EC 421, 004

Connor Lennon Spring 2021

Prologue

R showcase

R Markdown

- Simple mark-up language for for combining/creating documents, equations, figures, R, and more
- Basics of Markdown
- E.g., **I'm bold**, *I'm italic*, I \leftarrow "code"

Econometrics with R

- (Currently) free, online textbook
- Written and published using R (and probably R Markdown)
- Warning: I haven't read this book yet.

Related: Tyler Ransom has a great cheatsheet for econometrics.

Schedule

Last Time

We wrapped up our review.

Today

Heteroskedasticity

Schedule

This week

First assignment! It will be due in ~ 2 weeks on April 24th

Turn in 2 files[†]

- 1. Your write up (e.g., Word file).
- 2. The R script that generated your answers.

Important

- We should be able to easily find your answers for each question.
- **Do not copy.** (You will receive a zero.)

Let's write down our current assumptions

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- 5. The disurbances have **constant variance** σ^2 and **zero covariance**, *i.e.*,

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 eq j
- 6. The disturbances come from a **Normal** distribution, *i.e.*, $u_i \stackrel{ ext{iid}}{\sim} \mathrm{N} ig(0, \sigma^2 ig)$.

Today we're focusing on assumption #5:

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Heteroskedasticity: $\mathrm{Var}(u_i) = \sigma_i^2$ and $\sigma_i^2 \neq \sigma_j^2$ for some $i \neq j$.

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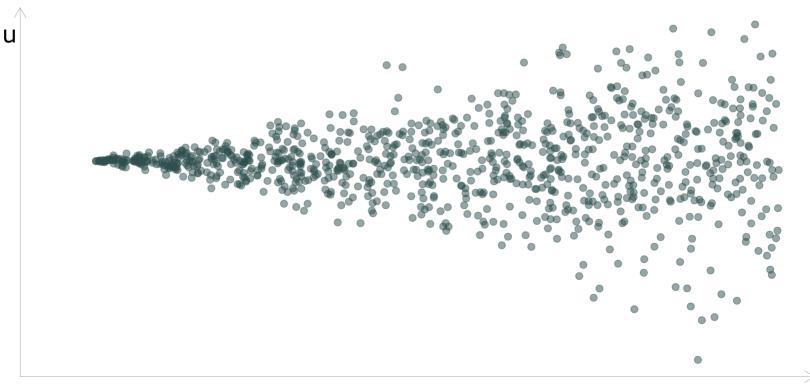
Violation of this assumption:

Heteroskedasticity: $\mathrm{Var}(u_i) = \sigma_i^2$ and $\sigma_i^2 \neq \sigma_j^2$ for some $i \neq j$.

In other words: Our disturbances have different variances.

Classic example of heteroskedasticity: The funnel

Variance of u increases with x



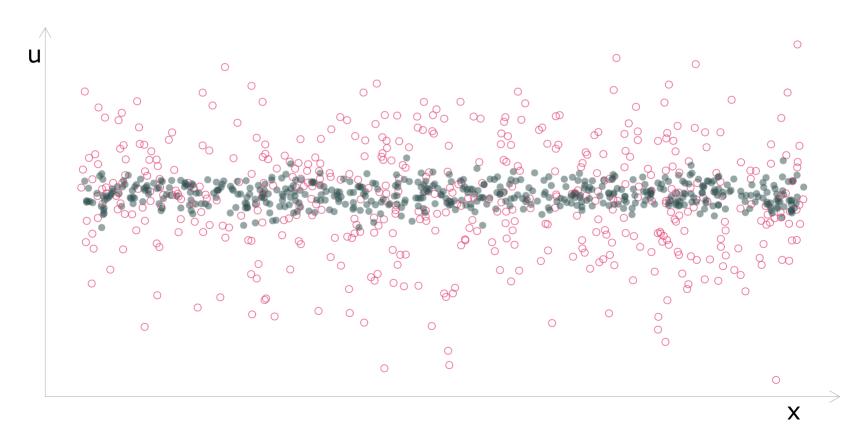
Another example of heteroskedasticity: (double funnel?)

Variance of u increasing at the extremes of x



Another example of heteroskedasticity:

Differing variances of u by group



Heteroskedasticity is present when the variance of u changes with any combination of our explanatory variables x_1 , through x_k (henceforth: X).

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Why we care: Heteroskedasticity shows us how small violations of our assumptions can affect OLS's performance.

Consequences

So what are the consquences of heteroskedasticity? Bias? Inefficiency?

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Recall₂: We previously showed
$$\hat{eta}_1 = rac{\sum_i \left(y_i - \overline{y}
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It will actually help us to rewrite this estimator as

$$\hat{eta}_1 = eta_1 + rac{\sum_i \left(x_i - \overline{x}
ight) u_i}{\sum_i \left(x_i - \overline{x}
ight)^2}$$

Proof: Assuming $y_i = \beta_0 + \beta_1 x_i + u_i$

$$egin{aligned} \hat{eta}_1 &= rac{\sum_i \left(y_i - \overline{y}
ight) \left(x_i - \overline{x}
ight)^2}{\sum_i \left(x_i - \overline{x}
ight)^2} \ &= rac{\sum_i \left(\left[eta_0 + eta_1 x_i + u_i
ight] - \left[eta_0 + eta_1 \overline{x} + \overline{u}
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Phew. **OLS is still unbiased** for the β_k .

Consequences: Efficiency

OLS's efficiency and inference do not survive heteroskedasticity.

• In the presence of heteroskedasticity, OLS is **no longer the most efficient** (best) linear unbiased estimator.

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- In the presence of heteroskedasticity, OLS is no longer the most efficient (best) linear unbiased estimator.
- It would be more informative (efficient) to **weight observations** inversely to their u_i 's variance.
 - \circ Downweight high-variance u_i 's (too noisy to learn much).
 - \circ Upweight observations with low-variance u_i 's (more 'trustworthy').
 - Now you have the idea of weighted least squares (WLS)

Consequences: Inference

OLS **standard errors are biased** in the presence of heteroskedasticity.

- Wrong confidence intervals
- Problems for hypothesis testing (both t and F tests)

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OLS **standard errors are biased** in the presence of heteroskedasticity.

