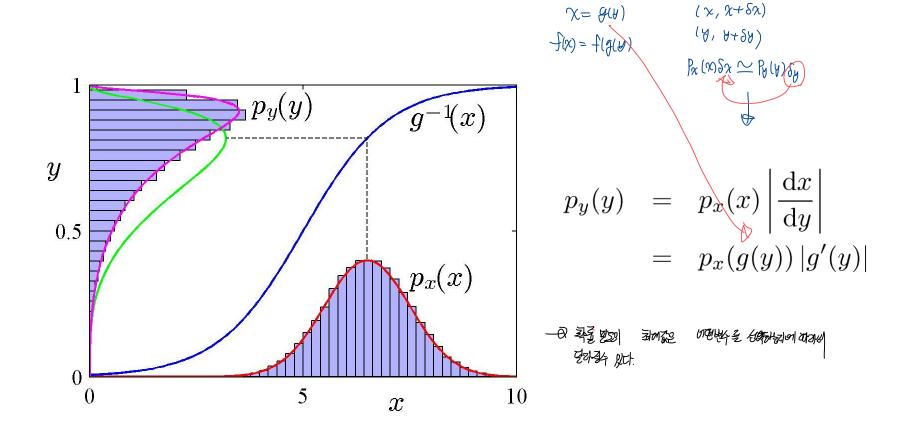
그렇게 상처가 길만시키 시교라되었다.

- 역시기라 외하일 고를 착중은

이번 상원 중하나 있다.



Transformed Densities



Expectations

$$\mathbb{E}[f] = \sum_{x} p(x) f(x) \qquad \qquad \mathbb{E}[f] = \int p(x) f(x) \, \mathrm{d}x$$

$$\mathbb{E}[f] = \sum_{x} p(x) f(x) \qquad \qquad \mathbb{E}[f] = \int p(x) f(x) \, \mathrm{d}x$$

$$\mathbb{E}[f] = \int p(x) f(x) \, \mathrm{d}x$$

$$\mathbb{E}_x[f|y] = \sum_x p(x|y)f(x)$$

Conditional Expectation (discrete)

$$\mathbb{E}[f] \simeq \frac{1}{N} \sum_{n=1}^{N} f(x_n)$$

Approximate Expectation (discrete and continuous)

Variances and Covariances

$$\operatorname{var}[f] = \mathbb{E}\left[\left(f(x) - \mathbb{E}[f(x)]\right)^{2}\right] = \mathbb{E}[f(x)^{2}] - \mathbb{E}[f(x)]^{2}$$

$$\operatorname{cov}[x, y] = \mathbb{E}_{x,y}\left[\left\{x - \mathbb{E}[x]\right\}\left\{y - \mathbb{E}[y]\right\}\right]$$

$$= \mathbb{E}_{x,y}[xy] - \mathbb{E}[x]\mathbb{E}[y]$$

$$\operatorname{cov}[\mathbf{x}, \mathbf{y}] = \mathbb{E}_{\mathbf{x},\mathbf{y}}\left[\left\{\mathbf{x} - \mathbb{E}[\mathbf{x}]\right\}\left\{\mathbf{y}^{\mathrm{T}} - \mathbb{E}[\mathbf{y}^{\mathrm{T}}]\right\}\right]$$

$$= \mathbb{E}_{\mathbf{x},\mathbf{y}}[\mathbf{x}\mathbf{y}^{\mathrm{T}}] - \mathbb{E}[\mathbf{x}]\mathbb{E}[\mathbf{y}^{\mathrm{T}}]$$

$$= \mathbb{E}_{\mathbf{x},\mathbf{y}}[\mathbf{x}\mathbf{y}^{\mathrm{T}}] - \mathbb{E}[\mathbf{x}]\mathbb{E}[\mathbf{y}^{\mathrm{T}}]$$

• $\mathrm{Cov}(X,\,Y)=\mathbb{E}(XY)-\mathbb{E}(X)\mathbb{E}(Y)^{[1]}$ [중명]

$$(X - \mu_X)(Y - \mu_Y) = XY - \mu_X Y - \mu_Y X + \mu_X \mu_Y$$
를 이용하면
$$\operatorname{Cov}(X,Y) = \mathbb{E}(XY) - \mathbb{E}(\mu_X Y) - \mathbb{E}(\mu_Y X) + \mathbb{E}(\mu_X \mu_Y)$$
$$= \mathbb{E}(XY) - \mu_X \mathbb{E}(Y) - \mu_Y \mathbb{E}(X) + \mu_X \mu_Y$$
$$= \mathbb{E}(XY) - \mu_X \mu_Y$$

- Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y) [증명]
 - \circ 일반화: $\operatorname{Var}\left(\sum_{k=1}^n X_k
 ight) = \sum_{k=1}^n \operatorname{Var}(X_k) + 2\sum_{i < j} \operatorname{Cov}(X_i,\,X_j)$
- $|Cov(X, Y)| \le \sqrt{Var(X) \cdot Var(Y)}$

× 2.1. 모공분산

[편집]

모공분산은 모집단의 공분산이다. $\mathrm{Cov}(X,Y)$ 또는 σ_{XY} 로 쓴다. X와 Y는 확률 변수, N은 모집단의 표본의 개수, X_i 와 Y_i 는 각 확률 변수의 도수, μ 는 모평균을 뜻한다.

$$ext{Cov}(X, Y) = \sigma_{XY}$$

= $\frac{1}{N} \sum_{i=1}^{n} (X_i - \mu_X)(Y_i - \mu_Y)$
= $\mathbb{E}\{(X - \mu_X)(Y - \mu_Y)\}$

 $_{Z}$, 모공분산이란 $_{X}$ 의 편차와 $_{Y}$ 의 편차의 곱의 평균이다.

∨ 2.2. 표본공분산

[편집]

표본공분산은 표본집단의 공분산이다. S_{XY} 로 쓴다. X와 Y는 확률 변수, n은 표본집단의 표본의 개수, X_i 와 Y_i 는 각 확률 변수의 도수, \bar{X} 와 \bar{Y} 는 표본평균을 뜻한다.

$$S_{XY} = \frac{1}{n-1} \sum_{i=1}^{n} (X_i - \bar{X})(Y_i - \bar{Y})$$
$$= \mathbb{E}\{(X - \bar{X})(Y - \bar{Y})\}$$

 \mathbf{Z} , 표본공분산이란 X의 편차와 Y의 편차의 \mathbf{Z} 의 평균이다. 주의할 점은 (표본의 개수)-1로 나눈다는 것이다. n이 아니라 n-1로 나누는 것은 오차를 줄이기 위함으로, 일반적인 표본 분산의 계산법과 같다.

The Gaussian Distribution

$$\mathcal{N}\left(x|\mu,\sigma^{2}\right) = \frac{1}{(2\pi\sigma^{2})^{1/2}} \exp\left\{-\frac{1}{2\sigma^{2}}(x-\mu)^{2}\right\}$$

$$\mathcal{N}(x|\mu,\sigma^{2})$$

$$\mathcal{N}(x|\mu,\sigma^{2}) > 0$$

$$\int_{-\infty}^{\infty} \mathcal{N}\left(x|\mu,\sigma^{2}\right) dx = 1$$

Gaussian Mean and Variance

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$

$$\mathbb{E}[x] = \int_{-\infty}^{\infty} \mathcal{N}(x|\mu, \sigma^2) x \, \mathrm{d}x = \mu$$

$$\mathbb{E}[x^2] = \int_{-\infty}^{\infty} \mathcal{N}(x|\mu, \sigma^2) x^2 dx = \mu^2 + \sigma^2$$

$$var[x] = \mathbb{E}[x^2] - \mathbb{E}[x]^2 = \sigma^2$$

From the definition of the Gaussian distribution, X has probability density function:

$$f_{X}\left(x
ight)=rac{1}{\sigma\sqrt{2\pi}}\mathrm{exp}igg(-rac{\left(x-\mu
ight)^{2}}{2\sigma^{2}}igg)$$

From the definition of the expected value of a continuous random variable:

$$\mathsf{E}\left(X
ight) = \int_{-\infty}^{\infty} x f_{X}\left(x
ight) \,\mathrm{d}x$$

So:

