#### General Linear Model

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#### Purpose of this lecture

 Give you the keys to understand how any statistics can be performed using the GLM

- > What are the maths behind (easy)
- > How is the GLM used in fMRI
- > Key concepts (design matrix, linearity, independence, orthogonality, variance, contrasts)

#### Overview

- General LINEAR model: what is linear?
- General linear MODEL: what is a model?
- GENERAL linear model: why is that general?
- Even more general:
- → General linear convolution model (fMRI)

# What is linearity?

Correlation analyses

#### What is a linear model?

■ An equation or a set of equations that models data and which corresponds geometrically to straight lines, plan, cubes, hypercubes ...

Linear models have two properties:

<u>scaling</u>: the magnitude of the system output is proportional to the system input (y = ax)

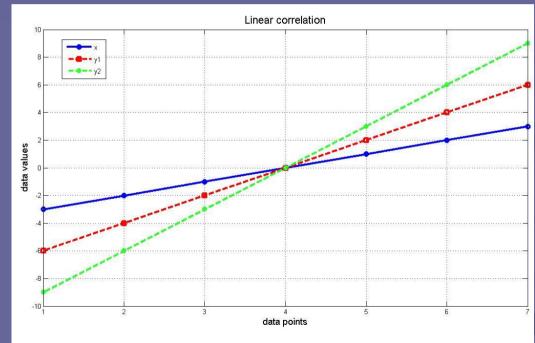
<u>superposition</u>: the total response to a set of inputs is the sum of individual inputs (y = x1+x2)

#### What is linear?

■ Example of a linear correlation

$$x = [-3 -2 -1 0 1 2 3];$$

Scaling: y1 = 2\*x;Superposition y2 = x+2\*x;



Pearson corr. = 1
corr([x' y'],'type','Pearson')

#### What is NOT linear?

■ Example of a non-linear correlation

$$x = [-3 -2 -1 \ 0 \ 1 \ 2 \ 3];$$

Non-linear scaling  $y1 = x.^2$ ; Non-linear superposition y2 = x + abs(x) Non Linear correlation

Non Linear correlation

A data points

Pearson corr. = 0 corr([x' y'], 'type', 'Spearman') = 1 / 0.9

# What is a linear model?

Regression analyses

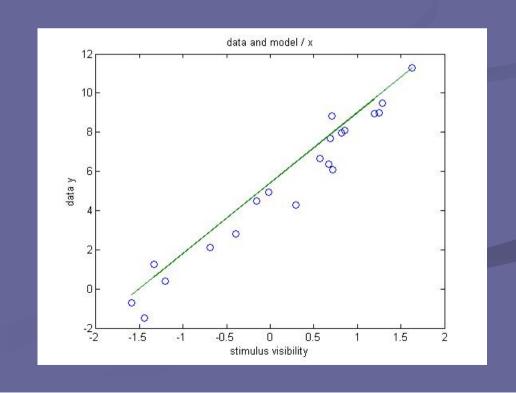
### Simple regression

■ An equation that models data and which corresponds geometrically a straight line

$$\rightarrow$$
 simple regression:  $y = \beta x + b$  (perf =  $\beta *age + b$ )

$$x = randn(20,1);$$
  
 $y = 3*x + 5 + e$ 

minimize the square distance between the line and each points = least squares fit



■ A equation that models data and which corresponds geometrically to a plan / cube ..

$$\Rightarrow$$
 y =  $\beta_1 x_1 + \beta_2 x_2 + b$  (perf =  $\beta_1 *age + \beta_2 *IQ + b$ )

- It is again solved by the least squares method, i.e. one looks for coefficients (Betas) that minimizes the error, i.e. the difference between  $(\beta_1 x_1 + \beta_2 x_2 + b)$  and y
- This time, instead of a line, we will have a plan as we have 2 regressors