- Wrong confidence intervals
- Problems for hypothesis testing (both t and F tests)
- It's hard to learn much without sound inference.

Heteroskedasticity

Solutions

- 1. **Tests** to determine whether heteroskedasticity is present.
- 2. **Remedies** for (1) efficiency and (2) inference

While we *might* have solutions for heteroskedasticity, the efficiency of our estimators depends upon whether or not heteroskedasticity is present.

- 1. The Goldfeld-Quandt test
- 2. The Breusch-Pagan test
- 3. The White test

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Each of these tests centers on the fact that we can **use the OLS residual** e_i **to estimate the population disturbance** u_i .

The Goldfeld-Quandt test

Focuses on a specific type of heteroskedasticity: whether the variance of u_i differs **between two groups**. †

Remember how we used our residuals to estimate the σ^2 ?

$$s^2 = rac{ ext{SSE}}{n-1} = rac{\sum_i e_i^2}{n-1}$$

We will use this same idea to determine whether there is evidence that our two groups differ in the variances of their disturbances, effectively comparing s_1^2 and s_2^2 from our two groups.

The Goldfeld-Quandt test

Operationally,

```
1. Order your the observations by x
2. Split the data into two groups of size n*
     \circ G<sub>1</sub>: The first third
     \circ G<sub>2</sub>: The last third
3. Run separate regressions of y on x for G_1 and G_2
4. Record SSE_1 and SSE_2
5. Calculate the G-O test statistic
```

The Goldfeld-Quandt test

The G-Q test statistic

$$F_{(n^\star-k,\,n^\star-k)} = rac{\mathrm{SSE}_2/(n^\star-k)}{\mathrm{SSE}_1/(n^\star-k)} = rac{\mathrm{SSE}_2}{\mathrm{SSE}_1}$$

follows an F distribution (under the null hypothesis) with $n^\star - k$ and $n^\star - k$ degrees of freedom.

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Notes

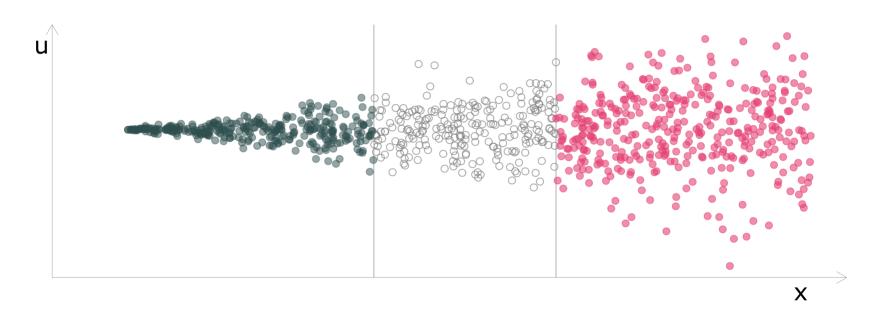
- The G-Q test requires the disturbances follow normal distributions.
- The G-Q assumes a very specific type/form of heteroskedasticity.
- Performs very well if we know the form of potentially heteroskedasticity.

[†]: Goldfeld and Quandt suggested n^* of (3/8)n. k gives number of estimated parameters (i.e., $\hat{\beta}_i$'s).

The Goldfeld-Quandt test



The Goldfeld-Quandt test



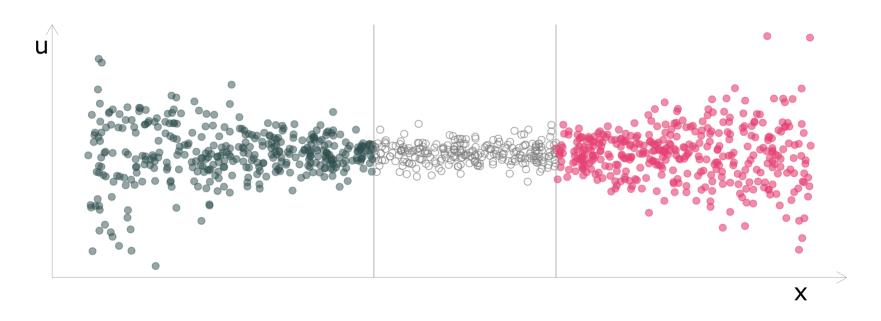
$$F_{375,\,375}=rac{ ext{SSE}_2=18,203.4}{ ext{SSE}_1=1,039.5}pprox 17.5 \implies ext{p-value} < 0.001$$

... We reject H_0 : $\sigma_1^2 = \sigma_2^2$ and conclude there is statistically significant evidence of heteroskedasticity.

The Goldfeld-Quandt test

The problem...

The Goldfeld-Quandt test



$$F_{375,\,375}=rac{ ext{SSE}_2=14,516.8}{ ext{SSE}_1=14,937.1}pprox 1\implies ext{p-value}pprox 0.609$$

 \therefore We fail to reject H_0 : $\sigma_1^2 = \sigma_2^2$ while heteroskedasticity is present.

The Breusch-Pagan test

Breusch and Pagan (1981) attempted to solve this issue of being too specific with the functional form of the heteroskedasticity.

- Allows the data to show if/how the variance of u_i correlates with X.
- If σ_i^2 correlates with X, then we have heteroskedasticity.
- Regresses e_i^2 on $X=[1,\,x_1,\,x_2,\,\ldots,\,x_k]$ and tests for joint significance.

The Breusch-Pagan test

How to implement:

- 1. Regress y on an intercept, x_1 , x_2 , ..., x_k .
- 2. Record residuals e.
- 3. Regress e^2 on an intercept, x_1 , x_2 , ..., x_k .

$$e_i^2=lpha_0+lpha_1x_{1i}+lpha_2x_{2i}+\cdots+lpha_kx_{ki}+v_i$$

- 4. Record R^2 .
- 5. Test hypothesis $exttt{H}_0: \ lpha_1=lpha_2=\dots=lpha_k=0$

The Breusch-Pagan test

The B-P test statistic[†] is

$$\mathrm{LM} = n imes R_e^2$$

where R_e^2 is the R^2 from the regression

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Under the null, LM is asymptotically distributed as χ_k^2 .

