

Bayesian Genetic Mark-Recapture Methods For Estimating Seasonal River Run Size Of Stock Populations

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Motivating Application



Figure: Sockeye Salmon (Photo by Kevin Phillips)



Figure: Taku River (Photo by Paul Vecsei)

Motivating Application

Genetic Mark-Recapture (GMR) Method

Genetic Data:

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Genetic Data:

Estimated Proportions for species
of interest

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Counts for some species

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$$(\sigma_1, \dots, \sigma_K)$$

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Count Data:

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$$N = \frac{N_1}{\mu_1}$$

Motivating Application

Genetic Stock Identification (GSI) Data

n_t , $t = 1, \dots, T$: Sample size of genetic tissues in week t

$(\mu_{k,t})$: In-sample posterior stock proportion estimate in week t for stock k

$(\sigma_{k,t})$: In-sample posterior standard deviation in week t for stock k

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Run Weight Data

w_t : The Sockeye Salmon run weight (relative abundance) in week t , with $\sum_{t=1}^T w_t = 1$

Method of Moments (MM)

In 2001, W. J. Gazey developed a method of moments estimator to estimate total abundance of specific Salmon stocks in the Alsek River system[2]. This method has been used to estimate the abundance of Sockeye Salmon in the Taku River system by Pestal et al.[1]:

$$\hat{N} = \frac{N_1}{\sum_{t=1}^T w_t \mu_{1,t}}, \quad (1)$$

with

$$\widehat{\text{Var}}(\hat{N}) = \sum_{t=1}^T (w_t \sigma_{1,t})^2 \left[\frac{\hat{N}}{\sum_{i=1}^T w_i \mu_{1,i}} \right]^2. \quad (2)$$

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- **The variance may be significantly underestimated.**

Proposed Model

- ① Let $X_{k,t}$ denote the number of fish of stock k in the GSI sample in week t , then:

$$(X_{1,t}, X_{2,t}, \dots, X_{K,t}) \sim \text{Multinom}(n_t; \pi_{1,t}, \pi_{2,t}, \dots, \pi_{K,t}), \quad (3)$$

where $\pi_{k,t}$ is the true proportion of stock k in the population in week t .

- ② The observed GSI proportions $\mu_{k,t}$ are described as

$$\mu_{k,t} \sim \text{TN}\left(\frac{X_{k,t}}{n_t}, \sigma_{k,t}^2\right) \quad (4)$$

where TN represents a truncated normal distribution with a range of $[0, \infty)$. The choice of the normal distribution is motivated by the Bernstein-von Mises theorem.

Proposed Model

We can then derive the total number of Sockeye Salmon as

$$N = \frac{N_1}{\sum_{t=1}^T w_t \pi_{1,t}} \quad (5)$$

and stock-specific abundances as $N_k = N \sum_{t=1}^T w_t \pi_{k,t}$, for $k = 2, \dots, K$.

Choice of Prior

Conjugate Dirichlet Prior (CDP)

$$(\pi_{1,t}, \dots, \pi_{K,t}) \sim \text{Dirichlet}(\theta_1, \dots, \theta_K) \quad (6)$$

When $\theta_k = 1$, it is a uniform prior; when $\theta_k = \frac{1}{2}$, it is a Jeffreys prior[3].

Choice of Prior

Transformed Normal Prior (TNP)[4]

$$\pi_{k,t} = \frac{\exp(\tau_{k,t})}{\sum_{i=1}^K \exp(\tau_{i,t})} \quad (7)$$

$$\tau_{k,t} \stackrel{ind}{\sim} N(0, \xi^2). \quad (8)$$

We consider various values for ξ : 0.5, 1, 2, ..., 5 as well as a uniform hyperprior $\xi \sim U(0, 10)$.

Choice of Prior

Time Series Prior (TSP)

$$\tau_{k,t} = \rho\tau_{k,t-1} + \epsilon_{k,t} \quad (9)$$

$$\tau_{k,1} \sim N\left(0, \frac{s^2}{1 - \rho^2}\right) \quad (10)$$

$$\epsilon_{k,t} \sim N(0, s^2) \quad (11)$$

$$\rho \sim \text{Unif}(-1, 1) \quad (12)$$

For s , we consider the values 0.5, 1, 2, ..., 5 as well as a uniform hyperprior $s \sim U(0, 5)$. We also consider the case where s and ρ vary per stock and week with independent priors.