# Digression into Linear Algebra

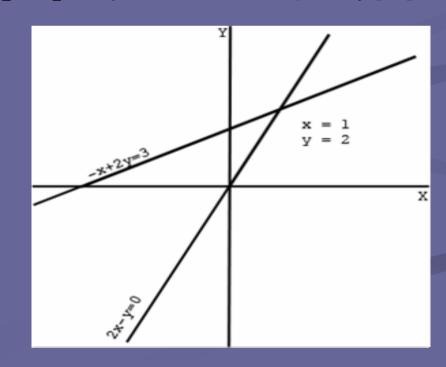
from equations to matrices

### Linear equations

- Linear algebra has to do with solving linear systems, i.e. a set of linear equations
- For instance we have observations (y) for a stimulus characterized by its property x such as  $y = \beta_1 x_1$

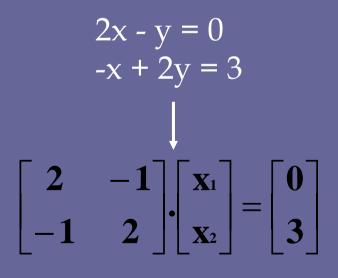
$$y = 2x$$
$$y = (3 + x) / 2$$

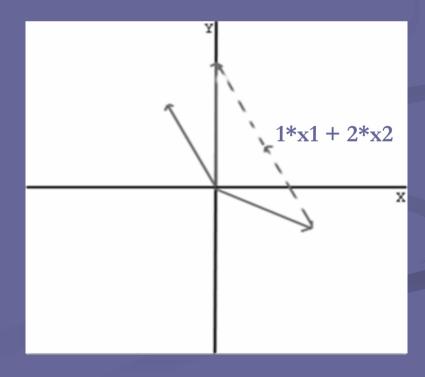
$$\frac{\text{Solve}}{2x - y} = 0$$
$$-x + 2y = 3$$



# Linear equations

■ With matrices, we change the perspective and try to combine columns instead of rows, i.e. we look for the coefficients with allow the linear combination of vectors



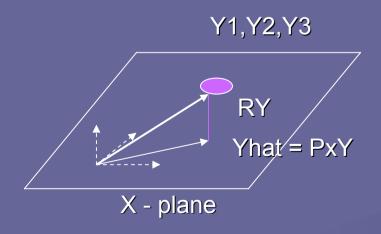


#### Linear equations

- The advantage of matrices is that one can handle a huge set a variables and numbers at once instead of working out the equations one by one; in addition, it offers a fast and reliable way to solve a linear system
- With a single equation  $x \beta = y$  the solution is  $\beta = y(1/x)$
- What about a square matrix X such as  $X \beta = y$ ?
- The solution is similar, the solution is Y \* by the inverse of X (for non square design matrix X, solution is inv(X<sup>T</sup>X)X<sup>T</sup>Y or pinv(X)Y)

#### Geometrical perspective

- Y and X $\beta$  are two points in  $\mathbb{R}^J$  and we want to minimize the distance between those points.
- In addition the point  $X\beta$  lies in a subspace of  $\mathbb{R}^J$  (dimension rank(X)) spanned by the columns of X
- The orthogonal projection of Y onto X is Yhat



$$Y = X\beta + e$$
 $Px = X(X^TX)^{-1}X^T$ 
 $e = RY$ 
 $Y = X\beta + e$ 
 $Y = X(X^TX)^{-1}X^T$ 

