# Workbook: Statistical Thinking for Forensic Practitioners

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# Contents

1	Intr	roduction	5
2	Stat	cistical Preliminaries	7
	2.1	Probability	8
	2.2	Probability to Statistical Inference	19
	2.3	Statistical Inference - Estimation	24
	2.4	Statistical Inference - Hypothesis Testing	29
		Overview of Statistical Preliminaries	
3	Stat	cistics for Forensic Science	35
	3.1	Brief Review of Probability and Statistics	35
		The Forensic Examination	
	3.3	Common Approaches to Assessing Forensic Evidence	35
$\mathbf{R}$	efere:	nces	37

4 CONTENTS

# Chapter 1

# Introduction

This workbook is intended to accompany the Statistical Thinking for Forensic Practitioners workshop taught by members of the Center for Statistics and Applications in Forensic Evidence (CSAFE). The slides for this workshop were constructed initially by Dr. Hal Stern of UC-Irvine and Dr. Alicia Carriquiry of Iowa State University.

When taking the workshop, please follow along with the slides handout (if given) and this workbook. The workbook contains the same material as the slides, with room for you to take notes and to fill in the missing material.

# Chapter 2

# Statistical Preliminaries

Briefly, this section contains a broad review of probability concepts and of statistical inference concepts, with examples from the forensic science context. We will cover probability, data collection, statistical distributions, estimation, and hypothesis testing.

#### 2.0.1 Definitions

population:	
sample:	
· · ·	be thought of as a type of deductive reasoning,
where we are applying general knowledge about the population.	llation of interest to make conclusions about a
statistics: Using knowledge about the Statistics can loosely b	to make statements describing the e thought of as a type of inductive reasoning,
where we are applying knowledge about a sample to spopulation generally.	state that something may be true about the

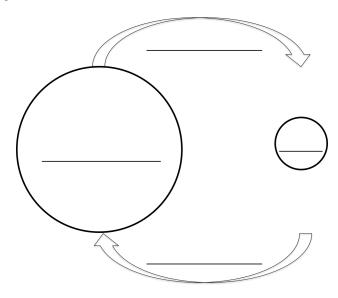


Figure 2.1: "The Big Picture"

### 2.0.2 Forensic Science Examples

• Suppose 100 1-pound bags of heroin are seized on the US-Mexico border, and the FBI want to know the chemical composition of the confiscated drugs to store in their database.
- Population:
- Sample:
• A window was broken in a robbery, and the suspect who was apprehended nearby had glass fragments lodged in the soles of their shoes. Do the fragments from the suspect's shoes have the same or similar chemical composition as the broken window?
- Population 1:
- Sample 1:
- Population 2:
- Sample 2:
<ul> <li>A city government employee is suspected of embezzling funds from the city's coffers. Forensic accountants examine a subset of the city's transactions to determine whether embezzling occurred and how much money was lost.</li> </ul>
- Population:
- Sample:
How do you think this pertains to pattern evidence? List some possible relevant populations and samples below.
• Population 1:
• Sample 1:
• Population 2:
• Sample 2:
• Population 3:
• Sample 3:
2.1 Probability
Probability concerns the <i>uncertainty</i> of outcomes. The set of all possible outcomes is called the space, and a particular outcome or set of outcomes of interest is referred to as an
2.1.1 Examples

- 1. Footwear

  - Event = Shoe of size 9

2.1. PROBABILITY 9

- 2. Footwear
  - Sample Space = Brand of shoe e.g. { Nike, Vans, Converse, ...}
  - Event = Nike sneaker
- 3. Firearms
  - Sample Space = CMS (consecutive matching striae) for a pair of bullets e.g.  $\{0, 1, 2, 3, 4, \dots\}$
  - Event = CMS of 10 or more

#### 2.1.2 Interpretation

The probability of observing an event in a sample space is a number less than or equal to 1 and greater than or equal to 0 that describes the \_\_\_\_\_\_ that the event will occur.

There are two primary interpretations of probability:

- 1. The long run \_\_\_\_\_\_ of occurrence of an event.
- 2. The \_\_\_\_\_\_ belief of likelihood of an event occurring.

#### 2.1.3 Basic Notation and Laws of Probability

Let an event of interest be denoted by \_\_\_\_\_\_. The probability of this event occurring is then denoted \_\_\_\_\_\_. Recall that the probability of an event is always between 0 and 1. When P(Y) = 0, the event Y will never happen. When P(Y) = 1, the event Y will always happen. The sum of the probabilities of all possible outcomes in the sample space always equal to \_\_\_\_\_\_.

The event of interest, Y, also has a complement event,  $\overline{Y}$ , which is read as "not Y". The complement,  $\overline{Y}$ , of an event, Y, is itself an event containing all outcomes in the sample space other than that initial event of interest, Y.

$$P(Y) + P(\overline{Y}) = \underline{\hspace{1cm}}$$

The above equation also gives us the following rules:

$$P(Y) = 1 - P(\overline{Y})$$

$$P(\overline{Y}) = 1 - P(Y)$$
(2.1)

#### 2.1.4 Probability and Odds

The probability of an event defines the odds of the event. The odds in favor of an event Y are defined as the probability of Y divided by the probability of everything except Y ("not Y"):

$$O(Y) = \frac{P(Y)}{P(\overline{Y})} = \frac{P(Y)}{1 - \underline{\hspace{1cm}}}.$$

Conversely, the odds against a event Y are defined as the probability of everything except Y ("not Y") divided by the probability of Y:

$$O(\overline{Y}) = \frac{P(\overline{Y})}{P(Y)} = \frac{1 - \underline{\hspace{1cm}}}{P(Y)}.$$

When we typically talk about odds, like in horse racing, the odds reported are the odds *against* the outcome of interest. Let's construct a horse race scenario using our probability notation to find the probability of a horse winning a race from the reported odds:

- Suppose you want to place a bet on a horse name Cleopatra winning the race. Odds for Cleopatra are reported as 4:1.
- Y =Cleopatra wins the race
- $\overline{Y}$  = Any horse in the race other than Cleopatra wins the race.
- $O(\overline{Y}) = \frac{P(\overline{Y})}{P(Y)} = \frac{4}{1} = 4$
- We know that  $P(Y)+P(\overline{Y})=1$ . With this information, we can determine P(Y), which is the probability that Cleopatra wins the race:

$$O(\overline{Y}) = \frac{P(\overline{Y})}{P(Y)} = 4$$

$$\Rightarrow \frac{P(\overline{Y})}{P(Y)} = 4$$

$$\Rightarrow \frac{1 - P(Y)}{P(Y)} = 4 \qquad (See Equation 2.1)$$

$$\Rightarrow \frac{1}{P(Y)} - 1 = 4$$

$$\Rightarrow \frac{1}{P(Y)} = 5$$

$$\Rightarrow P(Y) = \frac{1}{5} = 0.2$$

$$\Rightarrow P(\overline{Y}) = 0.8$$

- So, the odds for Cleopatra (4:1) mean that Cleopatra has a probability of 0.2 of winning the race. Because this outcome is not very likely (it will only happen in 1 race out of 5), you win money if Cleopatra wins simply because that is not a likely outcome.
- Betting: Suppose you bet \$1 on Cleopatra to win the race with 4:1 odds. You will win \$4 if Cleopatra wins, otherwise you've lost \$1.
- The amount you win (\$4) is determined so that you break even in the long run.
- Suppose 5 identical races are run. In 1 of those races, Cleopatra will win, and in the other 4, Cleopatra will lose. If you bet \$1 on Cleopatra in each race, you will lose that \$1 4 of 5 times. So, in order for you to break even, the designated amount you'll win when Cleopatra wins is \$4.
- This is a statistical concept known as *expected value*. Your expected value when placing the bet is \$0. We compute expected value by multiplying each possible outcome value by its probability and adding them all together:

$$\$4 \cdot P(Y) + (-\$1) \cdot P(\overline{Y}) = 0$$
$$\$4 \cdot 0.2 + (-\$1) \cdot 0.8 = 0$$
$$\$0.8 - \$0.8 = 0$$

#### 2.1.5 Probability Math

Up until now, we have only considered one event, Y. Now, suppose we have another event that we are interested in, Z.

