# Thoughts on Epidemic Compartment Models and their Markovian Origin

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#### Introduction

https://www.ncbi.nlm.nih.gov/pmc/articles/PMC6001967/

- Epidemics have been around since the beginning of recorded history:
  - "Spanish" flu epidemic of 1918–19 caused more than 50,000,000 deaths worldwide.
  - The Black Deaths (probably bubonic plague) spread from Asia throughout 1346-1666, and is estimated to have caused the death of as much as one-third of the population of Europe between 1346 and 1350.
  - Modern flu's: Birdflu (Avian) 2013, Pigflu
  - Corona: SARS (2003), MERS (2012), CoVID-19

#### Introduction

https://www.ncbi.nlm.nih.gov/pmc/articles/PMC6001967/

#### • Epidemic models:

- Bernoulli D. 1766. Essai d'une nouvelle analyse de la mortalité causeé par la petite vérole. Mem. Math. Phys. Acad. Roy. Sci. Paris. life expectancy
- In 1906 W.H. Hamer proposed that the spread of infection should depend on the number of susceptible individuals and the number of infective individuals (<a href="Hamer, 1906">Hamer, 1906</a>). He suggested a mass action law for the rate of new infections, and this idea has been basic in compartmental models since that time.
- COVID-19-CTRL: DTU, AAU, Novo Nordisk

## Agenda

- Introduction
- Model types
- Derivation of models
- Comparison examples
- Conclusion

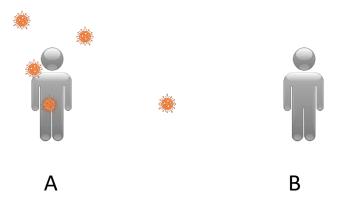
#### Analysis tools for epidemic spread

- CTMC: State dimension = number of desease states population size
- Aggregated CTMC = population size number of desease states
- Compartment models: SIR,SEIR,SIS,SIRS ... (SI)
   Differential equations: ODE -> mean value evolution
- Compartment models with Poisson process noise (TK)
   SDE > variance analysis
- Agent based simulation (KDH)
   Monte Carlo simulation, full CTMC

- Consider two individuals x and y, where y is infected and x is susceptible.
- Within a time interval [t,t+dt] the 2 individuals are within *infection* range and the probability that infection happens is assumed to be

b dt

(smooth, memoryless, time-invariant, CTMC)



- Consider a susceptible individual x within a group of I infected individuals all within infection range.
- The probability that infection happens within [t,t+dt] is for small dt (disregarding multiple infection)

$$I * b dt$$

- Next consider a group of S susceptible individuals all within the same infection range with a group of I infected individuals.
- The probability that one or more susceptibles are infected within the interval is

$$1 - (1 - I * b * dt)^{S} \sim b * S * I * dt$$

• The probability that 1 susceptible is infected within [t,t+dt] is

$$b * S * I * dt$$

• The probability that more than 1 susceptibles are infected is therefore at most  $O(n^2)$ 

- A group of S susceptible individuals all within the same infection range with a group of I infected individuals.
- The probability that 1 susceptible is infected within [t,t+dt] is

$$b * S * I * dt$$

- The probability that more than 1 susceptibles are infected is therefore at most  $O(n^2)$
- Considering conditional expectations (inf generator) we have

$$E[I(t + dt) - I(t)|I(t)] = b * S * I * dt$$

$$\lim \frac{E[I(t+dt)]-E[I(t)|I(t)]}{dt} = b * S * I$$

Taking expectations (Dynkin)

$$\frac{d}{dt} E[I](t) = b * E[S * I] \sim b * E[S] * E[I]$$

#### Example – 2 families

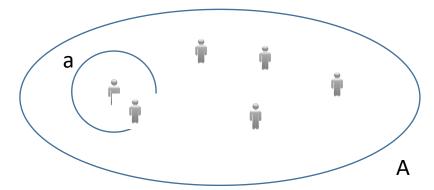
2 families each with M members

Assume E[I1]=E[I2], then with I=I1+I2,
 d/dt (E[I1] + E[I2]) = b\*E[I1]\*(M-E[I1]) + b\*E[I1]\*(M-E[I1]) = 2\*b\*E[I1]\*(M-E[I1]) =>
 d/dt E[I]= b\*E[I]\*(M-E[I]/2)

• Merge now the 2 families into a single family with 2M members, we have  $d/dt E[I] \sim b*E[I]*(2M-E[I]) = 2*b*E[I]*(M-E[I]/2)$ 

- The overall infection pressure (rate) is d/dt E[I]
- Thus, the infection rate is doubled by merging the families. !!
- Underlying assumption is that everybody are within infection range.