# Back to linear models (this time with matrices)

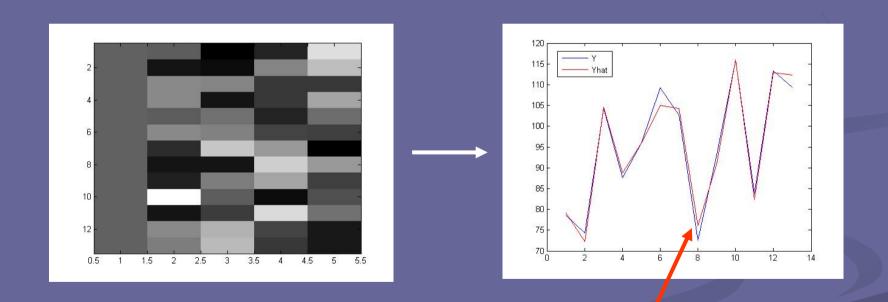
Regression analyses

■ Lets take an example

78.5000		C		$7 \beta_1$	$26 \beta_2$		$6 \beta_3$	$60 \beta_4$		e1
74.3000		C		$1 \beta_1$	$29 \beta_2$		$15  \hat{\beta}_3$	$52 \beta_4$		e2
104.3000		C		11 β <sub>1</sub>	$56 \beta_2$		$8 \beta_3$	$20 \beta_4$		e3
87.6000		C		$11  \beta_1$	$31 \beta_2$		$8 \beta_3$	$47  \beta_4$		e4
95.9000		C		$7  \beta_1$	$52 \beta_2$		$6 \beta_3$	$33 \beta_4$		e5
109.2000	=	C	+	11 β <sub>1</sub> +	$55 \beta_2$	+	$9 \beta_3$	$22 \beta_4$	+	e6
102.7000		C		$3 \beta_1$	$71 \beta_2$		$17  \beta_3$	$6 \beta_4$		e7
72.5000		C		$1  \beta_1$	$31 \beta_2$		$22 \beta_3$	$44 \beta_4$		e8
93.1000		C		$2 \beta_1$	$54 \beta_2$		$18  \beta_3$	$22 \beta_4$		e9
115.9000		C		$21 \beta_1$	$47 \beta_2$		$4 \beta_3$	$26 \beta_4$		e10
83.8000		C		$1  \beta_1$	$40  \beta_2$		$23 \beta_3$	$34 \beta_4$		e11
113.3000		C		$11  \beta_1$	$66 \beta_2$		$9 \beta_3$	$12 \beta_4$		e12
109.4000		C		$10  \beta_1$	$68 \beta_2$		$8 \beta_3$	$12 \beta_4$		e13

■ In a Matrix form this can be rewritten as

- >  $Y = X\beta + e$  ... we want  $\beta = X^{-1}Y$
- > The solution is  $\beta = (X^TX)^{-1}X^TY \rightarrow X^{\sim}Y$  (matlab inv, pinv)



The modelled data Yhat =  $X\beta$ 

- Stats:
- $Y = X \beta + e \rightarrow Yhat = X \beta \rightarrow e = Y-Yhat$
- R2 = SS effect / SS total =  $\Sigma$ (yhat i – mean(yhat))<sup>2</sup> /  $\Sigma$ (y i – mean(y))<sup>2</sup>
- F value = SS effect / SS error =  $\Sigma$ (yhat i - mean(yhat))<sup>2</sup> /  $\Sigma$ (e i - mean(e))<sup>2</sup>  $\tau$  rank(X) - 1 N - rank(X)

- Usual problems in multiple regression
- → Linear independence / colinearity
- Colinearity: indicates that a set of points are on a single straight line
- Multicollinearity is a statistical phenomenon in which two or more predictor variables are correlated
- Linear independence: in a family of vectors none can be describe as the linear combination of the others

 Since we describe a regression model as a set of vectors, multicolinearity and linear independence have similar meaning

$$X = \begin{cases} 1 & 3 & 5 \\ 3 & 8 & 14 \\ 4 & 5 & 13 \\ 8 & 6 & 22 \\ 7 & 8 & 22 \\ 4 & 1 & 9 \end{cases} \rightarrow x3 = 2x1 + x2$$

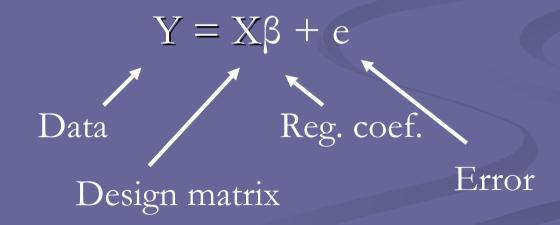
$$\Rightarrow corr([x1', x3']) = .9$$

$$\Rightarrow corr([x2' x3']) = .7$$

The matrix  $\underline{rank}$  is the number of independent columns (= effective df; usually df = rank – 1 because of the cst term)

#### Intermediate summary

- Linear models are equations that describe lines, cubes, hypercubes, etc..
- A regression is a linear model of the data
- It can be solve using linear algebra (matrices)