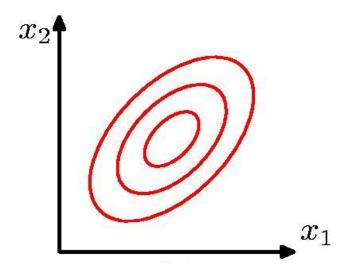
$$\begin{split} \mathsf{E}(X) &= \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} x \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) \mathrm{d}x \\ &= \frac{\sqrt{2}\sigma}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} (\sqrt{2}\sigma t + \mu) \exp(-t^2) \, \mathrm{d}t \qquad \text{substituting } t = \frac{x-\mu}{\sqrt{2}\sigma} \\ &= \frac{1}{\sqrt{\pi}} \left(\sqrt{2}\sigma \int_{-\infty}^{\infty} t \exp(-t^2) \, \mathrm{d}t + \mu \int_{-\infty}^{\infty} \exp(-t^2) \, \mathrm{d}t\right) \\ &= \frac{1}{\sqrt{\pi}} \left(\sqrt{2}\sigma \left[-\frac{1}{2}\exp(-t^2)\right]_{-\infty}^{\infty} + \mu\sqrt{\pi}\right) \qquad \text{Fundamental Theorem of Calculus, Gaussian Integral} \\ &= \frac{\mu\sqrt{\pi}}{\sqrt{\pi}} \qquad \qquad \text{Exponential Tends to Zero and Infinity} \\ &= \mu \end{split}$$

$$\begin{aligned} Var[X] &= \int_{-\infty}^{\infty} (x - \mu)^2 f_X(x) \ dx \\ &= \int_{-\infty}^{\infty} (x - \mu)^2 \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x - \mu)^2/2\sigma^2} \ dx \\ &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} z^2 e^{-z^2/2\sigma^2} \ dz \qquad (\because \ let \ z = x - \mu, \ dz = dx) \\ &= \left[-\frac{1}{\sqrt{2\pi\sigma^2}} z \sigma^2 e^{-z^2/2\sigma^2} \right]_{-\infty}^{\infty} + \frac{1}{\sqrt{2\pi\sigma^2}} \sigma^2 \int_{-\infty}^{\infty} e^{-z^2/2\sigma^2} \ dz \\ &= 0 + \sigma^2 \qquad (\because \int_{-\infty}^{\infty} e^{-z^2/2\sigma^2} \ dz = \sqrt{2\pi\sigma^2}) \end{aligned}$$

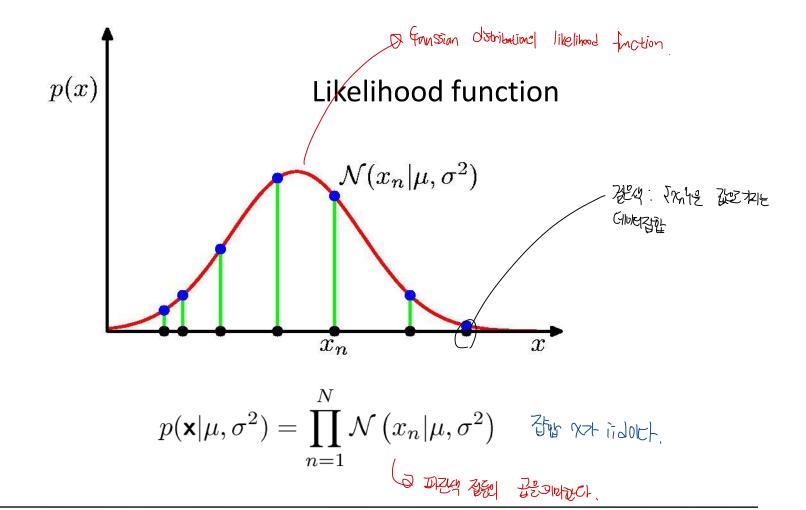
The Multivariate Gaussian

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$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right\}$$



Gaussian Parameter Estimation



M, 6-35mm 初多程 到对键则 对链 到哪 好 知.

Maximum (Log) Likelihood

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$$\underline{\ln p\left(\mathbf{x}|\mu,\sigma^2\right)} = -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (x_n - \mu)^2 - \frac{N}{2} \ln \sigma^2 - \frac{N}{2} \ln(2\pi)$$

$$\mu_{\mathrm{ML}} = \frac{1}{N} \sum_{n=1}^{N} x_{n} \qquad \sigma_{\mathrm{ML}}^{2} = \frac{1}{N} \sum_{n=1}^{N} (x_{n} - \mu_{\mathrm{ML}})^{2}$$

$$\frac{\partial L}{\partial \theta^{2}} = 0 \qquad \frac{\partial L}{\partial \theta^{2}} = 0$$

