Tsinghua-Berkeley Shenzhen Institute Information Theory and Statistical Learning Fall 2020

Homework 1

Chenyu Tian

September 19, 2020

- Acknowledgments: This template takes some materials from course CSE 547/Stat 548 of Washington University: https://courses.cs.washington.edu/courses/cse547/17sp/index.html.
- I certify that all solutions are entirely in my words and that I have not looked at another student's solutions. I have credited all external sources in this write up.

$$\begin{split} \mathbb{E}[\mathbb{E}[\mathbf{x}|\mathbf{y}\mathbf{z}]|\mathbf{y}] &= \mathbb{E}((\sum_{i} x_{i} P(\mathbf{x} = x_{i}|\mathbf{y}\mathbf{z}))|\mathbf{y}) \\ &= \sum_{j} (\sum_{i} x_{i} P(\mathbf{x} = x_{i}|\mathbf{y}, \mathbf{z} = z_{j})) P(\mathbf{z} = z_{j}) \\ &= \sum_{i} x_{i} (\sum_{j} P(\mathbf{x} = x_{i}|\mathbf{y}, \mathbf{z} = z_{j}) P(\mathbf{z} = z_{j})) \\ &= \sum_{i} x_{i} P(\mathbf{x} = x_{i}|\mathbf{y}) \\ &= \mathbb{E}(\mathbf{x}|\mathbf{y}) \end{split}$$

ii.

$$g(y) \mathbb{E}[x|y] = \sum_{i} g(y)x_i \Pr(x = x_i|y) = \mathbb{E}[xg(y)|y]$$

iii.

$$\mathbb{E}[\mathbf{x}\,\mathbb{E}[\mathbf{x}|\mathbf{y}]] = \mathbb{E}\left[\mathbf{x}\sum_{i}x_{i}\operatorname{Pr}(\mathbf{x} = x_{i}|\mathbf{y})\right]$$

$$= \sum_{j,k}(x_{j}\sum_{i}x_{i}\operatorname{Pr}(\mathbf{x} = x_{i}|\mathbf{y} = y_{k}))\operatorname{Pr}(\mathbf{x} = x_{j}, \mathbf{y} = y_{k})$$

$$= \sum_{i,j,k}x_{i}x_{j}\operatorname{Pr}(\mathbf{x} = x_{i}|\mathbf{y} = y_{k})\operatorname{Pr}(\mathbf{x} = x_{j}, \mathbf{y} = y_{k})$$

$$\begin{split} \mathbb{E}[(\mathbb{E}[\mathbf{x}|\mathbf{y}])^2] &= \mathbb{E}\left[(\sum_i x_i \Pr(\mathbf{x} = x_i|\mathbf{y}))^2\right] \\ &= \sum_k (\sum_i x_i \Pr(\mathbf{x} = x_i|\mathbf{y} = y_k))^2 \Pr(\mathbf{y} = y_k) \\ &= \sum_{i,j,k} x_i x_j \Pr(\mathbf{x} = x_i|\mathbf{y} = y_k) \Pr(\mathbf{x} = x_j|\mathbf{y} = y_k) \Pr(\mathbf{y} = y_k) \\ &= \sum_{i,j,k} x_i x_j \Pr(\mathbf{x} = x_i|\mathbf{y} = y_k) \Pr(\mathbf{x} = x_j, \mathbf{y} = y_k) \end{split}$$

Thus, $\mathbb{E}[(\mathbb{E}[x|y])^2] = \mathbb{E}[x \mathbb{E}[x|y]].$

$$\begin{split} & \mathbb{E}[\mathrm{Var}(\mathbf{x}\mid\mathbf{y})] + \mathrm{Var}(\mathbb{E}[\mathbf{x}\mid\mathbf{y}]) \\ & = \mathbb{E}[\mathbb{E}[\mathbf{x}^2|\mathbf{y}] - (\mathbb{E}[\mathbf{x}|\mathbf{y}])^2] + \mathbb{E}[(\mathbb{E}[\mathbf{x}|\mathbf{y}])^2] - (\mathbb{E}[\mathbb{E}[\mathbf{x}|\mathbf{y}]])^2 \\ & = \mathbb{E}[\mathbb{E}[\mathbf{x}^2|\mathbf{y}]] - (\mathbb{E}[\mathbb{E}[\mathbf{x}|\mathbf{y}]])^2 \\ & = \mathbb{E}[\mathbf{x}^2] - (\mathbb{E}[\mathbf{x}])^2 \\ & = \mathrm{Var}(\mathbf{x}) \end{split}$$