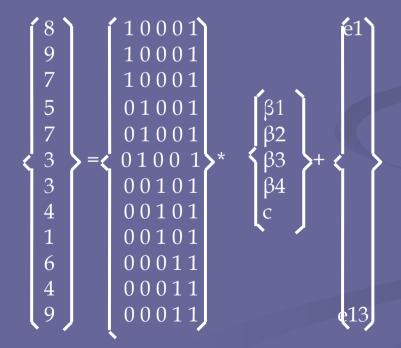
# General Linear Model

All the stats with the same basic algebra

# One-way ANOVA

Y	Gp
8	1
9	1
7	1
5	2
7	2
3	2
3	3
4	3
1	3
6	4
4	4
9	4

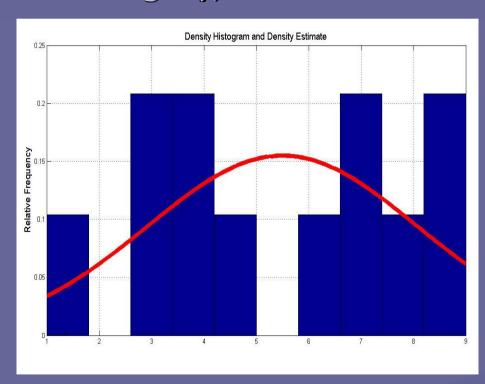
$$y(1..3)1=1x1+0x2+0x3+0x4+c+e11$$
  
 $y(1..3)2=0x1+1x2+0x3+0x4+c+e12$   
 $y(1..3)3=0x1+0x2+1x3+0x4+c+e13$   
 $y(1..3)4=0x1+0x2+0x3+1x4+c+e13$ 

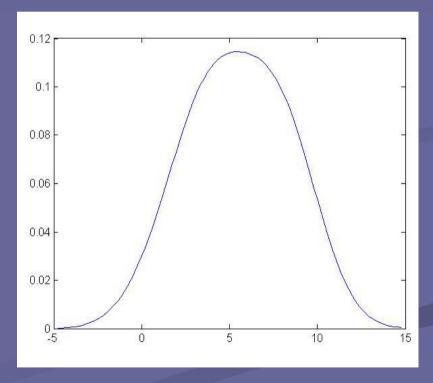


→ This is like the multiple regression except that we have ones and zeros instead of 'real' values

Y = [8 9 7 5 7 3 3 4 1 6 4 9]';

