

Lecture 2 – Feb. 2012

General linear model: theory of linear model & advanced applications in statistics

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Overview

- Linear algebra 2: Projection matrices
- ANOVAs using projections
- Multivariate Regressions
- Linear time invariant model (fMRI)
- A word on generalized linear model



Linear Algebra

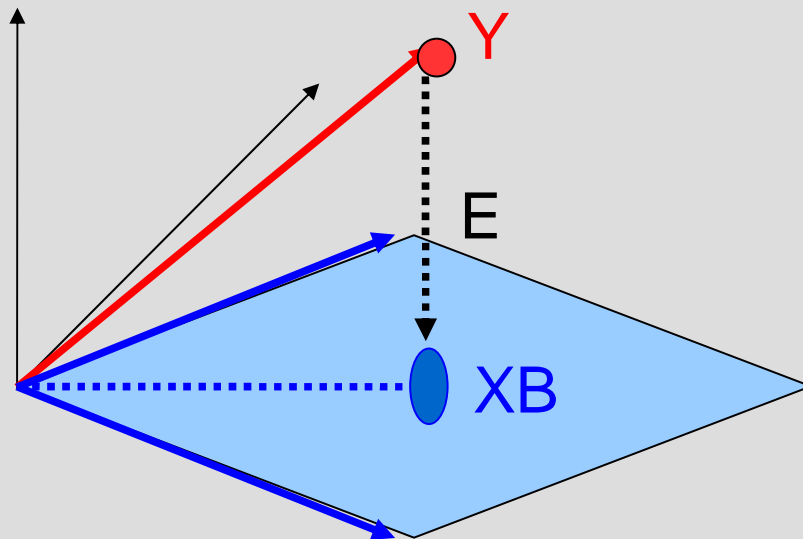
again!

Linear Algebra

- Lecture 1: we write equations as matrices – using matrices we can find all of the unknown (the Bs, or regression coefficients) in 1 single operation $B = \text{pinv}(X)Y$
- What we do with matrices, is to find how many of each vectors in X we need to be as close as possible to Y ($Y = XB + e$)

Linear Algebra and Statistics

- $Y = 3$ observations $X = 2$ regressors
- $Y = XB + E \rightarrow \hat{Y} = XB$



SS total = variance in Y

SS effect = variance in XB

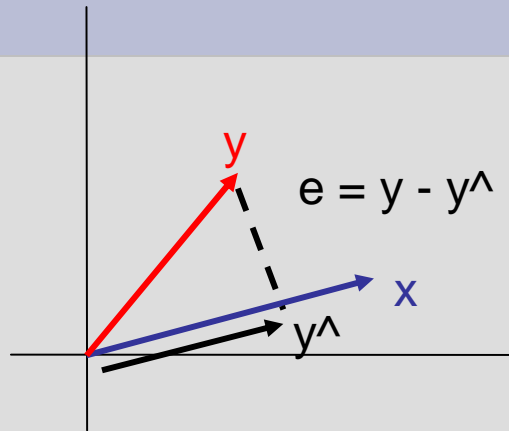
SS error = variance in E

$R^2 = \text{SS effect} / \text{SS total}$

$F = \text{SS effect/df} / \text{SS error/dfe}$

We can find \hat{Y} by computing B
Can we think of another way?

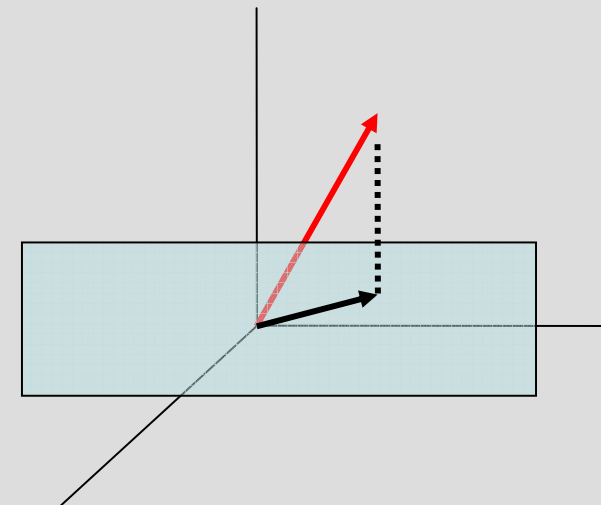
Linear Algebras: Projections



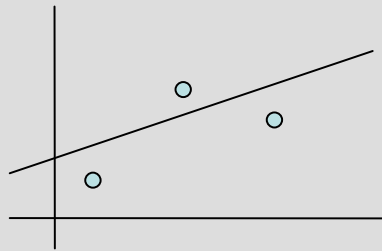
$$\begin{aligned}y^\wedge &= \beta x \\x'(y - \beta x) &= 0 \\ \beta x'x &= x'y \\ \beta &= x'y / x'x\end{aligned}$$

$$\begin{aligned}y^\wedge &= (x'y / x'x)x \\ y^\wedge &= Py \rightarrow P = xx' / x'x\end{aligned}$$

Why project? $XB = Y$ may have no solution, the closest solution is a vector located in X space that is the closest to Y . With a bit of math we can find $P = X \text{inv}(X'X)X'$