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This test statistic tests H_0 : $\alpha_1 = \alpha_2 = \cdots = \alpha_k = 0$.

Rejecting the null hypothesis implies evidence of heteroskedasticity.

[†]: This specific form of the test statistic actually comes form Koenker (1981).

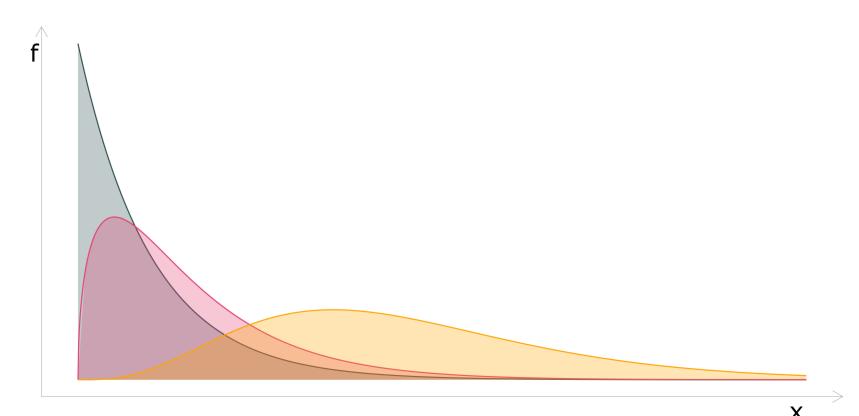
The χ^2 distribution

We just mentioned that under the null, the B-P test statistic is distributed as a χ^2 random variable with k degrees of freedom.

The χ^2 distribution is just another example of a common (named) distribution (like the Normal distribution, the t distribution, and the F).

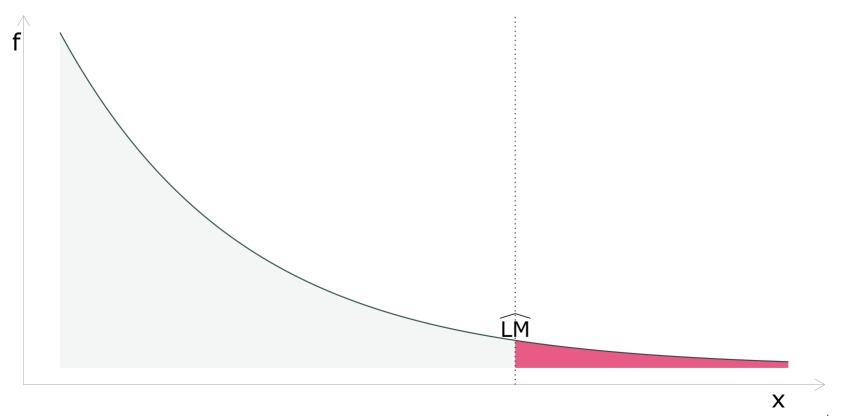
The χ^2 distribution

Three examples of χ_k^2 : k=1, k=2, and k=9



The χ^2 distribution

Probability of observing a more extreme test statistic \widehat{LM} under H₀



The Breusch-Pagan test

Problem: We're still assuming a fairly restrictive **functional form** between our explanatory variables X and the variances of our disturbances σ_i^2 .

The Breusch-Pagan test

Problem: We're still assuming a fairly restrictive **functional form** between our explanatory variables X and the variances of our disturbances σ_i^2 .

Result: B-P may still miss fairly simple forms of heteroskedasticity.

The Breusch-Pagan test

Breusch-Pagan tests are still **sensitive to functional form**.



$$egin{aligned} e_i^2 &= \hat{lpha}_0 + \hat{lpha}_1 x_{1i} \ e_i^2 &= \hat{lpha}_0 + \hat{lpha}_1 x_{1i} + \hat{lpha}_2 x_{1i}^2 \end{aligned}$$

$$\widehat{ ext{LM}} = 1.26$$

$$\widehat{ ext{LM}} = 185.8$$

$$p ext{-value}pprox 0.261$$

$$p$$
-value < 0.001

The White test

So far we've been testing for specific relationships between our explanatory variables and the variances of the disturbances, e.g.,

- H_0 : $\sigma_1^2 = \sigma_2^2$ for two groups based upon x_j (**G-Q**)
- extstyle e

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- H_0 : $\sigma_1^2 = \sigma_2^2$ for two groups based upon x_j (**G-Q**)
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However, we actually want to know if

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Q: Can't we just test this hypothesis?

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However, we actually want to know if

$$\sigma_1^2=\sigma_2^2=\cdots=\sigma_n^2$$

Q: Can't we just test this hypothesis? **A:** Sort of.

The White test

Toward this goal, Hal White took advantage of the fact that we can **replace the homoskedasticity requirement with a weaker assumption**:

- Old: $\operatorname{Var}(u_i|X) = \sigma^2$
- **New:** u^2 is uncorrelated with the explanatory variables (i.e., x_j for all j), their squares (i.e., x_j^2), and the first-degree interactions (i.e., x_jx_h).

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This new assumption is easier to explicitly test (hint: regression).

The White test

An outline of White's test for heteroskedasticity:

- 1. Regress y on x_1 , x_2 , ..., x_k . Save residuals e.
- 2. Regress squared residuals on all explanatory variables, their squares, and interactions.

$$e^2 = lpha_0 + \sum_{h=1}^k lpha_h x_h + \sum_{j=1}^k lpha_{k+j} x_j^2 + \sum_{\ell=1}^{k-1} \sum_{m=\ell+1}^k lpha_{\ell,m} x_\ell x_m + v_i$$

- 3. Record R_e^2 .
- 4. Calculate test statistic to test $exttt{H}_0\colon \ lpha_p=0$ for all p
 eq 0.

