



Variational Inference

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LECTURE 21

Roadmap

- Big-picture questions
- VI for LDA
- More content questions
- Walkthrough of VI for LDA (HW)

Big picture content questions

Big picture content questions

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Deriving Variational Inference for LDA

Joint distribution:

$$p(\theta, z, w \mid \alpha, \beta) = \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \prod_d \left[\prod_k \theta_{d,k}^{\alpha_k - 1} \left(\prod_n \prod_i \prod_j (\theta_{d,i} \beta_{i,j})^{w_{d,n}^j} \right) \right] \quad (1)$$

Deriving Variational Inference for LDA

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Variational distribution:

$$q(\theta, z) = q(\theta \mid \gamma) q(z \mid \phi) \quad (2)$$

Deriving Variational Inference for LDA

Joint distribution:

$$p(\theta, z, w \mid \alpha, \beta) = \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \prod_d \left[\prod_k \theta_{d,k}^{\alpha_k - 1} \left(\prod_n \prod_i \prod_j (\theta_{d,i} \beta_{i,j})^{w_{d,n}^j} \right) \right] \quad (1)$$

Variational distribution:

$$q(\theta, z) = q(\theta \mid \gamma) q(z \mid \phi) \quad (2)$$

ELBO:

$$\begin{aligned} L(\gamma, \phi; \alpha, \beta) = & \mathbb{E}_q [\log p(\theta \mid \alpha)] + \mathbb{E}_q [\log p(z \mid \theta)] + \mathbb{E}_q [\log p(w \mid z, \beta)] \\ & - \mathbb{E}_q [\log q(\theta)] - \mathbb{E}_q [\log q(z)] \end{aligned} \quad (3)$$

Expectation of log Dirichlet

- Most expectations are straightforward to compute
- Dirichlet is harder

$$\mathbb{E}_{\text{dir}} [p(\theta_i | \alpha)] = \Psi(\alpha_i) - \Psi\left(\sum_j \alpha_j\right) \quad (4)$$

Expectation 1

$$\mathbb{E}_q [\log p(\theta \mid \alpha)] = \mathbb{E}_q \left[\log \left\{ \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \prod_i \theta_i^{\alpha_i - 1} \right\} \right] \quad (5)$$

(6)

Expectation 1

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$$= \mathbb{E}_q \left[\log \left\{ \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \right\} + \sum_i \log \theta_i^{\alpha_i - 1} \right] \quad (6)$$

Log of products becomes sum of logs.

Expectation 1

$$\mathbb{E}_q [\log p(\theta | \alpha)] = \mathbb{E}_q \left[\log \left\{ \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \prod_i \theta_i^{\alpha_i - 1} \right\} \right] \quad (5)$$

$$\begin{aligned} &= \mathbb{E}_q \left[\log \left\{ \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \right\} + \sum_i \log \theta_i^{\alpha_i - 1} \right] \\ &= \log \Gamma(\sum_i \alpha_i) - \sum_i \log \Gamma(\alpha_i) + \mathbb{E}_q \left[\sum_i (\alpha_i - 1) \log \theta_i \right] \end{aligned} \quad (6)$$

Log of exponent becomes product, expectation of constant is constant

Expectation 1

$$\begin{aligned}\mathbb{E}_q [\log p(\theta | \alpha)] &= \mathbb{E}_q \left[\log \left\{ \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \prod_i \theta_i^{\alpha_i - 1} \right\} \right] \\ &= \mathbb{E}_q \left[\log \left\{ \frac{\Gamma(\sum_i \alpha_i)}{\prod_i \Gamma(\alpha_i)} \right\} + \sum_i \log \theta_i^{\alpha_i - 1} \right] \\ &= \log \Gamma(\sum_i \alpha_i) - \sum_i \log \Gamma(\alpha_i) + \mathbb{E}_q \left[\sum_i (\alpha_i - 1) \log \theta_i \right] \\ &= \log \Gamma(\sum_i \alpha_i) - \sum_i \log \Gamma(\alpha_i) \\ &\quad + \sum_i (\alpha_i - 1) \left(\Psi(\gamma_i) - \Psi \left(\sum_j \gamma_j \right) \right)\end{aligned}\tag{5}$$

Expectation of log Dirichlet

Expectation 2

$$\mathbb{E}_q [\log p(z \mid \theta)] = \mathbb{E}_q \left[\log \prod_n \prod_i \theta_i^{\mathbb{1}[z_n=i]} \right] \quad (6)$$

(7)

