

# Statistical Tests

Tuesday 29<sup>th</sup> November, 2016

1 Introduction

2 Tests

3 Conclusion

# Statistical Tests

- ▶ up to this point
- ▶ basic procedure
- ▶ in practice - specific tests used

# One and Two Tailed Tests

## Experiment I

- ▶ test hypothesis - crime rates are elevated when Internet services are down
- ▶ what are the appropriate alternative and null hypotheses ?

# One and Two Tailed Tests

## Experiment I

- ▶ test hypothesis - crime rates are elevated when Internet services are down
- ▶ what are the appropriate alternative and null hypotheses ?
- ▶  $C_r$  crime rate
- ▶  $C_r^{no-net}$  crime rate when Internet is down

$$H_0 : C_r = C_r^{no-net}$$

$$H_1 : C_r < C_r^{no-net}$$

# One and Two Tailed Tests

## Experiment II

- ▶ test hypothesis - crime rates are changed when Internet services are down
- ▶ what are the appropriate alternative and null hypotheses ?

# One and Two Tailed Tests

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- ▶ what are the appropriate alternative and null hypotheses ?
- ▶  $C_r$  crime rate
- ▶  $C_r^{no-net}$  crime rate when Internet is down

$$H_0 : C_r = C_r^{no-net}$$

$$H_1 : C_r \neq C_r^{no-net}$$

# One and Two Tailed Tests

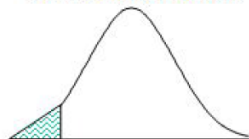
## One tailed test

$$H_0 : \theta \leq \theta_0$$

$$H_1 : \theta > \theta_0$$



Positive one-tailed test

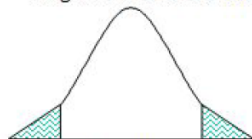


Negative one-tailed test

## Two tailed test

$$H_0 : \theta = \theta_0$$

$$H_1 : \theta \neq \theta_0$$



Two-tailed test



# The Two Tailed Coin inspector

test if a coin is fair with significance  $\alpha = 0.1$ :

- ▶ toss coin 10 times - 8 heads
- ▶ do you reject?
- ▶ what is the P-Value?

# The Two Tailed Coin inspector

if the test is two tailed

$H$  = number of heads

$$Pval = 2 \cdot \min \{P(H \geq 8), P(H \leq 8)\}$$

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$$\begin{aligned} Pval &= 2 \cdot \min \{ P(H \geq 8), P(H \leq 8) \} \\ &= 2 \cdot P(H \geq 8) \\ &= 2 \cdot \left( \binom{10}{2} + \binom{10}{1} + 1 \right) \cdot 0.5^{10} \end{aligned}$$

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# Rejection Region

## Automation

two tailed coin inspector

▶  $H = 8$  -  $Pval = 0.109$  - **not rejected**

▶  $H = 9$  -  $Pval = 0.021$  - **rejected**

instead of computing P-Value for each coin we can safely reject...

# Rejection Region

## Automation

two tailed coin inspector

▶  $H = 8 - Pval = 0.109$  - **not rejected**

▶  $H = 9 - Pval = 0.021$  - **rejected**

instead of computing P-Value for each coin we can safely reject...

$$\begin{aligned} R &= \{X : or (H_X \geq 9, H_X \leq 1)\} \\ &= \{X : |5 - H_X| \geq 4\} \end{aligned}$$

# Rejection Region

$$R = \{X : T(X) > c\}$$

- ▶  $T$  - test statistic
- ▶  $c$  - critical value
- ▶  $P(T(X) > c) = ?$



# Rejection Region

$$R = \{X : T(X) > c\}$$

- ▶  $T$  - test statistic
- ▶  $c$  - critical value
- ▶  $P(T(X) > c) = \alpha$

# One Tailed Test Example

- ▶  $X = X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2)$
- ▶  $\sigma$  is known
- ▶  $H_0 : \mu \leq 0$
- ▶  $H_1 : \mu > 0$

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- ▶  $T(X) = \frac{1}{n} \sum_{i=1}^n X_i$

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- ▶  $X \sim \mathcal{N}(\mu_X, \sigma_X^2)$ ,  $Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$
- ▶  $X$  and  $Y$  are independent
- ▶  $\alpha \in \mathbb{R}$

