2019302030053-胡哲-第一

次作业

第三章习题

H3 3-3

四 分析过程

由 题 意 $F(x)=1-(\frac{2}{x})^2, f(x)=8x^{1/3}$,利 用 逆 变 换 方 法 可 得 $F^{-1}(U)=\frac{2}{\sqrt{1-u}},\ u\sim U(0,1).$

```
n <- 1000

u <- runif(n)

x <- 2/sqrt(1-u)

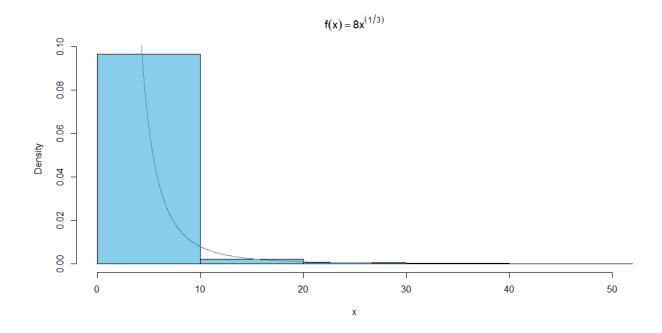
hist(x, prob = TRUE ,main = expression(f(x)==8*x^(1/3)),

col = "skyblue",xlim = c(0,50)) # 样本密度直方图

y <- seq(2, 30, 0.01)

lines(y, 8/(y^3),col="steelblue") #密度曲线 f(x)
```

四 结果



四 结果分析

样本密度直方图和密度曲线大致吻合,说明生成效果较好。

H3 3-4

四 分析过程

瑞利分布可以通过两个正态分布,生成以下内容摘自wiki:

Consider the two-dimensional vector Y=(U,V) which has components that are bivariate normally distributed, centered at zero, and independent. Then U and V have density functions

$$f_U(x;\sigma) = f_V(x;\sigma) = rac{e^{-x^2/(2\sigma^2)}}{\sqrt{2\pi\sigma^2}}.$$

Let X be the length of Y. That is, $X=\sqrt{U^2+V^2}$. Then X has cumulative distribution function

$$F_X(x;\sigma) = \iint_{D_x} f_U(u;\sigma) f_V(v;\sigma) \, dA,$$

where D_x is the disk

$$D_x = \left\{ (u,v) : \sqrt{u^2 + v^2} \leq x
ight\}.$$

Writing the double integral in polar coordinates, it becomes

$$F_X(x;\sigma) = rac{1}{2\pi\sigma^2} \int_0^{2\pi} \int_0^x r e^{-r^2/(2\sigma^2)} \, dr \, d heta = rac{1}{\sigma^2} \int_0^x r e^{-r^2/(2\sigma^2)} \, dr.$$

Finally, the probability density function for X is the derivative of its cumulative distribution function, which by the fundamental theorem of ca

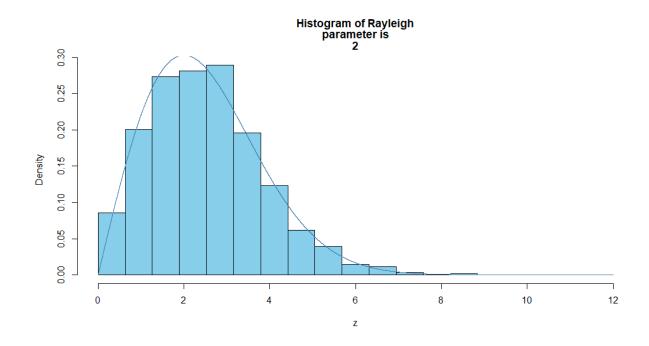
$$f_X(x;\sigma)=rac{d}{dx}F_X(x;\sigma)=rac{x}{\sigma^2}e^{-x^2/(2\sigma^2)},$$