$$\frac{\partial L}{\partial \theta^{2}} = 0 \qquad \text{The Comple mean}$$

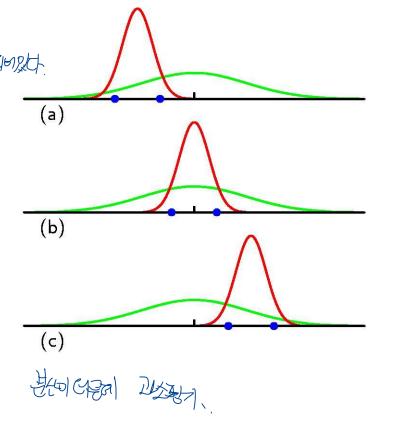
Properties of $\mu_{ m ML}$ and $\sigma_{ m ML}^2$

$$\mathbb{E}[\mu_{
m ML}] = \mu$$

$$\mathbb{E}[\sigma_{
m ML}^2] = \left(rac{N-1}{N}
ight)\sigma^2$$

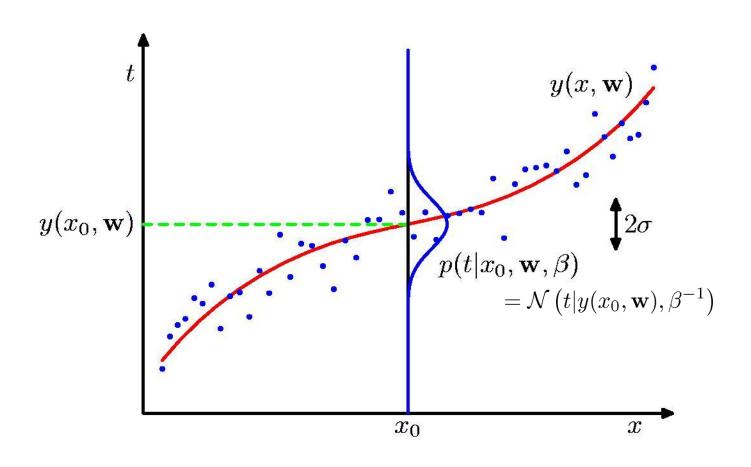
$$\widetilde{\sigma}^{2} = \frac{N}{N-1} \sigma_{\text{ML}}^{2}$$

$$= \frac{1}{N-1} \sum_{n=1}^{N} (x_{n} - \mu_{\text{ML}})^{2}$$



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Curve Fitting Re-visited



Maximum Likelihood

$$p(\mathbf{t}|\mathbf{x}, \mathbf{w}, \beta) = \prod_{n=1}^{N} \mathcal{N}\left(t_n | y(x_n, \mathbf{w}), \beta^{-1}\right)$$

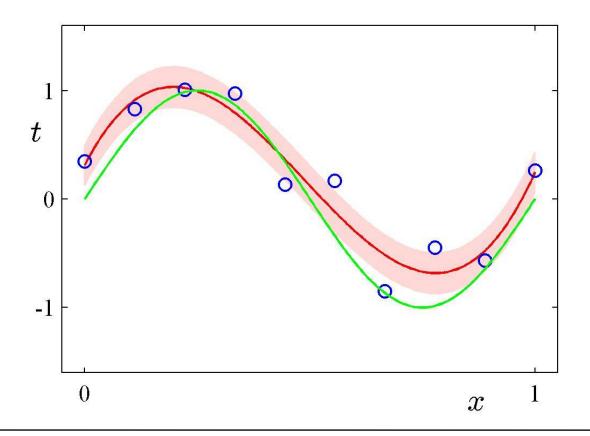
$$\ln p(\mathbf{t}|\mathbf{x}, \mathbf{w}, \beta) = -\underbrace{\frac{\beta}{2} \sum_{n=1}^{N} \left\{ y(x_n, \mathbf{w}) - t_n \right\}^2 + \frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi)}_{\beta E(\mathbf{w})}$$

Determine \mathbf{w}_{ML} by minimizing sum-of-squares error, $E(\mathbf{w})$.

$$\frac{1}{\beta_{\text{ML}}} = \frac{1}{N} \sum_{n=1}^{N} \{y(x_n, \mathbf{w}_{\text{ML}}) - t_n\}^2$$

Predictive Distribution

$$p(t|x, \mathbf{w}_{\mathrm{ML}}, \beta_{\mathrm{ML}}) = \mathcal{N}\left(t|y(x, \mathbf{w}_{\mathrm{ML}}), \beta_{\mathrm{ML}}^{-1}\right)$$



