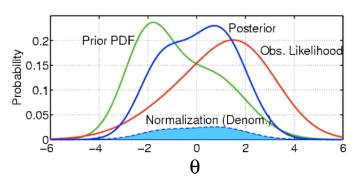
More About Priors I Models for Socio-Environmental Data

Chris Che-Castaldo, Mary B. Collins, N. Thompson Hobbs

May 23, 2019



The posterior distribution uses information from both the likelihood and prior distributions. *An informative prior is more influential than an vague prior*.



Outline

- Informative priors
- Vague priors
- Conjugate priors

Why use informative priors?

- They speed up convergence
- They reduce problems with identifiability
- They can allow you to estimate difficult/impossible quantities

Why don't we find people using informative priors more often?

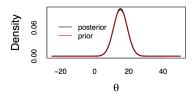
- Cultural: "All studies stand alone" argument
- Texts often use vague priors (including H&H)
- Hard work!
- Concerns about "excessive subjectivity"

If you wanted to use an informative prior, how would you do it?

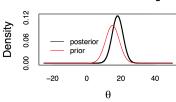
- Strong scholarship is the basis for strong priors
- Moment match, converting means and standard deviations to usable parameters
- Pilot studies
- Allometric relationships
- Deterministic models with parameters that have specific meaning

How much does an informative prior influence the posterior?

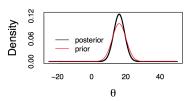




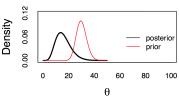
B. Moved mean + shrinkage



C. Shrinkage



D. Increased variance (rare)



Communicating your use of informative priors

Ecological Monographs, 85(4), 2015, pp. 525-556 © 2015 by the Ecological Society of America

State-space modeling to support management of brucellosis in the Yellowstone bison population

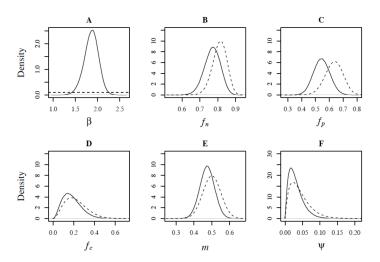
N. Thompson Hobbs, ^{1,5} Chris Geremia, ² John Treanor, ² Rick Wallen, ² P. J. White, ² Mevin B. Hooten, ³ and Jack C. Rhyan⁴

Communicating your use of informative priors

Table 3: Prior distributions for parameters in model of brucellosis in the Yellowstone bison population. Sources are given for informative priors.

Parameter	Definition	Distribution	Mean	SD	Source
β	Rate of transmission	uniform(0,50)	25	14.3	vague
	(yr^{-1})				
f_n	Number of offspring	beta(77,18)	.81	.04	Fuller et al., 2007
	recruited per				
	seronegative				
	(susceptible) female				
f_p	Number of offspring	beta(37,20)	.64	.06	Fuller et al., 2007
	recruited per				
	seropositive (recovered)				
	female				
f_c	Number of offspring	beta(3.2,11)	.22	.10	Fuller et al., 2007
	recruited per				
	seroconverting				
	/· c · · · · · · · · · · ·				

Communicating your use of informative priors



Vague Priors

A vague prior is a distribution with a range of uncertainty that is clearly wider than the range of reasonable values for the parameter (Gelman and Hill 2007:347).

Vague Priors

- Avoid calling a prior "uninformative" or "non-informative" rather:
 - difuse
 - ▶ flat
 - automatic
 - nonsubjective
 - ► locally uniform
 - objective

Issues With Vague Priors

- Computational: failure to converge, slicer errors, failure to calculate log density, etc.
- Sensitivity: changes in parameters of "vague" priors meaningfully changes the posterior when data sets are small or when they have high variance (e.g. $\tau \sim \text{gamma}(.001,.001)$ can really be problematic, this will come up in the multilevel modeling lab)

Conjugacy

- In special cases the posterior, $[\theta|y]$, has the same distributional form as the prior, $[\theta]$.
 - For example, if you had a prior, gamma(α, β), your posterior would be gamma($\alpha_{new}, \beta_{new}$)
- In these cases, the prior and the posterior are said to be conjugate.

Conjugacy is important for two reasons:

- Conjugacy minimizes computational work and, in more complicated cases, allows us to break down calculations into manageable chunks.
- ② Conjugacy plays an important role in Markov chain Monte Carlo (more on this later).

We know that the beta distribution is a conjugate prior for the binomial likelihood.

- Consider calculating the posterior distribution for the parameter θ .
- θ is the probability of a success conditional on n trials and y observed successes.

