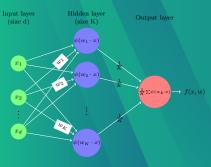
Log Concave Coupling for Sampling Neural Net Posteriors

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Problem Definition



- Wide single hidden layer neural network
- Outer weights fixed, linear combination of neurons from infinite dictionary
- Train inner weights *w* not by optimization:

$$w = \operatorname{argmax}_{w} L(\operatorname{Data}, w)$$

but via posterior sampling:

$$w \sim p(w|\mathsf{Data}) \propto e^{L(\mathsf{Data},w)}$$

Overall fit is posterior mean,

$$\hat{f}(x) = E[f(x, w)|Data]$$

Target Density

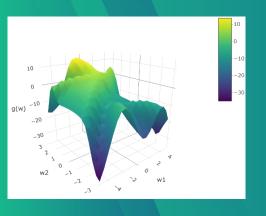
- Posterior density priorities:
 - i) Predictive risk control, $E[||\hat{f} f||^2]$
 - ii) Computational feasibility (i.e. rapid MCMC sampling)
- Train network in a Greedy fashion, add in one neuron at a time
- Previous fit $f_k(x)$ using k neurons, add in new neuron w_{k+1}
- Prior density $p_0(w)$ uniform over ℓ_1 ball, $||w||_1 \le 1$
- For some $\alpha \in [0, 1]$, residuals $r_i = y_i (1 \alpha)f_k(x_i)$
- For some gain $\beta > 0$,

$$p_{k+1}(w) \propto e^{\beta \sum_{i=1}^{n} r_i \psi(x_i \cdot w)} p_0(w)$$

$$f_{k+1}(x) = (1 - \alpha) f_k(x) + \beta E_{D_{k+1}} [\psi(x \cdot w)]$$

- Assume:
 - 1) Bounded data entries $|x_{i,j}| \le 1$ for all i,j
 - 2) Activation function ψ has bounded second derivative, $|\psi''(u)| \le c$, example tanh or squared ReLU.

Difficulties Sampling and Multi-Modal Posterior



Posterior log likelihood:

$$g(w) = \sum_{i=1}^{n} r_i \, \psi(x_i \cdot w)$$
$$p(w) \propto e^{\beta g(w)}$$

- (Jones, Ba. 93) greedy pursuit, maximize g(w) to find best next neuron
- However, g(w) is multi-modal and not concave
 - $abla^2 g(w) = \sum_{i=1}^n r_i \, \psi''(x_i \cdot w) x_i x_i^T$ is sum of rank 1 matrices with \pm scaling, not negative definite (i.e. not concave)

Auxiliary Random Variable and Joint Density

- Auxiliary r.v.'s $\xi_i \sim \text{Normal}(\sqrt{\beta |r_i|c} x_i \cdot w, 1), i = 1, \dots, n$
- Joint density $p(w,\xi) = p(w)p(\xi|w)$
- Pair $(w, \xi) \sim p(w, \xi)$ has correct w marginal $w \sim p(w)$
- By Bayes rule, also have $p(w,\xi) = p(\xi)p(w|\xi)$
- $p(\xi)$: marginal of auxiliary r.v., convolution with normal
- $p(w|\xi)$: reverse conditional, has addition negative definite terms in log likelihood Hessian (will make log concave)
- Goal: first sample a $\xi \sim p(\xi)$ and then sample a $w \sim p(w|\xi)$. If both densities are log concave, MCMC methods will be rapidly mixing

Log Concavity of Reverse Conditional Density

$$p(w|\xi) \propto p(w)p(\xi|w)$$

$$\propto \exp\left(\beta \sum_{i=1}^{n} r_{i}\psi(x_{i} \cdot w) - \frac{1}{2}(\xi_{i} - \sqrt{\beta|r_{i}|c} x_{i} \cdot w)^{2}\right) p_{0}(w)$$

$$\propto \exp\left(\beta \sum_{i=1}^{n} r_{i}\psi(x_{i} \cdot w) - \beta|r_{i}|c \frac{1}{2}(x_{i} \cdot w)^{2} + \sqrt{\beta|r_{i}|c} \xi_{i}x_{i} \cdot w\right) p_{0}(w)$$

$$\nabla^{2} \log p(w) = \beta \sum_{i=1}^{n} (r_{i}\psi''(x_{i} \cdot w)) x_{i}x_{i}^{T}$$

$$\nabla^{2} \log p(w|\xi) = \beta \sum_{i=1}^{n} (r_{i}\psi''(x_{i} \cdot w) - |r_{i}|c) x_{i}x_{i}^{T}$$

