统计计算作业4

7.2 参考"law"数据("bootstrap")。使用基于自助法的水手刀法 se(R)的自助法估计的标准误差。

获取 law 数据,使用 Jackknife-after-Bootstrap 方法估计其标准差,程序运行如下:

```
data(law, package = "bootstrap")
n \leftarrow nrow(law)
B <- 1000
COR <- numeric(B)
indices \leftarrow matrix(0, nrow = B, ncol = n)
for (b in 1:B) {
  i <- sample(1:n, size = n, replace = TRUE)
  LSAT <- law$LSAT[i]
  GPA <- law$GPA[i]
  COR[b] <- cor(LSAT, GPA)
  indices[b, ] <- i
se.jack <- numeric(n)</pre>
for (i in 1:n) {
  keep \leftarrow (1:B)[apply(indices, MARGIN = 1,
                        FUN = function(k) \{ !any(k == i) \} ) ]
  se.jack[i] <- sd(COR[keep])
print(sd(COR))
print(sqrt((n-1) * mean((se.jack - mean(se.jack))^2)))
得到结果:
> print(sd(COR))
[1] 0.135925
> print(sqrt((n-1) * mean((se.jack - mean(se.jack))^2)))
[1] 0.0789038
```

7.3 对例 7.2 中的相关性统计量("bootstrap"中的"law"数据)给出一个自助法 t 置信区间估计。

写一个函数来计算一元或者多元样本下 Bootstrap t 置信区间,默认的置信水平为 95%, Bootstrap 重复数为 500,估计标准差的重复次数默认为 100,程序运行如下:

```
boot.t.ci <-
  function(x, B = 500, R = 100, level = .95, statistic){
    x \leftarrow as.matrix(x)
    n \leftarrow nrow(x)
    stat <- numeric(B)</pre>
    se <- numeric(B)
    boot.se <- function(x, R, f) {</pre>
      x <- as.matrix(x)
      m \leftarrow nrow(x)
       th <- replicate(R, expr = {
         i <- sample(1:m, size = m, replace = TRUE)
         f(x[i, ])
       return(sd(th))
    for (b in 1:B) {
      j <- sample(1:n, size = n, replace = TRUE)</pre>
      y \leftarrow x[j,]
      stat[b] <- statistic(y)
      se[b] <- boot.se(y, R = R, f = statistic)
    stat0 <- statistic(x)
    t.stats <- (stat - stat0) / se
    se0 <- sd(stat)
    alpha <- 1 - level
    Qt <- quantile(t.stats, c(alpha/2, 1-alpha/2), type = 1)
    names(Qt) <- rev(names(Qt))</pre>
    CI <- rev(stat0 - Qt * se0)
    return(CI)
data(law, package = "bootstrap")
stat <- function(x) { cor(x[,1], x[,2]) }
ci <- boot.t.ci(law, statistic = stat, B=1000, R=200)
print(ci)
可得到结果:
> print(ci)
       2.5% 97.5%
-0.2074181 0.9926069
故自助法 t 置信区间估计为[-0.207,0.993]。
```

7.4 参考 "boot"程序包中提供的空调数据集 "aircondit", 观测值是空调设备两次故障之间的小时数。假设故障之间的时间服从指数模型 $Exp(\lambda)$,给出故障率 λ 的

极大似然估计,并使用自助法对该估计的偏差和标准误差进行估计。程序运行如下:

```
library(boot)
mle <- function(x,i){ 1/mean(x[i]) }
aircondit.boot <- boot(data = as.matrix(aircondit), statistic = mle, R = 2000)
print(aircondit.boot)</pre>
```

可得到结果如下:

```
Bootstrap Statistics:
```

original bias std. error t1* 0.00925212 0.001587651 0.0045378

偏差的估计为 0.0016, 标准差的估计为 0.0045.

7.5 分别使用标准正态、基本、百分位数和 BCa 方法计算故障间隔平均时间 1/λ 的 95%自助法置信区间。

程序运行如下:

```
library(boot)
aircondit1 <- as.matrix(aircondit)
aircondit.b <- boot(data = aircondit1, statistic <- function(x,i) {mean(x[i])}, R = 2000)
boot.ci(aircondit.b, type = c("basic","norm", "perc", "bca"))</pre>
```

得到结果如下:

```
Intervals:
Level Normal Basic
95% (33.9, 184.0) (27.6, 171.9)
```

```
Level Percentile BCa
95% (44.3, 188.6) (54.9, 230.8)
Calculations and Intervals on Original Scale
Some BCa intervals may be unstable
```

