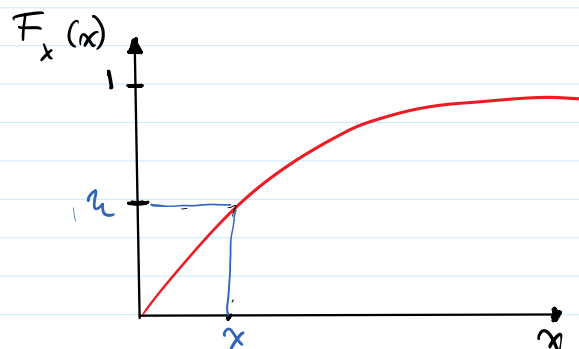
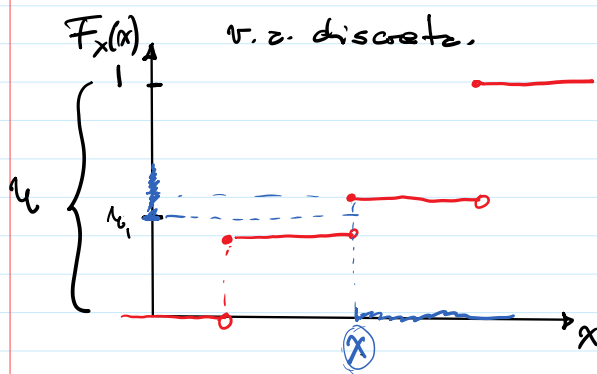


Definimos la inversa generalizada como:

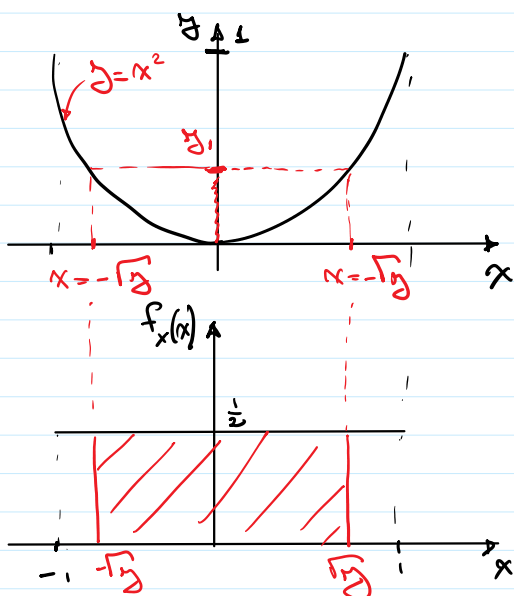
$$F_X^{-1}(u) = \min\{x \in \mathbb{R} : F_X(x) \geq u\}, u \in (0, 1)$$



$$u = F_X(x) \Rightarrow u = 1 - e^{-\lambda x} \Rightarrow x = \frac{\ln(1-u)}{-\lambda} = F^{-1}(u)$$

## Ejercicio 6

Sea  $X \sim U(-1, 1)$ , y sea  $Y = X^2$ . Hallar la función de densidad de  $Y$



$$\begin{aligned} F_Y(y) &= P(Y \leq y) = P(X^2 \leq y) = \\ &= P(-\sqrt{y} \leq X \leq \sqrt{y}) = \\ &= F_X(\sqrt{y}) - F_X(-\sqrt{y}) = \end{aligned}$$

$$F_X(x) = \begin{cases} 0 & \text{si } x < -1 \\ \frac{x+1}{2} & \text{si } -1 \leq x < 1 \\ 1 & \text{si } x \geq 1 \end{cases}$$

$$F_Y(y) = \frac{\sqrt{y}+1}{2} - \left( \frac{-\sqrt{y}+1}{2} \right) = \frac{2\sqrt{y}}{2} = \sqrt{y}$$

$$F_Y(y) = \begin{cases} 0 & \text{si } y < 0 \\ \sqrt{y} & \text{si } 0 \leq y < 1 \\ 1 & \text{si } y \geq 1 \end{cases} \rightarrow$$

$$f_Y(y) = \frac{dF_Y(y)}{dy} = \frac{1}{2\sqrt{y}}$$

$$f_Y(y) = \begin{cases} \frac{1}{2\sqrt{y}} & \text{si } 0 < y < 1 \\ 0 & \text{en otro caso.} \end{cases}$$

$$F_Y(y) = \sqrt{y} \mathbb{1}\{0 \leq y < 1\} + \mathbb{1}\{y \geq 1\}$$

$$f_Y(y) = \frac{1}{2\sqrt{y}} \mathbb{1}\{0 < y < 1\}$$

## Ejercicio 7

Sean  $X$  e  $Y$  dos v.a. con distribución de Poisson de parámetros  $\mu$  y  $\lambda$  respectivamente. Hallar la función de probabilidad de  $W = X + Y$ .

