

6.046- *Design and Analysis of Algorithms*

Lecture 09

Markov Chain Monte Carlo

**(supplementary material to these
slides has also been posted)**

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Menu

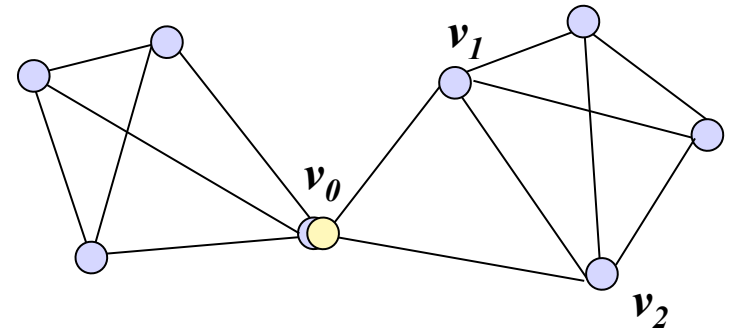
- Random walks on graphs
- Markov Chains
- Examples:
 - pagerank
 - card-shuffling
 - colorings

Menu

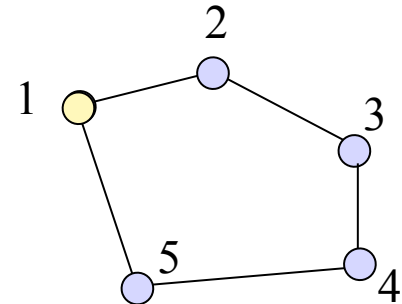
- **Random walks on graphs**
- Markov Chains
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Random Walks

- Given undirected graph $G = (V, E)$
- A squirrel stands at vertex v_0 :
- Squirrel ate fermented pumpkin so doesn't know what he's doing
- So jumps to random neighbor v_1 of v_0
- Then jumps to random neighbor v_2 of v_1
- etc
- Question: Where is squirrel after t steps?
- A: At some random location.
- OK, with what probability is squirrel at each vertex of the graph?
- Want to compute $x_t \in \mathbb{R}^n$, where
- $x_t(i)$: probability squirrel is at node i at time t .
- v_t : random variable representing location at time t .



$$\mathbf{x}_t \rightarrow \mathbf{x}_{t+1} ?$$



- Simplification: all nodes have same degree d , e.g.
- $x_0 = (1, 0, 0, 0, 0)$
- $x_0 \rightarrow x_1 ?$
- if u_1, u_2, \dots, u_d are the d neighbors of v_0 , then
- $v_1 = u_i$ with probability $1/d$
- so $x_1 = (0, 1/2, 0, 0, 1/2)$
- $x_2 = (1/2, 0, 1/4, 1/4, 0)$
- ...

$$A = \begin{pmatrix} 0 & 1/2 & 0 & 0 & 1/2 \\ 1/2 & 0 & 1/2 & 0 & 0 \\ 0 & 1/2 & 0 & 1/2 & 0 \\ 0 & 0 & 1/2 & 0 & 1/2 \\ 1/2 & 0 & 0 & 1/2 & 0 \end{pmatrix}$$

“the transition matrix”

(adjacency matrix divided by d)

A_{ij} : probability of jumping to j if squirrel is at i

formally $A_{ij} = \Pr[v_{t+1}=j \mid v_t = i]$

$$x_1 = x_0 A$$

$$x_2 = x_1 A = x_0 A^2$$

$$x_3 = x_2 A = x_0 A^3$$

\vdots

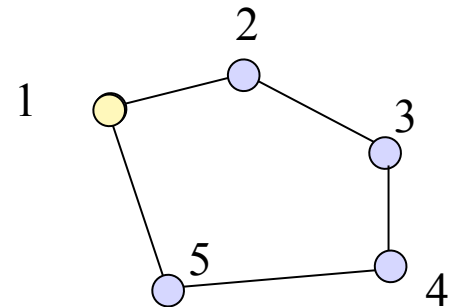
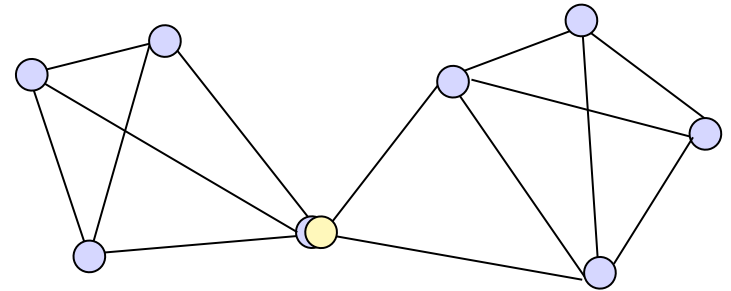
$$x_{t+1} = x_t A = x_0 A^{t+1} \quad (1)$$