Thus, $Var(x) = \mathbb{E}[Var(x \mid y)] + Var(\mathbb{E}[x \mid y]).$

(b) i.

$$\begin{split} & \operatorname{cov}(\underline{x}) = \mathbb{E}[(\underline{x} - \mathbb{E}[\underline{x}])(\underline{x} - \mathbb{E}[\underline{x}])^{\mathrm{T}}] \\ & = \mathbb{E}[\underline{x}\underline{x}^{\mathrm{T}}] - \mathbb{E}[\underline{x}] \, \mathbb{E}[\underline{x}]^{\mathrm{T}} \\ & = \mathbb{E}[\mathbb{E}[\underline{x}\underline{x}^{\mathrm{T}}|\mathbf{y}]] - \mathbb{E}[\mathbb{E}[\underline{x}|\mathbf{y}]] \, \mathbb{E}[\mathbb{E}[\underline{x}|\mathbf{y}]]^{\mathrm{T}} \\ & = \mathbb{E}[\operatorname{Cov}[\underline{x}|\mathbf{y}] + \mathbb{E}[\underline{x}|\mathbf{y}] \, \mathbb{E}[\underline{x}|\mathbf{y}]^{\mathrm{T}}] - \mathbb{E}[\mathbb{E}[\underline{x}|\mathbf{y}]] \, \mathbb{E}[\mathbb{E}[\underline{x}|\mathbf{y}]]^{\mathrm{T}} \\ & = \mathbb{E}[\operatorname{Cov}[\underline{x}|\mathbf{y}]] + \mathbb{E}[\mathbb{E}[\underline{x}|\mathbf{y}] \, \mathbb{E}[\underline{x}|\mathbf{y}]^{\mathrm{T}}] - \mathbb{E}[\mathbb{E}[\underline{x}|\mathbf{y}]] \, \mathbb{E}[\mathbb{E}[\underline{x}|\mathbf{y}]]^{\mathrm{T}} \\ & = \mathbb{E}[\operatorname{Cov}[\underline{x}|\mathbf{y}]] + \operatorname{Cov}[\mathbb{E}[\underline{x}|\mathbf{y}]] \end{split}$$

ii. If $\exists \underline{c} \in \mathbb{R}^k, c \neq \underline{0}, \underline{c}^{\mathrm{T}}\underline{x} = d$ and d is a constant. $\underline{c}^{\mathrm{T}}\mathbb{E}[\underline{x}] = d$ Then $\underline{c}^{\mathrm{T}}\underline{x}\underline{x}^{\mathrm{T}}\underline{c} = d^2$. Also, $\underline{c}^{\mathrm{T}}\mathbb{E}[\underline{x}\underline{x}^{\mathrm{T}}]\underline{c} = d^2$.

$$\begin{split} \underline{c}^{\mathrm{T}} \operatorname{cov}(\underline{x}) \underline{c} = & \underline{c}^{\mathrm{T}} \operatorname{\mathbb{E}}[\underline{x}\underline{x}^{\mathrm{T}}] \underline{c} - \underline{c}^{\mathrm{T}} \operatorname{\mathbb{E}}[\underline{x}] \operatorname{\mathbb{E}}[\underline{x}]^{\mathrm{T}} \underline{c} \\ = & \underline{c}^{\mathrm{T}} \operatorname{\mathbb{E}}[\underline{x}\underline{x}^{\mathrm{T}}] \underline{c} - \underline{c}^{\mathrm{T}} \operatorname{\mathbb{E}}[\underline{x}] \operatorname{\mathbb{E}}[\underline{x}]^{\mathrm{T}} \underline{c} \\ = & 0 \end{split}$$

