

# Beyond (?) the pinball loss: Quantile Methods for Calibrated Uncertainty Quantification

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## Definitions

## Methods

- A fancy loss

- A Proper Scoring Rule

- KDE

## Experiments

## Results

## Concluding Thoughts

## References

- ▶ Supervised learning
  - ⇒ Let  $\mathbf{X}, \mathbf{Y} \sim \mathbb{F}_{\mathcal{X}, \mathcal{Y}}$  denote random variables over  $\mathcal{X} \times \mathcal{Y}$ ,  $\mathcal{Y}$  an interval in  $\mathbb{R}$  (i.e regression setting)
- ▶ We assume there exists a true conditional distribution  $\mathbb{F}_{\mathbf{Y}|\mathbf{x}}$  over  $\mathcal{Y}$
- ▶  $Q_p(x)$  denotes the true p-th conditional quantile of this distribution i.e.  $\mathbb{F}_{\mathbf{Y}|\mathbf{x}}(Q_p(x)) = p$
- ▶ Conditional quantile estimator  $\hat{Q}_p : \mathcal{X} \times (0, 1) \rightarrow \mathcal{Y}$

- The  $p$ -th quantile minimizes the *pinball loss*  $\ell_p : \mathcal{Y} \times \mathcal{Y} \rightarrow \mathbb{R}$ .  
Given a target  $Q_p(x) = y$  and a prediction  $\hat{Q}_p(x) = \hat{y}$ :

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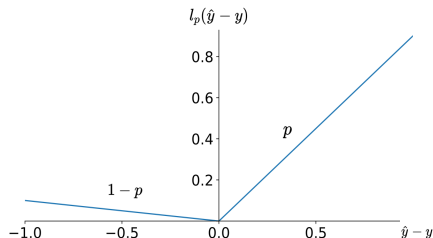
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$$\ell_p(y, \hat{y}) = (\hat{y} - y)(\mathbb{I}\{y \leq \hat{y}\} - p) = \begin{cases} (1 - p)(\hat{y} - y) & y < \hat{y} \\ -p(\hat{y} - y) & y \geq \hat{y} \end{cases}$$

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- ▶ Let  $R(\hat{y}) = \mathbb{E}_{y \sim \mathbb{F}_{Y|x}} \ell_p(y, \hat{y})$  denote the statistical risk
- ▶ Assuming  $R(\hat{y})$  is differentiable

$$\begin{aligned} \frac{\partial R(\hat{y})}{\partial \hat{y}} &= (1 - p)\mathbb{F}_{Y|x}(\hat{y}) - p(1 - \mathbb{F}_{Y|x}(\hat{y})) = \mathbb{F}_{Y|x}(\hat{y}) - p \\ \implies \left. \frac{\partial R(\hat{y})}{\partial \hat{y}} \right|_{\hat{y}=y} &= 0 \end{aligned}$$

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  - ▶ Proper scoring rules have tons of nice properties:
    - ▶ Form a non-negative convex cone
    - ▶ Admit an integral representation
    - ▶ Can define information measures and Bregman divergences under some conditions.
- ⇒ See [Buja et al., 2005] for a wonderful characterization of proper scoring rules for binary classification.

- ▶ Even Stronger Results:
- ▶ [Schervish et al., 2018] (Theorem 1): Any real valued quantile prediction loss  $g_p$  is a (strictly) proper scoring rule for quantile  $p$  iff there exists a (strictly) increasing function  $s$  such that:

$$g_p(y, \hat{y}) - g_p(y, y) = \begin{cases} p[s(y) - s(\hat{y})] & \text{if } y > \hat{y} \\ (1 - p)[s(\hat{y}) - s(y)] & \text{if } \hat{y} < y \end{cases}$$

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⇒ In fact all scoring rules of this form are also (strictly) proper scoring rules for probability prediction of binary variables [Buja et al., 2005].