- Consider S susceptible within a group of I infected confined to an area A.
- We denote the area around x within infection range as a.



- Thus, the expected number of infected within a is not E[I] but  $E[I] \frac{|a|}{|A|}$
- This gives the following mean value dynamics

$$\frac{d}{dt} E[I](t) \sim b * E[S] * E[I] \frac{|a|}{|A|}$$

- Consider S susceptibles within a group of I infected
- Assume that for a relevant range of population-densities the number of people within the infection range of any susceptible is fixed to L.
- Then the mean value dynamics become (with N=S+I)

$$\frac{d}{dt} E[I](t) \sim b * E[S] * E[I] \frac{L}{N} = \frac{b * L}{N} * E[S] * E[I] = \frac{\beta}{N} * E[S] * E[I]$$

Normalizing I = N\*i, S = N\*s leads to

$$\frac{d}{dt} E[i](t) = \beta * E[s] * E[i]$$

#### Example – 2 families

2 families each with M members

d/dt E[I1] 
$$\sim \frac{\beta}{M} * E[I1] * (M-E[I1])$$
  
d/dt E[I2]  $\sim \frac{\beta}{M} * E[I2] * (M-E[I2])$ 

• Assume E[I1]=E[I2], then with I=I1+I2,

d/dt (E[I1] + E[I2]) = 
$$\frac{\beta}{M}$$
\*E[I1]\*(M-E[I1]) +  $\frac{\beta}{M}$ \*E[I1]\*(M-E[I1]) = 2\* $\frac{\beta}{M}$ \*E[I1]\*(M-E[I1]) => d/dt E[I]= $\frac{\beta}{M}$ \*E[I]\*(M-E[I]/2)

• Merge now the 2 families into a single family with 2M members, we have

d/dt E[I] 
$$\sim \frac{\beta}{2M}$$
\*E[I]\*(2M-E[I]) = 2\* $\frac{\beta}{2M}$ \*E[I]\*(M-E[I]/2) = =  $\frac{\beta}{M}$ \*E[I]\*(M-E[I]/2)

- Thus, the infection rate is unchanged by merging the families. !!
- Underlying assumption is that a fixed number of people is within infection rate.

#### 2 different models

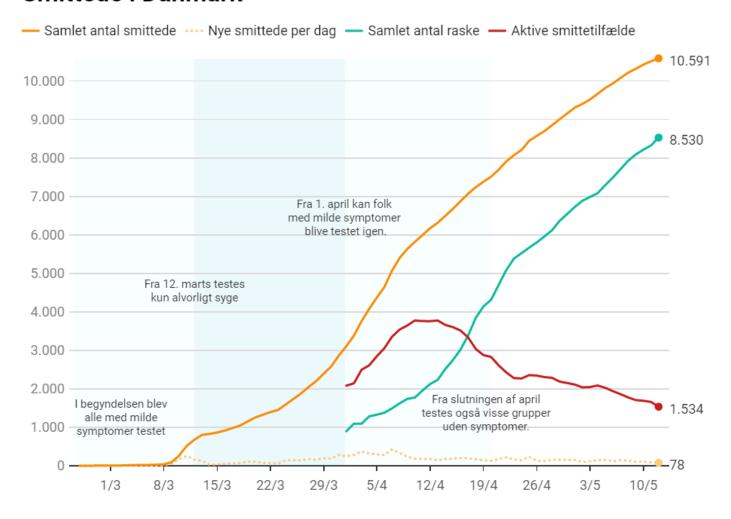
- Which model is correct?
- Which model should be used in various situations:
  - The entire DK population
  - Herning warm blood stallion show (HWSS)
  - Families
  - Schools
  - •
- Combined model

$$\beta = b * (L + \alpha \rho)$$

#### Claim

- 5 day event 70.000 participants
- HWSS is just merging 12.000/5 = 2400 families
- With the 1<sup>st</sup> model the infection rate is multiplied by 2400 !!!
- The number of people within infection range is not higher at HWSS than in a common family ..
- Thus merging 2400 families does not make a change
- Which is true?