- Negative definite term introduced to Hessian
- By assumption $r_i \overline{\psi''(x_i \cdot w) |r_i| c < 0}$
- Reverse conditional now log-concave
- Log concave density over convex set rapidly sampled by existing MCMC methods (Applegate, Kannan 91, Lovász, Vempala 07)

Marginal Density of Auxiliary RV

• Marginal density $p(\xi)$ is convolution of target p(w) with normal $p(\xi|w)$

$$p(\xi) = \int p(w)p(\xi|w)dw$$

• Denote $g(w) = \beta \sum_{i=1}^{n} r_i \psi(x_i \cdot w) - \beta |r_i| c_{\frac{1}{2}}^1 (x_i \cdot w)^2$ (concave modified target)

$$\log p(\xi) = -\frac{1}{2} \|\xi\|^2 + \log \left(\int p_0(w) e^{g(w)} e^{\sum_{i=1}^n \sqrt{\beta |r_i| c} \, \xi_i x_i w} dw \right)$$

- Negative quadratic (concave, neg def hessian)
- Cumulant generating function (convex, pos def hessian)
- Competition for overall log concavity

Score of Auxiliary RV

Score, define |R| as diagonal matrix of residuals:

$$\nabla \log p(\xi) = -\xi + E[\sqrt{\beta c}|R|^{\frac{1}{2}}Xw|\xi]$$

- MCMC algorithms need access to score (e.g. Langevin Diffusion)
- Score defined by expectation over log concave density $p(w|\xi)$
- Score estimated empirically via MCMC method on reverse conditional $p(w|\xi)$

Log Concavity of Auxiliary RV

Hessian:

$$\nabla^2 \log p(\xi) = -I + \operatorname{Cov}[\sqrt{\beta c} |R|^{\frac{1}{2}} X w |\xi]$$

• Log concave if for all unit vector a, direction $v = \sqrt{\beta c} |R|^{\frac{1}{2}} X^T a$

$$Var(v \cdot w|\xi) \leq 1$$

• Upper bound on maximum eigenvalue of $Cov[w|\xi]$, i.e. max variance in any direction sufficiently small

Covariance Under Prior

- Want Var[v⋅w|ξ]
- Consider first prior density Var_{p₀}[v·w]
- Prior $w \sim \text{Uniform}\{w \ s.t. \|w\|_1 \leq 1\}$
- $\operatorname{Cov}_{p_0}[w] \leq \frac{1}{d^2}I$
- If w drawn from prior, $\mathrm{Cov}_{p_0}[\sqrt{\beta c}|R|^{\frac{1}{2}}Xw] \preceq \frac{c\|r\|_{\infty}\beta n}{d}I$
- Intuition from prior: large dimension d can control covariance, d of order $c||r||_{\infty}\beta n$.

Two Proof Methods

1. Contraction Conjecture (not proven):

Conditional covariance LESS than prior,

$$Cov[w|\xi] \leq Cov_{p_0}[w]$$

- Condition for log concavity: $\frac{c\|r\|_{\infty}\beta n}{d} < 1$
- With $\beta = \frac{1}{\sqrt{n}}$ need dimension

$$(c||r||_{\infty})\sqrt{n} < d$$

2. Hölder Inequality Lemma:

• Any direction $v = \sqrt{\beta c} |R|^{\frac{1}{2}} X^T a$,

$$\operatorname{Var}(v \cdot w|\xi) < 20 \frac{(c\|r\|_{\infty}\beta n)^2}{d}$$

- Condition for log concavity, $20 \frac{(c||r||_{\infty}\beta n)^2}{d} < 1$
- With $\beta = \frac{1}{\sqrt{n}}$ need dimension

$$20(c||r||_{\infty})^2 n < d$$

Contraction Conjecture

- $p(w|\xi) \propto p_0(w)e^{\beta g_\xi(w)}, \, g_\xi(w)$ is concave in w
- Would seem to reduce variance in any direction
- When prior $p_0(w)$ is over all of \mathbb{R}^d , this is true
- One dimensional marginal in any direction $u = v \cdot w$

$$\frac{d^2}{du^2}\log p(u|\xi)<\frac{d^2}{du^2}\log p_0(u)$$

- One dimensional optimal transport map is a contraction
- Less variance in any direction
- Does NOT hold when $p_0(w)$ restricted to convex set (e.g. uniform over ℓ_1 ball)
- Examples where concave function in one direction increases variance in another
- Example,

