Stochastic Gradient Descent in NPIV estimation

Anonymous Author(s)

Affiliation Address email

Abstract

TODO

2 1 Problem setup

3 1.1 Basic definitions

4 Fix a probability space $(\Omega, \mathcal{A}, \mathbb{P})$. Given $X \in L^2(\Omega; \mathbb{X} \subseteq \mathbf{R}^p)$, we define

$$L^2(X) \triangleq \left\{ h : \mathbb{X} \to \mathbf{R} : \mathbb{E}[h(X)^2] < \infty \right\},\,$$

- that is, $L^2(X) = L^2(X, \mathcal{B}(X), \mathbb{P}_X)^1$, a Hilbert space equipped with the inner product $\langle h, g \rangle_{L^2(X)} =$
- 6 $\mathbb{E}[h(X)g(X)]$. The regression problem we are interested in has the form

$$Y = h^{\star}(X) + \varepsilon, \tag{1}$$

- where $h^* \in L^2(X)$ and ε is an integrable r.v. such that $\mathbb{E}[\varepsilon \mid X] \neq 0$. We assume there exists
- 8 $Z \in L^2(\Omega; \mathbb{Z} \subseteq \mathbf{R}^q)$ such that $Z \not\perp \!\!\! \perp X$ and $\mathbb{E}[\varepsilon \mid Z] = 0$. This variable is called the instrumental
- 9 variable. The problem consists of estimating h^* based on independent joint samples from X, Z and
- 11 Conditioning (1) in Z, we find

$$\mathbb{E}[Y \mid Z] = \mathbb{E}[h^*(X) \mid Z]. \tag{2}$$

12 This motivates us to introduce the operator $\mathcal{T}:L^2(X) o L^2(Z)$ defined by

$$T[h](z) \triangleq \mathbb{E}[h(X) \mid Z = z].$$

- Clearly \mathcal{T} is linear and, using Jensen's inequality, one may prove that it's bounded. It's also interesting
- to notice that its adjoint $\mathcal{T}^*:L^2(Z)\to L^2(X)$ satisfies

$$\mathcal{T}^*[g](x) = \mathbb{E}[g(Z) \mid X = x]. \tag{3}$$

- Define $r_0: \mathbb{Z} \to \mathbf{R}$ by $r_0(Z) = \mathbb{E}[Y \mid Z]$. Again by Jensen's inequality, we have $r_0 \in L^2(Z)$, and
- thus we can rewrite (2) as

$$T[h^*] = r_0. (4)$$

Hence, (1) can be formulated as an inverse problem, where we wish to invert the operator \mathcal{T} .

18 1.2 Risk measure

- 19 Let $\ell: \mathbf{R} \times \mathbf{R} \to \mathbf{R}_+$ be a pointwise loss function, which, with respect to its second argument, is
- 20 convex and differentiable. We use the symbol ∂_2 to denote a derivative with respect to the second

¹We denote by \mathbb{P}_X the distribution of the r.v. X and by $\mathcal{B}(\mathbb{X})$ the Borel σ -algebra in \mathbb{X} .

- argument. The example to keep in mind is the quadratic loss function $\ell(y,y')=(y-y')^2$. Given
- 22 $h \in L^2(X)$, we define the *populational risk* associated with it to be

$$\mathcal{R}(h) \triangleq \mathbb{E}[\ell(r_0(Z), \mathcal{T}[h](Z))].$$

23 We would like to solve

$$\inf_{h \in \mathcal{F}} \mathcal{R}(h),$$

where $\mathcal{F} \subseteq L^2(X)$ is a subspace such that $h^* \in \mathcal{F}$.

25 **2 Gradient computation**

We'd like to compute $\nabla \mathcal{R}(h)$ for $h \in L^2(X)$. We start by computing the directional derivative of \mathcal{R}