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$$\begin{aligned} X + Y &\sim \mathcal{N}(\mu_X + \mu_Y, \sigma_X^2 + \sigma_Y^2) \\ \alpha \cdot X &\sim \mathcal{N}(\alpha\mu_X, \alpha^2\sigma_X^2) \end{aligned}$$

# One Tailed Test Example

$$T(X) \sim \mathcal{N}\left(0, \frac{\sigma^2}{n}\right)$$

why?

$$T(X) = \frac{1}{n} [X_1 + X_2 + \dots + X_n]$$

$$T(X) \sim \mathcal{N}\left(0, \frac{1}{n^2} n \sigma^2\right)$$

# One Tailed Test Example

- ▶  $Z$  is standard normal -  $Z \sim \mathcal{N}(0, 1)$
- ▶  $\Phi$  the CDF of  $Z$

$$P(T > c) = P\left(\frac{T - \mu}{\sigma/\sqrt{n}} > \frac{c - \mu}{\sigma/\sqrt{n}}\right)$$



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$$\begin{aligned} P(T > c) &= P\left(\frac{T - \mu}{\sigma/\sqrt{n}} > \frac{c - \mu}{\sigma/\sqrt{n}}\right) \\ &= P\left(Z > \frac{\sqrt{n}(c - \mu)}{\sigma}\right) \end{aligned}$$

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# One Tailed Test Example

$$P(T > c) = 1 - \Phi\left(\frac{\sqrt{n}(c - \mu)}{\sigma}\right)$$

- ▶ test is most strict for  $\mu = 0$
- ▶ we require  $P(T > c) = \alpha$

why

# One Tailed Test Example

$$P(T > c) = 1 - \Phi\left(\frac{\sqrt{n}c}{\sigma}\right)$$

$$\alpha = 1 - \Phi\left(\frac{\sqrt{n}c}{\sigma}\right)$$

$$\frac{\sqrt{n}c}{\sigma} = \Phi^{-1}(1 - \alpha)$$

$$c = \frac{\sigma \Phi^{-1}(1 - \alpha)}{\sqrt{n}}$$

# One Tailed Test Example

$$P(T > c) = 1 - \Phi\left(\frac{\sqrt{n}c}{\sigma}\right)$$

$$\alpha = 1 - \Phi\left(\frac{\sqrt{n}c}{\sigma}\right)$$

$$\frac{\sqrt{n}c}{\sigma} = \Phi^{-1}(1 - \alpha)$$

$$c = \frac{\sigma Z_{\alpha}}{\sqrt{n}}$$

# Two Tailed Test Example

- ▶  $X = X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2)$
- ▶  $\sigma$  is known
- ▶  $H_0 : \mu = \mu_0$
- ▶  $H_1 : \mu \neq \mu_0$

# Two Tailed Test Example

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- ▶  $H_0 : \mu = \mu_0$
- ▶  $H_1 : \mu \neq \mu_0$
- ▶ **Test:**  $R = \{X : T(X) > c\}$
- ▶  $T(X) = \frac{1}{n} \sum_{i=1}^n X_i \dots$

# Two Tailed Test Example

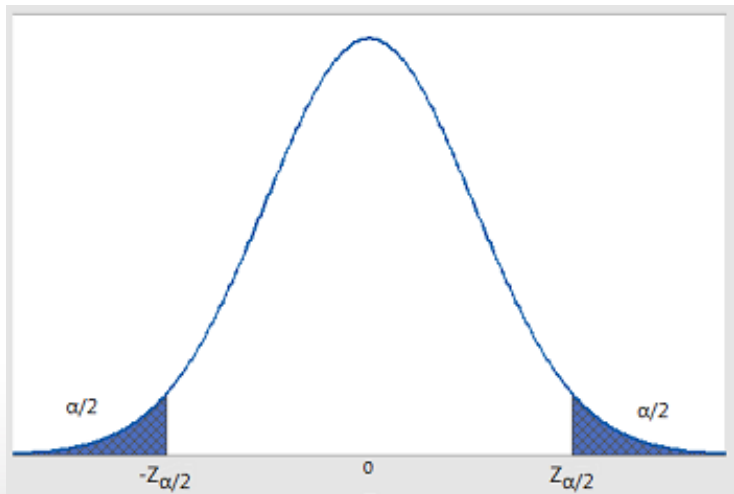
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# Two Tailed Test Example