The White test

Just as with the Breusch-Pagan test, White's test statistic is

$${
m LM} = n imes R_e^2 \qquad {
m Under} \ {
m H_0}, \ {
m LM} \stackrel{
m d}{\sim} \chi_k^2$$

but now the R_e^2 comes from the regression of e^2 on the explanatory variables, their squares, and their interactions.

$$e^2 = lpha_0 + \sum_{h=1}^k lpha_h x_h \ + \sum_{j=1}^k lpha_{k+j} x_j^2 + \sum_{\ell=1}^{k-1} \sum_{m=\ell+1}^k lpha_{\ell,m} x_\ell x_m + v_i$$
Expl. variables Squared terms

Note: The k (for our χ_k^2) equals the number of estimated parameters in the regression above (the α_i), excluding the intercept (α_0).

The White test

Practical note: If a variable is equal to its square (*e.g.*, binary variables), then you don't (can't) include it. The same rule applies for interactions.

The White test

Example: Consider the model $y = eta_0 + eta_1 x_1 + eta_2 x_2 + eta_3 x_3 + u$

Step 1: Estimate the model; obtain residuals (e).

Step 2: Regress e^2 on explanatory variables, squares, and interactions.

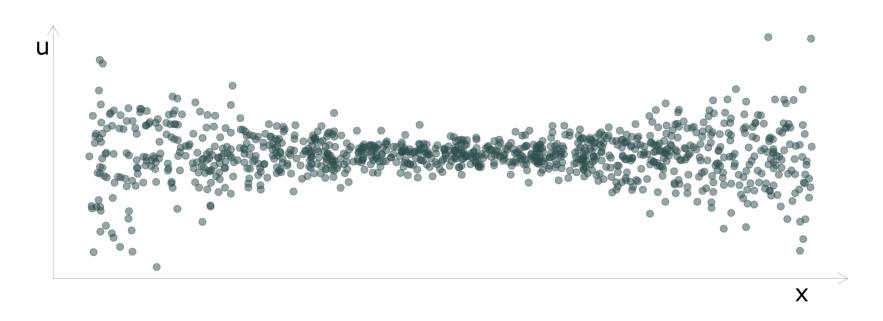
$$e^2 = lpha_0 + lpha_1 x_1 + lpha_2 x_2 + lpha_3 x_3 + lpha_4 x_1^2 + lpha_5 x_2^2 + lpha_6 x_3^2 \ + lpha_7 x_1 x_2 + lpha_8 x_1 x_3 + lpha_9 x_2 x_3 + v$$

Record the R^2 from this equation (call it R_e^2).

Step 3: Test
$$H_0$$
: $lpha_1=lpha_2=\dots=lpha_9=0$ using $\mathrm{LM}=nR_e^2\overset{\mathrm{d}}{\sim}\chi_9^2$.

[\dagger]: To simplify notation here, I'm dropping the i subscripts.

The White test



We've already done the White test for this simple linear regression.

$$e_i^2=\hat{lpha}_0+\hat{lpha}_1x_{1i}+\hat{lpha}_2x_{1i}^2\qquad \widehat{ ext{LM}}=185.8 \qquad ext{p-value} < 0.001$$

Examples

Examples

Goal: Estimate the relationship between standardized test scores (outcome variable) and (1) student-teacher ratio and (2) income, *i.e.*,

$$(\text{Test score})_i = \beta_0 + \beta_1 \text{Ratio}_i + \beta_2 \text{Income}_i + u_i$$
 (1)

Potential issue: Heteroskedasticity... and we do not observe u_i .

Solution:

- 1. Estimate the relationship in (1) using OLS
- 2. Use the residuals (e_i) to test for heteroskedasticity
 - Goldfeld-Quandt
 - Breusch-Pagan
 - White

Examples

We will use testing data from the dataset Caschool in the Ecdat R package.

```
# Load packages
library(pacman)
p_load(tidyverse, Ecdat)
# Select and rename desired variables; assign to new dataset
test_df ← select(Caschool, test_score = testscr, ratio = str, income = avginc)
# Format as tibble
test_df ← as_tibble(test_df)
# View first 2 rows of the dataset
head(test_df, 2)
```

Examples

Let's begin by estimating our model

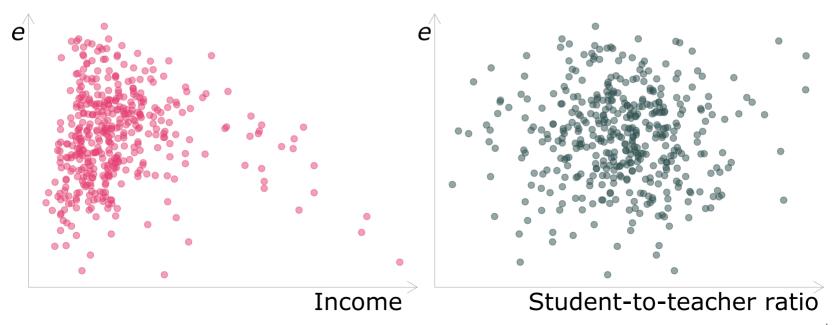
$$(\text{Test score})_i = \beta_0 + \beta_1 \text{Ratio}_i + \beta_2 \text{Income}_i + u_i$$

```
# Estimate the model
est_model ← lm(test_score ~ ratio + income, data = test_df)
# Summary of the estimate
tidy(est_model)
```

Examples

Now, let's see what the residuals suggest about heteroskedasticity

```
# Add the residuals to our dataset
test_df$e ← residuals(est_model)
```



Example: Goldfeld-Quandt

Income looks potentially heteroskedastic; let's test via Goldfeld-Quandt.

```
# Arrange the data by income
test_df ← arrange(test_df, income)
```

Example: Goldfeld-Quandt

Income looks potentially heteroskedastic; let's test via Goldfeld-Quandt.

```
# Arrange the data by income
test_df ← arrange(test_df, income)
# Re-estimate the model for the last and first 158 observations
est_model1 ← lm(test_score ~ ratio + income, data = tail(test_df, 158))
est_model2 ← lm(test_score ~ ratio + income, data = head(test_df, 158))
```

Example: Goldfeld-Quandt

Income looks potentially heteroskedastic; let's test via Goldfeld-Quandt.

```
# Arrange the data by income

test_df ← arrange(test_df, income)

# Re-estimate the model for the last and first 158 observations

est_model1 ← lm(test_score ~ ratio + income, data = tail(test_df, 158))

est_model2 ← lm(test_score ~ ratio + income, data = head(test_df, 158))

# Grab the residuals from each regression

e_model1 ← residuals(est_model1)

e_model2 ← residuals(est_model2)
```