Projection and Least square

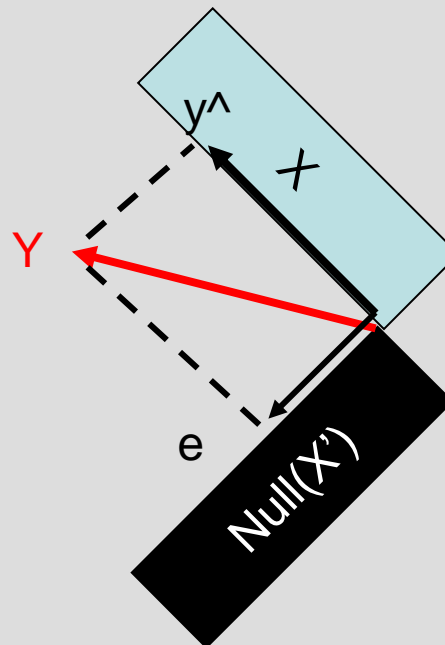


$$y = \beta x + c$$

P projects the points on the line

Minimizing the distance (\wedge^2)

is projecting at perpendicular angles



$$Y = \hat{y} + e$$

$$\hat{y} = PY$$

$$e = (I - P)Y$$



ANOVA revised

1 way ANOVA

- $u1 = \text{rand}(10,1) + 11.5$; $u2 = \text{rand}(10,1) + 7.2$; $u3 = \text{rand}(10,1) + 5$; $Y = [u1; u2; u3]$;
- $x1 = [\text{ones}(10,1); \text{zeros}(20,1)]$; $x2 = [\text{zeros}(10,1); \text{ones}(10,1); \text{zeros}(10,1)]$; $x3 = [\text{zeros}(20,1); \text{ones}(10,1)]$; $X = [x1 \ x2 \ x3 \ \text{ones}(30,1)]$;
- Lecture 1 solution:
- $B = \text{pinv}(X)*Y$ and $\hat{Y} = X*B$
- Now the solution:
- $P = X*\text{pinv}(X)$ and $\hat{Y}_2 = P*Y$

1 way ANOVA

- What to use? Both!
- Projections are great because we can now constrain \hat{Y} to move along any combinations of the columns of X
- Say you now want to contrast gp1 vs gp2? $C = [1 \ -1 \ 0 \ 0]$
- Compute B so we have XB based on the full model X then using $P(C(X))$ we project \hat{Y} onto the constrained model

1 way ANOVA

- $R = \text{eye}(Y) - P$; % projection on error space
- $C = \text{diag}([1 \ -1 \ 0 \ 0])$; % our contrast
- $C0 = \text{eye}(\text{size}(X,2)) - C * \text{pinv}(C)$; % the opposite of C
- $X0 = X * C0$; % the opposite of C into X
- $R0 = \text{eye}(\text{size}(Y,1)) - (X0 * \text{pinv}(X0))$; % projection on E
- $M = R0 - R$; % finally our projection matrix
- $SS_{\text{effect}} = (\text{Betas}' * X' * M * X * \text{Betas})$; % ~ 93.24
- $F = (SS_{\text{effect}} / \text{rank}(X) - 1) / (SS_{\text{error}} / \text{size}(Y,1) - \text{rank}(X))$

1 way ANOVA

- $SS_{\text{total}} = \text{norm}(Y - \text{mean}(Y))^2$;
- $SS_{\text{error}} = \text{norm}(\text{Res} - \text{mean}(\text{Res}))^2$;
- $F = SS_{\text{effect}} / df(C) / SS_{\text{error}} / dfe(X)$
- $df = \text{rank}(C) - 1$;
- $dfe = \text{length}(Y) - \text{rank}(X)$;

Code Summary

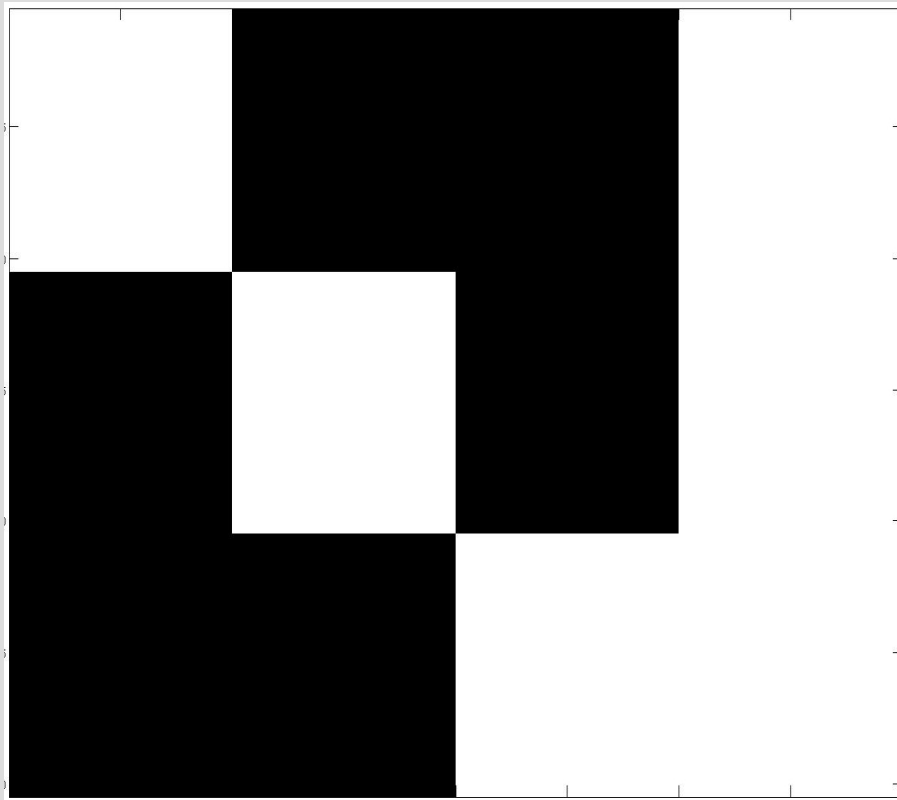
- $Y = XB + e$ % master equation
- $B = \text{pinv}(X)Y$ % betas
- $P = X\text{pinv}(X)$ % projection onto X
- $R = I - P$; % projection on $\text{null}(X)$
- $SS_{\text{Error}} = Y'R Y$; % makes sense
- $C = \text{diag}(\text{Constrast}) \rightarrow C0 \rightarrow R0$
- $M = R - R0$ % projection on $C(X)$
- $SS_{\text{Effect}} = B'X'MXB$ % our effect for C
- $F = (SS_{\text{Effect}}/\text{rank}(C)-1) / (SS_{\text{Error}}/\text{rank}(X)-1);$



