

Problem Set 0

Issued: Thursday 12th September, 2019

Due: Monday 16th September, 2019

Tips: It is not a formal homework and will not be graded. The primary goal is to help you remember those basic mathematics you have learnt before.

Probability Theory Part

- 0.1. **(Conditional Probability)** For discrete random variables, the conditional probability can be derived by Product Rule.

$$p(X, Y) = p(Y|X) p(X)$$

We can define the conditional expectation as

$$\mathbb{E}[Y|X = x] \triangleq \sum_{y \in \mathcal{Y}} y \cdot p(Y = y|X = x)$$

Explain that

- (a) $\mathbb{E}[X|X] = X$
- (b) $\mathbb{E}[\mathbb{E}[X|Y]] = \mathbb{E}(X)$
- (c) $\mathbb{E}[g(X)h(Y)|Y] = h(Y)\mathbb{E}[g(X)|Y]$ $g(x)$ and $h(y)$ are bounded functions

Solution:

(a)

$$\begin{aligned} p(X = x|X = x) &= 1 \\ \mathbb{E}[X|X] &= X \cdot p(X|X) = X \end{aligned}$$

(b)

$$\begin{aligned} \mathbb{E}[\mathbb{E}[X|Y]] &= \mathbb{E}[g(Y)] \\ &= \sum_{y \in \mathcal{Y}} p(Y = y) \cdot \left[\sum_{x \in \mathcal{X}} x \cdot p(X = x|Y = y) \right] \\ &= \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} x \cdot p(X = x, Y = y) \\ &= \mathbb{E}[X] \end{aligned}$$

(c)

$$\begin{aligned} p(X = x, Y = y|Y = y) &= p(X = x|Y = y) \\ \mathbb{E}[g(X)h(Y)|Y] &= \sum_{x \in \mathcal{X}} g(x) h(Y) p(X = x, Y|Y) \\ &= h(Y) \sum_{x \in \mathcal{X}} g(x) p(X = x|Y) \\ &= h(Y) \mathbb{E}[g(X)|Y] \end{aligned}$$

0.2. (Bayes) A city has a 50% chance to rain everyday and the weather report has a 90% chance to correctly forecast.

You will take an umbrella when the report says it will rain and you have a 50% chance to take an umbrella when the report says it will not rain.

Compute

- (a) the probability of raining when you don't take an umbrella
- (b) the probability of not raining when you take an umbrella

Solution: Let's evaluate the question. \bar{A} denotes the opposite events of A .
Let A be the event **Rain**.

$$p(A) = p(\bar{A}) = 0.5$$

Let B be the event **Forecasting Rain**.

$$p(B|A) = p(\bar{B}|\bar{A}) = 0.9$$

Let C be the event **Taking Umbrella**.

$$p(C|B) = 1$$

$$p(C|\bar{B}) = 0.5$$

OK, now let's come to the questions.

- (a) the probability of raining when you don't take an umbrella $= p(A|\bar{C})$

$$p(A|\bar{C}) = \frac{p(A)p(\bar{C}|A)}{p(A)p(\bar{C}|A) + p(\bar{A})p(\bar{C}|\bar{A})}$$

$$p(\bar{C}|A) = p(\bar{C}|AB)p(B|A) + p(\bar{C}|\bar{A}B)p(\bar{B}|A) = 0 * 0.9 + 0.5 * 0.1 = 0.05$$

$$p(\bar{C}|\bar{A}) = p(\bar{C}|\bar{A}B)p(\bar{B}|\bar{A}) + p(\bar{C}|AB)p(B|\bar{A}) = 0 * 0.1 + 0.5 * 0.9 = 0.45$$

Here, we use that

$$p(\bar{C}|AB) = p(\bar{C}|B)$$

$$p(A|\bar{C}) = \frac{0.5 * 0.05}{0.5 * 0.05 + 0.5 * 0.45} = 0.1$$

- (b) the probability of not raining when you take an umbrella $= p(\bar{A}|C)$

The deduction is the same, so let me omit some steps.

$$p(\bar{A}|C) = \frac{0.5 * 0.55}{0.5 * 0.55 + 0.5 * 0.95} = \frac{11}{30}$$

0.3. (Joint Distribution) Random Variables X and Y have a joint distribution with joint probability density function

$$f(x, y) = \begin{cases} Ce^{-(2x+y)} & x > 0, y > 0 \\ 0 & \text{ow.} \end{cases}$$

Please find C by

$$\int_0^\infty \int_0^\infty f(x, y) dx dy = 1$$

Solution:

$$C \int_0^\infty \int_0^\infty e^{-(2x+y)} dx dy = C \cdot \frac{1}{2} \cdot 1 = 1$$

$$C = 2$$

0.4. (Covariance) For two random variables X and Y , the covariance is defined by

$$\text{Cov}[X, Y] = \mathbb{E}[XY] - \mathbb{E}[X] \mathbb{E}[Y]$$

Now we have a joint pdf

$$f(x, y) = \begin{cases} 4xy & 0 < x < 1, 0 < y < 1 \\ 0 & \text{ow.} \end{cases}$$

Please show that the covariance of X and Y is 0.