Example: Goldfeld-Quandt

Income looks potentially heteroskedastic; let's test via Goldfeld-Quandt.

```
# Arrange the data by income
test_df ← arrange(test_df, income)
# Re-estimate the model for the last and first 158 observations
est_model1 ← lm(test_score ~ ratio + income, data = tail(test_df, 158))
est_model2 ← lm(test_score ~ ratio + income, data = head(test_df, 158))
# Grab the residuals from each regression
e_model1 ← residuals(est_model1)
e_model2 ← residuals(est_model2)
# Calculate SSE for each regression
(sse_model1 ← sum(e_model1^2))
```

#> [1] 19305.01

```
(sse_model2 ← sum(e_model2^2))
```

#> [1] 29537.83

Example: Goldfeld-Quandt

$$F_{n^\star-k,\,n^\star-k} = rac{ ext{SSE}_2}{ ext{SSE}_1}$$

Example: Goldfeld-Quandt

$$F_{n^{\star}-k,\,n^{\star}-k} = rac{ ext{SSE}_2}{ ext{SSE}_1} {pprox} \, rac{29,537.83}{19,305.01}$$

Example: Goldfeld-Quandt

$$F_{n^{\star}-k,\,n^{\star}-k} = rac{ ext{SSE}_2}{ ext{SSE}_1} \!\!pprox rac{29,537.83}{19,305.01} \!\!pprox 1.53$$

Example: Goldfeld-Quandt

$$F_{n^\star-k,\,n^\star-k} = rac{ ext{SSE}_2}{ ext{SSE}_1} \! pprox rac{29,537.83}{19,305.01} \! pprox 1.53$$
 Test via $F_{158-3,\,158-3}$

Example: Goldfeld-Quandt

$$F_{n^\star-k,\,n^\star-k} = rac{ ext{SSE}_2}{ ext{SSE}_1} \! pprox rac{29,537.83}{19,305.01} \! pprox 1.53$$
 Test via $F_{158-3,\,158-3}$

```
# G-Q test statistic
(f_gq ← sse_model2/sse_model1)
```

```
#> [1] 1.530061
```

Example: Goldfeld-Quandt

Remember the Goldfeld-Quandt test statistic?

$$F_{n^\star-k,\,n^\star-k} = rac{ ext{SSE}_2}{ ext{SSE}_1} \!pprox rac{29,537.83}{19,305.01} \!pprox 1.53$$
 Test via $F_{158-3,\,158-3}$

```
# G-Q test statistic
(f_gq ← sse_model2/sse_model1)
```

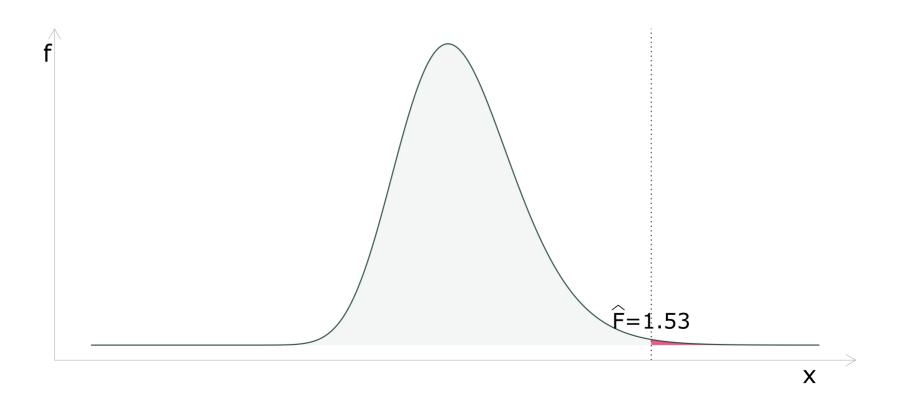
#> [1] 1.530061

```
# p-value
pf(q = f_gq, df1 = 158-3, df2 = 158-3, lower.tail = F)
```

#> [1] 0.004226666

Example: Goldfeld-Quandt

The Goldfeld-Quandt test statistic and its null distribution



Example: Goldfeld-Quandt

Putting it all together:

$$\mathsf{H}_0\!\!:\sigma_1^2=\sigma_2^2$$
 vs. $\mathsf{H}_\mathsf{A}\!\!:\sigma_1^2
eq\sigma_2^2$

Goldfeld-Quandt test statistic: F pprox 1.53

p-value pprox 0.00423

 \therefore Reject H₀ (p-value is less than 0.05).

Conclusion: There is statistically significant evidence that $\sigma_1^2 \neq \sigma_2^2$. Therefore, we find statistically significant evidence of heteroskedasticity (at the 5-percent level).

Example: Goldfeld-Quandt

What if we had chosen to focus on student-to-teacher ratio?

Example: Goldfeld-Quandt

What if we had chosen to focus on student-to-teacher ratio?

```
# Arrange the data by ratio

test_df ← arrange(test_df, ratio)

# Re-estimate the model for the last and first 158 observations

est_model3 ← lm(test_score ~ ratio + income, data = tail(test_df, 158))

est_model4 ← lm(test_score ~ ratio + income, data = head(test_df, 158))

# Grab the residuals from each regression

e_model3 ← residuals(est_model3)

e_model4 ← residuals(est_model4)

# Calculate SSE for each regression

(sse_model3 ← sum(e_model3^2))
```

#> [1] 26243.52

```
(sse_model4 ← sum(e_model4^2))
```

Example: Goldfeld-Quandt

$$F_{n^\star-k,\,n^\star-k} = rac{ ext{SSE}_4}{ ext{SSE}_3} pprox rac{29,101.52}{26,243.52} pprox 1.11$$

which has a *p*-value of approximately 0.2603.

Example: Goldfeld-Quandt

$$F_{n^\star-k,\,n^\star-k} = rac{ ext{SSE}_4}{ ext{SSE}_3} pprox rac{29,101.52}{26,243.52} pprox 1.11$$

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 \therefore We would have failed to reject H₀, concluding that we failed to find statistically significant evidence of heteroskedasticity.