Any ANOVAs

1 way

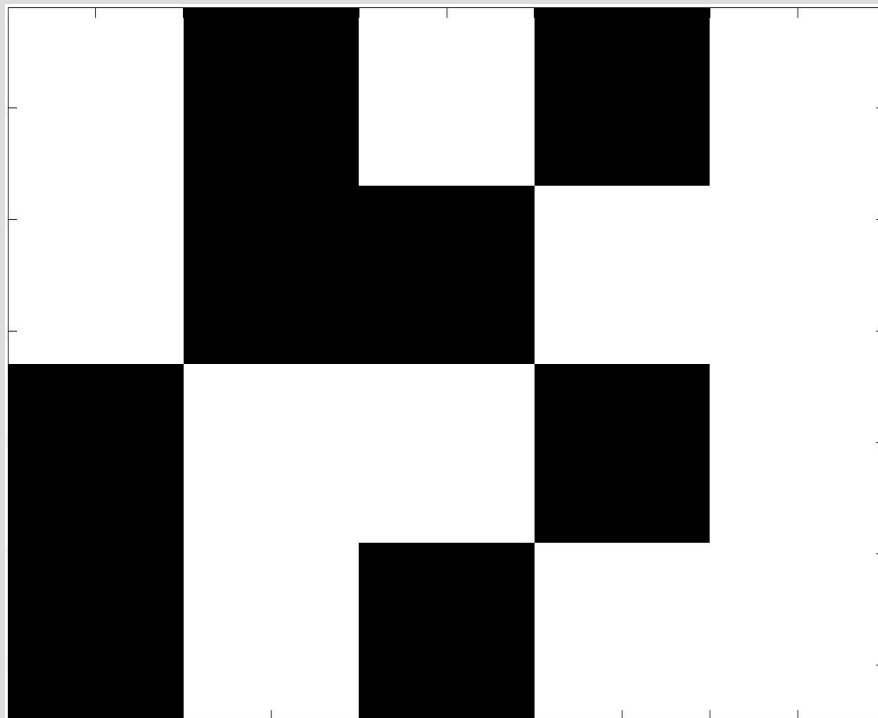
- $F \rightarrow C = \text{diag}([1 \ 1 \ 1 \ 0])$



How much the factor
explain of the data
(thus test against the
mean)

N way

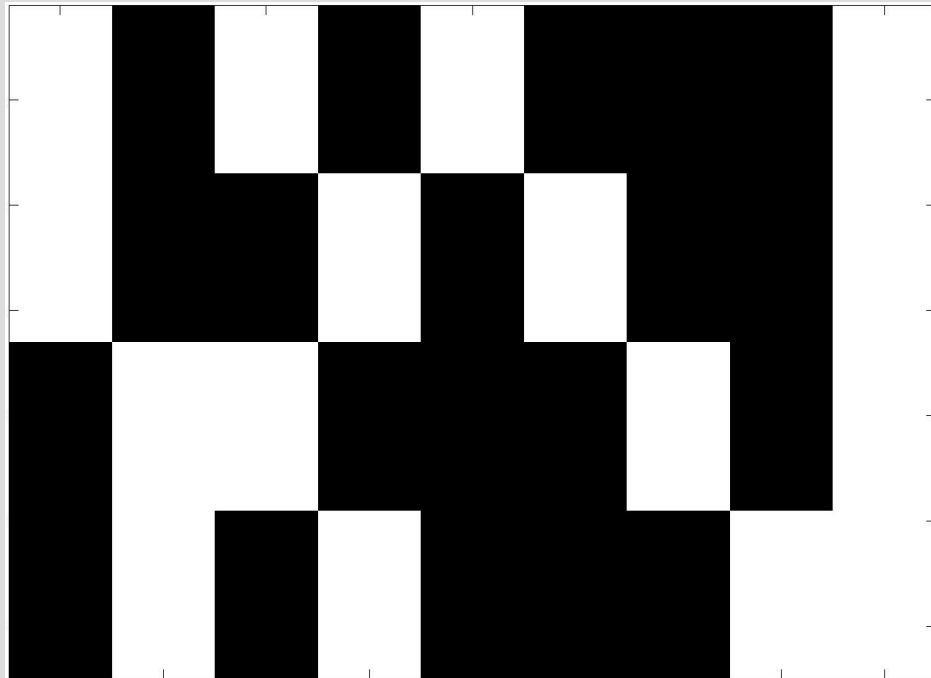
- $F1 \rightarrow C1 = \text{diag}([1 \ 1 \ 0 \ 0 \ 0])$
- $F2 \rightarrow C2 = \text{diag}([0 \ 0 \ 1 \ 1 \ 0])$