Using Bayes theorem:

$$[\phi \mid y, n] \propto \underbrace{\binom{n}{y} \phi^{y} (1 - \phi)^{n - y}}_{\text{binomial likelihood}} \underbrace{\frac{\gamma(\alpha + \beta)}{\gamma(\alpha)\gamma(\beta)} \phi^{\alpha - 1} (1 - \phi)^{\beta - 1}}_{\text{beta prior}}$$

where α and β are beta prior parameters.

$$[\phi \mid y, n] \propto \binom{n}{y} \phi^y (1-\phi)^{n-y} \frac{\gamma(\alpha+\beta)}{\gamma(\alpha)\gamma(\beta)} \phi^{\alpha-1} (1-\phi)^{\beta-1}$$

$$\begin{aligned} & [\phi \mid y, n] \propto \binom{n}{y} \phi^{y} (1 - \phi)^{n - y} \frac{\gamma(\alpha + \beta)}{\gamma(\alpha)\gamma(\beta)} \phi^{\alpha - 1} (1 - \phi)^{\beta - 1} \\ & [\phi \mid y, n] \propto \phi^{y} (1 - \phi)^{n - y} \phi^{\alpha - 1} (1 - \phi)^{\beta - 1} \end{aligned}$$

$$\begin{aligned} & [\phi \mid y, n] \propto \binom{n}{y} \phi^{y} (1 - \phi)^{n - y} \frac{\gamma(\alpha + \beta)}{\gamma(\alpha)\gamma(\beta)} \phi^{\alpha - 1} (1 - \phi)^{\beta - 1} \\ & [\phi \mid y, n] \propto \phi^{y} (1 - \phi)^{n - y} \phi^{\alpha - 1} (1 - \phi)^{\beta - 1} \\ & [\phi \mid y, n] \propto \phi^{y + \alpha - 1} (1 - \phi)^{\beta + n - y - 1} \end{aligned}$$

$$[\phi \mid y, n] \propto \phi^{y+\alpha-1} (1-\phi)^{\beta+n-y-1}$$

Let $\alpha_{new} = y + \alpha$ and $\beta_{new} = \beta + n - y$. Multiply by normalizing constant:

$$\frac{\gamma(\alpha_{new}+\beta_{new})}{\gamma(\alpha_{new})\gamma(\beta_{new})},$$

and the posterior of ϕ is a beta distribution:

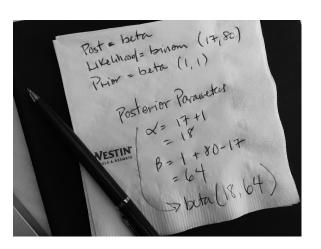
$$[\phi \mid y, n] = rac{\gamma(lpha_{new} + eta_{new})}{\gamma(lpha_{new})\gamma(eta_{new})} \phi^{lpha_{new} - 1} (1 - \phi)^{eta_{new} - 1},$$

with parameters α_{new} and β_{new} .

Conjugate priors

Table A.3: Table of conjugate distributions

T 21121 1	D. 1 11	D-4-1
Likelihood	Prior distribution	Posterior distribution
$y_i \sim \text{binomial}(n, \phi)$	$\phi \sim \text{beta}(\alpha, \beta)$	$\phi \sim \text{beta}\left(\sum y_i + \alpha, n - \sum y_i + \beta\right)$
$y_i \sim \text{Bernoulli}(\phi)$	$\phi \sim \text{beta}(\alpha, \beta)$	$\phi \sim \text{beta}(\sum_{i=1}^{n} y_i + \alpha, \sum_{i=1}^{n} (1 - y_i) + \beta)$
$y_i \sim \text{Poisson}(\lambda)$	$\lambda \sim \text{gamma}(\alpha, \beta)$	$\lambda \sim \text{gamma} (\alpha + \sum_{i=1}^{n} y_i, \beta + n)$
$y_i \sim \text{normal}(\mu, \sigma^2)$ σ^2 is known.	$\mu \sim \text{normal} \left(\mu_0, \sigma_0^2\right)$	$\mu \sim \text{normal}\left(\frac{\left(\frac{\mu_0}{\sigma_0^2} + \frac{\sum_{i=1}^n y_i}{\sigma^2}\right)}{\left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}\right)}, \left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}\right)^{-1}\right)$
$y_i \sim \text{normal}(\mu, \sigma^2)$	$\sigma^2 \sim$	$\sigma^2 \sim$
μ is known.	inverse gamma (α, β)	inverse gamma $\left(\alpha + \frac{n}{2}, \beta + \frac{\sum_{i=1}^{n} (y_i - \mu)^2}{2}\right)$
$y_i \sim \text{lognormal}(\mu, \sigma^2)$,	$\sigma^2 \sim$	$\sigma^2 \sim$
μ is known	inverse gamma (α, β) ,	inverse gamma $\left(n/2 + \alpha, \frac{(\log(y_i) - \mu)^2}{2} + \beta\right)$
$y_i \sim \text{lognormal}(\mu, \sigma^2)$ $\sigma^2 \text{ is known}$	$\mu \sim \text{normal} (\mu_0, \sigma_0^2)$	$\mu \sim \text{normal}\left(\frac{\left(\frac{\mu_0}{\sigma_0^2} + \frac{\sum_{i=1}^n \log y_i}{\sigma^2}\right)}{\left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}\right)}, \left(\frac{1}{\sigma_0^2} + \frac{n}{\sigma^2}\right)^{-1}\right)$



Why Use Conjugacy

- It is not necessary, conjugate priors will accelerate MCMC.
- For simple models, you can use conjugate priors to obtain the posterior distribution in closed form, without any simulation.

Things to remember

- There is no such thing as a uninformative prior, but certain priors influence the posterior distribution more than others.
- Informative priors, when properly justified, can be useful.
- Strong data overwhelms a prior.
- Priors represent current knowledge (or lack of), which is updated with data.
- We encourage you to think of vague priors as a provisional starting point.

Lab exercises.