x_t

- More general undirected graphs?
- *Transition Matrix:*

A = adjacency matrix where row i is divided by the degree d_i of i

- $x_t = x_0 A^t$
- Computing x_t fast?
 - repeated squaring!
 - compute $A \rightarrow A^2 \rightarrow A^4 \rightarrow \dots \rightarrow A^t$
(if t is a power of 2; if not see lecture 2)
 - then do vector-matrix product
- Limiting distribution x_t as $t \rightarrow \infty$?
- e.g. what is x_∞ in 5-cycle?
- $x_\infty = (1/5, 1/5, 1/5, 1/5, 1/5)$

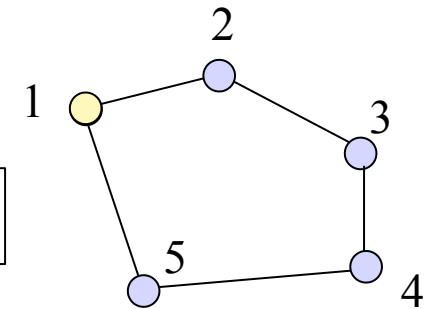


Verifying $x_t \rightarrow (1/5, 1/5, 1/5, 1/5, 1/5)$

- Recall

$$A = \begin{pmatrix} 0 & 1/2 & 0 & 0 & 1/2 \\ 1/2 & 0 & 1/2 & 0 & 0 \\ 0 & 1/2 & 0 & 1/2 & 0 \\ 0 & 0 & 1/2 & 0 & 1/2 \\ 1/2 & 0 & 0 & 1/2 & 0 \end{pmatrix}$$

$$x_t = x_0 A^t$$



• $x_0 = [1 \quad 0 \quad 0 \quad 0 \quad 0]$	$x_{15} = [0.1833 \quad 0.2135 \quad 0.1949 \quad 0.1949 \quad 0.2135]$
• $x_1 = [0 \quad 0.5000 \quad 0 \quad 0 \quad 0.5000]$	$x_{16} = [0.2135 \quad 0.1891 \quad 0.2042 \quad 0.2042 \quad 0.1891]$
• $x_2 = [0.5000 \quad 0 \quad 0.2500 \quad 0.2500 \quad 0]$	$x_{17} = [0.1891 \quad 0.2088 \quad 0.1966 \quad 0.1966 \quad 0.2088]$
• $x_3 = [0 \quad 0.3750 \quad 0.1250 \quad 0.1250 \quad 0.3750]$	$x_{18} = [0.2088 \quad 0.1929 \quad 0.2027 \quad 0.2027 \quad 0.1929]$
• $x_4 = [0.3750 \quad 0.0625 \quad 0.2500 \quad 0.2500 \quad 0.0625]$	$x_{19} = [0.1929 \quad 0.2058 \quad 0.1978 \quad 0.1978 \quad 0.2058]$
• $x_5 = [0.0625 \quad 0.3125 \quad 0.1562 \quad 0.1562 \quad 0.3125]$	$x_{20} = [0.2058 \quad 0.1953 \quad 0.2018 \quad 0.2018 \quad 0.1953]$
• $x_6 = [0.3125 \quad 0.1094 \quad 0.2344 \quad 0.2344 \quad 0.1094]$	$x_{21} = [0.1953 \quad 0.2038 \quad 0.1986 \quad 0.1986 \quad 0.2038]$
• $x_7 = [0.1094 \quad 0.2734 \quad 0.1719 \quad 0.1719 \quad 0.2734]$	$x_{22} = [0.2038 \quad 0.1969 \quad 0.2012 \quad 0.2012 \quad 0.1969]$
• $x_8 = [0.2734 \quad 0.1406 \quad 0.2227 \quad 0.2227 \quad 0.1406]$	$x_{23} = [0.1969 \quad 0.2025 \quad 0.1991 \quad 0.1991 \quad 0.2025]$
• $x_9 = [0.1406 \quad 0.2480 \quad 0.1816 \quad 0.1816 \quad 0.2480]$	$x_{24} = [0.2025 \quad 0.1980 \quad 0.2008 \quad 0.2008 \quad 0.1980]$
• $x_{10} = [0.2480 \quad 0.1611 \quad 0.2148 \quad 0.2148 \quad 0.1611]$	$x_{25} = [0.1980 \quad 0.2016 \quad 0.1994 \quad 0.1994 \quad 0.2016]$
• $x_{11} = [0.1611 \quad 0.2314 \quad 0.1880 \quad 0.1880 \quad 0.2314]$	
• $x_{12} = [0.2314 \quad 0.1746 \quad 0.2097 \quad 0.2097 \quad 0.1746]$	
• $x_{13} = [0.1746 \quad 0.2206 \quad 0.1921 \quad 0.1921 \quad 0.2206]$	
• $x_{14} = [0.2206 \quad 0.1833 \quad 0.2064 \quad 0.2064 \quad 0.1833]$	