How much the factor i explain of the data (thus test against the other factors and the mean)

N way

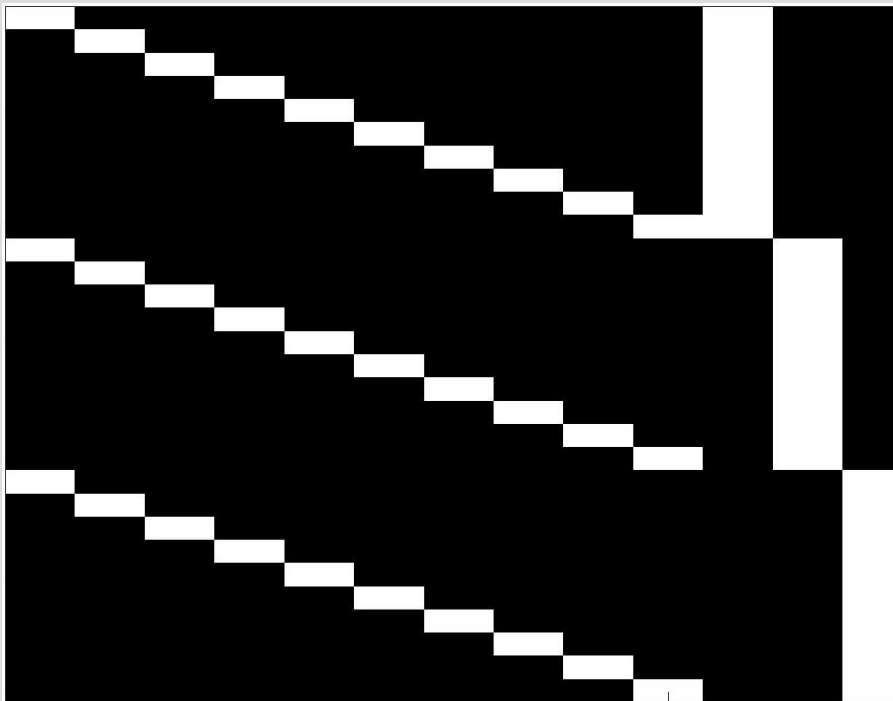
- Interaction: multiply columns
- $\rightarrow C = \text{diag}([0 \ 0 \ 0 \ 0 \ 1 \ 1 \ 1 \ 1 \ 0])$



How much the interaction explains of the data (thus test against the main effects and the mean)

Repeated Measure

- $S \rightarrow Cs = \text{diag}([\text{ones}(10,1) \ 0 \ 0 \ 0 \ 0])$
- $F \rightarrow C = \text{diag}([\text{zeros}(10,1) \ 1 \ 1 \ 1 \ 0])$



The specificity of repeated measures is the subject effect

Note in this model SS error is the SS subject – there is no more grand mean, but a mean per subject



Multivariate Stats

Is this really more difficult?

Multivariate stats

- Before we had one series of Y and a model X --> find B and various effects using C
- Now we have a set of Ys and a model X --> still want to find B and look at various effects
- IMPORTANT: the same model applies to all Ys

Multivariate regression

- We have one experimental conditions and plenty of measures – e.g. show emotional pictures and the subjects have to rate 1,2,3,4 – beside this subjective measurement the researcher measures RT, heart rate, pupil size
- Pblm: RT, heart rate, pupil size are likely to be correlated so doing 3 independent tests is not very informative.

Multivariate regression

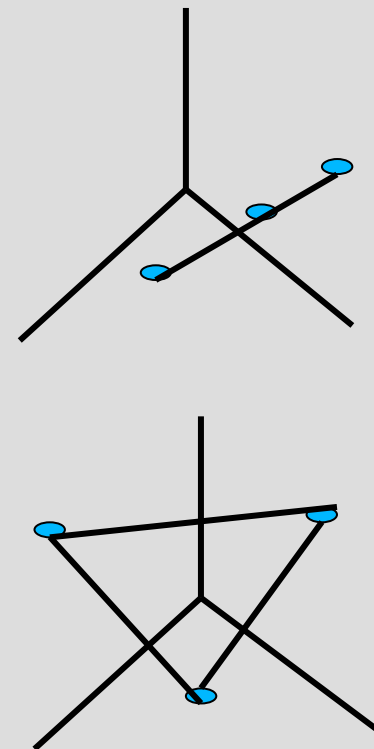
- This can be solved easily since we apply the same model X to all the data Y
- $Y = XB$ and $B = \text{pinv}(X)Y$
- B is now a matrix and the coef. for each Y are located on the diagonal of B
- $SStotal$, $SSeffect$ and $SSerror$ are also matrices called sum square (on the diagonal) and cross products matrices

Multivariate regression

- Load carbig
- $Y = [\text{Acceleration Displacement}]$
- $X = [\text{Cylinders Weights ones}(\text{length}(\text{Weight}), 1)]$
- Then apply the same technique as before
- Note the difference in results
- Multivariate test depends on eigen values ~ PCA
- Take a matrix and find a set orthogonal vectors (eigen vectors) and weights (eigen values) such as $Ax = \lambda x$

