STSCI 4780

Multivariate relationships, 2: The bivariate normal distribution and joint/conditional/marginal relationships

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Recap

Building models for multivariate data via joint/conditional/marginal distribution relationships:

- Types of studies of multivariate data:
 - ▶ Correlation/dependence: Study joint, p(x, y)
 - ▶ Regression: Study conditional, p(y|x)
- BT as posterior = (joint for everything)/(marginal for knowns)
- Directed acyclic graphs (DAGs)
- Conditional independence
- Example: binomial prediction: $n_2 \perp n_1 | \alpha$

Agenda

1 Bivariate normal distribution & regression

2 Joint from conditionals

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Regression perspective on the bivariate normal distribution

Outline of BVN development:

- Write a joint dist'n for two variables (x, y) factored as [marginal for x] \times [cond'l for y given x]:
 - Normal marginal for x; mean μ_x , variance σ_x^2
 - Normal cond'l for y; conditional variance $\tilde{\sigma}_y^2$
 - ► Simplest nontrivial regression function:

$$\mathbb{E}(y|x) = \beta_0 + \beta_1 x$$

$$\Rightarrow p(x, y) = \text{Norm}(x|\mu_x, \sigma_x) \times \text{Norm}(y|\beta_0 + \beta_1 x, \tilde{\sigma}_y)$$

Five parameters: μ_x, σ_x ; $\beta_0, \beta_1, \tilde{\sigma}_y$

- Rewrite it as [marginal for y] \times [cond'l for x given y]:
 - ▶ Find the marginal for y by completing the square in $x \rightarrow$ normal integral
 - Find the cond'l for x via p(x) = p(x, y)/p(y)

$$\Rightarrow p(x, y) = \text{Norm}(y|\mu_y, \sigma_y) \times \text{Norm}(x|\alpha_0 + \alpha_1 y, \tilde{\sigma}_y)$$

Five parameters: μ_y, σ_y ; $\alpha_0, \alpha_1, \tilde{\sigma}_x$,

$$\mu_{y} = \beta_{0} + \beta_{1}\mu_{x}, \qquad \sigma_{y}^{2} = \tilde{\sigma}_{y}^{2} + \beta_{1}^{2}\sigma_{x}^{2};$$

$$\alpha_{0} = \mu_{x} - \beta_{1}\frac{\sigma_{x}^{2}}{\sigma_{y}^{2}}\mu_{y}, \qquad \alpha_{1} = \beta_{1}\frac{\sigma_{x}^{2}}{\sigma_{y}^{2}};$$

$$\tilde{\sigma}_{x}^{2} = \sigma_{x}^{2}\frac{\tilde{\sigma}_{y}^{2}}{\sigma_{y}^{2}}$$

• Write the joint *symmetrically*, using the four marginal parameters $(\mu_x, \sigma_x, \mu_y, \sigma_y)$ and the *correlation coefficient*:

$$\rho = \beta_1 \frac{\sigma_x}{\sigma_y} = \alpha_1 \frac{\sigma_y}{\sigma_x}$$

$$\Rightarrow p(x, y) = \frac{1}{2\pi\sigma_x \sigma_y \sqrt{1 - \rho^2}} \exp\left[-\frac{Q'(x, y)}{2(1 - \rho^2)}\right]$$

$$Q'(x, y) = \left(\frac{x - \mu_x}{\sigma_x}\right)^2 + \left(\frac{y - \mu_y}{\sigma_y}\right)^2 - 2\rho \frac{x - \mu_x}{\sigma_x} \frac{y - \mu_y}{\sigma_y}$$

Regression functions in terms of joint parameters:

$$\mathbb{E}(y|x) = \mu_y + \beta_1(x - \mu_x)$$

$$= \mu_y + \rho\sigma_y \left(\frac{x - \mu_x}{\sigma_x}\right)$$

$$\mathbb{E}(x|y) = \mu_x + \alpha_1(y - \mu_y)$$

$$= \mu_x + \rho\sigma_x \left(\frac{y - \mu_y}{\sigma_y}\right)$$

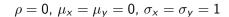
The regression lines track the conditional mean, which is the (conditional) mode for normal distributions; thus they intersect contours of p(x, y) where they are vertical (for y|x) or horizontal (for x|y)