Example 2: 4-cycle

- $x_t \rightarrow ??$
- in this case: $A = \begin{pmatrix} 0 & \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} & 0 \\ 0 & \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} & 0 \end{pmatrix}$

- Let's iterate $x_t = x_0 A^t$

- $x_0 = [1 \quad 0 \quad 0 \quad 0]$
- $x_1 = [0 \quad \frac{1}{2} \quad 0 \quad \frac{1}{2}]$
- $x_2 = [\frac{1}{2} \quad 0 \quad \frac{1}{2} \quad 0]$
- $x_3 = [0 \quad \frac{1}{2} \quad 0 \quad \frac{1}{2}]$
- $x_4 = [\frac{1}{2} \quad 0 \quad \frac{1}{2} \quad 0]$
- $x_5 = [0 \quad \frac{1}{2} \quad 0 \quad \frac{1}{2}]$
- $x_6 = [\frac{1}{2} \quad 0 \quad \frac{1}{2} \quad 0]$

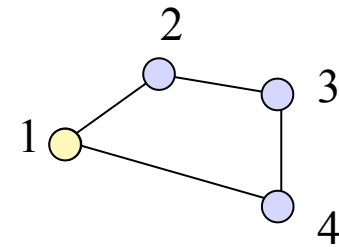
• ...

- Ouch:

$$x_{2t} \rightarrow [\frac{1}{2} \quad 0 \quad \frac{1}{2} \quad 0]$$

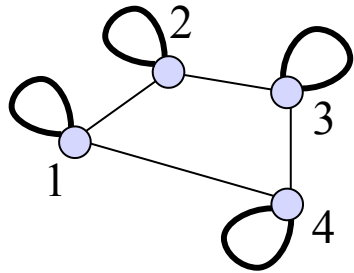
$$x_{2t+1} \rightarrow [0 \quad \frac{1}{2} \quad 0 \quad \frac{1}{2}]$$

- So no limiting distribution...
- Issue: **Periodicity**



Avoiding Periodicity

- What can we do to this graph to make the walk aperiodic?



Called a “*lazy random walk*”

- (one self-loop would suffice)

- in this case:
$$A_{\text{new}} = \begin{pmatrix} \frac{1}{3} & \frac{1}{3} & 0 & \frac{1}{3} \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} & 0 \\ 0 & \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \\ \frac{1}{3} & 0 & \frac{1}{3} & \frac{1}{3} \end{pmatrix} = \frac{2}{3} A_{\text{old}} + \frac{1}{3} I$$

- $x_t \rightarrow ??$

- Let's iterate $x_t = x_0 A_{\text{new}}^t$

-

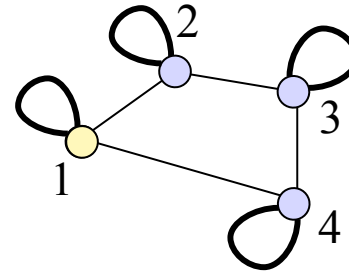
- $x_t \rightarrow [\frac{1}{4} \quad \frac{1}{4} \quad \frac{1}{4} \quad \frac{1}{4}]$

$$A_{\text{old}} = \begin{pmatrix} 0 & \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} & 0 \\ 0 & \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} & 0 \end{pmatrix}$$

Proving $x_t \rightarrow (1/4, 1/4, 1/4, 1/4)$?

