

Hidden Markov Models (HMM) continued

ECE/CS 498 DS U/G

Lecture 15

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Announcements

- MP2 Checkpoint 3 due on Wednesday, Mar 27
- MP3 will be released on Friday, Mar 29
- ICA 4 on HMMs today
- Graduate project details released (check Piazza, course website)
 - Two proposal ideas due on Friday, Mar 29
 - Encouraged to start early and share ideas before the deadline

Markov Models vs HMM

A **Markov Model** can be specified by the following components.

Component	Explanation
$x = \{1, 2, \dots, N\}; x_t \in x$	A set of N states that can be observed directly
$A = \begin{bmatrix} a_{11} & \dots & \dots & \dots & a_{N1} \\ \vdots & \ddots & & & \vdots \\ a_{1j} & \dots & a_{ij} & \dots & a_{Nj} \\ \vdots & \dots & \dots & \ddots & \dots \\ a_{1N} & \dots & \dots & \dots & a_{NN} \end{bmatrix}$	A transition probability matrix A , each a_{ij} representing the probability of moving from state i to state j , s. t. $\sum_{j=1}^N a_{ij} = 1 \forall i$
$\pi = \pi_1, \pi_2, \dots, \pi_N$	An initial probability distribution over states. π_i is the probability that the Markov chain will start in state i . Some states j may have $\pi_j = 0$, meaning that they cannot be initial states. Also, $\sum_{i=1}^N \pi_i = 1$

A **Markov Model** embodies the Markov Assumption:

$$P(x_{t+1}|x_0, \dots, x_t) = P(x_{t+1}|x_t)$$

Markov Models vs HMM

A **Hidden Markov Model (HMM)** can be specified by the following components.

Component	Explanation
$S = \{\sigma_1, \sigma_2, \dots, \sigma_n\}; S_t \in S$	A set of N states that are hidden and cannot be directly observed
$A = \begin{bmatrix} a_{11} & \dots & \dots & \dots & a_{N1} \\ \vdots & \ddots & & & \vdots \\ a_{1j} & \dots & a_{ij} & \dots & a_{Nj} \\ \vdots & \dots & \dots & \ddots & \dots \\ a_{1N} & \dots & \dots & \dots & a_{NN} \end{bmatrix}$	A transition probability matrix A, each a_{ij} representing the probability of moving from state i to state j , s. t. $\sum_{j=1}^N a_{ij} = 1 \forall i$
$E = \{\epsilon_1, \epsilon_2, \dots, \epsilon_M\}; E_t \in E$	A set of observable events
$O = E_1, E_2, \dots, E_T$	A sequence of T observations
$B = \begin{bmatrix} b_{11} & \dots & \dots & \dots & b_{M1} \\ \vdots & \ddots & & & \vdots \\ b_{1j} & \dots & b_{ij} & \dots & b_{Mj} \\ \vdots & \dots & \dots & \ddots & \dots \\ b_{1N} & \dots & \dots & \dots & b_{MN} \end{bmatrix}$	An observation matrix B. Each b_{ij} is referred to as an emission probability or observation likelihood. i. e $b_{ij} = P(E = \epsilon_j S = \sigma_i)$
$\pi = \pi_1, \pi_2, \dots, \pi_N$	An initial probability distribution over states. π_i is the probability that the Markov chain will start in state i . Some states j may have $\pi_j = 0$, meaning that they cannot be initial states. Also, $\sum_{i=1}^N \pi_i = 1$

A HMM embodies the **Markov Assumption**:

$$P(S_{t+1} | S_0, \dots, S_t) = P(S_{t+1} | S_t)$$

A HMM also follows **Output Independence**:

$$P(E_t | S_0, \dots, S_t, \dots, S_T, E_1, \dots, E_t, \dots, E_T) = P(E_t | S_t)$$