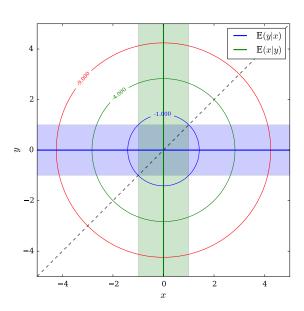
 Matrix forms (inverse covariance matrix vs. inverse correlation matrix):

$$Q'(x,y) = \begin{bmatrix} (x - \mu_x) & (y - \mu_y) \end{bmatrix} \begin{bmatrix} \frac{1}{\sigma_x^2} & -\frac{\rho}{\sigma_x \sigma_y} \\ -\frac{\rho}{\sigma_x \sigma_y} & \frac{1}{\sigma_y^2} \end{bmatrix} \begin{bmatrix} x - \mu_x \\ y - \mu_y \end{bmatrix}$$
$$= \begin{bmatrix} \frac{x - \mu_x}{\sigma_x} & \frac{y - \mu_y}{\sigma_y} \end{bmatrix} \begin{bmatrix} 1 & -\rho \\ -\rho & 1 \end{bmatrix} \begin{bmatrix} \frac{x - \mu_x}{\sigma_x} \\ \frac{y - \mu_y}{\sigma_y} \end{bmatrix}$$

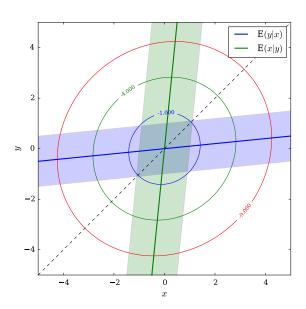
Covariance and correlation matrices:

$$\mathsf{Covar} \ = \begin{bmatrix} \sigma_{\mathsf{x}}^2 & \rho \sigma_{\mathsf{x}} \sigma_{\mathsf{y}} \\ \rho \sigma_{\mathsf{x}} \sigma_{\mathsf{y}} & \sigma_{\mathsf{y}}^2 \end{bmatrix}, \qquad \mathsf{Corr} \ = \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}$$

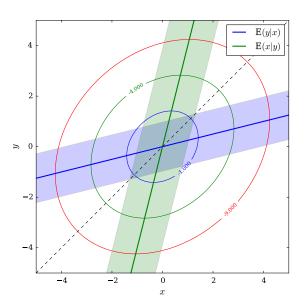




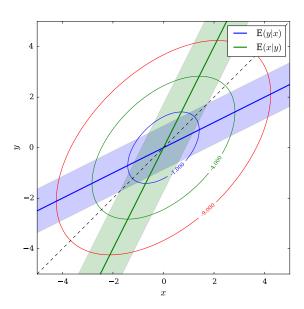




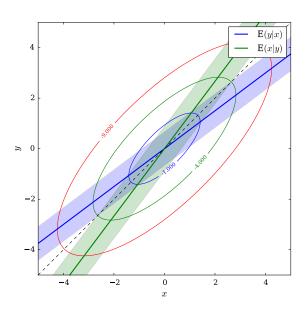




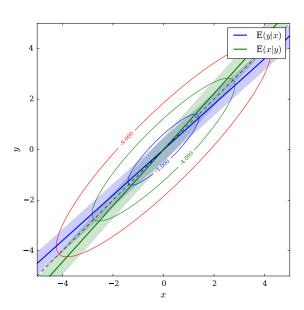




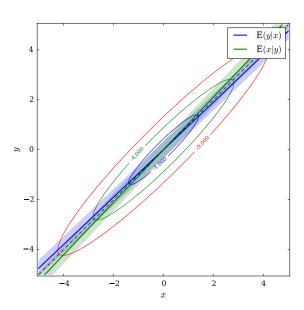












Features to note

- The regression lines track the conditional mean, which is the (conditional) mode for normal distributions; thus they intersect contours where they are vertical (for y|x) or horizontal (for x|y)
- The regression lines for y|x and x|y are not the same; i.e., to find $\mathbb{E}(x|y)$ (a value of x as a function of y) we *do not* take the (blue) line:

$$y(x) = \mathbb{E}(y|x) = \beta_0 + \beta_1 x$$

and solve for $x(y) = -\beta_0 + \frac{1}{\beta_1}y$ In fact, the slopes of *both* regression lines are $\propto \beta_1$

- Neither regression line is the symmetry axis of the elliptical countours, except when $\rho=\pm1$, when the ellipse collapses to a line
- The ellipse *always* has slope 1 (or -1), regardless of the value of ρ (or β_1); more generally its slope is σ_V/σ_X

Whence "regression"?