$$X \sim \text{Poi}(\mu) \rightarrow p_X(x) = \frac{\mu^x}{x!} e^{-\mu}, x \in \mathbb{N}_0$$

$X \sim \text{Po}(\mu)$ ,  $Y \sim \text{Po}(\lambda) \rightarrow X$  e  $Y$  son independientes

$$\begin{aligned} P_W(w) &= P(W=w) = P(X+Y=w) = \sum_{x=0}^w P(X=x, Y=w-x) = \\ &= \sum_{x=0}^w \frac{\mu^x \cdot e^{-\mu}}{x!} \cdot \frac{\lambda^{w-x} \cdot e^{-\lambda}}{(w-x)!} = e^{-(\mu+\lambda)} \sum_{x=0}^w \frac{\mu^x \cdot \lambda^{w-x}}{x! \cdot (w-x)!} = \\ &= \frac{e^{-(\mu+\lambda)}}{w!} \sum_{x=0}^w \frac{w!}{x! \cdot (w-x)!} \cdot \mu^x \cdot \lambda^{w-x} = \\ &= \frac{e^{-(\mu+\lambda)}}{w!} \sum_{x=0}^w \frac{w!}{x! \cdot (w-x)!} \cdot \left(\frac{\mu}{\mu+\lambda}\right)^x \cdot \left(\frac{\lambda}{\mu+\lambda}\right)^{w-x} \cdot (\mu+\lambda)^x \cdot (\mu+\lambda)^{w-x} = \textcircled{*} \end{aligned}$$

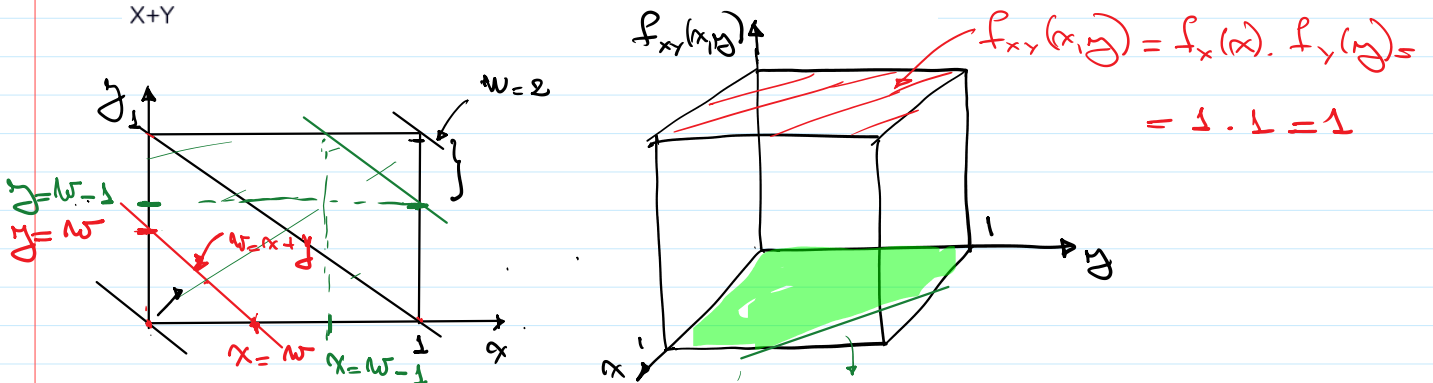
$$\frac{\mu}{\mu+\lambda} \quad 1 - \frac{\mu}{\mu+\lambda} = \frac{\mu+\lambda-\mu}{\mu+\lambda} = \frac{\lambda}{\mu+\lambda}$$

$$\textcircled{*} = \frac{e^{-(\mu+\lambda)}}{w!} \cdot (\mu+\lambda)^w \sum_{x=0}^w \frac{w!}{x! \cdot (w-x)!} \cdot \underbrace{\left(\frac{\mu}{\mu+\lambda}\right)^x \cdot \left(\frac{\lambda}{\mu+\lambda}\right)^{w-x}}_{=1} = 1, X \sim \text{Bi}(w, \frac{\mu}{\mu+\lambda})$$

$$P_W(w) = \frac{(\mu+\lambda)^w}{w!} \cdot e^{-(\mu+\lambda)} \rightarrow W \sim \text{Po}(\mu+\lambda)$$

## Ejercicio 8

Sean  $X, Y \sim U(0,1)$  e independientes. Hallar la función de densidad de  $W = X+Y$