Forwards Algorithm

1. Input: (A, B, π) and observed sequence E_1, \dots, E_n

2. $[\alpha_1, Z_1] = \text{normalize}(b_1 \otimes \pi)$

$$\alpha_t(j) = \frac{1}{Z_t} P(E_t | S_t = \sigma_j) \sum_{i=1}^N P(S_t = \sigma_j | S_{t-1} = \sigma_i) \alpha_{t-1}(i)$$

3. **for** $t = 2:n$ **do**

$$[\alpha_t, Z_t] = \text{normalize}(b_t \otimes (A^T \alpha_{t-1}))$$

$$Z_t = \sum_{j=1}^N \alpha_t(j)$$

4. return $\alpha_1, \dots, \alpha_n$ and $\log(P(E_1, \dots, E_n)) = \sum_t \log(Z_t)$

5. Subroutine: $[v, Z] = \text{normalize}(u)$: $Z = \sum_j u_j$; $v_j = u_j / Z$;

NOTE: \otimes represents elementwise product (Hadamard product)

Backwards Algorithm

1. Input: (A, B, π) and observed sequence E_1, \dots, E_n
2. $\beta_n = 1$; // initialize $\beta_n(j)$ to 1 for all states σ_j
3. **for** $t = n - 1 : 1$ **do**
 $\beta_{t-1} = A(b_t \beta_t)$
4. return β_1, \dots, β_n

Inference – using Forwards-Backwards expressions

$$P(S_t | E_1, E_2, \dots, E_n) = \frac{P(E_{t+1}, \dots, E_n | S_t) P(S_t | E_1, \dots, E_t)}{P(E_{t+1}, \dots, E_n | E_1, \dots, E_t)}$$

For $S_t = \sigma_j$ and $\gamma_t(j) = P(S_t = \sigma_j | E_1, E_2, \dots, E_n)$, the above equation is:

$$P(S_t = \sigma_j | E_1, E_2, \dots, E_n) = \frac{P(E_{t+1}, \dots, E_n | S_t = \sigma_j) P(S_t = \sigma_j | E_1, \dots, E_t)}{P(E_{t+1}, \dots, E_n | E_1, \dots, E_t)}$$

$$\gamma_t(j) = \frac{\beta_t(j) \alpha_t(j)}{P(E_{t+1}, \dots, E_n | E_1, \dots, E_t)} = \frac{\beta_t(j) \alpha_t(j)}{\sum_{i=1}^N \beta_t(j) \alpha_t(j)}$$

Theorem of total probability

$\gamma_t(j) \propto \beta_t(j) \alpha_t(j)$

Inference: Most likely state

- Forwards-backwards algorithm gives $P(S_t = \sigma_j | E_1, \dots, E_n)$ for all j
- Find the **individually most likely state** at time t given all observations

$$S_t^* = \operatorname{argmax}_{j \in \{1, \dots, N\}} \gamma_t(j)$$

HMM Security Example

- Suppose you are a security expert monitoring the NCSA system
- By monitoring the system events, you want to say whether the system is safe or not
 - System's safety is a hidden state
 - Events are observed
 - Events are related to the safety of the system
- Is the system safe?
 - **HMM** to the rescue!

