STSCI 4780 Relationships between distributions

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Sum of Bernoulli 0/1's is binomial

FFSSSFSSFS (n = 8 successes in N = 12 trials)

Bernoulli process

$$p(D|\alpha, M) = p(\text{failure}|\alpha, M) \times p(\text{failure}|\alpha, M) \times \cdots$$

= $\alpha^{n} (1 - \alpha)^{N-n}$
= $\mathcal{L}(\alpha)$

Binomial distribution

Let S = a sequence of flips with n heads.

$$p(n|\alpha, M) = \sum_{S} p(S|\alpha, M) p(n|S, \alpha, M)$$
$$= \alpha^{n} (1 - \alpha)^{N-n} C_{n,N}$$
$$C_{n,N} \equiv \binom{N}{n} = \frac{N!}{n!(N-n)!}$$

Sum of same- α binomial counts is binomial

 N_1 trials with n_1 successes, followed by N_2 trials with n_2 successes, gives total successes $n = n_1 + n_2$ in $N = N_1 + N_2$ trials

• The total corresponds to the successes from N trials, for which

$$p(n|\alpha, N) = \binom{N}{n} \alpha^n (1-\alpha)^{N-n}$$

• Show it explicitly using the LTP:

$$\begin{split} p(n|\alpha,N) &= \sum_{n_1} \sum_{n_2} p(n_1|\alpha,N_1) p(n_2|\alpha,N_2) p(n|n_1,n_2,\ldots) \\ &= \sum_{n_1} p(n_1|\alpha,N_1) p(n_2=n-n_1|\alpha,N_2) \\ &= \sum_{n_1} \binom{N_1}{n_1} \binom{N_2}{n-n_1} \alpha^n (1-\alpha)^{N-n} \\ &= \binom{N}{n} \alpha^n (1-\alpha)^{N-n} \end{split}$$

by a binomial coefficient identity

Sum of Poisson counts is Poisson

 n_1 counts in an interval with $r_1T_1=\lambda$, n_2 counts in an interval with $r_2T_2=\mu$, total counts $n=n_1+n_2$:

$$p(n|\lambda,\mu) = \sum_{n_1} \sum_{n_2} p(n_1|\lambda) p(n_2|\mu) p(n|n_1, n_2, ...)$$

$$= \sum_{n_1} p(n_1|\lambda) p(n_2 = n - n_1|\mu)$$

$$= \sum_{n_1} \frac{\lambda^{n_1}}{n_1!} e^{-\lambda} \frac{\mu^{n-n_1}}{(n-n_1)!} e^{-\mu}$$

$$= \frac{e^{-(\lambda+\mu)}}{n!} \sum_{n_1} \frac{n!}{n_1!(n-n_1)!} \lambda^{n_1} \mu^{n-n_1}$$

$$= \frac{(\lambda+\mu)^n}{n!} e^{-(\lambda+\mu)}$$

Poisson with summed rate \times interval

Sum of normally distributed values is normal

Suppose $x \sim \text{Norm}(0, \sigma^2)$ and $y \sim \text{Norm}(0, \sigma^2)$; what is the PDF for s = x + y?

$$p(s) \propto \exp\left[-rac{s^2}{4\sigma^2}
ight]$$

I.e., normal with $\sigma_s^2 = \sigma_x^2 + \sigma_y^2 = 2\sigma^2$

When the distribution for a sum is in the same family as the distributions for the components, the family is *infinitely divisible*

Poisson for large expected counts

Recall the Poisson distribution for n counts from a process with rate r over an interval T:

$$p(n|r, M) = \frac{(rT)^n}{n!}e^{-rT} = \frac{\mu^n}{n!}e^{-\mu}$$
 for $\mu \equiv rT$

Expectation value of n:

$$\mathbb{E}(n) \equiv \sum_{n=0}^{\infty} np(n|r, I)$$

$$= e^{-\mu} \sum_{n=1}^{\infty} \frac{\mu^n}{(n-1)!}$$

$$= \mu e^{-\mu} \sum_{m=0}^{\infty} \frac{\mu^m}{m!} \quad \text{for } m = n-1$$

$$= \mu$$

Similarly, $Var(n) = \mathbb{E}(n^2) - \mu^2 = \mu$, so the standard deviation $\sigma_n = \mu^{1/2}$