#### Smittede i Danmark



Det er ikke alle smittede, der har opsøgt lægehjælp eller er blevet testet. Langt flere kan derfor reelt være eller have været smittet. // Raske er defineret ved at være personer, som ikke er hverken indlagt eller døde, 14 dage efter de fik konstateret COVID-19 // Aktive smittetilfælde er det samlede antal registrerede smittede fratrukket raske og døde // Tallene for det seneste døgn er muligvis ikke endelige. Der kan komme nye og højere tal senere på dagen.

Grafik: Videnskab.dk • Kilde: Statens Serum Institut • Hent data • Lavet med Datawrapper

## Sanity check 1<sup>st</sup> model

- Newly infected per day in DK: 100-200 (tested)
- Infection rate pr capita (DK): 150/6e6 = 2.5e-5 (pr day)
- Infection rate pr capita (HWSS): 2400 \* 2.5e-5 = 0.06 (pr day)
- Expected no of infected in Herning pr day: 14000\*0.06 = 840!!

# Sanity check 2<sup>nd</sup> model

1.4 Antal testede, smittede, indlagte og dødsfald fordelt på region

Region	Testede	Bekræftede COVII	O-19 smittede Indla	agte* Dødsfald**
Nordjyll	and 34.205	421	99	21
Midtjylla	nd 64.607	1.544	275	69
Syddanı	nark 61.810	921	228	29
Hovedst	aden 125.135	5.958	1.082	2 303
Sjælland	47.484	1.689	448	105
I alt	333.241	10.533	2.132	2 527

Density or meeting different people seems to matter

#### More precise SIR model for low N

- State:  $x(t) = \{s_1, ..., s_N\} \in X$
- $s_i \in \{SS, II, RR\}$
- State space:  $X = \{SS, II, RR\}^N$  (large  $3^N$  elements)
- $S_x(t) = \#\{s \in x(t) \mid s = SS\}$
- $I_x(t) = \#\{s \in x(t) \mid s = II\}$
- $\bullet \ R_{\mathcal{X}}(t) = \#\{s \in \mathcal{X}(t) \mid s = RR\}$

#### Markovian property

- $X_t = \{ f: [0, t] \to \{SS, II, RR\} \}$
- y,  $x_t \in X_t$ ,  $x_t(\tau) = x(\tau)$  for  $\tau \in [0, t]$

$$P(x(t+dt)=\{s_1,...,s_i',...,s_N\} \mid x_t=y) = \Gamma(x(t)) =$$

$$P(x(t+dt)=\{s_1,...,s_i',...,s_N\} \mid x(t)=\{s_1,...,s_i,...,s_N\})$$

#### Transition probabilities

•  $P(x(t+dt)=\{s_1,...,s_i,...,s_N\} \mid x(t)=\{s_1,...,s_i,...,s_N\})$ 

- $P(x(t+dt)=\{s_1,..,II,..,s_N\} \mid x(t)=\{s_1,..,SS,..,s_N\}) = b/N * I_x(t) * dt$
- $P(x(t+dt)=\{s_1,...,RR,...,s_N\} \mid x(t)=\{s_1,...,II,...,s_N\}) = \alpha * I_x(t) * dt$

#### Aggregate model

- State space:  $X = \{SS, II, RR\}^N$  (large  $3^N$  elements)
- $A_{s,i,r} = \{x \in X \mid S_x(t) = s, I_x(t) = i, R_x(t) = r\}$
- $X_A = \{A_{s,i,r}\}$  is a partition of X with  $\frac{N^2}{2} + 2N + 1$  elements
- N=5 ->  $3^N = 243$ ,  $\frac{N^2}{2} + 2N + 1 \sim 34$

#### Transition probabilities

• (1) 
$$P(x(t+dt) = k \mid x(t) = p) = (g+b*i)*dt$$
  
for  $p \in A_{s,i,r}$  and  $k \in A_{s-1,i+1,r}$ 

• (2) 
$$P(x(t+dt)=k\mid x(t)=p)=\alpha*dt$$
 for  $p\in A_{s,i,r}$  and  $k\in A_{s,i-1,r+1}$ 