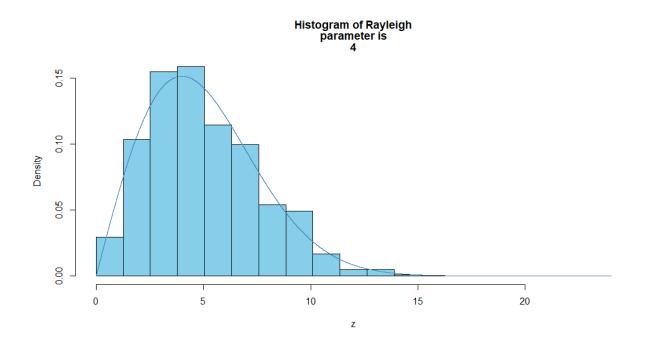
即令
$$Z=\sqrt{X^2+Y^2}, \underline{X,Y}\sim N(0,\sigma^2)$$
 且独立,则 $Z\sim \mathsf{Rayleigh}(\mathsf{G})$ 分布,

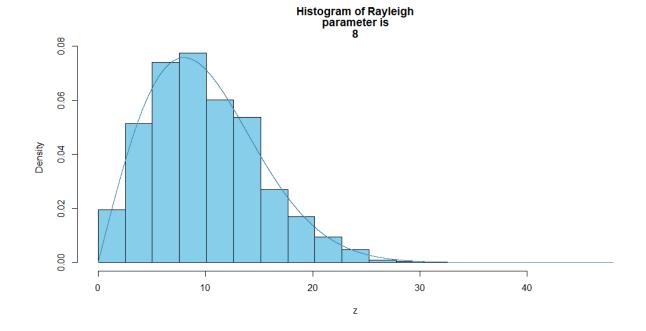
```
# 选取一系列sigma
sigma<-c(1,2,4,8,16,32)
for (i in 1:length(sigma)) {
  #设置种子以保证伪随机数的一致性
  set.seed(i)
  # 直方图的标题
  title<-c("Histogram of Rayleigh", "parameter is", sigma[i])</pre>
  # 生成两个正态分布
  x<-rnorm(1000,0,sigma[i])</pre>
  y<-rnorm(1000,0,sigma[i])</pre>
  # 生成瑞利分布的随机数
  z < -sqrt(x^2+y^2)
  #绘制并检查
  hist(z,prob=TRUE,breaks = seq(0,6*sigma[i],length.out = 20)
       ,main = title,col = "skyblue")
  # 绘制Rayleigh 密度函数
  x1 < -seq(0,6*sigma[i],length.out = 100000)
  y1<-(x1/sigma[i]^2)*exp(-(x1^2)/(2*sigma[i]^2))
  lines(x1,y1,col="steelblue")
}
```