Multivariate regression

- 4 tests in multivariate analyses, all relying on the eigen values λ of $\text{inv}(E)^*H$
- Roy $\theta = \lambda_1 / 1 + \lambda_1$
- Wilk $\Lambda = \prod (1/1 + \lambda_i)$
- Lawley-Hotelling $U = \sum \lambda_i$
- Pillai $V = \sum (\lambda_i / 1 + \lambda_i)$





Convolution model

Application to fMRI

General linear convolution model

- $y(t) = X(t)\beta + e(t)$
- $X(y) = u(t) \otimes h(\tau) = \int u(t-\tau) h(\tau) d\tau$
- The data y are expressed as a function of X which varies in time ($X(t)$) but β are time-invariant parameters (= linear time invariant model)
- X the design matrix describes the occurrence of neural events or stimuli $u(t)$ convolved by a function $h(\tau)$ with τ the peristimulus time

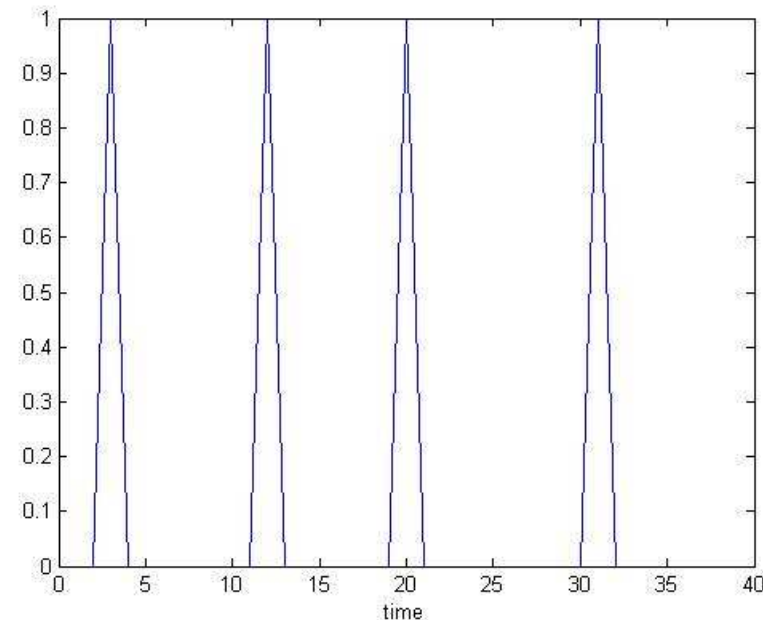
General linear convolution model

- $X(y) = u(t) \otimes h(\tau)$
- Say you have a stimulus (u) occurring every 8/12 sec and you measure the brain activity in between (y)
- If you have an a priori idea of how the brain response is (i.e. you have a function h which describes this response) then you can use this function instead of 1s and 0s

General linear convolution model

- $X(y) = u(t) \otimes h(\tau)$

```
u = [0 0 1 0 0 0 0 0 0 0 0 1 0 0 0 0 0 0 0 1 0 0 0
      0 0 0 0 0 0 0 1 0 0 0 0 0 0 0 0 0 0];
```

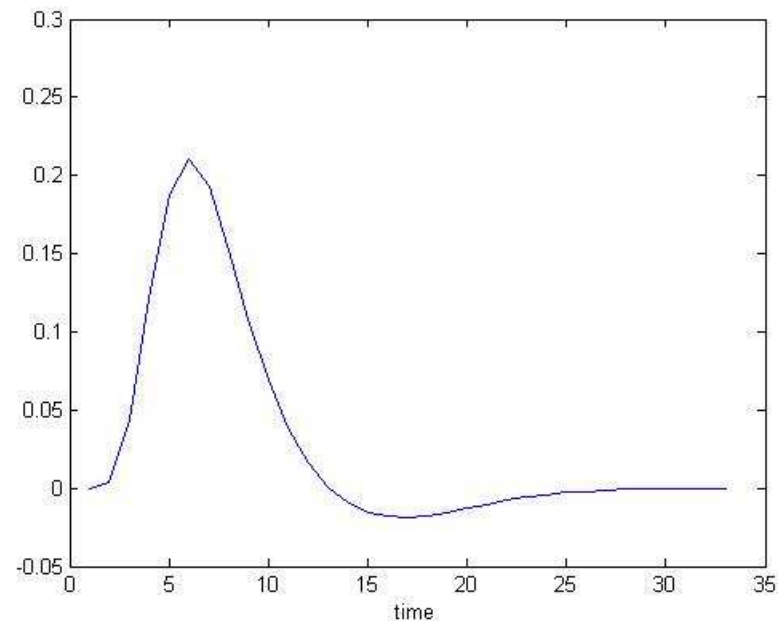


General linear convolution model

- $X(y) = u(t) \otimes h(\tau)$

$u = [0 \ 0 \ 1 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 1 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 1 \ 0 \ 0 \ 0$
 $0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 1 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0 \ 0];$