#### Transition probabilities — Infection

• 
$$P(x(t+dt) \in A_{s-1,i+1,r} | x(t) \in A_{s,i,r}) = P(x(t+dt) \in A_{s-1,i+1,r} | x(t) \in A_{s,i,r}) = P(x(t+dt) \in A_{s-1,i+1,r} | x(t) \in A_{s,i,r}) / P(x(t) \in A_{s,i,r})$$

$$= \sum_{p \in A_{s,i,r}} \sum_{k \in A_{s-1,i+1,r}} P(x(t+dt) = k | x(t) = p) P(x(t) = p) / P(x(t) \in A_{s,i,r})$$

$$= \sum_{p \in A_{s,i,r}} \sum_{k \in A_{s-1,i+1,r}} (g+b*i)*dt*\delta_{I}(p,k)*P(x(t) = p) / P(x(t) \in A_{s,i,r})$$

$$= (g+b*i)*dt*\sum_{p \in A_{s,i,r}} P(x(t) = p) \sum_{k \in A_{s-1,i+1,r}} \delta_{I}(p,k) / P(x(t) \in A_{s,i,r})$$

$$= (g+b*i)*dt*P(x(t) \in A_{s,i,r}) *s / P(x(t) \in A_{s,i,r})$$

$$= (g+b*i)*dt*s$$

#### Transition probabilities - Recovery

 $= \alpha * dt * i$ 

• 
$$P(x(t+dt) \in A_{s,i-1,r+1} \mid x(t) \in A_{s,i,r}) = P(x(t+dt) \in A_{s,i-1,r} \otimes x(t) \in A_{s,i,r}) / P(x(t) \in A_{s,i,r})$$

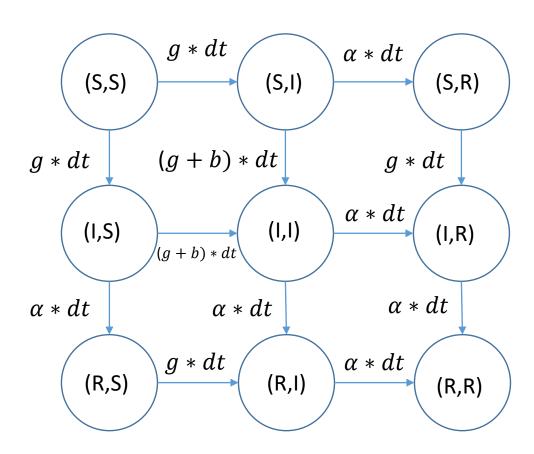
$$= \sum_{p \in A_{s,i,r}} \sum_{k \in A_{s,i-1,r+1}} P(x(t+dt) = k \mid x(t) = p) P(x(t) = p) / P(x(t) \in A_{s,i,r})$$

$$= \sum_{p \in A_{s,i,r}} \sum_{k \in A_{s,i-1,r+1}} \alpha * dt * \delta_{I}(p,k) * P(x(t) = p) / P(x(t) \in A_{s,i,r})$$

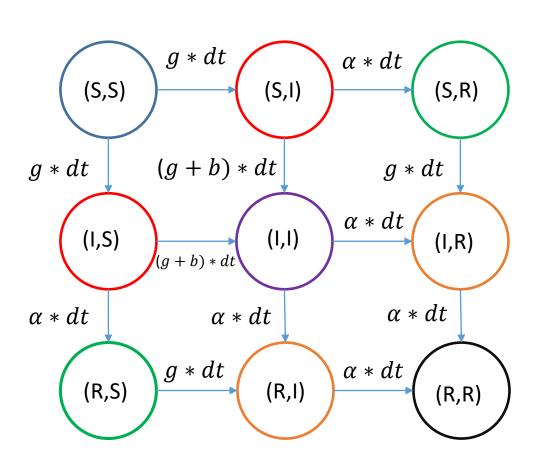
$$= \alpha * dt * \sum_{p \in A_{s,i,r}} P(x(t) = p) \sum_{k \in A_{s,i-1,r+1}} \delta_{I}(p,k) / P(x(t) \in A_{s,i,r})$$

$$= \alpha * dt * P(x(t) \in A_{s,i,r}) * i / P(x(t) \in A_{s,i,r})$$

# Example (N=2)



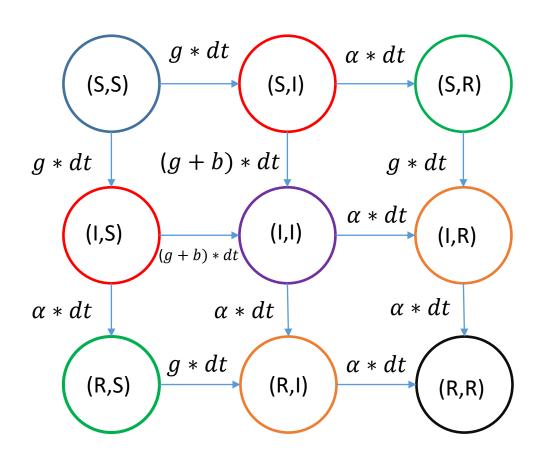
# Example (N=2)

