Introduction to Multiple Linear Regression

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Today's lecture

- Multiple Linear Regression
 - Assumptions
 - Interpretation
 - Notation

Motivation

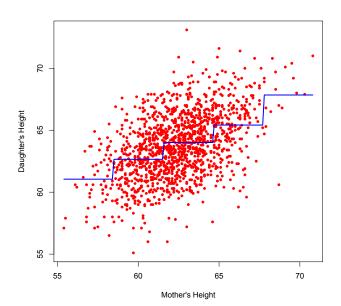
Most applications involve more that one covariate – if more than one thing can influence an outcome, you need multiple linear regression.

- Improved description of $y | \mathbf{x}$
- More accurate estimates and predictions
- Allow testing of multiple effects
- Includes multiple predictor types

Why not bin all predictors?

- Divide x_i into k_i bins
- Stratify data based on inclusion in bins across x's
- Find mean of the y_i in each category
- Possibly a reasonable non-parametric model

Why not bin all predictors?



Why not bin all predictors?

- More predictors = more bins
- If each x has 5 bins, you have 5^p overall categories
- May not have enough data to estimate distribution in each category
- Curse of dimensionality is a problem in a lot of non-parametric statistics

Multiple linear regression model

■ Observe data $(y_i, x_{i1}, ..., x_{ip})$ for subjects 1, ..., n. Want to estimate $\beta_0, \beta_1, ..., \beta_p$ in the model

$$y_i = \beta_0 + \beta_1 x_{i1} + \ldots + \beta_1 x_{ip} + \epsilon_i; \ \epsilon_i \stackrel{iid}{\sim} (0, \sigma^2)$$

- Assumptions (residuals have mean zero, constant variance, are independent) are as in SLR
- Impose linearity which (as in the SLR) is a big assumption
- Our primary interest will be $E(y|\mathbf{x})$
- Eventually estimate model parameters using least squares

Predictor types

- Continuous
- Categorical
- Ordinal

Interpretation of coefficients

$$\beta_0 = E(y|x_1 = 0, \dots, x = 0)$$

• Centering some of the x's may make this more interpretable

Interpretation of β_1

Example with two predictors

Suppose we want to regress weight on height and sex.

- Model is $y_i = \beta_0 + \beta_1 x_{i,age} + \beta_2 x_{i,sex} + \epsilon_i$
- Age is continuous starting with age 0; sex is binary, coded so that $x_{i,sex} = 0$ for men and $x_{i,sex} = 1$ for women

Example with two predictors

$$\beta_1 =$$

$$\beta_2 =$$

Omitted variable bias

What happens if the true regression model is

$$y_i = \beta_0 + \beta_1 x_{i,1} + \beta_2 x_{i,2} + \epsilon_i$$

but we ignore x_2 and fit the simple linear regression

$$y_i = \beta_0^* + \beta_1^* x_{i,1} + \epsilon_i^*$$

Does $\beta_1^* = \beta_1$?

Omitted variable bias

When should you be concerned?

If both of the following conditions are met, then $\beta_1^* = \beta_1$:

- The omitted variable is unrelated to the outcome
- The omitted variable is uncorrelated with the retained variable

Note: A Simpson's paradox can be explained by ommitted variable bias.

Matrix notation

■ Observe data $(y_i, x_{i1}, \dots, x_{ip})$ for subjects $1, \dots, n$. Want to estimate $\beta_0, \beta_1, \dots, \beta_p$ in the model

$$y_i = \beta_0 + \beta_1 x_{i1} + \ldots + \beta_1 x_{ip} + \epsilon_i; \ \epsilon_i \stackrel{iid}{\sim} (0, \sigma^2)$$

- Notation is cumbersome. To fix this, let
 - $\mathbf{x}_i = [1, x_{i1}, \dots, x_{ip}]$
 - $\beta^T = [\beta_0, \beta_1, \dots, \beta_p]$
 - Then $y_i = \mathbf{x}_i \boldsymbol{\beta} + \epsilon_i$

Multiple linear regressoion

Let

$$\mathbf{y} = \left[\begin{array}{c} y_1 \\ \vdots \\ y_n \end{array} \right], \quad \mathbf{X} = \left[\begin{array}{ccc} 1 & x_{11} & \dots & x_{1\rho} \\ \vdots & \vdots & & \vdots \\ 1 & x_{n1} & \dots & x_{n\rho} \end{array} \right], \quad \boldsymbol{\beta} = \left[\begin{array}{c} \beta_0 \\ \vdots \\ \beta_p \end{array} \right], \quad \boldsymbol{\epsilon} = \left[\begin{array}{c} \epsilon_1 \\ \vdots \\ \epsilon_n \end{array} \right]$$

■ Then we can write the model in a more compact form:

$$\mathbf{y}_{n imes 1} = \mathbf{X}_{n imes (p+1)} oldsymbol{eta}_{(p+1) imes 1} + \epsilon_{n imes 1}$$

■ **X** is called the *design matrix*

Matrix notation

$$y = X\beta + \epsilon$$

- ullet ϵ is a random vector rather than a random variable
- $E(\epsilon) = 0$ and $Cov(\epsilon) = \sigma^2 I$
- Note that *Cov* means the "variance-covariance matrix"

Mean, variance and covariance of a random vector

Let $\mathbf{y}^T = [y_1, \dots, y_n]$ be an *n*-component random vector. Then its mean and variance are defined as

$$E(\mathbf{y})^T = [E(y_1), \dots, E(y_n)]$$

 $Var(\mathbf{y}) = E[(\mathbf{y} - E\mathbf{y})(\mathbf{y} - E\mathbf{y})^T] = E(\mathbf{y}\mathbf{y}^T) - (E\mathbf{y})(E\mathbf{y})^T$

■ Let \mathbf{y} and \mathbf{z} be an n-component and an m-component random vector respectively. Then their covariance is an $n \times m$ matrix defined by

$$Cov(\mathbf{y}, \mathbf{z}) = E[(\mathbf{y} - E\mathbf{y})(\mathbf{z} - \mathbf{z})^T]$$

Coming up next...

Multiple linear regression models

- estimation (more least squares)
- more detailed model diagnostics
- ▶ inference