Let's consider the possibility of either of these two events, Y or Z, occurring. We'd write this as  $Y \cup Z$ , which is mathematical notation for "Y or Z occurs". There are two scenarios that arise:

2.1. PROBABILITY

- 1. Y and Z cannot occur together: they are \_\_\_\_\_
- 2. Y and Z can occur together.

In scenario #1, computing the probability of either Y or Z happening is easy: we just add their respective probabilities together:

$$Y, Z$$
 mutually exclusive  $\Rightarrow P(Y \cup Z) = P(Y) + P(Z)$ 

In scenario #2, computing the probability of either Y or Z happening is more complicated because we know there is a chance that Y and Z can happen together. We'd write this as  $Y \cap Z$ , which is mathematical notation for "Y and Z occurs". In scenario #1, this event never occurred, so  $P(Y \cap Z) = 0$  there. To compute the probability of Y or Z occurring in scenario #2, we have to consider the probability of Y, the probability of Y, and the probability of  $Y \cap Z$ . If we just add P(Y) + P(Z) as in scenario #1, we include the event  $Y \cap Z$  twice, so we have to subtract one instance of it:

$$Y, Z$$
 not mutually exclusive  $\Rightarrow P(Y \cup Z) = P(Y) + P(Z) - P(Y \cap Z)$ .

This probability is much easier to think about when illustrated. In Figure 2.2, we consider human blood types. There are four groups: A, B, O, and AB, and there are two RH types: + and -. We first consider the blood types A and B, represented by the two non-overlapping circles. Define:

- Event Y = a person has blood type A
- Event Z = a person has blood type B
- Event  $Y \cup Z =$  a person has blood type A or blood type B

These two events are *mutually exclusive* because one person cannot have both blood type A and blood type B. (The circles don't overlap in the venn diagram) So, the probability that a randomly selected person has blood type A or B is:

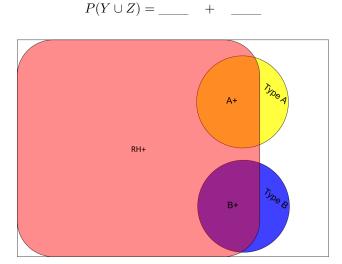


Figure 2.2: Probabilities of blood types in humans. Areas are approximate.

Return to Figure 2.2 and consider two other events: a person having blood type A or having the Rh factor (RH+). We see in Figure 2.2 that someone can have both type A blood and the Rh factor (blood type A+). Define:

- Event Y = a person has blood type A
- Event Z = a person has the Rh factor

- Event  $Y \cup Z =$  a person has blood type A or the Rh factor
- Event  $Y \cap Z$  = a person has blood type A and the Rh factor (they have A+ blood)

So, the probability that someone has either type A blood or has the Rh factor is the sum of probability of having type A blood (represented by the yellow circle) and the probability of having the Rh factor (represented by the red rectangle) minus the probability of having A+ blood (represented by the orange area of overlap that is counted twice) in Figure 2.2. So, the probability that a randomly selected person has blood type A or the Rh factor is:

$$P(Y \cup Z) = \underline{\qquad} + \underline{\qquad} - \underline{\qquad}$$

#### 2.1.6 Conditional Probability

Let's consider an event of interest Y which has probability P(Y). Then, suppose we learn of another event of interest Z that has occurred. Knowing that Z has occurred already may change our opinion about the likelihood of \_\_\_\_\_\_ occurring. The key idea here is that the probability of an event often depends on other information, leading us to the definition of *conditional probability*:

$$P(Y|Z)$$
,

which is the conditional \_\_\_\_\_\_ that Y occurs given that we know Z has occurred. Return to Figure 2.2. Suppose we want to know the probability of a person having type A blood, represented by the yellow circle. But, if we already know that a person has the Rh factor, we are only interested in the part of the type A circle that overlaps with the Rh+ rectangle. Thus the probability of having type A blood is different with different knowledge. The formula for calculating conditional probability is:

$$P(Y|Z) = \frac{P(Y \cap Z)}{P(Z)} \tag{2.2}$$

Returning to the venn diagram, the value  $P(Y \cap Z)$  is represented by the overlap of the type A circle and the Rh+ rectangle, and the value P(Z) is represented by the Rh+ rectangle. Then, the value P(Y|Z) is the ratio of the overlap (A+) to the Rh+ rectangle.

Equation 2.2 also gives us a multiplication rule for computing probabilities:

$$P(Y \cap Z) = P(Y|Z) \cdot P(Z) \tag{2.3}$$

Philosophically speaking, it can be helpful to think of *all* probabilities as conditional. It is just a question of what information is assumed to be \_\_\_\_\_\_.

#### **2.1.6.1** Examples

#### **Death Penalty Convictions**

A study of sentencing of 362 black people convicted of murder in Georgia in the 1980s found that 59 were sentenced to death (Baldus, Pulaski, and Woodworth (1983)). They also examined the race of the murder victim, either black or white, and found some disparities. In Table 2.1, DP means the defendant received the death penalty, NDP means the defendant did not receive the death penalty. The race of the victim (RV) is either black (B) or white (W).

Returning to Section 2.0.1, let's define the problem:

- Population: All black people convicted of murder in Georgia in the 1980s
- Sample: N/A (the whole population was studied)

2.1. PROBABILITY 13

RV	DP	NDP	Total
W	45	85	130
В	14	218	232
Total	59	303	362

Table 2.1: The results of the Baldus et al study for black defendants convicted of murder.

Using the numbers from Table 2.1, compute the following probabilities:

- P(DP) = -- = 0.\_\_\_\_
- P(DP|RV = W) = --- = 0.\_\_\_\_\_
- P(DP|RV = B) = --- = 0.\_\_\_\_

Note: These numbers are selected from the study, and should not be considered a comprehensive summary of its results. There are a number of things not discussed here. The entire publication can be found online<sup>1</sup>

#### Consecutive Matching Striae

In firearms and toolmark analysis, the number of consecutive matching striae (CMS) between a crime scene sample and a lab sample is often used to help determine a match. Generally speaking, the higher the maximum number of CMS found in a pair, the more likely the two samples came from the same source. Several known match (KM) pairs and known non-match (KNM) pairs of bullets were examined, and the results are shown in Figure 2.3 (Hare, Hofmann, and Carriquiry (2017)). What is the probability of seeing two known matches (or two known non-matches) given the maximum number of CMS? Here, we condition on \_\_\_\_\_\_\_\_\_\_. Again, we briefly return to Section 2.0.1, let's define the problem:

- Population: All pairs of fired bullets from unknown sources
- Sample: A sample of pairs of known matches and known non-matches

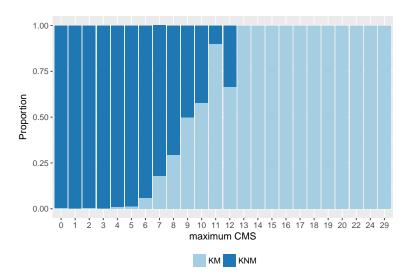


Figure 2.3: This bar chart represents the conditional probabilities of two bullets matching given the maximum number of CMS. The light blue represents known matches, while the dark blue represents known non-matches.