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- ▶  $\sigma$  is known
- ▶  $H_0 : \mu = \mu_0$
- ▶  $H_1 : \mu \neq \mu_0$
- ▶ **Test:**  $R = \{X : |T(X)| > c\}$
- ▶  $T(X) = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}, T \sim \mathcal{N}(0, 1)$
- ▶  $c = Z_{\alpha/2}$

# Two Tailed Test Example



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## z-test

- ▶  $X = X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2)$
- ▶  $\sigma$  is known
- ▶  $H_0 : \mu = \mu_0$
- ▶  $H_1 : \mu \neq \mu_0$
- ▶ **Test:**  $R = \{X : |T(X)| > c\}$
- ▶  $T(X) = \frac{\bar{X} - \mu_0}{\sigma / \sqrt{n}}$
- ▶  $T \sim \mathcal{N}(0, 1)$

## z-test

### **Z-Test**

a test where the test statistic  $T \sim \mathcal{N}(0, 1)$

what happens when  $\sigma$  is unknown?

## t-test

- ▶  $X = X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2)$
- ▶  $\sigma$  is unknown
- ▶  $H_0 : \mu = \mu_0$
- ▶  $H_1 : \mu \neq \mu_0$
- ▶  $T(X) = \frac{\bar{X} - \mu_0}{\hat{\sigma} / \sqrt{n}}$

## t-test

- ▶  $X = X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2)$
- ▶  $\sigma$  is unknown
- ▶  $H_0 : \mu = \mu_0$
- ▶  $H_1 : \mu \neq \mu_0$
- ▶  $T(X) = \frac{\bar{X} - \mu_0}{\hat{\sigma} / \sqrt{n}}$
- ▶  $\sigma$  is estimated - sample standard deviation
- ▶ for large  $n$ , the distribution of  $T$  under  $H_0$  tends to  $\mathcal{N}(0, 1)$

## t-test

- ▶  $X = X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2)$
- ▶  $\sigma$  is unknown
- ▶  $H_0 : \mu = \mu_0$
- ▶  $H_1 : \mu \neq \mu_0$
- ▶  $T(X) = \frac{\bar{X} - \mu_0}{\hat{\sigma} / \sqrt{n}}$
- ▶  $\sigma$  is estimated - sample standard deviation
- ▶ exact distribution of  $T$  under  $H_0$  is Student's t-distribution with  $n - 1$  degrees of freedom

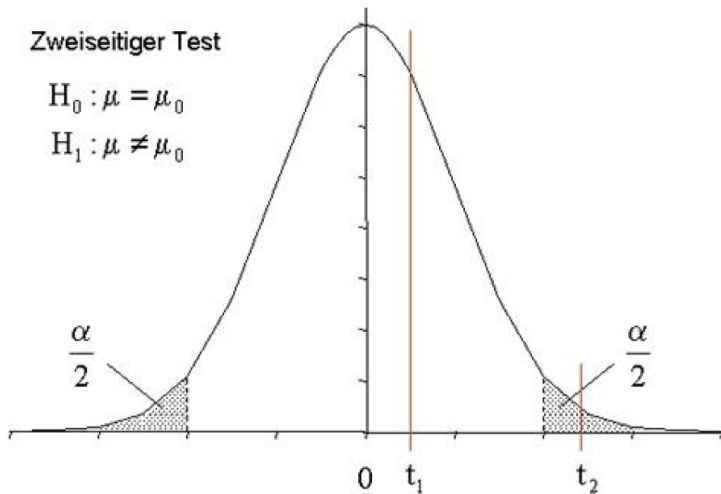


# Student's t-distribution

Zweiseitiger Test

$$H_0 : \mu = \mu_0$$

$$H_1 : \mu \neq \mu_0$$



# Likelihood Ratio Test

- ▶  $X = X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2)$
- ▶  $\sigma$  is known
- ▶  $H_0 : \mu = \mu_0$
- ▶  $H_1 : \mu \neq \mu_0$