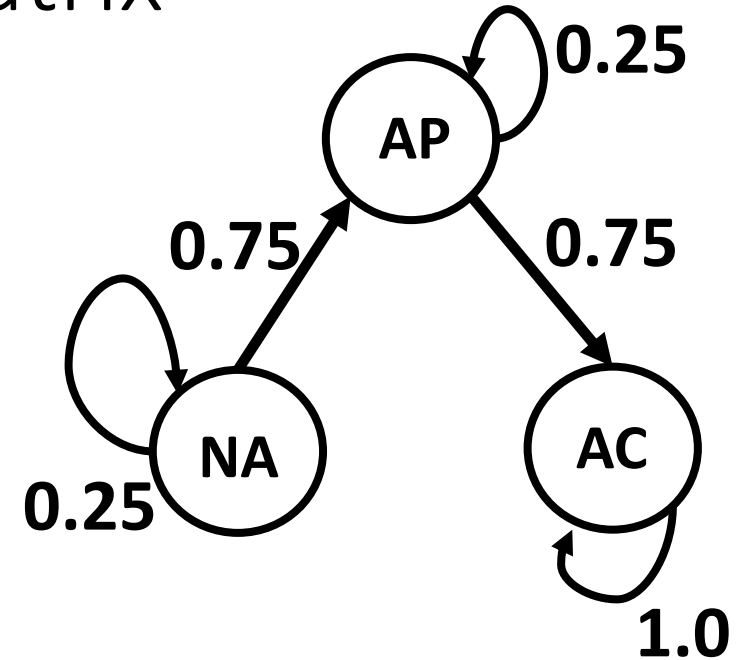
Security Example: Transition Matrix

Transition matrix (A)

The system has three distinct security states –

- (a) No Attack (**NA**),
- (b) Attack in Progress (**AP**), and
- (c) Attack Complete (**AC**).

- Every hour, the system is being attacked by attackers coordinating together around the world and trying to compromise the system.
- The system states always transition from **NA to AP** and **AP to AC**.
- An attacker is successful in changing the state of the system with probability of 0.75 and fails with a probability of 0.25.
- If the attack fails, the system stays in its current state.
- If the system state reaches **AC** the attack is complete, and the system stays in that state.



	NA	AP	AC
NA			
AP			
AC			

Transition Probability Matrix

Security Example: Emission matrix and initial distribution

Observation matrix (B)

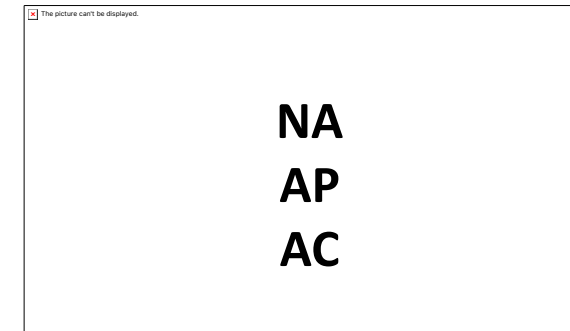
- Your monitoring system reports two types of events
 - Port Scan (**PS**)
 - Software Installation (**SI**)
- Monitors are always accurate and works. Attackers cannot compromise the monitors. Every hour, we get information from the monitors if the attackers are trying to do **PS or SI**.

Initial distribution (π)

- We have no idea about the initial state of the system.

$$\mathbf{B} = \begin{matrix} & \mathbf{PS} & \mathbf{SI} \\ \mathbf{NA} & 0.7 & 0.3 \\ \mathbf{AP} & 0.5 & 0.5 \\ \mathbf{AC} & 0.2 & 0.8 \end{matrix}$$