Generally, as seen in Figure 2.3, the probability of finding a match tends to increase with then number of maximum CMS. For \_\_\_\_\_ maximum CMS values is it much more likely that we have a pair.

 $<sup>^{1}</sup> http://scholarlycommons.law.northwestern.edu/cgi/viewcontent.cgi?article=6378\&context=jclc.$ 

#### 2.1.7 Independence

If the likelihood of one event is *not* affected by knowing whether a second has occured, then the two events are said to be \_\_\_\_\_\_\_. For example, the region of the country where you live and what color car you drive are (probably) not related.

The death penalty example from the previous section demonstrates that defendants receiving the death penalty is *not* independent of the race of the victim. In other words, a black defendant found guilty of murder in Georgia in the 1980s received a different penalty depending on the race of the victim.

Another example from DNA analysis relies on on independence across chromosomes. By using loci on different chromosomes, there is independence between the allele counts, allowing for simple calculation of random match probabilities.

#### 2.1.8 Probability Math...Again

Recall Equation 2.3, which gives us the probability of two events, Y and Z occurring together:

$$P(Y \cap Z) = P(Z) \cdot P(Y|Z) = P(Y) \cdot P(Z|Y)$$

If Y and Z are *independent*, there is a simple formula:

$$P(Y \cap Z) = \underline{\hspace{1cm}} \cdot \underline{\hspace{1cm}}$$

This is because Z occurring does not effect the probability of Y occurring, and vice versa. Thus,

$$P(Y|Z) = P(Y)$$
 and  $P(Z|Y) = P(Z)$ 

For example, the probability of being left-handed and from Florida is equal to the probability of being left-handed times the probability of being from Florida, assuming the events "being left-handed" and "being from Florida" are independent.

Multiplying probabilities of events directly like this is *only* applicable when the events are independent. When *dependent* events are treated as independent events, things can go terribly wrong. An infamous example of this in the courts is the case *People v. Collins*<sup>2</sup>. This was a robbery trial, where eyewitnesses described the robbers a "black male with a beard and a moustache, and a white female with a blonde ponytail, fleeing in a yellow car".

The prosecution provided estimated probabilities of each of these individual characteristic:

- P(black man with a beard) = \_\_\_\_\_
- P(black man with a moustache) = \_\_\_\_\_
- P(white woman with ponytail) = \_\_\_\_\_
- P(white woman with blonde hair) =
- P(yellow car) =
- P(interratial couple in a car) =

A mathematics "expert" talked about the so-called "multiplication rule for probability", and directly multiplied the above probabilities together without considering that the events could be *dependent*. i.e. a man with a beard probably has a much higher chance of having a moustache than a man with no beard. Due to this faulty math, the conviction was set aside and the statistical reasoning criticized for ignoring dependence among the characteristics.

 $<sup>^{2}</sup>People\ v.\ Collins,\ 68\ Cal.2d\ 319,\ 438\ P.2d\ 33\ (1968)$ 

2.1. PROBABILITY

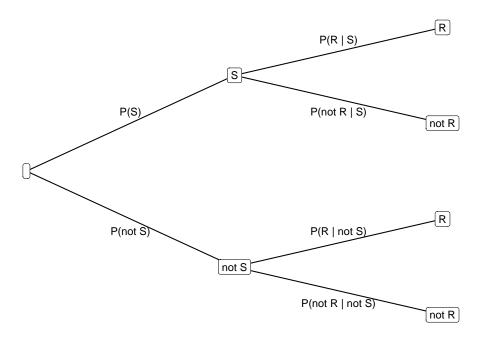


Figure 2.4: A probability tree showing the direction of flow when updating probabilities. Move from left to right on the tree through the events possible. Events are in boxes, probabilities are on the branches of the tree.

In a courtroom situation, let S be the event that the suspect was present at the scene of the crime and  $\overline{S}$  be the event that the suspect was not present at the scene. Assume that each juror has in mind an initial probability for the events S and  $\overline{S}$ . Then, a witness says they saw a tall Caucasian male running from the scene, and the defendant is a tall Caucasian male. After hearing the witness' testimony, the jurors their probabilities. Next, an expert witness testifies that fragments from a window broken during the crime and fragments found on the defendant's clothing match. Again, the jurors update their \_\_\_\_\_\_\_. This process continues throughout the trial. There are some key questions to consider:

- How should jurors update their probabilities?
- Do jurors actually think this way?

#### 2.1.9 Bayes' Rule

Bayes' Rule provides an \_\_\_\_\_\_ formula for probabilities. Like in the trial scenario above, suppose we have an initial estimate for the probability of event S, P(S). Then, we learn that an event R has occurred and we want to update or probability of event S. To do this, we need to know about the \_\_\_\_\_ of R and S. To update the probability of S, we can use Bayes' Rule, also called Bayes' \_\_\_\_\_:

$$P(S|R) = \frac{P(R \cap S)}{P(R)} = \frac{P(R|S)P(S)}{P(R)}$$

$$= \frac{P(R|S)P(S)}{P(R|S)P(S) + P(R|\overline{S})P(\overline{S})}$$
(2.4)

#### 2.1.9.1 Examples

Consider performing diagnostic tests for gunshot residue.

- Let G denote the presence of gunshot residue
- Let  $\overline{G}$  denote the \_\_\_\_\_ of gunshot residue
- ullet Let T denote a \_\_\_\_\_\_ diagnostic test
- Let  $\overline{T}$  denote a negative diagnostic test

Truth	T	$\overline{T}$
G	True Positive	False Negative
$\overline{G}$	False Positive	True Negative

Table 2.2: All potential outcomes of a diagnostic test for gunshot residue.

The values in the table can also be thought of as conditional probabilities:

- The value P(T|G) is the \_\_\_\_\_ rate, also called *sensitivity* of the test
- The value  $P(\overline{T}|\overline{G})$  is the \_\_\_\_\_ rate, also called the *specificity* of the test
- The value  $P(T|\overline{G})$  is the \_\_\_\_\_ rate, the Type I error rate
- The value  $P(\overline{T}|G)$  is the \_\_\_\_\_\_ rate, the Type II error rate

Studies of the diagnostics test usually tell us P(T|G), \_\_\_\_\_\_, and  $P(\overline{T}|\overline{G})$ , \_\_\_\_\_\_ of gunshot residue in a similar situation. What is most relevent for the case is the *postitive predictive value*, or in probability notation, \_\_\_\_\_\_ to obtain this value:

$$P(G|T) = \frac{P(T|G)P(G)}{P(T|G)P(G) + P(T|\overline{G})P(\overline{G})}$$

Generally speaking, the most important thing to remember is that, in general, P(T|G) \_\_\_\_\_ P(G|T).