# Likelihood Ratio Test

- ▶  $X = X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2)$
- ▶  $\sigma$  is known
- ▶  $H_0 : \mu = \mu_0$
- ▶  $H_1 : \mu \neq \mu_0$
- ▶ we can estimate  $\mu_{mle}$
- ▶ compute the likelihood ratio

rationale

$$\frac{\mathcal{L}(\mu_{mle})}{\mathcal{L}(\mu_0)}$$

# Likelihood Ratio Test

- ▶ test statistic

$$\lambda = 2 \log \frac{\mathcal{L}(\theta_{mle})}{\mathcal{L}(\theta_0)}$$

# Likelihood Ratio Test

in our case

$$\blacktriangleright \mu_{mle} = \sum_{i=1}^n X_i / n$$

$$\mathcal{L}(\mu) = \left( \frac{1}{\sqrt{2\pi\sigma^2}} \right)^n e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2}$$

$$\lambda = 2 \log \frac{e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu_{mle})^2}}{e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu_0)^2}}$$

# Likelihood Ratio Test

$$\begin{aligned}\lambda &= 2 \log \frac{e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu_{mle})^2}}{e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu_0)^2}} \\ &= 2 \log e^{\frac{1}{2\sigma^2} \left[ \sum_{i=1}^n (X_i - \mu_0)^2 - \sum_{i=1}^n (X_i - \mu_{mle})^2 \right]}\end{aligned}$$

$$\begin{aligned}\left[ \frac{(a-b)^2}{a^2 - 2ab + b^2} \right] &= \frac{1}{\sigma^2} \sum_{i=1}^n (\mu_0^2 - \mu_{mle}^2 - 2X_i\mu_0 + 2X_i\mu_{mle}) \\ &= \frac{1}{\sigma^2} \left( n(\mu_0^2 - \mu_{mle}^2) - 2(\mu_0 - \mu_{mle}) \sum_{i=1}^n X_i \right)\end{aligned}$$

# Likelihood Ratio Test

$$\lambda = \frac{1}{\sigma^2} \left( n (\mu_0^2 - \mu_{mle}^2) - 2 (\mu_0 - \mu_{mle}) \frac{n}{n} \sum_{i=1}^n X_i \right)$$

$$\begin{aligned} \left[ \frac{(a^2 - b^2)}{(a+b)(a-b)} \right] &= \frac{1}{\sigma^2} \left( \frac{n (\mu_0 - \mu_{mle}) (\mu_0 + \mu_{mle})}{2n (\mu_0 - \mu_{mle}) \mu_{mle}} \right) \\ &= \frac{n (\mu_0 - \mu_{mle})^2}{\sigma^2} \end{aligned}$$

# Likelihood Ratio Test

$$\lambda = \frac{n (\mu_0 - \mu_{mle})^2}{\sigma^2}$$

► under  $H_0$



# Likelihood Ratio Test

$$\lambda = \frac{n (\mu_0 - \mu_{mle})^2}{\sigma^2}$$

► under  $H_0$

$$\mu_{mle} \sim \mathcal{N}(\mu_0, \sigma^2/n)$$

# Likelihood Ratio Test

$$\lambda = \frac{n (\mu_0 - \mu_{mle})^2}{\sigma^2}$$

► under  $H_0$

$$\mu_{mle} \sim \mathcal{N}(\mu_0, \sigma^2/n)$$

$$\sqrt{\lambda} = \frac{\sqrt{n}}{\sigma} (\mu_0 - \mu_{mle}) \sim \mathcal{N}(0, 1)$$

# Likelihood Ratio Test

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$$\boxed{\lambda \sim \chi_1^2}$$