MAP: A Step towards Bayes

$$p(\mathbf{w}|\alpha) = \mathcal{N}(\mathbf{w}|\mathbf{0}, \alpha^{-1}\mathbf{I}) = \left(\frac{\alpha}{2\pi}\right)^{(M+1)/2} \exp\left\{-\frac{\alpha}{2}\mathbf{w}^{\mathrm{T}}\mathbf{w}\right\}$$

$$p(\mathbf{w}|\mathbf{x}, \mathbf{t}, \alpha, \beta) \propto p(\mathbf{t}|\mathbf{x}, \mathbf{w}, \beta)p(\mathbf{w}|\alpha)$$

$$\beta \widetilde{E}(\mathbf{w}) = \frac{\beta}{2} \sum_{n=1}^{N} \{y(x_n, \mathbf{w}) - t_n\}^2 + \frac{\alpha}{2} \mathbf{w}^{\mathrm{T}} \mathbf{w}$$

Determine $\mathbf{w}_{\mathrm{MAP}}$ by minimizing regularized sum-of-squares error, $\widetilde{E}(\mathbf{w})$.

Bayesian Curve Fitting

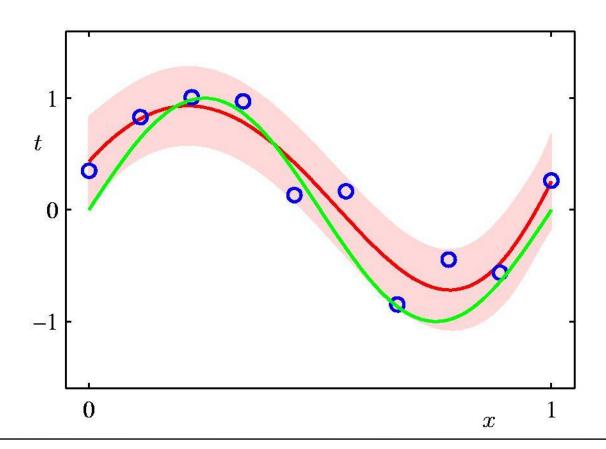
$$p(t|x, \mathbf{x}, \mathbf{t}) = \int p(t|x, \mathbf{w}) p(\mathbf{w}|\mathbf{x}, \mathbf{t}) d\mathbf{w} = \mathcal{N}(t|m(x), s^2(x))$$

$$m(x) = \beta \phi(x)^{\mathrm{T}} \mathbf{S} \sum_{n=1}^{N} \phi(x_n) t_n$$
 $s^2(x) = \beta^{-1} + \phi(x)^{\mathrm{T}} \mathbf{S} \phi(x)$

$$\mathbf{S}^{-1} = \alpha \mathbf{I} + \beta \sum_{n=1}^{N} \boldsymbol{\phi}(x_n) \boldsymbol{\phi}(x_n)^{\mathrm{T}} \qquad \boldsymbol{\phi}(x_n) = (x_n^0, \dots, x_n^M)^{\mathrm{T}}$$