The careful application of Bayes' Rule can sometimes lead to surprising, non-intuitive results. Continuing with the gunshot residue test example, assume

- sensitivity is 98% (P( | ) = 0.98)
- specificity is 96% (P( | ) = 0.96)
- prevalence is  $90\% (P(\ ) = 0.90)$
- Plug values into the Bayes' Rule formula to find P(G|T):

$$P(G|T) = \frac{P(T|G)P(G)}{P(T|G)P(G) + P(T|\overline{G})P(\overline{G})}$$

$$= \frac{0.98 \cdot 0.9}{0.98 \cdot 0.9 + (1 - 0.96) \cdot (1 - 0.9)}$$

$$= \frac{0.882}{0.882 + 0.004}$$

$$= 0.995$$
(2.5)

• Now assume prevalence is 10% (P(-)=0.10) and plug in the values again

2.1. PROBABILITY 17

$$P(G|T) = \frac{P(T|G)P(G)}{P(T|G)P(G) + P(T|\overline{G})P(\overline{G})}$$

$$= \frac{0.98 \cdot 0.1}{0.98 \cdot 0.1 + (1 - 0.96) \cdot (1 - 0.1)}$$

$$= \frac{0.098}{0.098 + 0.036}$$

$$= \frac{0.098}{0.134}$$

$$= 0.731$$
(2.6)

- So, even if there is a postive test, we are not really sure about whether gunshot residue is *actually* present.
- Why does this happen?? See Figure 2.5.

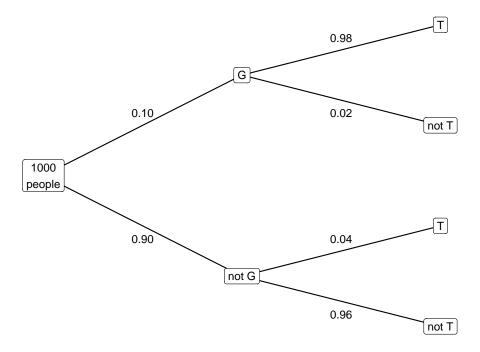


Figure 2.5: A probability tree showing the direction of flow when updating probabilities for the presence of gunshot residue. Suppose there are 1,000 people in the population you're considering. Write the number of people in the groups throughout the tree according to the probabilities indicated on the branches of the tree

#### 2.1.10 Bayes' Rule to the Likelihood Ratio

In the general forensic setting, let S denote the event that the evidence from the scene and comparison sample are from the same source. Let E denote the evidence found at the scene. The formulation of Bayes' Rule for this situation is:

$$P(S|E) = \frac{P(E|S)P(S)}{P(E|S)P(S) + P(E|\overline{S})P(\overline{S})}$$

We can rewrite Bayes' Rule in terms of odds:

$$\frac{P(S|E)}{P(\overline{S}|E)} = \frac{P(E|S)}{P(E|\overline{S})} \frac{P(S)}{P(\overline{S})}$$
(2.7)

Derivation of Equation 2.7 is shown in Equation 2.8. For now, just consider Equation 2.7:

- On the left,  $\frac{P(S|E)}{P(\overline{S}|E)}$  are the odds in favor of S given the evidence E.
- The last term on the right,  $\frac{P(S)}{P(\overline{S})}$  are the odds in favor of S before seeing the evidence E (the "prior odds")
- The first term on the right  $\frac{P(E|S)}{P(E|\overline{S})}$ , is known as the \_\_\_\_\_ ratio
- The likelihood ratio (LR) is the factor by which we \_\_\_\_\_ prior odds of two samples being from the same source to get \_\_\_\_\_ odds (after seeing evidence) of the same source.

$$P(S|E) = \frac{P(E|S)P(S)}{P(E|S)P(S) + P(E|\overline{S})P(\overline{S})}$$

$$\Rightarrow \frac{1}{P(S|E)} = \frac{P(E|S)P(S) + P(E|\overline{S})P(\overline{S})}{P(E|S)P(S)}$$

$$= 1 + \frac{P(E|\overline{S})P(\overline{S})}{P(E|S)P(S)}$$

$$\Rightarrow \frac{1}{P(S|E)} - 1 = \frac{P(E|\overline{S})P(\overline{S})}{P(E|S)P(S)}$$

$$\frac{1}{P(S|E)} - \frac{P(S|E)}{P(S|E)} =$$

$$\frac{1 - P(S|E)}{P(S|E)} =$$

$$\frac{P(\overline{S}|E)}{P(S|E)} = \frac{P(E|\overline{S})P(\overline{S})}{P(E|S)P(S)}$$

$$\Rightarrow \frac{P(S|E)}{P(\overline{S}|E)} = \frac{P(E|S)P(S)}{P(E|S)P(S)}$$

$$\Rightarrow \frac{P(S|E)}{P(\overline{S}|E)} = \frac{P(E|S)P(S)}{P(E|\overline{S})P(\overline{S})}$$

#### 2.1.10.1 Examples

Return to the gunshot residue (GSR) test example. Define:

- E = evidence = a positive test for (GSR)
- S = suspect has GSR on them

$$LR = \frac{P(E|S)}{P(E|\overline{S})} = \frac{0.98}{0.04} = 24.5$$

In a high prevalence case (P(G) = 0.9), the prior odds are  $\frac{0.9}{0.1} = 9$ . The posterior odds are  $LR \times \text{prior}$  odds  $= 24.5 \times 9 = 220.5 : 1$ .

In a low prevalence case (P(G) = 0.1), the prior odds are  $\frac{0.1}{0.9} = \frac{1}{9}$ . The posterior odds are  $LR \times \text{prior}$  odds  $= 24.5 \times \frac{1}{9} = 24.5 : 9 = 2.72 : 1$ .

We can also compute the likelihood ratio if the evidence were a negative test. This value turns out to be  $\frac{1}{48}$ , which is **not** the reciprocal of the LR for the positive test.

#### 2.1.11 Recap

•	Probability is the language of
•	Provides a common scale, from to, for describing the chance that an event will occur
	Conditional probability is a key concept! The probabilitity of an event depends on what is available
•	Independent events can be powerful! They allow us to probabilities of events directly, as is common in
•	is a mathematical result showing how we should
	our probabilities when available information changes.
	<ul> <li>This will later lead us to the likelihood ratio as a numerical of the evidence.</li> </ul>
	- Bayes' Rule does not necessarily describe how people operate in practice.

#### 2.1.12 Probability and the Courts

Sally Clark was the only person in the house when her first child died unexpectedly at 3 months old. The cause of death was determined to be SIDS, sudden infant death syndrome. One year later, Sally and her husband had a second child, who died at 2 months old under similar circumstances. Sally was convicted of murder.

During her trial, a pediatrician testified that the probability of a single SIDS death for a family like the Clarks (similar income, etc.) was  $\frac{1}{8500} \approx 0.0001$ , and thus the probability of two SIDS death in the family was  $\frac{1}{8500^2} = \frac{1}{73 \times 10^6} \approx 1.37 \times 10^{-8}$ . There are several problems with this approach to evidence. What do you think? Jot down a few ideas below:

Issues with the evidence presented by the pediatrician:

- 1. Is the probability of a child dying of SIDS given,  $\frac{1}{8500}$ , correct for "families like the Clarks"?
- 2. The use of direct multiplication of probabilities assumes independence of the two deaths in the family. (Independence within the family is not a reasonable assumption.)
- 3. Alternative hypotheses (causes of death of the infants) were not considered. Did something else with perhaps a higher likelihood cause the children's deaths?