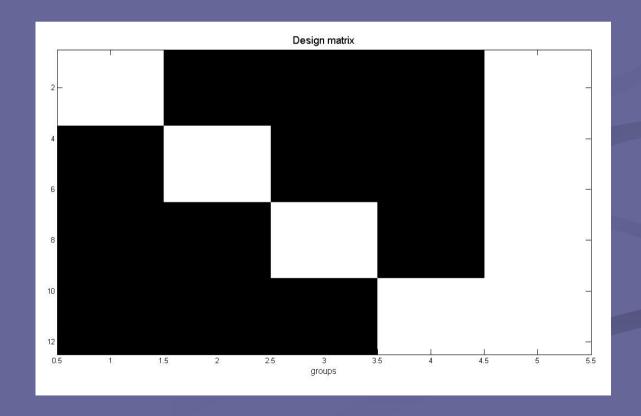
N = length(y); % = 12



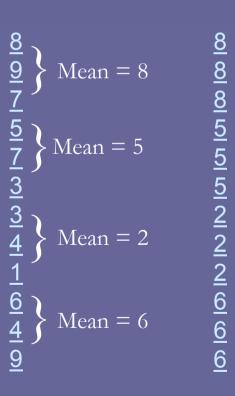


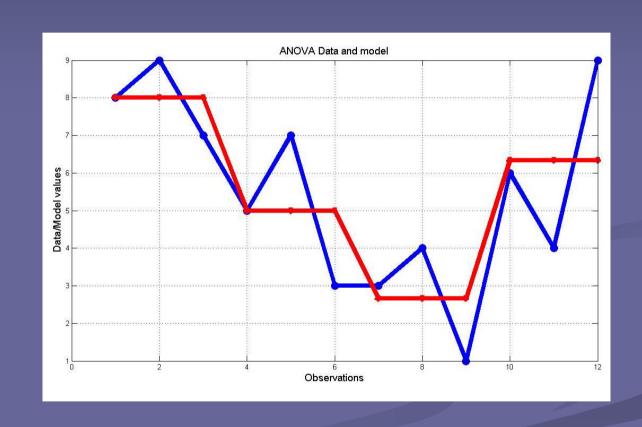
[f,z] = ksdensity(y); plot(z,f)

X = [kron(eye(4),ones(3,1)) ones(N,1)]; Imagesc(X); colormap(gray)



$$B = pinv(X)*Y;$$
  
 $Yhat = X*B;$ 





 $Y \longrightarrow Yhat$ 

```
SStotal = norm(y - mean(y)).^2; % another way to think about sum of squares is a squared distance in \mathbb{R}^J
```

SSeffect = norm(yhat – mean(yhat). $^2$ ; % same as above but in  $|R^X|$ 

SSerror = norm(e – mean(e)). $^2$ ; % this time in  $\mathbb{R}^{X^{-1}}$ 

R2 = SS effect / SS errorF = SS effect / rank(X)-1) / SS error / N-rank(x);

 $\rightarrow$  R square and F as for the regression  $\odot$ 

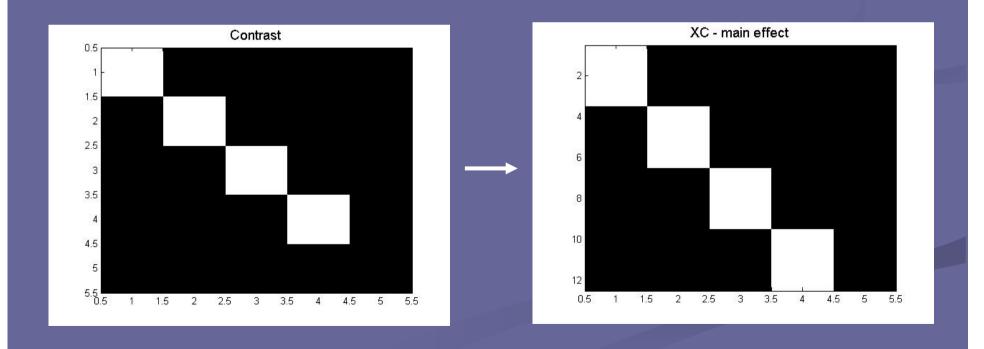
- Using a contrast C is like constraining the X space
- Then any test using a contrast matrix C will be like testing a combination of the columns again the error One can thus demonstrate that

$$F = \begin{array}{cccc} \beta^T X^T M X \beta \cdot J - p & M = R0 - R \\ R0 = I - X0pinv(X0) \\ X0 = XC0 \\ Y^T R Y & p1 & C0 = I - Cpinv(C) \end{array}$$

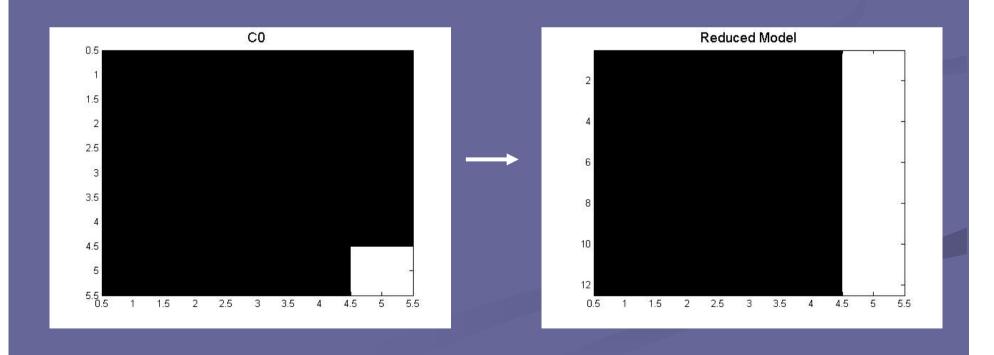
C = eye(5); C(5,5) = 0;