27 at h in the direction f, denoted by $D\mathcal{R}[h](f)$:

$$\begin{split} D\mathcal{R}[h](f) &= \lim_{\delta \to 0} \frac{1}{\delta} \left[\mathcal{R}(h + \delta f) - \mathcal{R}(f) \right] \\ &= \lim_{\delta \to 0} \frac{1}{\delta} \mathbb{E} \left[\ell(r_0(Z), \mathcal{T}[h + \delta f](Z)) - \ell(r_0(Z), \mathcal{T}[h](Z)) \right] \\ &= \lim_{\delta \to 0} \frac{1}{\delta} \mathbb{E} \left[\ell(r_0(Z), \mathcal{T}[h](Z) + \delta \mathcal{T}[f](Z)) - \ell(r_0(Z), \mathcal{T}[h](Z)) \right] \\ &= \lim_{\delta \to 0} \frac{1}{\delta} \mathbb{E} \left[\delta \partial_2 \ell(r_0(Z), \mathcal{T}[h](Z)) \mathcal{T}[f](Z) \right. \\ &\left. + \frac{\delta^2}{2} \partial_2^2 \ell(r_0(Z), \mathcal{T}[h + \theta f](Z)) \mathcal{T}[f](Z)^2 \right] \\ &= \mathbb{E} \left[\partial_2 \ell(r_0(Z), \mathcal{T}[h](Z)) \mathcal{T}[f](Z) \right] \\ &+ \lim_{\delta \to 0} \mathbb{E} \left[\frac{\delta}{2} \partial_2^2 \ell(r_0(Z), \mathcal{T}[h + \theta f](Z)) \mathcal{T}[f](Z)^2 \right] \\ &= \mathbb{E} \left[\partial_2 \ell(r_0(Z), \mathcal{T}[h](Z)) \mathcal{T}[f](Z) \right], \end{split}$$

where $\theta \in \mathbf{R}$ is due to Taylor's formula and can be assumed to be inside a fixed interval $(-\theta_0, \theta_0)$, Assumption with θ are historially small. We have assumed at the last start that there exists $\theta > 0$ such that

with θ_0 arbitrarily small. We have assumed at the last step that there exists $\theta_0 > 0$ such that

$$\sup_{|\theta|<\theta_0} \mathbb{E}\left[\partial_2^2 \ell(r_0(Z), \mathcal{T}[h+\theta f](Z)) \mathcal{T}[f](Z)^2\right] < \infty.$$

- 30 This is a mild integrability condition which can be shown to hold in the quadratic case.
- We can in fact expand the calculation a bit more, as follows:

$$D\mathcal{R}[h](f) = \mathbb{E}\left[\partial_2 \ell(r_0(Z), \mathcal{T}[h](Z))\mathcal{T}[f](Z)\right]$$

= $\langle \partial_2 \ell(r_0(\cdot), \mathcal{T}[h](\cdot)), \mathcal{T}[f] \rangle_{L^2(Z)}$
= $\langle \mathcal{T}^*[\partial_2 \ell(r_0(Z), \mathcal{T}[h](\cdot))], f \rangle_{L^2(X)},$

- where we are assuming that $\partial_2 \ell(r_0(\cdot), \mathcal{T}[h](\cdot)) \in L^2(Z)$. This shows that \mathcal{R} is Gateux-differentiable, \square Assumption
- with Gateux derivative at h given by

$$D\mathcal{R}[h] = \mathcal{T}^*[\partial_2 \ell(r_0(\cdot), \mathcal{T}[h](\cdot))].$$

- If we assume² that $h \mapsto D\mathcal{R}[h]$ is a continuous mapping from $L^2(Z)$ to $L^2(Z)$, then \mathcal{R} is also
- Fréchet-differentiable, and both derivatives coincide. Therefore, under this assumption, which we
- henceforth make, $\nabla \mathcal{R}(h) = \mathcal{T}^*[\partial_2 \ell(r_0(\cdot), \mathcal{T}[h](\cdot))].$

Talk about which conditions ℓ can satisfy so that this is continuous.

Assumption

Assumption

3 Unbiased estimator of the gradient

38 We have found that

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$$\nabla \mathcal{R}(h)(x) = \mathcal{T}^*[\partial_2 \ell(r_0(\cdot), \mathcal{T}[h](\cdot))](x) = \mathbb{E}[\partial_2 \ell(r_0(Z), \mathcal{T}[h](Z)) \mid X = x].$$

- 39 This turns out to be hard to estimate in practice, as we have two nested conditional expectation
- operators. Our objective in this section is to find a random element $u_h \in L^2(X)$ such that $\mathbb{E}[u_h(x)] = 1$
- $\nabla \mathcal{R}(h)(x)$, so we can replace $\nabla \mathcal{R}(h)(x)$ by $u_h(x)$ in a gradient descent algorithm, obtaining a

stochastic version which will be easier to compute.