### 2.2 Probability to Statistical Inference

Probability is important, but it is only one tool in our toolbox. Another, more powerful tool is statistical inference.

#### 2.2.1 Collecting Data

First, we consider data collection. Where do data come from? One data source is an *experiment*. An investigator designs a study and collects information on and maybe applies treatments to a *sample*, a subset of

the population of interest. or choose a		s a great deal abour	t how to design an
The area of statistics conceliterature is extensive (see			ental design. The experimental design crucial points:
• The goal of an exper	iment is to compare		
• Those	must be	assigned	to units
conclusions - Blinding plays an in	aportant role in avoiding	e.g.	enough t obe able to make informed "double-blind" studies in medicine, ent know which treatment the patient
How is experimental designation	n relevant to forensic scien	nce?	
_	d to evaluate process impublished box" studies, where		know ground truth
Experiments almost alway	s involve sampling from the	he population of in	terest. Why?
• We sample because i	t is too or	ſ	to study the <i>entire</i> population
• A sa	ample allows us to use the	laws of	to describe how certain we
	answer reflections answer reflections failures (cautionary ta		sampling. (See
How is sampling relevant t	to forensic science?		
• Sampling techniques crime scene to test.	used to determine which	and how many bag	gs of suspect powder collected from a
All data collected can be d	livided into one of two gro	oups: qualitative or	quantitative.
	escribe qualities about the on. There are two subcate		or example, the race of a suspect, or we data:
	type (A, B, AB, or O)	to one of a discrete	e number of groups or categories. For
	of a teacher (poor, average		a set of ordered values. For example, categories have an inherent ordering,
	describe quantities that ca two subcategories of quan		the observations. These are numerical
	ions: $\{0, 1, 2, 3, 4,\}$ . A olmarks. (See Figure 2.3)	forensic science ex	te. An easy-to-understand example is cample is consecutive matching striae value in a finite or infinite interval.
Continuous valuindex of a glass	ies fall anywhere on the n		nsic science example is the refractive
2.2.2 Probability I	Distributions		
Suppose we are to collect delement concentration). A			dividuals or objects (e.g. weight, trace escribe these possible values and how



Figure 2.6: This picture from the US presidential election of 1948 shows President Harry Truman, who won the election, holding a newspaper that went to print with the headline "Dewey Defeats Truman!" The headline was based on biased sampling that favored typically Republican demographics. Image Source: https://blogs.loc.gov/loc/2012/11/stop-the-presses/

#### 2.2.2.1 Normal

You may already be familiar with this distribution with the bell-shaped curve. Measurement error is one example of something often assumed to follow a normal distribution. The normal distribution is described by two parameters: the \_\_\_\_\_\_, denoted by  $\mu$ , and the \_\_\_\_\_\_, denoted by  $\sigma$ . If we have a variable (say, something observed in our data like weight), we give the variable a capital letter, typically X. If this variable X is normally distributed with mean  $\mu$  and standard deviation  $\sigma$ , we write this as:

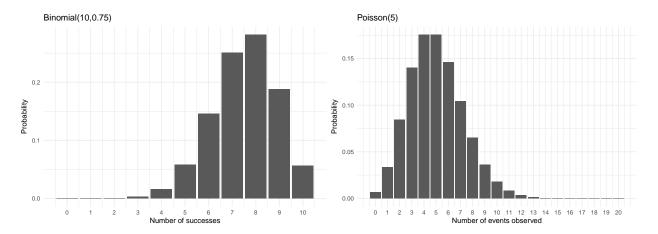


Figure 2.7: On the left, the probability of each possible outcome for variable with binomial distribution with 10 trials and probability of success 0.75. On the right, the probability of each possible outcome for variable with Poisson distribution with mean value 5.

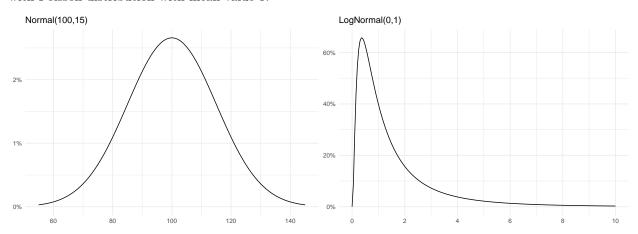


Figure 2.8: On the left, the probability distribution curve of possible outcomes for a variable with Normal distribution with mean value 100 and standard deviation 15. On the right, the probability distribution curve of possible outcomes for a variable with Lognormal distribution with mean value 0 and standard deviation 1.

$$X \sim N($$
, )

In measurement error, for example, we typically assume that the mean is 0. So, if X represents measurement error, we'd write  $X \sim N(0, \sigma)$ .

There are many nice properties of the normal distribution. For instance, we know that \_\_\_\_\_\_% of observable values lie within \_\_\_\_\_\_ standard deviations of the mean  $(\mu \pm 2\sigma)$ , and also that \_\_\_\_\_% of observable values lie within \_\_\_\_\_ standard deviations of the mean  $(\mu \pm 3\sigma)$ . When working with the normal distribution, we use software (such as Excel, Matlab, R, SAS, etc.), tables<sup>3</sup>, or websites like onlinestatbook.com or stattrek.com to compute probabilities of events.

#### 2.2.2.2 Lognormal

We often act as if everything is normally distributed, but of course this is not true. For instance, a quantity that is certain to be \_\_\_\_\_\_ (greater than or equal to zero) cannot possible be normally distributed.

<sup>&</sup>lt;sup>3</sup>See for example http://www.stat.ufl.edu/~athienit/Tables/Ztable.pdf

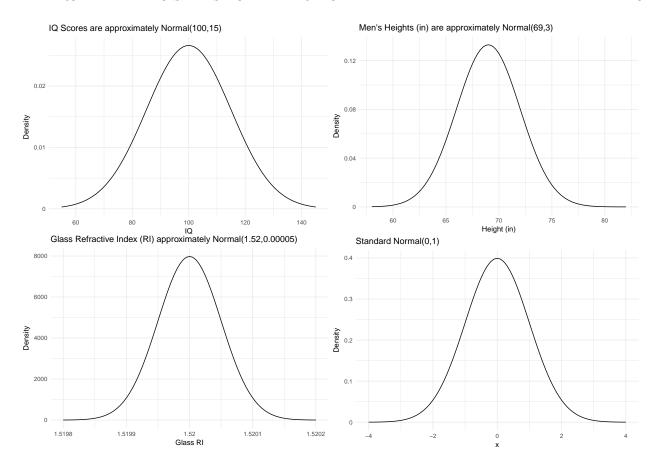


Figure 2.9: Four examples of the probability distribution functions for normally distributed variables

Consider trace element concetration: either none is detected, or there is some amount greater than 0 detected. In cases where nonnegative values are not possible, we may believe that the (natural) \_\_\_\_\_\_ of the quantity is normal, which gives us a \_\_\_\_\_\_ distribution for the quantity itself. The lognormal distribution, like the normal, has two parameters: mean (on the log scale), denoted \_\_\_\_\_, and standard deviation (on the log scale), denoted \_\_\_\_\_.

#### 2.2.2.3 Discrete

Some quantities take on very few possible values. These are  $\it discrete$  data.