- ▶ All proper scoring rules minimise both **Calibration and Sharpness** because, by definition they are minimised under the true distribution.
  - ⇒ However, this **balance/trade-off** (between penalising Calibration and Sharpness) is fixed, which according to [Chung et al., 2020] becomes relevant when minimising it empirically.

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- ▶ We can re-write the pinball loss as:

$$\ell_p(y, \hat{y}) = p\hat{y} + (y - \hat{y})\mathbb{I}_{y \leq \hat{y}} - py$$

- ▶  $p\hat{y}$  penalizes larger quantile predictions, i.e. sharpness
- ▶  $(y - \hat{y})\mathbb{I}_{y \leq \hat{y}}$  penalizes calibration
- ▶  $py$  does not depend on predictions.

[Chung et al., 2020] propose a “tunable” loss function.



- ▶ Same Notions of calibration as in probability forecasts.
- ▶  $\mathcal{F}(\mathcal{Y})$  that maps an input  $x \in \mathcal{X}$  to a continuous CDF  $h(y)$  over  $\mathcal{Y}$ .
- ▶ *Perfect Probability Forecast* outputs the true conditional CDF  $h^*(x) = \mathbb{F}_{Y|x}$
- ▶ Conditional Quantile regression “at all quantiles” (for all  $p$ ) is equivalent to Inverse CDF estimation

⇒ We will refer to the family of quantile estimates as

$$\hat{Q} : \mathcal{X} \times (0, 1) \rightarrow \mathcal{Y} = \{\hat{Q}_p(x), p \in [0, 1]\}$$

⇒ This is what all the experiments in [Chung et al., 2020] actually estimate/model

- ▶ What we defined in class:

- ▶ Classifier  $\hat{y} : \mathcal{X} \rightarrow \mathcal{Y}$
- ▶ Confidence predictor  $\hat{p} : \mathcal{X} \rightarrow [0, 1]$

$$P(\hat{y}(x) = y \mid \hat{p}(x) = c) = c$$

- ▶ Quantile Regression

for all  $p \in (0, 1)$ ,  $x \in \mathcal{X}$

$$\hat{Q}_p(x) = Q_p(x)$$

$$\Leftrightarrow \mathbb{F}_{\mathbf{Y}|x}(\hat{Q}_p(x)) = p$$

- ▶ This is equivalent to marginal coverage.

$\Rightarrow$  For  $\hat{Q}$  to be calibrated, this has to hold for all  $p$

$\Rightarrow$  “Individual Fairness” [Kearns et al., 2019]

- ▶ Impossibility results cited

- ▶ [Vovk, 2012]: (we saw it in class) Conditional Conformal Inference: at almost all nonatomic points of  $\mathbf{x}$ , the prediction interval has infinite expected length.
- ▶ [Vovk et al., 2005]: Probabilistic prediction (without assumptions on distributions) is impossible under Finite  $\mathcal{X}$  and  $\mathcal{Y}$ , and Finite training set without repetition of  $\mathbf{x}$ .
- ▶ [Zhao et al., 2020] Individual calibration for probability forecasters is impossible (and unverifiable) for finite datasets (without assumptions on distributions).

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$$p_{\text{obs}}(p) := \mathbb{E}_{\mathbf{Y}|x} \left( \hat{Q}_p(x) \right)$$

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- ▶ Consider groups, i.e. measurable subsets  $\mathcal{S}_i \subset \mathcal{X}$ ,  $i = 0, \dots, k$ .  
We want mean calibration to hold when conditioning on each group:

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- ▶ **Adversarial Group Calibration** “Average calibration within all subsets of the dataset with sufficiently many points” as a proxy for

Group Calibration for all  $\mathcal{S} \subset \mathcal{X}$  s.t.  $P_{x \sim \mathbb{F}_X}(x \in \mathcal{S}) > 0$

