

# Likelihood

## Models for Socio-Environmental Data

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# Why Likelihood?

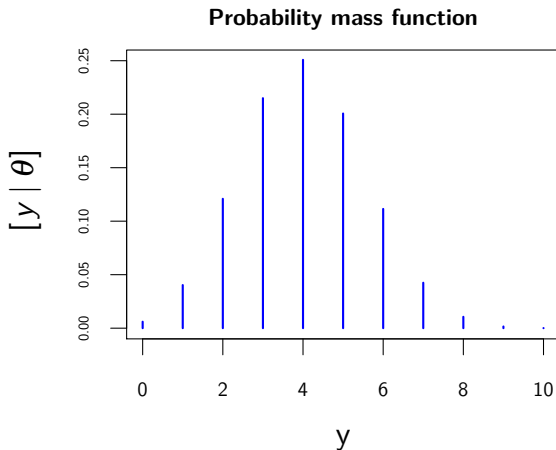
- Maximum likelihood is an important statistical approach.
- Likelihood is a component of all Bayesian models.

# Learning objectives for lecture and lab

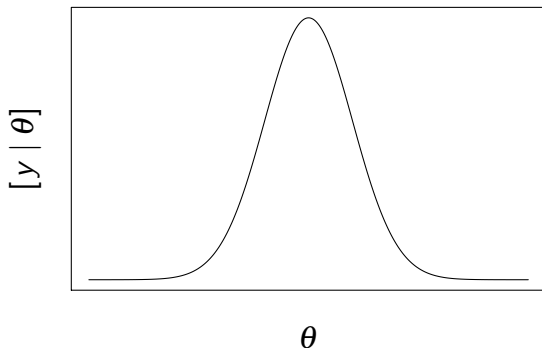
- Understand the concept of likelihood and its relationship to the probability of data conditional on parameters.
- Describe a likelihood profile and how it differs from a probability density function.
- Compose likelihoods for multiple parameters and multiple observations.

## Course progression so far...

$$[y]$$
$$[y \mid \theta]$$
$$y \sim \text{binomial}(\theta, 10)$$



Inference from likelihood is based on  $[y | \theta]$



- Likelihood allows us to compare alternative parameter values by calculating the probability (or probability density) of the data conditional on parameters  $[y | \theta]$ .

# The key idea in likelihood

- In a probability mass or density function, the parameter  $\theta$  is constant (fixed) and the outcomes  $\mathbf{y}$  vary (these outcomes represent data we may observe). The function sums or integrates to 1 over its support.
- In a likelihood function, the data  $\mathbf{y}$  are constant (fixed) and the parameter  $\theta$  varies. We use  $[y | \theta]$  to assess the likelihood of different values of  $\theta$  in light of the data. In this case, the function does not sum or integrate to one over all possible values of the parameter.
- As you will see, all evidence based on likelihood is relative.

$$\underbrace{L(\theta | y)}_{\text{likelihood function}} \propto \underbrace{[y | \theta]}_{\text{PDF or PMF}}$$

Likelihood is *proportional* to probability or probability density

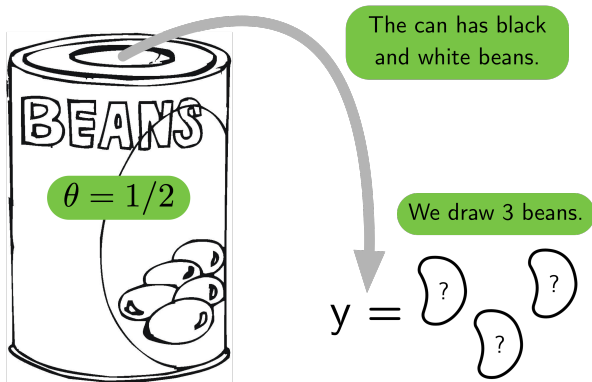
## Discuss notation

$$L(\theta | y) \propto [y | \theta] \quad (1)$$

$$L(\theta | y) = c[y | \theta] \quad (2)$$

$$L(\theta | y) = [y | \theta] \quad (3)$$

## The parameter is known and the data are random



- What are the possible outcomes?
- What probability mass function would you use to model these data?
- What is the probability of each outcome?
- What is the sum of the individual probabilities?