Recall the two common discrete distributions from section 2.2.2:

- Binomial:
  - Data are \_\_\_\_\_ (two categories: "success" or "failure")
  - Data are a result of n independent
  - -P(success) = p on each trial. (Same \_\_\_\_\_\_ of success each time)
  - Expected number of successes you expect to see out of n trials:  $\underline{\phantom{a}} \times \underline{\phantom{a}}$
  - Example: Suspect a student of cheating on an exam, response is the number of correct answers.
- Poisson:
  - Data are counts: number of events occurring in a \_\_\_\_\_ time

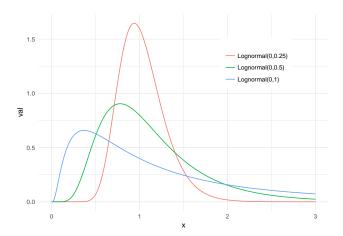


Figure 2.10: Three lognormal distributions with the same mean (on the log scale) and different standard deviations (on the log scale)

- The mean and the \_\_\_\_\_\_ of this distribution are the same, so the variability in responses increases as the \_\_\_\_\_ increases.
- Example: number of calls to 911 between 10:00 and midnight on Friday nights. See Figure 2.11 for a forensics example.

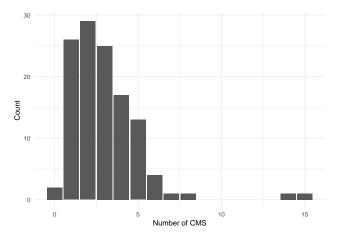


Figure 2.11: Distribution of the maximum number of CMS for a randomly selected bullet compared to 118 known lands approximately follows a Poisson distribution.

### 2.3 Statistical Inference - Estimation

Recall from Section 2.0.1:

- The \_\_\_\_\_ is the universe of objects of interest.
- The \_\_\_\_\_ is comprised of the objects available for study.
- \_\_\_\_\_ is deductive: use knowledge about the population to make statements describing the sample
- \_\_\_\_\_ is inductive: use knowledge about the sample to make statements describing the population

• Probability and statistics are used together!
1. Build or assume a for a population
2. Assess the using the model
3. Refine the model, return to step 2.
2.3.1 Background
A is a numerical characteristic of the population, e.g. the population mean. Statistical methods are usually concerned with learning about population parameters from
Note: The mean of a <i>sample</i> and the mean of a <i>population</i> are differenct concepts. The mean of a sample can be calculated exactly, while the mean of a population is (usually) unknown, because there are too many objects in the population to record and calculate the mean.
The idea underlying statistical inference is that we can apply laws of probability to drawabout a population from a sample. This process is briefly summarized below:
• Observe mean
• If we have a "good" sample this sample mean should be close to the mean.
• The laws of tell us how close we can expect them to be.
For example, suppose we are interested in the average height of the adult population in the U.S.
• Population:
• Sample:
<ul> <li>We can take the average height of everyone here and use this sample to make about the mean of all U.S. adults.</li> <li>Note: This approach will work if our sample is a sample from the population. This assumption may be questionable, so it should be verified.</li> </ul>
The goal of statistical inference is about a Different possible parameters are:
<ul><li>Mean</li><li>Variance</li><li>Proportion</li></ul>
We can also make different types of inferential statements, depending on what question we are trying to answer and how we are going to report our results. We will talk about:
• estimate: an estimate of a parameter value
• estimate: a range of plausible values for a parameter
• Hypothesis: examine a specific hypothesis about the true value of a parameter
When you want to do statistical inference, it is always inportant to look at your sample data before proceeding directly to inference. We do this because we want to
1. See general in the data
2. Get an idea of the of the distribution of the data
3 Identify values and/or errors

How we look at our data to check for these three things? If our data are \_\_\_\_ table of frequencies or a bar chart of the different outcomes. If our data are histograms of the values, or numerical summaries such as mean, median, standard deviation, or percentiles.

A quick example shows why it can be important to examine your data before a formal statistical analysis:

- Suppose the data are (19,20,21,22,23). Then, the mean is  $\frac{19+20+21+22+23}{5}=21$ , the median is 21, and the standard deviation is 1.58.
- But what if the data you receive are (19,20,21,22,93)? Then, the mean is  $\frac{19+20+21+22+93}{5} = 35$ , the median is 21, and the standard deviation is 32.4.
- There could have been a typo, or someone interpreted some handwriting wrong, etc. The moral of the story is ALWAYS look at your data first!

#### 2.3.2 **Point Estimation**

	is a rule for estime estimator by considering	nating a population g two key properties:	_ from a sample.	We evaluate the
• Bias: h	now close	an estimator is to the tr	ue population me	ean
• Variabi	ility: how	_ is the estimate?		
For the popusuch as:		is There are other		

- The median (good for skewed data or data with outliers)
- The midrange (max + min/2)
   47 (obviously this is just guessing and is not advised)

Let  $\theta$  denote an unknown population parameter that we wish to estimate. The letter  $\theta$  represents the true value of the parameter. In Figure 2.12, we see what would happen in many repeated attempts to estimate  $\theta$ using estimators with different properties.

#### 2.3.3 Standard Errors

One limitation of just providing a point estimate is that it doesn't give us any indication of
As we saw in Figure 2.12, a point estimate alone can be very different
from the true mean. We can do better than this!
The of an estimator measures the uncertainty in our esti-
mate. When looking at a summary statistic, like mean, median, or percentiles, that statistic is also a
quantity. This means that if we had observed a different set of sample values
we would observe different values of the summary statistics. The idea of standard error is similar to the idea
of standard deviation. Both are measures of spread, or variability. The difference is that standard deviation
is a measure of variability of a sample or population, while standard error is a measure of variability of an

Consider a population with that is normally distributed with mean 100 and standard deviation 15. Recall from Section 2.2.2.1 that 95% of observations from a normal distribution fall within two standard deviations of the mean. So, in this example we expect 95% of observations to be between 70 and 130. This distribution is shown on left in Figure 2.13. Now suppose we want to look at the distribution of the estimates of the population mean from several samples. This will demonstrate the idea of standard error, using sample size of n=25. The formula to compute standard error is:



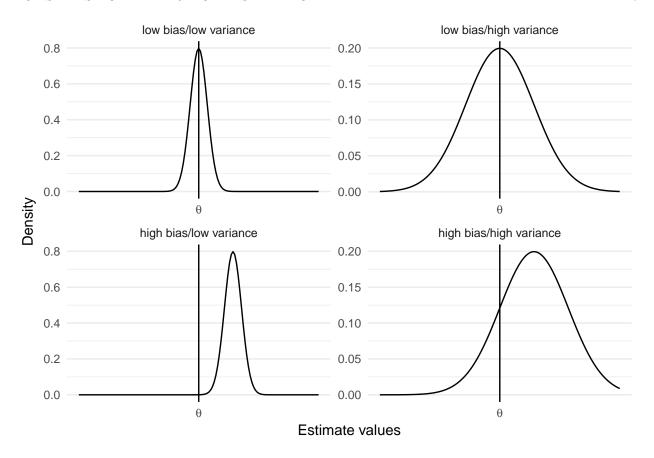


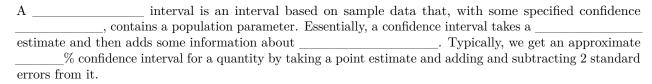
Figure 2.12: The curve in each plot shows the distribution of estimates we would see under each condition shown.