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which has a p-value of approximately 0.2603.

 \therefore We would have failed to reject H₀, concluding that we failed to find statistically significant evidence of heteroskedasticity.

Lesson: Understand the limitations of estimators, tests, etc.

Example: Breusch-Pagan

Let's test the same model with the Breusch Pagan.

Recall: We saved our residuals as e in our dataset, i.e.,

```
test_df$e ← residuals(est_model)
```

Example: Breusch-Pagan

In B-P, we first regress e_i^2 on the explanatory variables,

```
# Regress squared residuals on explanatory variables
bp_model ← lm(I(e^2) ~ ratio + income, data = test_df)
```

Example: Breusch-Pagan

and use the resulting \mathbb{R}^2 to calculate a test statistic.

```
# Regress squared residuals on explanatory variables
bp_model ← lm(I(e^2) ~ ratio + income, data = test_df)
# Grab the R-squared
(bp_r2 ← summary(bp_model)$r.squared)
```

```
#> [1] 3.23205e-05
```

Example: Breusch-Pagan

The Breusch-Pagan test statistic is

$${
m LM}=n imes R_e^2$$

Example: Breusch-Pagan

The Breusch-Pagan test statistic is

$$\mathrm{LM} = n imes R_e^2 pprox 420 imes 0.0000323$$

Example: Breusch-Pagan

The Breusch-Pagan test statistic is

$$\mathrm{LM} = n imes R_e^2 pprox 420 imes 0.0000323 pprox 0.0136$$

which we test against a χ^2_k distribution (here: k=2).

Example: Breusch-Pagan

The Breusch-Pagan test statistic is

$$\mathrm{LM} = n imes R_e^2 pprox 420 imes 0.0000323 pprox 0.0136$$

which we test against a χ^2_k distribution (here: k=2).

```
# B-P test statistic
bp_stat ← 420 * bp_r2
# Calculate the p-value
pchisq(q = bp_stat, df = 2, lower.tail = F)
```

```
#> [1] 0.9932357
```

[\dagger]: k is the number of explanatory variables (excluding the intercept).

Example: Breusch-Pagan

$$\mathsf{H}_0$$
: $lpha_1=lpha_2=0$ vs. H_A : $lpha_1
eq 0$ and/or $lpha_2
eq 0$

for the model
$$u_i^2 = lpha_0 + lpha_1 \mathrm{Ratio}_i + lpha_2 \mathrm{Income}_i + w_i$$

Example: Breusch-Pagan

$$\mathsf{H}_0$$
: $lpha_1=lpha_2=0$ vs. H_A : $lpha_1
eq 0$ and/or $lpha_2
eq 0$

for the model $u_i^2 = lpha_0 + lpha_1 \mathrm{Ratio}_i + lpha_2 \mathrm{Income}_i + w_i$

Breusch-Pagan test statistic: $\widehat{LM} pprox 0.014$

Example: Breusch-Pagan

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: $lpha_1=lpha_2=0$ vs. H_A : $lpha_1
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Breusch-Pagan test statistic: $\widehat{LM} pprox 0.014$

p-value pprox 0.993

Example: Breusch-Pagan

$$\mathsf{H}_0$$
: $lpha_1=lpha_2=0$ vs. H_A : $lpha_1
eq 0$ and/or $lpha_2
eq 0$

for the model $u_i^2 = lpha_0 + lpha_1 \mathrm{Ratio}_i + lpha_2 \mathrm{Income}_i + w_i$

Breusch-Pagan test statistic: $\widehat{LM}\approx 0.014$

p-value pprox 0.993

 \therefore Fail to reject H₀ (the p-value is greater than 0.05)

Example: Breusch-Pagan

$$\mathsf{H}_0$$
: $lpha_1=lpha_2=0$ vs. H_A : $lpha_1
eq 0$ and/or $lpha_2
eq 0$

for the model
$$u_i^2 = lpha_0 + lpha_1 \mathrm{Ratio}_i + lpha_2 \mathrm{Income}_i + w_i$$

Breusch-Pagan test statistic: $\widehat{LM}\approx 0.014$

p-value pprox 0.993

 \therefore Fail to reject H₀ (the p-value is greater than 0.05)

Conclusion: We do not find statistically significant evidence of heteroskedasticity at the 5-percent level.

Example: Breusch-Pagan

$$\mathsf{H}_0$$
: $lpha_1=lpha_2=0$ vs. H_A : $lpha_1
eq 0$ and/or $lpha_2
eq 0$

for the model
$$u_i^2 = lpha_0 + lpha_1 \mathrm{Ratio}_i + lpha_2 \mathrm{Income}_i + w_i$$

Breusch-Pagan test statistic: $\widehat{LM}\approx 0.014$

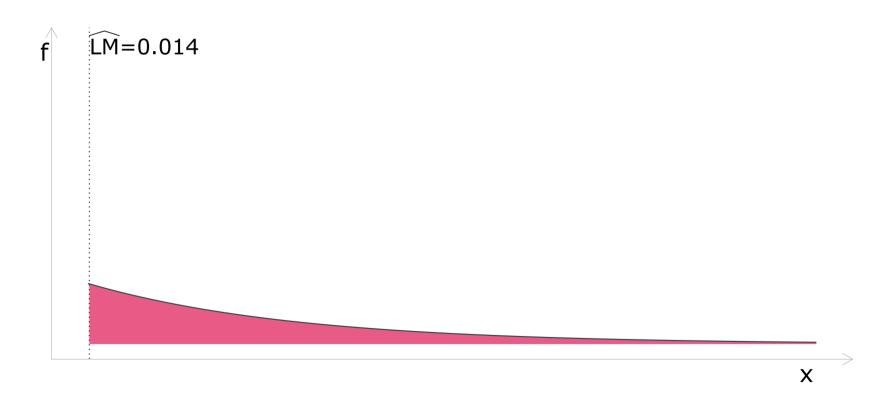
p-value pprox 0.993

 \therefore Fail to reject H₀ (the p-value is greater than 0.05)

Conclusion: We do not find statistically significant evidence of heteroskedasticity at the 5-percent level. (We find no evidence of a *linear* relationship between u_i^2 and the explanatory variables.)