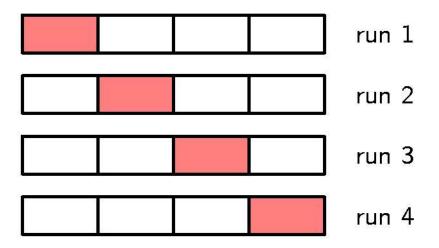
Bayesian Predictive Distribution

$$p(t|x, \mathbf{x}, \mathbf{t}) = \mathcal{N}\left(t|m(x), s^2(x)\right)$$

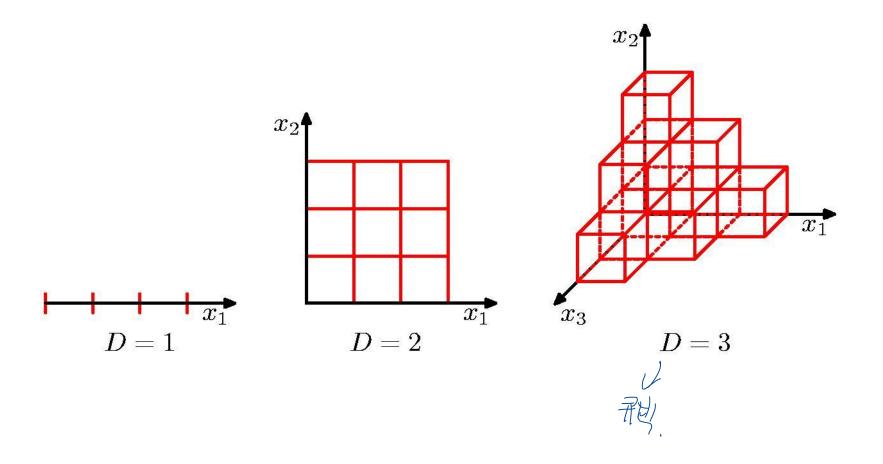


Model Selection

Cross-Validation



Curse of Dimensionality



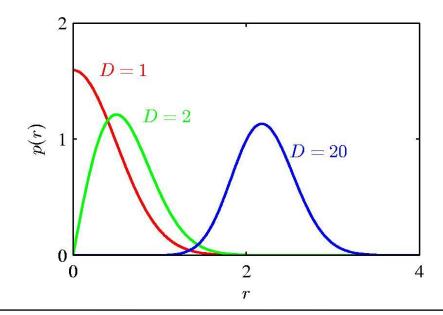
Curse of Dimensionality

Polynomial curve fitting, M=3

$$y(\mathbf{x}, \mathbf{w}) = w_0 + \sum_{i=1}^{D} w_i x_i + \sum_{i=1}^{D} \sum_{j=1}^{D} w_{ij} x_i x_j + \sum_{i=1}^{D} \sum_{j=1}^{D} \sum_{k=1}^{D} w_{ijk} x_i x_j x_k$$

Gaussian Densities in higher dimensions

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Decision Theory

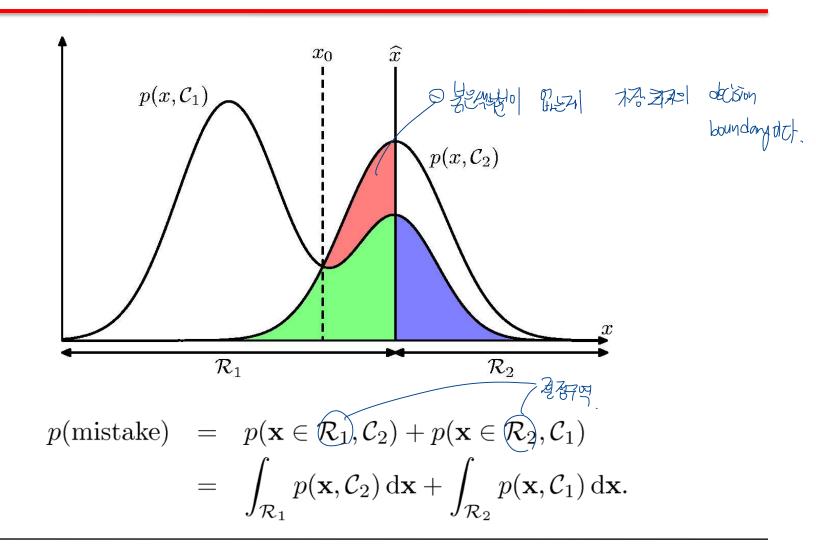
Inference step

Determine either $p(t|\mathbf{x})$ or $p(\mathbf{x},t)$.

Decision step

For given x, determine optimal t.

Minimum Misclassification Rate



Minimum Expected Loss

Example: classify medical images as 'cancer' or 'normal'

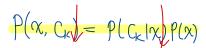
Decision
$$\operatorname{cancer normal} \quad \begin{array}{c} \\ \\ \\ \\ \\ \\ \end{array} \begin{array}{c} \\ \\ \\ \\ \end{array} \begin{array}{c} \\ \\ \\ \\ \end{array} \begin{array}{c} \\ \\ \\ \\ \end{array} \begin{array}{c} \\ \\ \\ \end{array} \begin{array}{c$$