And.

$$\underline{c}^{\mathrm{T}} \operatorname{cov}(\underline{x})\underline{c} = \underline{c}^{\mathrm{T}} \mathbb{E}[(\underline{x} - E[\underline{x}])(\underline{x} - E[\underline{x}])^{\mathrm{T}}]\underline{c}$$

$$= \mathbb{E}[\underline{c}^{\mathrm{T}}(\underline{x} - E[\underline{x}])(\underline{x} - E[\underline{x}])^{\mathrm{T}}\underline{c}]$$

$$= 0 \Longleftrightarrow (\underline{x} - E[\underline{x}])(\underline{x} - E[\underline{x}])^{\mathrm{T}}\underline{c} = \underline{0}$$

It has $cov(\underline{x})c = \underline{0}$. Thus, $det(cov(\underline{x})) = 0$.

1.2. (a)

$$\mathbb{E}[(y - ax - b)^2] == a^2 \, \mathbb{E}[x^2] + \mathbb{E}[y^2] + b^2 - 2a \, \mathbb{E}[xy] - 2b \, \mathbb{E}[y] + 2ab \, \mathbb{E}[x]$$

Take the derivatives with respect to a and b to find the minimum when the derivatives equal 0. It has

$$\begin{split} a^* \operatorname{\mathbb{E}}[x^2] + b^* \operatorname{\mathbb{E}}[x] &= \operatorname{\mathbb{E}}[xy] \\ a^* \operatorname{\mathbb{E}}[x] + b^* &= \operatorname{\mathbb{E}}[y] \end{split}$$

Considering
$$\operatorname{var}(x) = \operatorname{var}(y) = \sqrt{\operatorname{var}(x)\operatorname{var}(y)}$$
, it has $a^* = \frac{\mathbb{E}[xy] - \mathbb{E}[x] \operatorname{\mathbb{E}}[y]}{\operatorname{var}(x)} = \rho(x,y)$.

(b) If $x \perp y$,

$$\begin{split} \mathbb{E}[f(x)g(y)] &= \sum_{i,j} f(x_i)g(y_j) \Pr(x = x_i, y = y_j) \\ &= \sum_{i,j} f(x_i)g(y_j) \Pr(x = x_i) \Pr(y = y_j) \\ &= \left(\sum_i f(x_i) \Pr(x = x_i)\right) \left(\sum_i g(y_j) \Pr(y = x_j)\right) \\ &= \mathbb{E}[f(x)] \, \mathbb{E}[g(y)] \\ &\rho(g(y), g(y)) = \frac{\mathbb{E}[(f(x) - \mathbb{E}[f(x)])(g(y) - \mathbb{E}[g(y)])]}{\sqrt{\text{var}(f(x)) \, \text{var}(g(y))}} \\ &= \frac{\mathbb{E}[f(x)g(y)] - \mathbb{E}[f(x)]\mathbb{E}[g(y)]}{\sqrt{\text{var}(f(x)) \, \text{var}(g(y))}} \end{split}$$

So, $\forall f, g, \rho(f(x), g(y)) = 0$. On the other hand, if $\forall f, g, \rho(f(x), g(y)) = 0$. Let f(x) = x, g(y) = y. It exists

$$\mathbb{E}[xy] = \mathbb{E}[x]\mathbb{E}[y] \Longleftrightarrow \Pr(x,y) = \Pr(x)\Pr(y)$$

So, $x \perp y$.

1.3. Denote y = exp(x). Because exp(x) is monotonically increasing, it has

$$f_Y(y) = \frac{\mathrm{d}}{\mathrm{d}y} \Pr(Y \le y) = \frac{\mathrm{d}}{\mathrm{d}y} \Pr(\ln Y \le \ln y) = \frac{\mathrm{d}}{\mathrm{d}y} \Phi(\ln y)$$
$$= \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(\ln y)^2}{2}\right) \frac{\mathrm{d}\ln y}{\mathrm{d}y}$$
$$= \frac{1}{y} \cdot \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(\ln y)^2}{2}\right)$$