- Recall

$$A = \begin{pmatrix} 1/3 & 1/3 & 0 & 1/3 \\ 1/3 & 1/3 & 1/3 & 0 \\ 0 & 1/3 & 1/3 & 1/3 \\ 1/3 & 0 & 1/3 & 1/3 \end{pmatrix}$$

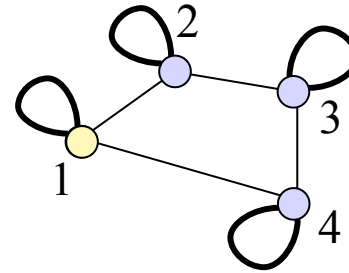


- Idea: look at the *eigenvalues* of A
- A symmetric so it has 4 real eigenvalues
- $\lambda_1 = 1, \lambda_2 = \lambda_3 = 1/3, \lambda_4 = -1/3$ (thanks Matlab)
- coincidence: $\lambda_2 = \lambda_3$ and $\lambda_4 = -\lambda_2$ (4-cycle is a special graph)
- non-coincidence (true for any *lazy* r.w. on an undirected *connected* graph):
 - largest eigenvalue = 1
 - all others have absolute value < 1
- left eigenvector corresponding to $\lambda_1 = 1$?
- $e_1 = (1/4, 1/4, 1/4, 1/4)$ is a left eigenvector for λ_1
- Wow. Why would $x_t \rightarrow e_1$ as $t \rightarrow \infty$?

Proving $x_t \rightarrow (1/4, 1/4, 1/4, 1/4)$?

- Recall

$$A = \begin{pmatrix} 1/3 & 1/3 & 0 & 1/3 \\ 1/3 & 1/3 & 1/3 & 0 \\ 0 & 1/3 & 1/3 & 1/3 \\ 1/3 & 0 & 1/3 & 1/3 \end{pmatrix}$$



- $\lambda_1 = 1, \lambda_2 = \lambda_3 = 1/3, \lambda_4 = -1/3$
- $e_1 = (1/4, 1/4, 1/4, 1/4)$
- Proof (see also notes): choose eigenvectors e_2, e_3, e_4 for $\lambda_2, \lambda_3, \lambda_4$, so that $\{e_1, e_2, e_3, e_4\}$ is a basis of \mathbb{R}^4 (guaranteed by the spectral theorem, since A is symmetric)
- so $x_0 = a_1 e_1 + a_2 e_2 + a_3 e_3 + a_4 e_4$, for some a_1, a_2, a_3, a_4
- Now $x_t = x_0 A^t =$

$$\begin{aligned} &= a_1 e_1 A^t + a_2 e_2 A^t + a_3 e_3 A^t + a_4 e_4 A^t \\ &= a_1 e_1 \lambda_1^t + a_2 e_2 \lambda_2^t + a_3 e_3 \lambda_3^t + a_4 e_4 \lambda_4^t \\ &\rightarrow a_1 e_1, \text{ as } t \rightarrow \infty \end{aligned}$$

exact same proof for any
lazy r.w. on any connected
undirected graph.

- since $e_1 = (1/4, 1/4, 1/4, 1/4)$ is a distribution, it must be that $a_1 = 1$
- Hence $x_t \rightarrow (1/4, 1/4, 1/4, 1/4)$, as $t \rightarrow \infty$.

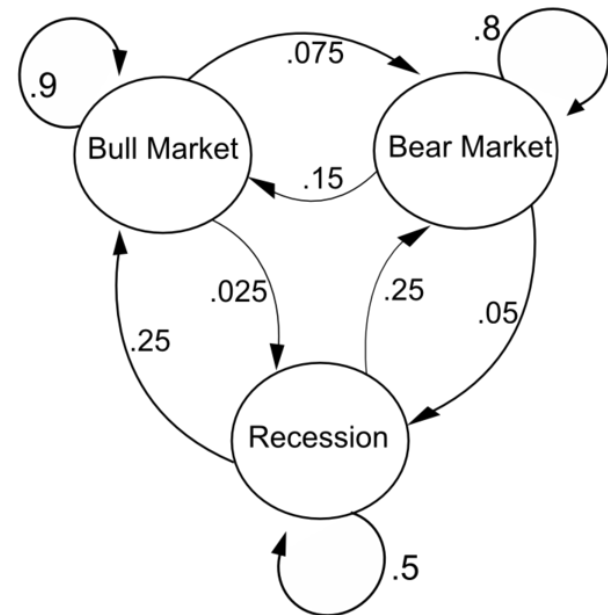
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General Setup