## The parameter is known and the data are random

We draw three beans from a can with equal numbers of white and black beans. The possible outcomes are:

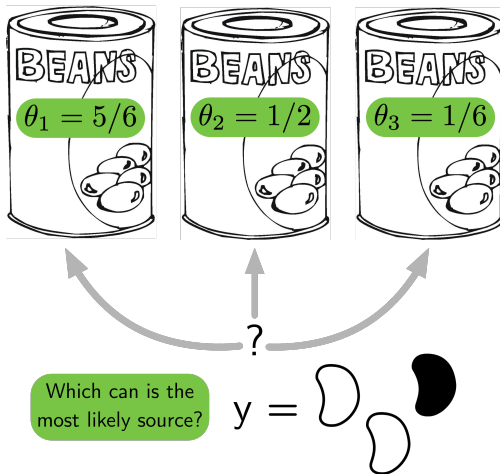
$\theta$	number of white beans, $y_i$	$[y_i \mid \theta = ?]$
	$\sum_{i=1}^4 [y_i \mid \theta = ?]$	?

## The parameter is known and the data are random

We draw three beans from a can with equal numbers of white and black beans. The possible outcomes are:

$\theta$	number of white beans, $y_i$	$[y_i \mid \theta = .5]$
.5	0	.125
.5	1	.375
.5	2	.375
.5	3	.125
	$\sum_{i=1}^4 [y_i \mid \theta = .5]$	1

The data are known and the parameter is random



- What is the likelihood of each parameter value?

## The data are known and the parameter is random

We have three hypothesized parameter values,  $(5/6, 1/2, 1/6)$ . Data in hand are 2 white beans on three draws. The likelihood of each parameter value is:

hypothesis, $\theta_i$	number of white beans, $y$	$[y = ? \mid \theta_i]$
	$\sum_{i=1}^3 [y = ? \mid \theta_i]$	?

## The data are known and the parameter is random

We have three hypothesized parameter values,  $(5/6, 1/2, 1/6)$ . Data in hand are 2 white beans on three draws. The likelihood of each parameter value is:

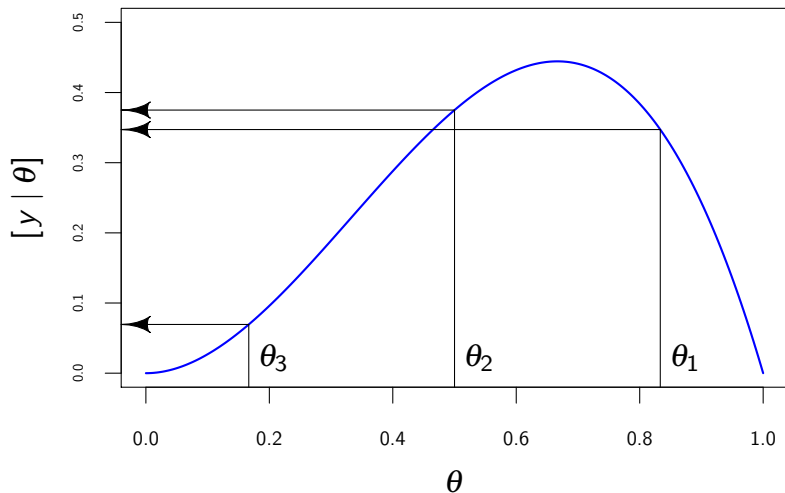
hypothesis, $\theta_i$	number of white beans, $y$	$[y = 2 \mid \theta_i]$
$\theta_1 = 5/6$	2	0.35
$\theta_2 = 1/2$	2	0.38
$\theta_3 = 1/6$	2	0.069
	$\sum_{i=1}^3 [y = 2 \mid \theta_i]$	0.79

## The data are known and the parameter is random (rescaled)

We have three hypothesized parameter values,  $(5/6, 1/2, 1/6)$ . Data in hand are 2 white beans on three draws. The likelihood of each parameter value is:

hypothesis, $\theta_i$	number of white beans, $y$	$[y = 2 \mid \theta_i]$
$\theta_1 = 5/6$	2	$0.35 / 0.38 = 0.92$
$\theta_2 = 1/2$	2	$0.38 / 0.38 = 1.0$
$\theta_3 = 1/6$	2	$0.069 / 0.38 = 0.18$
	$\sum_{i=1}^3 [y = 2 \mid \theta_i]$	2.1

