#### 1

# AI1103 Assignment-7

### SRIVATSAN T - CS20BTECH11062

Download all python codes from

https://github.com/CS20BTECH11062/AI1103/tree/main/Assignment-7/codes

and latex-tikz codes from

https://github.com/CS20BTECH11062/AI1103/tree/main/Assignment-7/Assignment-7.tex

## QUESTION (CSIR UGC NET June 2013 Q.59)

Let  $U_1, U_2, \dots, U_n$  be independent and identically distributed random variables each having a uniform distribution on (0,1). Then,

$$\lim_{n\to+\infty} \Pr\left(U_1+U_2\ldots,U_n\leq \frac{3}{4}n\right)$$

- 1) does not exist
- 2) exists and equals 0
- 3) exists and equals 1
- 4) exists and equals  $\frac{3}{4}$

#### **SOLUTION**

We use Weak law for large numbers to solve this problem. Let the collection of identically distributed random variables  $U_1, U_2 ..., U_n$  have a finite mean  $\mu$  and finite variance  $\sigma^2$ .

$$\mu = E(U_i) \text{ for } i \in (1,2,3...,n)$$
 (0.0.1)

Since the distribution is uniform on (0,1),  $\mu = 0.5$ . Let  $M_n$  be the sample mean

$$M_n = \frac{U_1 + U_2 + U_3 \dots + U_n}{n} \tag{0.0.2}$$

Expected value of  $M_n$  (using (0.0.2) and (0.0.1))is

$$E(M_n) = \frac{E(U_1 + U_2 + U_3 + \dots + U_n)}{E(n)}$$
(0.0.3)

$$= \frac{E(U_1) + E(U_2) + \dots + E(U_n)}{n} \qquad (0.0.4)$$

$$=\frac{n\times\mu}{n}\tag{0.0.5}$$

$$=\mu \tag{0.0.6}$$

Variance of M

$$Var(M_n) = \frac{Var(U_1 + U_2 + U_3 \dots + U_n)}{n^2}$$
 (0.0.7)

$$=\frac{Var(U_1) + Var(U_2) \cdots + Var(U_n)}{n^2}$$
(0.0.8)

$$=\frac{n\times\sigma^2}{n^2}\tag{0.0.9}$$

$$=\frac{\sigma^2}{n}\tag{0.0.10}$$

From Chebyshev inequality, for any  $\epsilon > 0$ 

$$\Pr(|M_n - \mu| \ge \epsilon) \le \frac{Var(M_n)}{\epsilon^2}$$
 (0.0.11)

From (0.0.1) and (0.0.10)

$$\Pr\left(\left|\frac{U_1 + U_2 \cdots + U_n}{n} - \mu\right| \ge \epsilon\right) \le \frac{\sigma^2}{n \times \epsilon^2}$$

$$\lim_{n \to \infty} \Pr\left(\left|\frac{U_1 + U_2 \cdots + U_n}{n} - \mu\right| \ge \epsilon\right)$$

$$\le \lim_{n \to \infty} \frac{\sigma^2}{n \times \epsilon^2} \le 0 \text{ for fixed } \epsilon > 0$$
(0.0.12)

But since Probabilities are always non-negative,

$$\lim_{n \to \infty} \Pr\left(\left|\frac{U_1 + U_2 \cdots + U_n}{n} - \mu\right| \ge \epsilon\right) \to 0 \quad (0.0.13)$$

The inverse of (0.0.13) is also true

$$\lim_{n \to \infty} \Pr\left(\left|\frac{U_1 + U_2 \dots + U_n}{n} - \mu\right| \le \epsilon\right) \to 1 \quad (0.0.14)$$

$$\left|\frac{U_1 + U_2 \dots + U_n}{n} - \mu\right| \le \epsilon \text{ as } n \to \infty \qquad (0.0.15)$$

From  $\epsilon$ , n definition of limits, it is clear that

$$\frac{U_1 + U_2 \cdots + U_n}{n} \to \mu \tag{0.0.16}$$

$$U_1 + U_2 \dots U_n \to n \times \mu \text{ as } n \to \infty$$
 (0.0.17)

Since 
$$\mu = \frac{1}{2}$$
,

$$\lim_{n \to \infty} U_1 + U_2 \dots U_n = \frac{1}{2}n < \frac{3}{4}n \qquad (0.0.18)$$

So

$$\lim_{n \to +\infty} \Pr\left(U_1 + U_2 \dots, U_n \le \frac{3}{4}n\right) = 1 \qquad (0.0.19)$$

Correct Option : C