- Let $G=(V, E)$ be a:
 - (i) ***strongly-connected*** weighted directed graph (i.e. there is a path from every vertex to every vertex),
 - with (ii) ***self-loop*** on every vertex (to avoid periodicities)
- The weight p_{ij} of every edge $(i, j) \in E$ represents a probability, namely:
 - p_{ij} : probability of transitioning from i to j .
 - Hence for all i : $\sum_{(i,j) \in E} p_{ij} = 1$.
- Transition matrix: $A = [p_{ij}]$.
- In the example on the left:

$$A = \begin{pmatrix} .9 & .075 & .025 \\ .15 & .8 & .05 \\ .25 & .25 & .5 \end{pmatrix}$$



General Setup

- Transition matrix: $A = [p_{ij}]$.
- Still true that, if x_0 is the distribution over vertices at time 0, then $x_t = x_0 A^t$.
- **Claim:** A has eigenvalue 1 (with multiplicity 1), and there is a unique positive vector x such that:

- $x \cdot A = A$ (*)
- $\sum_i x_i = 1$ (**)

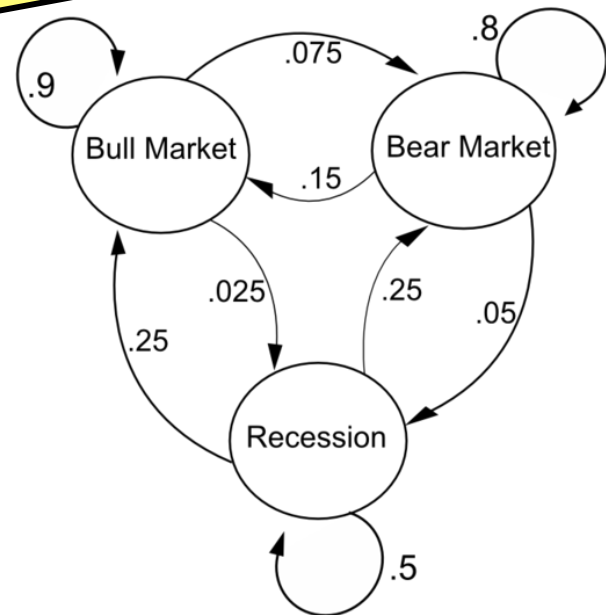
Note that if a limiting distribution x exists, then it definitely needs to satisfy (*), o.w. it wouldn't be a limiting distribution. The point of the theorem is that the Markov Chain indeed converges to the unique eigenvector of A satisfying (*) and (**), no matter what x_0 is.

Theorem: For any x_0 , $x_t \rightarrow x$ as $t \rightarrow \infty$.

- x is called the “stationary distribution of G ”

Two obvious Questions:

- why is x_∞ interesting?
- how fast does $x_t \rightarrow x_\infty$?



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Pagerank

- No better proof that something is useful than having cool applications ☺
- It turns out that Markov Chains have a famous one: *PageRank*.
- PageRank of a webpage $w \approx$ Probability that a web-surfer starting from some central page (e.g. Yahoo!) and clicking random links arrives at webpage w .
- How to compute this probability?
- Form graph $G =$ the hyperlink graph;



- Namely, G has a node for every webpage, and there is an edge from webpage w_1 to webpage w_2 *iff* there is a hyperlink from w_1 to w_2 .
- All outgoing links from a webpage w are equally probable.
- Compute stationary distribution x_∞ , i.e. the left eigenvector of the transition matrix A of G , corresponding to eigenvalue 1.
- Pagerank of page $w = x_\infty(w)$.

Computing Pagerank

- Graph G = the hyperlink graph
- Compute stationary distribution of G , i.e. the left eigenvector of the (normalized by out-degrees) adjacency matrix A of G , corresponding to eigenvalue 1.
- How to compute stationary distribution?
- **Idea 1:** Crawl the web, create giant A , solve eigenvalue problem.
- Runtime $O(n^3)$ using Gaussian elimination
- this is too much for n = size of the web.
- **Idea 2:** Simulate the walk sufficiently many times (theory meets practice)
 - Start at some central page and do random walk for sufficiently many steps;
 - Restart and repeat sufficiently many times;
 - then take $\text{PageRank}(w) \approx$ empirical frequency that random walk ended at w .
- Hope that empirical distribution is good approximation to stationary distribution for the right choice of “sufficiently many” above...
- or at least for the top components of the stationary distribution, which are the most important for ranking the top results.