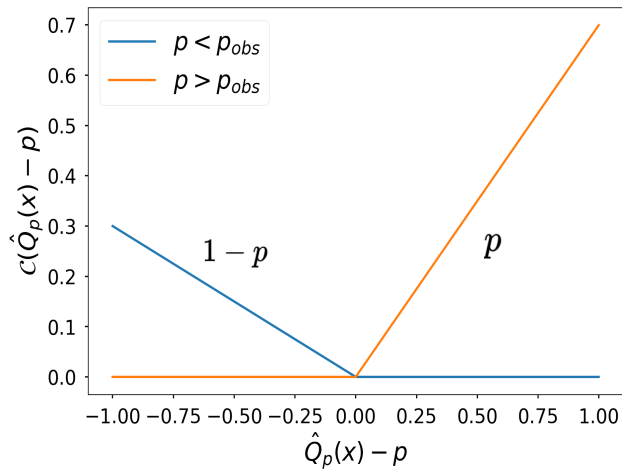
- ▶ From mean to adversarial Groups: Increasingly stringent (in  $x$ ) definitions of calibration, may be suitable for different applications.
- ▶ There are also PAC notions of calibration in QR [Zhao et al., 2020].
- ▶ Other definitions in other contexts. E.g.: [Kearns et al., 2019] average over different tasks.



- Calibration objective:

$$\begin{aligned} \mathcal{C}(\hat{Q}_p) = & \mathbb{I}\{p_{obs} < p\} * \mathbb{E}\left[Y - \hat{Q}_p \mid Y > \hat{Q}_p\right] * (1 - p_{obs}) \\ & + \mathbb{I}\{p_{obs} > p\} * \mathbb{E}\left[\hat{Q}_p - Y \mid \hat{Q}_p > Y\right] * p_{obs} \end{aligned}$$

# A Tunable Loss function



- ▶ Calibration objective:
- ▶ It is minimised by the true quantiles.
- ▶ It is not decomposable in individual samples.
- ▶ It is not a proper scoring function

- ▶ Sharpness objective
- ▶ Predict quantiles at  $p$  and  $1 - p$

$$\mathcal{P}(\hat{Q}_p, p) = \mathbb{E} \left[ \left| \hat{Q}_p - \hat{Q}_{1-p} \right| \right]$$

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- ▶ You should only penalise

$$\mathbb{E} \left[ \left| \hat{Q}_p - \hat{Q}_{1-p} \right| \right] \neq \mathbb{E} [|Q_p - Q_{1-p}|] = 1 - 2p$$

- Tunable loss:

$$\mathcal{L}(\hat{Q}_p, \hat{Q}_{1-p}) = (1 - \lambda) [\mathcal{C}(\hat{Q}_p) + \mathcal{C}(\hat{Q}_{1-p})] + \lambda \mathcal{P}(\hat{Q}_p, \hat{Q}_{1-p})$$

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- ▶  $\lambda$  is set by doing cross validation
- ▶ [Chung et al., 2020] train a model that outputs all quantiles by optimizing

$$\mathbb{E}_{p \sim \text{Unif}(0,1)} \mathcal{L}(\hat{Q}_p, \hat{Q}_{1-p})$$

- ▶ This loss not a proper scoring rule
- ▶ It is not decomposable in individual samples
- ▶ It is not minimised under the true quantiles.



- The interval (Winkler) Score [Winkler, 1972]:

$$S_{\alpha}(\hat{l}, \hat{u}; y) = (\hat{u} - \hat{l}) + \frac{2}{\alpha} \left[ (\hat{l} - y)\mathbb{I}\{y < \hat{l}\} + (y - \hat{u})\mathbb{I}\{y > \hat{u}\} \right]$$

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⇒ It is the scaled sum of pinball losses:

$$S_{\alpha}(\hat{l}, \hat{u}; y) = \frac{2}{\alpha} \left[ \ell_{1-\alpha}(y, \tilde{u}) + \ell_{\alpha}(y, \tilde{l}) \right]$$

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⇒ “bring to light a proper scoring rule that has largely been neglected for the purpose of learning quantiles. While some previous works utilize the interval score to evaluate interval predictions [some citations], to the best of our knowledge, no previous work has focused on simultaneously optimizing it and shown a thorough experimental evaluation as we provide”