Expectation 2

$$\mathbb{E}_q [\log p(z \mid \theta)] = \mathbb{E}_q \left[\log \prod_n \prod_i \theta_i^{\mathbb{1}[z_n=i]} \right] \quad (6)$$

$$= \mathbb{E}_q \left[\sum_n \sum_i \log \theta_i^{\mathbb{1}[z_n=i]} \right] \quad (7)$$

$$(8)$$

Products to sums

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$$\mathbb{E}_q [\log p(z \mid \theta)] = \mathbb{E}_q \left[\log \prod_n \prod_i \theta_i^{\mathbb{1}[z_n=i]} \right] \quad (6)$$

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$$(9)$$

Linearity of expectation

Expectation 2

$$\mathbb{E}_q [\log p(z | \theta)] = \mathbb{E}_q \left[\log \prod_n \prod_i \theta_i^{\mathbb{1}[z_n == i]} \right] \quad (6)$$

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$$= \sum_n \sum_i \mathbb{E}_q \left[\log \theta_i^{\mathbb{1}[z_n == i]} \right] \quad (8)$$

$$= \sum_n \sum_i \phi_{ni} \mathbb{E}_q [\log \theta_i] \quad (9)$$

$$(10)$$

Independence of variational distribution, exponents become products

Expectation 2

$$\mathbb{E}_q [\log p(z \mid \theta)] = \mathbb{E}_q \left[\log \prod_n \prod_i \theta_i^{\mathbb{1}[z_n=i]} \right] \quad (6)$$

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$$= \sum_n \sum_i \phi_{ni} \mathbb{E}_q [\log \theta_i] \quad (9)$$

$$= \sum_n \sum_i \phi_{ni} \left(\Psi(\gamma_i) - \Psi \left(\sum_j \gamma_j \right) \right) \quad (10)$$

Expectation of log Dirichlet

Complete objective function

$$\begin{aligned} L(\gamma, \phi; \alpha, \beta) = & \log \Gamma \left(\sum_{j=1}^k \alpha_j \right) - \sum_{i=1}^k \log \Gamma(\alpha_i) + \sum_{i=1}^k (\alpha_i - 1) \left(\Psi(\gamma_i) - \Psi \left(\sum_{j=1}^k \gamma_j \right) \right) \\ & + \sum_{n=1}^N \sum_{i=1}^k \phi_{ni} \left(\Psi(\gamma_i) - \Psi \left(\sum_{j=1}^k \gamma_j \right) \right) \\ & + \sum_{n=1}^N \sum_{i=1}^k \sum_{j=1}^V \phi_{ni} w_n^j \log \beta_{ij} \\ & - \log \Gamma \left(\sum_{j=1}^k \gamma_j \right) + \sum_{i=1}^k \log \Gamma(\gamma_i) - \sum_{i=1}^k (\gamma_i - 1) \left(\Psi(\gamma_i) - \Psi \left(\sum_{j=1}^k \gamma_j \right) \right) \\ & - \sum_{n=1}^N \sum_{i=1}^k \phi_{ni} \log \phi_{ni}, \end{aligned}$$

Note the entropy terms at the end (negative sign)

Deriving the algorithm

- Compute partial wrt to variable of interest
- Set equal to zero
- Solve for variable

Update for ϕ

Derivative of ELBO:

$$\frac{\partial L}{\partial \phi_{ni}} = \Psi(\gamma_i) - \Psi\left(\sum_{j=1}^k \gamma_j\right) + \log \beta_{iv} - \log \phi_{ni} - 1 + \lambda.$$

Solution:

$$\phi_{ni} \propto \beta_{iv} \exp \left(\Psi(\gamma_i) - \Psi \left(\sum_j \gamma_j \right) \right) \quad (11)$$

Update for γ

Derivative of ELBO:

$$\frac{\partial L}{\partial \gamma_i} = \Psi'(\gamma_i) (\alpha_i + \sum_{n=1}^N \phi_{ni} - \gamma_i) - \Psi'(\sum_{j=1}^k \gamma_j) \sum_{j=1}^k (\alpha_j + \sum_{n=1}^N \phi_{nj} - \gamma_j).$$

Solution:

$$\gamma_i = \alpha_i + \sum_{n=1}^N \phi_{ni}.$$

Update for β

Slightly more complicated (requires Lagrange parameter), but solution is obvious:

$$\beta_{ij} \propto \sum_d \sum_n \phi_{dni} w_{dn}^j \quad (12)$$

Detail content questions

Detail content questions

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Administrivia

- End of flipped classroom
 - Short content session at start
 - Use the time to meet with teammates
- First deliverable soon!