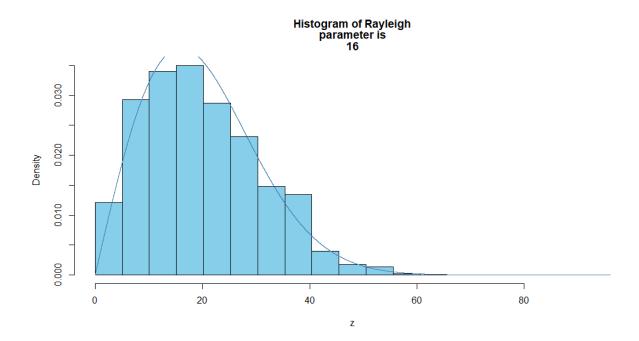
₩ 结果

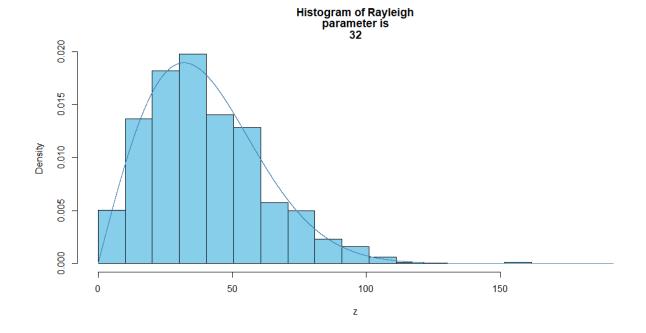
以下分别为不同取值σ对应直方图和密度曲线图











四 结果分析

样本密度直方图和密度曲线大致吻合,说明生成效果较好。

H3 3-9

四 分析过程

令 $y=rac{1+x}{2},f(y)=rac{3}{4}(rac{3}{4}-rac{1}{4}y^2-rac{1}{2}y)$,可见 $Y\sim Beta(2,2)$,可以利用该变换方便的生成理论直方图和概率密度图像。

另一方面利用题目所给算法也可生成直方图和概率密度图像,二者可进行直观的比较。

₩ 代码

```
#### 3.9 ####

n <- 1000 # 随机数个数

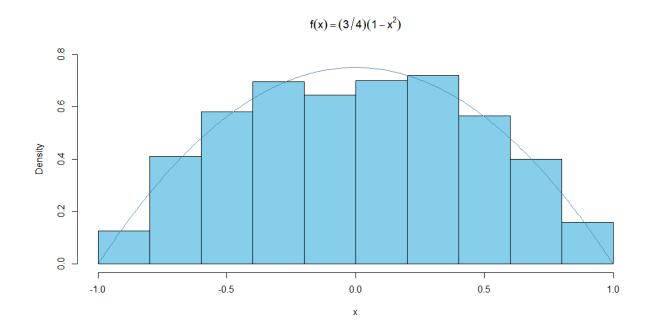
y <- rbeta(n,2,2) #令Y=(X+1)/2, Y~Be(2,2),由Y产生随机数

x <- 2*y-1 #将Y产生的随机数结果回代

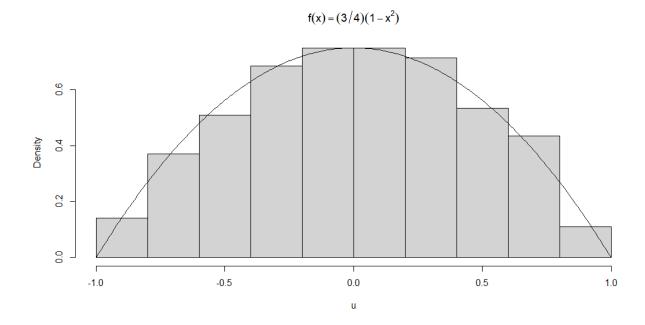
hist(x, prob = TRUE,main = expression(f(x)==(3/4)(1-x^2)),col = "skyblue") #
由此得到直方图
```

₩ 结果

理论图像:



算法生成图像:



四 结果分析

样本密度直方图和密度曲线大致吻合,而且与理论直方图图像较为接近,说明生成效果较好。

E3 3-11

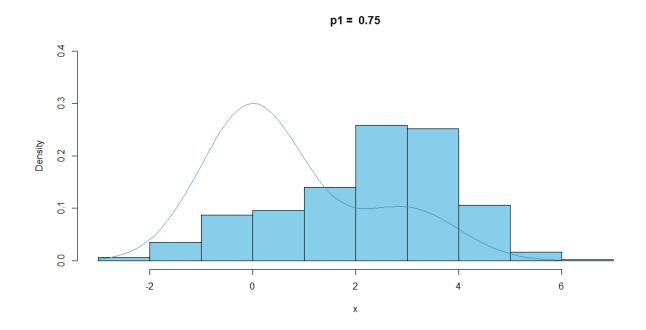
四 分析过程

首先生成 $p\sim Bio(0.75), X=pX_1+(1-p)X_2, \ X_1\sim N(0,1), X_2\sim N(3,1)$,为了探究p对双峰的影响,可以写一个把p当成参数的函数。

