

ℓ_2 -learning of Gaussian graphical models – miscellanea

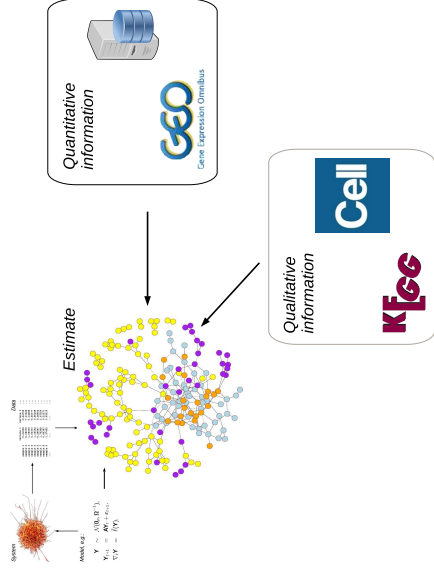
Wessel N. van Wieringen

Dept. of Epidemiology & Data Science, Amsterdam UMC
Dept. of Mathematics, Vrije Universiteit Amsterdam
Amsterdam, the Netherlands

IBC pre-conference course, Riga, 10.07.2022



Aim



Illustration

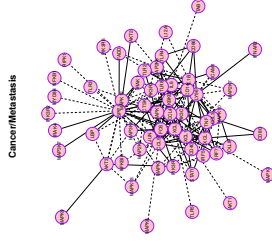
Context:

- Early detection is the best strategy against cancer.
- More information on primary than metastasized cancers.
- Primary cancer is precursor stage to metastasis.

Lung cancer:

- TCGA study,
- 111 cancers / 87 metastases,
- Toll-like receptor signalling pathway of 87 genes,
- Expression and copy number.

Regular ridge estimate of the metastasis precision is identical to the cancer precision estimate, i.e. the target is perfect.



Network reconstruction

Network

- *CIG* : graph $\mathcal{G} = (\mathcal{V}, \mathcal{E})$ of edges \mathcal{E} that represent CIs among variables $\mathcal{V} = \{1, \dots, p\}$.
- *GGM* : $\mathbf{Y} \sim \mathcal{N}(\mathbf{0}_p, \Omega^{-1})$ with precision matrix Ω .
- *CIs* : $(\Omega)_{1,2} = 0 \iff \mathbf{Y}_1 \perp\!\!\!\perp \mathbf{Y}_2 \mid \mathbf{Y}_3, \dots, \mathbf{Y}_p$.

Reconstruction

- Data : $Y_1, \dots, Y_n \sim_{i.i.d.} \mathcal{N}(\mu_p, \Omega^{-1})$.
- Estimation : $\hat{\Omega}_{ML} = S^{-1}$ with $S = n^{-1} \sum_{i=1}^n Y_i Y_i^T$.
- Inference : obtain $\hat{\rho}(Y_1, Y_2 | Y_3, \dots, Y_p)$ from $\hat{\Omega}_{ML}$ and evaluate its likelihood under H_0 .

Problem

The generalized ridge estimator of Ω , combining

- i) non-zero target shrinkage, and
- ii) element-wise penalization,

is:

$$\hat{\Omega}(\Lambda, \mathbf{T}) = \arg \max_{\Omega \succ 0} \log(|\Omega|) - \text{tr}(S\Omega) - \|\sqrt{\Lambda} \circ (\Omega - \mathbf{T})\|_F^2,$$

with

- a) target $T \in S_+^p$,
- b) positive and symmetric penalty matrix Λ , and
- c) \sqrt{B} the Hadamard square root.

If $\Lambda = \lambda 1_{pp}$, we obtain the regular ridge precision estimator.

In general, no analytic expression available. The estimator is evaluated by row-by-row updating.

Illustration: structure of Λ

Incorporate molecular biology assumptions in penalty matrix:

$$\Lambda = \begin{pmatrix} 0 & \lambda_{ge} & \dots & \lambda_{ge} & \lambda_{cn} & \dots & \infty & \dots & \infty \\ \lambda_{ge} & 0 & \dots & \lambda_{ge} & \infty & \dots & \infty & \dots & \infty \\ \dots & \dots & \dots & \dots & \lambda_{cn} & \dots & \infty & \dots & \infty \\ \lambda_{ge} & \lambda_{ge} & \dots & 0 & \infty & \dots & \infty & \dots & 0 \\ \lambda_{cn} & \dots & \dots & \dots & 0 & \dots & \infty & \dots & \infty \\ \infty & \lambda_{cn} & \dots & \lambda_{cn} & \infty & \dots & \infty & \dots & \infty \\ \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots \\ \infty & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \infty \end{pmatrix}$$

Choose λ_{cn} and λ_{ge} by 5-fold cross-validation.