Example: Breusch-Pagan

The Breusch-Pagan test statistic and its null distribution



Example: White

The White test adds squared terms and interactions to the **B-P test**.

$$egin{aligned} u_i^2 = & lpha_0 + lpha_1 \mathrm{Ratio}_i + lpha_2 \mathrm{Income}_i \ & + lpha_3 \mathrm{Ratio}_i^2 + lpha_4 \mathrm{Income}_i^2 + lpha_5 \mathrm{Ratio}_i imes \mathrm{Income}_i \ & + w_i \end{aligned}$$

which moves the null hypothesis from

$$\mathsf{H}_0$$
: $lpha_1=lpha_2=0$ to H_0 : $lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$

Example: White

The White test adds squared terms and interactions to the **B-P test**.

$$egin{aligned} u_i^2 = & lpha_0 + lpha_1 \mathrm{Ratio}_i + lpha_2 \mathrm{Income}_i \ & + lpha_3 \mathrm{Ratio}_i^2 + lpha_4 \mathrm{Income}_i^2 + lpha_5 \mathrm{Ratio}_i imes \mathrm{Income}_i \ & + w_i \end{aligned}$$

which moves the null hypothesis from

$$extsf{H}_0$$
: $lpha_1=lpha_2=0$ to $extsf{H}_0$: $lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$

So we just need to update our R code, and we're set.

Example: White

Aside: R has funky notation for squared terms and interactions in lm():

- Squared terms use I(), e.g., $lm(y \sim I(x^2))$
- **Interactions** use : between the variables, *e.g.*, lm(y ~ x1:x2)

Example: Regress y on quadratic of x1 and x2:

```
# Pretend quadratic regression w/ interactions lm(y \sim x1 + x2 + I(x1^2) + I(x2^2) + x1:x2, data = pretend_df)
```

Example: White

Step 1: Regress e_i^2 on 1st degree, 2nd degree, and interactions

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
   I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
   data = test_df
)
# Grab the R-squared
(white_r2 ← summary(white_model)$r.squared)
```

Example: White

Step 2: Collect R_e^2 from the regression.

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
   I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
   data = test_df
)
# Grab the R-squared
(white_r2 ← summary(white_model)$r.squared)
```

```
#> [1] 0.07332222
```

Example: White

Step 3: Calculate White test statistic ${
m LM}=n imes R_e^2pprox 420 imes 0.073$

```
# Regress squared residuals on quadratic of explanatory variables
white_model ← lm(
   I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
   data = test_df
)
# Grab the R-squared
white_r2 ← summary(white_model)$r.squared
# Calculate the White test statistic
(white_stat ← 420 * white_r2)
```

```
#> [1] 30.79533
```

Example: White

Step 4: Calculate the associated p-value (where LM $\stackrel{d}{\sim} \chi_k^2$); here, k=5

```
# Regress squared residuals on quadratic of explanatory variables
white_model \( - \text{lm(} \)
    I(e^2) ~ ratio + income + I(ratio^2) + I(income^2) + ratio:income,
    data = test_df
)
# Grab the R-squared
white_r2 \( - \text{summary(white_model)} \)*r.squared
# Calculate the White test statistic
white_stat \( - \text{420} \text{* white_r2} \)
# Calculate the p-value
pchisq(q = white_stat, df = 5, lower.tail = F)
```

```
#> [1] 1.028039e-05
```

Example: White

Example: White

$$\mathsf{H}_0$$
: $\alpha_1=\alpha_2=\alpha_3=\alpha_4=\alpha_5=0$

Example: White

$$\mathsf{H}_0$$
: $lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$ vs. H_A : $lpha_i
eq 0$ for some $i\in\{1,\,2,\,\ldots,\,5\}$

Example: White

H
$$_0$$
: $lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$ vs. H $_{ ext{A}}$: $lpha_i
eq 0$ for some $i\in\{1,\,2,\,\ldots,\,5\}$ $u_i^2=lpha_0+lpha_1 ext{Ratio}_i+lpha_2 ext{Income}_i \ +lpha_3 ext{Ratio}_i^2+lpha_4 ext{Income}_i^2 \ +lpha_5 ext{Ratio}_i imes ext{Income}_i+w_i$

Example: White

Putting everything together...

H₀:
$$\alpha_1=\alpha_2=\alpha_3=\alpha_4=\alpha_5=0$$
 vs. H_A: $\alpha_i\neq 0$ for some $i\in\{1,\,2,\,\ldots,\,5\}$
$$u_i^2=\alpha_0+\alpha_1\mathrm{Ratio}_i+\alpha_2\mathrm{Income}_i$$

$$+\alpha_3\mathrm{Ratio}_i^2+\alpha_4\mathrm{Income}_i^2$$

$$+\alpha_5\mathrm{Ratio}_i\times\mathrm{Income}_i+w_i$$

Our White test statistic: LM $= n imes R_e^2 pprox 420 imes 0.073 pprox 30.8$

Example: White

Putting everything together...

H₀:
$$lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$$
 vs. H_A: $lpha_i
eq 0$ for some $i\in\{1,\,2,\,\ldots,\,5\}$
$$u_i^2=lpha_0+lpha_1\mathrm{Ratio}_i+lpha_2\mathrm{Income}_i \\ +lpha_3\mathrm{Ratio}_i^2+lpha_4\mathrm{Income}_i^2 \\ +lpha_5\mathrm{Ratio}_i imes\mathrm{Income}_i+w_i$$

Our White test statistic: L ${
m M}=n imes R_e^2pprox 420 imes 0.073pprox 30.8$

Under the χ^2_5 distribution, this $\widehat{\mathrm{LM}}$ has a p-value less than 0.001.

Example: White

Putting everything together...