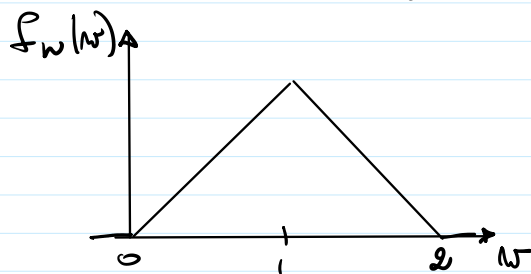
$$F_W(w) = \frac{w^2}{2} \rightarrow \text{si } 0 < w < 1$$

$$F_W(w) = 1 - \frac{[1-(w-1)]^2}{2} \rightarrow \text{si } 1 \leq w < 2$$

$$F_w(w) = 1 - \frac{[1 - (w-1)]^2}{2} \rightarrow \text{si } 1 \leq w < 2$$

$$F_w(w) = \frac{w^2}{2} \mathbb{I}\{0 < w < 1\} + 1 - \frac{(2-w)^2}{2} \mathbb{I}\{1 \leq w < 2\} + \mathbb{I}\{w \geq 2\}$$

$$f_w(w) = w \mathbb{I}\{0 < w < 1\} + (2-w) \mathbb{I}\{1 \leq w < 2\}$$



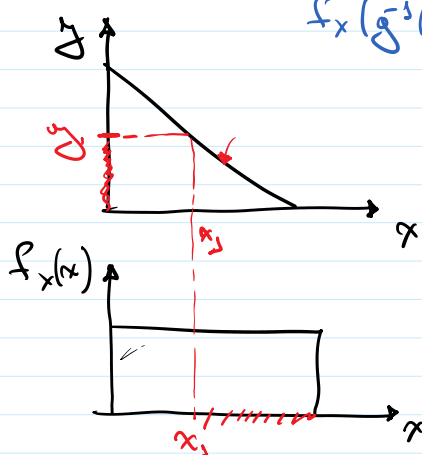
- Sea  $X$  una v.a.c. con función de densidad  $f_X(x)$ ,
- Sea  $Y=g(X)$ .
- $g(x)$  es una función 1 a 1 (existe  $g^{-1}(y)$ )

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{dg^{-1}(y)}{dy} \right| \quad \checkmark$$

$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) = P(X \leq g^{-1}(y)) = F_X(g^{-1}(y))$$

$$f_Y(y) = \frac{dF_X(g^{-1}(y))}{dx} \cdot \frac{dg^{-1}(y)}{dy}$$

$f_X(g^{-1}(y))$



$$F_Y(y) = P(Y \leq y) = 1 - P(X \leq g^{-1}(y)) = 1 - F_X(g^{-1}(y))$$

$$f_Y(y) = \frac{d[1 - F_X(g^{-1}(y))]}{dx} \cdot \frac{dg^{-1}(y)}{dy}$$

$$f_Y(y) = -f_X(g^{-1}(y)) \cdot \frac{dg^{-1}(y)}{dy}$$

no es negativo porque  $X$  e  $Y$  tienen una relación inversa

Entonces en genl.:  $f_Y(y) = f_X(g^{-1}(y)) \cdot \left| \frac{dg^{-1}(y)}{dy} \right|$

## Método del Jacobiano

### Ejercicio 1

Sean  $X_1, X_2 \stackrel{i.i.d}{\sim} \mathcal{E}(\lambda)$  y sean  $U = X_1 + X_2$  y  $V = \frac{X_1}{X_1 + X_2}$ . Hallar  $f_{U,V}(u, v)$  ¿Qué puede decir al respecto?

$$u > 0 \text{ y } v > 0$$

$$X_1 \sim \mathcal{E}(\lambda), X_2 \sim \mathcal{E}(\lambda) \rightarrow \text{ind.}, f_{X_1, X_2}(x_1, x_2) = \lambda e^{-\lambda x_1} \cdot \lambda e^{-\lambda x_2} = \lambda^2 \cdot e^{-\lambda(x_1 + x_2)}$$

$$u = h_1(x_1, x_2), v = h_2(x_1, x_2) \rightarrow x_1 = h_1^{-1}(u, v), x_2 = h_2^{-1}(u, v)$$

$$\begin{cases} u = x_1 + x_2 \Rightarrow x_1 = u - x_2 = u - u + u \cdot v = u \cdot v \\ v = \frac{x_1}{x_1 + x_2} \Rightarrow v = \frac{u - x_2}{u - x_2 + x_2} = \frac{u - x_2}{u} = 1 - \frac{x_2}{u} \Rightarrow \end{cases}$$

$$\frac{dh_1^{-1}(u, v)}{du} \Rightarrow x_1 = (1 - v)u = u - u \cdot v$$

$$J = \det \begin{pmatrix} v & u \\ \frac{dh_2^{-1}(u, v)}{du} & -u \end{pmatrix} = -u \cdot v - u(1 - v) = -u$$

$$f_{UV}(u, v) = f_{X_1, X_2}(x_1, x_2) \bigg|_{\substack{x_1 = h_1^{-1}(u, v) \\ x_2 = h_2^{-1}(u, v)}} \cdot |J|$$

$$f_{UV}(u, v) = \lambda^2 \cdot e^{-\lambda(x_1 + x_2)} \mathbb{I} \{x_1 > 0, x_2 > 0\} \bigg|_{\substack{x_1 = u \cdot v \\ x_2 = u - u \cdot v}} \cdot |J|$$