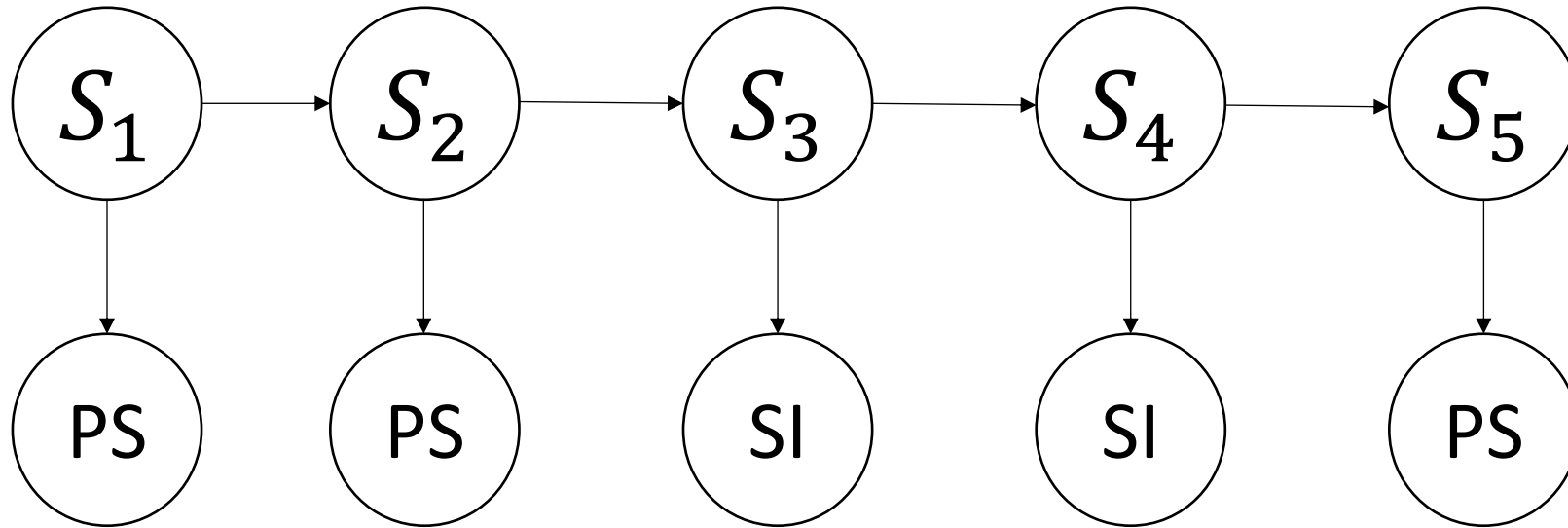
Observation Matrix



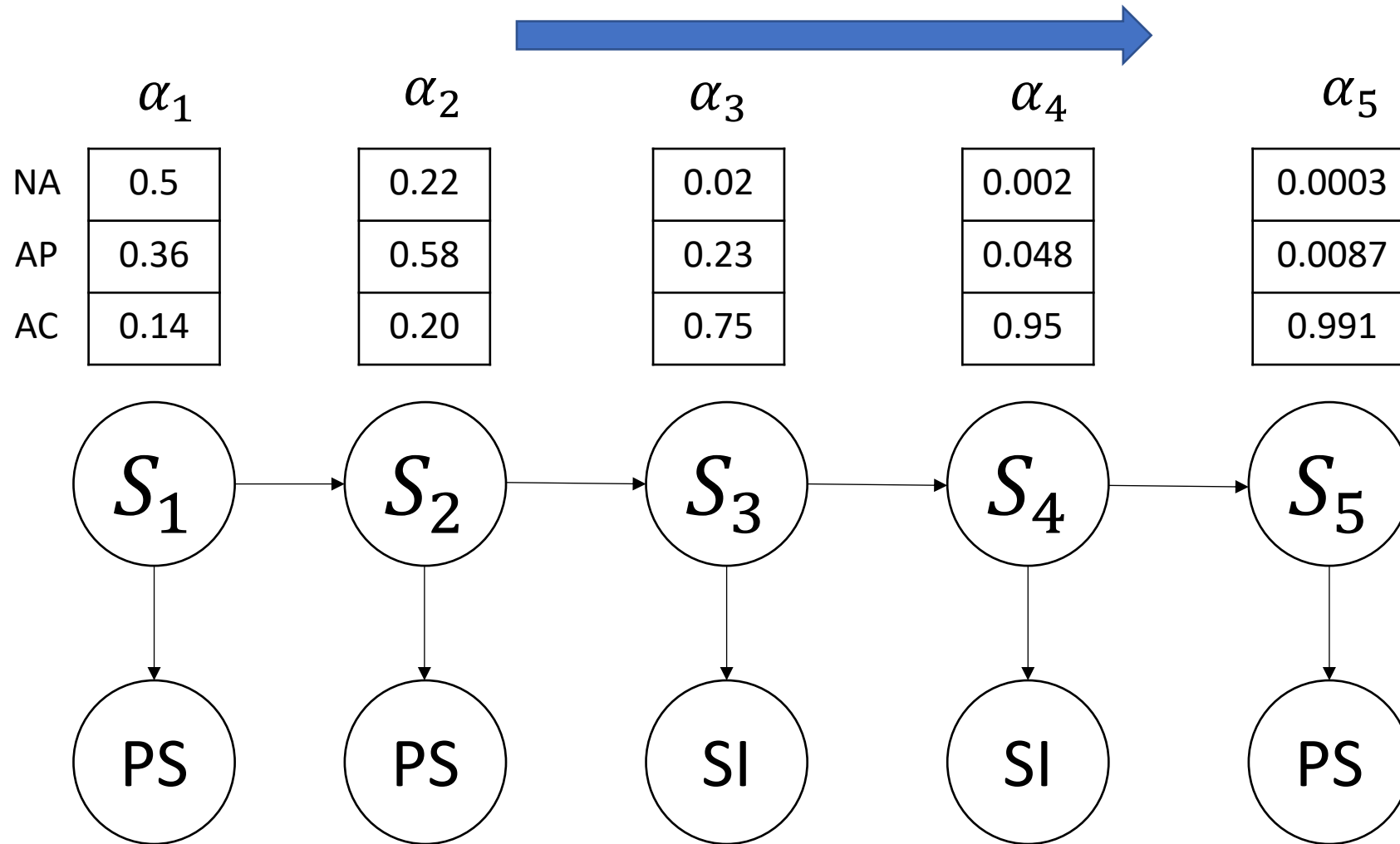
Initial state
distribution/prior

Security Example – Observed Sequence

Find S_1, \dots, S_5 given the observed sequence PS, PS, SI, SI, SI.



Forward Algorithm



$$\alpha_3 \propto b_3 \odot (A^T \alpha_2)$$

$$= \begin{bmatrix} 0.3 \\ 0.5 \\ 0.8 \end{bmatrix} \odot \left(\begin{bmatrix} 0.25 & 0 & 0 \\ 0.75 & 0.25 & 0 \\ 0 & 0.75 & 1 \end{bmatrix} \begin{bmatrix} 0.22 \\ 0.58 \\ 0.20 \end{bmatrix} \right)$$

$$= \begin{bmatrix} 0.3 \\ 0.5 \\ 0.8 \end{bmatrix} \odot \begin{bmatrix} 0.055 \\ 0.31 \\ 0.635 \end{bmatrix}$$

$$= \begin{bmatrix} 0.0165 \\ 0.155 \\ 0.508 \end{bmatrix}$$

Normalizing, we get:

$$\alpha_3 = \frac{1}{0.6795} \begin{bmatrix} 0.0165 \\ 0.155 \\ 0.508 \end{bmatrix}$$