H₀:
$$lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$$
 vs. H_A: $lpha_i
eq 0$ for some $i\in\{1,\,2,\,\ldots,\,5\}$
$$u_i^2=lpha_0+lpha_1\mathrm{Ratio}_i+lpha_2\mathrm{Income}_i \\ +lpha_3\mathrm{Ratio}_i^2+lpha_4\mathrm{Income}_i^2 \\ +lpha_5\mathrm{Ratio}_i imes\mathrm{Income}_i+w_i$$

Our White test statistic: LM $= n imes R_e^2 pprox 420 imes 0.073 pprox 30.8$

Under the χ^2_5 distribution, this $\widehat{\mathrm{LM}}$ has a p-value less than 0.001.

∴ We reject H₀

Example: White

Putting everything together...

H₀:
$$lpha_1=lpha_2=lpha_3=lpha_4=lpha_5=0$$
 vs. H_A: $lpha_i
eq 0$ for some $i\in\{1,\,2,\,\ldots,\,5\}$
$$u_i^2=lpha_0+lpha_1\mathrm{Ratio}_i+lpha_2\mathrm{Income}_i \\ +lpha_3\mathrm{Ratio}_i^2+lpha_4\mathrm{Income}_i^2 \\ +lpha_5\mathrm{Ratio}_i imes\mathrm{Income}_i+w_i$$

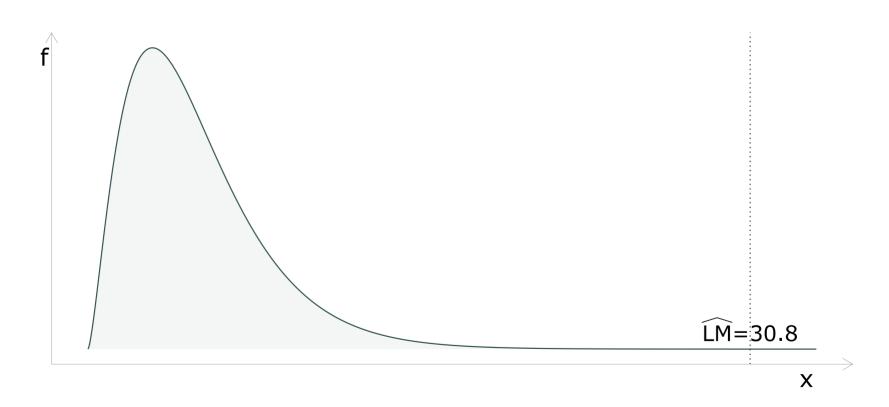
Our White test statistic: LM $= n imes R_e^2 pprox 420 imes 0.073 pprox 30.8$

Under the χ^2_5 distribution, this $\widehat{\mathrm{LM}}$ has a p-value less than 0.001.

... We **reject H₀** and conclude there is **statistically significant evidence of heteroskedasticity** (at the 5-percent level).

Example: White

The White test statistic and its null distribution



- **Q:** What is the definition of heteroskedasticity?
- **Q:** Why are we concerned about heteroskedasticity?
- **Q:** Does plotting y against x, tell us anything about heteroskedasticity?
- **Q:** Does plotting e against x, tell us anything about heteroskedasticity?
- **Q:** Since we cannot observe the u_i 's, what do we use to *learn about* heteroskedasticity?
- **Q:** Which test do you recommend to test for heteroskedasticity? Why?

Review questions

• **Q:** What is the definition of heteroskedasticity?

Review questions

- **Q:** What is the definition of heteroskedasticity?
- A:

Math: $\operatorname{Var}(u_i|X) \neq \operatorname{Var}(u_i|X)$ for some $i \neq j$.

Words: There is a systematic relationship between the variance of u_i and our explanatory variables.

- **Q:** What is the definition of heteroskedasticity?
- **Q:** Why are we concerned about heteroskedasticity?

- **Q:** What is the definition of heteroskedasticity?
- **Q:** Why are we concerned about heteroskedasticity?
- **A:** It biases our standard errors—wrecking our statistical tests and confidence intervals. Also: OLS is no longer the most efficient (best) linear unbiased estimator.

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- **Q:** What is the definition of heteroskedasticity?
- Q: Why are we concerned about heteroskedasticity?
- **Q:** Does plotting y against x, tell us anything about heteroskedasticity?
- **A:** It's not exactly what we want, but since y is a function of x and u, it can still be informative. If y becomes more/less disperse as x changes, we likely have heteroskedasticity.

- **Q:** What is the definition of heteroskedasticity?
- Q: Why are we concerned about heteroskedasticity?
- Q: Does plotting y against x, tell us anything about heteroskedasticity?
- **Q:** Does plotting e against x, tell us anything about heteroskedasticity?

- **Q:** What is the definition of heteroskedasticity?
- Q: Why are we concerned about heteroskedasticity?
- Q: Does plotting y against x, tell us anything about heteroskedasticity?
- **Q:** Does plotting e against x, tell us anything about heteroskedasticity?
- **A:** Yes. The spread of *e* depicts its variance—and tells us something about the variance of *u*. Trends in this variance, along *x*, suggest heteroskedasticity.

- **Q:** What is the definition of heteroskedasticity?
- Q: Why are we concerned about heteroskedasticity?
- Q: Does plotting y against x, tell us anything about heteroskedasticity?
- Q: Does plotting e against x, tell us anything about heteroskedasticity?
- **Q:** Since we cannot observe the u_i 's, what do we use to *learn about* heteroskedasticity?

- **Q:** What is the definition of heteroskedasticity?
- Q: Why are we concerned about heteroskedasticity?
- **Q:** Does plotting y against x, tell us anything about heteroskedasticity?
- **Q:** Does plotting *e* against *x*, tell us anything about heteroskedasticity?
- **Q:** Since we cannot observe the u_i 's, what do we use to *learn about* heteroskedasticity?
- **A:** We use the e_i 's to predict/learn about the u_i 's. This trick is key for almost everything we do with heteroskedasticity testing/correction.

- **Q:** What is the definition of heteroskedasticity?
- Q: Why are we concerned about heteroskedasticity?
- **Q:** Does plotting y against x, tell us anything about heteroskedasticity?
- Q: Does plotting e against x, tell us anything about heteroskedasticity?
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- **Q:** Since we cannot observe the u_i 's, what do we use to *learn about* heteroskedasticity?
- **Q:** Which test do you recommend to test for heteroskedasticity? Why?
- **A:** I like White. Fewer assumptions. Fewer issues.