- Our strategy to obtain u_h will be to write $\nabla \mathcal{R}(h)(x) = \mathbb{E}[\Phi(x,Z)\partial_2\ell(r_0(Z),\mathcal{T}[h](Z))]$, for some
- suitable kernel Φ . To ease the notation, define $\xi_h(z) \triangleq \partial_2 \ell(r_0(z), \mathcal{T}[h](z))$. Assuming that X and
- Z have a joint distribution which is absolutely continuous with respect to Lebesgue measure in \mathbf{R}^{p+q} ,
- 46 we can write

$$\nabla \mathcal{R}(h)(x) = \mathbb{E}[\xi_h(Z) \mid X = x]$$

$$= \int_{\mathbb{Z}} p(z \mid x) \xi_h(z) \, dz$$

$$= \int_{\mathbb{Z}} p(z) \frac{p(z \mid x)}{p(z)} \xi_h(z) \, dz$$

$$= \mathbb{E}\left[\frac{p(Z \mid x)}{p(Z)} \xi_h(Z)\right].$$

47 Thus, we must take

$$\Phi(x,z) = \frac{p(z \mid x)}{p(z)} = \frac{p(x \mid z)}{p(x)} = \frac{p(x,z)}{p(x)p(z)}.$$

With this choice, setting $u_h(x) = \Phi(x, Z)\xi_h(Z)$ we clearly have $\mathbb{E}[u_h(x)] = \nabla \mathcal{R}(h)(x)$.

Must discuss why $u_h \in L^2(X)$.

Should we discuss this

further?

49 4 Algorithm

Having an unbiased estimator of the gradient, we can construct an SGD algorithm for estimating h^* .

Algorithm 1: SGD-NPIV

input: Datasets $\mathcal{D}_{r_0} = \{(y_i, z_i)\} \overset{\text{iid}}{\sim} \mathbb{P}_{YZ}, \mathcal{D}_{\Phi} = \{(x_i, z_i)\} \overset{\text{iid}}{\sim} \mathbb{P}_{XZ}, \mathcal{D}_{\Phi} =$

output: $\left\{\widehat{h}(x_k)\right\}_{k=1}^K$

Compute $\{\widehat{r_0}(\boldsymbol{z}_m; \mathcal{D}_{r_0})\}_{m=1}^M$;

1 Compute $\widehat{\Phi}(m{x}, m{z}; \mathcal{D}_{\Phi})$;

for $1 \le m \le M$ do

 $\begin{aligned} & \text{Compute } \widehat{\mathcal{T}[\widehat{h}_{m-1}]}(\boldsymbol{z}_m; \mathcal{D}_{\mathcal{T}}) \;; \\ & \text{Set } u_m(\boldsymbol{x}_k) = \widehat{\Phi}(\boldsymbol{x}_k, \boldsymbol{z}_m) \partial_2 \ell \left(\widehat{r_0}(\boldsymbol{z}_m, \mathcal{D}_{r_0}), \widehat{\mathcal{T}[\widehat{h}_{m-1}]}(\boldsymbol{z}_m; \mathcal{D}_{\mathcal{T}}) \right) \quad \text{ for } 1 \leq k \leq K \;; \\ & \text{Set } \widehat{h}_m(\boldsymbol{x}_k) = \widehat{h}_{m-1}(\boldsymbol{x}_k) - \alpha_m u_m(\boldsymbol{x}_k) \quad \text{ for } 1 \leq k \leq K \;; \end{aligned}$

end

Set $\widehat{h} = \frac{1}{M} \sum_{m=1}^{M} \widehat{h}_m$;

52 **Proof of convergence**

- 53 The first problem is proving our sequence of estimates is, in fact, contained in $L^2(X)$. This amounts
- to proving $u_m \in L^2(X)$ for every m. It's not even immediate why $u_h(x) = \Phi(x,Z)\xi_h(Z)$ (the
- to proving $u_m \in L^{*}(X)$ for every m. It's not even infinediate why $u_h(x) = \Psi(x, Z)\zeta_h(Z)$ (in unbiased gradient when we know r_0 , Φ and \mathcal{T}) belongs to $L^2(X)$.

Discuss everything we don't know and must estimate.

Comment on exactly what is needed to estimate each unknown (samples from which r.v.'s).

Discuss necessity of discretizing X.

Do this.

²It is if ℓ is quadratic.