Solution:

$$\mathbb{E}[X] = \int_0^1 \int_0^1 x \cdot 4xy dx dy = \frac{2}{3}$$

$$\mathbb{E}[Y] = \int_0^1 \int_0^1 y \cdot 4xy dx dy = \frac{2}{3}$$

$$\mathbb{E}[XY] = \int_0^1 \int_0^1 xy \cdot 4xy dx dy = \frac{4}{9}$$

Thus,

$$\text{Cov}[X, Y] = 0$$

Of course, if you are clever enough, you will see that they are independent.

0.5. (Uncorrelated and independent RVs) We have a uniform distribution of X and Y on a disk. The pdf is

$$f(x, y) = \frac{1}{\pi} \quad x^2 + y^2 \leq 1$$

When the covariance of X and Y is 0, we call them uncorrelated variables.

For continuous random variables, when the joint pdf can be written as the product of two RVs' pdf

$$f(x, y) = f_X(x) f_Y(y),$$

we call them independent.

Please show that X and Y are uncorrelated but not independent.

Solution:

$$f_X(x) = \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} dy = \frac{2\sqrt{1-x^2}}{\pi} \quad -1 < x < 1$$

Similarly,

$$f_Y(y) = \frac{2\sqrt{1-y^2}}{\pi} \quad -1 < y < 1$$

Obviously,

$$f(x, y) \neq f_X(x)f_Y(y) \Rightarrow \text{Not Independent}$$

$$\mathbb{E}[X] = \int_{-1}^1 x \frac{2\sqrt{1-x^2}}{\pi} dx = 0$$

$$\mathbb{E}[Y] = \int_{-1}^1 y \frac{2\sqrt{1-y^2}}{\pi} dy = 0$$

$$\mathbb{E}[XY] = \int_{x^2+y^2 \leq 1} \frac{xy}{\pi} dx dy = 0$$

$$\mathbb{E}[XY] = \mathbb{E}[X] \mathbb{E}[Y] \Rightarrow \text{Uncorrelated}$$

0.6. (Gaussian Distribution) There is a famous integral here

$$\int_{-\infty}^{\infty} e^{-x^2} dx = \sqrt{\pi}$$

It is called Gaussian Integral. Based on it, please find some results of the Normal Distribution

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) \quad -\infty < x < \infty$$

- (a) Prove it is a pdf ($\sigma > 0$)
- (b) Compute the expectation and variance

Solution:

(a)

$$\int_{-\infty}^{\infty} f(x) dx = \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\left(\frac{x-\mu}{\sqrt{2}\sigma}\right)^2\right) d\left(\frac{x-\mu}{\sqrt{2}\sigma}\right) = 1$$

(b)

$$\begin{aligned} \int_{-\infty}^{\infty} x f(x) dx &= \int_{-\infty}^{\infty} (x - \mu + \mu) \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\left(\frac{x-\mu}{\sqrt{2}\sigma}\right)^2\right) d\left(\frac{x-\mu}{\sqrt{2}\sigma}\right) \\ &= \mu + \int_{-\infty}^{\infty} \frac{\sigma}{\sqrt{2\pi}} \exp\left(-\left(\frac{x-\mu}{\sqrt{2}\sigma}\right)^2\right) d\left(\frac{x-\mu}{\sqrt{2}\sigma}\right) \\ &= \mu \end{aligned}$$

$$\begin{aligned}
 \int_{-\infty}^{\infty} x^2 e^{-x^2} dx &= \int_{-\infty}^{\infty} \frac{1}{2} x e^{-x^2} dx^2 \\
 &= - \int_{-\infty}^{\infty} \frac{1}{2} x de^{-x^2} \\
 &= \int_{-\infty}^{\infty} \frac{1}{2} e^{-x^2} dx - \frac{1}{2} x e^{-x^2} \Big|_{-\infty}^{+\infty} \\
 &= \frac{\sqrt{\pi}}{2}
 \end{aligned}$$

$$\begin{aligned}
 \int_{-\infty}^{\infty} x^2 f(x) dx &= \int_{-\infty}^{\infty} \frac{(x - \mu)^2 + 2\mu x - \mu^2}{\sqrt{\pi}} \exp\left(-\left(\frac{x - \mu}{\sqrt{2}\sigma}\right)^2\right) d\left(\frac{x - \mu}{\sqrt{2}\sigma}\right) \\
 &= \frac{\sqrt{\pi}}{2} \cdot \frac{2\sigma^2}{\sqrt{\pi}} + 2\mu \cdot \mu - \mu \cdot \mu \\
 &= \sigma^2 + \mu^2
 \end{aligned}$$