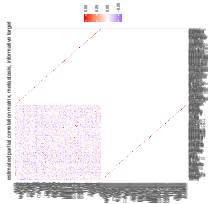
Illustration: comparison

Cross-validated optimal penalty parameters:

- cancer : $(\lambda_{\text{ge, opt}}, \lambda_{\text{cn, opt}}) = (0.001495, 0.000752)$,
- metastasis : $(\lambda_{\text{ge, opt}}, \lambda_{\text{cn, opt}}) = (0.013933, 14.76048)$.

Difference in λ 's due to:

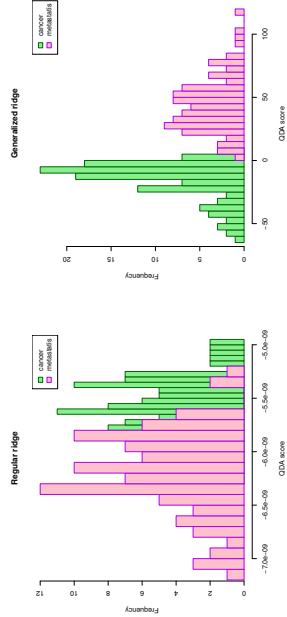
- Sample size,
- DNA-RNA interaction more stable than within expression level,
- Target indeed informative?



Navigation icons

Illustration: results

Quadratic discriminant analysis (QDA) scores with Ω estimates:



Navigation icons

The Alzheimer metabolite study revisited

A re-analysis

```
# activate packages
library("rags2ridges")
library("porridge")

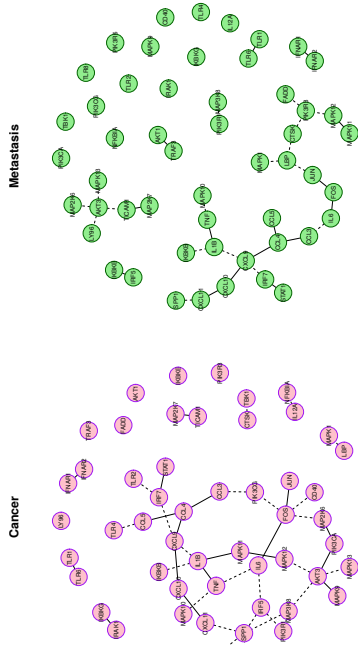
# load, extract and scale data for AD Class 2
data("ADdata")
ADclass1 <- scale(t(ADmetabolites[,
  sampleInfo$ApoECClass=="Class_1"]))
ADclass2 <- scale(t(ADmetabolites[,
  sampleInfo$ApoECClass=="Class_2"]))

# precision matrix estimation with L00CV
Phat2 <- optPenalty.kCVauto(ADclass2, ...) $optPrec
```

One could sparsify $\hat{\Omega}_{AD \text{ class } 2}$.

Navigation icons

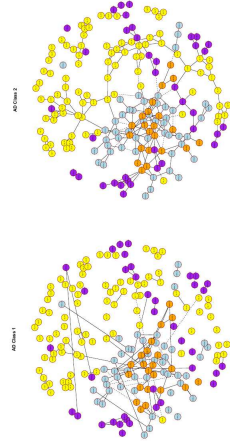
Illustration: results



Navigation icons

The Alzheimer metabolite study revisited

Result from fused ridge estimation:



Take-away : compound-type matters.

Sample size : 40 (# AD class 1) vs. 87 (# AD class 2).

Thought : can $\hat{\Omega}_{AD \text{ class } 2}$ inform $\hat{\Omega}_{AD \text{ class } 1}$?

Navigation icons

The Alzheimer metabolite study revisited

Find optimal group-specific penalty parameters:

```
# make group indicator
groups <- colnames(ADclass1)
groups[grep("Amine", colnames(ADclass1))] <- 1
...
groups <- as.integer(groups)

# find optimal lambdas
optLambdas <- optPenaltyPgen.kCVauto.groups(ADclass1,
  target = Phat2,
  groups = groups,
  penalize.diag = FALSE,
  ...)
```

Warning: takes forever. Spoiler provided later.