Recall the regression function for y given x

$$\mathbb{E}(y) = \mu_y + \rho \, \sigma_y \, \frac{x - \mu_x}{\sigma_x}$$

$$\to \frac{\mathbb{E}(y - \mu_y)}{\sigma_y} = \rho \, \frac{x - \mu_x}{\sigma_x}$$

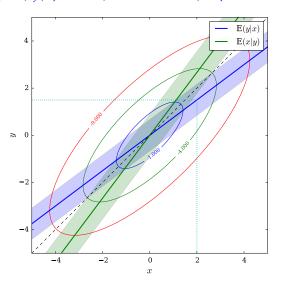
Since $|\rho| \le 1$, so long as $\rho \ne 0$,

For a given (observed) x, we expect y to deviate from its mean by a smaller relative amount than x does from its mean

Galton, in a study of inheritance of stature, referred to this effect as "regression to mediocrity"

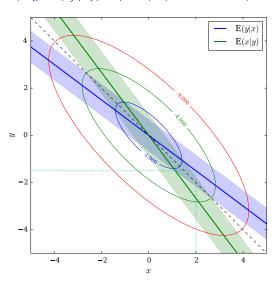
$$\rho = 0.75, \; \mu_{x} = \mu_{y} = 0, \; \sigma_{x} = \sigma_{y} = 1$$

$\mathbb{E}(y - \mu_y | x) < x - \mu_x$ when $x > \mu_x$ (else reverse)



Negative correlation: $\rho = -0.75$

$$|\mathbb{E}(y - \mu_y|x)| < |x - \mu_x|$$
 when $x > \mu_x$



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Joint distribution from conditionals?

The symmetric parameterization of the BVN has 5 parameters:

- Marginal means: μ_x, μ_y
- Marginal standard deviations: σ_x, σ_y
- Correlation coefficient: ρ

If we fix $(\mu_x, \sigma_x, \mu_y, \sigma_y)$ and vary ρ , we generate a family of distributions with *identical marginals but different joint distributions*

Specifying marginals does not uniquely determine the joint

Specifying one marginal and its associated conditional does give the joint:

$$p(x, y) = p(x) p(y|x)$$

= $p(y) p(x|y)$

What about specifying the two conditionals?

Hammersly-Clifford theorem

We'll be evaluating joint, marginal, and conditional distributions for multiple choices of (x, y), so we introduce notation distinguishing the various functions (instead of using p() for everything):

$$f(x,y) \equiv p(x,y)$$

$$m_1(x) \equiv p(x) = \int dy \, p(x,y)$$

$$m_2(y) \equiv p(y) = \int dx \, p(x,y)$$

$$c_{12}(x;y) \equiv p(x|y)$$

$$c_{21}(y;x) \equiv p(y|x)$$

From the product rule, for any choice of a, b,

$$f(a,b)=m_1(a)\,c_{21}(b;a)$$
 $ightarrow m_1(a)=rac{f(a,b)}{c_{21}(b;a)}, ext{ for any } b$
similarly $m_2(b)=rac{f(a,b)}{c_{12}(a;b)}, ext{ for any } a$

Now use the product rule for p(x, y), replacing marginals:

$$f(x,y) = m_1(x) c_{21}(y;x)$$

$$= \frac{f(x,b)}{c_{21}(b;x)} c_{21}(y;x), \text{ for any } b$$

$$= \frac{m_2(b)c_{12}(x;b)}{c_{21}(b;x)} c_{21}(y;x)$$

$$= f(a,b)\frac{c_{12}(x;b)}{c_{12}(a;b)} \frac{c_{21}(y;x)}{c_{21}(b;x)}$$

for any choice (a, b) (requires a *positivity condition*: support of joint = cartesian product of supports of marginals)

$$f(x,y) = f(a,b) \frac{c_{12}(x;b)}{c_{12}(a;b)} \frac{c_{21}(y;x)}{c_{21}(b;x)}$$

Here f(a, b) is independent of (x, y), playing the role of a normalization constant for the remaining (x, y)-dependent factors

Knowing all the conditionals uniquely determines the joint

Uses of this result (and its generalizations):

- Pseudo-likelihood methods
- Complex graphical models—Markov random fields
- Gibbs sampling: Using conditionals to build a MH proposal distribution