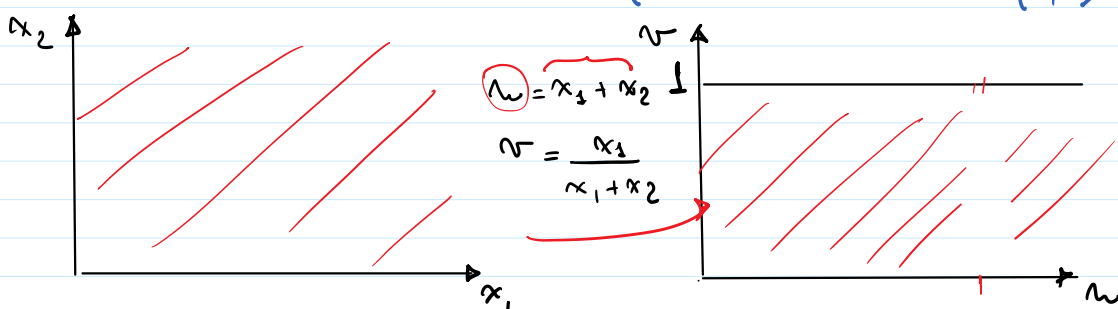
$$f_{UV}(u, v) = \lambda^2 \cdot e^{-\lambda(u \cdot v + u - u \cdot v)} \cdot u \mathbb{I} \{u \cdot v > 0, u - u \cdot v > 0\}$$

$$f_{UV}(u, v) = \underbrace{u \cdot \lambda^2 \cdot e^{-\lambda u}}_{\Gamma(2, \lambda)} \mathbb{I} \{u > 0, 0 < v < 1\}$$

$$Y \sim \Gamma(k, \lambda) \rightarrow f_Y(y) = \frac{\lambda^k}{\Gamma(k)} \cdot y^{k-1} \cdot e^{-\lambda y}$$

$\Gamma(k) = (k-1)!$   
↓  
si es entero

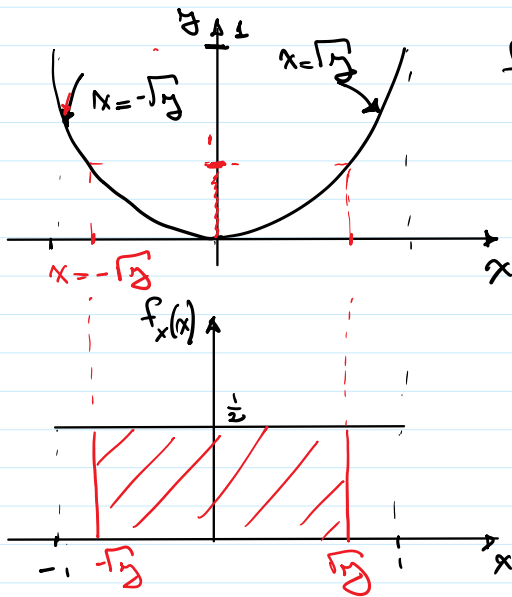
$$f_{UV}(u, v) = \underbrace{\lambda^2 \cdot u \cdot e^{-\lambda u}}_{U \sim \Gamma(2, \lambda)} \mathbb{I} \{u > 0\} \cdot \underbrace{1 \mathbb{I} \{0 < v < 1\}}_{V \sim U(0, 1)}$$



## Ejercicio 2

Sea  $X \sim U(-1,1)$  e  $Y=X^2$ . Hallar por el método del Jacobiano generalizado la función de densidad de  $Y$ .

$$Y = g(X) \rightarrow \sim \Leftrightarrow \text{biyectivo}$$



$$f_Y(y) = \sum_{i=1}^2 \frac{f_X(x_i)}{\left| \frac{dg(x_i)}{dx} \right|} \Big|_{x=g_i^{-1}(y)}$$

$$f_Y(y) = \frac{\frac{1}{2}}{|2x|} \Big|_{x=-\sqrt{y}} + \frac{\frac{1}{2}}{|2x|} \Big|_{x=\sqrt{y}} =$$

$$= \frac{1}{4\sqrt{y}} + \frac{1}{4\sqrt{y}} \mathbb{I}\{0 < y < 1\} =$$

$$= \frac{1}{2\sqrt{y}} \mathbb{I}\{0 < y < 1\}$$