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Card Shuffling

Q: Why shuffle the deck?

A: well, to start from a uniformly random permutation of the cards

Q2: How many permutations are there?

A: $52! \approx 2^{257} \approx 10^{77}$ – how large is that?

Q3: Getting a random permutation?

- soln1: dice 10^{77} faces
- soln2: shuffle \approx dice

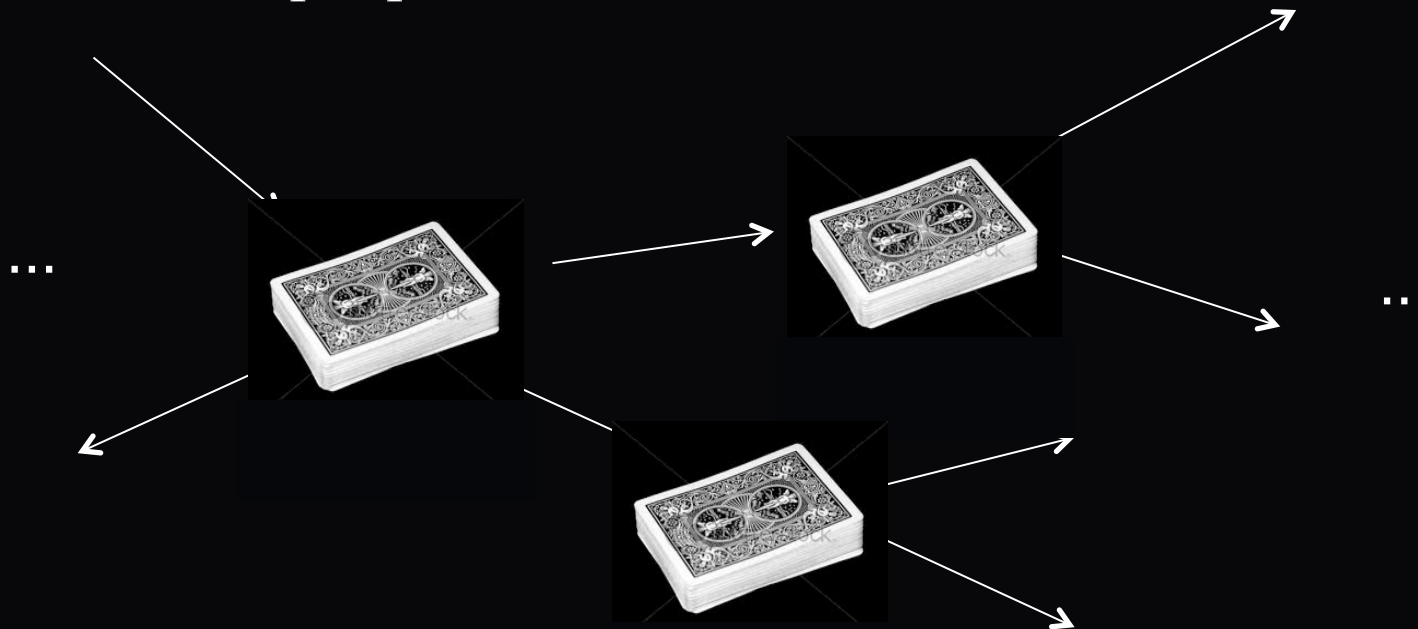
Example Shuffles:

- top-in-at-random
- riffle-shuffle



Shuffling as a Markov Chain

Graph: - one node per permutation of the deck.



- edge (u,v) : v is reachable in one move from u (specific to shuffle)

While performing the shuffle we jump from node to node.

Q: Stationary distribution of a correct shuffle?

Probability $1/52!$ on each permutation.

Mixing Time

Q: How many steps suffice for uniformity?

A1: The question is meaningless as no matter how long you shuffle there is always a small trace of what permutation you started from in x_t .

Q: OK, how many steps suffice to be close to uniform?

- Top-in-at-Random: ~ 300 repetitions suffice

namely $\text{distance}(x_{300}, \text{uniform}) < 1\%$

- Riffle Shuffle: ~ 10 repetitions suffice

namely $\text{distance}(x_{10}, \text{uniform}) < 1\%$

i.e. different shuffles have different graphs, and as a result converge to uniform distribution at different speeds.