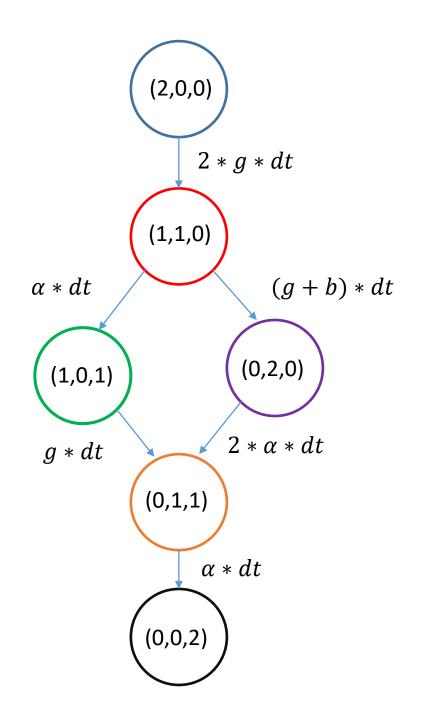


• 
$$3^N = 3^2 = 9$$

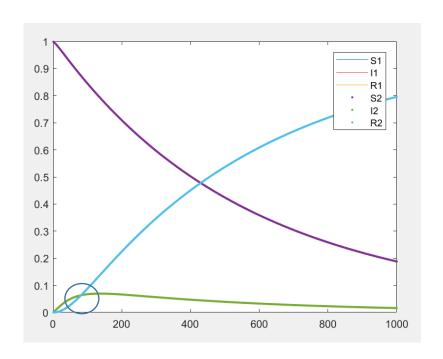
• 
$$\frac{N^2}{2} + 2N + 1 = 7$$

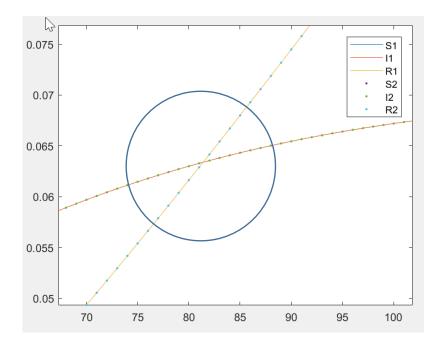
# Example (N=2)