$h = \text{spm_hrf}(1)$



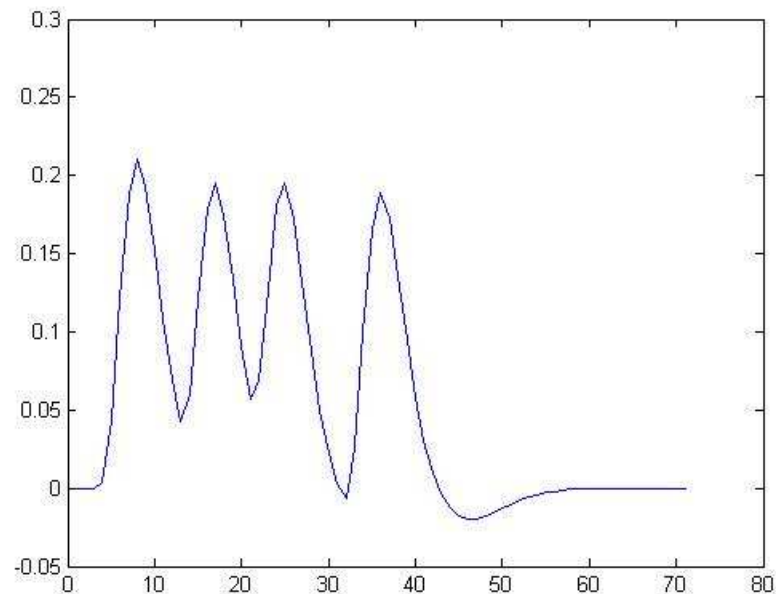
General linear convolution model

- $X(y) = u(t) \otimes h(\tau)$

```
u = [0 0 1 0 0 0 0 0 0 0 0 1 0 0 0 0 0 0 0 1 0 0 0  
     0 0 0 0 0 0 0 1 0 0 0 0 0 0 0 0 0 0];
```

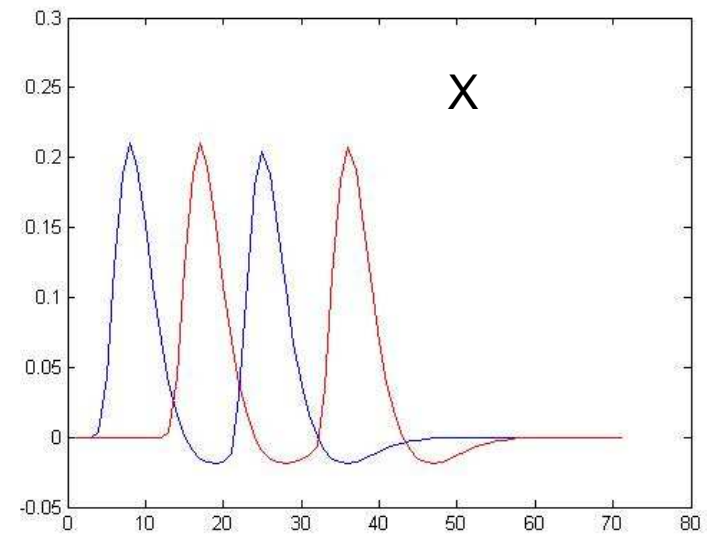
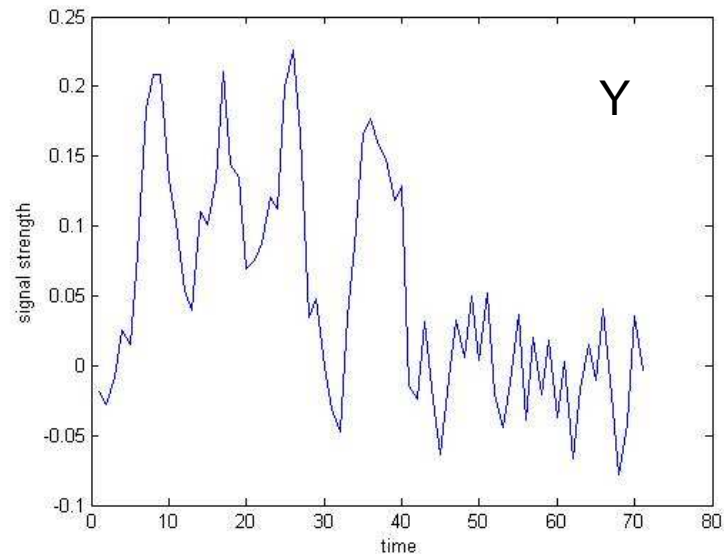
```
h = spm_hrf(1)
```

```
X = conv(u,h);
```

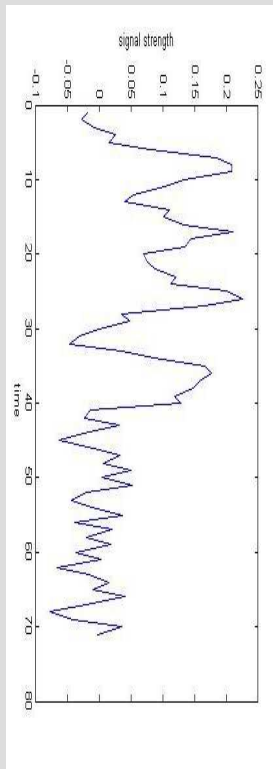


General linear convolution model

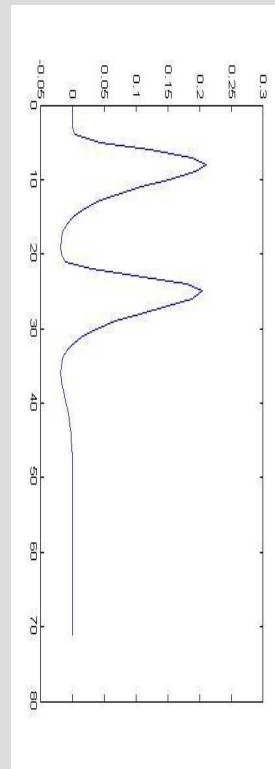
- $y(t) = X(t)\beta + e(t)$
- X has now two conditions u_1 and u_2 ...
- And we search the beta parameters to fit Y