Minimum Expected Loss

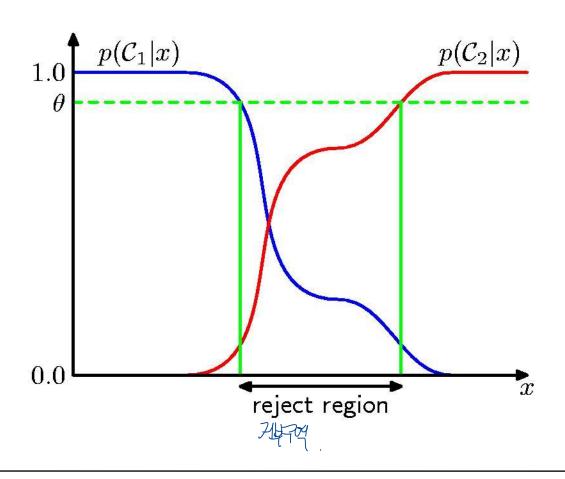
$$\mathbb{E}[L] = \sum_{k} \sum_{j} \int_{\mathcal{R}_{j}} L_{kj} p(\mathbf{x}, \mathcal{C}_{k}) d\mathbf{x}$$

Regions \mathcal{R}_i are chosen to minimize

$$\mathbb{E}[L] = \sum_{k} L_{kj} p(\mathcal{C}_k | \mathbf{x})$$



Reject Option



Why Separate Inference and Decision?

- Minimizing risk (loss matrix may change over time)
- Reject option
- Unbalanced class priors
- Combining models

Decision Theory for Regression

Inference step

Determine $p(\mathbf{x}, t)$.

Decision step For given \mathbf{x} , make optimal prediction, $y(\mathbf{x})$, for t.

Loss function:
$$\mathbb{E}[L] = \iint \underline{L(t, y(\mathbf{x}))} p(\mathbf{x}, t) \, d\mathbf{x} \, dt$$

The Squared Loss Function

$$\mathbb{E}[L] = \iint \{y(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) \, d\mathbf{x} \, dt$$

$$\{y(\mathbf{x}) - t\}^2 = \{y(\mathbf{x}) - \mathbb{E}[t|\mathbf{x}] + \mathbb{E}[t|\mathbf{x}] - t\}^2$$
$$= \{y(\mathbf{x}) - \mathbb{E}[t|\mathbf{x}]\}^2 + 2\{y(\mathbf{x}) - \mathbb{E}[t|\mathbf{x}]\}\{\mathbb{E}[t|\mathbf{x}] - t\} + \{\mathbb{E}[t|\mathbf{x}] - t\}^2$$

$$\mathbb{E}[L] = \int \{y(\mathbf{x}) - \mathbb{E}[t|\mathbf{x}]\}^2 p(\mathbf{x}) d\mathbf{x} + \int \operatorname{var}[t|\mathbf{x}] p(\mathbf{x}) d\mathbf{x}$$

$$y(\mathbf{x}) = \mathbb{E}[t|\mathbf{x}]$$

$$\frac{\int \mathbb{E}[t]}{\int y(x)} = 2 \int \{y(x) - t \} p(x,t) dt = 0, \quad y(x) = \frac{\int t p(x,t) dt}{p(x)} = \int t p(t|x) dt - \mathbb{E}_t[t|x]$$

Generative vs Discriminative

Generative approach:

Model
$$p(t, \mathbf{x}) = p(\mathbf{x}|t)p(t)$$

Use Bayes' theorem $p(t|\mathbf{x}) = \frac{p(\mathbf{x}|t)p(t)}{p(\mathbf{x})}$

Discriminative approach:

Model $p(t|\mathbf{x})$ directly

Entropy: 製物 细胞 细胞 驱

$$\begin{array}{l} h(x) = -l_{x} p(x) \\ h(x) = h(x) + h(y) \\ p(x,y) = p(x) p(y) \end{array} \qquad \begin{array}{l} H[x] = -\sum_{x} p(x) \log_2 p(x) \\ \\ f(x,y) = f(x) p(y) \end{array}$$

Important quantity in

- coding theory
- statistical physics
- machine learning

Coding theory: x discrete with 8 possible states; how many bits to transmit the state of x?

All states equally likely

$$H[x] = -8 \times \frac{1}{8} \log_2 \frac{1}{8} = 3 \text{ bits.}$$

$$H[x] = -\frac{1}{2}\log_2\frac{1}{2} - \frac{1}{4}\log_2\frac{1}{4} - \frac{1}{8}\log_2\frac{1}{8} - \frac{1}{16}\log_2\frac{1}{16} - \frac{4}{64}\log_2\frac{1}{64}$$
= 2 bits

average code length =
$$\frac{1}{2} \times 1 + \frac{1}{4} \times 2 + \frac{1}{8} \times 3 + \frac{1}{16} \times 4 + 4 \times \frac{1}{64} \times 6$$