$$p(w_1, w_2) \propto e^{-\beta w_2^2} \mathbf{1}\{|w_1| + |w_2| \leq 1\}$$

Increased β INCREASES variance in w_1 direction

Hölder Inequality Proof Sketch

• $Var(v \cdot w|\xi)$ is not more than

$$\int (v\cdot w)^2 e^{\beta \tilde{g}_{\xi}(w) - \Gamma_{\xi}(\beta)} p_0(w) dw$$

where $\tilde{g}_{\xi}(w)$ is $g_{\xi}(w)$ minus it's mean under $p_{0}(w)$

- $\Gamma_{\xi}(w)$ is cumulant generating function of $\tilde{g}_{\xi}(w)$
- $\Gamma_{\xi}(w)$ concave, $\Gamma_{\xi}(0)=0, \Gamma_{\xi}'(0)=0$
- By Hölders inequality variance is not more than

$$(\mathcal{E}_{\rho_0}[(v\cdot w)^{2\ell}])^{\frac{1}{\ell}} \exp\left\{\frac{\ell}{\ell-1}\Gamma_{\xi}(\frac{\ell}{\ell-1}\beta) - \Gamma_{\xi}(\beta)\right\}$$

- Moments of prior analysis, $(E_{
 ho_0}[(v\cdot w)^{2\ell}])^{rac{1}{\ell}} < c\|r\|_\infty eta n rac{4\ell}{ed}$
- CGF analysis, $rac{\ell}{\ell-1} \Gamma_\xi(rac{\ell}{\ell-1}eta) \Gamma_\xi(eta) < c \|r\|_\infty eta n_{\ell}^5$
- Optimize over ℓ to get bound $20 \frac{(c\|r\|_{\infty} \beta n)^2}{d}$

Summary

Train single hidden layer NN via sequential Greedy Bayes, residuals $r_i = y_i - (1 - \alpha)f_k(x_i)$

$$p_{k+1}(w) \propto e^{\beta \sum_{i=1}^{n} r_i \psi(x_i \cdot w)}$$

$$f_{k+1}(x) = (1 - \alpha) f_k(x) + \beta E_{p_k} [\psi(x \cdot w)]$$

- Define joint density $p(w,\xi)$ with correct w marginal
- $p(w|\xi)$ log concave and can be sampled via MCMC
- Marginal $p(\xi)$ has score which can be estimated via MCMC sampling over $p(w|\xi)$:

$$\nabla \log p(\xi) = -\xi + E[\sqrt{\beta c}|R|^{\frac{1}{2}}Xw|\xi]$$

- $p(\xi)$ is log concave for large dimension d:
 - (Conjectured) Log-concave when $c||r||_{\infty}\beta n < d$
 - (Proven) Log-concave when $20(c||r||_{\infty}\beta n)^2 \leq d$

Future Work: Predictive Risk Control

- Statistical risk or generalization squared error: $E[\|\hat{f} f\|^2]$
- Use multiple subsets of data $1 \le n \le N$
 - Predictive density: $p_n(y|x) = \int p(y|x, w)p(w|x^n, y^n)dw$
 - Predictive mean: $\hat{f}_n(x) = \int f(x, w) p(w|x^n, y^n) dw$
 - Cumulative estimator: $\hat{f}(x) = \frac{1}{N} \sum_{n=1}^{N} \hat{f}_n(x)$
- Upper bound with Kullback divergence,

$$E[\|\hat{\hat{f}} - f\|^2] \le \frac{1}{N} D(P_{Y^N, X^N}^* \| P_{Y^N, X^N})$$

Index of resolvability (Ba. 87, 98):

$$\begin{split} \frac{1}{N} D(P_{Y^N,X^N}^* \| P_{Y^N,X^N}) &= \frac{1}{N} E[\log \frac{p^*(y^N,x^N)}{\int p(y^N,x^N|w) p_0(w) dw}] \\ &\leq \frac{1}{N} E[\log \frac{p^*(y^N,x^N)}{\int_A p(y^N,x^N|w) p_0(w) dw}] \\ &\leq D_A + \frac{1}{N} \log \frac{1}{P_0(A)} \end{split}$$

- Prior probability control $P_0(A)$, Gaussian not sufficient, need uniform prior
- Best bounds to date: $E[\|\hat{\hat{t}} f\|^2] \leq C \left(\frac{(\log 2d)}{N} \right)^{\frac{1}{3}}$



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