Simple Simulation

- Settings:

- ▶ 5 weeks, 3 stocks.
- ▶ 500 GSI datasets with arbitrary preset values of $\pi_{k,t}$.
- ▶ The true value of N was set as 80,000 and the weir count data N_1 was calculated from

$$N = \frac{N_1}{\sum_{t=1}^T w_t \pi_{1,t}}$$

as a function of N , $\pi_{k,t}$ and w_t , where $w_t = 0.2$ for $t = 1, \dots, 5$.

Simple Simulation

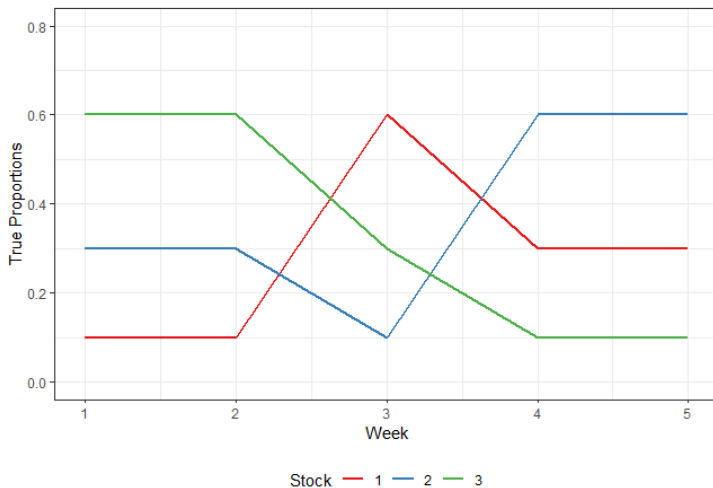


Figure: Parameter values ($\pi_{k,t}$) of the simple simulation.

Simple Simulation

- Results:
 - ▶ Close to the true value: absolute RBs < 0.01, RRMSEs < 0.07.
 - ▶ CP: 34.2% (MM) vs. 94.6% ~ 96% (Proposed)
 - ▶ SD of \hat{N} : Proposed = $4 \times$ MM
- The MM estimator underestimates the amount of uncertainty in the estimate.
- The MM estimator does not account for the fact that $\mu_{k,t}$ is the stock proportion in the GSI samples, instead of in the population.

Data-Based Simulation

- Settings:

- ▶ 12 weeks, 17 stocks (4 lake-types vs. 13 river-types).
- ▶ 500 GSI datasets with the observed $\mu_{k,t}$ in the GSI data as the values of $\pi_{k,t}$.
- ▶ The true value of N was set as 60,000 based on the MM estimator in the PSC report[1].
- ▶ The values of w_t and $\sigma_{k,t}$ were set as the observed values from the Taku River dataset.

Data-Based Simulation

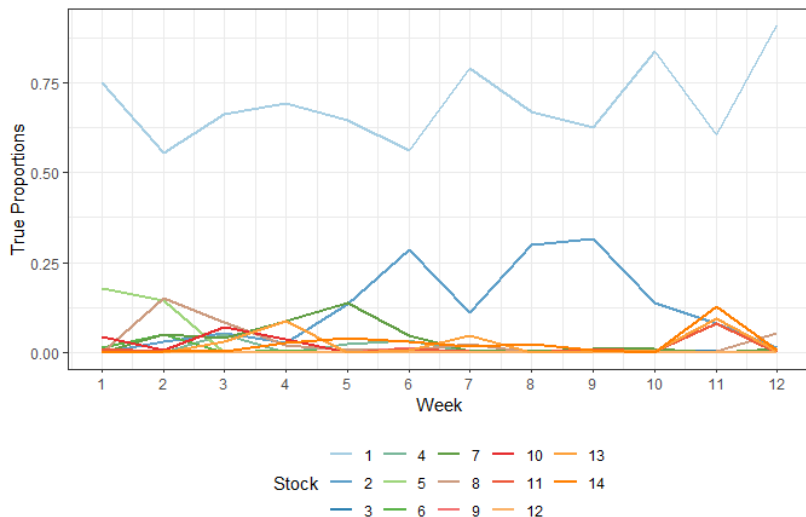


Figure: Parameter values ($\pi_{k,t}$) of the data-based simulation.