The standard error for this example is  $\frac{15}{\sqrt{25}} = 3$ . The mean of a sample of size 25 from this population should be about \_\_\_\_\_\_. And about 95% of the time, the sample mean will be between \_\_\_\_\_ and \_\_\_\_\_. This distribution is shown on right in Figure 2.13.

#### 2.3.4 Sample Size

The size of a sample plays a *critical* role in determining how accurate we can be. Again, consider a population with distribution N(100, 15). We can use \_\_\_\_\_\_ to examine the effect of sample size. We simulate samples from a normal distribution with mean 100 and standard deviation 15. We use four different sample sizes: 10, 25, 50, and 100. We take 500 samples of each size and compute the mean for each sample, leaving us with 2,000 means that we have calculated. We show histograms of the means of these samples for each sample size in Figure 2.14.

#### 2.3.5 Interval Estimation



The most well-established procedures for finding confidence intervals are those related to drawing conclusion

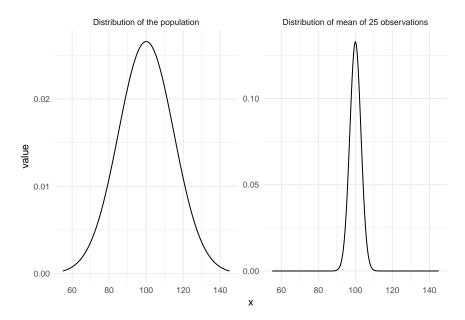


Figure 2.13: On the left, the distribution of the population, which is distributed N(100,15). On the right, the distribution of the sample mean for samples of size 25 from the populations, which is distributed N(100,3).

about the mean of a  $\_$ \_\_\_\_\_ population. Suppose that we have acquired a random sample of n observations from a normal population.

• **Point** estimate: the natural point estimate of the population mean is the sample mean, as we've seen already. The sample mean is often denoted by \_\_\_\_\_\_. To calculate the sample mean, simple add up all the values and divide by the number of values you have, n. In mathematical notation this is written:

$$\overline{X} = \frac{1}{n} \cdot \sum_{i=1}^{n} X_i$$

• Interval estimate: denote the standard error of the sample mean by  $SE(\overline{X})$ , and the standard deviation of the population by SD(population). Then, compute the standard error by:

$$SE(\overline{X}) = \frac{SD(\text{population})}{\sqrt{n}}$$

Then, an approximate 95% confidence interval is computed:

$$\overline{X} \pm 2 \cdot SE(\overline{X}) \tag{2.9}$$

A key result is that these procedures for point and interval estimation work well even if the population does **not** follow a \_\_\_\_\_\_ distribution **as long as the sample is** \_\_\_\_\_\_.

For example, suppose there are 10 glass fragments found at a crime scene, and the concentration of aluminum in each one is measured. The mean aluminum concentration of the sample was 0.73 and the standard deviation was 0.04. The standard error is thus:

$$SE(\overline{X}) = \frac{}{\sqrt{}} = 0.013$$

The approximate 95% confidence interval for the mean aluminum concentration in the crime scene window is:

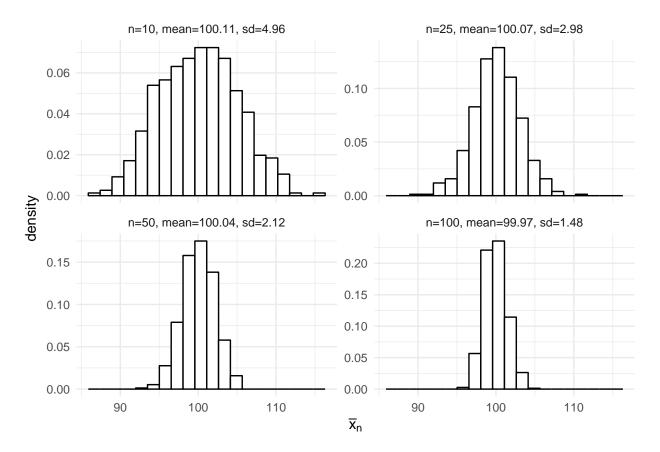


Figure 2.14: How does sample size affect the sampling distribution? As sample size increases, the standard error decreases, so the distribution of sample means becomes more narrow.

0	73 +	1.96	$\cdot 0.013$	= (	

The *interpretation* of the confidence interval is important: 95% of the intervals \_\_\_\_\_\_ in this way will contain the \_\_\_\_\_ population parameter.

### 2.4 Statistical Inference - Hypothesis Testing

Sometimes we wish t	o formally	a hypoth	nesis about a popu	lation parameter.	The hypothesis
to be evaluated is kn	own as the	hypothesis. '	This hypothesis is	usually the status	quo, or what is
assumed to be true.	When hypothesis	testing we look for	or evidence	the null.	There is also a
an	or research hy	pothesis that help	os us design the to	est. If we	the null
hypothesis, then we	say that we have a	a statistically		result.	
As with anything in care about:	life, errors are pos	sible in hypothesis	s testing. There are	re two main typs o	f errors that we
1. Type I:	the nul	l hypothesis when	ı it is	(false positive	e)
2. Type II: negative)		the r	null hypothesis w	when it is	(false

Of these two errors, Type I error is often considered more serious. We only want to null hypothesis is there is strong evidence against it. These statistical testing ideas are closely related to • The null hypothesis: the defendant is • The alternative hypothesis: the defendant is • In court, a Type I error is to find guilty when the defendant is \_\_\_ • In court, a Type II error is to find innocent when the defendant is Ultimately, the basic idea of hypothesis testing is to compute a "distance" between the \_\_\_\_\_ we have collected and what we expect under the \_\_\_\_ hypothesis. Typically, we use a test statistic of the form: point estimate – null hypothesis value (2.10)SE(estimate)This test statistic can be interpreted as the number of \_\_ \_\_\_\_\_ the sample estimate is from the \_\_\_\_\_ value under the null hypothesis. A common way to to summarize hypothesis tests is by attaching a \_\_\_\_\_\_ to the test statistic. This probability is called a \_\_\_\_\_\_. The p-value of a hypothesis test gives the probability that we would get data like that in our sample (or something even more \_\_\_\_\_\_), given our assumption that the null hypothesis is \_\_\_\_\_. This idea is demonstrated in Figure 2.15. Small *p*-values mean that we have observed \_\_\_\_\_ data that lead us to question the hypothesis, which we have assumed to be true. Small p-values tell us that the sample data are unlikely to happen by chance under the null. A p-value, however, only addresses the \_\_\_\_\_\_ hypothesis. It does not speak to the likelihood of the \_\_\_\_\_ hypothesis being true. Fail to reject the null hypothesis Reject the null hypothesis p-value large p-value small 0.4 test statistic density 0.2 p-value=0.212 0.1 p-value=0.024

Figure 2.15: On the left, an example of a p-value that is large (test statistic small), leading us to fail to reject the null hypothesis. On the right, an example of a p-value that is small (test statistic large), leading us to reject the null hypothesis.