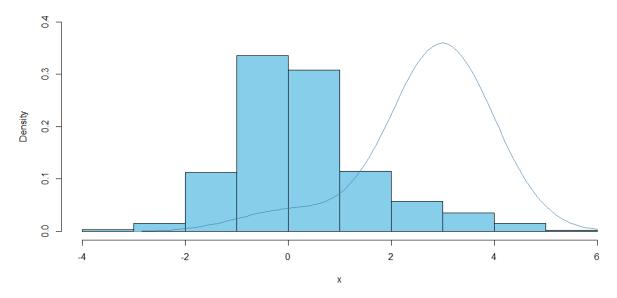
```
#### 3.11 ####
mixturehist<-function(p1){
set.seed(1012)
p<-rbinom(1000,1,prob = p1)
x1<-rnorm(1000)
x<-p*x1+(1-p)*x2 #Generate samples
title<-paste('p1 = ',p1)
hist(x,probability = T,main = title,col = "skyblue")</pre>
```

```
y<-seq(-3,6,0.1)
densityplot<-function(x){
    p1*dnorm(x)+(1-p1)*dnorm(x,3,1)
    }
lines(y,densityplot(y),col='steelblue')
}
density (densityplot(y)
mixturehist(0.75)
for (p1 in seq(0.1,0.9,0.1)){
    mixturehist(p1)
}</pre>
```