7.7 五维成绩数据的 5× 5 协方差矩阵Σ具有正的特征值λ1 > > λ5。在主成分分析中 $\theta = \frac{\lambda 1}{\sum_{i=1}^5 \lambda_i}$,给出了第一主成分解释的方差所占比例。

首先进行主成分分析:

```
library(bootstrap)
theta <- function(data, center = TRUE, scale = TRUE) {
   Sigma.hat <-
      cov(scale(as.matrix(data),center = center, scale = scale))
   e.hat <- eigen(Sigma.hat)
   theta.hat <- e.hat$values[1]/sum(e.hat$values)
   return(theta.hat)
}
theta.all <- theta(scor)</pre>
```

```
使用自助法估计偏差和标准误差:
B < -2000
n <- nrow(scor)
theta.B <- numeric(B)
theta.B <- replicate(B, expr = {
  i \leftarrow sample(1:n, size = n, replace = TRUE)
  scor.B <- as.matrix(scor)[i,]</pre>
  theta.B <- theta(scor.B)
})
bias <- mean(theta.B-theta.all)</pre>
std <- sd(theta.B)</pre>
result <- c(theta.all,bias,std)
names(result) <- c("original", "bias.B", "std.B")</pre>
print(result)
结果如下:
> print(result)
      original
                       bias.B
                                        std.B
 0.6361960298 -0.0009985678 0.0434058298
偏差为-0.000999,标准差为 0.0434。
7.8 参考上题,利用水手刀法估计偏差和标准误差。
n <- nrow(scor)
theta.jack <- numeric(n)
for (i in 1:n)
  theta.jack[i] <- theta(as.matrix(scor)[-i,])
bias.jack <- (n-1)*mean(theta.jack-theta.all)
std.jack <- sqrt((n-1)*mean((theta.jack-mean(theta.jack))^2))</pre>
result.jack <- c(theta.all,bias.jack,std.jack)
names(result.jack) <- c("original","bias.jack", "std.jack")</pre>
print(result.jack)
得到结果如下:
> print(result.jack)
     original bias.jack std.jack
 0.6361960298 -0.0003778535 0.0446668393
```

7.10 将双对数模型转换为三次多项式模型并重复分析过程,使用交叉验证法。程序运行如下:

```
install.packages("DAAG")
par(ask=TRUE)
library(DAAG)
attach(ironslag)
n <- length(magnetic)</pre>
e1 <- e2 <- e3 <- e4 <- numeric(n)
for (k in 1:n) {
  y <- magnetic[-k]</pre>
  x <- chemical[-k]
  J1 \leftarrow lm(y \sim x)
  yhat1 \leftarrow J1$coef[1] + J1$coef[2] * chemical[k]
  e1[k] <- magnetic[k] - yhat1
  J2 \leftarrow lm(y \sim x + l(x^2))
  yhat2 \leftarrow J2$coef[1] + J2$coef[2] * chemical[k] +
    J2$coef[3] * chemical[k]^2
  e2[k] <- magnetic[k] - yhat2
  J3 <- lm(y \sim x + I(x^2) + I(x^3))
  yhat3 \leftarrow J3$coef[1] + J3$coef[2] * chemical[k] +
    J3$coef[3] * chemical[k]^2 + J3$coef[4] * chemical[k]^3
  e3[k] <- magnetic[k] - yhat3
  J4 <- lm(log(y) \sim x)
  logyhat4 <- J4$coef[1] + J4$coef[2] * chemical[k]
  yhat4 <- exp(logyhat4)</pre>
  e4[k] <- magnetic[k] - yhat4
crossv \leftarrow c(mean(e1^2), mean(e2^2), mean(e3^2), mean(e4^2))
best.cross.ind <- which(crossv == min(crossv), arr.ind = T)</pre>
L1 <- summary(lm(magnetic ~ chemical))
L2 <- summary(lm(magnetic ~ chemical + I(chemical^2)))
L3 <- summary(lm(magnetic \sim chemical + I(chemical^2) + I(chemical^3)))
L4 <- summary(lm(log(magnetic) ~ chemical))
adj.r <-
  c(L1$adj.r.squared,L2$adj.r.squared,L3$adj.r.squared,L4$adj.r.squared)
best.adj.ind <- which(adj.r == max(adj.r),arr.ind = T)
best.ind <- c(best.cross.ind, best.adj.ind)</pre>
names(best.ind) <- c('best.cross.ind', 'best.adj.ind')</pre>
print(best.ind)
得到结果如下:
best.cross.ind best.adj.ind
故交叉验证法和根据 R2 的最大调整都会选择第二个模型。
```