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⇒ If training over all quantiles the  $\alpha$ -th pinball loss gets scaled by:

$$\frac{2}{\alpha} + \frac{2}{1 - \alpha}$$

- ▶ Conditional Density estimation
  - ▶ Assumes smoothness i.e.  $x_j \approx x_k$  then  $\mathbb{F}_{\mathbf{Y}|x_j} \approx \mathbb{F}_{\mathbf{Y}|x_k}$
- ▶ Under assumptions on the bandwidth the kernel density estimation Converges Uniformly to CDF [Stute, 1986]
- ▶ MAQR:
  - ▶ Utilize conditional density estimators to collect a dataset of quantile estimates
  - ▶ Fit a regressor on those quantiles (to get their inverses)

- ▶ Expected Calibration Error:

- ▶ Quantile predictor  $\hat{Q}_p$
- ▶ For N samples the empirical observed probability is:

$$\hat{p}_{obs}(p) = \frac{1}{N} \sum_{i=1}^N \mathbb{I} \left\{ y_i \leq \hat{Q}_p(x_i) \right\}$$

$$ECE(\hat{Q}_p) = |p_{obs} - p|$$

- ▶ Family of Quantile predictors  $\hat{Q} = \hat{Q}_p, p \in [0, 1]$
- ▶ [Chung et al., 2020] Average over m quantiles

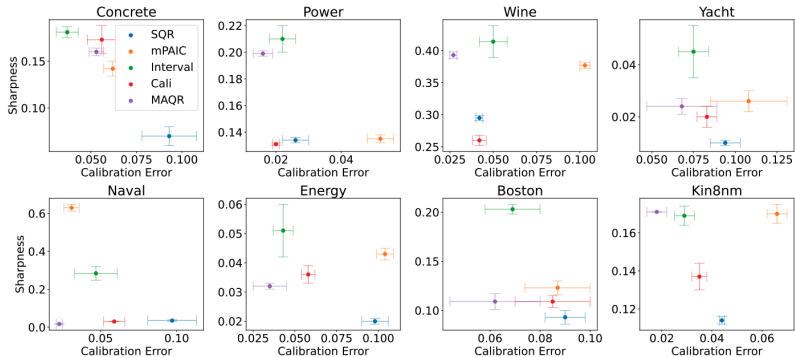
$$ECE(Q) = \frac{1}{m} \sum_{j=1}^m |\hat{p}_{obs}(p_j) - p_j|, \text{ where } p_j \sim \text{Unif}(0, 1)$$

- ▶ Models trained for all quantiles sampling  $p$  uniformly
  - ▶ Cali: Using their loss
  - ▶ SQR [Tagasovska and Lopez-Paz, 2019]: Pinball Loss
  - ▶ Interval Score
  - ▶ MAQR: KDE + Regressor
  - ▶ mPAIC [Zhao et al., 2020]: Randomised Quantile predictors trained on NLL+ECE

- ▶ UCI:
  - ▶ 8 regression datasets
  - ▶ Tiny and Low dimensional: dimension  $\leq 20$
- ▶ Nuclear fission: windowed time series. (iid assumption does not hold)
  - ▶ 16 regression tasks/outputs
  - ▶ Better but still low input dimension: 468