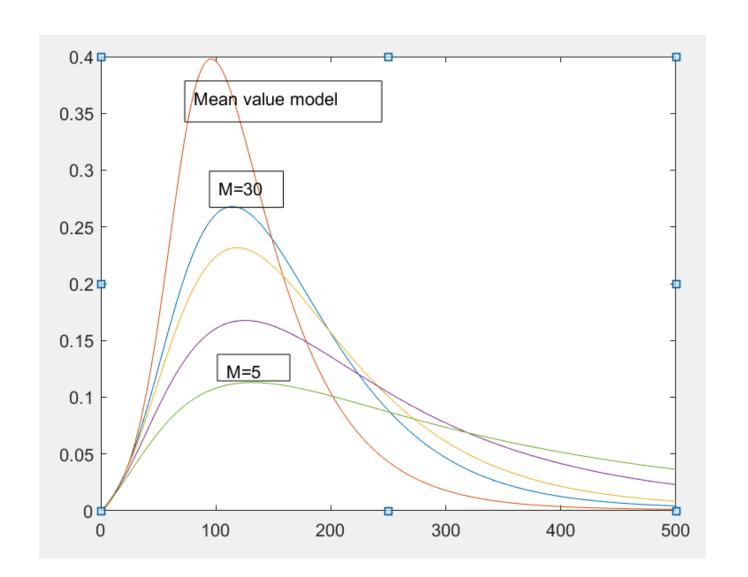


# Comparison

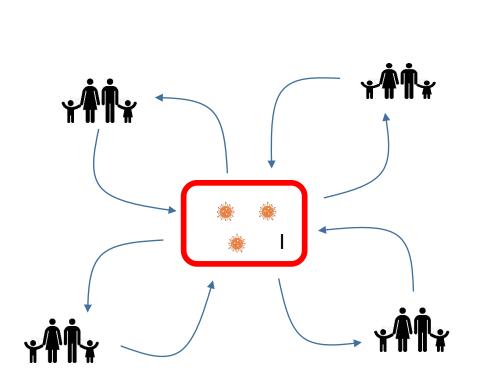


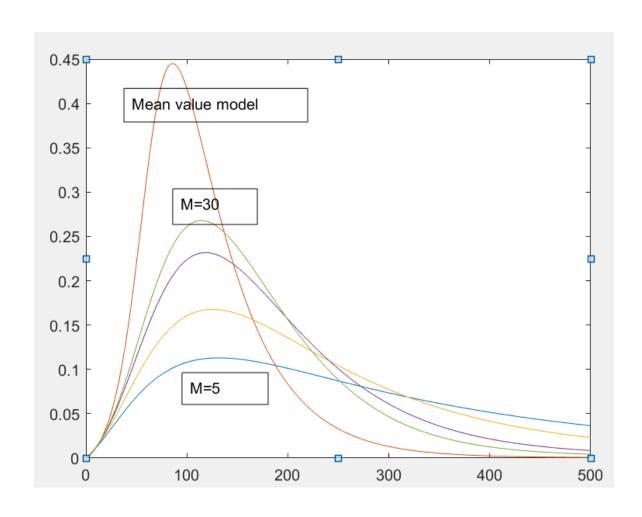


#### Mean value model vs aggregated model



# Mean value model vs aggregated model with FeedBack





#### Conclusions

- For mean value models infection rate heavily depends on model assumption especially the number N of people within infection rate
- Model with proportional N gives highly exaggerated results for large assemblies
- Model with scaled N gives counter-intuitive results for small groups
- For small groups both mean-value models overestimate infection rate
- Feedback seems to amplify the model discrepancy

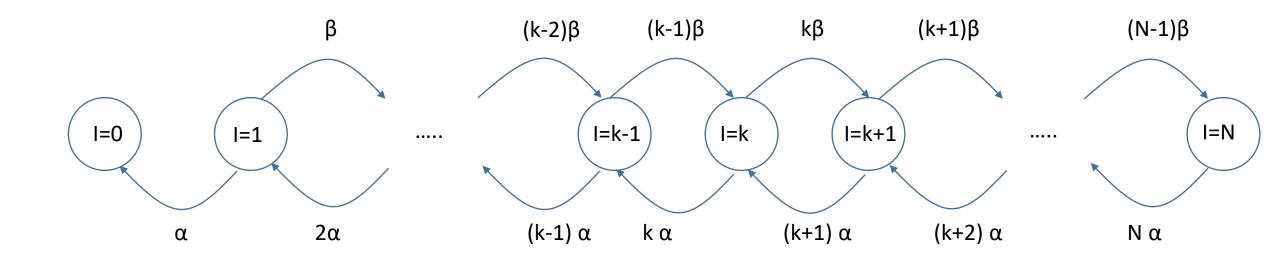
# Simpler model for Epidemic birth- and infant-death

- The number S of susceptibles is large and approximately N
- The number I of infected is small << N</li>
- Infected are spread so IL/N within infection range
- SIR model
- Transition probabilities

$$P(x(t+dt)=\{s_1,...,l_1,...,s_N\} \mid x(t)=\{s_1,...,SS,...,s_N\}) = b\frac{L}{N}*I_x(t)*dt = \beta*I_x(t)*dt$$

$$P(x(t+dt)=\{s_1,...,RR,...,s_N\} \mid x(t)=\{s_1,...,l_1,...,s_N\}) = \alpha*I_x(t)*dt$$

# Chain for only I



#### Measurement model

- Sequencing performed on a sample of positives p.
- c5 found among this

$$P(c5 \& p) = P(c5 | p) P(p) = P(c5) \Rightarrow$$
$$P(c5 | p) = P(c5)/P(p)$$

- P(c5) ~ no of c5 / no of tests (maybe X 2 for dark cases) ~ I / 6e6
- $P(p) \sim 1.4 \%$