## A likelihood profile: 2 white beans on 3 draws



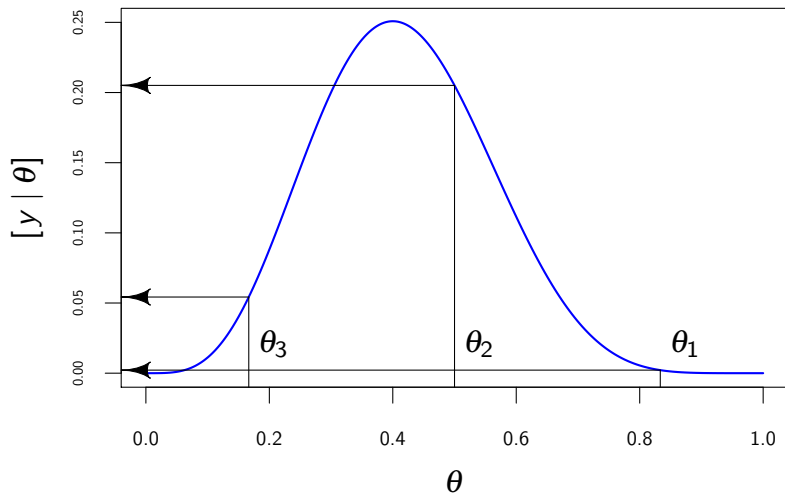
## Class exercise: Can of beans

An adventurous person takes a draw of 10 beans from one of the cans where the identity of the can is unknown. Of the 10 beans drawn from the mystery can, 4 are white.

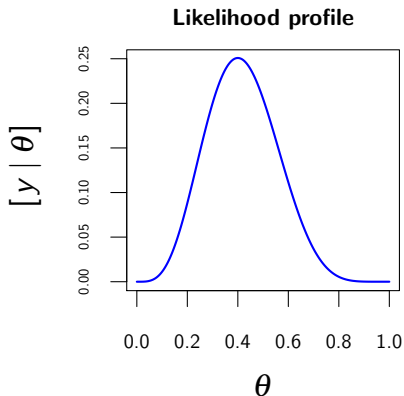
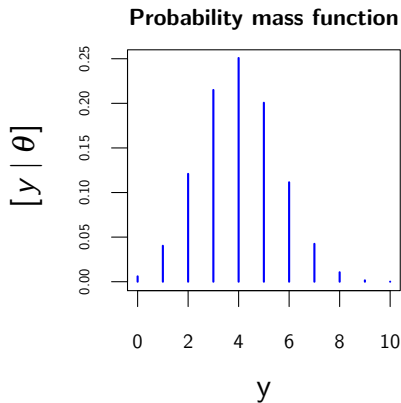
- Which of the three cans is the most likely to have produced this draw?
- How much more likely is this can than the other two?



## A likelihood profile: 4 white beans on 10 draws



## Likelihood vs Probability:



$$[y | \theta] = \binom{10}{y} \theta^y (1 - \theta)^{n-y}$$

## How do we compute likelihoods for multiple parameters?

- In the example we had a single parameter,  $\theta$ , one set of observation, 4 successes on 10 draws, and a binomial likelihood.
- However, we could have  $\theta$  a function of covariates. Procedurally, we could replace  $\theta$  with  $g(\beta, x_i)$ , which would allow  $\theta$  to vary with respect to  $x_i$ .
- More generally, we often wish to predict the mean response of an RV and then embed this mean function in a probability mass or density function, either directly or using moment matching.

$$\begin{aligned}\mu_i &= g(\alpha, x_i) \\ L(\alpha, \sigma^2 \mid y_i) &\propto \underbrace{[y_i \mid \mu_i, \sigma^2]}_{\text{PDF or PMF}}\end{aligned}$$

# Likelihood Surfaces

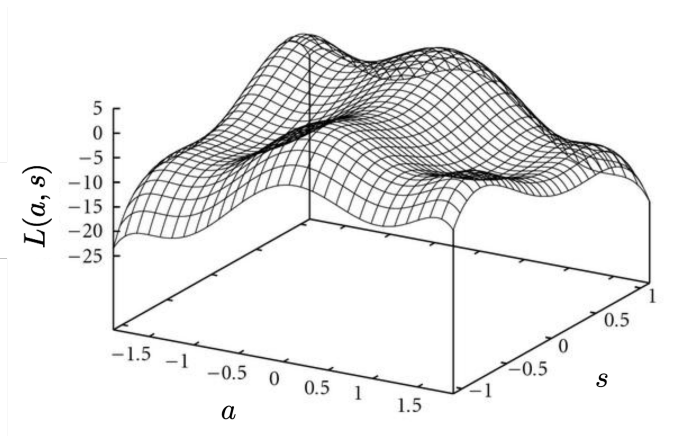


Figure courtesy of [Sergiy Nesterko](#).