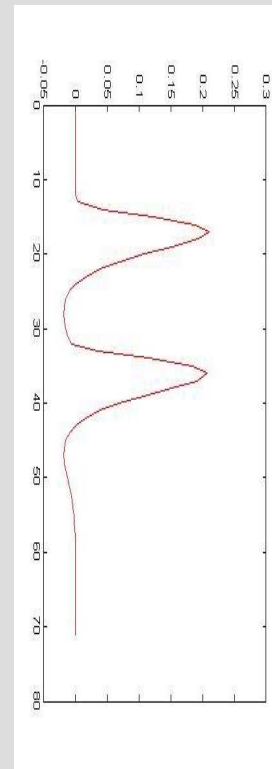
General linear convolution model



=



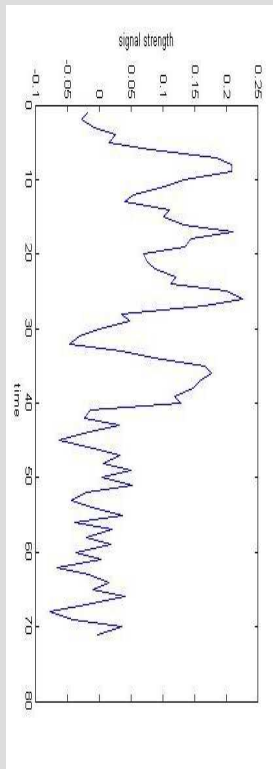
β_1 +



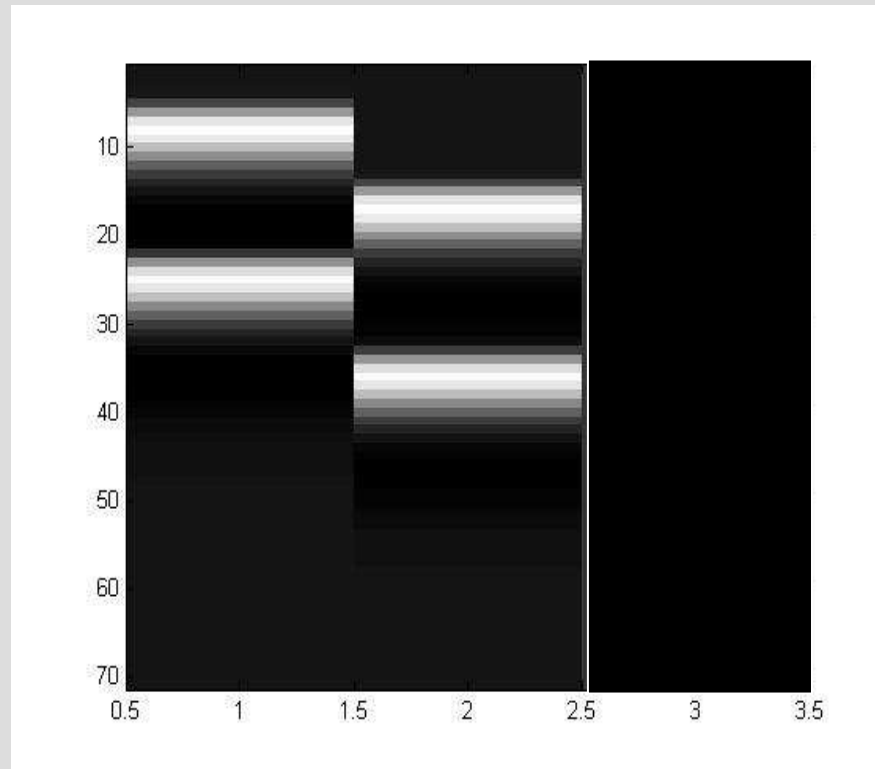
$\beta_2 + u + e$

$$\text{Data} = \text{cond. 1} * \beta_1 + \text{cond. 2} * \beta_2 + \text{cst} + e$$