% main effect = all columns except the grand avg

Xc = X\*C; % this is the model to test



C = eye(5); C(5,5) = 0; Xc = X\*C C0 = eye(rank(X)+1) - C\*pinv(C); % this is the opposite of C X0 = X\*C0; % this is thus the reduced model



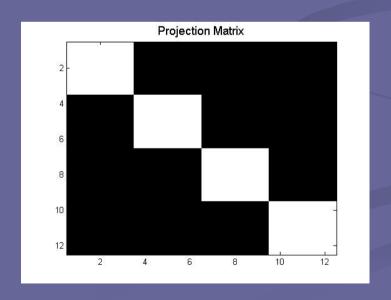
```
C = eye(5); C(5,5) = 0; Xc = X*C;

C0 = eye(rank(X)+1) - C*pinv(C); X0 = X*C0;

R = eye(length(Y)) - (X*pinv(X)); % residual matrix for X

R0 = eye(length(Y)) - (X0*pinv(X0)); % residual matrix for X0

M = R0 - R; % M is the projection matrix onto Xc
```

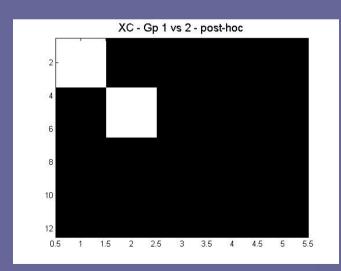


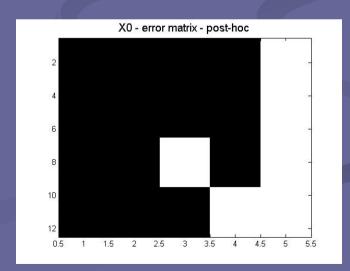
```
C = eye(5); C(5,5) = 0; Xc = X*C;
C0 = eye(rank(X)+1) - C*pinv(C); X0 = X*C0;
R = eye(length(Y)) - (X*pinv(X));
R0 = eye(length(Y)) - (X0*pinv(X0));
M = R0 - R; dfe = length(Y) - rank(X);
F = ((beta'*X'*M*X*beta)/(rank(C)-1))/((Y'*R*Y)/dfe);
  = 4.45 \odot
                                     df = rank(C) - 1
```

= Yhat'\*Yhat projected onto Xc

■ We can also contrast only 2 columns for a post-hoc test – note that the error matrix will contains all the other columns, i.e. ≠ t-test between these 2 columns only

$$C = eye(5); C(:,3:5) = 0; Xc = X*C$$
  
 $C0 = eye(rank(X)+1) - C*pinv(C); X0 = X*C0;$ 





#### GLM: contrasts

```
C = eye(5); C(:,3:5) = 0; Xc = X*C

C0 = eye(rank(X)+1) - C*pinv(C); X0 = X*C0;

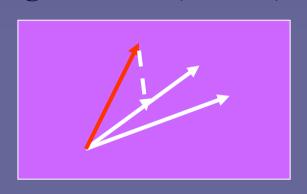
var = ((R*Y)'*(R*Y)) / dfe;

T = (C*Betas) ./ sqrt(var.*(C*pinv(X'*X)*C'));
```

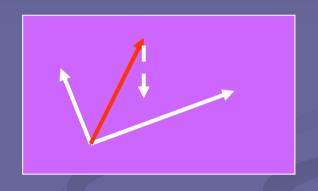
- → In SPM, you specify your design matrix and then you enter your contrasts
- $\rightarrow$  The T test is unilateral (A>B  $\neq$  A<B)

### Orthogonalization

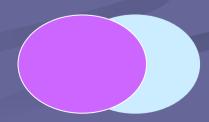
■ It sometimes happens that regressors have to be orthogonalized (~PCA)



Lot of variance shared - because we look for the unique part of variance, the shared part goes into the error



Orthogonalization removes shared variance BUT order matters! (like step by step regression)