# Likelihood Ratio Test

$$\lambda = \frac{n (\mu_0 - \mu_{mle})^2}{\sigma^2}$$

- ▶ under  $H_0$

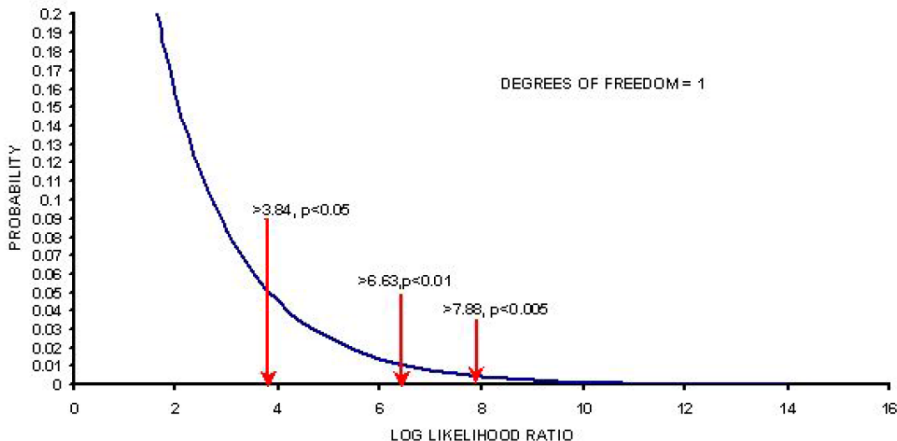
$$\mu_{mle} \sim \mathcal{N}(\mu_0, \sigma^2/n)$$

$$\boxed{\lambda \sim \chi_1^2}$$

- ▶ The Chi-Distribution with 1 degree of freedom
- ▶ asymptotically true for a wide range of cases

# Chi Squared Distribution

CHI-SQUARED DISTRIBUTION



# Testing Independence

$X$  and  $Y$  two random variables

- ▶ let's test independence

# Testing Independence

$X$  and  $Y$  two random variables

- ▶  $H_0$   $X$  and  $Y$  are independent
- ▶  $H_1$   $X$  and  $Y$  are **not** independent

# Testing Independence

under  $H_0$  what can we say about  $P(X, Y)$ ?



# Testing Independence

under  $H_0$  what can we say about  $P(X, Y)$ ?

$$P(X, Y) = P(X) P(Y)$$

# Testing Independence

- ▶  $X$  and  $Y$  are binary
- ▶  $(X_1, Y_1), \dots, (X_n, Y_n)$  samples of  $(X, Y)$
- ▶ number of observations for each value

	$Y = 0$	$Y = 1$	Total
$X = 0$	$z_{00}$	$z_{01}$	$z_{0.}$
$X = 1$	$z_{10}$	$z_{11}$	$z_{1.}$
Total	$z_{.0}$	$z_{.1}$	$n$

# Testing Independence

- ▶  $X$  and  $Y$  are binary
- ▶  $(X_1, Y_1), \dots, (X_n, Y_n)$  samples of  $(X, Y)$
- ▶ probabilities

	$Y = 0$	$Y = 1$	Total
$X = 0$	$p_{00}$	$p_{01}$	$p_{0\cdot}$
$X = 1$	$p_{10}$	$p_{11}$	$p_{1\cdot}$
Total	$p_{\cdot 0}$	$p_{\cdot 1}$	1

# Testing Independence

## estimation

$$\hat{p}_{ij} = \frac{z_{ij}}{n}$$

$$\hat{p}_{i\cdot} = \frac{z_{i\cdot}}{n}$$

$$\hat{p}_{\cdot j} = \frac{z_{\cdot j}}{n}$$

# Testing Independence

## example

we sample  $(X, Y)$   $n = 3$  times

$(0, 1), (0, 0), (1, 0)$

number of observations for each value

	$Y = 0$	$Y = 1$	Total
$X = 0$	1	1	2
$X = 1$	1	0	1
Total	2	1	3

# Testing Independence

## example

we sample  $(X, Y)$   $n = 3$  times

$(0, 1), (0, 0), (1, 0)$

our estimation of the probabilities

	$Y = 0$	$Y = 1$	Total
$X = 0$	$1/3$	$1/3$	$2/3$
$X = 1$	$1/3$	$0$	$1/3$
Total	$2/3$	$1/3$	$1$

# Testing Independence

## example

we sample  $(X, Y)$   $n = 3$  times

$$(0, 1), (0, 0), (1, 0)$$

---

$$P(X = 0, Y = 1) = ?$$

► under  $H_0$

$$\begin{aligned} P(X = 0, Y = 1) &= P(X = 0) P(Y = 1) \\ &= 2/3 \cdot 1/3 \\ &= 2/9 \end{aligned}$$