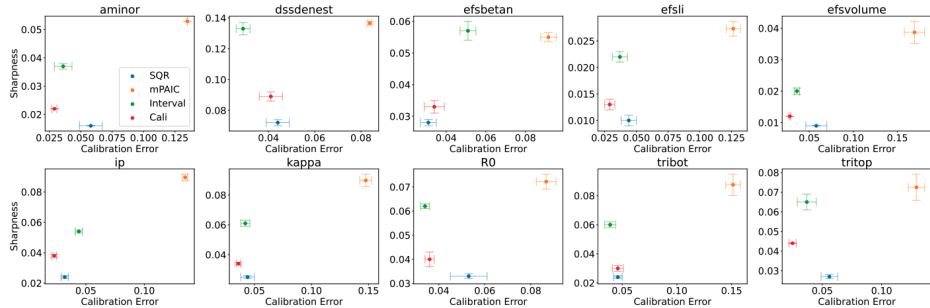


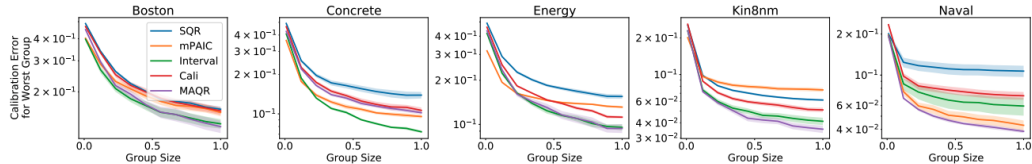
- ▶ UCI:
  - ▶ 2 layer Neural Network with 64 hidden neurons
- ▶ Nuclear fission
  - ▶ Deep learning: 3 hidden layers w/100 hidden neurons each



	<i>SQR</i>	<i>mPAIC</i>	<i>Interval</i>	<i>Cali</i>	<i>MAQR</i>
concrete	$2.038 \pm 0.225$	$1.157 \pm 0.069$	$0.943 \pm 0.053$	$1.465 \pm 0.086$	<b><math>0.672 \pm 0.118</math></b>
power	$0.834 \pm 0.022$	$0.917 \pm 0.021$	$0.620 \pm 0.010$	$0.699 \pm 0.019$	<b><math>0.592 \pm 0.009</math></b>
wine	$3.242 \pm 0.166$	$3.168 \pm 0.019$	$2.197 \pm 0.045$	$2.498 \pm 0.135$	<b><math>2.052 \pm 0.052</math></b>
yacht	$0.314 \pm 0.061$	$0.197 \pm 0.036$	$0.190 \pm 0.021$	$0.298 \pm 0.063$	<b><math>0.086 \pm 0.016</math></b>
naval	$0.097 \pm 0.011$	$3.112 \pm 0.053$	$0.620 \pm 0.114$	$1.560 \pm 0.268$	<b><math>0.044 \pm 0.001</math></b>
energy	$0.290 \pm 0.016$	$0.223 \pm 0.017$	$0.182 \pm 0.026$	$0.204 \pm 0.018$	<b><math>0.101 \pm 0.006</math></b>
boston	$1.833 \pm 0.299$	$1.395 \pm 0.176$	$1.010 \pm 0.118$	$1.449 \pm 0.259$	<b><math>0.864 \pm 0.287</math></b>
kin8nm	$1.241 \pm 0.041$	$1.347 \pm 0.031$	$0.776 \pm 0.017$	$1.121 \pm 0.072$	<b><math>0.691 \pm 0.015</math></b>





Figure 10: **UCI Interval Score** Full interval score results of UCI experiments from Section 4.1. Mean score across 5 trials is given, along with  $\pm 1$  standard error. The best mean has been bolded. *MAQR* tends to achieve the best interval score, which is surprising given that *Interval* utilizes the same model class to optimize the interval score directly.









- ▶ QR is a widely used Uncertainty Quantification method
- ▶ Can be used to obtain prediction regions.
- ▶ But unlike the conformal setting iid-ness is assumed.
- ▶ Estimating at all quantiles is just estimating the inverse conditional CDF.
- ▶ How does this relate to Takeuchi's optimal regions in the conformal setting?

- ▶ Not so happy about
  - ▶ Toy datasets and models are used for benchmarking
    - ⇒ Although the usual computer vision benchmarks are used in other papers. And some fairness and causal inference related datasets.
  - ▶ KDE for those problems seems to work better but won't work in high dimensional settings.
  - ▶ fitting all quantiles may not be traditional QR
  - ▶ Designing penalised losses that may work in practice but without much connection to theory

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