Data-Based Simulation

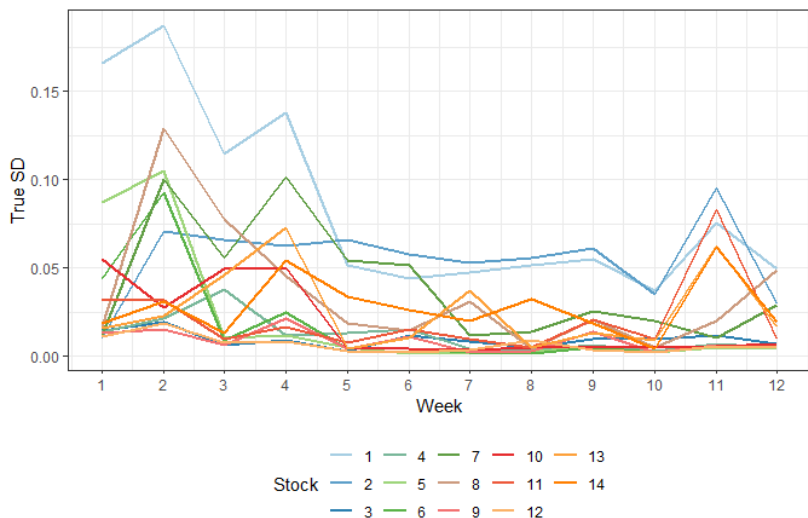


Figure: Standard deviation values ($\sigma_{k,t}$) of the data-based simulation.

Data-Based Simulation

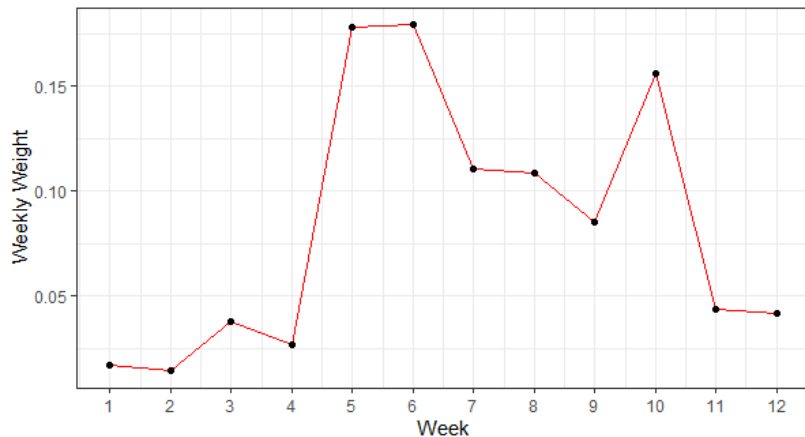


Figure: Run weights (w_t) of the data-based simulation.

Data-Based Simulation

Table 1: Results of the data-based simulation.

Setting	Relative Bias	RRMSE	SD	CP
MM	0.00	0.032	1,570	0.88
DCP $\theta_k = 1$	0.18	0.1868	2,334	0.00
DCP $\theta_k = 0.5$	0.09	0.092	2,042	0.18
TNP $\xi = 0.5$	0.25	0.253	2,621	0.00
TNP $\xi = 1$	0.06	0.064	1,983	0.58
TNP $\xi = 2$	-0.01	0.030	1,760	0.93
TNP $\xi = 3$	-0.03	0.038	1,694	0.85
TNP $\xi = 4$	-0.03	0.045	1,658	0.78
TNP $\xi = 5$	-0.04	0.049	1,634	0.69
TNP $\xi \sim U(0, 10)$	-0.05	0.054	1,633	0.63
TSP $s \sim U(0, 5)$	-0.04	0.052	1,640	0.64

Discussion

- The proposed method is particularly sensitive to the choice of prior, may due to some of the $\pi_{k,t}$'s being very close to zero.
- One possible explanation could be that when simulating the data, we did not constrain $\sum_{k=1}^K \mu_{k,t} = 1$ for $t = 1, \dots, T$.
- The TNP model may still be reasonable to use if various values of ξ are considered and model selection criteria such as the DIC or cross-validation are applied.
- Future Work:
 - ▶ Propose priors that would not be so sensitive.
 - ▶ Use spline method to explore the dynamics through time.
 - ▶ ...

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Thank You