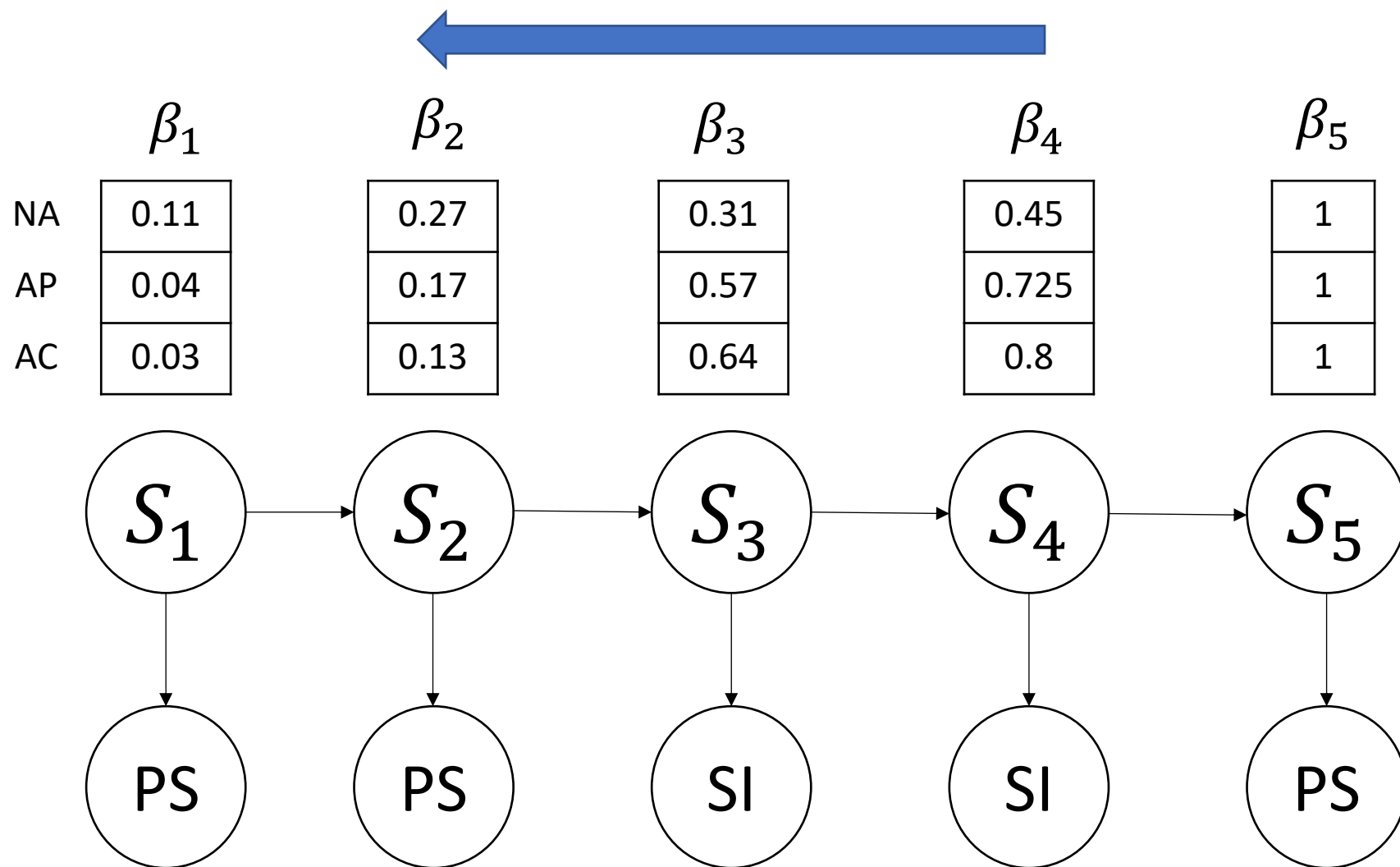
$$= \begin{bmatrix} 0.02 \\ 0.23 \\ 0.75 \end{bmatrix}$$

Forward Algorithm

$$\mathbf{B} = \begin{matrix} & \begin{matrix} \text{PS} & \text{SI} \end{matrix} \\ \begin{matrix} \text{NA} \\ \text{AP} \\ \text{AC} \end{matrix} & \begin{pmatrix} 0.7 & 0.3 \\ 0.5 & 0.5 \\ 0.2 & 0.8 \end{pmatrix} \end{matrix} \quad \mathbf{A} = \begin{matrix} & \begin{matrix} \text{NA} & \text{AP} & \text{AC} \end{matrix} \\ \begin{matrix} \text{NA} \\ \text{AP} \\ \text{AC} \end{matrix} & \begin{pmatrix} 0.25 & 0.75 & 0 \\ 0 & 0.25 & 0.75 \\ 0 & 0 & 1 \end{pmatrix} \end{matrix}$$