## How do we compute likelihoods for multiple parameters and multiple data points?

The total likelihood is the product of the individual likelihoods, assuming the data are *conditionally independent*:

$$L(\boldsymbol{\mu}, \sigma^2 \mid \mathbf{y}) = c \prod_{i=1}^n [y_i \mid g(\boldsymbol{\theta}, x_i), \sigma^2]$$

## What does conditionally independent mean?

Independence is an assumption! Remember from the chain rule:

$$Pr(y_1, \dots, y_n) = Pr(y_1 | y_2 \dots y_n) Pr(y_2 | y_3 \dots y_n) \dots Pr(y_n).$$

However, by assuming that these random variables are independent, you can simplify the joint probability into:

$$Pr(y_1, \dots, y_n) = Pr(y_1) Pr(y_2) \dots Pr(y_n),$$

such that the total likelihood is a product of the individual likelihoods.

## Log likelihoods:

We often use the sum of the log likelihoods to get the total log likelihood as a basis for fitting models:

$$\log(L(\boldsymbol{\theta}, \sigma^2 \mid y)) = \log(c) + \sum_{i=1}^n \log([y_i \mid g(\boldsymbol{\theta}, x_i), \sigma^2])$$

## Class Exercise: Likelihood profile of $\lambda$ for tadpole counts

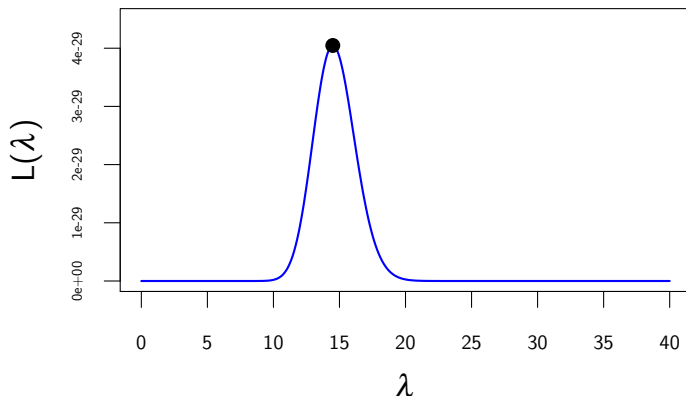
Assume we have observed the following counts of tadpoles in funnel traps from 6 different ponds: 7, 11, 54, 12, 2, 1 individuals.

Assuming these counts were generated using a Poisson process governed by  $\lambda$ , compute a likelihood profile using R.

Eyeball this profile to determine the MLE.



## Likelihood profile of $\lambda$ for tadpole counts



```
lambda <- seq(0, 40, length = 1000)
y <- NA
for(i in 1:length(lambda)) {y[i] <- prod(dpois(c(7, 11, 54, 12, 2, 1), lambda[i]))}
```

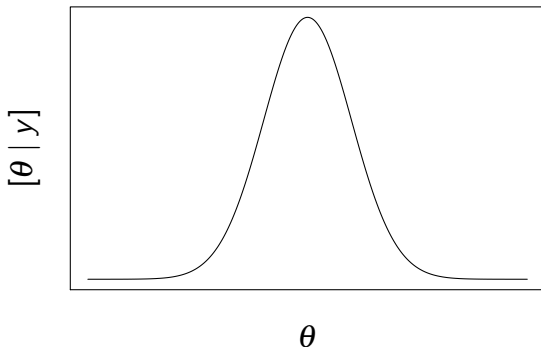
# Things to remember

- Likelihood allows us to evaluate the relative strength of evidence for one parameter or model relative to another.
- Likelihood is a component of all Bayesian models.
- The data are fixed and the parameters are variable in likelihood functions. These functions do not integrate or sum to one over the range of values of the parameter.
- The outcomes are variable and the parameters are fixed in probability mass functions and density functions. These functions sum or integrate to one over the support of the random variable.

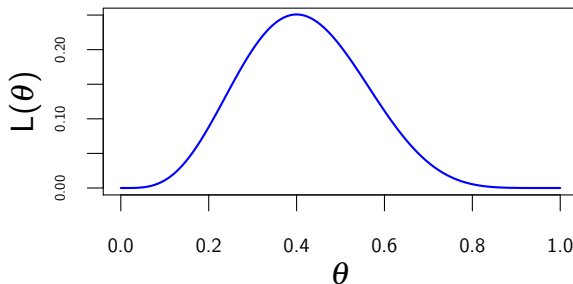
# End of Lecture

The remaining slides are in anticipation of questions or to provide additional information.