Example

- Three topics, same documents as last time

$$\beta = \begin{bmatrix} \text{cat} & \text{dog} & \text{hamburger} & \text{iron} & \text{pig} \\ .26 & .185 & .185 & .185 & .185 \\ .185 & .185 & .26 & .185 & .185 \\ .185 & .185 & .185 & .26 & .185 \end{bmatrix} \quad (13)$$

- Assume uniform γ : (2.0, 2.0, 2.0)
- Compute update for ϕ

$$\phi_{ni} \propto \beta_{iv} \exp \left(\Psi(\gamma_i) - \Psi \left(\sum_j \gamma_j \right) \right) \quad (14)$$

- For the first word (dog) in the document: **dog cat cat pig**

Example

- Three topics, same documents as last time

$$\beta = \begin{bmatrix} \text{cat} & \text{dog} & \text{hamburger} & \text{iron} & \text{pig} \\ .26 & .185 & .185 & .185 & .185 \\ .185 & .185 & .26 & .185 & .185 \\ .185 & .185 & .185 & .26 & .185 \end{bmatrix} \quad (13)$$

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- For a the first word (dog) in the document: **dog cat cat pig**

Update ϕ for dog

- $\gamma = (2.000, 2.000, 2.000)$

Update ϕ for dog

- $\gamma = (2.000, 2.000, 2.000)$
- $\phi(0) \propto 0.185 \times \exp(\Psi(2.000) - \Psi(2.000 + 2.000 + 2.000)) = 0.051$

Update ϕ for dog

- $\gamma = (2.000, 2.000, 2.000)$
- $\phi(0) \propto 0.185 \times \exp(\Psi(2.000) - \Psi(2.000 + 2.000 + 2.000)) = 0.051$
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- After normalization: $\{0.333, 0.333, 0.333\}$

Update ϕ for pig

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- After normalization: $\{0.333, 0.333, 0.333\}$

Update ϕ for cat

- $\gamma = (2.000, 2.000, 2.000)$

Update ϕ for cat

- $\gamma = (2.000, 2.000, 2.000)$
- $\phi(0) \propto 0.260 \times \exp(\Psi(2.000) - \Psi(2.000 + 2.000 + 2.000)) = 0.072$

Update ϕ for cat

- $\gamma = (2.000, 2.000, 2.000)$
- $\phi(0) \propto 0.260 \times \exp(\Psi(2.000) - \Psi(2.000 + 2.000 + 2.000)) = 0.072$
- $\phi(0) \propto 0.185 \times \exp(\Psi(2.000) - \Psi(2.000 + 2.000 + 2.000)) = 0.051$

Update ϕ for cat

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- $\phi(0) \propto 0.185 \times \exp(\Psi(2.000) - \Psi(2.000 + 2.000 + 2.000)) = 0.051$
- $\phi(0) \propto 0.185 \times \exp(\Psi(2.000) - \Psi(2.000 + 2.000 + 2.000)) = 0.051$
- After normalization: $\{0.413, 0.294, 0.294\}$

Update γ

- Document: dog cat cat pig
- Update equation

$$\gamma_i = \alpha_i + \sum_n \phi_{ni} \quad (15)$$

- Assume $\alpha = (.1, .1, .1)$

Update γ

- Document: dog cat cat pig
- Update equation

$$\gamma_i = \alpha_i + \sum_n \phi_{ni} \quad (15)$$

- Assume $\alpha = (.1, .1, .1)$

	ϕ_0	ϕ_1	ϕ_2
dog	.333	.333	.333
cat	.413	.294	.294
pig	.333	.333	.333
α	0.1	0.1	0.1
sum	1.592	1.354	1.354

- Note: **do not normalize!**

Update γ

- Document: dog cat cat pig
- Update equation

$$\gamma_i = \alpha_i + \sum_n \phi_{ni} \quad (15)$$

- Assume $\alpha = (.1, .1, .1)$

	ϕ_0	ϕ_1	ϕ_2	
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cat	.413	.294	.294	x2
pig	.333	.333	.333	
α	0.1	0.1	0.1	
sum	1.592	1.354	1.354	

- Note: **do not normalize!**

Update β

- Count up all of the ϕ across all documents
- For each topic, divide by total
- Corresponds to maximum likelihood of expected counts

Update β

- Count up all of the ϕ across all documents
- For each topic, divide by total
- Corresponds to maximum likelihood of expected counts
- Unlike Gibbs sampling, no Dirichlet prior