States	< PS > (t = 1)		Normalize
NA	$\alpha_1(NA)$	$P(NA) \times P(PS NA) = \frac{1}{3} \times 0.7 = 0.23$	$= \frac{0.23}{0.464} = 0.5$
AP	$\alpha_1(AP)$	$P(AP) \times P(PS AP) = \frac{1}{3} \times 0.5 = 0.167$	$= \frac{0.167}{0.464} = 0.36$
AC	$\alpha_1(AC)$	$P(AC) \times P(PS AC) = \frac{1}{3} \times 0.2 = 0.067$	$= \frac{0.067}{0.464} = 0.14$
	< PS, PS > (t = 2)		
NA	$\alpha_2(NA)$	$(\alpha_1(NA) \times P(NA NA) + \alpha_1(AP) \times P(NA AP) + \alpha_1(AC) \times P(NA AC)) \times P(PS NA) =$ $(0.5 \times 0.25 + 0.36 \times 0 + 0.14 \times 0) \times 0.7 = 0.0875$	$= \frac{0.0875}{0.402} = 0.22$
AP	$\alpha_2(AP)$	$(\alpha_1(NA) \times P(AP NA) + \alpha_1(AP) \times P(AP AP) + \alpha_1(AC) \times P(AP AC)) \times P(PS AP) =$ $(0.5 \times 0.75 + 0.36 \times 0.25 + 0.14 \times 0) \times 0.5 = 0.2325$	$= \frac{0.2325}{0.402} = 0.58$
AC	$\alpha_2(AC)$	$(\alpha_1(NA) \times P(AC NA) + \alpha_1(AP) \times P(AC AP) + \alpha_1(AC) \times P(AC AC)) \times P(PS AC) =$ $(0.5 \times 0 + 0.36 \times 0.75 + 0.14 \times 1) \times 0.2 = 0.082$	$= \frac{0.082}{0.402} = 0.20$