Mixing Time (formally)

- Let $x_0, x_1, x_2, \dots, x_t, \dots$ be a Markov Chain with stationary distribution x_∞ .
- **Def:** The *mixing time* of the Markov Chain is the minimum τ such that $d(x_\tau, x_\infty) < .01$.

where for two distributions x, y : $d(x, y) = \sum_i |x(i) - y(i)|$.

- Captures rate at which $x_t \rightarrow x_\infty$. Namely:
- **Claim:** $d(x_t, x_\infty) \leq \exp(-t/\tau_{\text{mix}})$.
- When we design Markov chains we desire τ_{mix} to be small.
- τ_{mix} depends on the connectivity of the Markov Chain.

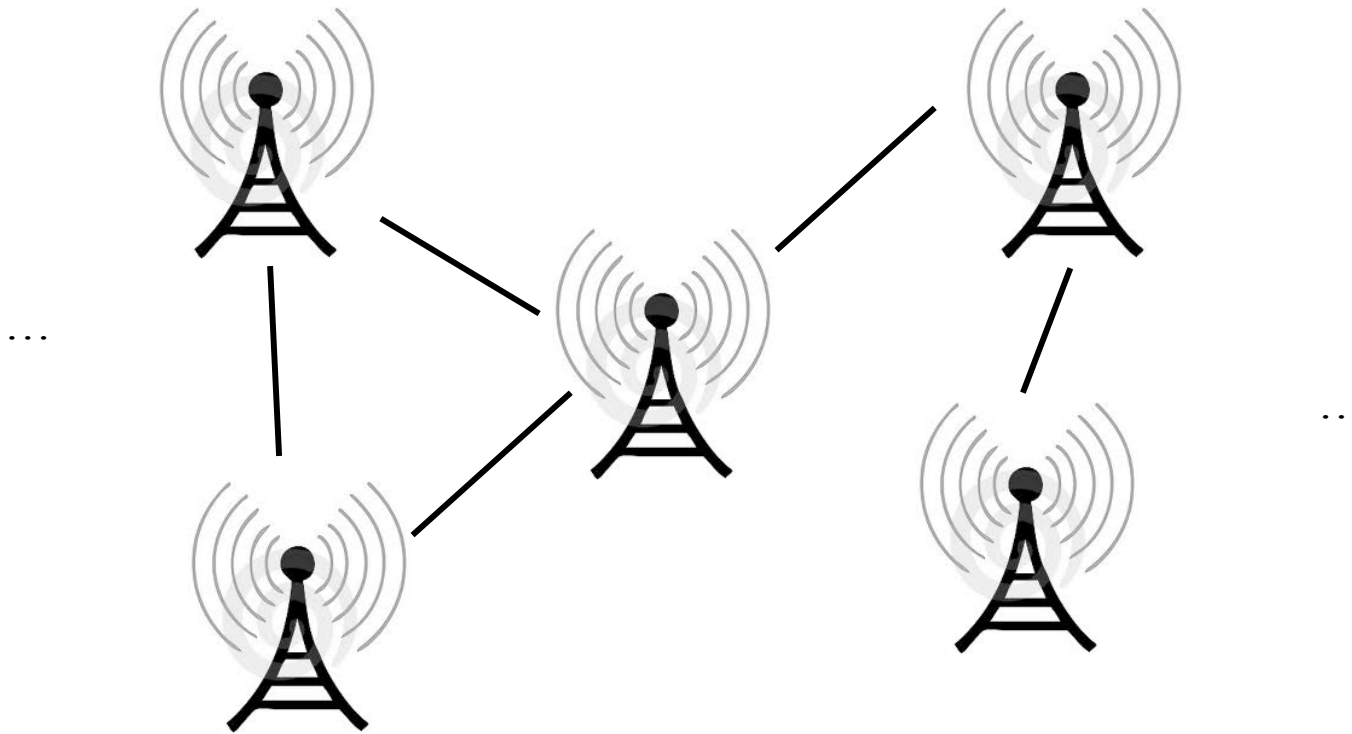
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Allocating Frequencies to Stations

Input: - Graph $G = (V, E)$, V : radio stations, E : interferences
- A set $F = \{f_1, f_2, \dots, f_q\}$ of frequencies.

Goal: Assign frequencies to stations so that no two interfering stations get the same frequency.



continued on the board (see notes)