General linear convolution model



=



β_1
 $\beta_2 + e$
 u

fMRI (one voxel) = Design matrix * Betas + error



Generalized linear model

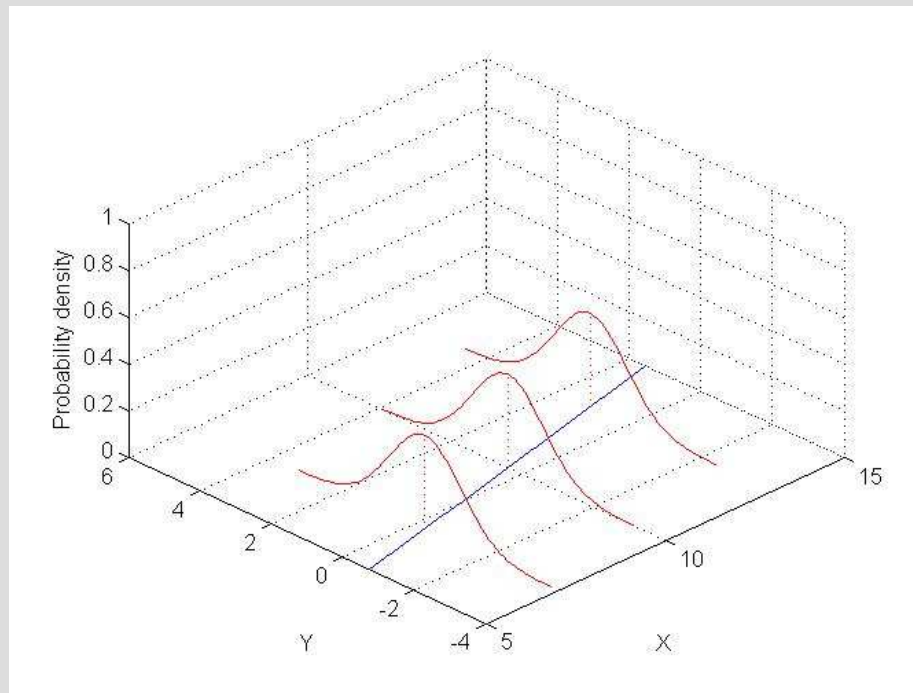
Move between
distributions

Generalized linear models

- You still have your responses Y but they follow a distribution that may be binomial, poisson, gamma, etc..
- You still make a model (design matrix) X and you search for a coefficient vector β
- Here in addition, there is a link function $f(.)$ such as $f(Y)=X\beta$

Generalized linear models

- Usually, Y is a normally distributed response variable with a cste variance and for a simple case can be represented as a straight line ($y = c + ax$) with Gaussian distributions about each point



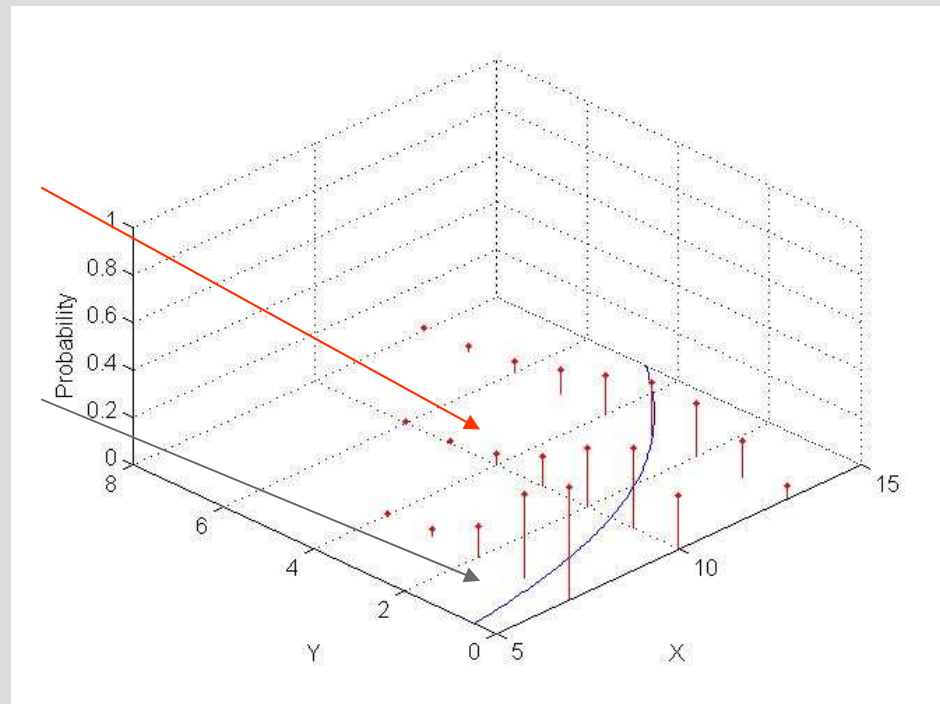
Generalized linear models

- In a generalized linear model, the mean of the response is modelled as a monotonic nonlinear transformation of a linear function of the predictors, $g(b_0 + b_1x_1 + \dots)$. The inverse of the transformation g is known as the "link" function.

Gamma distributed data

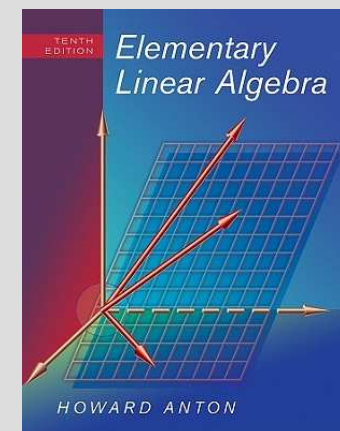
Link function

See glmfit



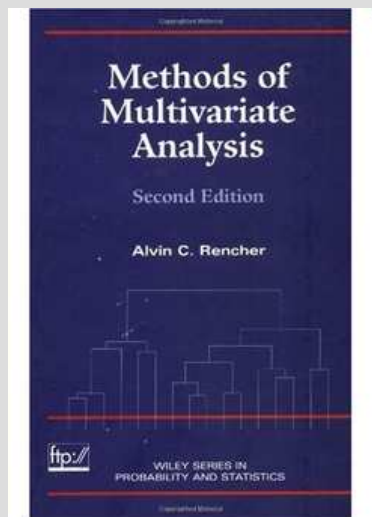
References

- **Linear algebra**
- MIT open course by prof. Strang (best course ever!)
- <http://ocw.mit.edu/courses/mathematics/18-06-linear-algebra-spring-2010/>
- Elementary Linear Algebra
By Howard Anton (best book ever! 10th edition)



References

- Stats with matrices
- Plane Answers to complex questions by Ronald Christensen



Methods of Multivariate
Analysis by Alvin Rencher