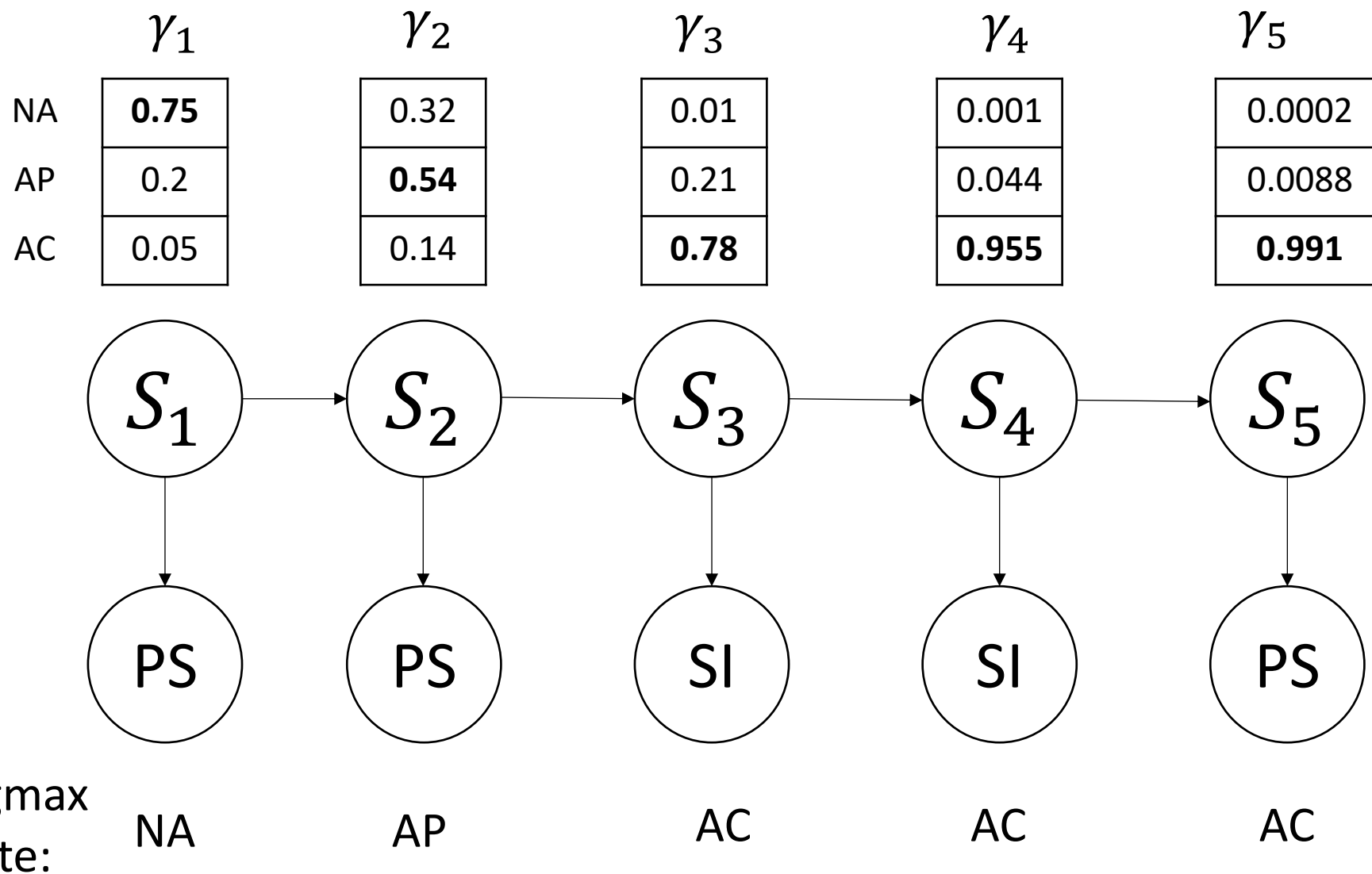
Backward Algorithm



$$\begin{aligned}
 \beta_3 &= A(b_3 \odot \beta_4) \\
 &= \begin{bmatrix} 0.25 & 0.75 & 0 \\ 0 & 0.25 & 0.75 \\ 0 & 0 & 1 \end{bmatrix} \left(\begin{bmatrix} 0.3 \\ 0.5 \\ 0.8 \end{bmatrix} \odot \begin{bmatrix} 0.45 \\ 0.725 \\ 0.8 \end{bmatrix} \right) \\
 &= \begin{bmatrix} 0.25 & 0.75 & 0 \\ 0 & 0.25 & 0.75 \\ 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} 0.135 \\ 0.3625 \\ 0.64 \end{bmatrix} \\
 &= \begin{bmatrix} 0.31 \\ 0.57 \\ 0.64 \end{bmatrix}
 \end{aligned}$$

Note that $\sum_j \beta_t(j)$ is not necessarily 1.

Gamma calculation (using forwards-backwards)



$$\gamma_3 \propto \alpha_3 \odot \beta_3$$

$$= \begin{bmatrix} 0.02 \\ 0.23 \\ 0.75 \end{bmatrix} \odot \begin{bmatrix} 0.31 \\ 0.57 \\ 0.64 \end{bmatrix}$$

$$= \begin{bmatrix} 0.0062 \\ 0.1311 \\ 0.48 \end{bmatrix}$$

Normalizing, we get:

$$\gamma_3 = \frac{1}{0.6173} \begin{bmatrix} 0.0062 \\ 0.1311 \\ 0.48 \end{bmatrix}$$

$$= \begin{bmatrix} 0.01 \\ 0.21 \\ 0.78 \end{bmatrix}$$

$$S_3^* = \operatorname{argmax}_{\text{NA AP AC}} \{0.01, 0.21, 0.78\}$$

$$S_3^* = \text{AC}$$