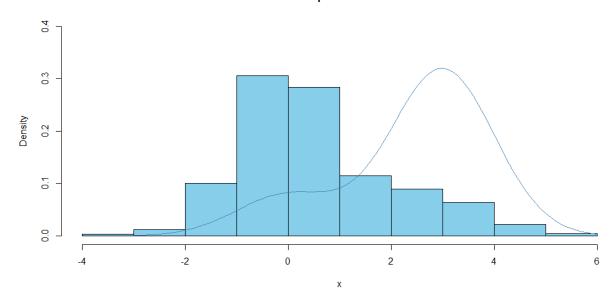
₩ 结果



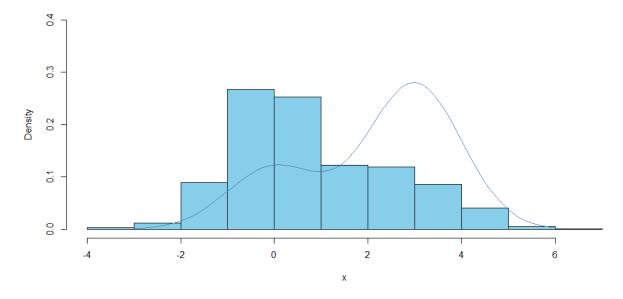




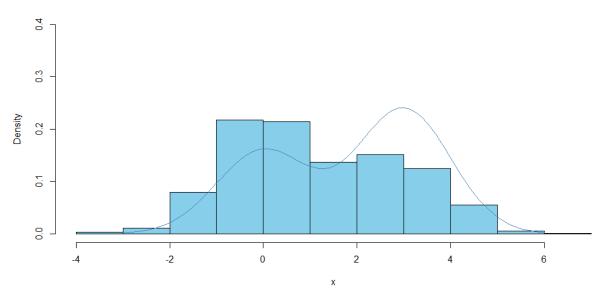


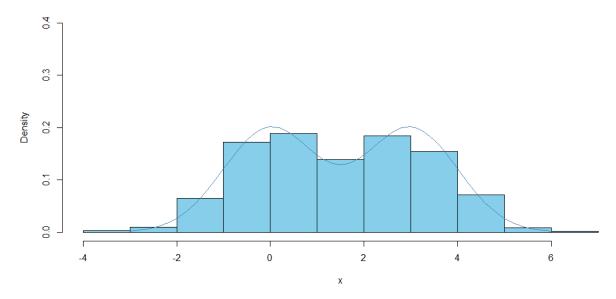




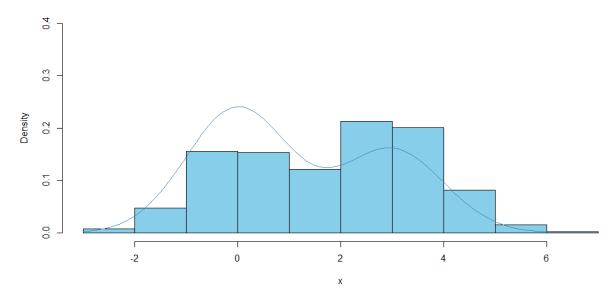


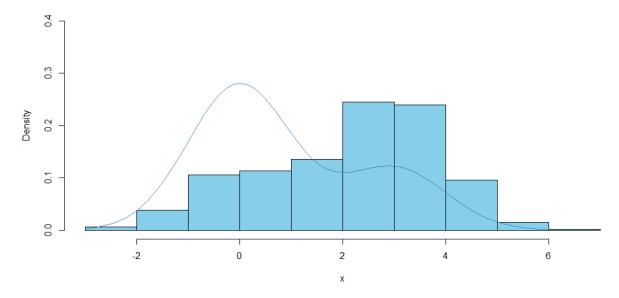




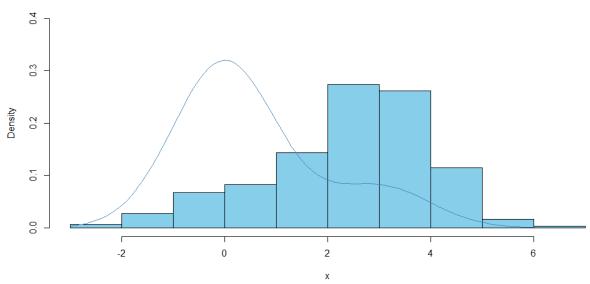


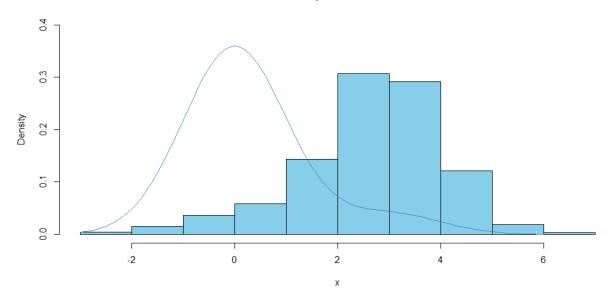












四 结果分析

当 $p \in (0.2, 0.8)$ 时会出现双峰现象

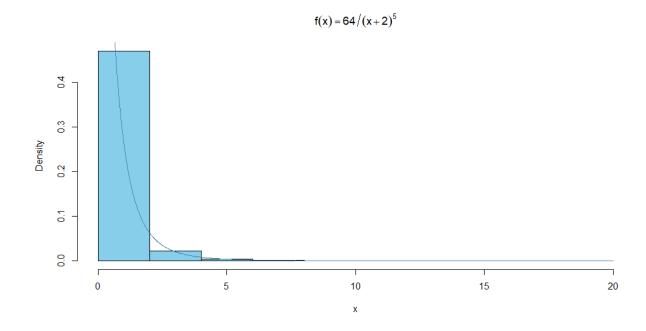
E3-13

四 分析过程

利 用 逆 变 换 方 法 可 以 进 行 模 拟 , 由 题 意 $r=4,\beta=2, F(y)=1-(\frac{2}{2+y})^4, F^{-1}(u)=(\frac{2}{1-u})^{1/4}, f(y)=\frac{64}{(y+2)^5}$

```
n <- 1000
u <- runif(n)
x <- (2/(1-u)^(1/4))-2
hist(x, prob = TRUE, main = bquote(f(x)==64/(x+2)^5),col = "skyblue") # 求导得密度函数
y <- seq(0, 20, 0.01) # 得到密度函数曲线
lines(y, 64/(y+2)^5,col="steelblue")
```