Navigation icons

The Alzheimer metabolite study revisited

Have a look at the penalty parameter matrix.

```
# construct generalized penalty matrix
lambdaMat <- c(rep(optLambdas[1], sum(groups==1)),
...)
lambdaMat <- matrix(lambdaMat,
  ncol=length(groups),
  nrow=length(groups))
lambdaMat <- (lambdaMat + t(lambdaMat))/2
diag(lambdaMat) <- 0
# plot generalized penalty matrix
edgeHeat(lambdaMat)
```

Navigation icons

The Alzheimer metabolite study revisited

Q: plot penalty matrix.

$(\lambda_{\text{amine}}^{(\text{opt})}, \lambda_{\text{org.acid}}^{(\text{opt})}, \lambda_{\text{lipid}}^{(\text{opt})}, \lambda_{\text{ox.stress}}^{(\text{opt})}) = (1.7095, 4.4945, 22.1870, 12.3487)$

Q: fit a GGM to the AD class 1 data by means of the generalized ridge estimate.

Q: sparsify the estimated precision.

Q: plot the CIG.

Q: compare to the CIGs of the two classes.

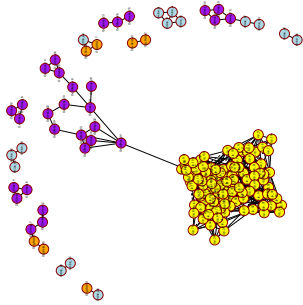
Q: also use a zero target, i.e. $T = 0_{pp}$.

$(\lambda_{\text{amine}}^{(\text{opt})}, \lambda_{\text{org.acid}}^{(\text{opt})}, \lambda_{\text{lipid}}^{(\text{opt})}, \lambda_{\text{ox.stress}}^{(\text{opt})}) = (11.9672, 12.6888, 0.2882, 6.4313)$

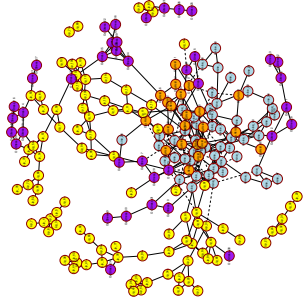
Navigation icons

The Alzheimer metabolite study revisited

CIG, $T = 0_{pp}$



CIG, $T = \hat{\Omega}_{AD \text{ class } 2}$.



Navigation icons

The Alzheimer metabolite study revisited

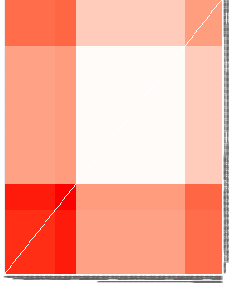
Find $\hat{\Omega}_{AD \text{ class } 1}$, sparsify, and plot CIG.

```
# fit precision matrix
Phat1 <- ridgePgen(covML(ADclass1),
  lambda=lambdaMat,
  target=Phat2)
# extract the CIG
Phat1 <- sparsify(Phat1, ...)
# visualize the CIG
PcorP1 <- pruneMatrix(Phat1$sparseParCor)
Colors <- rownames(PcorP1)
Colors[grep("Amine", rownames(PcorP1))] <- "lightblue"
...
# plot network
Ugraph(PcorP1, ...)
```

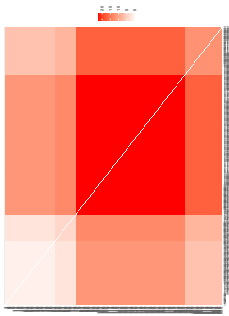
Navigation icons

The Alzheimer metabolite study revisited

Penalty matrix, $T = 0_{pp}$



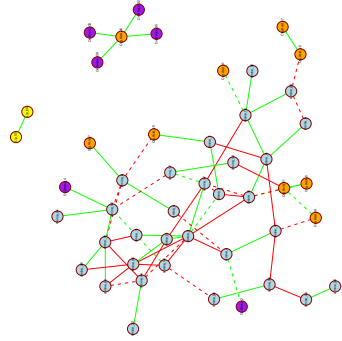
Penalty matrix, $T = \hat{\Omega}_{AD \text{ class } 2}$.



Navigation icons

The Alzheimer metabolite study revisited

Differential Network



Differences appear to be among the amines (the lightblue folks).

Navigation icons

Are reconstructed molecular networks reproducible?

Network reconstruction only considers sampling variation, while:

“An element of chance enters into every measurement; hence every set of measurements is inherently a sample of certain more or less unknown conditions. Even in those few instances where we believe that the objective reality being measured is constant, the measurements of this constant are influenced by chance or unknown causes.”

– Shewhart, 1931, p. 378.

Additive error model

Replicated data of i -th sample: $\{Y_{i,k_j}\}_{k_j=1}^{K_j}$.