## Orthogonalization

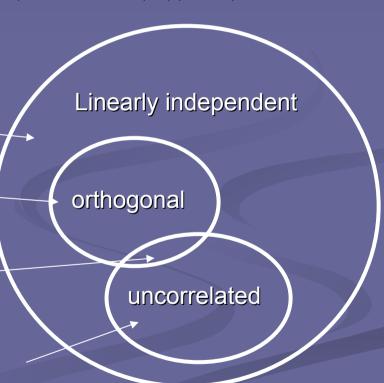
Linearly independent  $(Y \neq aX)$ , orthogonal (X'Y = 0) and uncorrelated (X-mean(X))'(Y-mean(Y))=0) variables

[1 1 2 3] and [2 3 4 5] Independent, correlated, not orthogonal

[1 -5 3 -1] and [5 1 1 3] Independent, correlated and orthogonal

[-1 -1 1 1] and [1 -1 1 -1] Independent, uncorrelated and orthogonal

[0 0 1 1] and [1 0 1 0] Independent, uncorrelated, not orthogonal



# Intermediate summary 2

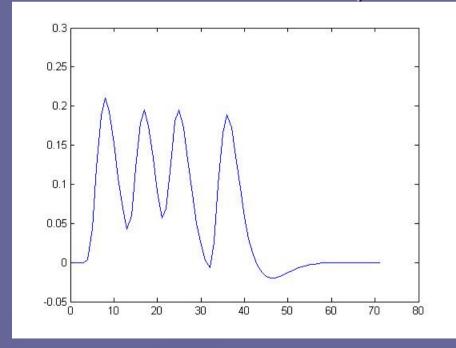
- Any statistics can be performed using the GLM
- The design matrix describes your experiment but is also the model of your statistics (gp, repeated measures, covariates, interactions ...)
- GLM links arithmetic and geometry (nice), meaning that we can combine vectors to create subspaces (contrasts) to test various effects
- Regressors (vectors) have to be independent but can still be correlated and not orthogonal (although the more uncorrelated and close to orthogonal the better)

# General convolution model (fMRI)

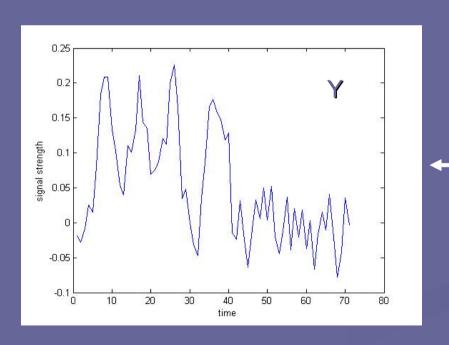
- $y(t) = X(t)\beta + e(t)$
- The data y are expressed as a function of X which varies in time (X(t)) but  $\beta$  are time-invariant parameters (= linear time invariant model)
- X the design matrix describes the occurrence of neural events or stimuli u(t) convolved by a function h(τ) with τ the peristimulus time

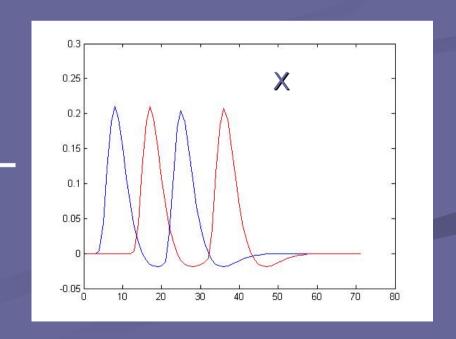
- **u** $(y) = u(t) \otimes h(\tau)$
- Say you have a stimulus (u) occurring every 8/12 sec and you measure the brain activity in between (y)
- If you have an a priori idea of how the brain response is (i.e. you have a function h which describes this response) then you can use this function instead of 1s and 0s