四结果



四 结果分析

样本密度直方图和密度曲线大致吻合,说明生成效果较好。

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四 分析过程

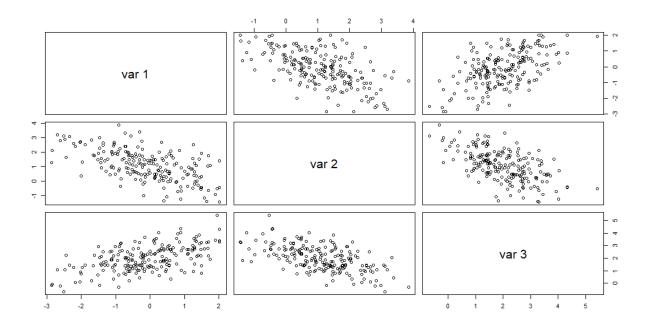
这题还需要分析吗

```
mu=c(0,1,2)
Sigma=matrix(c(1,-.5,.5,-.5,1,-.5,.5,-.5,1),ncol = 3,byrow = F)
rmvn.Choleski <-
function(n, mu, Sigma) {
    # generate n random vectors from MVN(mu, Sigma)
    # dimension is inferred from mu and Sigma
    d <- length(mu)</pre>
```

```
Q <- chol(Sigma) # Choleski factorization of Sigma
Z <- matrix(rnorm(n*d), nrow=n, ncol=d)
X <- Z %*% Q + matrix(mu, n, d, byrow=TRUE)
X
}

X <- rmvn.Choleski(200, mu, Sigma)
# 绘图
pairs(X)
```

四 结果



四 结果分析

上图所示的样本数据散点图展示了多元正态分布的椭圆对称性,说明位置和相关性与相应二元正态分布的理论参数大致一致。

E3 3-16

四 分析过程

按考试类型标准化,再计算协方差矩阵。

四 代码

```
#### 3.16 ###
library(bootstrap)# 没有就先下载
cov(scale(scor[,1:2]))
cov(scale(scor[,3:5]))
```

四 结果

```
# 闭卷

mec vec

alg ana sta

mec 1.0000000 0.5534052

vec 0.5534052 1.0000000 ana 0.7108059 1.0000000 0.6071743

sta 0.6647357 0.6071743 1.0000000
```

3-20

四 分析过程

利 用 书 上 给 的 算 法 生 成 $N(t)\sim Poss(\lambda), Then$ $Y_i\sim Gamma(shape,scale), X_i=\sum_{i=1}^{N(t)}Y_i$,之后再与理论值的期望和方差作比较

```
#### 3.20 ####
# shape: Gamma分布形状参数; scale:Gamma分布尺度参数
comp_poss <- function(lambda, shape, scale, size = 1000 ,t = 10) {</pre>
 # 到达间隔时间随速率λ呈指数分布。
  pp.exp = function (t0) {
   Tn = rexp(1000, lambda)
   Sn = cumsum(Tn)
   return(min(which(Sn > t0)) - 1)
 }
 # 生成服从泊松分布的N (t)
 ns = replicate(size, expr={ pp.exp(t)})
 # 生成题目描述的X(t)
 xs = sapply(ns, function (n) {
   ys = c(rgamma(n = n, shape = shape, scale = scale))
   sum(ys[1:n])
 })
 # 计算模拟值和理论值的差别
 # 样本
 mean.s = mean(xs)
 var.s = var(xs)
 # 理论
 mean.t = lambda * t * shape * scale
 var.t = (shape + 1) * shape * scale^2*lambda*t
 df = matrix(c(mean.s,mean.t,var.s,var.t),ncol = 4,
             dimnames =
list(c("value"),c("mean.s","mean.t","var.s","var.t")))
}
comp_poss(3,5,4)
```

四 结果

```
# 3 5 4

mean.s mean.t var.s var.t

value 595.898 600 14881.48 14400

# 1 2 3

mean.s mean.t var.s var.t

value 59.80842 60 511.6086 540

# 4 5 6

mean.s mean.t var.s var.t

value 1203.058 1200 42209.32 43200
```

咝 结果分析

结果显示理论值和样本值的期望和方差都很接近,说明生成效果良好。