A “signal + noise”-model:

$$\begin{aligned} Y_{i,k_j} &= Z_i + \epsilon_{i,k_j} \\ Z_i &\sim_{i.i.d.} \mathcal{N}(0_p, \Omega_z^{-1}), \\ \epsilon_{i,k_j} &\sim_{i.i.d.} \mathcal{N}(0_p, \Omega_\epsilon^{-1}), \end{aligned}$$

and $Z_i \perp \epsilon_{i,k_j}$.

Hence,

$$Y_{i,k_j} \sim_{i.i.d.} \mathcal{N}(0_p, \Omega_z^{-1} + \Omega_\epsilon^{-1}).$$

Estimation

Penalized maximum likelihood estimators:

$$\widehat{\Omega}_z(\lambda_z), \widehat{\Omega}_\epsilon(\lambda_\epsilon) = \arg \max_{\Omega_z, \Omega_\epsilon} \mathcal{L}(Y; \Omega_z, \Omega_\epsilon) - \frac{1}{2} \lambda_z \|\Omega_z\|_2^2 - \frac{1}{2} \lambda_\epsilon \|\Omega_\epsilon\|_2^2,$$

with a diagonal Ω_ϵ :

$$\widehat{\Omega}_z(\lambda_z), \widehat{\Omega}_\epsilon = \arg \max_{\Omega_z, \Omega_\epsilon} \mathcal{L}(Y; \Omega_z, \Omega_\epsilon) - \frac{1}{2} \lambda_z \|\Omega_z\|_2^2,$$

a lasso penalty:

$$\widehat{\Omega}_z(\lambda_z), \widehat{\Omega}_\epsilon = \arg \max_{\Omega_z, \Omega_\epsilon} \mathcal{L}(Y; \Omega_z, \Omega_\epsilon) - \lambda_z \|\Omega_z\|_1.$$

All found by a penalized EM algorithm (= linear algebra fun).

Effect of noise

Ignorance of the noise introduces false positive/negative edges.
For instance, take $\Omega_\epsilon^{-1} = \mathbf{I}_3$ and

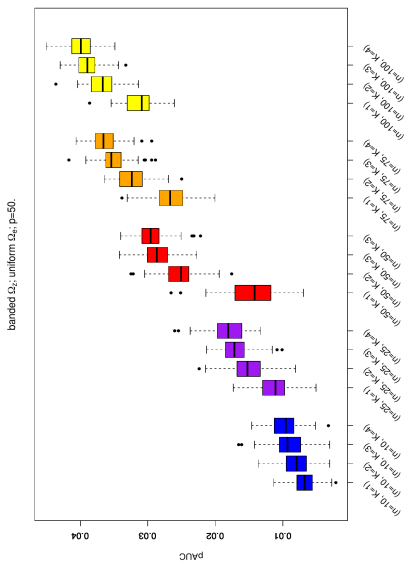
$$\Omega_z^{-1} = \begin{pmatrix} 3 & -1 & 2 \\ -1 & 3 & -2 \\ 2 & -2 & 4 \end{pmatrix}.$$

Then, $(\Omega_z)_{1,2} = 0$ but $[(\Omega_z^{-1} + \Omega_\epsilon^{-1})^{-1}]_{1,2} \neq 0$.

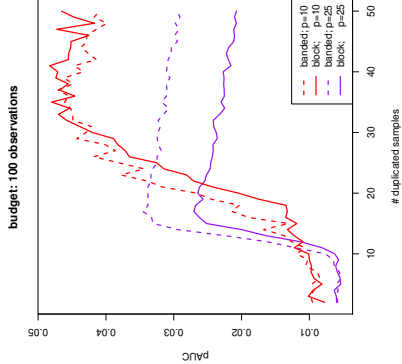
Generally,

$$\Omega_y = (\Omega_z^{-1} + \Omega_\epsilon^{-1})^{-1} = \underbrace{\Omega_z - (\mathbf{I}_{pp} + \Omega_z \Omega_\epsilon^{-1})^{-1} \Omega_z \Omega_\epsilon^{-1} \Omega_z}_{\propto \text{edge strengths diff. between } \Omega_y \text{ and } \Omega_z}.$$