For $\mu \gg 1$, expect $n \approx \mu \pm \mu^{1/2}$ so

$$\delta \equiv \frac{n-\mu}{\mu} \ll 1$$

In terms of δ , we can write $n = \mu(\delta + 1)$

Look for the leading order term in $\log p(n)$; use two approximations, Stirling's approximation:

$$\log(n!) \approx n \log(n) - n$$

and a Taylor expansion for the logarithm:

$$\log(1+x)\approx x-\frac{x^2}{2}+\cdots$$

$$\log p(n) \approx n \log \mu - n \log n + n - \mu$$

$$= -n \log \frac{n}{\mu} + n - \mu$$

$$= -\mu(\delta + 1) \log(1 + \delta) + \mu(\delta + 1) - \mu$$

$$\approx -\mu(\delta + 1) \left(\delta + \frac{\delta^2}{2}\right) + \mu\delta$$

$$\approx -\mu\delta^2 - \mu\delta + \frac{\mu\delta^2}{2} + \mu\delta$$

$$= -\frac{\mu\delta^2}{2}$$

So

$$p(n|\mu) \propto \exp\left(-rac{(n-\mu)^2}{2\mu}
ight)$$

a Gaussian distribution (evaluated at integer values of n)

Binomial for rare events

How many Cornell students share your birthday?

- $N \approx 24,000 \gg 1$
- $\alpha \approx \frac{1}{365} \ll 1$
- Expected number $\mu \equiv \mathbb{E}(n) = \alpha N \approx 66 \gg 1$

1000 bacteria are mixed in a liter of water. How many are in a 0.1 ml sample?

- $N = 1000 \gg 1$
- $\alpha = 10^{-4}$
- Expected number $\mu \equiv \mathbb{E}(n) = \alpha N = 0.1 \ll 1$

Seek an approximation for $p(n|\ldots)$ for small α , but not necessarily small μ (or n): A rare event can happen many times in a very large sample

Recall the binomial sampling distribution for n successes in N trials, given success probability α :

$$p(n|\alpha,M) = \frac{N!}{n!(N-n)!}\alpha^n(1-\alpha)^{N-n}$$

Expected number of successes $\mu \equiv \mathbb{E}(n) = \alpha N$

Recursion relation:

$$\frac{p(n)}{p(n-1)} = \frac{N!}{n!(N-n)!} \frac{(n-1)!(N-n+1)!}{N!} \frac{\alpha}{1-\alpha}$$
$$= \frac{N-n+1}{n} \frac{\alpha}{1-\alpha}$$

Consider the limit where $N \to \infty$ and $\alpha \to 0$, but with $\mu = \alpha N$ fixed and not necessarily small (but $\mu \ll N$); focus on $n \sim \mu$ so $n \ll N$ as well:

$$\frac{p(n)}{p(n-1)} \approx \frac{N\alpha}{n} = \frac{\mu}{n}$$

In that same limit, writing α in terms of μ and N,

$$p(0) = (1 - \alpha)^N = \left(1 - \frac{\mu}{N}\right)^N \approx e^{-\mu}$$

Now evaluate p(n) using the recurrence relation:

$$p(1) = \frac{\mu}{1} \times p(0) = \mu e^{-\mu}$$

$$p(2) = \frac{\mu}{2} \times p(1) = \frac{\mu^{2}}{2} e^{-\mu}$$

$$p(n) = \frac{\mu}{n} \times p(n-1) = \frac{\mu^{n}}{n!} e^{-\mu} = Poisson$$

The Poisson limit theorem or law of rare events (events rare in proportion, though possibly numerous)

Binomial for large N: de Moivre-Laplace theorem

For N Bernoulli outcomes with probability α and n successes, note that

$$\mathbb{E}(n) = \alpha N$$
; $\operatorname{Var}(n) = \alpha (1 - \alpha) N$

Define

$$\Delta = \frac{n - \alpha N}{\sqrt{\alpha (1 - \alpha) N}}$$

Then for large N,

$$p(n|\alpha, N) \propto \exp\left(-\frac{\Delta^2}{2}\right)$$

This is a special case of a more general sum-of-outcomes result, the *central limit theorem* (CLT)