Looking ahead: Bayesian inference is based on  $[\theta | y]$



## Looking ahead: Getting to $[\theta | y]$ from a likelihood



- What must be done to insure that the area under the curve = 1?
- If you do this, is this now a probability density function for  $\theta$ ?

## What does conditionally independent mean?

We evaluate the independence assumption by examining the residuals ( $\varepsilon$ ) from a model, where ( $\varepsilon_i = y_i - g(\theta, x_i)$ ).

The independence assumption holds if knowing a residual tells you nothing about the other residuals.

We assess this by ensuring that the residuals:

- do not show a trend, meaning they should be centered on 0 throughout the range of fitted values,
- and are not be autocorrelated.

## Likelihood ratio confidence intervals

Find the upper and lower bounds of an interval where all  $\lambda$  values within that interval are as consistent with the data as  $\lambda_{MLE}$ .

We compute the likelihood ratio statistic:

$$R = 2 \log \left( \frac{L(\lambda_{MLE} | y)}{L(\lambda_0 | y)} \right) \sim \chi_{k=1}^2$$

which is distributed  $\chi^2$  with 1 degree of freedom. Note that we fail to reject  $H_0$  that  $\lambda = \lambda_0$  if  $R < \chi_{k=1}^2(1 - \alpha)$ .

## Likelihood ratio confidence intervals

We determine the  $(1 - \alpha = 0.95)$  likelihood ratio confidence interval by finding the upper and lower bounds for all values of  $\lambda_0$  where we would fail to reject  $H_0$ .

$$2 \log \left( \frac{L(\lambda_{MLE} | y)}{L(\lambda | y)} \right) < \chi_{k=1}^2(0.95)$$

$$L(\lambda_{MLE}) - \frac{3.84}{2} < L(\lambda | y)$$

$$L(\lambda | y) > L(\lambda_{MLE}) - 1.92$$



# Likelihood ratio confidence intervals

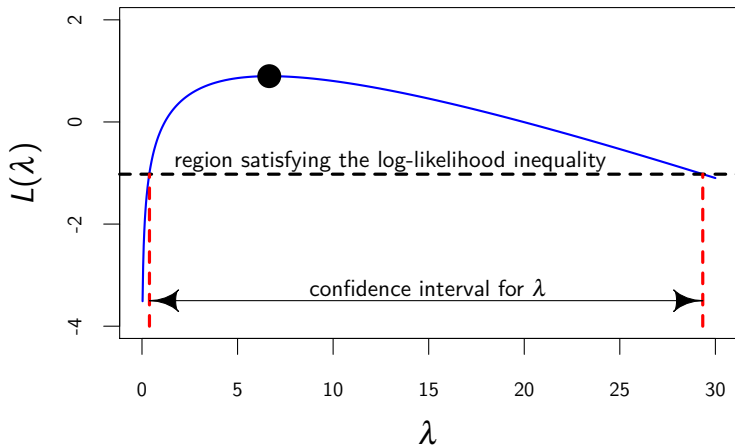
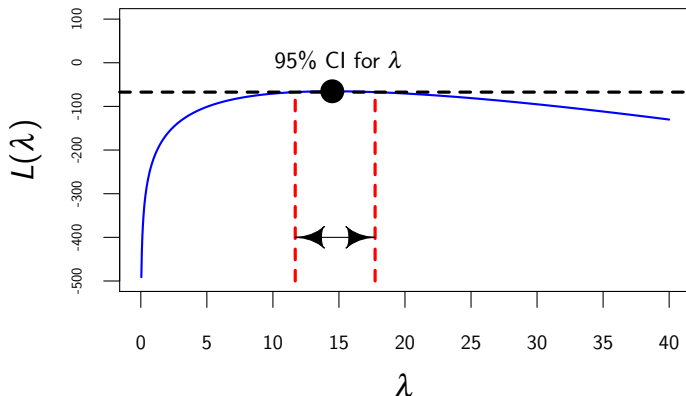


Figure courtesy of the [UNC Biology Department](#).

## Likelihood profile of $\lambda$ for tadpoles



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y <- NA
for(i in 1:length(lambda)) {y[i] <- log(prod(dpois(c(7, 11, 54, 12, 2, 1), lambda[i])))}
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