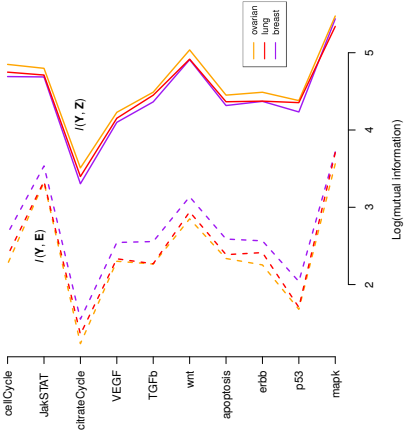
Simulation: effect of K and n



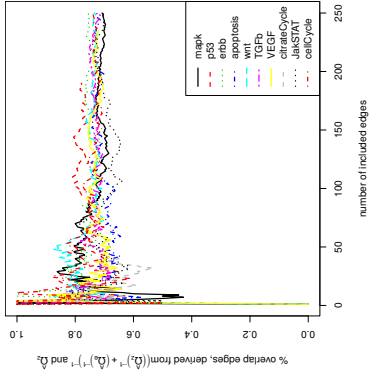
Fixed # measurements



a) Is there any signal?



a) Intersection $\hat{\mathcal{E}}_y$ and $\hat{\mathcal{E}}_z$



Illustration

Issues studied:

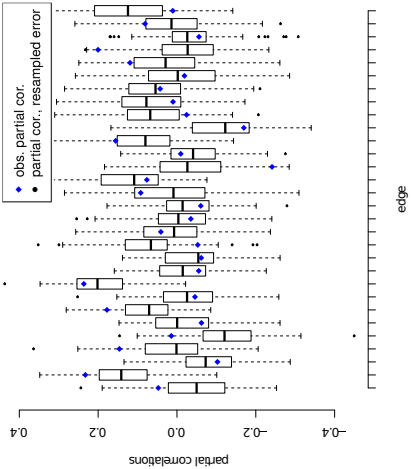
- Effect on reconstruction of an error-diluted signal.
- Tenability of the diagonal error assumption.
- Network differences between replicated and non-replicated data.

Means:

- Three TCGA studies
- Ten pathways: $p \in [29, 247]$
- Technical replicates: microarray and RNAseq

tissue	sample size n
lung	151
breast	526
ovarian	294

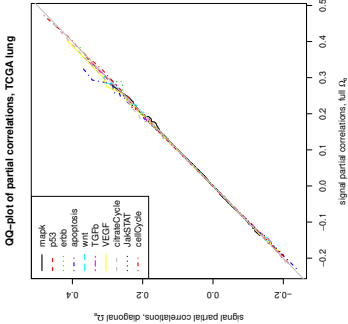
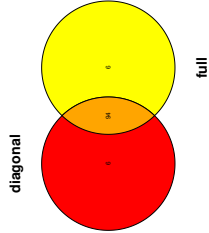
a) Error effect on partial correlations



b) Full vs. diag. Ω_ε

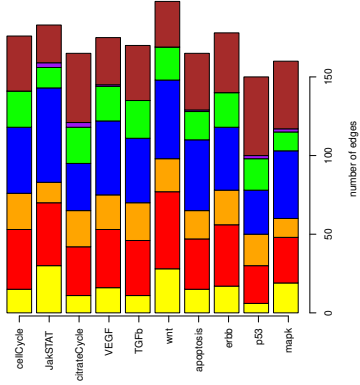
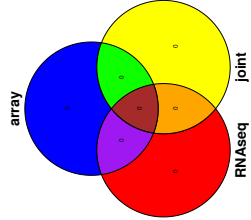
QQ-plot of TCGA lung \rightarrow

\downarrow top 100 $\hat{\mathcal{E}}_{full} \cap$ top 100 $\hat{\mathcal{E}}_{diag}$



c) Intersection $\hat{\mathcal{E}}_{seq}$, $\hat{\mathcal{E}}_{array}$ and $\hat{\mathcal{E}}_{joint}$

TCGA lung
→ 100 edges per $\hat{\mathcal{E}}$.



References

- These slides are based on:
- van Wieringen, W. N. (2019). The generalized ridge estimator of the inverse covariance matrix *Journal of Computational and Graphical Statistics*, 28(4), 932-942.
 - van Wieringen, W.N., Chen, Y. (2021). Penalized estimation of the Gaussian graphical model from data with replicates. *Statistics in Medicine*, 40(19), 4279-4293.
 - van Wieringen, W.N. (2022) porridge: Ridge-Type Estimation of a Potpourri of Models. R package version 0.3.1. <https://CRAN.R-project.org/package=porridge>.

Conclusion

The generalized ridge precision estimator provides ways to incorporate detailed and structured prior knowledge, both:
→ *quantitatively*, via the target matrix; and
→ *qualitatively*, by the parametrization of the penalty matrix.

But ...
... reconstructed networks should be taken with reservation, if measurement error is ignored.

License

This material is provided under the Creative Commons Attribution / Share-Alike / Non-Commercial License.



See <http://www.creativecommons.org> for details.