#### 2.4.1 Normal Data

Test Statistic

The most well-established hypothesis testing procedure is for testing a hypothesis about the \_\_\_\_\_\_ of a \_\_\_\_\_ population. This population parameter is denoted  $\mu$ . We can test the hypothesis that the population mean,  $\mu$ , is equal to some specified values  $\mu_0$ . The hypotheses for this test are:

- Null hypothesis  $H_0: \mu = \mu_0$
- Alternative hypothesis  $H_A: \mu \neq \mu_0$

The test statistic, call it T, to perform this hypothesis follows the form of Equation 2.10:

$$t = \frac{\phantom{-}}{SE(\underline{\phantom{}})} \tag{2.11}$$

The p-value for this hypothesis test is obtained from a t distribution (see Figure 2.16) using software, a table of values, or an online calculator. These hypothesis testing procedures work well even if the population is not normally distributed as long as the sample size is

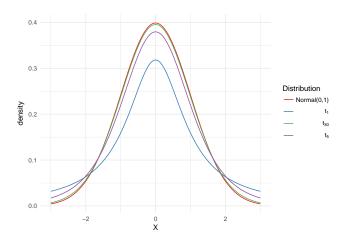


Figure 2.16: The t distribution is similar to the Normal distribution (red line) is shape, but the tails are higher. As the sample size increases, the t distribution approaches the Normal distribution.

#### 2.4.1.1Example

Suppose we want to estimate the mean amount of a trace element for the population of all bullets in Iowa. We get a random sample of 400 bullets from the state.

- The sample mean concentration is  $\overline{X} = 55.5$
- The sample standard deviation is s = 22.0
- The standard error of the mean is  $SE(\overline{X}) = \frac{22}{\sqrt{400}} = \frac{22}{20} =$ \_\_\_\_\_

Suppose we have reason to believe that Remington (mean = 58) is the main producer in this area. We can check this idea with a hypothesis test.

- Null hypothesis  $H_0: \mu = 58$
- Alternative hypothesis  $H_A$ :  $\mu \neq 58$  Test statistic  $t=\frac{\overline{X}-\mu_0}{SE(\overline{X})}=\frac{55.5-58}{1.10}=-2.27$

The value of the test statistic is more than two standard errors away from the mean under the null hypothesis. The exact p-value is 0.023. This means that if the null hypothesis is true, then observing a value 2.27 standard errors or more away from the mean happens only 2.3% of the time. So, we reject our assumption that the population mean concentration is 58.

We can also calculate a 95% confidence interval for the mean using Equation 2.9:

$$55.5 \pm 1.96 \cdot 1.10 = (53.3, 57.7)$$

The hypothesized value (58) is is not in the 95% confidence interval, which also suggests that population mean concentration equal to 58 is not possible.

#### 2.4.2 Confidence Intervals

There is a very close relationship between tests and interval estimates. Recall that a confidence interval (CI) gives a range of plausible values for the true population parameter, which here is the mean. A hypothesis test evaluates whether a specified ( $\mu_0$ ) is a \_\_\_\_\_\_ value for the mean. A CI collects all values of  $\mu_0$  that we would find plausible in a test.

Statistical hypothesis test are *very* popular in practice. Sometimes, they address the scientific question of interest, but often they do not.<sup>4</sup>

#### 2.4.3 Comparing Two Means

In section 2.4.1, we discussed hypothesis testing methods for one sample. In practice, we are often interested in comparing \_\_\_\_\_\_ samples from \_\_\_\_\_ different populations. For now, assume we have random samples from each of the two populations that we are interested in. The test we want to do is a test for \_\_\_\_\_ of parameters in the two populations.

#### 2.4.3.1 Example

Suppose, for example, that we have collected broken glass at a crime scene, and glass fragments on a suspect. Define  $\mu_{scene}$  to be the mean trace element level for population of glass at the scene. Define  $\mu_{suspect}$  to be the mean element level for the population of glass on the suspect. We can compare the means to address the question of whether or not the glass fragments on the suspect could plausibly have come from the crime scene.

Hypotheses:

- $H_0: \mu_{scene}$  \_\_\_\_  $\mu_{suspect}$
- $H_A: \mu_{scene}$  \_\_\_\_  $\mu_{suspect}$

Suppose 10 glass fragments are taken from the glass found at the scene (denote these by Y), and 9 fragments are found on the suspect (denote these by X). Concentrations of a trace element were measured in each fragment of glass. Summary values from the samples are:

- $\overline{X} = 5.3$
- $s_X = 0.9$
- $SE(\overline{X}) = \frac{s_X}{\sqrt{n_X}} = \frac{0.9}{\sqrt{10}} = 0.28$
- $\overline{Y} = 5.9$
- $s_Y = 0.85$
- $SE(\overline{Y}) = \frac{s_Y}{\sqrt{n_Y}} = \frac{0.85}{\sqrt{9}} = 0.28$
- Obeserved difference =  $\overline{X} \overline{Y} = = 0.6$
- The standard error for the difference,  $\overline{X} \overline{Y}$ , is

<sup>&</sup>lt;sup>4</sup>For more reading on this topic, consider this article from *Nature*: http://www.nature.com/news/psychology-journal-bans-p-values-1.17001

• The test statistic for this hypothesis test is:

$$t = \frac{(\overline{X} - \overline{Y}) - 0}{SE(\{\overline{X} - \overline{Y}\})} = \frac{0.6}{0.4} = 1.5$$

- The corresponding p-value for this statistic is 0.15.
- So, we fail to reject the null hypothesis that the two glass population means are equal.
- Interpretation is a *key* issue. When we say we fail to reject the null, we are saying there is a possibility of a common source.

#### 2.4.4 Discussion

There are three key points for you to take away:

1.	Hypothesis testing does <i>not</i> treat the two hypotheses The null hypothesis is
	given priority. This is appropriate when there is reason to the null hypothesis until
	there is significant evidence it. We don't necessarily always want this to be the
	case. (We will discuss this more later on in a forensic context.)
2.	The $p$ -values that result from hypothesis tests depend heavily on the sample size. If you have the same and standard deviation, but the sample size, the result will me more significant, due to the sample size alone.
3.	Interpreting the results of the hypothesis test can be tricky. If we the null hypothesis, this does not necessarily mean that we have found an important difference in the context of our problem. In addition the null hypothesis does not necessarily mean
	the null hypothesis is true.

#### 2.5 Overview of Statistical Preliminaries

- We reviewed the basics of probability
  - Probability is the language of uncertainty
  - It is important to understand what is being assumed when talking about probability
  - For instance, the probability of having disease given a positive test is different than the probability of having a positive test given the disease
  - Probability distributions describe the variability in a population or in a series of measurements
- We reviewed basics of statistical inference
  - Statistical inference uses sample data to draw conclusions about a population
  - Point estimation, interval estimation, and hypothesis tests are main tools
  - It is critical that our procedures account for variation that could be observed due to chance

# Chapter 3

# Statistics for Forensic Science

- 3.1 Brief Review of Probability and Statistics
- 3.2 The Forensic Examination
- 3.3 Common Approaches to Assessing Forensic Evidence
- 3.3.1 Significance Testing / Coincidence Probability
- 3.3.2 Likelihood Ratio

# References

Baldus, David C., Charles Pulaski, and George Woodworth. 1983. "Comparative Review of Death Sentences: An Empirical Study of the Georgia Experience." *Journal of Criminal Law and Criminology*.

Hare, Eric, Heike Hofmann, and Alicia Carriquiry. 2017. "Automatic Matching of Bullet Land Impressions." *The Annals of Applied Statistics* Upcoming.

Morris, Max D. 2011. Design of Experiments: An Introduction Based on Linear Models. Chapman; Hall.