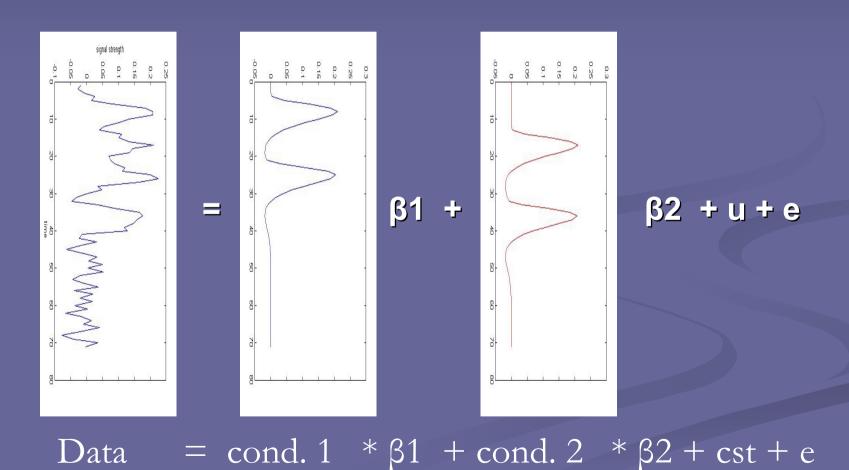
 $h = spm_hrf(1)$ X = conv(u,h);

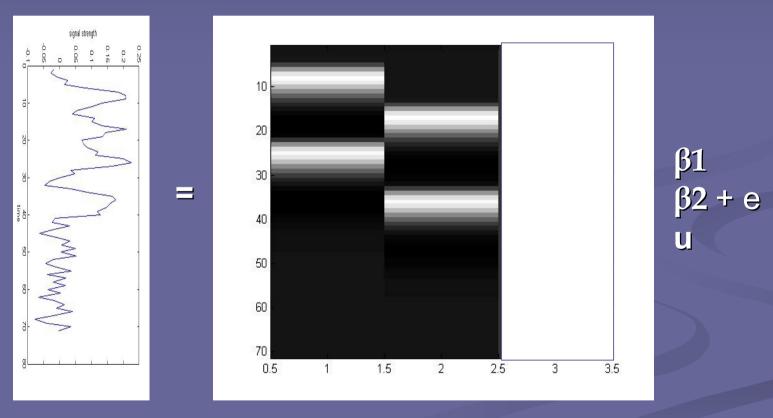


- $y(t) = X(t)\beta + e(t)$
- X has now two conditions u1 and u2 ...
- And we search the beta parameters to fit Y

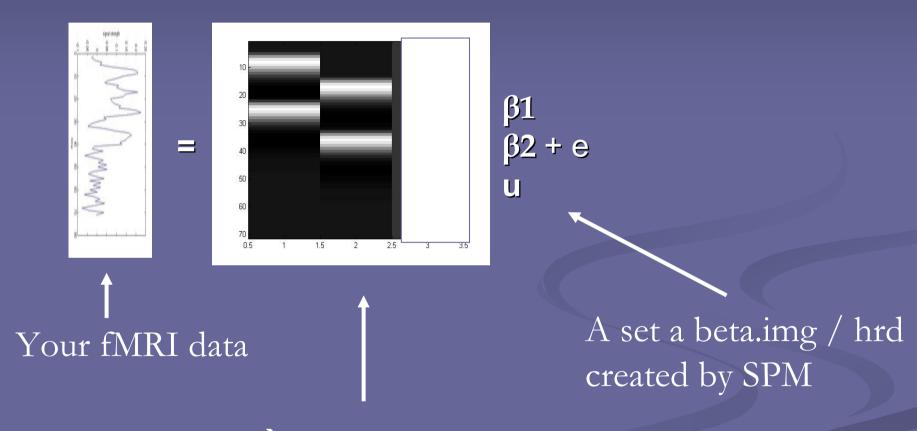








fMRI (one voxel) = Design matrix \* Betas + error



SPM.mat  $\rightarrow$  created by you by entering the different conditions

# Final Summary

#### General linear model

- Y=XB+e is the master equation
- Describing X is the most important thing you can model any design (Regression, ANOVA, ANCOVA, etc)
- ?? Well having good fMRI Y data matters too ②
- SPM uses a convolution but after all, you can still thing about this as a linear regression