After doing this, the first steps in the proof are the same as in the previous paper. We show that \mathcal{R} is convex in \mathcal{F} and then simple algebraic manipulation allows us to write

$$\sum_{n=1}^{M} \left[\mathcal{R}(\hat{h}_{m-1}) - \mathcal{R}(h^{\star}) \right] \leq \sum_{m=1}^{M} \frac{1}{2\alpha_{m}} \left(\left\| \hat{h}_{m-1} - h^{\star} \right\|_{L^{2}(X)}^{2} - \left\| \hat{h}_{m} - h^{\star} \right\|_{L^{2}(X)}^{2} \right) + \sum_{m=1}^{M} \frac{\alpha_{m}}{2} \left\| u_{m} \right\|_{L^{2}(X)}^{2} - \left\| \hat{h}_{m} - h^{\star} \right\|_{L^{2}(X)}^{2} - \sum_{m=1}^{M} \left\langle u_{m} - \nabla \mathcal{R}(\hat{h}_{m-1}), \hat{h}_{m-1} - h^{\star} \right\rangle_{L^{2}(X)}$$

We then treat each term separately:

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• The first term is bounded using the assumption that $\operatorname{diam} \mathcal{F} = D < \infty$.

Assumption

- The bound on the second term depends on bounding $\mathbb{E}\left[\|u_m\|_{L^2(X)}^2\right]$ by a constant independent of m.
- The third term must vanish because of the unbiasedness of u_m , but we don't know that our u_m is unbiased, and it may very well not be.
- Second term Define \mathcal{D} to be the set of all observed data, that is, all of the variables in $\mathcal{D}_{\Phi}, \mathcal{D}_{r_0}, \mathcal{D}_{\mathcal{T}}$ and $\{z_i\}_{i=1}^M$. Let's evaluate $\mathbb{E}\left[\|u_m\|_{L^2(X)}^2\right]$:

$$\mathbb{E}\left[\left\|u_m\right\|_{L^2(X)}^2\right] = \mathbb{E}_{\mathcal{D}}\left[\mathbb{E}_X\left[\widehat{\Phi}(X,\boldsymbol{z}_m;\mathcal{D}_{\Phi})^2\partial_2\left(\widehat{r_0}(\boldsymbol{z}_m;\mathcal{D}_{r_0}),\widehat{\mathcal{T}[\widehat{h}_{m-1}]}(\boldsymbol{z}_m;\mathcal{D}_{\mathcal{T}})\right)^2\right]\right],$$

where the second expectation is with respect to a copy of X which is independent of \mathcal{D} . Continuing:

$$\begin{split} \mathbb{E}_{\mathcal{D}}\left[\mathbb{E}_{X}\left[\widehat{\Phi}(X,\boldsymbol{z}_{m};\mathcal{D}_{\Phi})^{2}\partial_{2}\left(\widehat{r_{0}}(\boldsymbol{z}_{m};\mathcal{D}_{r_{0}}),\widehat{\mathcal{T}[\widehat{h}_{m-1}]}(\boldsymbol{z}_{m};\mathcal{D}_{\mathcal{T}})\right)^{2}\right]\right] \\ &= \mathbb{E}_{\mathcal{D}}\left[\partial_{2}\left(\widehat{r_{0}}(\boldsymbol{z}_{m};\mathcal{D}_{r_{0}}),\widehat{\mathcal{T}[\widehat{h}_{m-1}]}(\boldsymbol{z}_{m};\mathcal{D}_{\mathcal{T}})\right)^{2}\mathbb{E}_{X}\left[\widehat{\Phi}(X,\boldsymbol{z}_{m};\mathcal{D}_{\Phi})^{2}\right]\right]. \end{split}$$

If ℓ is quadratic, we have

$$\begin{split} \mathbb{E}_{\mathcal{D}} \left[\partial_2 \left(\widehat{r_0}(\boldsymbol{z}_m; \mathcal{D}_{r_0}), \widehat{\mathcal{T}[\widehat{h}_{m-1}]}(\boldsymbol{z}_m; \mathcal{D}_{\mathcal{T}}) \right)^2 \mathbb{E}_X \left[\widehat{\Phi}(X, \boldsymbol{z}_m; \mathcal{D}_{\Phi})^2 \right] \right] \\ = \mathbb{E}_{\mathcal{D}} \left[\left(\widehat{T[\widehat{h}_{m-1}]}(\boldsymbol{z}_m; \mathcal{D}_{\mathcal{T}}) - \widehat{r_0}(\boldsymbol{z}_m; \mathcal{D}_{r_0}) \right)^2 \mathbb{E}_X \left[\widehat{\Phi}(X, \boldsymbol{z}_m; \mathcal{D}_{\Phi})^2 \right] \right]. \end{split}$$