Therefore,

$$\begin{aligned}
 \mathbb{E}[X] &= \mu \\
 \text{Var}[X] &= \sigma^2 + \mu^2 - \mu^2 = \sigma^2
 \end{aligned}$$

Calculus & Linear Algebra

0.7. (Chain rule) $x \in \mathbb{R}$ is a scalar, we have

$$\begin{aligned}
 y &= ax + b \\
 z &= \frac{1}{1 + e^{-y}}
 \end{aligned}$$

Please give the $\frac{\partial z}{\partial x}$.

Solution: According to the chain rule, we have

$$\frac{\partial z}{\partial x} = \frac{\partial z}{\partial y} \times \frac{\partial y}{\partial x} = z(1 - z) \times a$$

0.8. (Orthogonal) The $Q \in \mathbb{R}^{n \times n}$ is said to be **orthogonal** if its columns are pairwise orthogonal, which implies that

$$QQ^\top = Q^\top Q = I$$

Please show that $\|Q\mathbf{x}\|_2 = \|\mathbf{x}\|_2$.

Solution:

$$\begin{aligned}\|Q\mathbf{x}\|_2 &= (Q\mathbf{x})^\top Q\mathbf{x} \\ &= \mathbf{x}^\top Q^\top Q\mathbf{x} \\ &= \|\mathbf{x}\|_2\end{aligned}$$

0.9. (Inner product) If $\mathbf{x} \in \mathbb{R}^n$ is orthogonal to $\mathbf{y} \in \mathbb{R}^n$, please show that

$$\|\mathbf{x} + \mathbf{y}\|^2 = \|\mathbf{x}\|^2 + \|\mathbf{y}\|^2$$

Solution:

$$\begin{aligned}\|\mathbf{x} + \mathbf{y}\|^2 &= (\mathbf{x} + \mathbf{y})^\top (\mathbf{x} + \mathbf{y}) \\ &= \mathbf{x}^\top \mathbf{x} + \mathbf{y}^\top \mathbf{y} + \mathbf{x}^\top \mathbf{y} + \mathbf{y}^\top \mathbf{x} \\ &= \|\mathbf{x}\|^2 + \|\mathbf{y}\|^2\end{aligned}$$

0.10. (Determinant) If a matrix $A \in \mathbb{R}^{n \times n}$ is invertible, the A^* is said to be the **adjoint matrix** of A where $A^{-1} = \frac{A^*}{\det A}$. Please prove that if $\det A = 0$, then we have $\det A^* = 0$.

Solution: Here we use proof by contradiction, that is, assume that when $\det A = 0$ and the $\det A^* \neq 0$, such that the A^* is invertible. We know

$$A^* A = \det A \cdot I$$

hence $A^* \cdot \frac{A}{\det A} = I$, which means $\frac{A}{\det A}$ is the invertible matrix of A^* , and further we get that A is invertible. It contradicts to the assumption that $\det A = 0$, so the $\det A^* = 0$ must hold.

0.11. (Invertibility) Given a matrix $A \in \mathbb{R}^{n \times n}$ and $A^3 = 4I$, please give the invertible matrix of $A - I$.

Solution:

$$\begin{aligned}A^3 - 4I &= 0 \\ A^3 - I &= 3I \\ A^3 - I^3 &= 3I \\ (A - I)(A^2 + A + I) &= 3I \\ (A - I) \cdot \frac{1}{3}(A^2 + A + I) &= I\end{aligned}$$

Therefore we know the inversion of $A - I$ is $\frac{1}{3}(A^2 + A + I)$.

0.12. (Trace) The **trace** of a matrix $A \in \mathbb{R}^{n \times n}$ is defined as sum of diagonal elements of A :

$$\text{tr}(A) = \sum_{i=1}^n A_{ii}$$

(a) Show that $\text{tr}(AB) = \text{tr}(BA)$.

(b) Show that $\nabla_A \text{tr}(AB) = B^\top$.