#### Measurement model

Probability of k observations of c5 given I = m

• 
$$P(c5 \mid p) = \frac{P(c5)}{P(p)} = \frac{\frac{m}{6e6}}{1.4\%}$$

- $P(k \ obs \ of \ c5 \ | I = m) = binom(k, P(c5 \ | \ p))$
- $P(0 \text{ obs of } c5 | I = m) = binom(0, P(c5 | p)) = (1 P(c5 | p))^k$

#### Bayesian Filter

Markovian dynamics

$$P(x(1) = i(1), ...x(N) = i(N)) = P(x(1) = i(1)) P(x(2) = i(2)|x(1) = i(1)) ... P(x(N) = i(N)|x(N-1) = i(N-1))$$

Independent observation model

$$P(x(1) = i(1), ...x(N) = i(N), y(1) = j(1), ..., y(N) = j(N)) =$$
  
 $P(x(1) = i(1)) P(x(2) = i(2)|x(1) = i(1)) ... P(x(N) = i(N)|x(N-1) = i(N-1))$   
 $P(y(1) = j(1)|x(1) = i(1)) ... P(y(N) = j(N)|x(N) = i(N))$ 

#### Bayesian Filter

• 
$$P(x(1) = i(1), ...x(N) = i(N), y(1) = j(1), ..., y(N) = j(N)) = P(x(1) = i(1)) P(x(2) = i(2)|x(1) = i(1)) ... P(x(N) = i(N)|x(N-1) = i(N-1)) P(y(1) = j(1)|x(1) = i(1)) ... P(y(N) = j(N)|x(N) = i(N)) = P(x(1) = i(1)) H(i(1), i(2)) ... H(i(N-1), i(N)) G(i(1), j(1)) ... G(i(N), j(N))$$

$$P(x(N) = i(N), y(1) = j(1), ..., y(N) = j(N))$$

$$= \sum_{i(N-1)} ... \sum_{i(1)} P(x(1) = i(1), ...x(N) = i(N), y(1) = j(1), ..., y(N) = j(N))$$

$$= G(i(N), j(N)) \sum_{i(N-1)} H(i(N-1), i(N)) G(i(N-1), j(N-1))$$

$$\sum_{i(N-2)} \dots \sum P(x(1))$$

#### Bayesian Filter

$$P(x(N) = i(N), y(1) = j(1), ..., y(N) = j(N))$$

$$= \sum_{i(N-1)} ... \sum_{i(1)} P(x(1) = i(1), ..., x(N) = i(N), y(1) = j(1), ..., y(N) = j(N))$$

$$= G(i(N), j(N)) \sum_{i(N-1)} H(i(N-1), i(N))G(i(N-1), j(N-1))$$

$$\sum_{i(N-2)} ... \sum_{i(1)} P(x(1) = i(1)) H(i(1), i(2)) ... H(i(N-2), i(N-1)) G(i(1), j(1)) ... G(i(N-1), j(N-1))$$

$$= G(i(N), j(N)) \sum_{i(N-1)} H(i(N-1), i(N))G(i(N-1), j(N-1))$$

$$P(x(N-1) = i(N-1), y(1) = j(1), ..., y(N-1) = j(N-1))$$

### Bayesian Filter – conditional by normalization

$$P(x(N) = i(N), y(1) = j(1), ..., y(N) = j(N)) =$$

$$G(i(N), j(N)) \sum_{i(N-1)} H(i(N-1), i(N))G(i(N-1), j(N-1))$$

$$P(x(N-1) = i(N-1), y(1) = j(1), ..., y(N-1) = j(N-1))$$

$$P(x(N) = i(N) | y(1) = j(1),...,y(N) = j(N))) = P(x(N) = i(N),y(1) = j(1),...,y(N) = j(N)) / \sum_{i(N)} P(x(N) = i(N),y(1) = j(1),...,y(N) = j(N))$$

# Bayesian Filter – observation Likelihood by summation of joint distribution

$$P(x(N) = i(N), y(1) = j(1), ..., y(N) = j(N)) =$$

$$G(i(N), j(N)) \sum_{i(N-1)} H(i(N-1), i(N))G(i(N-1), j(N-1))$$

$$P(x(N-1) = i(N-1), y(1) = j(1), ..., y(N-1) = j(N-1))$$

$$P(y(1) = j(1), ..., y(N) = j(N)) =$$

$$\sum_{i(N)} P(x(N) = i(N), y(1) = j(1), ..., y(N) = j(N))$$