► under  $H_1$

$$P(X = 0, Y = 1) = 1/3$$

# Testing Independence

$$\begin{aligned}\lambda &= 2 \log \frac{\mathcal{L}(\hat{p})}{\mathcal{L}(\hat{p}_0)} \\&= 2 \sum_{i=1}^n \log \frac{\hat{p}_{X_i Y_i}}{\hat{p}_{X_i \cdot} \hat{p}_{\cdot Y_i}} \\&= 2 \sum_{i=0}^1 \sum_{j=0}^1 z_{ij} \log \frac{\hat{p}_{ij}}{\hat{p}_{i \cdot} \hat{p}_{\cdot j}} \\&= 2 \sum_{i=0}^1 \sum_{j=0}^1 z_{ij} \log \frac{z_{ij} n}{z_{i \cdot} z_{\cdot j}}\end{aligned}$$



# Testing Independence

$$\lambda = 2 \sum_{i=0}^1 \sum_{j=0}^1 z_{ij} \log \frac{z_{ij}n}{z_{i\cdot}z_{\cdot j}}$$

- ▶ under  $H_0$ ,  $\lambda \sim \chi_1^2$

# Testing Independence

In the general case

- ▶  $X$  -  $r$  values observed
- ▶  $Y$  -  $c$  values observed

$$\lambda = 2 \sum_{i=0}^1 \sum_{j=0}^1 z_{ij} \log \frac{z_{ij}n}{z_{i.}z_{.j}}$$

- ▶ under  $H_0$ ,  $\lambda \sim \chi_k^2$
- ▶  $k = (r - 1)(c - 1)$

# Testing Independence

## Pearson's Test

- ▶  $U = \sum_{i=1}^r \sum_{j=1}^c \frac{(z_{ij} - z_{i.}z_{.j}/n)^2}{z_{i.}z_{.j}/n}$
- ▶ under  $H_0$ ,  $U \sim \chi^2_{(r-1)(c-1)}$

# Testing Correlation

- ▶  $X, Y$  two random variables
- ▶  $(X_1, Y_1), \dots, (X_n, Y_n)$  samples of  $(X, Y)$

how should we test?

# Testing Correlation

- ▶  $X, Y$  two random variables
- ▶  $(X_1, Y_1), \dots, (X_n, Y_n)$  samples of  $(X, Y)$
- ▶  $H_0 - \rho_{XY} = 0$
- ▶  $H_1 - \rho_{XY} \neq 0$
- ▶  $r$  the sample correlation

# Testing Correlation

- ▶ no knowledge of  $P(r)$  under  $H_0$
- ▶ maybe we need a different approach

# Testing Correlation

- ▶ let's generate an empiric  $P(r)$
- ▶  $p$  a random permutation of  $1, \dots, n$

$$p : \{1, \dots, n\} \mapsto \{1, \dots, n\}$$

- ▶ example

$$p(\{1, 2, 3\}) = \{3, 1, 2\}$$

- ▶ a "new" (permuted) set of samples

$$(X_1, Y_{p(1)}) , \dots , (X_1, Y_{p(n)})$$

# Testing Correlation

- ▶ generate  $p_1, \dots, p_m$
- ▶  $\forall j = 1, \dots, m$  compute  $r_j$  from

$$(X_1, Y_{p_j(1)}), \dots, (X_1, Y_{p_j(n)})$$

- ▶ the P-Value for this empiric distribution

$$P_{val} = \frac{1}{m} \sum_{j=1}^m \mathbb{1} \{ |r_j| > |r| \}$$



# Important Terms

- ▶ **power/sensitivity/recall** - probability of correctly rejecting  $H_0$  when  $H_1$  is true

$$P(\text{reject } H_0 \mid H_1)$$

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- ▶ **specificity** - probability of correctly not rejecting  $H_0$  when it is true

$$P(\text{not reject } H_0 \mid H_0)$$

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- ▶ **precision** - probability of a detection being true

$$P(H_1 \mid \text{reject } H_0)$$

1 Introduction

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3 Conclusion

# Conclusion

- ▶ from P-Values to test statistics
- ▶ one and two tailed tests
- ▶ the statistical test zoo
- ▶ tests for specific cases of interest - independence, correlation

figures

- ▶ two tailed test - [onlinecourses.science.psu.edu/stat500/node/44](https://onlinecourses.science.psu.edu/stat500/node/44)