Solution:

(a) From the definition of trace, we know

$$\begin{aligned} \text{tr}(AB) &= \sum_{i=1}^n \sum_{j=1}^n a_{ij} b_{ji} \\ &= \sum_{i=1}^n \sum_{j=1}^n b_{ij} a_{ji} \\ &= \text{tr}(BA) \end{aligned}$$

(b) We know $\text{tr}(AB) = \sum_{i=1}^n \sum_{j=1}^n a_{ij} b_{ji}$, for an arbitray element a_{kl} , we have

$$\frac{\partial \text{tr}(AB)}{\partial a_{kl}} = b_{lk}$$

Hence we have

$$\nabla_A \text{tr}(AB) = \begin{bmatrix} b_{11} & b_{21} & \cdots & b_{n1} \\ b_{12} & b_{22} & \cdots & b_{n2} \\ \vdots & \vdots & \ddots & \vdots \\ b_{1n} & b_{2n} & \cdots & b_{nn} \end{bmatrix} = B^\top$$

0.13. (Eigenthings) Let \mathbf{x} be an eigenvector of a matrix A with corresponding eigenvalue λ , then

(a) Show that for any $\gamma \in \mathbb{R}$, the \mathbf{x} is an eigenvector of $A + \gamma I$ with eigenvalue $\lambda + \gamma$.

(b) If A is invertible, then \mathbf{x} is an eigenvector of A^{-1} with eigenvalue λ^{-1} .

(c) $A^k \mathbf{x} = \lambda^k \mathbf{x}$ for any $k \in \mathbb{Z}$ ($A^0 = I$ by definition)

Solution:

(a) We have

$$(A + \gamma I)\mathbf{x} = A\mathbf{x} + \gamma\mathbf{x} = (\lambda + \gamma)\mathbf{x}$$

(b) Suppose A is invertible, then

$$\mathbf{x} = A^{-1}A\mathbf{x} = A^{-1}(\lambda\mathbf{x}) = \lambda A^{-1}\mathbf{x}$$

such that we have $A^{-1}\mathbf{x} = \frac{1}{\lambda}\mathbf{x}$.

(c) The case $k > 0$ follows immediately by induction on k , as

$$\begin{aligned} A\mathbf{x} &= \lambda\mathbf{x} \\ A^2\mathbf{x} &= A \cdot A\mathbf{x} = \lambda^2\mathbf{x} \\ A^3\mathbf{x} &= A \cdot A^2\mathbf{x} = \lambda^3\mathbf{x} \\ &\dots \end{aligned}$$

0.14. (Matrix derivative) $\mathbf{x}, \mathbf{w} \in \mathbb{R}^n$, and $A \in \mathbb{R}^{n \times n}$. We have $f : \mathbb{R}^n \rightarrow \mathbb{R}$ as

$$f(\mathbf{x}) = \mathbf{x}^\top A\mathbf{x} + \mathbf{w}^\top \mathbf{x}$$

Please give the $\nabla_{\mathbf{x}} f(\mathbf{x})$.

Solution: The standard solution is, first, we give the differential of $f(\mathbf{x})$:

$$\begin{aligned} df(\mathbf{x}) &= \sum_{i=1}^n \frac{\partial f(\mathbf{x})}{\partial x_i} dx_i \\ &= \frac{\partial f(\mathbf{x})}{\partial \mathbf{x}}^\top d\mathbf{x} \\ &= \text{tr} \left(\frac{\partial f(\mathbf{x})}{\partial \mathbf{x}}^\top d\mathbf{x} \right) \end{aligned}$$

Here we use the trace trick, that is, for a scalar a we have $\text{tr}(a) = a$. Then, for the function above we derive its differential

$$\begin{aligned} df(\mathbf{x}) &= d\mathbf{x}^\top A\mathbf{x} + \mathbf{x}^\top A d\mathbf{x} + \mathbf{w}^\top d\mathbf{x} \\ &= \text{tr}(d\mathbf{x}^\top A\mathbf{x}) + \text{tr}(\mathbf{x}^\top A d\mathbf{x} + \mathbf{w}^\top d\mathbf{x}) \\ &= \text{tr}(\mathbf{x}^\top A^\top d\mathbf{x}) + \text{tr}(\mathbf{x}^\top A d\mathbf{x} + \mathbf{w}^\top d\mathbf{x}) \\ &= \text{tr}((\mathbf{x}^\top A^\top + \mathbf{x}^\top A + \mathbf{w}^\top) d\mathbf{x}) \end{aligned}$$

Refer to the above two equations, we have

$$\frac{\partial f(\mathbf{x})}{\partial \mathbf{x}}^\top d\mathbf{x} = (\mathbf{x}^\top A^\top + \mathbf{x}^\top A + \mathbf{w}^\top) d\mathbf{x}$$

which means

$$\frac{\partial f(\mathbf{x})}{\partial \mathbf{x}} = A\mathbf{x} + A^\top \mathbf{x} + \mathbf{w}$$

Or simply, you can remember the result for convenience

$$\begin{aligned} \frac{\partial \mathbf{x}^\top A\mathbf{x}}{\partial \mathbf{x}} &= (A^\top + A)\mathbf{x} \\ \frac{\partial \mathbf{w}^\top \mathbf{x}}{\partial \mathbf{x}} &= \mathbf{w} \end{aligned}$$