## Tsinghua-Berkeley Shenzhen Institute Information Theory and Statistical Learning Fall 2020

## Problem Set 5

**Notations**: We use x, y, w and  $\underline{x}, y, \underline{w}$  to denote random variables and random vectors.

5.1. Cramer-Rao inequality with a bias term. Let  $y \sim f(y; x)$  and let  $\hat{x}(y)$  be an estimator for x. Let  $b(x) = \mathbb{E}[\hat{x}(y)] - x$  be the bias of the estimator. Show that

$$\mathbb{E}\left[(\hat{x}(y) - x)^2\right] \ge \frac{[1 + b'(x)]^2}{J_{y}(x)} + b^2(x)$$

5.2. (a) Let

$$p_{y}(y;x) = \begin{cases} x & \text{if } 0 \le y \le 1/x \\ 0 & \text{otherwise} \end{cases}$$

for x > 0. Show that there exist no unbiased estimators  $\hat{x}(y)$  for x. (Note that because only x > 0 are possible values, an unbiased estimator need only be unbiased for x > 0 rather than all x.)

(b) Suppose instead that

$$p_{y}(y;x) = \begin{cases} 1/x & \text{if } 0 \le y \le x \\ 0 & \text{otherwise} \end{cases}$$

for x > 0. Does a minimum-variance unbiased estimator for x based on y exist? If your answer is yes, determine  $\hat{x}_{MVU}(y)$ . If your answer is no, explain.

5.3. Suppose, for i = 1, 2

$$y_i = x + w_i$$

where x is an unknown but non-zero constant,  $w_1$  and  $w_2$  are statistically independent, zero-mean Gaussian random variables with

$$\operatorname{var}(\mathbf{w}_1) = 1$$
$$\operatorname{var}(\mathbf{w}_2) = \begin{cases} 1 & x > 0 \\ 2 & x < 0 \end{cases}.$$

(a) Calculate the Cramér-Rao bound for unbiased estimators of x based on observation of

$$\underline{\mathbf{y}} = \left[ \begin{array}{c} \mathbf{y}_1 \\ \mathbf{y}_2 \end{array} \right].$$

(b) Show that a minimum variance unbiased estimator  $\hat{x}_{MVU}(\underline{y})$  does not exist. *Hint*: Consider the estimators

$$\hat{x}_1(\underline{y}) = \frac{1}{2}y_1 + \frac{1}{2}y_2,$$
  
 $\hat{x}_2(\underline{y}) = \frac{2}{3}y_1 + \frac{1}{3}y_2.$ 

- 5.4. Let  $\underline{y} = [y_1 \ y_2]^T$  be a vector random variable whose components are i.i.d. Bernoulli random variables with parameter x, 0 < x < 1, i.e.,  $\mathbb{P}(y_i = 1) = x$ , i = 1, 2.
  - (a) Show that  $t(y) = y_1 + y_2$  is a sufficient statistic.
  - (b) Let  $\hat{x}(\underline{y}) = y_1$  be an estimator of the parameter x from the observation  $\underline{y}$ . Find  $MSE_{\hat{x}}(x)$ , the mean-square error of this estimator.
  - (c) Let  $\hat{x}'(t) = \mathbb{E}[\hat{x}(\underline{y})|t=t]$  be an estimator of the parameter x that uses the sufficient statistic t instead of the observations y.
    - i. Show that  $\hat{x}'(t)$  is a valid estimator, i.e., it does not depend on x.
    - ii. Show that  $MSE_{\hat{x}'}(x) = \gamma MSE_{\hat{x}}(x)$  and find the constant  $\gamma$ .
  - (d) We now consider a generalization of this problem. Let  $\underline{y}$  be a random variable generated by a distribution  $p_{\underline{y}}(\cdot;x)$  and  $\underline{t}(\underline{y})$  be a sufficient statistic. Let  $\hat{x}(\underline{y})$  be an estimator of the parameter x based on the observation  $\underline{y}$ . We define an alternate estimator  $\hat{x}'(\underline{t}) = \mathbb{E}[\hat{x}(\underline{y})|\underline{t} = \underline{t}]$ .
    - i. Show that  $\hat{x}'(\underline{t})$  is a valid estimator, i.e., it does not depend on x.
    - ii. Show that for any cost function  $C(x,\hat{x})$  that is convex in  $\hat{x}$ , the following inequality holds:

$$\mathbb{E}[C(x, \hat{x}'(\underline{\mathbf{t}}))] \le \mathbb{E}[C(x, \hat{x}(\mathbf{y}))].$$

- 5.5. For a non-bayesian case  $p_{y}(y;x)$ , we do a binary hypothesis testing where  $x \in \{H_0, H_1\}$ . Please prove that  $t(y) = \frac{p_{y}(y;H_1)}{p_{y}(y;H_0)}$  is a complete sufficient statistics.
- 5.6. In class we developed the EM algorithm for maximum likelihood estimation (EM-ML). That is, we gave an iterative procedure to compute

$$\hat{x}_{ML}(y) = \arg\max_{a} p_{\mathsf{y}}(y; a).$$

and showed that the likelihood was non-decreasing with each iteration. Please develop the EM-MAP algorithm for MAP estimation:

$$\hat{x}_{MAP}(y) = \arg\max_{a} p_{\mathsf{x}|\mathsf{y}}(a|y)$$

where the complete data z is an arbitrary random vector. (Please follow the procedures in the lecture note)