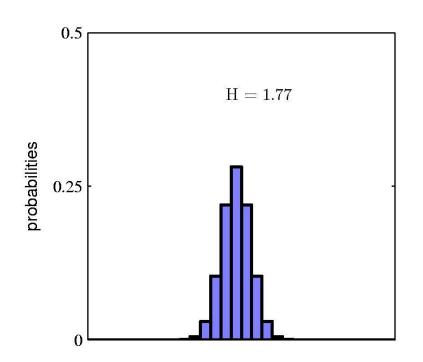
= 2 bits

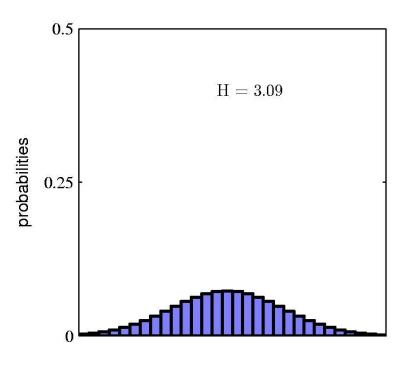
In how many ways can N identical objects be allocated M bins?

$$W = \frac{N!}{\prod_i n_i!}$$

$$H = \frac{1}{N} \ln W \simeq -\lim_{N \to \infty} \sum_{i} \left(\frac{n_i}{N}\right) \ln \left(\frac{n_i}{N}\right) = -\sum_{i} p_i \ln p_i$$

Entropy maximized when $\forall i: p_i = \frac{1}{M}$





Differential Entropy

Put bins of width Δ along the real line

$$\lim_{\Delta \to 0} \left\{ -\sum_{i} p(x_i) \Delta \ln p(x_i) \right\} = -\int p(x) \ln p(x) dx$$

Differential entropy maximized (for fixed σ^2) when

$$p(x) = \mathcal{N}(x|\mu, \sigma^2)$$

in which case

$$H[x] = \frac{1}{2} \left\{ 1 + \ln(2\pi\sigma^2) \right\}.$$

67号号 图图扩大 圣科学的经验 以上 乳性的 鬼仆

Conditional Entropy

$$H[\mathbf{y}|\mathbf{x}] = -\iint p(\mathbf{y}, \mathbf{x}) \ln p(\mathbf{y}|\mathbf{x}) d\mathbf{y} d\mathbf{x}$$

$$H[\mathbf{x}, \mathbf{y}] = H[\mathbf{y}|\mathbf{x}] + H[\mathbf{x}]$$

一义,然紧 导动 明明 要证 智知 紀 《已 四日 早秋川 州川 要配 智斯 图上 入了上 子以来的 安全 平秋川 子州 电路记 不知 是 教授的一名十一

The Kullback-Leibler Divergence

$$\begin{split} \operatorname{KL}(p\|q) &= -\int p(\mathbf{x}) \ln q(\mathbf{x}) \, \mathrm{d}\mathbf{x} - \left(-\int p(\mathbf{x}) \ln p(\mathbf{x}) \, \mathrm{d}\mathbf{x}\right) \\ &= -\int p(\mathbf{x}) \ln \left\{\frac{q(\mathbf{x})}{p(\mathbf{x})}\right\} \, \mathrm{d}\mathbf{x} \end{split}$$

$$\mathrm{KL}(p||q) \simeq \frac{1}{N} \sum_{n=1}^{N} \left\{ -\ln q(\mathbf{x}_n|\boldsymbol{\theta}) + \ln p(\mathbf{x}_n) \right\}$$

$$\mathrm{KL}(p\|q)\geqslant 0 \qquad \qquad \mathrm{KL}(p\|q)\not\equiv \mathrm{KL}(q\|p)$$

$$\mathrm{KL}(p\|y)=0 \text{ for } p(x)=p(x)$$

Mutual Information

$$I[\mathbf{x}, \mathbf{y}] \equiv KL(p(\mathbf{x}, \mathbf{y}) || p(\mathbf{x}) p(\mathbf{y}))$$

$$= -\iint p(\mathbf{x}, \mathbf{y}) \ln \left(\frac{p(\mathbf{x}) p(\mathbf{y})}{p(\mathbf{x}, \mathbf{y})} \right) d\mathbf{x} d\mathbf{y}$$

$$I[\mathbf{x}, \mathbf{y}] = H[\mathbf{x}] - H[\mathbf{x}|\mathbf{y}] = H[\mathbf{y}] - H[\mathbf{y}|\mathbf{x}]$$